

A superpopulation model...

Z_i = Treatment indicator

$\gamma_{i(0)}, \gamma_{i(1)}$ = Potential outcomes.

X_i = Covariate vector [e.g. age, gender, ...].

We usually assume that we observe i.i.d. samples.

$(Z_i, \gamma_{i(0)}, \gamma_{i(1)}, X_i)$ drawn from a superpopulation.

[i.e. a density over (z, y_0, y_1, x) .]

The average treatment effect (ATE) is now an expectation:

$$\tau = \mathbb{E}[\gamma_{i(1)}] - \mathbb{E}[\gamma_{i(0)}].$$

What goes wrong?

We observe only $\mathbb{E}[\gamma_{i(1)} | Z=1]$ and $\mathbb{E}[\gamma_{i(0)} | Z=0]$.

In general, these are different from $\mathbb{E}[\gamma_{i(1)}]$ and $\mathbb{E}[\gamma_{i(0)}]$ respectively, i.e. the prima facie treatment effect defined by

$$\tau_{PF} = \mathbb{E}[\gamma_{i(1)} | Z=1] - \mathbb{E}[\gamma_{i(0)} | Z=0]$$

is not the same as τ .

Can write τ as

$$\begin{aligned}\tau &= \mathbb{E}[\gamma_{i(1)}] - \mathbb{E}[\gamma_{i(0)}] \\ &= \mathbb{E}[\gamma_{i(1)} | Z=1] \cdot P(Z=1) + \mathbb{E}[\gamma_{i(0)} | Z=0] \cdot P(Z=0) \\ &\quad - (\mathbb{E}[\gamma_{i(1)} | Z=1] P(Z=1) + \mathbb{E}[\gamma_{i(0)} | Z=0] P(Z=0)).\end{aligned}$$

We do not observe the highlighted terms, so τ

is not identifiable unless we make further assumptions....

What does randomization give us?

$$Z \perp\!\!\!\perp \{Y_U, Y_O\}.$$

Remi: This is not saying that $Z \perp\!\!\!\perp Y_{obs}$.

→ This implies. $E[Y_U | Z=1] = E[Y_U | Z=0]$.

$$\Rightarrow E[Y_U] = E[Y_U | Z=1].$$

Analogously, $E[Y_O] = E[Y_O | Z=0]$.

$$\Rightarrow \tau = \tau_{PF}$$

How to identify AT&E in observational studies.

Unconfoundedness / ignorability / exchangeability:

$$\{Y_U, Y_O\} \perp\!\!\!\perp Z | X$$

Define $\tau(x) = E[Y_U | X=x] - E[Y_O | X=x]$.

$$\tau_{PR}(x) = E[Y_U | X=x, Z=1] - E[Y_O | X=x, Z=0].$$

If unconfoundedness holds, then.

$$E[Y_U | Z=1, X=x] = E[Y_U | X=x].$$

$$E[Y_O | Z=0, X=x] = E[Y_O | X=x].$$

$$\Rightarrow \tau(x) = \tau_{PR}(x).$$

If X is discrete, then.

$$\begin{aligned}\tau &= \mathbb{E}[\mathbb{E}[Y(0) - Y(1)|X]] \\ &= \mathbb{E}[\tau(X)]. \\ &= \sum_x \tau(x) \cdot P(X=x). \\ &= \sum_x \hat{\tau}_{pp}(x) \cdot P(X=x).\end{aligned}$$

We know $\hat{\tau}_{pp}(x)$ and $P(X=x)$.

\Rightarrow use this to identify τ .

When we have finitely many samples, use plug-in estimators for everything.

$$\hat{\tau}_{pp}(x) = \frac{1}{n_{1,x}} \sum_{\substack{i: X_i=x \\ Z_i=1}} Y_{i,obs} - \frac{1}{n_{0,x}} \sum_{\substack{i: X_i=x \\ Z_i=0}} Y_{i,obs},$$

$$\text{where } n_{1,x} = \#\{i : X_i=x, Z_i=1\},$$

$$n_{0,x} = \#\{i : X_i=x, Z_i=0\}.$$

$$\hat{P}(X=x) = \frac{1}{n} \sum_{i=1}^n \mathbb{I}(X_i=x)$$

$$\Rightarrow \hat{\tau} = \sum_x \hat{\tau}_{pp}(x) \cdot \hat{P}(X=x).$$

This relies on assumption that X is discrete

When X is cts, or X has many levels, then need other methods...

3 methods:

① Regression

- ② Propensity score weighting
- ③ Matching.

Back to kidney stones.

$$Z = \mathbb{1}(\text{Treatment B}).$$

$\gamma(u)$ = Recovery under treatment B.

$\gamma(o)$ = Recovery under treatment A.

$$X = \mathbb{1}(\text{small kidney stones}).$$

	Treatment A helps.	Treatment B helps.
large kidney stones.	69%	73%
small kidney stones.	87%	93%
All patients	83%	78%

Proportion of patients
with small kidney stones = 51%

$$\begin{aligned}\tau(1) &= \mathbb{E}[\gamma(u) | X=1] - \mathbb{E}[\gamma(o) | X=1]. \\ &\stackrel{\text{unconfoundedness}}{=} \underbrace{\mathbb{E}[\gamma(u) | X=1, Z=1]}_{= 0.93} - \underbrace{\mathbb{E}[\gamma(o) | X=1, Z=0]}_{= 0.87}. \\ &= 0.93 - 0.87 \\ &= 0.06.\end{aligned}$$

$$\tau(o) = 0.73 - 0.69$$

$$= 0.04.$$

$$\begin{aligned}\tau &= \tau(1) \cdot P(X=1) + \tau(0) \cdot P(X=0) \\ &= 0.06 \cdot 0.51 + 0.04 \cdot 0.49 \\ &\approx 0.05.\end{aligned}$$

Methods for estimating ATE under unconfoundedness.

Method 1: Outcome regression.

$$\text{ATE } \tau = E[\tau(x)] = \int \tau(x) \cdot p(x) dx,$$

where $p(x)$ = density of X , and.

$$\tau(x) = E[Y(1)|X=x] - E[Y(0)|X=x].$$

is called the Conditional Average Treatment Effect.
(CATE). function.

Outcome regression comprises 2 steps:

Step 1: Estimate $\tau(x)$ via $\hat{\tau}(x)$.

Step 2: Use plugin estimator for τ . i.e. we

substitute $p(x)$ with empirical distribution

$\hat{p}(x) = \text{uniform distribution on } \{x_1, \dots, x_n\}.$

$$\text{Hence, } \hat{\tau} = \frac{1}{n} \sum_{i=1}^n \hat{\tau}(x_i).$$

To estimate τ is tricky.

Many strategies, we will state 2 of them.

Strategy 1: Fit joint model for $y(x, z) = E[Y_{obs}|X=x, Z=z]$.

$$\text{observe that } \tau(x) = y(x, 1) - y(x, 0).$$

Strategy 2: Fit 2 models, one each for

- $\mu_{\text{ex}}(x) = \mathbb{E}[Y_{\text{obs}} | X=x, Z=1]$
- $\mu_{\text{eo}}(x) = \mathbb{E}[Y_{\text{obs}} | X=x, Z=0]$.

Observe that $\tau(x) = \mu_{\text{ex}}(x) - \mu_{\text{eo}}(x)$.

What can we use to estimate $\mu, \mu_{\text{ex}}, \mu_{\text{eo}}$?

- Linear model.
- GLM.
- ML methods. \rightarrow problems with overfitting
or extrapolation.

Fact: Use Strategy 1 with a linear model,

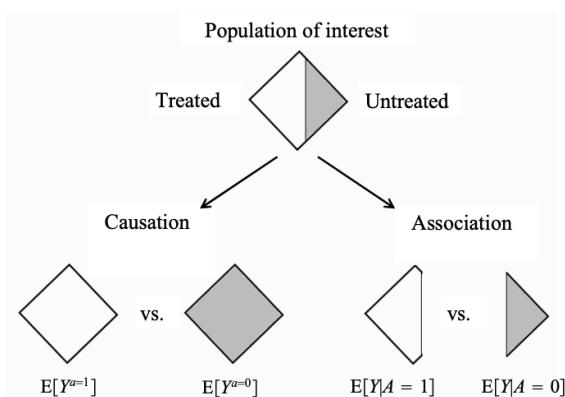
i.e. regress $Y \sim X + Z$, get is a model.

$$Y = \hat{\alpha} + \hat{\beta}^T X + \hat{\gamma} Z + \varepsilon.$$

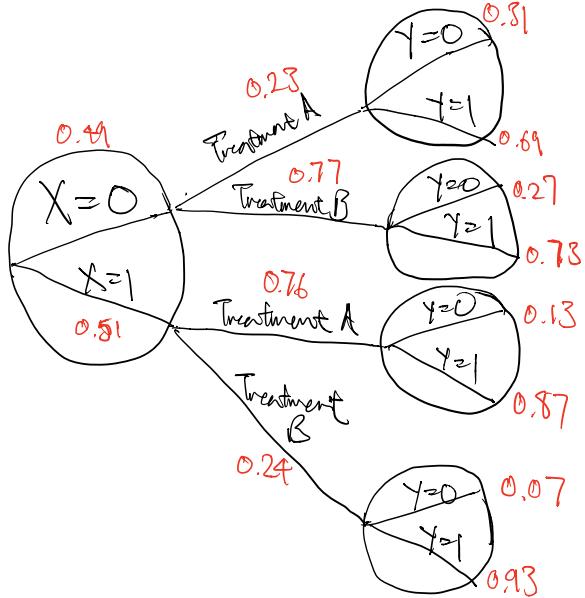
Then $\hat{\tau} = \hat{\gamma}$.

$$\begin{aligned} \text{Pf. } \hat{\tau}(x) &= \hat{\mu}_{\text{ex}}(x, 1) - \hat{\mu}_{\text{eo}}(x, 0). \quad [\hat{\mu}_{\text{ex}}(x, 1) = \hat{\alpha} + \hat{\beta}^T x + \hat{\gamma} \cdot 1] \\ &= \hat{\gamma}. \\ \Rightarrow \hat{\tau} &= \hat{\gamma}. \quad \square \end{aligned}$$

Method 2: Inverse propensity score weighting.



Observed:



See slides / lecture video for explanation of above.

Then. Assume $\{Y(0), Y(1)\} \perp\!\!\!\perp Z | X$, then.

$$\textcircled{1} \quad E\left[\frac{Y_{\text{obs}} \cdot Z}{e(X)}\right] = E[Y(1)].$$

$$\textcircled{2} \quad E\left[\frac{Y_{\text{obs}}(1-Z)}{1 - e(X)}\right] = E[Y(0)].$$

Hence, $\tau = E\left[\frac{Y_{\text{obs}} \cdot Z}{e(X)}\right] - E\left[\frac{Y_{\text{obs}}(1-Z)}{1 - e(X)}\right]$.

For finite samples, we have the inverse propensity weighting (IPW) estimator.

$$\hat{\tau}_{\text{IPW}} = \underbrace{\frac{1}{n} \sum_{i=1}^n \frac{Y_{\text{obs},i} Z_i}{\hat{e}(X_i)}}_{\text{Term 1}} - \underbrace{\frac{1}{n} \sum_{i=1}^n \frac{Y_{\text{obs},i}(1-Z_i)}{1 - \hat{e}(X_i)}}_{\text{Term 2}}$$

where $\hat{e}(x)$ is an estimator of $e(x)$.

Fact. Assume $\{Y(0), Y(1)\} \perp\!\!\!\perp Z | X$; then

$$\{Y(0), Y(1)\} \perp\!\!\!\perp Z | e(X)$$

MAP, THIS IS A MAP

Proofs of Theorem are easy.