# MA 575 HW 9

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#### Exercise 9.1

(a)

First, to find the eigenvalues of  $X^TX$  we solve

$$X^T X v = \gamma v$$
$$(X^T X - \gamma I)v = 0$$

Here  $\gamma$  is the eigenvalue and v is the eigenvector corresponding to  $X^TX$ . Now, let  $Y^TY = X^TX + \lambda I$ . Then we can find the eigenvalues of  $Y^TT$  by solving the equation

$$Y^T Y u = \theta u$$
$$(Y^T Y - \theta I)u = 0$$

Here  $\theta$  is the eigenvalue and u is the eigenvector corresponding to  $Y^TY$ . But notice, by the way we defined  $Y^TY$  we can reduce this further by

$$(Y^{T}Y - \theta I)u = 0$$
$$(X^{T}X + \lambda I - \theta I)u = 0$$
$$(X^{T}X - (\theta - \lambda)I)u = 0$$

Notice that this was just the system we were trying to solve above. That is the eigenvalues and eigenvectors  $(\gamma_i, v_i)$  corresponding to  $X^T X$  are in direct correspondence to the  $(\theta_i - \lambda, u_i)$  pairs. But since  $\lambda$  is constant in all of these equations, we have  $\lambda_i(X^T X) = \lambda_i(Y^T Y) - \lambda$ . This gives

$$\lambda_i(X^TX + \lambda I) = \lambda_i(X^TX) + \lambda$$

Therefore, we can write the condition number as

$$\kappa(X^TX + \lambda I) = \frac{\lambda_m(X^TX + \lambda I)}{\lambda_1(X^TX + \lambda I)} = \frac{\lambda_m(X^TX) + \lambda}{\lambda_1(X^TX) + \lambda}$$

(b)

```
#read in/format data
dat = read.csv("~/Desktop/Courses/MA 575/book_data/reducedbikedata2011.csv")
#form the p + 1 data matrix
X = as.matrix(dat[,-c(1,2)])
#make the design matrix
design = t(X) %*% X
#get eigen values from design
evals = eigen(design)$values
max = max(evals)
min = min(evals)
#define data structures
lambda \leftarrow 10^{\circ}seq(20, -2, length = 100)
Y = length(lambda)
#get condition values
for(i in 1:length(lambda)){
  val = (max + lambda[i])/(min + lambda[i])
  Y[i] = log(val, 10)
}
#make plot of condition numbers
plot(log(lambda, 10), Y, type = "l", xlab = "Lambda", ylab = "Condition Number", main = "", lt
      10
Condition Number
      \infty
      9
      ^{\circ}
      0
                   0
                                  5
                                                 10
                                                                15
                                                                               20
                                           Lambda
```

From a numerical point of view, larger  $\lambda$  will stabilize the inversion of  $X^TX$  by adding larger and

larger values to the main diagnose via the correction  $\lambda I$ . But notice as we let  $\lambda \to \infty$ , we require that all  $\beta$  values are zero. This is because as we let any  $\beta > 0$ , then the penalty will inflate the  $RSS_{RIDGE}$  to infinity. Therefore, we face a trade off - numerical stability and penalization for problems with several predictors (i.e. n << p) and the inferential task of relating the covariates to Y. Therefore, we should choose a  $\lambda$  value that penalizes enough to ensure stability and provide structure to high dimensional problems while still allowing the model to fit the trends in the data.

#### Exercise 9.2

We will fit all three models then comment below.

```
#-----
#
       Data Prep
#load in data
swiss <- datasets::swiss
#create model matrix/response vector
x = model.matrix(Fertility~., swiss)[,-1]
y = swiss$Fertility
#set up training/testing sets
set.seed(489)
train = sample(1:nrow(x), nrow(x)/2)
test = (1:nrow(x))[-train]
#
       OLS
#-----
#OLS model/prediction
ols = lm(y~., data = swiss[,-1], subset = train)
#Prediction
pred_ols = as.numeric(predict(ols, newdata = swiss[test,-1]))
#Get MSE
mse_ols = mean((y[test] - pred_ols)^2)
       Ridge
```

```
#
#-----
#fit Ridge
library("glmnet")

## Loading required package: Matrix

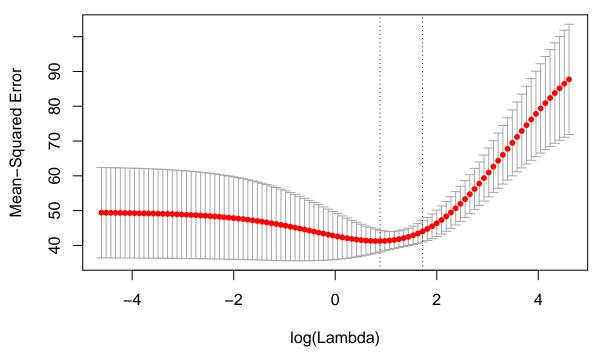
## Loading required package: foreach

## Loaded glmnet 2.0-13

lam = 10^seq(2, -2, length = 100)
ridge = glmnet(x[train,], y[train], lambda = lam, alpha = 0)

#CV to find lambda_opt
cv = cv.glmnet(x[train,], y[train], lambda = lam, nfolds = 5, alpha = 0)
plot(cv)
```

### 

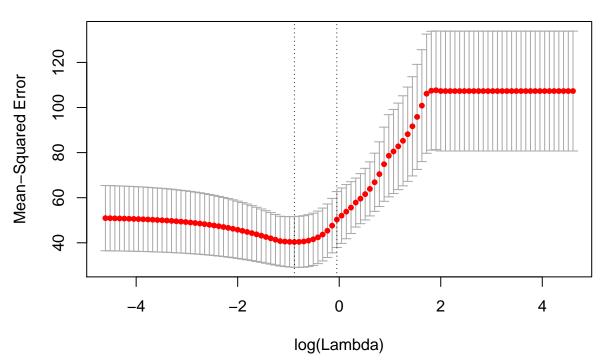


```
lam_opt_ridge = cv$lambda.min

#find predictions
pred_ridge = predict(ridge, s= lam_opt_ridge, newx = x[test,])

#MSE calculations
mse_ols = mean((y[test] - pred_ols)^2)
mse_ridge = mean((y[test] - pred_ridge)^2)
```

5 5 5 5 5 5 5 5 5 5 5 4 3 3 0 0 0 0 0 0



```
message(paste("MSE OLS:", round(mse_ols,3),"\n MSE Lasso:", round(mse_lasso,3),"\n MSE Ridge:"
## MSE OLS: 106.009
## MSE Lasso: 103.738
## MSE Ridge: 93.136
#get best coef
betas = data.frame(B.OLS = coef(ols))
betas$B.Ridge = predict(ridge, type = "coefficients", s = lam_opt_ridge)[1:6,]
betas$B.LASSO = predict(lasso, type = "coefficients", s = lam_opt_lasso)[1:6,]
#print all coefficents
print(betas)
##
                          B.OLS
                                    B.Ridge
                                               B.LASSO
## (Intercept)
                    74.63669064 62.45392671 68.0347328
## Agriculture
                    -0.27810752 -0.12381476 -0.1931092
## Examination
                    -0.93921916 -0.48843634 -0.7672501
## Education
                    -0.35970838 -0.33455989 -0.2666706
## Catholic
                     0.06498258 0.06201879 0.0569043
## Infant.Mortality 1.37617843 1.23132670 1.3273270
#plot coefficents (except for intercept)
plot(1:5,betas[2:6,1], lty = 1, type = "l", ylim = c(-2,2), ylab = "Beta_i", xlab = "i")
points(1:5,betas[2:6,2],lty = 2, type = "1")
points(1:5,betas[2:6,3], lty = 4, type = "1")
legend("topleft", legend = c("OLS", "Ridge", "LASSO"), lty = c(1,2,4))
                 OLS
                 Ridge
                 LASSO
           1
                           2
                                          3
                                                         4
                                                                         5
```

Here we see that the MSE decreases as we increase our penalty. That is we have

$$MSE_{OLS} > MSE_{Lasso} > MSE_{Ridge}$$

Ridge performs very well, with sizable MSE reduction. I included a table of the final  $\beta$  estimates from each model. Also included is a plot for the  $\beta$  estimates except for the intercept. We notice as  $\lambda$  increases, the  $\beta$  terms tend to zero. This matches our intuition/interpretation as Lasso and Ridge as shrinkage estimators. For instance for  $\beta_2$  we see great shrinkage, but for more important terms in the model (i.e. infant mortality) that the coefficients remain the same, even though it is statistically different than zero. This is a case where the penalization schemes improve on the OLS models significantly.