

Stochastic Processes: Assignment 1

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Importing libraries

```
#> Package:  markovchain
#> Version:  0.8.5-2
#> Date:     2020-09-07
#> BugReport: https://github.com/spedygiorgio/markovchain/issues
#>
#> Attaching package: 'dplyr'
#> The following objects are masked from 'package:stats':
#>
#>     filter, lag
#> The following objects are masked from 'package:base':
#>
#>     intersect, setdiff, setequal, union
```

Problem 1

```
#> [1] "n bnq ndrsw awornoc aslnxs rzs dowfr ksfrswc ljicrwq rj nirzjwoms n ljwjengowif gnllocs aworofz"
```

Problem 2

(a)

Let $N(t)$ be the number of cars arriving at a parking lot by time t , according to the proposed scenario, we can model $N(t)$ as a non-homogenous Poisson process. Such process has almost the same process as any other Poisson process, however, its rate is a function of time.

$N(t), t \in [0, \infty)$ is the non-homogenous Poisson process with rate $\lambda(t)$ where:

- $N(0) = 0$
- $N(t)$ has independent increments

We define 8:00 as $t = 0$ with the following integrable function and each unit of t equals to 1 hour:

$$\lambda(t) = \begin{cases} 100 & 0 \leq t \leq \frac{1}{2} \\ 600t - 200 & \frac{1}{2} < t \leq \frac{3}{4} \\ 400t - 50 & \frac{3}{4} < t \leq 1 \\ -500t + 850 & 1 < t \leq 1.5 \end{cases}$$

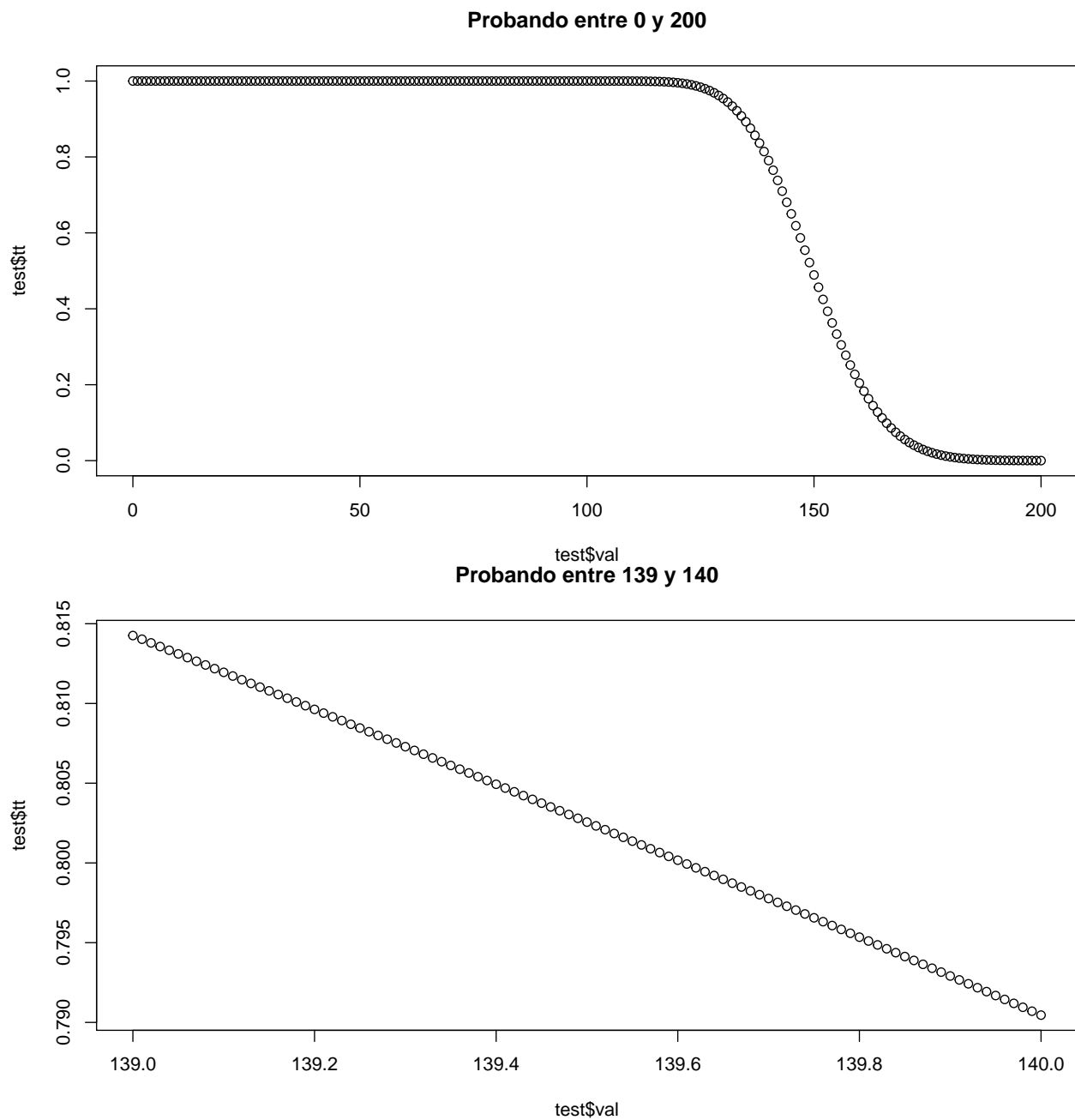
So,

$$E[N(t)] = \begin{cases} \int_0^t 100 dt = 100t & 0 \leq t \leq \frac{1}{2} \\ \int_{\frac{1}{2}}^t 600t - 200 dt + 50 = 300(t^2 - \frac{1}{4}) - 200(t - \frac{1}{2}) + 50 & \frac{1}{2} < t \leq \frac{3}{4} \\ \int_{\frac{3}{4}}^t 400t - 50 dt + 93.75 = 25(8t^2 - 2t - 3) + 93.75 & \frac{3}{4} < t \leq 1 \\ \int_1^t -500t + 850 dt + 168.75 = -50(5t^2 - 17t + 12) + 168.75 & 1 < t \leq 1.5 \end{cases}$$

Given that there is a limit of 150 vehicles:

$$E[N(t)] = \begin{cases} 100t & 0 \leq t \leq \frac{1}{2} \\ 300(t^2 - \frac{1}{4}) - 200(t - \frac{1}{2}) + 50 & \frac{1}{2} < t \leq \frac{3}{4} \\ 25(8t^2 - 2t - 3) + 93.75 & \frac{3}{4} < t < 0.94468 \\ 150 & t \geq 0.94468 \end{cases}$$

(b)



Luego de hacer las pruebas para $\lambda(t)$ obtenemos lo siguiente:

```
lambda = 139.6  
t = 0.91232  
# 8:44 AM
```

Por lo que $t = 0.91232$ horas.

(c)

The following function simulates a non-homogenous poisson process from a homogenous poisson process:

```

non_hom_poisson <- function(fun,l,a,b,start=0) {
  # This function generates a non-homogenous poisson
  # process from a homogenous poisson process
  # PARAMS:
  # fun:    if the non-homogenous poisson process has
  #         multiple functions per time subinterval
  #         this parameters represents such function
  # l:      lambda for the homogenous poisson process
  # a:      lower bound for the time subinterval
  # b:      upper bound for the time subinterval
  # start:  this parameter is used to keep track of
  #         the process count.

  # We generate the homogenous poisson process
  # arrival times
  val <- rpois(1,l*(b-a))
  intervals <- (b-a) * sort(runif(val)) + a

  # Non-homogenous poisson process
  evs <- length(intervals) # lenght of arrival times
  nh_val <- 0 + start # start of the event count
  nh_intervals <- c() # arrival times for the NHPP
  for (i in 1:evs) {
    if (runif(1) < fun(intervals[i])/l) {
      # only including intervals from the HPP which
      # match with fun(intervals[i])/l probability
      nh_intervals <- c(nh_intervals, intervals[i])
      nh_val <- nh_val+1 # adding one to the event count
    }
  }
  nh_events <- seq(1+start,nh_val,1) # events since the previous group
  return(list(arrival_times=nh_intervals, events=nh_events))
}

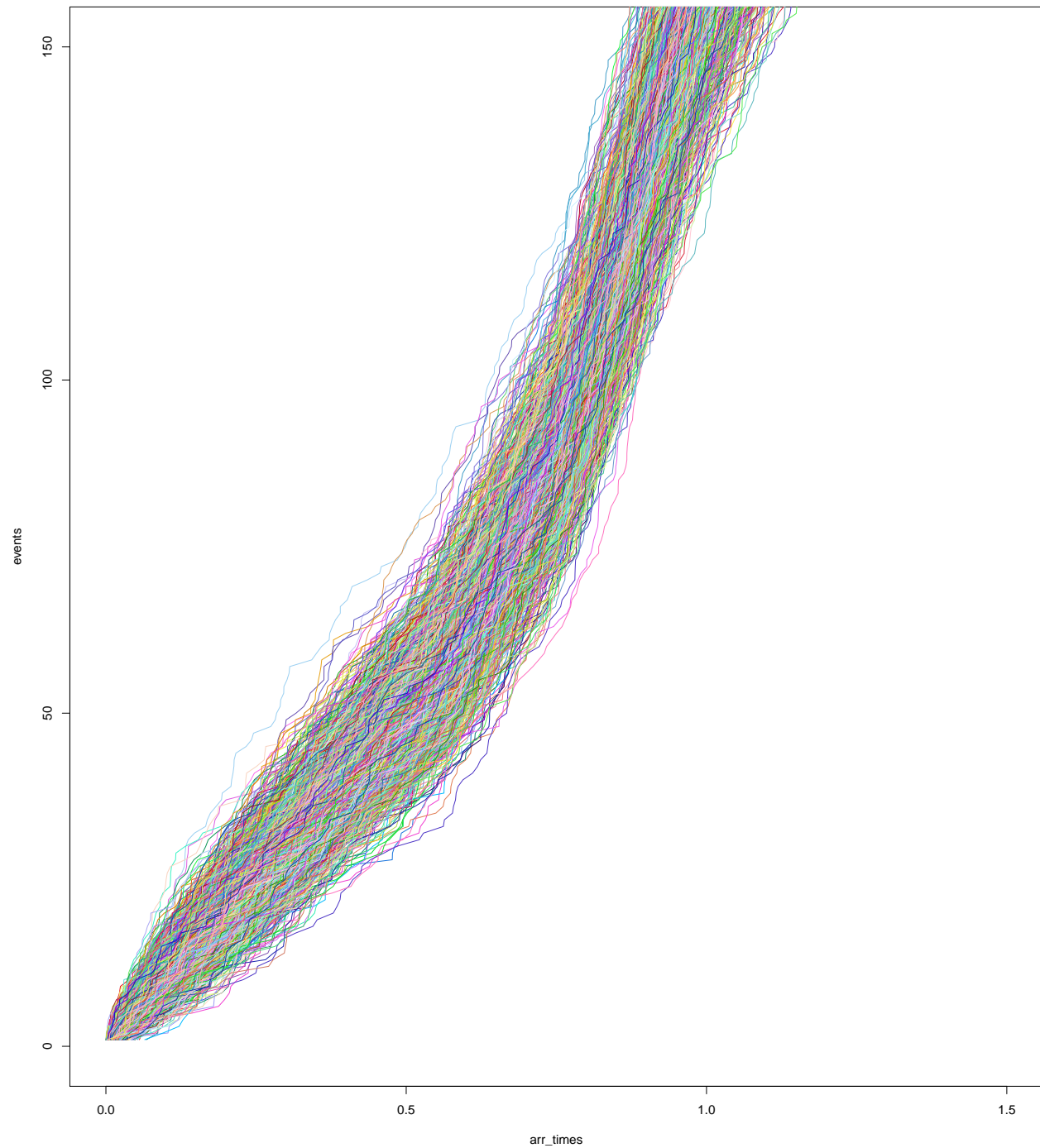
```

```

simulation <- function(iters, functions, lambdas, ints) {
  # This function simulates from the NHPP
  # iters:      number of iterations to plot and add to the list of
  #            dataframes
  # functions:  list of functions corresponding to the lambda function
  # lambdas:    list of lambdas for each subinterval
  # ints:      lists of vectors of 2 elements each containing the intervals
  #            that correspond to each element of lambdas and functions lists
  p <- list()
  for (i in 1:iters) {
    maximum <- 0 # start for the next NHPP simulation to continue count
    arr_times <- c() # arrival times
    events <- c() # event counts
    for (k in 1:4) {
      int <- non_hom_poisson(lambda_funs[[k]], lambdas[[k]],
                             ints[[k]][1], ints[[k]][2],
                             start=maximum)
      maximum <- max(int$events) # remembering last event count
      arr_times <- c(arr_times, int$arrival_times)
      events <- c(events, int$events)
    }
    p[[i]] <- data.frame(arrival_times=arr_times, events=events)
    # plots
    if (i == 1) {plot(arr_times, events, cex=0.5, pch='.',
                      col=randomColor(), xlim=c(0,1.5),
                      ylim=c(0,150))}
    lines(arr_times, events, col=randomColor())
  }
  return(p)
}

```

```
data <- simulation(1000, lambda_funs, lambdas, ints)
```



```
ratio <- 0
for (i in 1:length(data)) {
  df <- data.frame(data[[i]])
  cnt <- df %>% filter(arrival_times < 0.91232 & events >= 150) %>% dplyr::count()
  if (cnt[1] >= 1) {
    ratio <- ratio + 1
  }
}
```

```
ratio/1000
#> [1] 0.097
```

Problem 3

(a)

Our infinitesimal generator is the following:

$$Q = \begin{pmatrix} -\lambda & \lambda & 0 & 0 & \dots \\ \mu & -(\lambda + \mu) & \lambda & 0 & \dots \\ 0 & 2\mu & -(2\mu + \lambda) & \lambda & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{pmatrix}$$

(b)

We solve the following system:

$$\begin{cases} \sum_{i=1}^{\infty} \pi_i = 1 \\ \lambda \pi_0 = \mu \pi_1 \\ \lambda \pi_1 = 2\mu \pi_2 \\ \vdots \\ \lambda \pi_{n-1} = 2\mu \pi_n \\ \vdots \end{cases}$$

First we have:

$$\begin{aligned} \pi_1 &= \frac{\lambda \pi_0}{\mu} \\ \pi_2 &= \frac{\lambda^2 \pi_0}{2\mu^2} \\ \pi_3 &= \frac{\lambda^3 \pi_0}{2^2 \mu^3} \\ &\vdots \\ \pi_n &= \frac{\lambda^n \pi_0}{2^{n-1} \mu^n} \\ &\vdots \end{aligned}$$

Then:

$$\sum_{i=0}^{\infty} \pi_i = \pi_0 + \frac{\lambda \pi_0}{\mu} + \frac{\lambda^2 \pi_0}{2\mu^2} + \dots + \frac{\lambda^n \pi_0}{2^{n-1} \mu^n} + \dots = 1$$

And so factoring π_0 we get:

$$\pi_0 \left(1 + \frac{\lambda}{\mu} + \frac{\lambda^2}{2\mu^2} + \dots \right) = \pi_0 \left(1 + \sum_{i=1}^{\infty} \frac{1}{2^{i-1}} \left(\frac{\lambda}{\mu} \right)^i \right)$$

Then multiplying $\frac{2}{2}$ to the summation:

$$\pi_0 \left(1 + \frac{2}{2} \sum_{i=1}^{\infty} \frac{1}{2^{i-1}} \left(\frac{\lambda}{\mu} \right)^i \right)$$

$$= \pi_0 \left(2 \sum_{i=0}^{\infty} \left(\frac{\lambda}{2\mu} \right)^i - 1 \right)$$

$$\pi_0 \left(2 \left(\frac{1}{1 - \frac{\lambda}{2\mu}} \right) - 1 \right) = 1$$

$$\pi_0 = \frac{1}{2 \left(\frac{1}{1 - \frac{\lambda}{2\mu}} \right) - 1}$$

⋮

$$\pi_n = \frac{\lambda^n}{2^{n-1} \mu_n \pi_0} \frac{1}{2(\frac{1}{1-\frac{\lambda}{2\mu}})-1}$$

finally:

$$\pi_n = \frac{1}{2^{n-1}} \left(\frac{\lambda}{\mu}\right)^n \pi_0$$