



## Simple linear regression

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- Regression analysis is the most widely used statistical tool for understanding relationships among variables
- It provides a conceptually simple method for investigating functional relationships between one or more factors and an outcome of interest
- The relationship is expressed in the form of an equation or a model connecting the response or dependent variable and one or more explanatory or predictor variable

# Why?



Straight prediction questions:

- For how much will my house sell?
- How many runs per game will the Red Sox score this year?
- Will this person like that movie? (Netflix rating system)

Explanation and understanding:

- What is the impact of MBA on income?
- How do the returns of a mutual fund relate to the market?
- Does Walmart discriminate against women regarding salaries?

# 1st Example: Predicting house prices



## Problem:

- Predict market price based on observed characteristics

## Solution:

- Look at property sales data where we know the price and some observed characteristics.
- Build a decision rule that predicts price as a function of the observed characteristics.



## What characteristics do we use?

We have to define the variables of interest and develop a specific quantitative measure of these variables

- Many factors or variables affect the price of a house
  - size
  - number of baths
  - garage, air conditioning, etc
  - neighborhood

# Predicting house prices



To keep things super simple, let's focus only on size.

The value that we seek to predict is called the **dependent (or output)** variable, and we denote this:

- $Y$  = price of house (e.g. thousands of dollars)

The variable that we use to guide prediction is the **explanatory (or input)** variable, and this is labelled

- $X$  = size of house (e.g. thousands of square feet)

# Predicting house prices



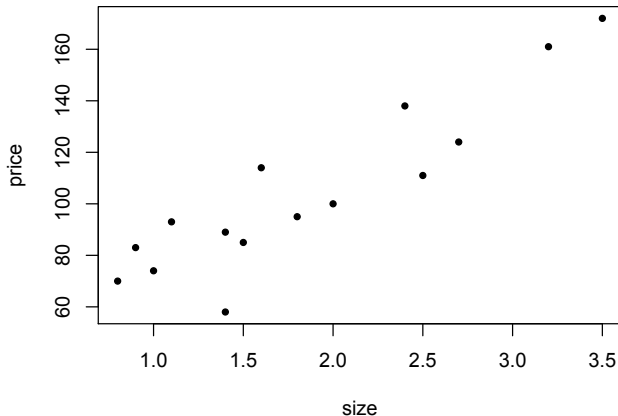
What does this data look like?

Size	Price
0.80	70
0.90	83
1.00	74
1.10	93
1.40	89
1.40	58
1.50	85
1.60	114
1.80	95
2.00	100
2.40	138
2.50	111
2.70	124
3.20	161
3.50	172

# Predicting house prices



It is much more useful to look at a scatterplot



In other words, view the data as points in the  $X \times Y$  plane.



# Regression model



$Y$  = response or outcome variable

$X$  = explanatory or input variables

A linear relationship is written

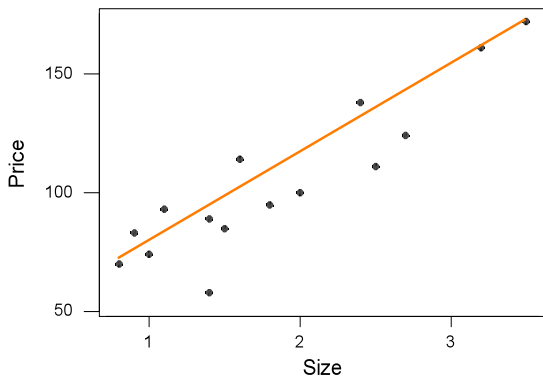
$$Y = b_0 + b_1X + e$$

# Linear prediction



There seems to be a linear relationship between price and size:

As size goes up, price goes up.



The line shown was fit by the “eyeball” method.



Recall that the equation of a line is:

$$Y = b_0 + b_1X + e$$

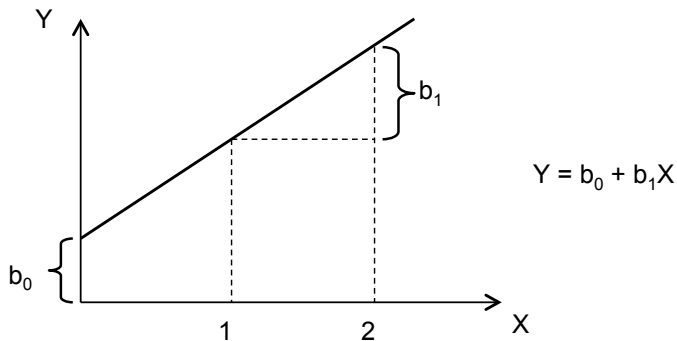
Where  $b_0$  is the **intercept** and  $b_1$  is the **slope**.

The intercept value is in units of  $Y$  (\$1,000).

The slope is in units of  $Y$  *per* units of  $X$  (\$1,000/1,000 sq ft).

The residual  $e$  is in units of  $Y$  (\$1,000).

# Linear prediction



Our “eyeball” line has  $b_0 = 35$ ,  $b_1 = 40$ .



We can now predict the price of a house when we know only the size; just read the value off the line that we've drawn.

For example, given a house with of size  $X = 2.2$ .

Predicted price  $\hat{Y} = 35 + 40(2.2) = 123$ .

Note: Conversion from 1,000 sq ft to \$1,000 is done for us by the slope coefficient ( $b_1$ )



Can we do better than the eyeball method?

We desire a strategy for estimating the slope and intercept parameters in the model  $\hat{Y} = b_0 + b_1X$

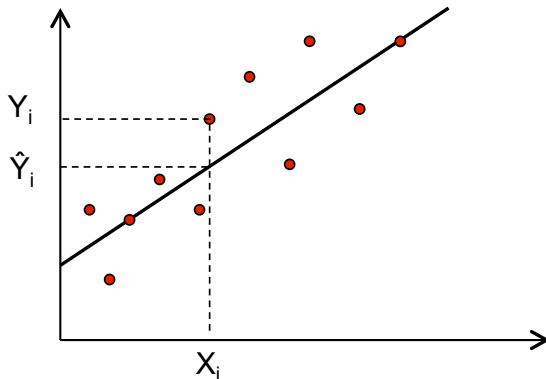
A reasonable way to fit a line is to minimize the amount by which the **fitted value** differs from the actual value.

This amount is called the **residual**.

# Linear prediction



What is the “fitted value”?

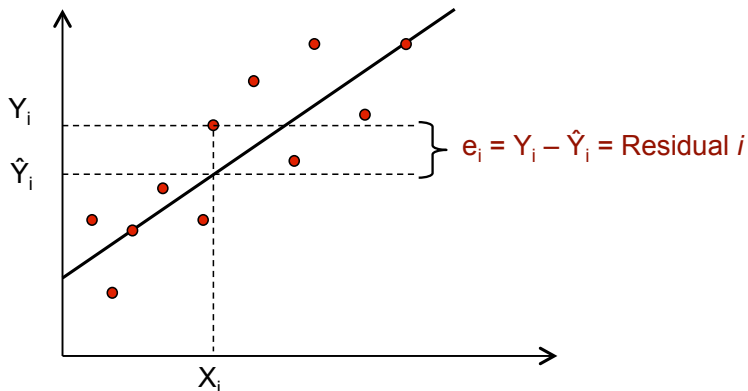


The dots are the observed values and the line represents our fitted values given by  $\hat{Y}_i = b_0 + b_1 X_i$ .

# Linear prediction



What is the “residual” for the  $i$ th observation?



We can write  $Y_i = \hat{Y}_i + (Y_i - \hat{Y}_i) = \hat{Y}_i + e_i$ .





Ideally we want to minimize the size of all residuals:

- If they were all zero we would have a perfect line.
- Trade-off between moving closer to some points and at the same time moving away from other points.

The line fitting process:

- Give weights to all of the residuals.
- Minimize the “total” of residuals to get best fit.

Least Squares chooses  $b_0$  and  $b_1$  to minimize  $\sum_{i=1}^N e_i^2$

$$\sum_{i=1}^N e_i^2 = e_1^2 + e_2^2 + \cdots + e_N^2 = (Y_1 - \hat{Y}_1)^2 + (Y_2 - \hat{Y}_2)^2 + \cdots + (Y_N - \hat{Y}_N)^2$$

# Least squares – excel output



## SUMMARY OUTPUT

<i>Regression Statistics</i>	
Multiple R	0.909209967
R Square	0.826662764
Adjusted R Square	0.81332913
Standard Error	14.13839732
Observations	15

## ANOVA

	<i>df</i>	<i>SS</i>	<i>MS</i>	<i>F</i>	<i>Significance F</i>
Regression	1	12393.10771	12393.10771	61.99831126	2.65987E-06
Residual	13	2598.625623	199.8942787		
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	<i>Coefficients</i>	<i>Standard Error</i>	<i>t Stat</i>	<i>P-value</i>	<i>Lower 95%</i>	<i>Upper 95%</i>
Intercept	38.88468274	9.09390389	4.275906499	0.000902712	19.23849785	58.53086763
Size	35.38596255	4.494082942	7.873900638	2.65987E-06	25.67708664	45.09483846



Remember how we get the slope ( $b_1$ ) and intercept ( $b_0$ ). We minimize the sum of squared prediction errors.

The formulas for  $b_0$  and  $b_1$  that minimize the least squares criterion are:

$$b_1 = r_{xy} \times \frac{s_y}{s_x} \quad b_0 = \bar{Y} - b_1 \bar{X}$$

where,

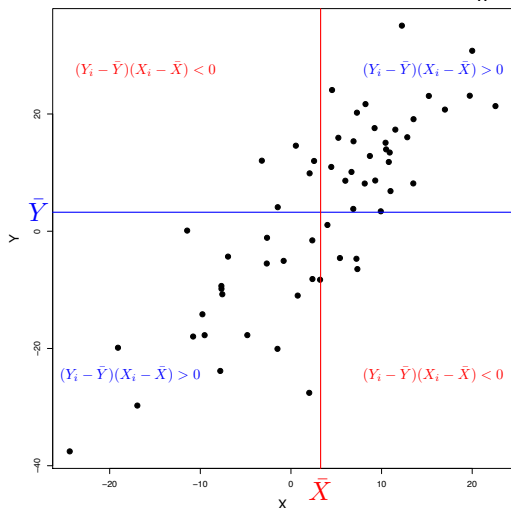
- $\bar{X}$  and  $\bar{Y}$  are the sample mean of  $X$  and  $Y$
- $\text{corr}(x, y) = r_{xy}$  is the sample correlation
- $s_x$  and  $s_y$  are the sample standard deviation of  $X$  and  $Y$

# Visual: Covariance



Measure the **direction** and **strength** of the linear relationship between  $Y$  and  $X$

$$\text{cov}(Y, X) = \frac{\sum_{i=1}^n (Y_i - \bar{Y})(X_i - \bar{X})}{n - 1}$$



–  $s_y = 15.98$ ,  $s_x = 9.7$

–  $\text{cov}(X, Y) = 125.9$

How do we interpret that?

# A standardized measure: Correlation



Correlation is the standardized covariance:

$$\text{corr}(X, Y) = \frac{\text{cov}(X, Y)}{\sqrt{s_x^2 s_y^2}} = \frac{\text{cov}(X, Y)}{s_x s_y}$$

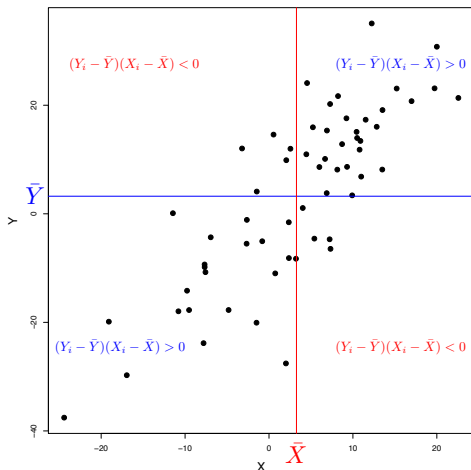
The correlation is scale invariant and the units of measurement don't matter: It is always true that  $-1 \leq \text{corr}(X, Y) \leq 1$ .

This gives the direction (negative or positive) and strength ( $0 \rightarrow 1$ ) of the linear relationship between  $X$  and  $Y$ .

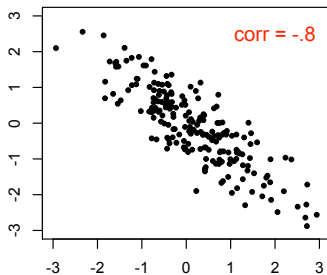
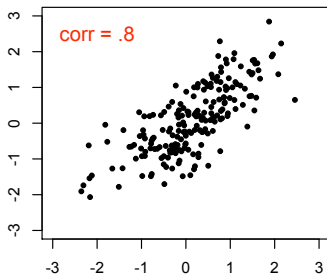
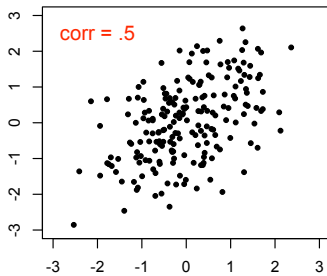
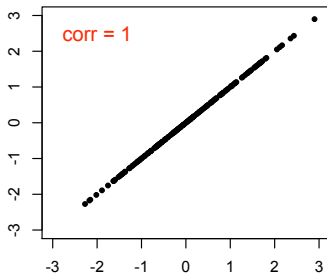
# Correlation



$$\text{corr}(Y, X) = \frac{\text{cov}(X, Y)}{\sqrt{s_x^2 s_y^2}} = \frac{\text{cov}(X, Y)}{s_x s_y} = \frac{125.9}{15.98 \times 9.7} = 0.812$$



# Correlation

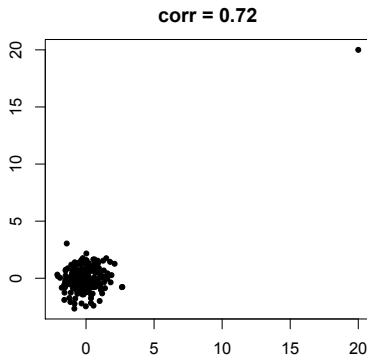
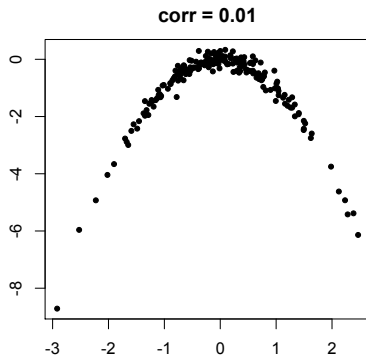


# Correlation



Only measures **linear** relationships:

$\text{corr}(X, Y) = 0$  does not mean the variables are not related!



Also be careful with influential observations!

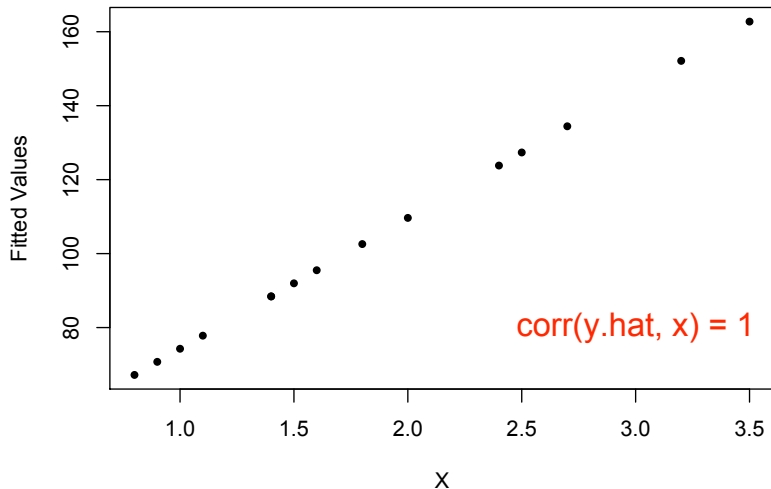




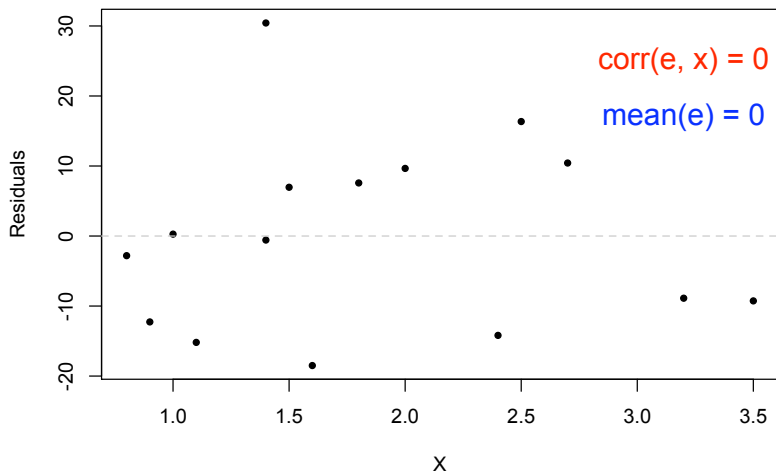
From now on, terms “fitted values” ( $\hat{Y}_i$ ) and “residuals” ( $e_i$ ) refer to those obtained from the least squares line.

The fitted values and residuals have some special properties. Let's look at the housing data analysis to figure out what these properties are...

# The fitted values and X



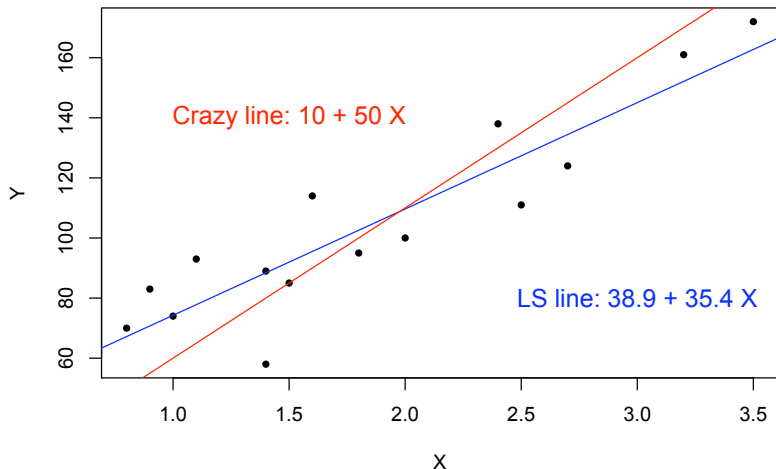
# The residuals and X



# Why?



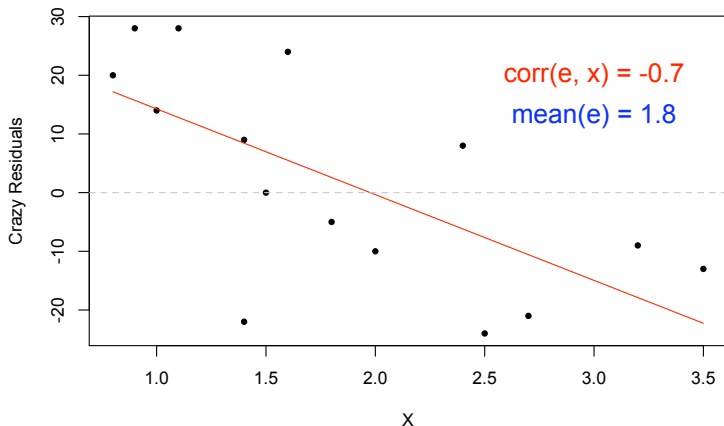
What is the intuition for the relationship between  $\hat{Y}$  and  $e$  and  $X$ ?  
Let's consider some "crazy" alternative line:



# Fitted values and residuals



**This is a bad fit!** We are underestimating the value of small houses and overestimating the value of big houses.



Clearly, we have left some predictive ability on the table!



As long as the correlation between  $e$  and  $X$  is non-zero, we could always adjust our prediction rule to do better.

We need to exploit all of the predictive power in the  $X$  values and put this into  $\hat{Y}$ , leaving no “ $X$ ness” in the residuals.

**In summary:**  $Y = \hat{Y} + e$  where:

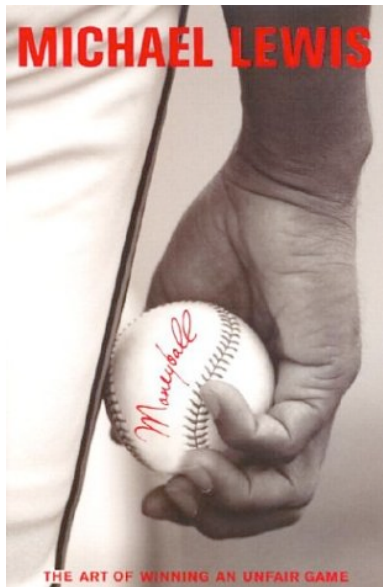
- $\hat{Y}$  is “made from  $X$ ”;  $\text{corr}(X, \hat{Y}) = 1$ .
- $e$  is unrelated to  $X$ ;  $\text{corr}(X, e) = 0$ .

## 2nd Example: Offensive performance in baseball



- Problems:
  - Evaluate/compare traditional measures of offensive performance
  - Help evaluate the worth of a player
- Solutions:
  - Compare *prediction rules* that forecast runs as a function of either AVG (batting average), SLG (slugging percentage) or OBP (on base percentage)

## 2nd Example: Offensive performance in baseball

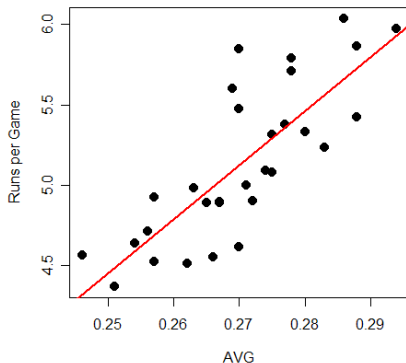




## Baseball data – using AVG



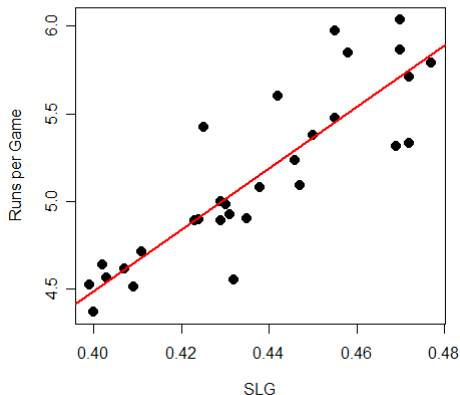
Each observation corresponds to a team in MLB. Each quantity is the average over a season.



–  $Y$  = runs per game;  $X$  = AVG (average)

LS fit:  $\text{Runs/Game} = -3.93 + 33.57 \text{ AVG}$

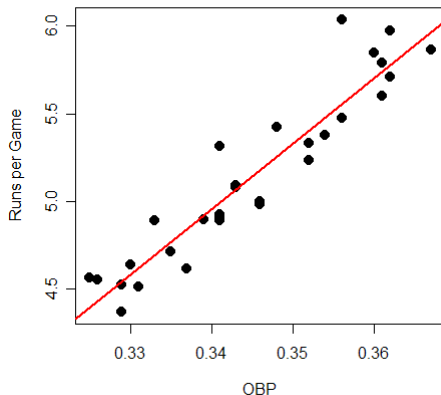
## Baseball data – using SLG



- $Y$  = runs per game
- $X$  = SLG (slugging percentage)

LS fit:  $\text{Runs/Game} = -2.52 + 17.54 \text{ SLG}$

## Baseball data – using OBP



- $Y$  = runs per game
- $X$  = OBP (on base percentage)

LS fit:  $\text{Runs/Game} = -7.78 + 37.46 \text{ OBP}$



- What is the best prediction rule?
- Let's compare the predictive ability of each model using the average squared error

$$\sqrt{\frac{1}{N} \sum_{i=1}^N e_i^2} = \sqrt{\frac{\sum_{i=1}^N \left( \widehat{Runs_i} - Runs_i \right)^2}{N}}$$

Place your money on OBP!!



Root Average Squared Error	
AVG	0.29
SLG	0.23
OBP	0.16



We can quantify the variability around the line by computing the variance of the residuals:

$$s = \sqrt{\frac{1}{n-2} \sum_{i=1}^n e_i^2} = \sqrt{\frac{\sum (Y_i - \hat{Y}_i)^2}{n-2}}$$

This is measured in the same units as  $Y$ . It's also called the **regression standard error**.

(Don't worry about this  $n - 2$  yet!).



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Total	14	14991.73333			

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$$R^2 = 0.82$$



# The simple linear regression model



The power of statistical inference comes from the ability to make precise statements about the accuracy of the forecasts.

In order to do this we must invest in a **probability model**.

Simple Linear Regression Model:  $Y = \beta_0 + \beta_1 X + \varepsilon$

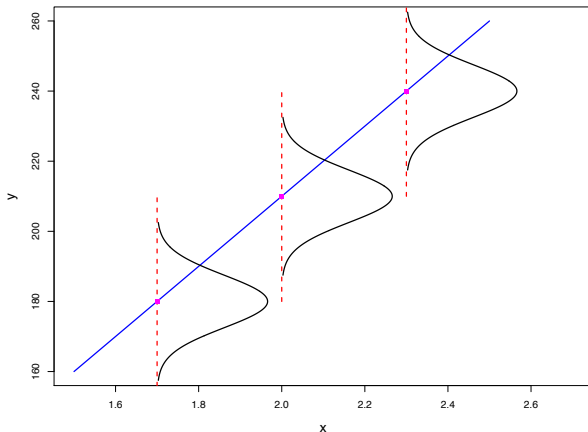
$$\varepsilon \sim N(0, \sigma^2)$$

- $\beta_0 + \beta_1 X$  represents the “true line”; The part of  $Y$  that depends on  $X$ .
- The error term  $\varepsilon$  is independent “idiosyncratic noise”; The part of  $Y$  not associated with  $X$ .

# The simple linear regression model



$$Y = \beta_0 + \beta_1 X + \varepsilon$$



The conditional distribution for  $Y$  given  $X$  is Normal:

$$Y|X = x \sim N(\beta_0 + \beta_1 x, \sigma^2).$$

# Back to the house data



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How much do our estimates depend on the particular random sample that we happen to observe? Imagine:

- Randomly draw different samples of the same size.
- For each sample, compute the estimates  $b_0$ ,  $b_1$ , and  $s$ .

If the estimates don't vary much from sample to sample, then it doesn't matter which sample you happen to observe.

If the estimates do vary a lot, then it matters which sample you happen to observe.

## Sampling distribution of $b_1$



The sampling distribution of  $b_1$  describes how estimator  $b_1 = \hat{\beta}_1$  varies over different samples with the  $X$  values fixed.

It turns out that  $b_1$  is normally distributed (approximately):  
 $b_1 \sim N(\beta_1, s_{b_1}^2)$ .

- $b_1$  is unbiased:  $E[b_1] = \beta_1$ .
- $s_{b_1}$  is the **standard error of  $b_1$** . In general, the standard error is the standard deviation of an estimate. It determines **how close**  $b_1$  is to  $\beta_1$ .
- This is a number directly available from the regression output.



Since  $b_1 \sim N(\beta_1, s_{b_1}^2)$ , Thus:

- 68% Confidence Interval:  $b_1 \pm 1 \times s_{b_1}$
- 95% Confidence Interval:  $b_1 \pm 2 \times s_{b_1}$
- 99% Confidence Interval:  $b_1 \pm 3 \times s_{b_1}$

Same thing for  $b_0$

- 95% Confidence interval:  $b_0 \pm 2 \times s_{b_0}$

The confidence interval provides you with a set of plausible values for the parameters

# Back to the house data



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$$[b_1 - 2 \times s_{b_1}; b_1 + 2 \times s_{b_1}] \approx [25.67; 45.09]$$



Suppose we want to assess whether or not  $\beta_1$  equals a proposed value  $\beta_1^0$ . This is called **hypothesis testing**.

Formally we test the null hypothesis:

$$H_0 : \beta_1 = \beta_1^0$$

vs. the alternative

$$H_1 : \beta_1 \neq \beta_1^0$$



# Testing, just like last time!



That are 2 ways we can think about testing:

1. Building a test statistic... the **t-stat**,

$$t = \frac{b_1 - \beta_1^0}{s_{b_1}}$$

This quantity measures how many standard deviations the estimate ( $b_1$ ) from the proposed value ( $\beta_1^0$ ).

If the absolute value of  $t$  is greater than 2, we need to worry (why?)... we **reject** the hypothesis.



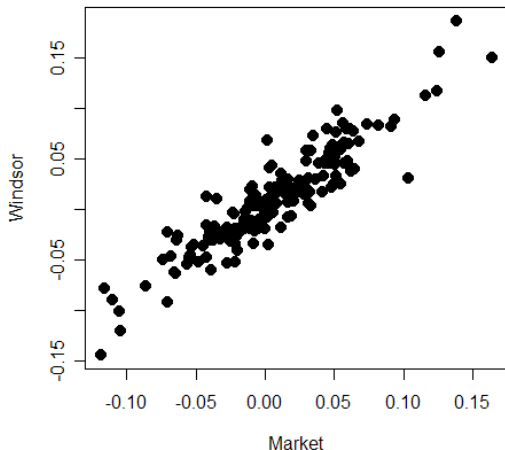
2. Looking at the **confidence interval**. If the proposed value is outside the confidence interval you **reject** the hypothesis.

Notice that this is equivalent to the t-stat. An absolute value for  $t$  greater than 2 implies that the proposed value is outside the confidence interval... therefore reject.

## Example: Mutual funds



Let's investigate the performance of the Windsor Fund, an aggressive large cap fund by Vanguard...



The plot shows monthly returns for Windsor vs. the S&P500

## Example: Mutual funds



Consider a CAPM regression for the Windsor mutual fund.

$$r_w = \beta_0 + \beta_1 r_{sp500} + \epsilon$$

Let's first test  $\beta_1 = 0$

$H_0 : \beta_1 = 0$ . Is the Windsor fund related to the market?

$H_1 : \beta_1 \neq 0$

# Example: Mutual funds



Regression Statistics	
Multiple R	0.923417768
R Square	0.852700374
Adjusted R Square	0.851872848
Standard Error	0.018720015
Observations	180

ANOVA					
	df	SS	MS	F	Significance F
Regression	1	0.3611	0.361099761	1030.421266	6.0291E-76
Residual	178	0.062378	0.000350439		
Total	179	0.423478			

	Coefficients	Standard Error	t Stat	P-value	Lower 95%	Upper 95%	Lower 95.0%	Upper 95.0%
Intercept	0.003646881	0.001409	2.587596412	0.010462425	0.000865657	0.006428	0.000866	0.006428
X Variable 1	0.935717012	0.02915	32.10017549	6.0291E-76	0.878193151	0.993241	0.878193	0.993241

$$b_1$$

$$s_{b_1}$$

$$\frac{b_1}{s_{b_1}}$$

- $t = 32.10...$  reject  $\beta_1 = 0!!$
- the 95% confidence interval is  $[0.87; 0.99]...$  again, reject!!

## Example: Mutual funds



Now let's test  $\beta_1 = 1$ . What does that mean?

$H_0 : \beta_1 = 1$  Windsor is as risky as the market.

$H_1 : \beta_1 \neq 1$  and Windsor softens or exaggerates market moves.

We are asking whether or not Windsor moves in a different way than the market (e.g., is it more conservative?).

# Example: Mutual funds



Regression Statistics	
Multiple R	0.923417768
R Square	0.852700374
Adjusted R Square	0.851872848
Standard Error	0.018720015
Observations	180

ANOVA					
	df	SS	MS	F	Significance F
Regression	1	0.3611	0.361099761	1030.421266	6.0291E-76
Residual	178	0.062378	0.000350439		
Total	179	0.423478			

	Coefficients	Standard Error	t Stat	P-value	Lower 95%	Upper 95%	Lower 95.0%	Upper 95.0%
Intercept	0.003646881	0.001409	2.587596412	0.010462425	0.000865657	0.006428	0.000866	0.006428
X Variable 1	0.935717012	0.02915	32.10017549	6.0291E-76	0.878193151	0.993241	0.878193	0.993241

$$b_1$$

$$s_{b_1}$$

$$\frac{b_1}{s_{b_1}}$$

$$- t = \frac{b_1 - 1}{s_{b_1}} = \frac{-0.0643}{0.0291} = -2.205... \text{ reject.}$$

– the 95% confidence interval is [0.87; 0.99]... again, reject,  
but...



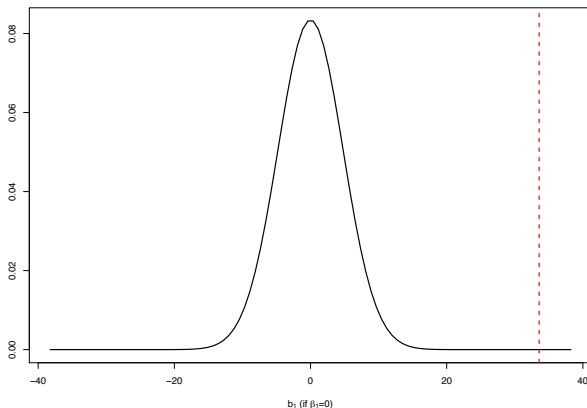
- The  $p$ -value provides a measure of how weird your estimate is if the null hypothesis is true
- Small  $p$ -values are evidence against the null hypothesis
- In the AVG vs. Runs per Game example...  $H_0 : \beta_1 = 0$ . How weird is our estimate of  $b_1 = 33.57$ ?
- Remember:  $b_1 \sim N(\beta_1, s_{b_1}^2)$ ... If the null was true ( $\beta_1 = 0$ ),  
 $b_1 \sim N(0, s_{b_1}^2)$



# P-values



- Where is 33.57 in the picture below?



The  $p$ -value is the probability of seeing  $b_1$  equal or greater than 33.57 in absolute terms. Here,  $p\text{-value}=0.000000124!!$

Small  $p$ -value = bad null!



–  $H_0 : \beta_1 = 0 \dots$  **p-value = 1.24E-07...** reject!

## SUMMARY OUTPUT

<i>Regression Statistics</i>	
Multiple R	0.798496529
R Square	0.637596707
Adjusted R Square	0.624653732
Standard Error	0.298493066
Observations	30

## ANOVA

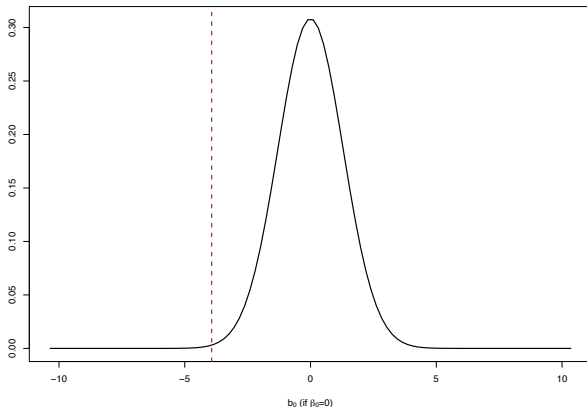
	<i>df</i>	<i>SS</i>	<i>MS</i>	<i>F</i>	<i>significance F</i>
Regression	1	4.38915033	4.38915	49.26199	1.239E-07
Residual	28	2.494747094	0.089098		
Total	29	6.883897424			

	<i>Coefficients</i>	<i>Standard Error</i>	<i>t Stat</i>	<i>P-value</i>	<i>Lower 95%</i>	<i>Upper 95%</i>
Intercept	-3.936410446	1.294049995	-3.04193	0.005063	-6.587152	-1.2856692
AVG	33.57186945	4.783211061	7.018689	1.24E-07	23.773906	43.369833

# P-values



- How about the intercept,  $H_0 : \beta_0 = 0$ ? How weird is  $b_0 = -3.936$ ?



The  $p$ -value (the probability of seeing  $b_0$  equal or greater than -3.936 in absolute terms) is **0.005**.

Small  $p$ -value = bad null!



- $H_0 : \beta_0 = 0 \dots$  **p-value = 0.005...** we still reject, but not with the same strength.

## SUMMARY OUTPUT

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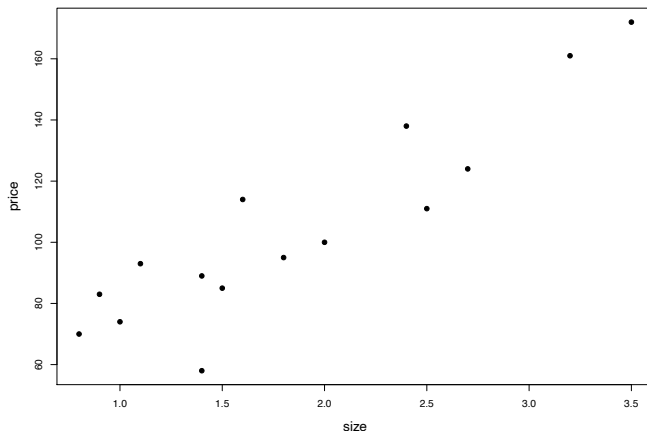


- Large  $t$  or small  $p$ -value mean the same thing...
- $p$ -value  $< 0.05$  is equivalent to a  $t$ -stat  $> 2$  in absolute value
- Small  $p$ -value means something weird happen if the null hypothesis was true...
- Bottom line, **Small  $p$ -value  $\rightarrow$  REJECT! Large  $t \rightarrow$  REJECT!**
- But remember, **always look at the confidence interval!**

# House data – one more time!



- $R^2 = 82\%$
- Great  $R^2$ , we are happy using this model to predict house prices, right?



# House data – one more time!



- But,  $s = 14$  leading to a predictive interval width of about US\$60,000!! How do you feel about the model now?
- As a practical matter,  $s$  is a much more relevant quantity than  $R^2$ . Once again, *intervals* are your friend!

