

## A superpopulation model...

$Z_i$  = Treatment indicator

$\gamma_{i(0)}, \gamma_{i(1)}$  = Potential outcomes.

$X_i$  = Covariate vector [e.g. age, gender, ...].

We usually assume that we observe i.i.d. samples.

$(Z_i, \gamma_{i(0)}, \gamma_{i(1)}, X_i)$  drawn from a superpopulation.

[i.e. a density over  $(z, y_0, y_1, x)$ .]

The average treatment effect (ATE) is now an expectation:

$$\tau = \mathbb{E}[\gamma_{i(1)}] - \mathbb{E}[\gamma_{i(0)}].$$

What goes wrong?

We observe only  $\mathbb{E}[\gamma_{i(1)} | Z=1]$  and  $\mathbb{E}[\gamma_{i(0)} | Z=0]$ .

In general, these are different from  $\mathbb{E}[\gamma_{i(1)}]$  and  $\mathbb{E}[\gamma_{i(0)}]$  respectively, i.e. the prima facie treatment effect defined by

$$\tau_{PF} = \mathbb{E}[\gamma_{i(1)} | Z=1] - \mathbb{E}[\gamma_{i(0)} | Z=0]$$

is not the same as  $\tau$ .

Can write  $\tau$  as

$$\begin{aligned}\tau &= \mathbb{E}[\gamma_{i(1)}] - \mathbb{E}[\gamma_{i(0)}] \\ &= \mathbb{E}[\gamma_{i(1)} | Z=1] \cdot P(Z=1) + \mathbb{E}[\gamma_{i(0)} | Z=0] \cdot P(Z=0) \\ &\quad - (\mathbb{E}[\gamma_{i(1)} | Z=1] P(Z=1) + \mathbb{E}[\gamma_{i(0)} | Z=0] P(Z=0)).\end{aligned}$$

We do not observe the highlighted terms, so  $\tau$

is not identifiable unless we make further assumptions....

What does randomization give us?

$$Z \perp\!\!\!\perp \{Y_U, Y_O\}.$$

Remi: This is not saying that  $Z \perp\!\!\!\perp Y_{obs}$ .

→ This implies.  $E[Y_U | Z=1] = E[Y_U | Z=0]$ .

$$\Rightarrow E[Y_U] = E[Y_U | Z=1].$$

Analogously,  $E[Y_O] = E[Y_O | Z=0]$ .

$$\Rightarrow \tau = \tau_{PF}$$

How to identify AT&E in observational studies.

Unconfoundedness / ignorability / exchangeability:

$$\{Y_U, Y_O\} \perp\!\!\!\perp Z | X$$

Define  $\tau(x) = E[Y_U | X=x] - E[Y_O | X=x]$ .

$$\tau_{PR}(x) = E[Y_U | X=x, Z=1] - E[Y_O | X=x, Z=0].$$

If unconfoundedness holds, then.

$$E[Y_U | Z=1, X=x] = E[Y_U | X=x].$$

$$E[Y_O | Z=0, X=x] = E[Y_O | X=x].$$

$$\Rightarrow \tau(x) = \tau_{PR}(x).$$

If  $X$  is discrete, then.

$$\begin{aligned}\tau &= \mathbb{E}[\mathbb{E}[Y(0) - Y(1)|X]] \\ &= \mathbb{E}[\tau(X)]. \\ &= \sum_x \tau(x) \cdot P(X=x). \\ &= \sum_x \hat{\tau}_{pp}(x) \cdot P(X=x).\end{aligned}$$

We know  $\hat{\tau}_{pp}(x)$  and  $P(X=x)$ .

$\Rightarrow$  use this to identify  $\tau$ .

When we have finitely many samples, use plug-in estimators for everything.

$$\hat{\tau}_{pp}(x) = \frac{1}{n_{1,x}} \sum_{\substack{i: X_i=x \\ Z_i=1}} Y_{i,obs} - \frac{1}{n_{0,x}} \sum_{\substack{i: X_i=x \\ Z_i=0}} Y_{i,obs},$$

$$\text{where } n_{1,x} = \#\{i : X_i=x, Z_i=1\},$$

$$n_{0,x} = \#\{i : X_i=x, Z_i=0\}.$$

$$\hat{P}(X=x) = \frac{1}{n} \sum_{i=1}^n \mathbb{I}(X_i=x)$$

$$\Rightarrow \hat{\tau} = \sum_x \hat{\tau}_{pp}(x) \cdot \hat{P}(X=x).$$

This relies on assumption that  $X$  is discrete

When  $X$  is cts, or  $X$  has many levels, then need other methods...

3 methods:

① Regression

- ② Propensity score weighting
- ③ Matching.

Back to kidney stones.

$$Z = \mathbb{1}(\text{Treatment B}).$$

$\gamma(u)$  = Recovery under treatment B.

$\gamma(o)$  = Recovery under treatment A.

$$X = \mathbb{1}(\text{small kidney stones}).$$

	Treatment A helps.	Treatment B helps.
large kidney stones.	69%	73%
small kidney stones.	87%	93%
All patients	83%	78%

Proportion of patients  
with small kidney stones = 51%

$$\begin{aligned}\tau(1) &= \mathbb{E}[\gamma(u) | X=1] - \mathbb{E}[\gamma(o) | X=1]. \\ &\stackrel{\text{unconfoundedness}}{=} \underbrace{\mathbb{E}[\gamma(u) | X=1, Z=1]}_{= 0.93} - \underbrace{\mathbb{E}[\gamma(o) | X=1, Z=0]}_{= 0.87}. \\ &= 0.93 - 0.87 \\ &= 0.06.\end{aligned}$$

$$\tau(o) = 0.73 - 0.69$$

$$= 0.04.$$

$$\begin{aligned}\tau &= \tau(1) \cdot P(X=1) + \tau(0) \cdot P(X=0) \\ &= 0.06 \cdot 0.51 + 0.04 \cdot 0.49 \\ &\approx 0.05.\end{aligned}$$

Methods for estimating ATE under unconfoundedness.

Method 1: Outcome regression.

$$\text{ATE } \tau = E[\tau(x)] = \int \tau(x) \cdot p(x) dx,$$

where  $p(x)$  = density of  $X$ , and.

$$\tau(x) = E[Y(1)|X=x] - E[Y(0)|X=x].$$

is called the Conditional Average Treatment Effect.  
(CATE). function.

Outcome regression comprises 2 steps:

Step 1: Estimate  $\tau(x)$  via  $\hat{\tau}(x)$ .

Step 2: Use plugin estimator for  $\tau$ . i.e. we

substitute  $p(x)$  with empirical distribution

$\hat{p}(x) = \text{uniform distribution on } \{x_1, \dots, x_n\}.$

$$\text{Hence, } \hat{\tau} = \frac{1}{n} \sum_{i=1}^n \hat{\tau}(x_i).$$

To estimate  $\tau$  is tricky.

Many strategies, we will state 2 of them.

Strategy 1: Fit joint model for  $y(x, z) = E[Y_{obs}|X=x, Z=z]$ .

$$\text{observe that } \tau(x) = y(x, 1) - y(x, 0).$$

Strategy 2: Fit 2 models, one each for

- $\mu_{\text{ex}}(x) = \mathbb{E}[Y_{\text{obs}} | X=x, Z=1]$
- $\mu_{\text{eo}}(x) = \mathbb{E}[Y_{\text{obs}} | X=x, Z=0]$ .

Observe that  $\tau(x) = \mu_{\text{ex}}(x) - \mu_{\text{eo}}(x)$ .

What can we use to estimate  $\mu, \mu_{\text{ex}}, \mu_{\text{eo}}$ ?

- Linear model.
- GLM.
- ML methods.  $\rightarrow$  problems with overfitting  
or extrapolation.

Fact: Use Strategy 1 with a linear model,

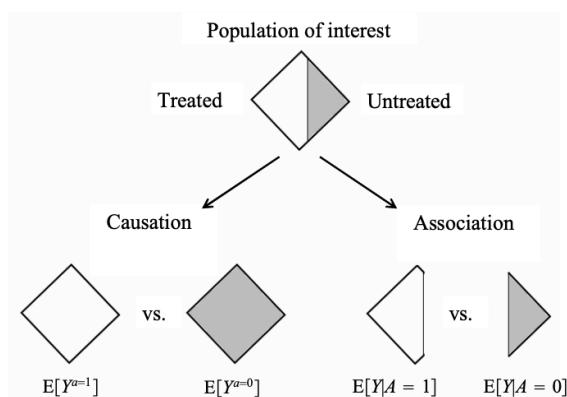
i.e. regress  $Y \sim X + Z$ , get is a model.

$$Y = \hat{\alpha} + \hat{\beta}^T X + \hat{\gamma} Z + \varepsilon.$$

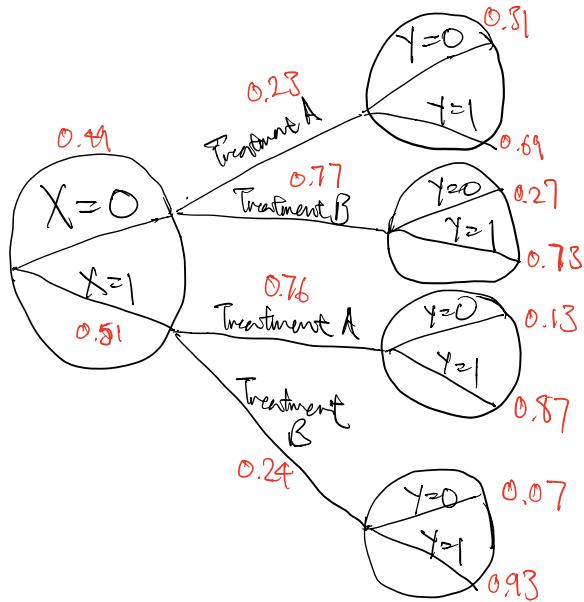
Then  $\hat{\tau} = \hat{\gamma}$ .

$$\begin{aligned} \text{Pf. } \hat{\tau}(x) &= \hat{\mu}_{\text{ex}}(x, 1) - \hat{\mu}_{\text{eo}}(x, 0). \quad [\hat{\mu}_{\text{ex}}(x, 1) = \hat{\alpha} + \hat{\beta}^T x + \hat{\gamma} \cdot 1] \\ &= \hat{\gamma}. \\ \Rightarrow \hat{\tau} &= \hat{\gamma}. \quad \square \end{aligned}$$

Method 2: Inverse propensity score weighting.



Observed:



See slides / lecture video for explanation of above.

Then. Assume  $\{Y(0), Y(1)\} \perp\!\!\!\perp Z | X$ , then.

$$\textcircled{1} \quad E\left[\frac{Y_{\text{obs}} \cdot Z}{e(X)}\right] = E[Y(1)].$$

$$\textcircled{2} \quad E\left[\frac{Y_{\text{obs}}(1-Z)}{1 - e(X)}\right] = E[Y(0)].$$

Hence,  $\tau = E\left[\frac{Y_{\text{obs}} \cdot Z}{e(X)}\right] - E\left[\frac{Y_{\text{obs}}(1-Z)}{1 - e(X)}\right]$ .

For finite samples, we have the inverse propensity weighting (IPW) estimator.

$$\hat{\tau}_{\text{IPW}} = \frac{1}{n} \sum_{i=1}^n \frac{Y_{\text{obs}} Z_i}{\hat{e}(X_i)} - \frac{1}{n} \sum_{i=1}^n \frac{Y_{\text{obs}}(1-Z_i)}{1 - \hat{e}(X_i)}$$

where  $\hat{e}(x)$  is an estimator of  $e(x)$ .

Fact. Assume  $\{Y(0), Y(1)\} \perp\!\!\!\perp Z | X$ ; then

$$\{Y(0), Y(1)\} \perp\!\!\!\perp Z | e(X)$$

MAP, THIS IS A MAP

Proofs of Theorem are easy.