

Stochastic Processes: Assignment 1

Group 1: Javier Esteban Aragoneses, Mauricio Marcos Fajgenbaun, Danyu Zhang, Daniel Alonso

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Importing libraries

```
#> Package:  markovchain
#> Version:  0.8.5-2
#> Date:     2020-09-07
#> BugReport: https://github.com/spedygiorgio/markovchain/issues
#>
#> Attaching package: 'dplyr'
#> The following objects are masked from 'package:stats':
#>
#>     filter, lag
#> The following objects are masked from 'package:base':
#>
#>     intersect, setdiff, setequal, union
```

Problem 1

```
# importing libraries
import numpy as np
from copy import deepcopy
import pandas as pd
import matplotlib.pyplot as plt
from nltk.corpus import words

# Importing the matrix with the frequencies
freq = np.loadtxt('./data/Englishcharacters.txt', usecols=range(27))

# transformation function for table values
def transf_log(table):
    return np.log(table + 1)

# applying the transformation function
freq = transf_log(freq)

# importing the messages and only selecting the
# message with index 0, as this is the one corresponding
# to group 1
with open('./data/messages.txt') as f:
    message = f.readlines()[0].replace('\n', '')

# decode function
def decode(message, freqs, iters):
```

```

"""
Function to decode a message encoded
using a substitution cipher utilizing the
Metropolis-Hastings algorithm.

Params:
message = string to decode
freqs = frequency table for transitions of letters
        to input
iters = amount of iterations to perform
"""

# dictionary to organize the iterations
# the score and the result of the attempt
# to decode the message corresponding to that
# iteration
msg_iters = {}

# defining the identity function
# all letters to be used excluding spaces
letters = ["a", "b", "c", "d",
           "e", "f", "g", "h",
           "i", "j", "k", "l",
           "m", "n", "o", "p",
           "q", "r", "s", "t",
           "u", "v", "w", "x",
           "y", "z"]

# creating a copy of the original letters
# to use as key for the dictionaries
init_letters = deepcopy(letters)
# creating a dictionary with init_letters as keys
# and letters as values
cd = {l:d for l,d in zip(init_letters,letters)}
# every time we update with {' ':' '} we add the space
# to the dictionary
cd.update({' ':' '})
# define a function that just uses the previous
# dictionary to seek the letters
def f(c):
    return cd[c]

# this dictionary and subsequent function maps each letter
# to a column/row in the freq matrix, ex: 'a':0, 'b':1
fvals = np.array([x for x in range(len(letters))])
cd_map = {l:v for l,v in zip(letters,fvals)}
cd_map.update({' ':26})
def f_map(c):
    return cd_map[c]

# score function uses sum of logs
def score(fun):
    p = 0
    for i in range(1,len(msg)):
        p = p + freq[f_map(fun(msg[i-1])),f_map(fun(msg[i]))]

```

```

return p

# converting the message to a list in order to
# go through the letters in pairs
msg = list(message)

# letters list, this one shall be modified
# every time the score passes the test
letters_n = deepcopy(letters)

# loop iters amount of times
for i in range(iters):
    # randomly choose 2 numbers and replace the 2 chosen
    # vals in a copy of letters
    ch1 = np.random.randint(0,len(letters))
    ch2 = np.random.randint(0,len(letters))
    plc1 = deepcopy(letters_n[ch1])
    plc2 = deepcopy(letters_n[ch2])
    letters_n[ch1] = plc2
    letters_n[ch2] = plc1

    # create the dictionary for the f* function
    cd_n = {l:v for l,v in zip(init_letters,letters_n)}
    # add the space to it after scramble
    cd_n.update({' ':' '})
    # f* definition
    def f_n(c):
        return cd_n[c]

    # calculating the score for each function and its ratio
    scr_f = score(f)
    scr_fn = score(f_n)
    a = scr_fn/scr_f
    # test if a random number is lower than min(a, 1)
    cond = np.random.rand() <= min(a,1)
    # if condition is true
    if cond:
        # replacing the letters list with the one from f*
        letters = deepcopy(letters_n)
        # updating the dictionary with the new letters list after replacing
        cd = {l:v for l,v in zip(init_letters,letters)}
        cd.update({' ':' '})
        # f re-definition
        def f(c):
            return cd[c]

        # replacing the letters in the message using
        # the new f replaced by f*
        for k in range(len(msg)):
            msg[k] = f(msg[k])

        # adding score and joining the message to the dictionary
        # to then transform into a dataframe
        msg_iters[i] = (a, ''.join([x for x in msg]))
    # if condition is false

```

```

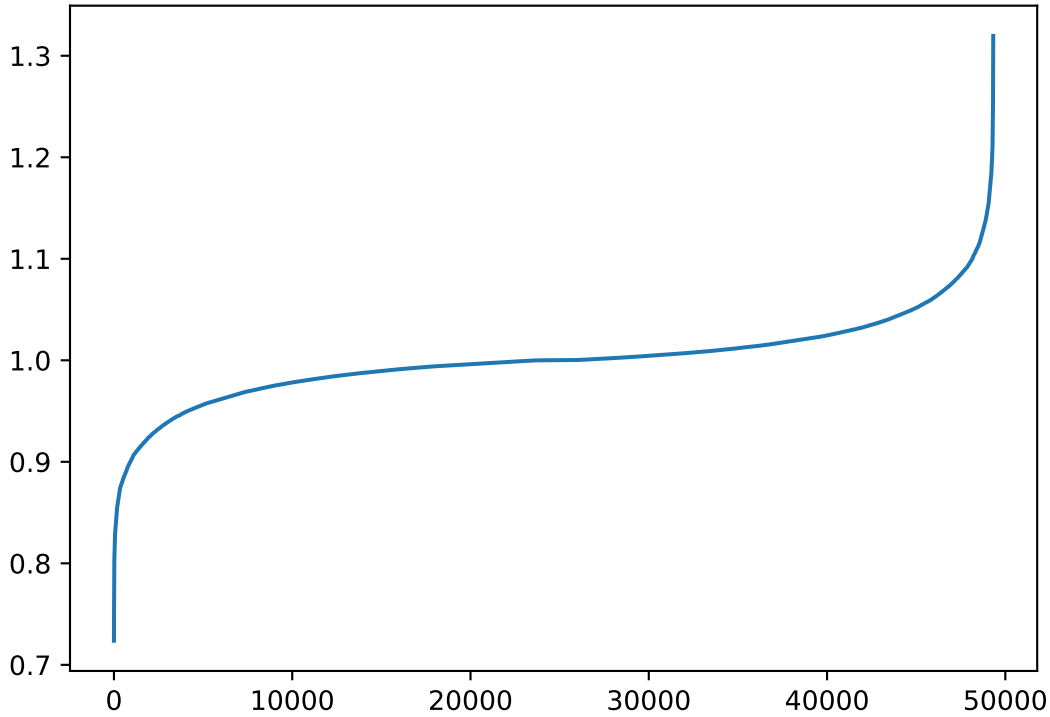
else:
    # apply the f function to the message instead
    for k in range(len(msg)):
        msg[k] = f(msg[k])
    # reset the message
    msg = list(message)

    # break the loop if 2 words are found in the english language corpus
    try:
        msg_list = msg_iters[i].split(' ')
        fw = {'w1':np.random.choice(msg_list),
              'w2':np.random.choice(msg_list)}
        conds = {w_n:(w in words.words()) for w_n,w in fw.items()}
        vals = list(conds.values())
        if False not in vals:
            print(f'found at iteration: {i}')
            print(msg_iters[i])
            print(f'words found: {list(conds.keys())}')
            break
        # otherwise continue
    except:
        continue

    # put the information in a dataframe, iters, score and the messages
    df = {'iter':[it for it in msg_iters.keys()],
          'score':[msg[0] for msg in msg_iters.values()],
          'msg':[msg[1] for msg in msg_iters.values()]}
    # return the dataframe
    return pd.DataFrame(df)

# we run the function
result = decode(message,freq, 50000)
plt.plot(result.sort_values('score')['score'].reset_index(drop=True))
plt.show()

```



```
print(result[result['score'] == max(result['score'])])
```

```
#>      iter      score      msg
#> 30938 31353 1.319558 v svt vwfoc hcqfvqp hozudo flo wqcbf mobfocp z...
```

```
#> [1] "n bnq ndrsw awornoc aslnxs rzs dowfr ksfrswc ljicrwq rj nirzjwoms n ljwjcngowif gnllocs aworofz"
```

Problem 2

(a)

Let $N(t)$ be the number of cars arriving at a parking lot by time t , according to the proposed scenario, we can model $N(t)$ as a non-homogenous Poisson process. Such process has almost the same process as any other Poisson process, however, its rate is a function of time.

$N(t), t \in [0, \infty)$ is the non-homogenous Poisson process with rate $\lambda(t)$ where:

- $N(0) = 0$
- $N(t)$ has independent increments

We define 8:00 as $t = 0$ with the following integrable function and each unit of t equals to 1 hour:

$$\lambda(t) = \begin{cases} 100 & 0 \leq t \leq \frac{1}{2} \\ 600t - 200 & \frac{1}{2} < t \leq \frac{3}{4} \\ 400t - 50 & \frac{3}{4} < t \leq 1 \\ -500t + 850 & 1 < t \leq 1.5 \end{cases}$$

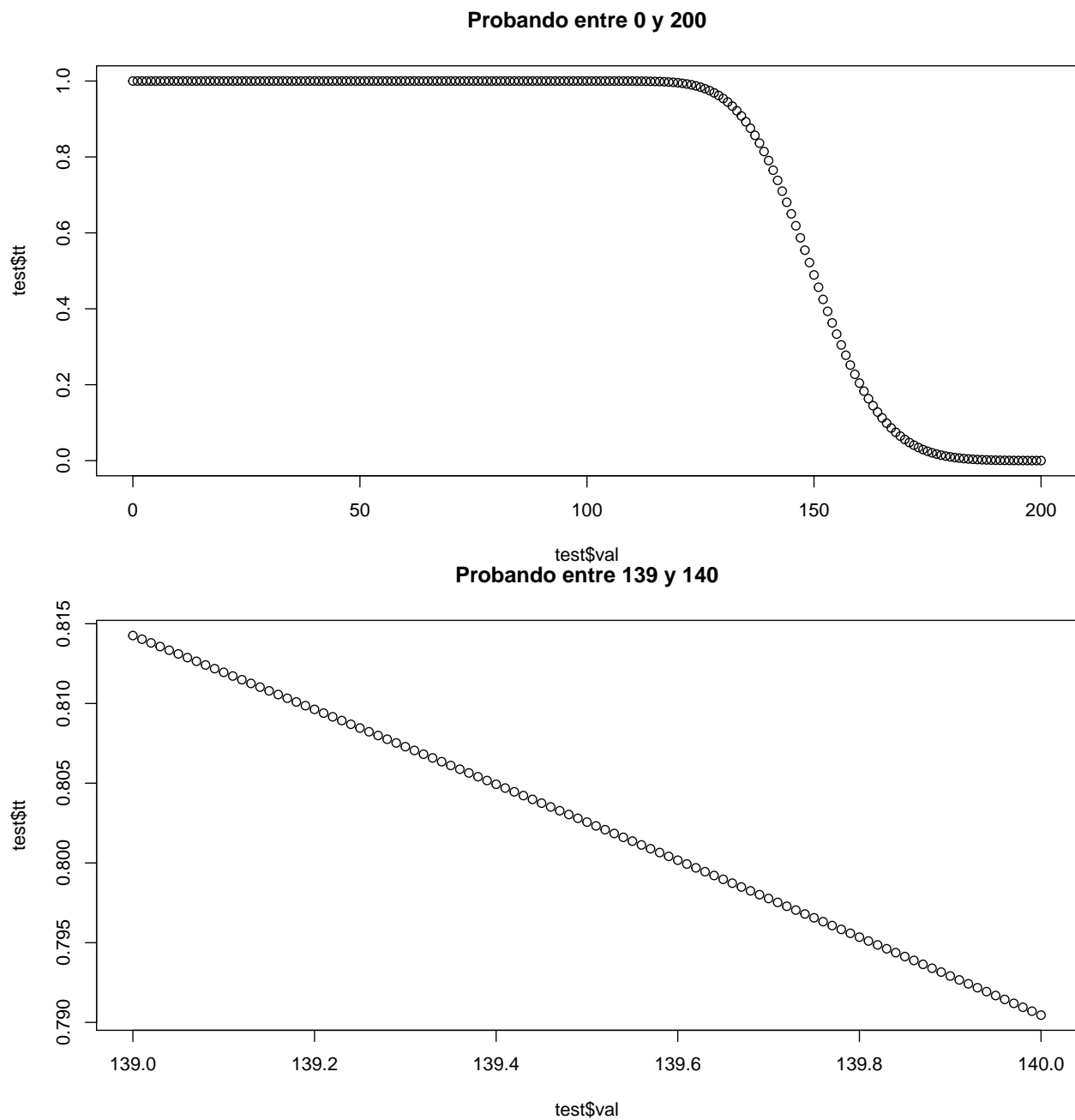
So,

$$E[N(t)] = \begin{cases} \int_0^t 100 \, dt = 100t & 0 \leq t \leq \frac{1}{2} \\ \int_{\frac{1}{2}}^t 600t - 200 \, dt + 50 = 300(t^2 - \frac{1}{4}) - 200(t - \frac{1}{2}) + 50 & \frac{1}{2} < t \leq \frac{3}{4} \\ \int_{\frac{3}{4}}^t 400t - 50 \, dt + 93.75 = 25(8t^2 - 2t - 3) + 93.75 & \frac{3}{4} < t \leq 1 \\ \int_1^t -500t + 850 \, dt + 168.75 = -50(5t^2 - 17t + 12) + 168.75 & 1 < t \leq 1.5 \end{cases}$$

Given that there is a limit of 150 vehicles:

$$E[N(t)] = \begin{cases} 100t & 0 \leq t \leq \frac{1}{2} \\ 300(t^2 - \frac{1}{4}) - 200(t - \frac{1}{2}) + 50 & \frac{1}{2} < t \leq \frac{3}{4} \\ 25(8t^2 - 2t - 3) + 93.75 & \frac{3}{4} < t < 0.94468 \\ 150 & t \geq 0.94468 \end{cases}$$

(b)



Luego de hacer las pruebas para $\lambda(t)$ obtenemos lo siguiente:

```
lambda = 139.6  
t = 0.91232  
# 8:44 AM
```

Por lo que $t = 0.91232$ horas (aproximadamente a las 8:54 de la mañana).

(c)

The following function simulates a non-homogenous poisson process from a homogenous poisson process:

```

non_hom_poisson <- function(fun,l,a,b,start=0) {
  # This function generates a non-homogenous poisson
  # process from a homogenous poisson process
  # PARAMS:
  # fun:    if the non-homogenous poisson process has
  #         multiple functions per time subinterval
  #         this parameters represents such function
  # l:      lambda for the homogenous poisson process
  # a:      lower bound for the time subinterval
  # b:      upper bound for the time subinterval
  # start:  this parameter is used to keep track of
  #         the process count.

  # We generate the homogenous poisson process
  # arrival times
  val <- rpois(1,l*(b-a))
  intervals <- (b-a) * sort(runif(val)) + a

  # Non-homogenous poisson process
  evs <- length(intervals) # lenght of arrival times
  nh_val <- 0 + start # start of the event count
  nh_intervals <- c() # arrival times for the NHPP
  for (i in 1:evs) {
    if (runif(1) < fun(intervals[i])/l) {
      # only including intervals from the HPP which
      # match with fun(intervals[i])/l probability
      nh_intervals <- c(nh_intervals, intervals[i])
      nh_val <- nh_val+1 # adding one to the event count
    }
  }
  nh_events <- seq(1+start,nh_val,1) # events since the previous group
  return(list(arrival_times=nh_intervals, events=nh_events))
}

```

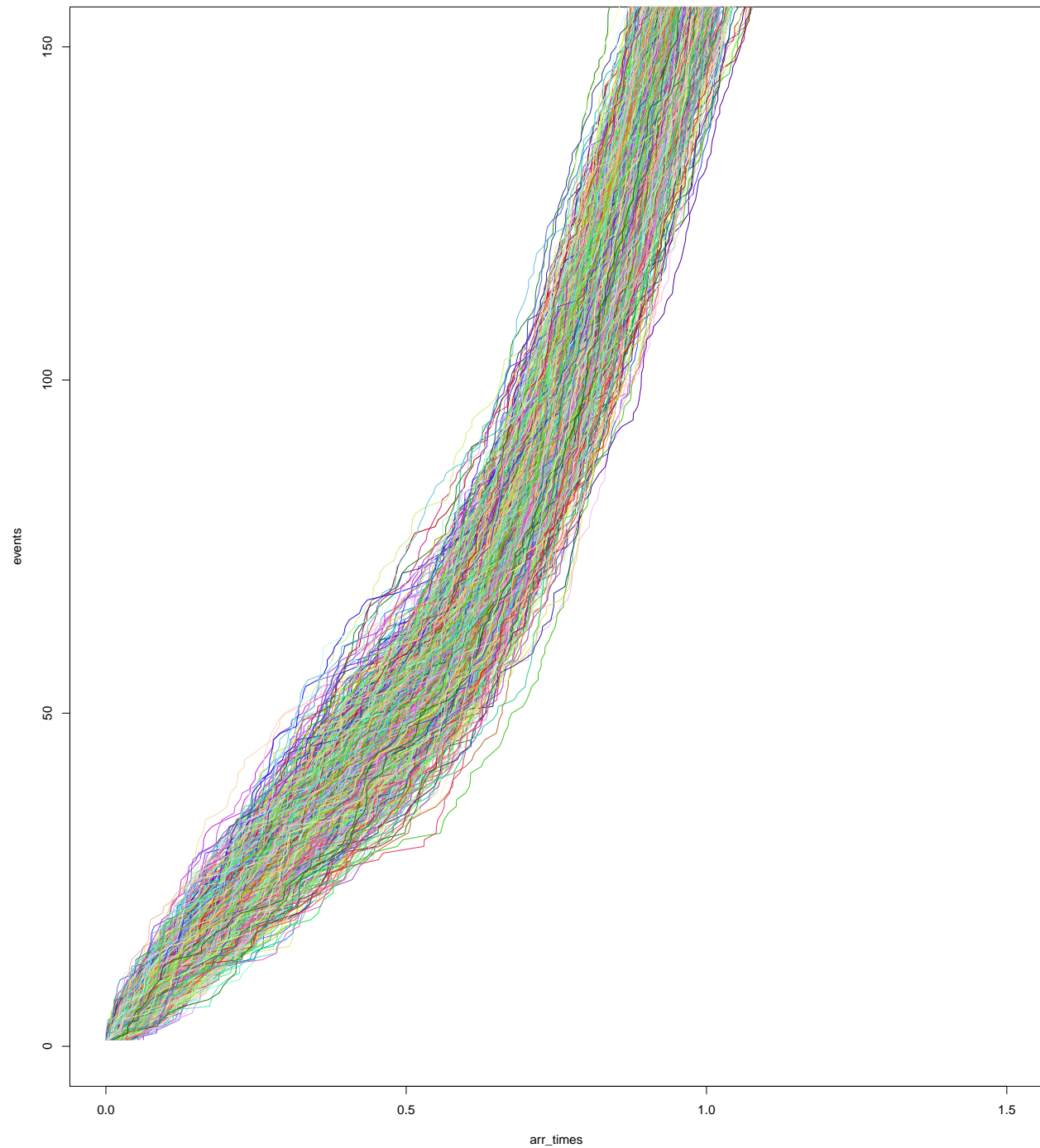


```

simulation <- function(iters, functions, lambdas, ints) {
  # This function simulates from the NHPP
  # iters:      number of iterations to plot and add to the list of
  #            dataframes
  # functions:  list of functions corresponding to the lambda function
  # lambdas:    list of lambdas for each subinterval
  # ints:      lists of vectors of 2 elements each containing the intervals
  #            that correspond to each element of lambdas and functions lists
  p <- list()
  for (i in 1:iters) {
    maximum <- 0 # start for the next NHPP simulation to continue count
    arr_times <- c() # arrival times
    events <- c() # event counts
    for (k in 1:4) {
      int <- non_hom_poisson(lambda_funs[[k]], lambdas[[k]],
                            ints[[k]][1], ints[[k]][2],
                            start=maximum)
      maximum <- max(int$events) # remembering last event count
      arr_times <- c(arr_times, int$arrival_times)
      events <- c(events, int$events)
    }
    p[[i]] <- data.frame(arrival_times=arr_times, events=events)
    # plots
    if (i == 1) {plot(arr_times, events, cex=0.5, pch='.',
                      col=randomColor(), xlim=c(0,1.5),
                      ylim=c(0,150))}
    lines(arr_times, events, col=randomColor())
  }
  return(p)
}

```

```
data <- simulation(1000, lambda_funs, lambdas, ints)
```



```
ratio <- 0
for (i in 1:length(data)) {
  df <- data.frame(data[[i]])
  cnt <- df %>% filter(arrival_times < 0.91232 & events >= 150) %>% dplyr::count()
  if (cnt[1] >= 1) {
    ratio <- ratio + 1
  }
}
```

```
ratio/1000
#> [1] 0.192
```

Problem 3

(a)

Our infinitesimal generator is the following:

$$Q = \begin{pmatrix} -\lambda & \lambda & 0 & 0 & \dots \\ \mu & -(\lambda + \mu) & \lambda & 0 & \dots \\ 0 & 2\mu & -(2\mu + \lambda) & \lambda & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{pmatrix}$$

(b)

We solve the following system:

$$\begin{cases} \sum_{i=1}^{\infty} \pi_i = 1 \\ \lambda \pi_0 = \mu \pi_1 \\ \lambda \pi_1 = 2\mu \pi_2 \\ \vdots \\ \lambda \pi_{n-1} = 2\mu \pi_n \\ \vdots \end{cases}$$

First we have:

$$\begin{aligned} \pi_1 &= \frac{\lambda \pi_0}{\mu} \\ \pi_2 &= \frac{\lambda^2 \pi_0}{2\mu^2} \\ \pi_3 &= \frac{\lambda^3 \pi_0}{2^2 \mu^3} \\ &\vdots \\ \pi_n &= \frac{\lambda^n \pi_0}{2^{n-1} \mu^n} \\ &\vdots \end{aligned}$$

Then:

$$\sum_{i=0}^{\infty} \pi_i = \pi_0 + \frac{\lambda \pi_0}{\mu} + \frac{\lambda^2 \pi_0}{2\mu^2} + \dots + \frac{\lambda^n \pi_0}{2^{n-1} \mu^n} + \dots = 1$$

And so factoring π_0 we get:

$$\pi_0(1 + \frac{\lambda}{\mu} + \frac{\lambda^2}{2\mu^2} + \dots) = \pi_0(1 + \sum_{i=1}^{\infty} \frac{1}{2^{i-1}} (\frac{\lambda}{\mu})^i)$$

Then multiplying $\frac{2}{2}$ to the summation:

$$\pi_0(1 + \frac{2}{2} \sum_{i=1}^{\infty} \frac{1}{2^{i-1}} (\frac{\lambda}{\mu})^i)$$

$$= \pi_0(2 \sum_{i=0}^{\infty} (\frac{\lambda}{2\mu})^i - 1)$$

$$\pi_0(2(\frac{1}{1 - \frac{\lambda}{2\mu}}) - 1) = 1$$

$$\pi_0 = \frac{1}{2(\frac{1}{1 - \frac{\lambda}{2\mu}}) - 1}$$

⋮

$$\pi_n = \frac{\lambda^n}{2^{n-1} \mu_n \pi_0} \frac{1}{2(\frac{1}{1-\frac{\lambda}{2\mu}})-1}$$

finally:

$$\pi_n = \frac{1}{2^{n-1}} \left(\frac{\lambda}{\mu}\right)^n \pi_0$$

The infinite sum converges when $|\frac{\lambda}{2\mu}| < 1$ in which case the stationary distribution P exists.

Then:

$$L = \sum_{n=0}^{\infty} \pi_n * n$$

Using the following sum:

$$\sum_{i=0}^{n-1} i a^i = \frac{a - n a^n + (n-1) a^{n+1}}{(1-a)^2}$$

As n approaches infinity:

$$\sum_{i=0}^{\infty} i a^i = \frac{a}{(1-a)^2}$$

We get the following:

$$L = \pi_0 [1 + \sum_{n=0}^{\infty} \left(\frac{\lambda}{\mu}\right)^n * \frac{n}{2^{n-1}}]$$

$$L = \pi_0 [1 + 2 \sum_{n=0}^{\infty} \left(\frac{\lambda}{2\mu}\right)^n * n]$$

$$L = \pi_0 [1 + 2 \frac{\frac{\lambda}{2\mu}}{(1-\frac{\lambda}{2\mu})^2}]$$

(c)

Let's consider the probabilities conditioned on the number of customers in the system that are present once our specific subject l gets into the system.

If there are no other customers when l gets into the system, there is no chance of overtaking.

$$P(N^{OV} = 0 | N^{PR} = 0) = 1$$

With N^{OV} being the number of customers that l overtakes and N^{PR} the number of customers present in the system (queuing) when l gets in the system.

If $N^{PR} = 1$, then l can overtake only 1 customer, if the time it takes to be served is shorter than the time it takes the other customers to be served. Because of the memoryless property we can assert the following:

$$P(N^{OV} = 0 | N^{PR} = 1) = \frac{\mu}{\mu + \mu} = \frac{1}{2}$$

Actually, in general:

$$P(N^{OV} = k | N^{PR} = n) = \frac{1}{n+1}, n \leq c-1, k = 0, 1$$

As in this case $c=2$, our l subject can't overtake more than one customer.

Now, if $n \geq c$, that is, l has to get in queue and wait to be served. When l gets served, there is also one more customer getting served. Because, again, of the memoryless property.

$$P(N^{OV} = k | N^{PR} = n) = \frac{1}{c}, n = c, k = 0, 1$$

In our case, it does not matter how many customers are in the system, the probability of overtaking, conditioned to the number of customers already in the system, is $\frac{1}{2}$.

Now, using Bayes' theorem and the total probability rule, we can find the probability of l overtaking another customer.

$$\begin{aligned}
P(A|B) &= \frac{P(A \cap B)}{P(B)} \\
\frac{1}{2} \sum_{i=1}^{\infty} \pi_i &= \frac{1}{2} \sum_{i=1}^{\infty} \left(\frac{1}{2^i-1}\right) \left(\frac{\lambda}{\mu}\right)^i \pi_0 \\
&= \sum_{i=1}^{\infty} \left(\frac{\lambda}{2\mu}\right)^i \pi_0 = \left(\sum_{i=0}^{\infty} \left(\frac{\lambda}{2\mu}\right)^i - 1\right) \pi_0 \\
&= \left(\frac{1}{1-\frac{\lambda}{2\mu}} - 1\right) \pi_0 = \frac{1}{2\left(\frac{1}{1-\frac{\lambda}{2\mu}} - 1\right)} \left(\frac{1}{1-\frac{\lambda}{2\mu}} - 1\right) \\
&= \frac{1}{1-\frac{\lambda}{2\mu}} - 1 = \frac{1-(1-\frac{\lambda}{2\mu})}{1-\frac{\lambda}{2\mu}} = \frac{\frac{\lambda}{2\mu}}{1-\frac{\lambda}{2\mu}} \\
&= 2\left(\frac{1}{1-\frac{\lambda}{2\mu}}\right) - 1 = \frac{2}{1-\frac{\lambda}{2\mu}} - 1 = \frac{2-(1-\frac{\lambda}{2\mu})}{1-\frac{\lambda}{2\mu}} \\
&= \frac{1+\frac{\lambda}{2\mu}}{1-\frac{\lambda}{2\mu}}
\end{aligned}$$

Then:

$$\frac{\frac{\lambda}{2\mu}}{1-\frac{\lambda}{2\mu}} = \frac{\frac{\lambda}{2\mu}}{\frac{2\mu-\lambda}{2\mu}} = \frac{\lambda}{2\mu-\lambda}$$

So then we get:

$$P(N^{OV} = k) = \frac{\lambda}{2\mu-\lambda}, k = c-1 = 1$$

(d)

We define the following function to simulate the queueing system:

```

# Simulation of the System (M/M/2)
q <- function(customers, l, m) {
  # This function generates a M/M/2 queue system and returns a matrix
  # with columns of ArrivalTimes, exit=ExitTimes and service=ServiceTimes
  # PARAMS:
  # customers: number of customers in the supermarket
  # l:        lambda for the homogenous poisson process
  #           (customers arrive to the unique cashiers waiting
  #           line according this rate)
  # m:        mu for the exponential distribution
  #           (The times to be served are independent and
  #           distributed as exponential with rate mu)

  # interval for arrival times
  exp_at <- rexp(customers,l)

  # we have as many arrival times as the number of customers
  ArrivalTimes <- rep(0,customers)

  ArrivalTimes[1] <- exp_at[1]
  for (i in 2:customers) {
    # arrival time = the previous arrival time + the interval between 2 customers
    ArrivalTimes[i] <- ArrivalTimes[i-1] + exp_at[i]
  }

  # service time distributed according to mu
  ServiceTimes <- rexp(customers, m)

  # we have as many exit times as the number of customers

```

```

ExitTimes <- rep(0,customers)

# the first two exit time is equal to the service time due to we have 2 cashiers
ExitTimes[1:2] <- ServiceTimes[1:2]
for (i in 3:customers) {
  # we sort exit time from larger to smaller
  SortedTimes <- sort(ExitTimes[1:(i-1)], decreasing=T)

  # all of the two cashiers are occupied, then the new customer will have to
  # wait until at least one of them leaves the supermarket (the faster one)
  # then the exit time of the new customer is exited time of the previous faster customer plus
  # the service time of this new customer
  if (ArrivalTimes[i] < SortedTimes[2]) {ExitTimes[i] <- SortedTimes[2] + ServiceTimes[i]}

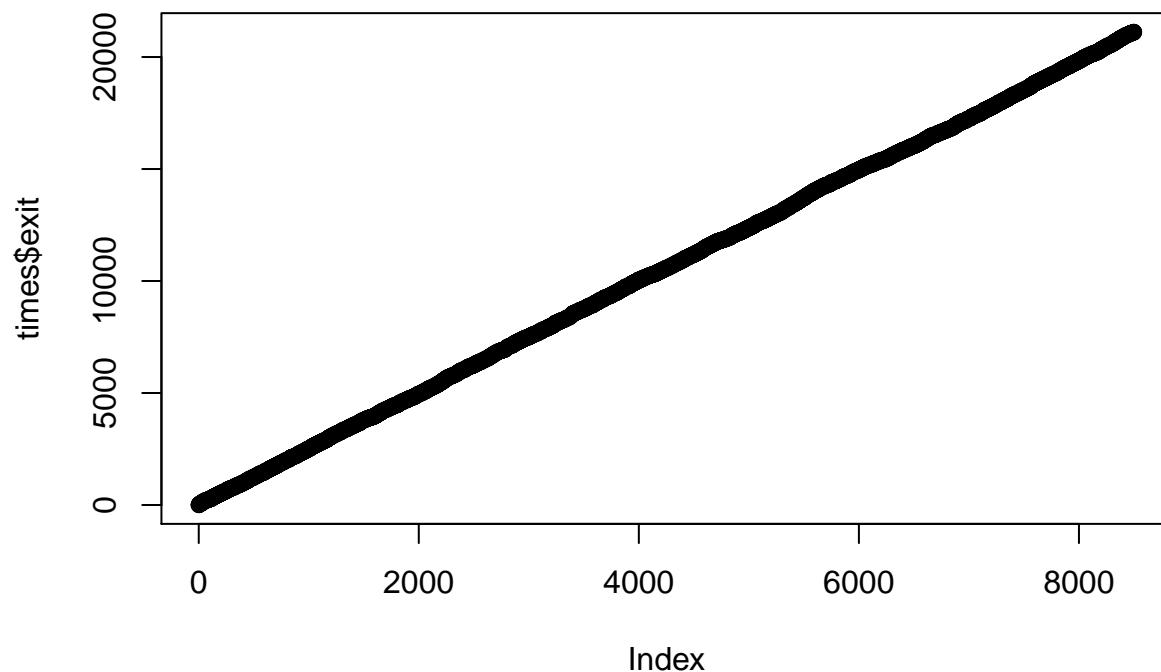
  # one or two cashiers are free, then the exit time is the arrival time plus the service time
  else {ExitTimes[i] <- ArrivalTimes[i] + ServiceTimes[i]}
}

# create a matrix and return it
times <- data.frame(arrival=ArrivalTimes,
                    exit=ExitTimes,
                    service=ServiceTimes)

return(times)
}

number_of_customers <- 8500
times <- q(number_of_customers, l=0.4, m=0.25)
plot(times$exit)

```



(e)

We define the following function to calculate the overtaking probability:

```

d <- function(customers, queue, ExitTimes) {
  # This function generates a M/M/2 queue system and returns
  # the probability of overtaking
  # PARAMS:
  # customers:    number of customers in the supermarket
  # queue:        number of people in the queue
  # ExitTimes:    the exited times of each customer

  # we define the number of people of overtaking
  ot <- 0
  for (i in queue:(customers - queue)) {

    # the number of people overtaking of all customers is the previous
    # number plus 1 if the exit time of the i-th customer is smaller (earlier)
    # than the (i-1)th customer
    ot <- ot + sum(ExitTimes[i] < ExitTimes[1:(i-1)])
  }
  r <- customers-2*queue
  ot_prob <- ot/r
  return(ot_prob)
}

```

Write the R code necessary to simulate the system (provide the code) and generate the times customers leave the supermarket

```

avgs <- c() # general average of multiple simulations
number_of_customers <- 8500 # simulations done with 8500 people
for (i in 1:10) {
  times <- q(number_of_customers, l=0.4, m=0.25) # running the simulation
  avg <- c() # estimation (usually includes an extra person)
  avg_p <- c() # pessimistic estimation (usually excludes that extra person)
  for (i in 0:99) {
    # assuming the person number number_of_customers - i is the last one
    last_arrival_time <- times$arrival[length(times$arrival)-i]
    # checking all exit times mayores al arrival time of the person number number_of_customers - i
    val <- sum(times$exit[1:(length(times$arrival)-i-1)] > last_arrival_time)
    avg <- c(avg, val)
    avg_p <- c(avg, val-1)
  }
  # joining all estimations
  total_avg <- mean(c(avg, avg_p))
  # adding to vector with all averages
  avgs <- c(avgs, total_avg)
}
# printing averages
avgs
#> [1] 2.278607 4.228856 2.636816 11.771144 2.064677 3.621891 2.696517
#> [8] 5.905473 2.179104 10.149254
# calculating the average of averages
mean(avgs)
#> [1] 4.753234

```

Calculation by hand:

```

lambda <- 0.4
mu <- 0.25
lambda2mu <- lambda/(2*mu)

pi_0 <- 1/(2*(1/(1-lambda2mu)) - 1)
L <- pi_0 * (1 + 2*(lambda2mu/((1-lambda2mu)^2)))

```

$$\frac{\lambda}{2\mu} = \frac{0.4}{2(0.25)} = 0.8$$

$$\pi_0 = \frac{1}{2(\frac{1}{1-0.8})-1} = \frac{1}{9} \approx 0.\bar{1}$$

$$L = \pi_0(1 + 2\frac{0.8}{(1-0.8)^2}) = 4.\bar{5}$$