

Suppose that we have $X_1, X_2, \dots, X_n | \mu, \sigma^2 \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$

Goal: deduce a conjugate prior

$$P(X|\mu, \sigma^2) \propto (\sigma^2)^{-n/2} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^n (x_i - \mu)^2 \right\}$$

$$\propto (\sigma^2)^{-n/2} \exp \left\{ -\frac{1}{2\sigma^2} \left[\underbrace{\sum_i (x_i - \bar{x})^2}_{S_x^2} + n(\bar{x} - \mu)^2 \right] \right\}$$

$$\propto (\sigma^2)^{-n/2} \exp \left\{ -\frac{S_x^2}{\sigma^2} - \frac{(\bar{x} - \mu)^2}{2\sigma^2/n} \right\}$$

Idea: define $P(\mu, \sigma^2) = P(\mu | \sigma^2) P(\sigma^2)$ prior as a hierarchy

$$P(\mu, \sigma^2) = P(\sigma^2 | \mu) P(\mu)$$

Conditional on σ^2 : $P(\mu | \sigma^2)$ and we know $\bar{X} | \mu, \sigma^2 \sim N(\mu, \sigma^2/n)$

$$\text{which suggests } P(\mu | \sigma^2) \propto \exp \left\{ -\frac{(\bar{x} - \mu)^2}{2\sigma^2/n} \right\}$$

$$\text{and } \mu | \sigma^2 \sim N(\mu_0, \sigma^2/k_0)$$

Thus the conditional likelihood is

$$P(X, \mu | \sigma^2) \propto (\sigma^2)^{-n/2} \exp \left\{ -\frac{S_x^2}{2\sigma^2} - \frac{(\bar{x} - \mu)^2}{2\sigma^2/n} \right\} (\sigma^2)^{-1/2} \exp \left\{ -\frac{(\mu - \mu_0)^2}{2\sigma^2/k_0} \right\}$$

$$= (\sigma^2)^{-n/2} (\sigma^2)^{-1/2} \exp \left\{ -\frac{1}{2\sigma^2} \left(S_x^2 + (\bar{x} - \mu)^2 + k_0 (\mu - \mu_0)^2 \right) \right\}$$

With $n(\bar{x} - \mu)^2 + k_0(\mu - \mu_0)^2 = (k_0 + n)(\mu - \mu_n)^2 + \frac{k_0 n}{k_0 + n}(\bar{x} - \mu_n)^2$

where $\mu_n = \frac{n\bar{x} + k_0\mu_0}{n + k_0} = \frac{n}{n + k_0}\bar{x} + \frac{k_0}{n + k_0}\mu_0$ Acts as "prior sample size"

$$P(x, \mu | \sigma^2) \propto (\sigma^2)^{-n/2} \exp \left\{ -\frac{1}{2\sigma^2} \left[s_x^2 + \frac{k_0 n}{n + k_0} (\bar{x} - \mu_n)^2 \right] \right\} (\sigma^2)^{-1/2} \exp \left\{ -\frac{(n + k_0)(\mu - \mu_n)^2}{2\sigma^2} \right\}$$

Thus $\mu | x, \sigma^2 \sim N(\mu_n, \sigma^2 / (n + k_0))$.

But we want

$$P(x | \sigma^2) = \int P(x, \mu | \sigma^2) d\mu = (\sigma^2)^{-n/2} \exp \left\{ -\frac{1}{2\sigma^2} \left[s_x^2 + \frac{k_0 n}{n + k_0} (\bar{x} - \mu_n)^2 \right] \right\}$$

Thus for conjugacy

$$P(\sigma^2) \propto (\sigma^2)^{-(r_0/2 + 1)} \exp \left(-\frac{r_0 \sigma_0^2}{2\sigma^2} \right) \sim \text{Inv. } \chi^2(v_0, \sigma_0^2)$$

$$P(x, \sigma^2) \propto (\sigma^2)^{-(\frac{n+r_0}{2} + 1)} \exp \left\{ -\frac{1}{2\sigma^2} \left(s_x^2 + \frac{k_0 n}{n + k_0} (\bar{x} - \mu_n)^2 + r_0 \sigma_0^2 \right) \right\} \frac{n + r_0}{n + v_0}$$

$$\sigma^2 | x \sim \text{Inv } \chi^2 \left(n + r_0, \frac{1}{n + r_0} \left(s_x^2 + \frac{k_0 n}{n + k_0} (\bar{x} - \mu_n)^2 + r_0 \sigma_0^2 \right) \right)$$

We write $\mu, \sigma^2 \sim N\text{-Inv}\chi^2(\mu_0, k_0; r_0, \sigma_0^2)$

meaning $\mu | \sigma^2 \sim N(\mu_0, \sigma_0^2 / k_0)$

$$\sigma^2 \sim \text{Inv}\chi^2(r_0, \sigma_0^2)$$

with posterior

$$\mu, \sigma^2 | X \sim N\text{-Inv}\chi^2(\mu_n, k_n; r_n, \sigma_n^2)$$

It can be shown $\frac{\mu - \mu_n}{\sigma_n / \sqrt{k_n}} \sim t_{r_n}$, $P(\mu | X) = \int P(\mu | X, \sigma^2) P(\sigma^2 | X) d\sigma^2$

Consider the following study:

$\log(\text{Dose})$	# Animals	# deaths
x_1	n_1	y_1
x_2	n_2	y_2
x_3	n_3	y_3
\vdots	\vdots	\vdots
x_m	n_m	y_m

We can assume $y_i | \theta_i \sim \text{Binom}(n_i, \theta_i)$

$$\theta_i = f(x_i)$$

Choose to model - logistic regression

$$\text{logit}(\theta_i) = \log \frac{\theta_i}{1 - \theta_i} = \alpha + \beta x_i$$

Under this setup $y_i | \alpha, \beta \sim \text{Binom}(n_i, \text{logit}^{-1}(\alpha + \beta x_i))$

$$P(y | \alpha, \beta) \propto \prod_i \text{logit}^{-1}(\alpha + \beta x_i)^{y_i} [1 - \text{logit}^{-1}(\alpha + \beta x_i)]^{n_i - y_i}$$

Conjugacy hard \implies numerical methods

Two common choices

$$(a) P(\alpha, \beta) \propto 1 \quad (b) \begin{pmatrix} \alpha \\ \beta \end{pmatrix} \sim N\left(\begin{pmatrix} \tilde{\alpha}_0 \\ \tilde{\beta}_0 \end{pmatrix}, \Sigma\right)$$

Then, gridding over \oplus and numerically calculate

$$P(\alpha, \beta | y) \propto \frac{P(y | \alpha, \beta) P(\alpha, \beta)}{\sum_{\tilde{\alpha}, \tilde{\beta}} P(y | \tilde{\alpha}, \tilde{\beta}) P(\tilde{\alpha}, \tilde{\beta})}$$