

**Exercise 3.3.2** Let  $Y_n$  be the number of balls in urn  $A$  at time  $n$ . Now, we have three cases; (1)  $Y_{n+1} = Y_n + 1$  when the ball is selected from  $B$  and urn  $A$  is selected (2)  $Y_{n+1} = Y_n - 1$  when the ball is selected from  $A$  and urn  $B$  is selected, and finally (3)  $Y_{n+1} = Y_n$  when the ball is selected from  $A$  and urn  $A$  is selected *or* the ball is selected from  $B$  and urn  $B$  is chosen. Choosing a ball from  $A$  at time  $n$  has probability  $Y_n/N$  and choosing a ball from  $B$  at time  $N$  happens with probability  $1 - Y_n/N$ . Using this, with the fact that the choice of ball and urn is independent we have

$$\mathbf{P}_{ij} = \begin{cases} (1 - i/N)p & j = i + 1 \\ (i/N)p + (1 - i/N)q & j = i \\ (i/N)q & j = i - 1 \\ 0 & |i - j| \geq 2 \end{cases}$$

**Exercise 3.4.3** (a) Define  $T = \min\{n : T \in \{0, 3\}\}$  and let  $u_i = \mathbb{P}[X_T = 0 | X_0 = i]$ . Our goal is to then find  $u_1$ . Trivially,  $u_0 = 1$  and  $u_3 = 0$ . Moreover, using the law of total probability

$$u_1 = \mathbb{P}[X_T | X_0 = 1] = \sum_{j=0}^3 \mathbb{P}[X_T | X_1 = j] P_{1j} = \sum_{j=0}^3 u_j P_{1j} = P_{10} + P_{11}u_1 + P_{12}u_2$$

$$u_2 = \mathbb{P}[X_T | X_0 = 2] = \sum_{j=0}^3 \mathbb{P}[X_T | X_1 = j] P_{2j} = \sum_{j=0}^3 u_j P_{2j} = P_{20} + P_{21}u_1 + P_{22}u_2$$

Solving the second equation for  $u_2$ , we see that

$$u_2 = \frac{P_{20} + P_{21}u_1}{1 - P_{22}}$$

Now using the first equation to solve for  $u_1$ , we have

$$\begin{aligned} u_1 &= P_{10} + P_{11}u_1 + \frac{P_{20} + P_{21}u_1}{1 - P_{22}} \\ u_1 \left( 1 - P_{11} - \frac{P_{12}P_{21}}{1 - P_{22}} \right) &= P_{10} + \frac{P_{12}P_{20}}{1 - P_{22}} \\ u_1 &= \frac{P_{10}(1 - P_{22}) + P_{12}P_{20}}{(1 - P_{11})(1 - P_{22}) - P_{12}P_{21}} \end{aligned}$$

Finally, using the values from the transition matrix, we have  $u_1 \approx 0.3809524$

(b) Let  $v_i = \mathbb{E}[T | X_0 = i]$ . Seeing that our chain starts in state 1, we seek to find  $v_1$ . First note, trivially, that  $v_0 = v_3 = 0$ . Focusing on the two transient states, we have

$$v_1 = \mathbb{E}[T | X_0 = 1] = \sum_{j=0}^3 \mathbb{E}[T + 1 | X_1 = j] P_{1j} = 1 + \sum_{j=0}^3 v_j P_{1j} = 1 + P_{11}v_1 + P_{12}v_2$$

$$v_2 = \mathbb{E}[T|X_0 = 2] = \sum_{j=0}^3 \mathbb{E}[T + 1|X_1 = j]P_{2j} = 1 + \sum_{j=0}^3 v_j P_{2j} = 1 + P_{21}v_1 + P_{22}v_2$$

Solving the second equation for  $v_2$ , we get

$$v_2 = \frac{1 + P_{21}v_1}{1 - P_{22}}$$

Plugging this into the first equation we see

$$\begin{aligned} v_1 &= 1 + P_{11}v_1 + \frac{1}{1 - P_{22}} + \frac{P_{21}}{1 - P_{22}}v_1 \\ v_1 \left(1 - P_{11} - \frac{P_{21}}{1 - P_{22}}\right) &= 1 + \frac{1}{1 - P_{22}} \\ v_1 &= \frac{(1 - P_{22}) + P_{12}}{(1 - P_{22})(1 - P_{11}) - P_{12}P_{21}} \end{aligned}$$

Finally, using the quantities from the transition probability matrix, we have  $v_1 \approx 3.33$ .

**Exercise 3.4.7** Let  $\delta_i(X_n) = \mathbf{1}_{\{X_n=i\}}$  and define  $T := \min\{n : X_T \in \{0, 3\}\}$ . Then the value  $W_{ik} = \mathbb{E}\left[\sum_{t=0}^{T-1} \delta_i(X_n)|X_0 = k\right]$  is the expected time we visit state  $i$  given we start in state  $k$ . Our goal is to find  $W_{11}$  and  $W_{21}$ . First note that  $W_{i0} = W_{i3} = 0$  for  $i = 1, 2$ . Now, by first step analysis we have that

$$\begin{aligned} W_{11} &= \mathbb{E}\left[\sum_{t=0}^{T-1} \delta_1(X_n)|X_0 = 1\right] = 1 + \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_1(X_n)|X_0 = 1\right] \\ &= 1 + \sum_{j=0}^3 \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_1(X_n)|X_1 = j\right] P_{1j} = 1 + \sum_{j=0}^3 W_{1j}P_{1j} \\ &= 1 + W_{11}P_{11} + W_{12}P_{12} \end{aligned}$$

$$\begin{aligned} W_{12} &= \mathbb{E}\left[\sum_{t=0}^{T-1} \delta_1(X_n)|X_0 = 2\right] = \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_1(X_n)|X_0 = 2\right] \\ &= \sum_{j=0}^3 \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_1(X_n)|X_1 = j\right] P_{2j} = \sum_{j=0}^3 W_{1j}P_{2j} \\ &= W_{11}P_{21} + W_{12}P_{22} \end{aligned}$$

Now, solving the second equation for  $W_{12} = \frac{P_{21}}{1 - P_{22}}W_{11}$ . Using this in our first equa-

tion, we arrive at the following.

$$\begin{aligned}
 W_{11} &= 1 + W_{11}P_{11} + \frac{P_{21}P_{12}}{1 - P_{22}}W_{11} \\
 W_{11}\left(1 - P_{11} - \frac{P_{21}P_{12}}{1 - P_{22}}\right) &= 1 \\
 W_{11} &= \frac{(1 - P_{22})}{(1 - P_{11})(1 - P_{22}) + P_{21}P_{12}}
 \end{aligned}$$

Using the transition probability matrix, we then see that  $W_{11} \approx 1.8182$ . Proceeding in the same way as above, we have the following two equations

$$\begin{aligned}
 W_{21} &= \mathbb{E}\left[\sum_{t=0}^{T-1} \delta_2(X_n) | X_0 = 1\right] = \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_2(X_n) | X_0 = 1\right] \\
 &= \sum_{j=0}^3 \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_2(X_n) | X_1 = j\right] P_{1j} = \sum_{j=0}^3 W_{2j}P_{1j} \\
 &= W_{21}P_{11} + W_{22}P_{12}
 \end{aligned}$$

$$\begin{aligned}
 W_{22} &= \mathbb{E}\left[\sum_{t=0}^{T-1} \delta_2(X_n) | X_0 = 2\right] = 1 + \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_2(X_n) | X_0 = 2\right] \\
 &= 1 + \sum_{j=0}^3 \mathbb{E}\left[\sum_{t=1}^{T-1} \delta_2(X_n) | X_1 = j\right] P_{2j} = 1 + \sum_{j=0}^3 W_{2j}P_{2j} \\
 &= 1 + W_{21}P_{21} + W_{22}P_{22}
 \end{aligned}$$

Solving the second equation, we see that  $W_{22} = \frac{1+W_{21}P_{21}}{1-P_{22}}$ . Now using this in the first equation we see the following.

$$\begin{aligned}
 W_{21} &= W_{21}P_{11} + \frac{P_{12}}{1 - P_{22}} + \frac{P_{12}P_{21}}{1 - P_{22}}W_{21} \\
 W_{21}\left(1 - P_{11} - \frac{P_{12}P_{21}}{1 - P_{22}}\right) &= \frac{P_{12}}{1 - P_{22}} \\
 W_{21} &= \frac{P_{12}}{(1 - P_{11})(1 - P_{22}) - P_{12}P_{21}}
 \end{aligned}$$

Using the values from the transition probability matrix, we have  $W_{21} \approx 2.2728$ .

To find the time until absorption let  $w_i = \mathbb{E}[T | X_0 = i]$ . Then we have  $w_0 = w_3 = 0$  and moreover

$$w_1 = \mathbb{E}[T | X_0 = 1] = \sum_{j=0}^3 \mathbb{E}[T + 1 | X_1 = j] P_{1j} = 1 + P_{11}w_1 + P_{12}w_2$$

$$w_2 = \mathbb{E}[T|X_0 = 2] = \sum_{j=0}^3 \mathbb{E}[T + 1|X_1 = j]P_{2j} = 1 + P_{21}w_1 + P_{22}w_2$$

Now, solving the first equation for  $w_2$ , we see that  $w_2 = \frac{1}{1-P_{22}} + \frac{P_{21}}{1-P_{22}}w_1$ . Using this in the first equation we have

$$\begin{aligned} w_1 &= 1 + P_{11}w_1 + \frac{P_{12}}{1-P_{22}} + \frac{P_{12}P_{21}}{1-P_{22}}w_1 \\ w_1 \left(1 - P_{11} - \frac{P_{12}P_{21}}{1-P_{22}}\right) &= 1 + \frac{P_{12}}{1-P_{22}} \\ w_1 &= \frac{(1-P_{22}) + P_{12}}{(1-P_{11})(1-P_{22}) - P_{12}P_{21}} \end{aligned}$$

Now notice that from above,

$$w_1 = \frac{1-P_{22}}{(1-P_{11})(1-P_{22}) - P_{12}P_{21}} + \frac{P_{12}}{(1-P_{11})(1-P_{22}) - P_{12}P_{21}} = W_{11} + W_{21}$$

Numerically, we see that  $w_1 \approx 4.091 \approx 2.2728 + 1.8182$

**Problem 3.2.4** First note that  $Z_n = (X_{n-1}, X_n)$  and  $Z_{n+1} = (X_n, X_{n+1})$  so any transition of the form  $(a, b) \rightarrow (c, d)$  will have zero probability if  $b \neq c$  and positive probability for  $b = c$ . In particular, this chain can be characterized at considered the transition probabilities for  $(a, b) \rightarrow (b, c)$  by considering the probability transition matrix of  $X$  corresponding to  $b \rightarrow c$ . With this, we arrive at the transition matrix as follows

$$\mathbf{P} = \begin{matrix} & \begin{matrix} (0,0) & (0,1) & (1,0) & (1,1) \end{matrix} \\ \begin{matrix} (0,0) \\ (0,1) \\ (1,0) \\ (1,1) \end{matrix} & \begin{pmatrix} \alpha & 1-\alpha & 0 & 0 \\ 0 & 0 & 1-\beta & \beta \\ \alpha & 1-\alpha & 0 & 1 \\ 0 & 0 & 1-\beta & \beta \end{pmatrix} \end{matrix}$$

**Problem 3.3.8** Suppose there are two urns,  $A$  and  $B$ , and at time  $t$  the number of balls in  $A$  is  $k$  and the number of balls in  $B$  is  $N - k$ . First an urn is selected with probability  $p_A = \frac{k}{N}$  and  $p_B = 1 - \frac{k}{N}$ . Then a ball is selected from either  $A$  with probability  $p$  or  $B$  with probability  $q$  and placed in the urn selected above. Let  $Y_n$  be the number of balls in urn  $A$  at time  $n$ . Then using the above we have that  $Y_{n+1} = Y_n + 1$  if urn  $A$  is chosen and we select a ball from urn  $B$ .  $Y_{n+1} = Y_n - 1$  if we select urn  $B$  then a ball from urn  $A$ . Finally,  $Y_{n+1} = Y_n$  if we select from urn  $A$  and a ball from  $A$  or select  $B$  and a ball from  $B$ . This corresponds to the following transition matrix

$$P_{ij} = \begin{cases} \frac{i}{N}q & j = i + 1 \\ \frac{i}{N}p + (1 - \frac{Y_n}{N})q & j = i \\ (1 - \frac{i}{N})p & j = i - 1 \\ 0 & |i - j| \geq 2 \end{cases}$$

**Problem 3.4.6** First, define  $T = \min\{n : X_n = 4\}$  and let  $v_i = \mathbb{E}[T|X_0 = i]$ . Our goal to find  $v_0$ . First note that trivially we have  $v_4 = 0$ . Moreover, we have

$$\begin{aligned} v_0 &= \mathbb{E}[T|X_0 = 0] = 1 + \sum_{j=0}^4 \mathbb{E}[T|X_1 = j]P_{0j} = 1 + \sum_{j=0}^4 v_j P_{0j} = qv_0 + pv_1 + 1 \\ v_1 &= \mathbb{E}[T|X_0 = 1] = 1 + \sum_{j=0}^4 \mathbb{E}[T|X_1 = j]P_{1j} = 1 + \sum_{j=0}^4 v_j P_{1j} = qv_0 + pv_2 + 1 \\ v_2 &= \mathbb{E}[T|X_0 = 2] = 1 + \sum_{j=0}^4 \mathbb{E}[T|X_1 = j]P_{2j} = 1 + \sum_{j=0}^4 v_j P_{2j} = qv_0 + pv_3 + 1 \\ v_3 &= \mathbb{E}[T|X_0 = 3] = 1 + \sum_{j=0}^4 \mathbb{E}[T|X_1 = j]P_{3j} = 1 + \sum_{j=0}^4 v_j P_{3j} = qv_0 + 1 \end{aligned}$$

Now seeing  $p = 1 - q$ ,  $v_0 = qv_0 + pv_1 + 1$  implies  $v_0 = v_1 + \frac{1}{p}$ . Using this relation recursively, we have

$$\begin{aligned} v_0 &= v_1 + \frac{1}{p} = qv_0 + pv_2 + 1 + \frac{1}{p} = v_2 + \frac{1}{p} + \frac{1}{p^2} = qv_0 + pv_3 + 1 + \frac{1}{p} + \frac{1}{p^2} \\ &= v_3 + \frac{1}{p} + \frac{1}{p^2} + \frac{1}{p^3} = qv_0 + 1 + \frac{1}{p} + \frac{1}{p^2} + \frac{1}{p^3} = \frac{1}{p} + \frac{1}{p^2} + \frac{1}{p^3} + \frac{1}{p^4} \\ &= \sum_{k=1}^4 \frac{1}{p^k} \end{aligned}$$

**Problem 3.4.17** Let  $T = \min\{n : X_n = 2\}$  be the time of the first failure and define  $\phi_i(s) := \mathbb{E}[s^T|X_0 = i]$ . Given that our system starts in state 0 (fully operational), we look to evaluate the probability generating function  $\phi_0(s)$  for  $0 < s < 1$ . First note, that trivially  $X_0 = 2$  implies  $T = 0$  and  $\phi_2(s) = 1$ . Using this, we can use first step analysis to write

$$\begin{aligned} \phi_0(s) &= \mathbb{E}[s^T|X_0 = 0] = \sum_{j=0}^2 \mathbb{E}[s^{T+1}|X_1 = j]P_{0j} = s \sum_{j=0}^2 \phi_j(s)P_{0j} = s[P_{00}\phi_0(s) + P_{01}\phi_1(s)] \\ \phi_1(s) &= \mathbb{E}[s^T|X_0 = 1] = \sum_{j=0}^2 \mathbb{E}[s^{T+1}|X_1 = j]P_{1j} = s \sum_{j=0}^2 \phi_j(s)P_{1j} = s[P_{11}\phi_1(s) + P_{12}] \end{aligned}$$

Solving the second equation for  $\phi_1(s)$ , we have  $\phi_1(s) = \frac{sP_{12}}{1-sP_{11}}$ . Using this in the first

equation, we have

$$\begin{aligned}\phi_0(s) &= sP_{00}\phi_0(s) + sP_{01}\left(\frac{sP_{12}}{1 - sP_{11}}\right) \\ \phi_0(s) &= \frac{s^2P_{01}P_{12}}{(1 - sP_{11})(1 - sP_{00})}\end{aligned}$$

Nowing using the values from the transition probability matrix, we arrive at the following form.

$$\phi_0(s) = \frac{.12s^2}{(1 - .6s)(1 - .7s)}$$