1. (a)

$$\begin{split} E[Y_{(1)}] &= \int_0^\infty (y * \frac{d}{dy} F_{Y_{(1)}}(y)) dy \\ &= \int_0^\infty (\frac{-ny}{\theta} e^{\frac{-ny}{\theta}}) dy \\ &= y e^{\frac{-ny}{\theta}}]_0^\infty - \int_0^\infty (e^{\frac{-ny}{\theta}}) dy \\ &= 0 - 0 + \frac{-\theta}{n} e^{\frac{-ny}{\theta}}]_0^\infty \\ &= 0 - \frac{-\theta}{n} \\ &= \frac{\theta}{n} \end{split}$$

Unbiased after a sample size of one.

(b) Multiply by the sample size to receive an unbiased estimator for θ .

$$E[Y_{(1)}] = \int_0^\infty (n * y * \frac{d}{dy} F_{Y_{(1)}}(y)) dy$$
$$= \int_0^\infty (\frac{-n^2 y}{\theta} e^{\frac{-n y}{\theta}}) dy$$
$$= n(0 + 0 - 0 - \frac{-\theta}{n})$$
$$= \theta$$

2. If we show:

$$\lim_{n \to \infty} Pr(|\hat{p} - p| > \epsilon) = 0, \forall \epsilon > 0$$

then \hat{p} is consistent for p.

 \bar{X} is consistent for μ , therefore we can write;

$$\lim_{n \to \infty} Pr(|n\hat{p} - np| > \epsilon) = 0, \forall \epsilon > 0$$
$$\lim_{n \to \infty} Pr(|\hat{p} - p| |n| > \epsilon) = 0, \forall \epsilon > 0$$

n is always a natural number, therefore;

$$\lim_{n \to \infty} Pr(|\hat{p} - p| \, n > \epsilon) = 0, \forall \epsilon > 0$$

$$\lim_{n \to \infty} Pr(|\hat{p} - p| > \frac{\epsilon}{n}) = 0, \forall \epsilon > 0$$

Then without loss of generality for ϵ ;

$$\lim_{n \to \infty} Pr(|\hat{p} - p| > \epsilon) = 0, \forall \epsilon > 0$$

 \hat{p} is consistent for p, then.

3. (a)

$$\begin{split} \int_0^1 (y\theta y^{\theta-1}) dy &= E[Y] \\ \int_0^1 (\theta y^{\theta}) dy &= E[Y] \\ \int_0^1 (\theta y^{\theta}) dy &= E[Y] \\ \frac{\theta y^{\theta+1}}{\theta+1}]_0^1 &= E[Y] \\ \frac{\theta}{\theta+1} &= E[Y] \end{split}$$

- (b) \bar{X} is consistent for μ by WLoLN So \bar{Y} is consistent for $\frac{\theta}{\theta+1}$
- 4. (a) It looks normal with the proper symmetry and the fifteen percent variance.
 - (b)

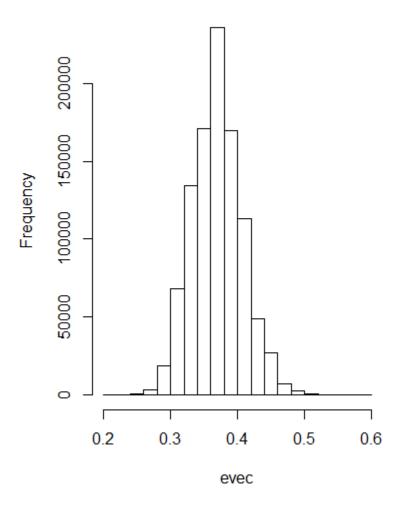
$$\sqrt{n}(\bar{X} - \lambda) \to N(0, \sigma^2(x))$$

$$\sqrt{100}(\bar{X} - \lambda) \to N(1, 1)$$

$$(\bar{X} - \lambda) \to N(1, \frac{1}{100})$$

- (c) $\bar{X}=1.00004195$ $\bar{X}-X\approx 0$ Var[X]=.009978These match.
- 5. (a) It sure looks normal to me. It's even more symmetric than the last one with a similar quartile range.

Histogram of evec



(b)
$$\begin{split} \sqrt{n}(\bar{X}-\lambda) &\to N(e^{-1},\sigma^2(x)\frac{d}{dx}(e^{-x})^2) \\ (\bar{X}-\lambda) &\to N(e^{-1},\frac{1}{100}(-e^{-1})^2) \\ (\bar{X}-\lambda) &\to N(e^{-1},\frac{e^{-2}}{100}) \end{split}$$

(c)
$$\bar{X} = .369698$$

 $\bar{X} - \lambda \approx 0$

$$\begin{split} Var[X] &= .001357471 \\ \text{These match.} \end{split}$$

