FCDs for robust regression model

Denis Valle

December 2017

Here are the full conditional distributions (FCD's) for this problem:

- For β_0 :

$$\begin{split} p(\beta_0|\ldots) &\propto \left[\prod_i N\left(y_i|\beta_0 + \beta_1 x_i, \frac{\sigma^2}{\tau_i}\right)\right] N(\beta_0|0,10) \\ &\propto \left[\prod_i \exp\left(-\frac{\tau_i}{2\sigma^2}(y_i - \beta_1 x_i - \beta_0)^2\right)\right] \exp\left(-\frac{1}{2\times 10}\beta_0^2\right) \\ &\propto \exp\left(-\frac{1}{2\sigma^2}\sum_i \tau_i(-2\beta_0[y_i - \beta_1 x_i] + \beta_0^2)\right) \exp\left(-\frac{1}{2\times 10}\beta_0^2\right) \\ &\propto \exp\left(-\frac{1}{2}\bigg\{\beta_0^2\left[\frac{1}{\sigma^2}\sum_i \tau_i + \frac{1}{10}\right] - 2\beta_0\left(\frac{1}{\sigma^2}\right)\sum_i \tau_i[y_i - \beta_1 x_i]\bigg\}\right) \\ p(\beta_0|\ldots) &= N\left(\left[\frac{1}{\sigma^2}\sum_i \tau_i + \frac{1}{10}\right]^{-1}\frac{1}{\sigma^2}\sum_i \tau_i[y_i - \beta_1 x_i], \left[\frac{1}{\sigma^2}\sum_i \tau_i + \frac{1}{10}\right]^{-1}\right) \end{split}$$

- For β_1 :

$$\begin{split} p(\beta_1|\dots) &\propto \left[\prod_i N \left(y_i | \beta_0 + \beta_1 x_i, \frac{\sigma^2}{\tau_i} \right) \right] N(\beta_1 | 0, 10) \\ &\propto \left[\prod_i \exp \left(-\frac{\tau_i}{2\sigma^2} (y_i - \beta_0 - \beta_1 x_i)^2 \right) \right] \exp \left(-\frac{1}{2 \times 10} \beta_1^2 \right) \\ &\propto \exp \left(-\frac{1}{2} \frac{1}{\sigma^2} \sum_i \tau_i \left(-2\beta_1 x_i [y_i - \beta_0] + \beta_1^2 x_i^2 \right) \right) \exp \left(-\frac{1}{2 \times 10} \beta_1^2 \right) \\ &\propto \exp \left(-\frac{1}{2} \left[\beta_1^2 \left[\frac{1}{\sigma^2} \sum_i \tau_i x_i^2 + \frac{1}{10} \right] - 2\beta_1 \frac{1}{\sigma^2} \sum_i \tau_i x_i [y_i - \beta_0] \right] \right) \\ p(\beta_1|\dots) &= N \left(\left[\frac{1}{\sigma^2} \sum_i \tau_i x_i^2 + \frac{1}{10} \right]^{-1} \frac{1}{\sigma^2} \sum_i \tau_i x_i [y_i - \beta_0], \left[\frac{1}{\sigma^2} \sum_i \tau_i x_i^2 + \frac{1}{10} \right]^{-1} \right) \end{split}$$

- For τ_i :

$$p(\tau_i|\dots) \propto N\left(y_i|\beta_0 + \beta_1 x_i, \frac{\sigma^2}{\tau_i}\right) Gamma\left(\tau_i|\frac{v}{2}, \frac{v}{2}\right)$$

$$\propto \frac{1}{\sqrt{\frac{1}{\tau_i}}} \exp\left(-\frac{\tau_i}{2\sigma^2} (y_i - \beta_0 - \beta_1 x_i)^2\right) \tau_i^{\frac{v}{2} - 1} \exp\left(-\tau_i \left(\frac{v}{2}\right)\right)$$

$$\propto \tau_i^{\frac{1+v}{2} - 1} \exp\left(-\tau_i \left[\frac{1}{2\sigma^2} (y_i - \beta_0 - \beta_1 x_i)^2 + \frac{v}{2}\right]\right)$$

$$p(\tau_i | \dots) = Gamma\left(\frac{1+v}{2}, \frac{\frac{1}{\sigma^2} (y_i - \beta_0 - \beta_1 x_i)^2 + v}{2}\right)$$

- For *v*:

$$p(v|...) \propto \left[\prod_{i} Gamma\left(\tau_{i}|\frac{v}{2},\frac{v}{2}\right)\right] DiscUnif(v,1,30)$$

This is the only tricky part of our code. To sample v, we rely on an independent Metropolis-Hastings algorithm. This entails choosing a new v (a number from 1,...,30) and accepting/rejecting it based on our threshold probability $p_t = min\left\{1, \frac{\prod_i Gamma\left(au_i|\frac{v_new}{2}, \frac{v_new}{2}\right)}{\prod_i Gamma\left(au_i|\frac{v_old}{2}, \frac{v_old}{2}\right)}\right\}$.

- For σ^2 :

$$p\left(\frac{1}{\sigma^{2}}\big|\dots\right) \propto \left[\prod_{i} N\left(y_{i}|\beta_{0} + \beta_{1}x_{i}, \frac{\sigma^{2}}{\tau_{i}}\right)\right] Gamma\left(\frac{1}{\sigma^{2}}\big|a, b\right)$$

$$\propto \left[\prod_{i} \left(\frac{1}{\sigma^{2}}\right)^{\frac{1}{2}} \exp\left(-\frac{\tau_{i}}{2\sigma^{2}}(y_{i} - \beta_{0} - \beta_{1}x_{i})^{2}\right)\right] \left(\frac{1}{\sigma^{2}}\right)^{a-1} \exp\left(-b\frac{1}{\sigma^{2}}\right)$$

$$\propto \left(\frac{1}{\sigma^{2}}\right)^{\frac{N+2a}{2}-1} \exp\left(-\frac{1}{\sigma^{2}}\left[\frac{1}{2}\sum_{i} \tau_{i}(y_{i} - \beta_{0} - \beta_{1}x_{i})^{2} + b\right]\right)$$

$$p\left(\frac{1}{\sigma^{2}}\big|\dots\right) = Gamma\left(\frac{N+2a}{2}, \frac{1}{2}\sum_{i} \tau_{i}(y_{i} - \beta_{0} - \beta_{1}x_{i})^{2} + b\right)$$