

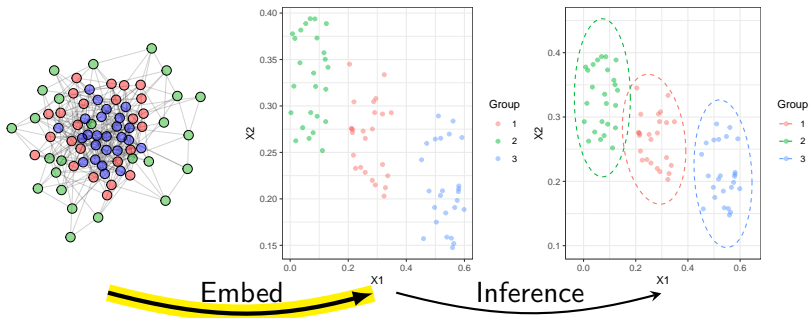
Joint Spectral Embeddings of Random Dot Product Graphs

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Network Embeddings

- Goal: Obtain a low dimensional representation of the network.
- Reason: Address network analysis questions with methods from multivariate statistics and machine learning.



Motivation - Multiplex Networks

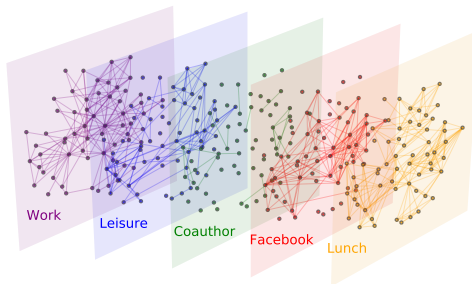
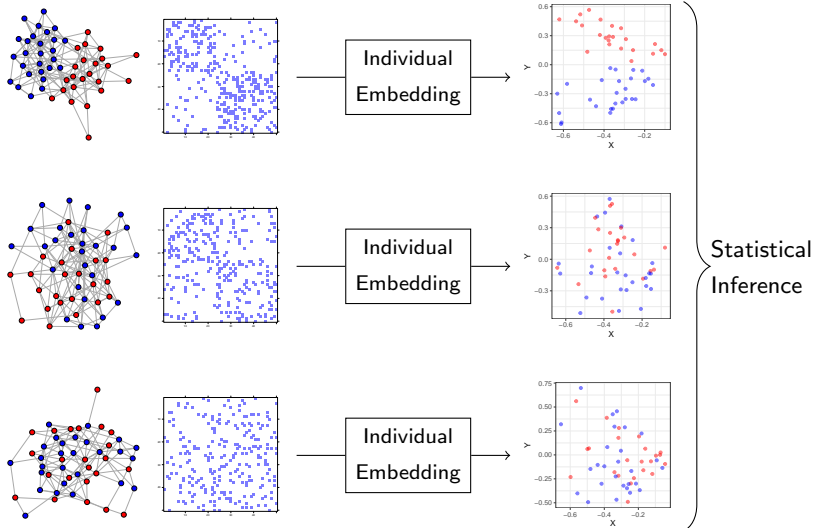


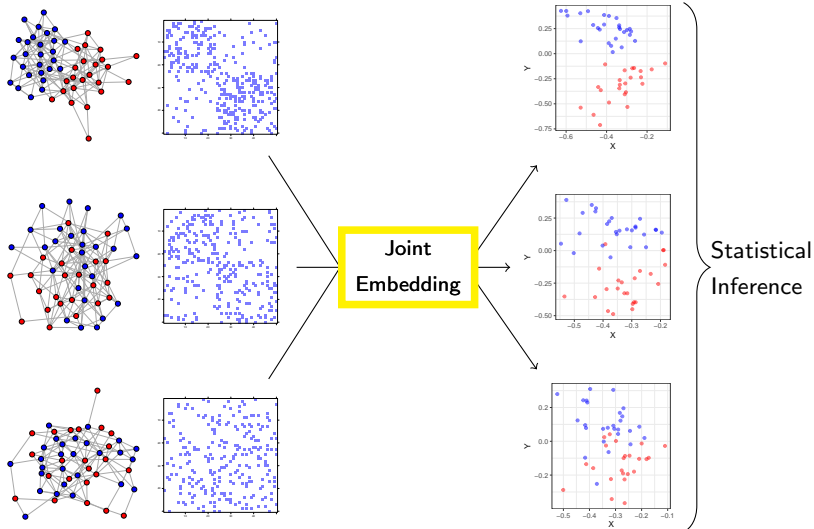
Figure: Multiplex Network of Aarhus Computer Science Department. Vertices are members of the department and each layer encodes a different type of interaction.(Magnani, Micenkova, and Rossi 2013).

- Multiplex networks - set of networks over a common vertex set.
- Adjacency matrices $A^{(g)} \in [0, 1]^{n \times n}$ often share structure.

Multiple Network Embeddings



Multiple Network Embeddings



Multiplex Random Dot Product Graph

- Consider m graphs over a common vertex set \mathcal{V} of size n .
- We extend the Random Dot Product Graph (Young and Scheinerman 2007) to multiplex networks.

Multiplex Random Dot Product Graph

- For $g \in [m]$, suppose that F_g is a probability distribution on \mathbb{R}^{d_g} such that for all $x, y \in \text{supp}(F_g)$, $x^T y \in [0, 1]$.
- Let $\{X_i^{(g)}\}_{i=1}^n \stackrel{i.i.d.}{\sim} F_g$ and $X^{(g)} = [X_1^{(g)}, \dots, X_n^{(g)}]^T \in \mathbb{R}^{n \times d_g}$ be the *latent positions* for graph $g \in [m]$.
- We say $\{(A^{(g)}, X^{(g)})\}_{g=1}^m \sim \text{MRDPG}(F, n)$ iff $\{A_{ij}^{(g)}\}_{i < j, g \in [m]}$ are conditionally independent with

$$\mathbb{P}(A_{ij}^{(g)} = 1 | X^{(g)}) = \langle X_i^{(g)}, X_j^{(g)} \rangle$$

Submodels & Past Work

- Each vertex in each layer is described by a latent position $X_i^{(g)}$.
- In essence, $A^{(g)} | X^{(g)} \sim \text{Bern}(P^{(g)})$ where $P^{(g)} \equiv X^{(g)} X^{(g)T}$.

Model	Latent Positions $X^{(g)}$	Source
Joint RDPG	X	Levin et al. (2017)
Multiple RDPG	$U\sqrt{D^{(g)}}$	Nielsen and Witten (2018)
Eigen-Scaling RDPG	$X\sqrt{D^{(g)}}$	Draves and Sussman (2020+)
MREG	$V\sqrt{D^{(g)}}$	Wang et al. (2017)
COSIE	$U\sqrt{R^{(g)}}$	Arroyo et al. (2019)
Multilayer RDPG	$X\sqrt{R^{(g)}}$	Jones and Rubin-Delanchy (2020)

Table: Global structure: U is matrix with orthogonal columns, V has columns with unit Euclidean norm, and X is a latent position matrix. Layer variation: matrix square root of diagonal matrix $D^{(g)}$ or symmetric matrix $R^{(g)}$.

- Key Point:** *Network Embeddings are latent position estimators.*

Adjacency Spectral Embedding

- Ignore joint structure and embed $A^{(g)}$ individually.

Adjacency Spectral Embedding (Sussman et al. 2012)

Let $A^{(g)}$ have eigendecomposition

$$A^{(g)} = [U_{A^{(g)}} | \tilde{U}_{A^{(g)}}][S_{A^{(g)}} \oplus \tilde{S}_{A^{(g)}}][U_{A^{(g)}} | \tilde{U}_{A^{(g)}}]^T$$

where $U_{A^{(g)}} \in \mathbb{R}^{n \times d}$ and $S_{A^{(g)}} \in \mathbb{R}^{d \times d}$ contains the top d eigenvalues of $A^{(g)}$. Then the ASE of $A^{(g)}$ is defined by $\text{ASE}(A^{(g)}, d) = U_{A^{(g)}} S_{A^{(g)}}^{1/2}$.

- The i -th row of $\text{ASE}(A^{(g)}, d_g)$ is consistent for $X_i^{(g)}$ up to Gaussian error for each $g \in [m]$ and $i \in [n]$ (Athreya et al. 2017).

Multiple Adjacency Spectral Embedding

- Embed $A^{(g)}$ individually then map to a common space.

Multiple Adjacency Spectral Embedding (Arroyo et al. 2019)

Let $\hat{X}_{ASE}^{(g)} = ASE(A^{(g)}, d_g)$ and define $\hat{X}_{ASE} = [\hat{X}_{ASE}^{(1)} \hat{X}_{ASE}^{(2)} \dots \hat{X}_{ASE}^{(m)}]$. Let \hat{X}_{ASE} has singular value decomposition

$$\hat{X}_{ASE} = \hat{U} \hat{\Sigma} \hat{V}^T + \hat{U}_\perp \hat{\Sigma}_\perp \hat{V}_\perp^T$$

where $\hat{\Sigma} \in \mathbb{R}^{d \times d}$. Partition $\hat{V} = [\hat{V}^{(1)T} \hat{V}^{(2)T} \dots \hat{V}^{(m)T}]^T$. Then for all

$$g \in [m], \quad \hat{X}_{MASE}^{(g)} = [MASE(\{A^{(g)}\}_{g=1}^m, d)]_g = \hat{U} \hat{\Sigma} \hat{V}^{(g)T}.$$

- Under a submodel of the MRDPG, \hat{U} is consistent estimator for global model parameter (Arroyo et al. 2019).

Omnibus Embedding

- Simultaneously embed $\{A^{(g)}\}_{g=1}^m$ into a common space.

Omnibus Embedding (Levin et al. 2017)

The *Omnibus matrix* is given by

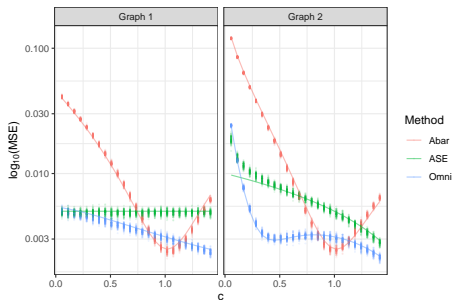
$$\tilde{A} = \begin{bmatrix} A^{(1)} & \frac{1}{2}[A^{(1)} + A^{(2)}] & \dots & \frac{1}{2}[A^{(1)} + A^{(m)}] \\ \frac{1}{2}[A^{(2)} + A^{(1)}] & A^{(2)} & \dots & \frac{1}{2}[A^{(2)} + A^{(m)}] \\ \vdots & \vdots & \ddots & \vdots \\ \frac{1}{2}[A^{(m)} + A^{(1)}] & \frac{1}{2}[A^{(m)} + A^{(2)}] & \dots & A^{(m)} \end{bmatrix} \in \mathbb{R}^{nm \times nm}.$$

The omnibus embedding of graph $g \in [m]$ is $\hat{X}_{\text{OMNI}}^{(g)} = [\text{ASE}(\tilde{A}, d)]_g$ where $[\cdot]_g$ denotes the g -th, $n \times d$ block.

- Under a submodel of the MRDPG, the i -th row of $\hat{X}_{\text{OMNI}}^{(g)}$ is a biased estimator of $X_i^{(g)}$ up to Gaussian error (Draves and Sussman 2020+).

Bias Variance Tradeoffs in Joint Spectral Embeddings

- Assess embedding techniques as latent positions estimators.
- Draves and Sussman 2020+ analyzed bias-variance tradeoff and ramifications of utilizing biased latent position estimates in subsequent inference.



- Unbiased latent position estimation *not necessarily* needed to achieve accurate statistical inference.

Network Embeddings as Dim. Reduction Techniques

- If biased latent position estimates can still lead to accurate inference, how should we assess embedding techniques?
- Goal: Study network embeddings *as their own statistical method*.
- **Key Point:** *Network Embeddings are latent position estimators and dimensionality reduction techniques.*
- Observation: ASE implicitly imitates *uncentered* principal component analysis (UPCA) on the matrix $A^{(g)}$.
- Question: Can we interpret MASE and OMNI as dimensionality reduction techniques?

Multiple Population Dimensionality Reduction

- Focus on dimensionality reduction of the matrix

$$X = [X^{(1)} X^{(2)} \dots X^{(m)}] \in \mathbb{R}^{n \times (\sum_{g=1}^m d_g)}.$$

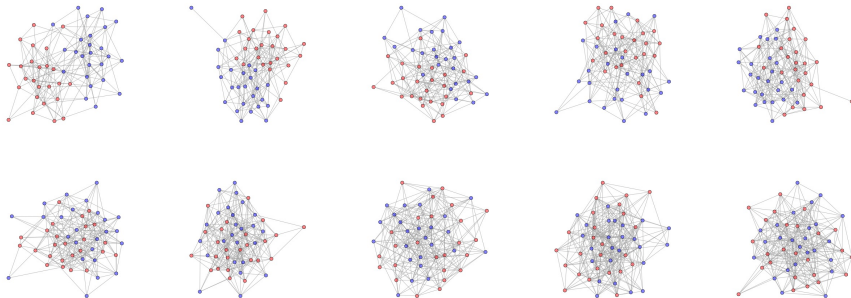
Uncentered PCA of X

- Suppose $r = \text{rank}(X)$ and let X have singular value decomposition $X = U\Sigma V^T + U_{\perp}\Sigma_{\perp}V_{\perp}^T$ where $\Sigma \in \mathbb{R}^{d \times d}$ for $d \in [r]$.
- The *uncentered principal components* are the rows of $\tilde{X} = U\Sigma$.
- Partition $V = [V^{(1)T} \dots V^{(m)T}]^T$ so that $V^{(g)} \in \mathbb{R}^{d_g \times d}$. Then $\tilde{X}^{(g)} = U\Sigma V^{(g)T}$ is the total least squares fit of the latent positions for graph $g \in [m]$.
- UPCA of X is identical to performing total least squares (TLS) regression on the latent positions $\{X^{(g)}\}_{g=1}^m$.

Example: Hierarchical Stochastic Block Model

- Suppose that $\{A^{(g)}\}_{g=1}^{25}$ are SBMs with 2 groups of even size, $n = 25$.
- Community assignment fixed and block probabilities vary.
- For $b_g \in [0, 0.25]$, model parameters and latent positions are

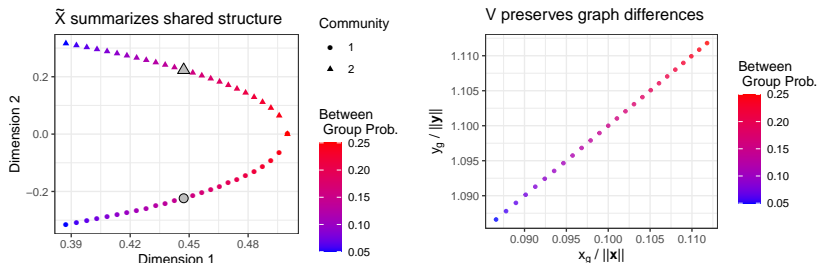
$$Z = \begin{bmatrix} 1_n & 0_n \\ 0_n & 1_n \end{bmatrix} \quad B^{(g)} = \begin{bmatrix} 0.25 & b_g \\ b_g & 0.25 \end{bmatrix} \quad L^{(g)} = Z \begin{bmatrix} x_g & -y_g \\ x_g & y_g \end{bmatrix}.$$



Example: Hierarchical Stochastic Block Model

- Let $\mathbf{x} = (x_1 x_2 \dots x_m)^T \in \mathbb{R}^{m \times 1}$ and $\mathbf{y} = (y_1 y_2 \dots y_m)^T \in \mathbb{R}^{m \times 1}$.
- Let $\bar{\mathbf{B}} = m^{-1} \sum_{g=1}^m \mathbf{B}^{(g)}$ and let $\bar{\mathbf{L}} = \text{ASE}(\bar{\mathbf{B}}, 2)$
- Then the UPCA estimates are $\tilde{\mathbf{X}}$ and the $\{\mathbf{V}^{(g)}\}_{g=1}^m$ as

$$\tilde{\mathbf{X}} = \sqrt{m} \mathbf{Z} \bar{\mathbf{L}} \quad \mathbf{V}^{(g)} = \text{diag} \left(\frac{x_g}{\|\mathbf{x}\|}, \frac{y_g}{\|\mathbf{y}\|} \right).$$



Preliminary Results

- \tilde{X} captures layer similarities by implicitly embedding \bar{P} into \mathbb{R}^d .

Lemma - \tilde{X} Summarizes Shared Structure

There exists $W \in \mathcal{O}_d$ such that $\tilde{X}W = m^{1/2}\text{ASE}(\bar{P}, d)$.

- $V^{(g)} : \mathbb{R}^d \rightarrow \mathbb{R}^{d_g}$ maps the rows of \tilde{X} to the TLS fit of the latent positions $\{X^{(g)}\}_{g=1}^m$.

Lemma - $V^{(g)}$ describes graph differences

Let $\{\tilde{X}^{(g)}\}_{g=1}^m$ be the TLS fit of the latent positions $\{X^{(g)}\}_{g=1}^m$. Then $V^{(g)}\tilde{x}_i = \tilde{x}_i^{(g)}$.

- UPCA of X succinctly segregates the common structure shared across layers and layer specific variation.

Preliminary Results: MASE & OMNI

- Let $X_{\text{MASE}}^{(g)}$ and $X_{\text{OMNI}}^{(g)}$ be the population equivalent of $\hat{X}_{\text{MASE}}^{(g)}$ and $\hat{X}_{\text{OMNI}}^{(g)}$ by replacing $\{A^{(g)}\}_{g=1}^m$ with $\{P^{(g)}\}_{g=1}^m$.

Lemma - MASE

There exists $W \in \mathcal{O}_{d_g}$ such that $X_{\text{MASE}}^{(g)} W = \tilde{X}^{(g)} = \tilde{X} V^{(g)T}$.

Lemma - OMNI

Let $\tilde{Y} = m^{1/2} \text{ASE}(\bar{P}, r)$ and \tilde{V} be the omnibus matrix of $\{V^{(g)T} V^{(g)}\}_{g=1}^m$. Let $B = \text{OMNI}(\tilde{V}, d)$ and partition $B^T = [B^{(1)} B^{(2)} \dots B^{(m)}]$ so that $B_g \in \mathbb{R}^{r \times d}$. Then there exists $W \in \mathcal{O}_d$ such that $X_{\text{OMNI}}^{(g)} W = \tilde{Y} B^{(g)T}$.

- MASE implicitly imitates rank- d UPCA of X .
- OMNI implicitly imitates full rank UPCA of X then maps the rows into \mathbb{R}^d .

Simulation Design

Consider SBM with $K = 3$ groups of size $n_k = 100$ so that $Z = I_{3 \times 3} \otimes \mathbf{1}_{n_k}$.

$$B = \begin{bmatrix} 0.25 & 0.05 & 0.05 \\ 0.05 & 0.25 & 0.05 \\ 0.05 & 0.05 & 0.25 \end{bmatrix} \quad L_B = \begin{bmatrix} 0.34 & -0.37 & 0 \\ 0.34 & 0.18 & -0.32 \\ 0.34 & 0.18 & 0.32 \end{bmatrix} \quad \Delta = \begin{bmatrix} 0 & 0 & 0 \\ 0 & 0 & 0.2 \\ 0 & 0.2 & 0 \end{bmatrix}$$

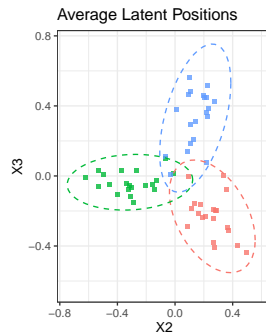
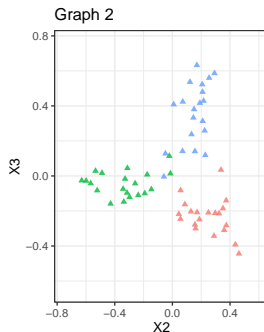
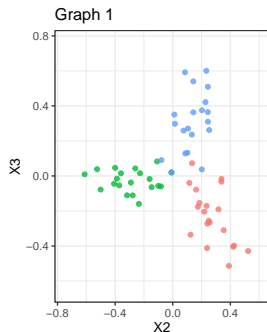
Anomalous Vertices

- 1 Set $X^{(1)} = X^{(2)} = ZL_B$
- 2 $A^{(1)} \sim \text{Bern}(X^{(1)}X^{(1)T})$
- 3 For $\alpha \in [0, 0.1]$, replace $(100 \times \alpha)\%$ rows of $X^{(2)}$ sample from $\text{Dir}(1, 1, 1)$
- 4 $A_{\alpha}^{(2)} \sim \text{Bern}(X_{\alpha}^{(2)}X_{\alpha}^{(2)T})$

Hierarchical SBM

- 1 Set $X^{(1)} = X^{(2)} = ZL_B$
- 2 $A^{(1)} \sim \text{Bern}(X^{(1)}X^{(1)T})$
- 3 For $\lambda \in [0, 0.25]$, let $L_{\lambda} = \text{ASE}(B + \lambda\Delta, 3)$ and set $X_{\lambda}^{(2)} = ZL_{\lambda}$
- 4 $A_{\lambda}^{(2)} \sim \text{Bern}(X_{\lambda}^{(2)}X_{\lambda}^{(2)T})$

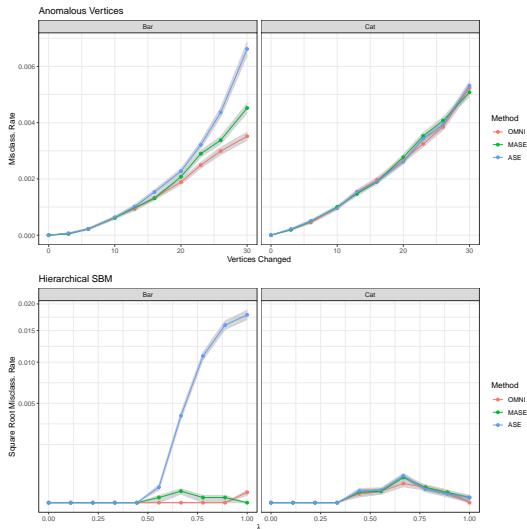
Multiplex Network Analysis: Community Detection



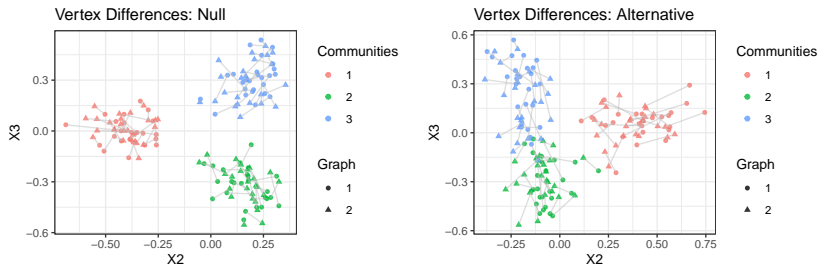
Community Detection

- 1 Embed $(\hat{X}^{(1)}, \hat{X}^{(2)}) \leftarrow \text{EMBED}(A^{(1)}, A^{(2)}, d)$
- 2 Apply clustering algorithms to rows of $[\hat{X}^{(1)} \hat{X}^{(2)}]$ or $\bar{X} = 2^{-1}(\hat{X}^{(1)} + \hat{X}^{(2)})$.

Inference Ramifications: Community Detection



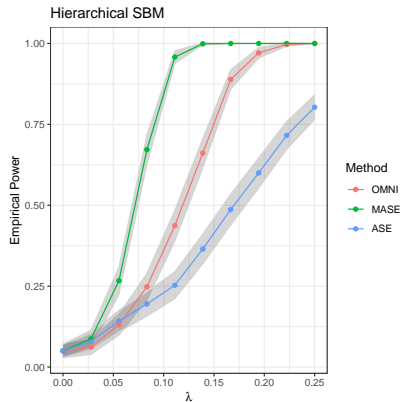
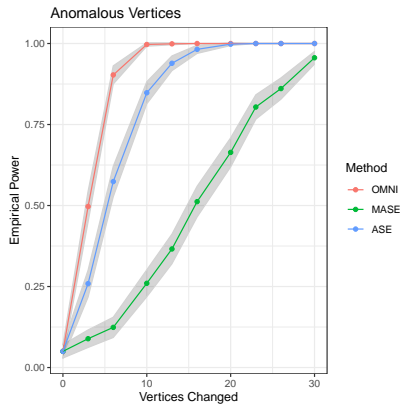
Multiplex Network Analysis: Hypothesis Testing



Hypothesis Testing

- 1 Embed $(\hat{X}^{(1)}, \hat{X}^{(2)}) \leftarrow \text{EMBED}(A^{(1)}, A^{(2)}, d)$
- 2 Set test statistic $T = \|\hat{X}^{(1)} - \hat{X}^{(2)}\|_F$
- 3 Bootstrap reference distribution $\{T_b : b \in [B]\}$ (Levin et al. 2017).
- 4 Return $p \leftarrow |\{b \in [B] : T_b \geq T\}|/B$

Inference Ramifications: Hypothesis Testing



Conclusion

- By pooling information across networks, joint spectral embeddings can learn more representative node embeddings.
- MASE and OMNI complete multiple population dimensionality reduction on the latent positions.
- Differences between these dimensionality reduction techniques make MASE and OMNI node embeddings more/less useful in certain statistical applications.

Questions?

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