# Using Joint Models to Estimate Causal Effects for Salvage Therapy after Prostatectomy

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# Aims, Models & Estimands

## 1 Background & Aim



- Setting Patients treated with surgery after diagnosis of Prostate Cancer (PCa)
  - > remain at risk of metastasis
- Follow-up
  - > PSA levels at frequent intervals

  - > ST androgen deprivation therapy, radiation therapy, chemotherapy, and combinations



- Important questions regarding Salvage Therapy

  - ▶ when to start?
  - ▷ does it work?



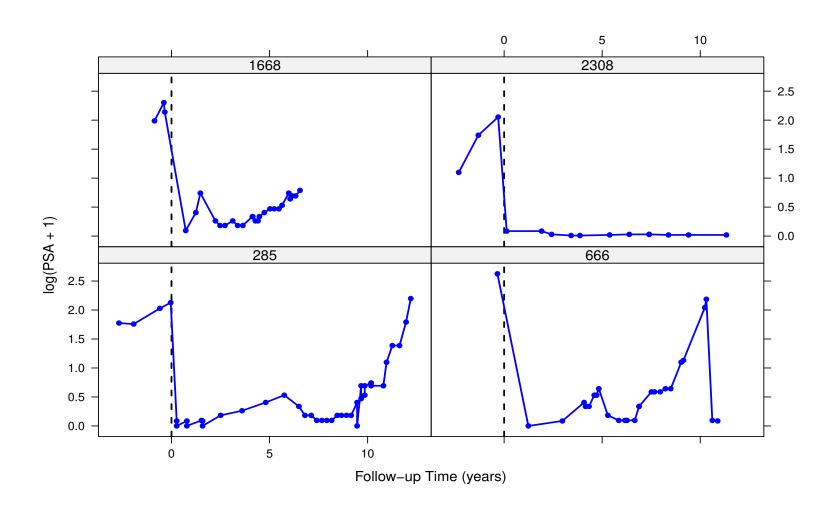
Quantify the amount by which Salvage Therapy reduces the risk of metastasis



#### University of Michigan Prostatectomy Data

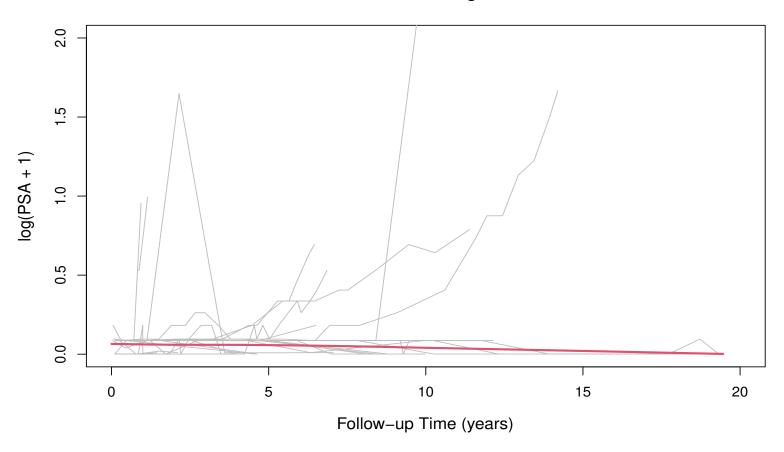
- ⇒ 3672 PCa patients treated with prostatectomy 1994–2013
- baseline variables: PSA, Gleason, T-stage, age, race, gland volume, perineural invasion, planned adjuvant therapy
- - \* post-surgery PSA values (median = 6)
  - \* post-surgery salvage therapy (n = 324)
  - \* PSA values also after salvage (median = 3)
  - \* metastasis (n = 108)
  - \* death information (n = 212)





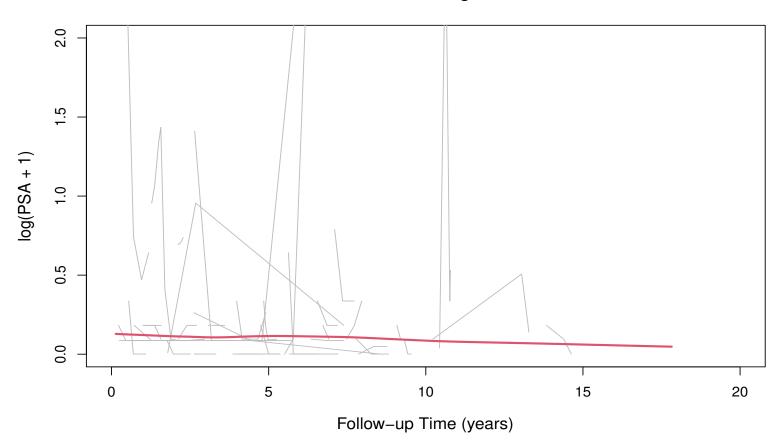


#### **Before Salvage**

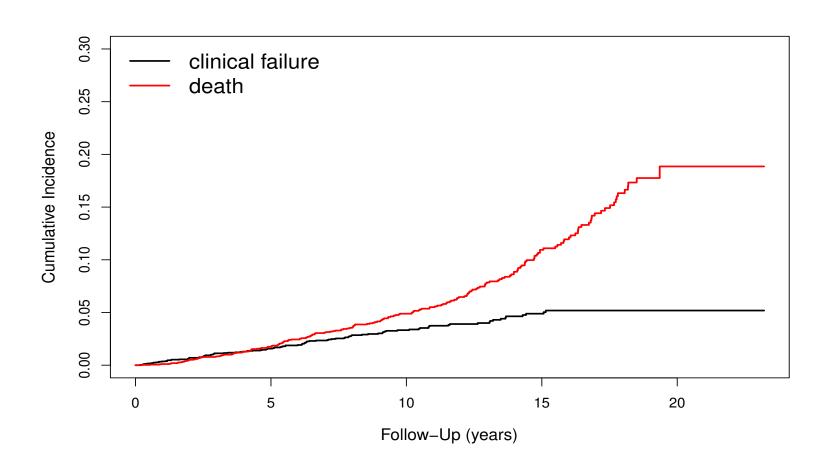




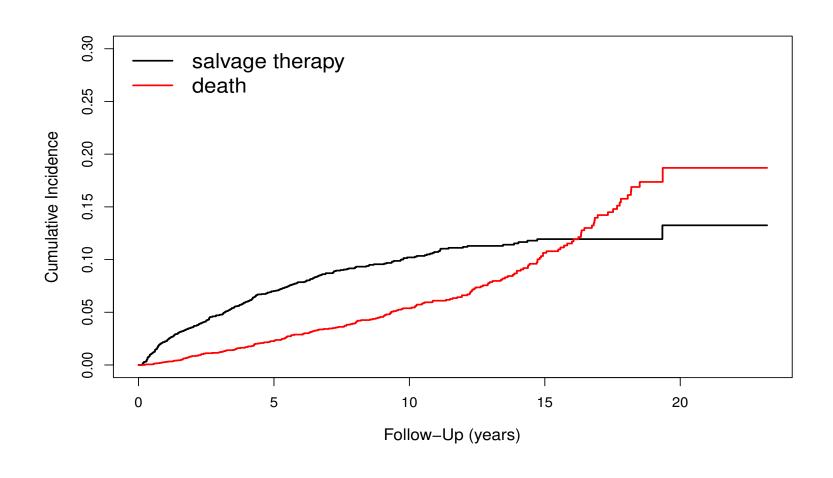














## Challenges

- ▷ Observational Data no RCT
  - \* selection bias
  - \* ascertainment bias
- ▷ Time-Varying Salvage Therapy
  - \* depends on previous PSA
  - \* PSA time-dependent confounder
  - \* endogeneity

#### 2 Causal ST Effects

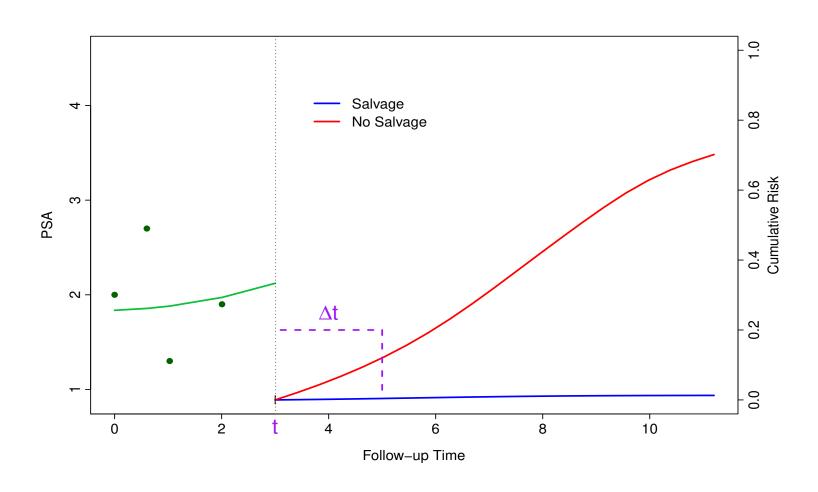


- Standard assumptions for Causal Inference
  - *Consistency:* Observed outcomes equal the counterfactual outcomes for the actually assigned treatment
  - ▷ Sequential Exchangeability: The counterfactual outcomes are independent of the assigned treatment conditionally on the history of PSA measurements and baseline covariates
  - $\triangleright$  *Positivity:* Each patient has a nonzero probability of receiving ST at each time point t



- Setting
  - $\triangleright$  PSA measurements up to t
  - $\triangleright$  no Salvage Therapy given up to t
  - $\triangleright$  we compare cumulative risk of metastasis in the medically-relevant interval  $[t,t+\Delta t]$
  - □ under the two regimes
    - 1. if Salvage Therapy is **not** given in the interval  $[t, t + \Delta t]$
    - 2. if Salvage Therapy is given at t







#### Which is the target group?

#### Notation

 $\triangleright T_m$ : time to metastasis

 $\triangleright T_d$ : time to death

 $\triangleright \mathcal{H}^*(t)$ : a version of the PSA history up to t

 $hd T_m^{(a)}$  and  $T_d^{(a)}$  counterfactual outcomes

\* a = 1, ST given at t

\* a=0, ST was not given in  $[t,t+\Delta t]$ 



# Marginal Salvage Therapy Effect

 $\triangleright$  we average over all PSA histories, i.e.,  $\mathcal{H}^*(t) = \emptyset$ 

$$ST^{M}(t + \Delta t, t) = \Pr\{T_{m}^{(1)} \le t + \Delta t \mid T_{m} > t, T_{d} > t\} - \Pr\{T_{m}^{(0)} \le t + \Delta t \mid T_{m} > t, T_{d} > t\}$$

#### • Notes:

 $\triangleright$  of lesser relevance to the urologists because they decide who gets ST based on PSA  $\Rightarrow$  more bias

 $\triangleright$  averages over a big groups of patients  $\Rightarrow$  **less variance** 



## Conditional Salvage Therapy Effect

 $\triangleright$  we condition on the PSA history of a specific patient, i.e.,  $\mathcal{H}^*(t) = \mathcal{H}_i(t)$ 

$$ST^{C}(t + \Delta t, t) = \Pr\{T_{m}^{(1)} \le t + \Delta t \mid T_{m} > t, T_{d} > t, \mathcal{H}_{i}(t)\}$$
$$-\Pr\{T_{m}^{(0)} \le t + \Delta t \mid T_{m} > t, T_{d} > t, \mathcal{H}_{i}(t)\}$$

#### • Notes:

- $\triangleright$  much more relevant to the urologists  $\Rightarrow$  **less bias**
- $\triangleright$  averages over a narrow group of patients identified via modeling assumptions  $\Rightarrow$  **more variance**



# Marginal-Conditional Salvage Therapy Effect

 $\triangleright$  consider ST for patients who had PSA levels above the threshold value c at their last visit, i.e.,  $\mathcal{H}^*(t) = \{Y(t): Y(t) > c\}$ 

$$ST^{MC}(t + \Delta t, t) = \Pr\{T_m^{(1)} \le t + \Delta t \mid T_m > t, T_d > t, \mathcal{H}^*(t)\}$$
$$-\Pr\{T_m^{(0)} \le t + \Delta t \mid T_m > t, T_d > t, \mathcal{H}^*(t)\}$$

#### • Notes:

- $\triangleright$  relevant to the urologists  $\Rightarrow$  compromised bias
- ▷ averages over a bigger group of patients ⇒ compromised variance

#### 3 Structural Models



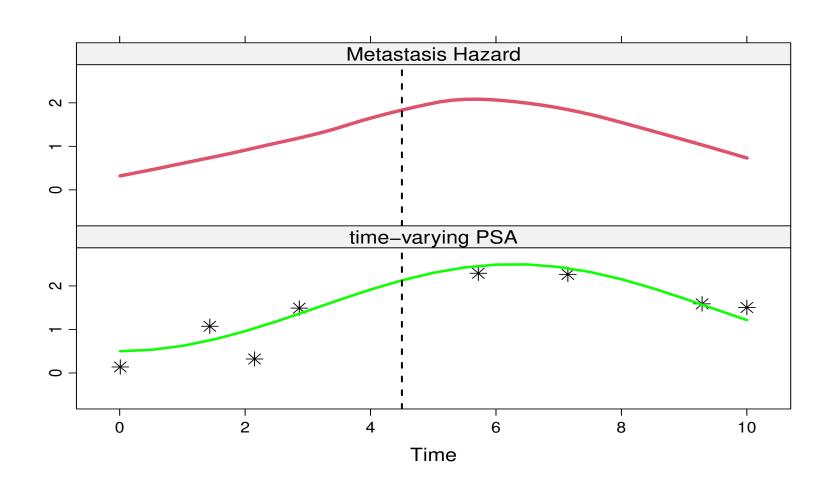
Standard Cox models not appropriate



Joint Models for Longitudinal and Time-to-Event Data

# 3 Structural Models (cont'd)





# 3 Structural Models (cont'd)



- Because joint models use all available data,
  - > they account for the time-varying confounding
  - ▷ no extra weighting is necessary
  - > they provide valid causal effects

#### 4 PSA Sub-Model

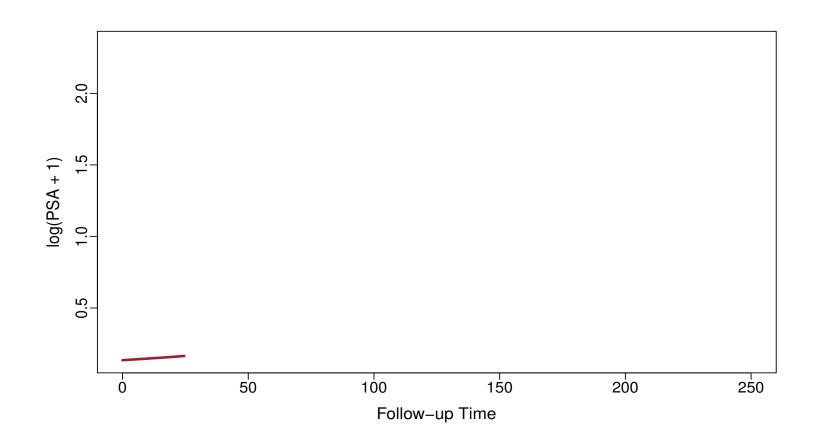


- As PSA increases, patient may receive ST
- We let  $S_i$  denote the time a patient initiated ST
  - $\triangleright$  for patients who did not initiate ST,  $S_i = \infty$
- After ST, PSA levels are expected to drop
  - but may rise again before metastasis

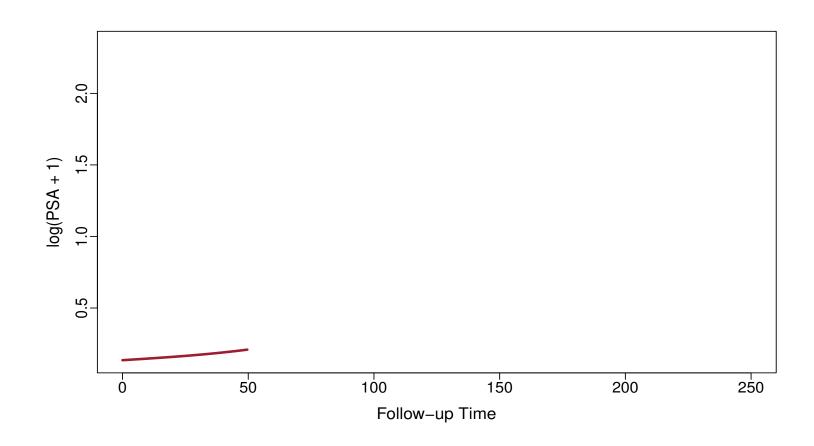


$$\log\{\mathsf{PSA}_i(t)+1\} = \begin{cases} \eta_i(t) + \varepsilon_i(t) = \boldsymbol{x}_i(t)\boldsymbol{\beta} + \boldsymbol{z}_i(t)\boldsymbol{b}_i + \varepsilon_i(t), \ t < S_i \\ \\ \tilde{\eta}_i(t) + \varepsilon_i(t) = \\ \\ \eta_i(t) + \left\{\tilde{\boldsymbol{x}}_i(\tilde{t})\tilde{\boldsymbol{\beta}} + \tilde{\boldsymbol{z}}_i(t)\tilde{\boldsymbol{b}}_i\right\} + \varepsilon_i(t), \quad t \geq S_i, \end{cases}$$

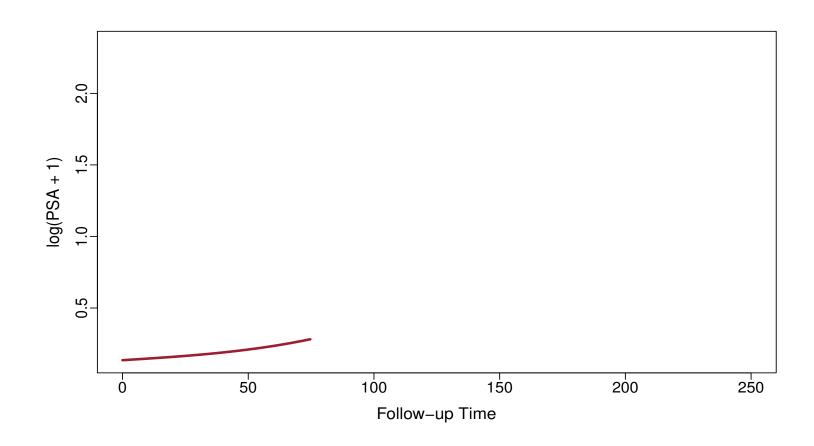




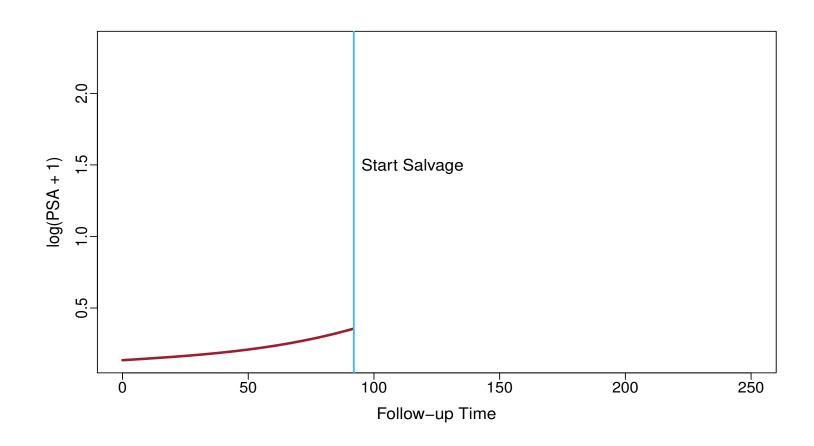




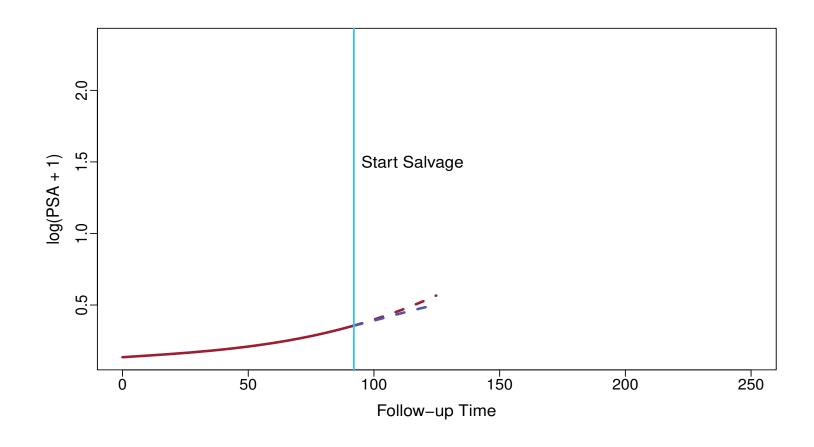




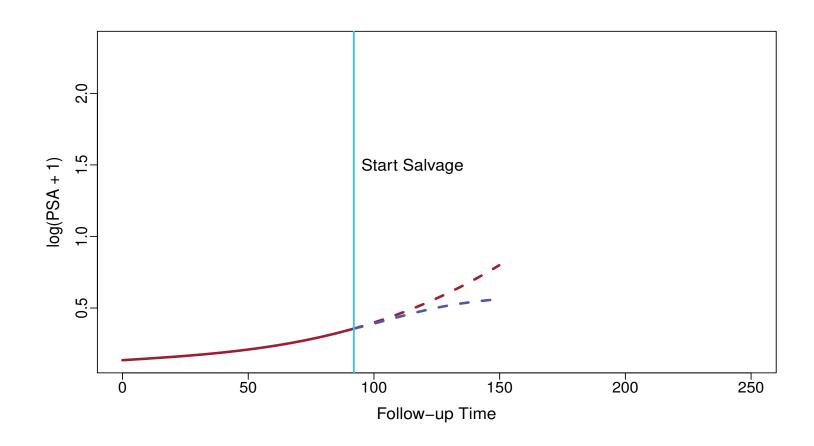




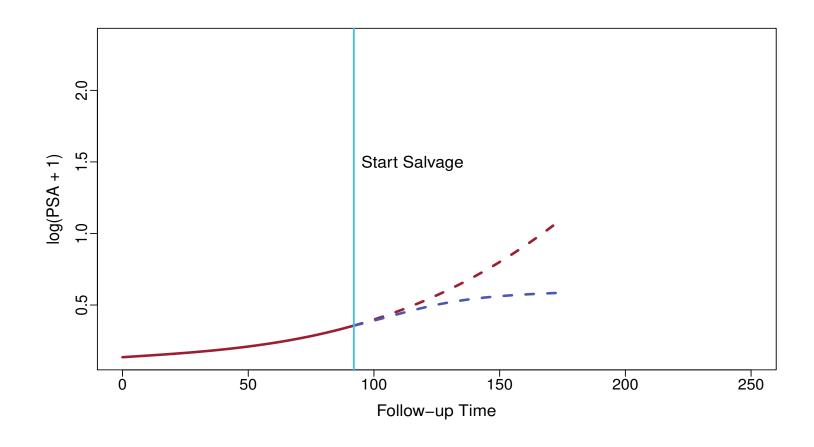




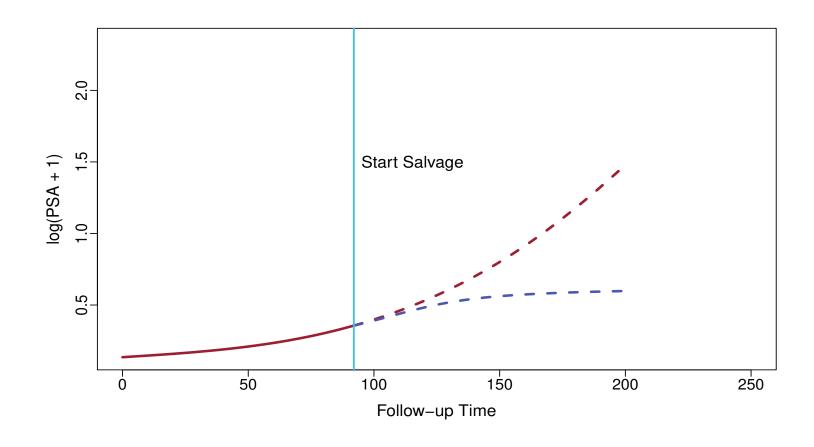




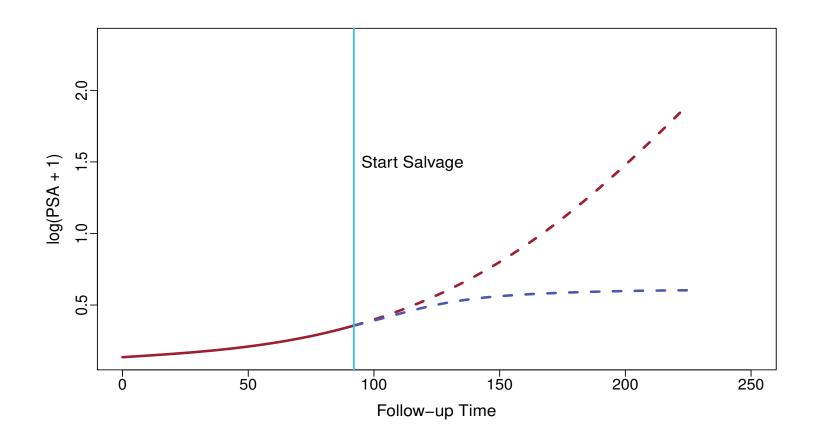




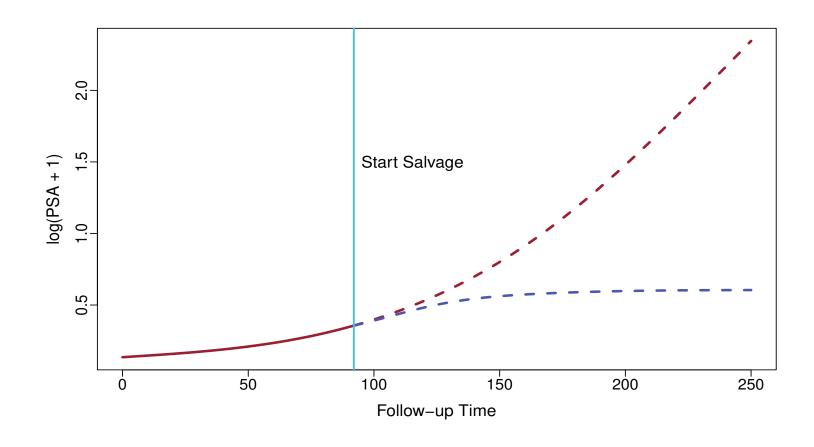














#### The model used in the UM data

- Fixed effects
  - ▷ Before Salvage: Nonlinear PSA evolution (B-spline with 6 internal knots)
  - ▷ After Salvage: Drop in PSA, and linear evolution
  - ▷ baseline covariates: Age, Gleason score, Charlson comorbidity index
- Random effects
  - > the same time effect as in the fixed part

#### 5 Metastasis and Death Sub-Models



- Metastasis and Death treated as *Competing Risks*
- Separate hazard models for metastasis and death

  - ▷ baseline covariates

## 5 Metastasis and Death Sub-Models (cont'd)



• Metastasis Sub-Model linked to baseline covariates, Salvage and PSA

$$h_i^m(t) = \begin{cases} h_0^m(t) \exp\left(\boldsymbol{\psi}_m^{\top} \boldsymbol{w}_i + \boldsymbol{\alpha}_m^{\top} f\{\eta_i(t)\}\right), & t < S_i \\ h_0^m(t) \exp\left(\boldsymbol{\psi}_m^{\top} \boldsymbol{w}_i + \gamma_m(t - S_i) + \boldsymbol{\xi}_m^{\top} g\{\tilde{\eta}_i(t)\}\right), & t \ge S_i \end{cases}$$



- ullet Functions  $f(\cdot)$  and  $g(\cdot)$  specify the functional form
  - before and after Salvage is linked to metastasis

• Some options are. . .



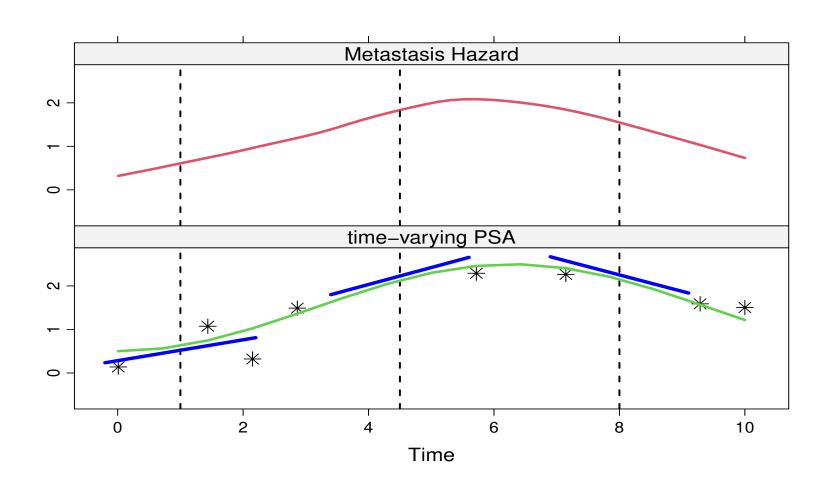
• *Time-dependent Slopes:* The hazard of metastasis at t is associated with both the current value and the slope of the PSA trajectory at t:

$$h_i^m(t \mid \mathcal{H}_i(t)) = h_0^m(t) \exp\{\boldsymbol{\psi}_m^{\mathsf{T}} \boldsymbol{w}_i + \alpha_{m1} \eta_i(t) + \alpha_{m2} \eta_i'(t)\},$$

where

$$\eta_i'(t) = \frac{d}{dt} \{ x_i^{\top}(t)\beta + z_i^{\top}(t)b_i \}$$





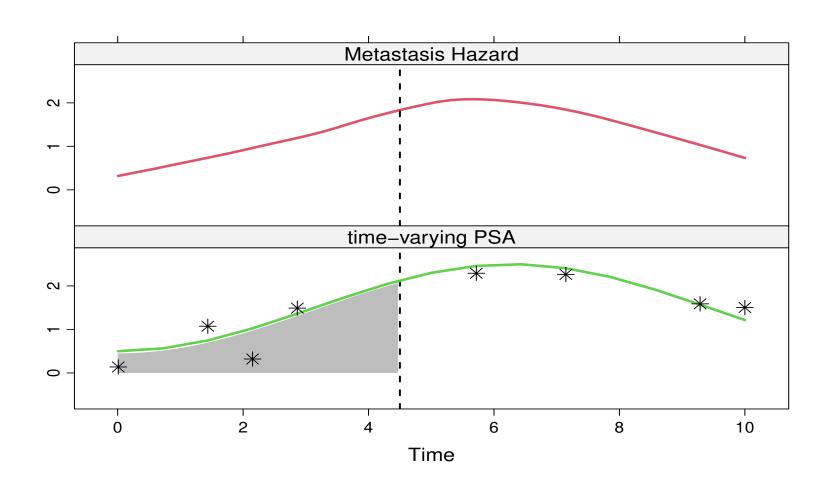


• *Cumulative Effects:* The hazard of metastasis at t is associated with the whole area under the trajectory up to t:

$$h_i(t \mid \mathcal{M}_i(t)) = h_0(t) \exp\left\{ \gamma^{\top} w_i + \alpha \frac{\int_0^t m_i(s) ds}{t} \right\}$$

We account for the observation period







#### Models used in the UM data

- Functional forms
  - ▷ Before Salvage: Nonlinear PSA evolution (B-spline with 6 internal knots)
    - \* value
    - \* value + slope
    - \* value + cumulative effect
  - ▷ After Salvage: Drop in PSA, and linear evolution
    - \* drop in PSA
    - \* slope
  - ▷ baseline covariates: Age, Gleason score, Charlson comorbidity index



• Death Sub-Model linked to baseline covariates, Salvage and but not PSA

$$h_i^d(t) = \begin{cases} h_0^d(t) \exp(\boldsymbol{\psi}_d^{\top} \boldsymbol{w}_i), & t < S_i \\ h_0^d(t) \exp(\boldsymbol{\psi}_d^{\top} \boldsymbol{w}_i + \gamma_d), & t \ge S_i \end{cases}$$

### 6 Causal Effect Estimation



• From the joint model, we can obtain the conditional causal effect

$$\Pr\{T_{mi}^{(a)} \leq t + \Delta t \mid T_{mi} > t, T_{di} > t, \mathcal{H}_{i}(t), \mathcal{X}_{i}\} =$$

$$\int \int \Pr\{T_{mi}^{(a)} \leq t + \Delta t \mid T_{mi} > t, T_{di} > t, \boldsymbol{u}_{i}, \mathcal{X}_{i}, \boldsymbol{\theta}\}$$

$$\times p\{\boldsymbol{u}_{i} \mid T_{mi} > t, T_{di} > t, \mathcal{H}_{i}(t), \mathcal{X}_{i}, \boldsymbol{\theta}\} \ p(\boldsymbol{\theta} \mid \mathcal{D}) \ d\boldsymbol{u}_{i} d\boldsymbol{\theta}$$



- ullet Monte Carlo scheme to estimate  $\mathrm{ST}_i^C(t+\Delta t,t)$ 
  - riangle sample  $reve{m{ heta}}^{(l)}$  from the posterior of the parameters  $[m{ heta} \mid \mathcal{D}]$
  - ightharpoonup sample  $m{ar{u}}_i^{(l)}$  from the posterior of the random effects  $[m{u}_i \mid T_{mi} > t, T_{di} > t, \mathcal{H}_i(t), \mathcal{X}_i, m{ar{ heta}}^{(l)}]$

$$ho$$
 calculate  $\pi_i^{(l)}(t + \Delta t \mid t, a) = \Pr\{T_{mi}^{(a)} \leq t + \Delta t \mid T_{mi} > t, T_{di} > t, \boldsymbol{\check{u}}_i^{(l)}, \boldsymbol{\mathcal{X}}_i, \boldsymbol{\check{\theta}}^{(l)}\}$ 

• We repeat L times and get

$$\widehat{\mathsf{ST}}_i^C(t + \Delta t, t) = \frac{1}{L} \sum_{l=1}^L \pi_i^{(l)}(t + \Delta t \mid t, a = 1) - \pi_i^{(l)}(t + \Delta t \mid t, a = 0)$$



- ullet Estimation of  $\mathrm{ST}^M(t+\Delta t,t)$  and  $\mathrm{ST}^{MC}(t+\Delta t,t)$  proceeds by averaging the conditional effects over the respective groups of patients
- ullet For example, for  $\mathrm{ST}^M(t+\Delta t,t)$ 
  - $\triangleright \mathcal{R}(t)$  the subset of patients at risk at time t and who have not initiated ST by t
  - ho for each patient in  $\mathcal{R}(t)$ , we calculate  $\widehat{\mathsf{ST}}_i^C(t+\Delta t,t)$

$$\widehat{\mathsf{ST}}^M(t+\Delta t,t) = n_r^{-1} \sum_{i:i \in R(t)} \widehat{\mathsf{ST}}_i^C(t+\Delta t,t),$$



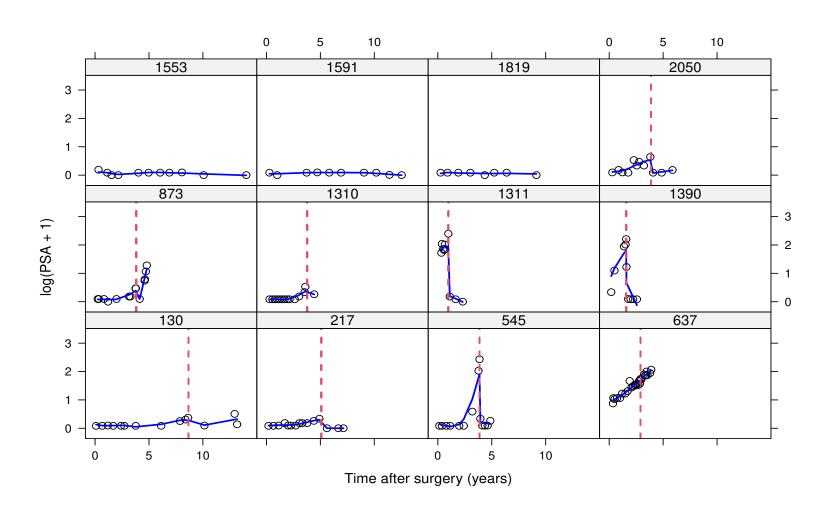
ullet To estimate the variance of the causal effects, we need to take into account that they are a function of both the parameters ullet and the data  $\mathcal D$ 

$$\mathsf{Var}_{\mathcal{D}}\big\{\widehat{\mathsf{ST}}^{M}\big(t+\Delta t,t;\boldsymbol{\theta},\mathcal{D}\big)\big\} = \mathsf{Var}_{\mathcal{D}}\bigg[E_{\boldsymbol{\theta}|\mathcal{D}}\Big\{\mathsf{ST}^{M}\big(t+\Delta t,t;\boldsymbol{\theta},\mathcal{D}\big)\Big\}\bigg]$$

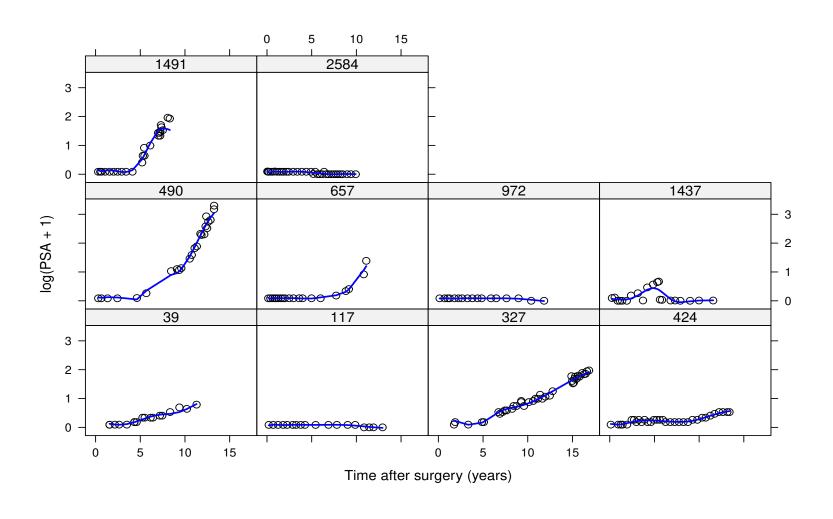
• We achieve this using an adaptation of the procedure of Antonelli et al. (2021)

### 7 Results

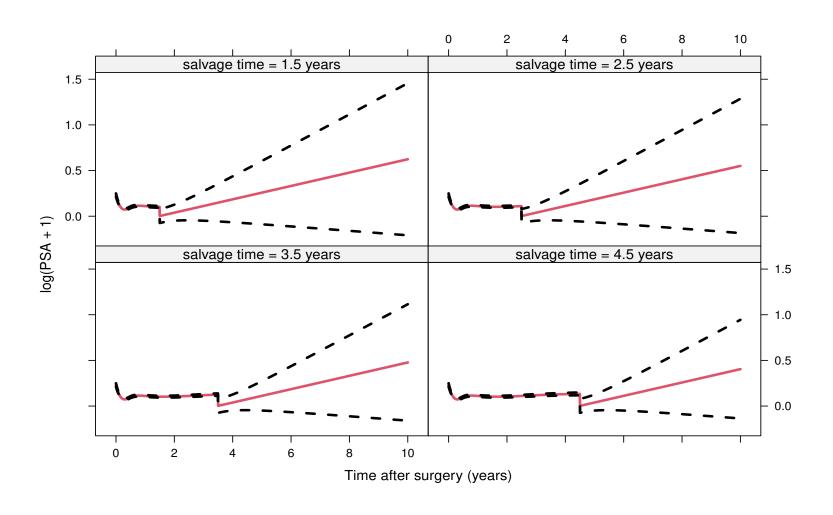








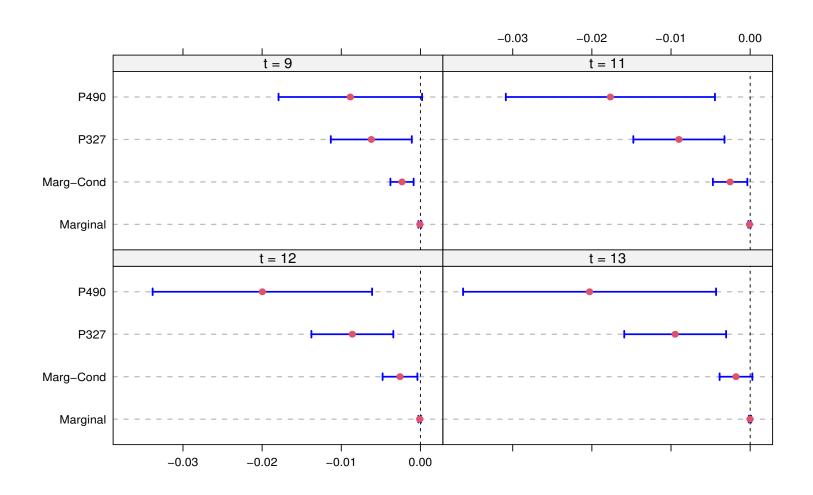






https://emcbiostatistics.shinyapps.io/Plots\_PSA/





#### 8 Extensions & Discussion



• Competing Risks ⇒ Multi-State

#### Competing Risks

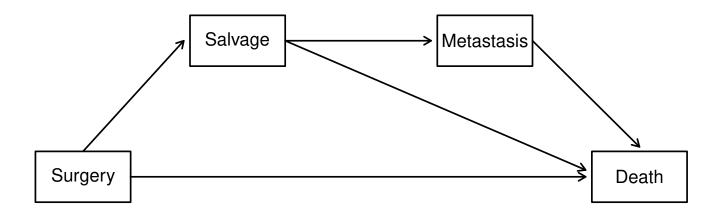
- > salvage as a time-varying covariate

#### • Multi-State

- $\triangleright$  metastasis  $\rightarrow$  death transition

## 8 Extensions & Discussion (cont'd)





## 8 Extensions & Discussion (cont'd)



- Implementation available in **JMbayes2** 
  - > predict() cumulative incidence risks
- Shiny app...

## Thank for your attention!

https://www.drizopoulos.com/



Where the first term is written as

$$\Pr\{T_{mi}^{(a)} \leq t + \Delta t, | T_{mi} > t, T_{di} > t, \boldsymbol{u}_i, \boldsymbol{\mathcal{X}}_i, \boldsymbol{\theta}\} =$$

$$\frac{\int_{t}^{t+\Delta t} h_{i}^{m(a)}(v) \exp \left(-\int_{t}^{v} \left\{h_{i}^{m(a)}(s) + h_{i}^{d(a)}(s)\right\} \, \mathrm{d}s - \int_{0}^{t} \left\{h_{i}^{m(0)}(s) + h_{i}^{d(0)}(s)\right\} \, \mathrm{d}s\right) \, \mathrm{d}v}{\exp \left(-\int_{0}^{t} \left\{h_{i}^{m(0)}(s) + h_{i}^{d(0)}(s)\right\} \, \mathrm{d}s\right)}$$



• Using telescoping we get:

$$p(\boldsymbol{\theta}, \boldsymbol{u}, \boldsymbol{\theta}_{N} \mid T, \delta, \boldsymbol{Y}, \boldsymbol{N})$$

$$\propto \prod_{i=1}^{n} \prod_{j=1}^{n_{i}} p\{Y_{i}(t_{ij}), T_{i}, \delta_{i} \mid \mathcal{Y}_{i}(t_{i,j-1}), \mathcal{N}_{i}(t_{i,j-1}), \mathcal{X}_{i}, \boldsymbol{\theta}, \boldsymbol{u}_{i}\}$$

$$\times \prod_{j=1}^{n_{i}} p\{N_{i}(t_{ij}) \mid \mathcal{Y}_{i}(t_{i,j-1}), \mathcal{N}_{i}(t_{i,j-1}), Y_{i}(t_{ij}), T_{i}, \delta_{i}, \mathcal{X}_{i}, \boldsymbol{\theta}_{N}, \boldsymbol{u}_{i}\}$$

$$\times p(\boldsymbol{u}_{i} \mid \boldsymbol{\theta}) \times p(\boldsymbol{\theta}) \times p(\boldsymbol{\theta}_{N})$$



• Under sequential exchangeability, we have that

$$p\{N_i(t_{ij}) \mid \mathcal{Y}_i(t_{i,j-1}), \mathcal{N}_i(t_{i,j-1}), Y_i(t_{ij}), T_i, \delta_i, \mathcal{X}_i, \boldsymbol{\theta}_N, \boldsymbol{u}_i\} = p\{N_i(t_{ij}) \mid \mathcal{Y}_i(t_{i,j-1}), \mathcal{N}_i(t_{i,j-1}), \mathcal{X}_i, \theta_N\}$$

ullet  $\Rightarrow$  inference can be based on the first term (i.e., the observed data model) and ignore the second term

## 8 Computational Details (cont'd)



- Custom-made and tailored MCMC algorithm
  - □ Gibbs sampling (hierarchical centering for fixed effects)
  - ▷ adaptive Metropolis-Hastings
  - ▷ (Metropolis-adjusted Langevin algorithm for certain parameter)
  - > centered design matrices
- Speed via parallel sampling of random effects
- Chains run in parallel