

Exercise 1

Solutions

David Puelz

9/12/2016

Exercise 1

David Puelz

Linear Regression

Consider $y = X\beta + \varepsilon$

$$y \in \mathbb{R}^N, \quad X \in \mathbb{R}^{N \times p}, \quad \beta \in \mathbb{R}^p$$

• Principle of weighted least squares:

$$\hat{\beta} = \underset{\beta \in \mathbb{R}^p}{\operatorname{argmin}} \sum_{i=1}^N \frac{w_i}{2} (y_i - x_i^T \beta)^2$$

$$A) (i) \sum_{i=1}^N \frac{w_i}{2} (y_i - x_i^T \beta)^2$$

$$= \frac{1}{2} \sum_{i=1}^N (y_i - x_i^T \beta)^T w_i (y_i - x_i^T \beta)$$

$$= \frac{1}{2} \sum_{i=1}^N (y_i w_i y_i - 2 y_i w_i x_i^T \beta + x_i^T \beta w_i x_i^T \beta)$$

$$= \frac{1}{2} (y^T W y - 2 y^T W X \beta + (X \beta)^T W X \beta)$$

$$= \frac{1}{2} (y - X \beta)^T W (y - X \beta)$$



(ii) Need to find minimum of

$$f(\beta) = y^T W y - 2 \beta^T X^T W y + \beta^T X^T W X \beta$$

$$\Rightarrow \frac{\partial f(\beta)}{\partial \beta} = -2 X^T W y + 2 X^T W X \beta \quad (\text{matrix derivative})$$

$$\text{FOC is: } \frac{\partial f(\beta)}{\partial \beta} = 0 = -2 X^T W y + 2 X^T W X \hat{\beta}$$

$\Rightarrow \hat{\beta}$ must satisfy:

$$0 = -2 X^T W y + 2 X^T W X \hat{\beta}$$

\Leftrightarrow

$$X^T W y = (X^T W X) \hat{\beta} \quad \blacksquare$$

(iii) let's look at the QR factorization.

Theorem: Given problem

$$\min_{\beta} \|y - X\beta\|_2, \text{ its solution set is}$$

equivalent to
residual set
of

$$R\beta = Q^T y \quad \text{where } Q \text{ is}$$

orthonormal, R is upper triangular, and both come from the QR factorization of X .

} one approach

A second approach would be to solve the set of normal equations: (FOL from least-squares)

$$(X^T W X) \beta = X^T W y \quad \text{where the unknown is } y.$$

Write: $C = X^T W X$

$$d = X^T W y$$

$LL^T = C$ as Cholesky factorization
(L upper triangular).

Algorithm: Solve $Lz = d$, forward substitution
for z since L is triangular.

Solve $L^T \beta = z$, backward substitution
for β

Class notes:

3 standard factorizations

- | | | |
|------------------------|---|---|
| 1) Cholesky (LL^T) | } | → fast, unstable (the faster) back. |
| 2) QR | | → fast, stable. |
| 3) SVD | | → matrix that is close to rank deficient. |

↓

Focus on QR:

$$W^{1/2}X = QR \quad (\text{where } W^{1/2}W^{1/2} = W)$$

$$\begin{array}{ccc} N \times N & N \times P & N \times P \\ & \swarrow & \searrow \\ & \text{Square} & \text{upper} \\ & P \times P & \Delta \\ & \text{orthonormal} & \\ & \text{columns} & \end{array} \quad \left. \vphantom{\begin{array}{ccc} N \times N & N \times P & N \times P \\ & \swarrow & \searrow \\ & \text{Square} & \text{upper} \\ & P \times P & \Delta \\ & \text{orthonormal} & \\ & \text{columns} & \end{array}} \right\} \text{reduced QR factorization.}$$

Looking back at normal equations:

$$\Rightarrow X^T W Y = X^T W X \beta$$

$$\Rightarrow X^T W^{1/2} W^{1/2} Y = X^T W^{1/2} W^{1/2} X \beta$$

$$\Rightarrow (QR)^T W^{1/2} Y = (QR)^T QR \beta$$

$$\Rightarrow \cancel{R^T} Q^T W^{1/2} Y = \cancel{R^T} Q^T QR \beta$$

$$Q^T W^{1/2} Y = R \beta \quad \leftarrow \text{now we have an upper triangular system!!!}$$

Generalized Linear Models

Let $y_i \sim \text{Bin}(m_i, w_i)$

with $w_i(\beta) = \frac{1}{1 + \exp\{-x_i^T \beta\}}$.

A) Likelihood is:

$$\ell(\beta) = -\log \left\{ \prod_{i=1}^N p(y_i | \beta) \right\}$$

where $p(y_i | \beta) = \binom{m_i}{y_i} w_i^{y_i} (1-w_i)^{m_i-y_i}$

$$\Rightarrow \ell(\beta) = -\log \left\{ \prod_{i=1}^N \binom{m_i}{y_i} w_i^{y_i} (1-w_i)^{m_i-y_i} \right\}$$

$$\propto -\sum_{i=1}^N \left[y_i \log w_i + (m_i - y_i) \log(1-w_i) \right]$$

Let's define

$$\ell_i(\beta) = y_i \log w_i + (m_i - y_i) \log(1-w_i)$$

so that

$$\ell(\beta) = -\sum_{i=1}^N \ell_i(\beta)$$

]

→ for a general function, $f(x)$

$$\frac{d}{dx} \log f(x) = \frac{f'(x)}{f(x)}$$

→ Similarly,

$$\nabla_{\beta} \log w_i(\beta) = \frac{\nabla_{\beta} w_i(\beta)}{w_i(\beta)}$$

Solving for gradient of $w_i(\beta)$:

$$\nabla_{\beta} w_i(\beta) = \frac{-\exp\{\bar{x}_i^T \beta\}}{(1 + \exp\{-x_i^T \beta\})^2} x_i = w_i(1 - w_i) x_i$$

$$\begin{aligned} \Rightarrow \frac{\nabla_{\beta} w_i(\beta)}{w_i(\beta)} &= \frac{-\exp\{-x_i^T \beta\}}{1 + \exp\{-x_i^T \beta\}} x_i \\ &= -(1 - w_i) x_i \end{aligned}$$

Now, looking back at the likelihood, we have

$$\begin{aligned} \nabla \ell(\beta) &= \nabla (y_i \log w_i + (m_i - y_i) \log(1 - w_i)) \\ &= -y_i(1 - w_i)x_i + (m_i - y_i)w_i x_i \\ &= -y_i x_i + \cancel{y_i w_i x_i} + m_i w_i x_i - \cancel{y_i w_i x_i} \\ &= -(y_i - m_i w_i) x_i \end{aligned}$$

where

$$\nabla_{\beta} \log(1 - w_i(\beta)) = w_i x_i$$

by similar
argument to
above.

Therefore,

$$\nabla \ell(\beta) = - \sum_{i=1}^N \nabla \ell_i(\beta)$$

$$= - \sum_{i=1}^N (y_i - m_i w_i) x_i$$

$$= - X^T S$$

X matrix, S is vector with
 N components $\{$
 i^{th} component
 $\{y_i - m_i w_i\}$.

c) Let $\beta_0 \in \mathbb{R}^P$, we call that we write

$\ell(\beta)$ as

$$\ell(\beta) = - \sum_{i=1}^N \ell_i(\beta)$$

with $\ell_i(\beta) = y_i \log w_i + (n_i - y_i) \log(1 - w_i)$

Taylor's theorem states (approximating $\ell(\beta)$ near β_0)

$$\ell(\beta) \approx \ell(\beta_0) + \underbrace{\nabla \ell(\beta_0)^T (\beta - \beta_0)}_{\text{1st order term}} + \underbrace{\frac{1}{2} (\beta - \beta_0)^T \nabla^2 \ell(\beta_0) (\beta - \beta_0)}_{\text{2nd order term}}$$

Let's solve for the Hessian, $\nabla^2 \ell(\beta_0)$.

Recall:

$$\begin{aligned} \nabla \ell(\beta) &= -X^T S, \quad S \text{ vector with } i\text{th component } y_i - n_i w_i \\ &= -X^T (y - Mw), \quad M \text{ is matrix (diagonal) with } M_{ii} = n_i \\ &= -X^T y + X^T Mw \end{aligned}$$

Now, we aim to calculate

$$\begin{aligned} \nabla^2 \ell(\beta) &= \nabla (-X^T y + X^T Mw) \\ &= \nabla (X^T Mw) \end{aligned}$$

$$= X^T M \nabla w$$

where the gradient of a vector field is a second order tensor (i.e. matrix).

$$\nabla w = \frac{\partial w_i}{\partial \beta_j} e_i \otimes e_j \quad \text{where } \{e_i\}_{i=1}^N \text{ is standard basis for } \mathbb{R}^N.$$

$$= \begin{bmatrix} \frac{\partial w_1}{\partial \beta_1} & \dots & \frac{\partial w_1}{\partial \beta_N} \\ \vdots & & \vdots \\ \frac{\partial w_N}{\partial \beta_1} & \dots & \frac{\partial w_N}{\partial \beta_N} \end{bmatrix}$$

$$= \left[\nabla_{\beta} w_1(\beta) \dots \nabla_{\beta} w_N(\beta) \right]^T$$

We've shown!

$$\nabla_{\beta} w_i(\beta) = w_i(1-w_i) x_i$$

\Rightarrow

$$\nabla w = \left[w_1(1-w_1)x_1 \dots w_N(1-w_N)x_N \right]^T$$

$$= \tilde{W} X$$

, W diagonal matrix
with $w_{ii} = w_i(1-w_i)$

Then for

So simple.

$$\nabla^2 \ell(\beta) = \nabla (X^T M w) = X^T M \nabla w = X^T M \tilde{W} X$$

So, by Taylor's theorem

$$l(\beta) \approx$$

$$l(\beta_0) - (X^T S)^T (\beta - \beta_0) + \frac{1}{2} (\beta - \beta_0)^T X^T A X (\beta - \beta_0)$$

$$\begin{cases} \text{with } A = M \tilde{W} \\ \tilde{W}_{ii} = w_i (1 - w_i) \\ M_{ii} = m_i \end{cases}$$

=

$$l(\beta_0) - S^T (X\beta - X\beta_0) + \frac{1}{2} (X\beta - X\beta_0)^T A (X\beta - X\beta_0)$$

=

$$l(\beta_0) - S^T (X\beta - \tilde{z}) + \frac{1}{2} (X\beta - \tilde{z})^T A (X\beta - \tilde{z})$$

completing
the
square!

wrt to
 $X\beta - \tilde{z}$

$$\propto \frac{1}{2} (X\beta - \tilde{z} - A^{-1}S)^T A (X\beta - \tilde{z} - A^{-1}S) + c$$

Flipping
around
signs

$$= \frac{1}{2} (\{\tilde{z} + A^{-1}S\} - X\beta)^T A (\{\tilde{z} + A^{-1}S\} - X\beta) + c$$

$$= \frac{1}{2} (z - X\beta)^T A (z - X\beta) + c$$

$$z = X\beta_0 + (M\tilde{W})^{-1}S, \quad A = M\tilde{W}$$