

Model Selection & Hypothesis Testing

Recall we want to test $H_0: \theta \in \Theta_0$ $H_1: \theta \in \Theta_1$

and want to infer $I(\theta \in \Theta_0)$

We can define a decision $\phi(y) = \underset{\phi \in \{0,1\}}{\operatorname{argmin}} \mathbb{E}_{\theta|y}[L(\theta, \phi)]$

We considered the loss type I: δ type II: 1 and found

$$\phi(y) = I\left[\frac{P(\theta \in \Theta_0|y)}{P(\theta \in \Theta_1|y)} \leq \delta\right] = I\left[P(\theta \in \Theta_0|y) \geq \frac{1}{1+\delta}\right]$$

If we have a point hypothesis s.t. $P(\theta \in \Theta_0) = 0$ $P(\theta \in \Theta_1) = 0$ and θ is cont. then this setup can breakdown.

One way to accommodate simple nulls then use a hierarchical model to test

$P(H_0) = P(\theta \in \Theta_0)$ and $P(H_1) = 1 - P(H_0)$ and set up the prior

$$P(\theta) = P(H_0)P(\theta | \theta \in \Theta_0) + P(H_1)P(\theta | \theta \in \Theta_1)$$

$$\text{This way, } \frac{P(\theta \in \Theta_1|y)}{P(\theta \in \Theta_0|y)} = \frac{P(\theta \in \Theta_1, y)}{P(\theta \in \Theta_0, y)} = \frac{\int P(y|\theta)P(\theta)I(\theta \in \Theta_1) d\theta}{\int P(y|\theta)P(\theta)I(\theta \in \Theta_0) d\theta}$$

$$= \frac{\int P(y|\theta)P(\theta|\theta \in \Theta_1)I(\theta \in \Theta_1) d\theta P(H_1)}{\int P(y|\theta)P(\theta|\theta \in \Theta_0)I(\theta \in \Theta_0) d\theta P(H_0)} = \underbrace{\frac{P(y|\theta \in \Theta_1)}{P(y|\theta \in \Theta_0)}}_{\substack{\text{Bayes Factor} \\ BF_{10} \\ \text{one to zero}}} \times \frac{P(H_1)}{P(H_0)}$$

$$BF_{10}(y) = \frac{\frac{P(\theta \in \Theta_1 | y)}{P(\theta \in \Theta_0 | y)}}{\frac{P(\theta \in \Theta_1)}{P(\theta \in \Theta_0)}} \quad \text{Post-Prior Odds Ratio.}$$

A useful table by Kass/Raftery

BF	Evidence
1-3	Barely motion
3-10	positive
10-150	strong
>150	v. strong

Variable Selection

Suppose we regress $Y \sim X$, where $X \in \mathbb{R}^{n \times p}$. Want to compare all possible models from including or excluding predictors. 2^p models

We indicate which model we are considering with $\theta \in \{0, 1\}^p$

Want to make inferences on $\hat{\theta} = \arg \min_{\theta \in \{0, 1\}^p} \mathbb{E}_{\theta|y} [L(\theta, \phi)]$

two common choices:

$$(i) L_m(\theta, \phi) = I(\theta \neq \phi) = 1 - \prod_{j=1}^p I(\theta_j = \phi_j)$$

$$\phi_m(y) = \arg \min_{\phi \in \{0, 1\}^p} \mathbb{E}_{\theta|y} [I(\theta \neq \phi)] = \arg \min_{\phi} [1 - P(\theta = \phi | y)]$$

$$= \arg \max_{\phi} P(\theta = \phi | y) = \hat{\theta}_{\text{map}}$$

$$(ii) L_h(\theta, \phi) = \sum_{j=1}^p \underbrace{L(\theta_j, \phi_j)}_{\substack{L: \\ 0 \mid \theta \\ 1 \mid \phi}} = \sum_{j=1}^p [\tau(1-\phi_j)\theta_j + \phi_j(1-\theta_j)]$$

$$\phi_H(y) = \arg \min_{\phi} \mathbb{E}_{\theta|y} \left[\sum_{j=1}^p L(\theta_j, \phi_j) \right] = \arg \min_{\phi} \sum_{j=1}^p \mathbb{E}_{\theta_j|y} [L(\theta_j, \phi_j)]$$

$$\Rightarrow (\phi_H(y))_{j=1}^p = \arg \min_{\phi_j} \mathbb{E}_{\theta_j|y} [L(\theta_j, \phi_j)] = \mathbb{I} \left[\underbrace{\mathbb{E}[\theta_j|y]}_{\substack{P(\theta_j=1|y) \\ \text{post. predictive}}} \geq \frac{1}{1+\delta} \right]$$

"posterior prob. of inclusion"

Remark: Hard to compute.

ΣX: For simplicity assume $\gamma | \beta, \sigma^2, \theta \sim N(\underbrace{X_0 \beta_0}_{X[\theta=1]}, \underbrace{\sigma^2 I_n}_{\beta[\theta=1]})$

$$\beta_0 | \sigma^2 \sim N(\beta_{0,\theta}, \sigma^2 \Sigma_\theta) \quad \beta_{0,\theta} = 0, \quad \Sigma_\theta = g(X_0^T X_0)^{-1}$$

— typically $g=n$.

$$\sigma^2 \sim \text{Inv} \chi^2(\nu, \tau^2)$$

For prior on θ by $\theta_j \stackrel{\text{iid}}{\sim} \text{Bern}(\alpha)$

From the model

$$\gamma | \sigma^2, \theta \sim N(X_0 \beta_{0,\theta}, \sigma^2 (I_n + X_0 \Sigma_\theta X_0^T))$$

$$\gamma | \theta \sim \text{tr}(X \beta_{0,\theta}, \tau^2 (I_n + X_0 \Sigma_\theta X_0^T))$$

$$P(\theta|y) = \frac{P(y|\theta) P(\theta)}{\sum_{\tilde{\theta} \in \{0,1\}^p} P(y|\tilde{\theta}) P(\tilde{\theta})}$$