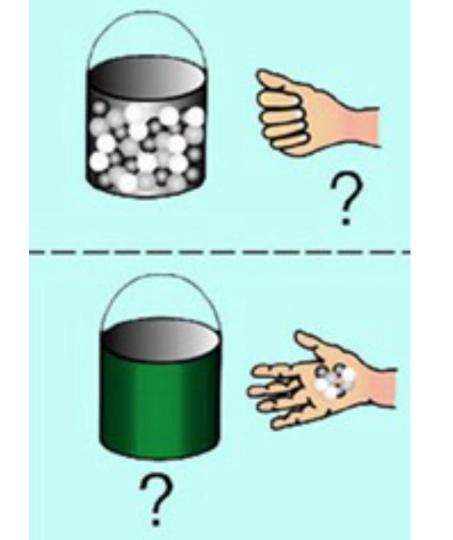


#### **Introduction to Data Science**

# Lecture 11: Statistics Point Estimation: Regression Analysis

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## Recap



#### **Probability**

 Knowing the true model, you can predict the samples outcome.

#### **Statistics**

 Knowing the generated samples, you can estimate the true model.

#### **Statistics**

Target: extract some useful information about the data.

- What you need?
  - Samples
  - A set of possible models
  - Criterion for quantifying model performance

### Maximum likelihood estimate (MLE)

- Target: estimate  $\theta$  of a model
- Samples:  $X_1$ ,  $X_2$  ...,  $X_n$
- Possible models:  $\theta \in \Theta$  (additional information)
- Model performance (probability):  $L(\theta) = P(X_1, X_2, ..., X_n | \theta)$
- Estimator:  $\hat{\theta}$  such that L( $\theta$ ) is maximized at  $\theta = \hat{\theta}$ .

#### **Formal Definition**

- Given a model with an unknown parameter  $\theta$
- Given samples:  $X_1$ ,  $X_2$  ...,  $X_n$
- The probability that the models generates the samples is called **likelihood**.  $L(\theta) = P(X_1, X_2, ..., X_n | \theta)$
- To determine the best  $\theta$ , we choose  $\hat{\theta}$  such that  $L(\theta)$  is maximized at  $\theta = \hat{\theta}$ . Maximum likelihood estimate (MLE)

#### Likelihood function

- Given a model with an unknown parameter  $\theta$
- Given samples:  $X_1$ ,  $X_2$  ...,  $X_n$

#### Continuous RV model:

- Likelihood:  $L(\theta) = \Pi_i f(X_i | \theta)$
- Log-Likelihood:  $l(\theta) = \sum_{i} \log(f(X_i | \theta))$

f: the model's probability density function (PDF)

#### Discrete RV model:

- Likelihood:  $L(\theta) = \Pi_i P(X_i | \theta)$
- Log-Likelihood:  $l(\theta) = \sum_{i} \log(P(X_i | \theta))$

P: the model's probability mass function (PMF)

#### Likelihood function

- ullet Given a model with an unknown parameter heta
- Given samples:  $X_1$ ,  $X_2$  ...,  $X_n$

#### Continuous RV model:

- Likelihood:  $L(\theta) = \prod_i f(X_i \mid \theta)$
- Log-Likelihood:  $l(\theta) = \sum_{i} \log(f(X_i | \theta))$

#### Discrete RV model:

- Likelihood:  $L(\theta) = \prod_{i} P(X_i | \theta)$
- Log-Likelihood:  $l(\theta) = \sum_{i} \log(P(X_i | \theta))$

Why do we multiply the pdfs or pmfs here?

Because we assume those samples are independent observations

### **Example: no additional information**

- For drug experiments,
  - N experiments
  - M successes

Possible models: Bernoulli distribution with probability p.

- Given one model, the probability of generating such samples:  $L(p) = p^{M}(1-p)^{N-M}$
- Most likely: the model maximizes the probability L(p)

## **Example: no additional information**

• Given one model, the probability of generating such samples:  $L(p) = p^{M} (1 - p)^{N-M}$ 

Maximize the probability



We often maximize the **logarithm of** the probability (usually easier to maximize) of generating such samples

$$l(p) = log L(p) = M * log p + (N-M) * log (1-p)$$

### **Example: no additional information**

- To maximize a continuous function I(p)
  - Find the point l'(p) = 0 and l''(p) < 0
  - $p \in [0,1]$

$$l(p) = \log L(p) = M * \log p + (N-M) * \log (1-p)$$

• 
$$l'(p) = \frac{M}{p} + \frac{N-M}{1-p} = 0$$
 p = M/N (sample cure rate)

## **Example: with additional information**

- You may have additional information.
  - $\circ$  Drug 1: p = 0.8
  - $\circ$  Drug 2: p = 0.9
  - $\circ$  Samples: M=81 successes from N = 100 experiments.
- Potential models: Bernoulli with  $p \in \{0.8, 0.9\}$

$$I = Log L = M*log p+(N-M)*log (1-p)$$

- Drug 1: -48.6539
- Drug 2: -52.2833

- **Conclusion:** drug 1 is more
- likely to be experimented.

## An Additional Example of MLE Baseball Team

• The weights for a baseball team players are {150, 143, 132, 160, 175, 190, 123, 154}

• Assume their weights are uniformly distributed over an interval [a, b]

What are good estimators for a? for b?

This example will show that the **MLE could be complicated to solve**, e.g., the equation  $l'(\theta) = 0$  may be difficult to solve, or it may **not** always be possible to use calculus methods **directly** to find the maximum of  $L(\theta)$ .

#### **MLE: Uniform**

Let X be a Uniform random variable on the interval  $[0, \theta]$ 

$$f(x;\theta) = \begin{cases} \frac{1}{\theta}, & \text{for } 0 \le x \le \theta, \\ 0, & \text{otherwise,} \end{cases} = \frac{1}{\theta} \mathbf{1}_{\{0 \le x \le \theta\}}$$

indicator function  $1_A(x)$ 

$$\mathbf{1}_{A}(x) = \begin{cases} 1, & \text{if } x \in A, \\ 0, & \text{otherwise} \end{cases}$$

The likelihood function of a random sample of size *n* is:

$$L(\theta) = \prod_{i=1}^{n} f(x_i; \theta) = \frac{1}{\theta^n} \prod_{i=1}^{n} \mathbf{1}_{\{0 \le x_i \le \theta\}} = \begin{cases} \frac{1}{\theta^n}, & \text{if } \theta \ge \max\{X_1, X_2, ..., X_n\} \\ 0, & \text{if } \theta < \max\{X_1, X_2, ..., X_n\} \end{cases}$$

$$\hat{\theta} = \max\{X_1, X_2, ..., X_n\}$$

L(a) O  $Max(x_i)$  a

Calculus methods don't work here because  $L(\theta)$  is maximized at the **discontinuity**. Clearly,  $\theta$  cannot be smaller than  $\max(x_i)$ , thus the MLE is  $\max\{X_1, X_2, ..., X_n\}$ .

#### More discussion on MLE

 Is MLE of the mean of a distribution always the sample average?

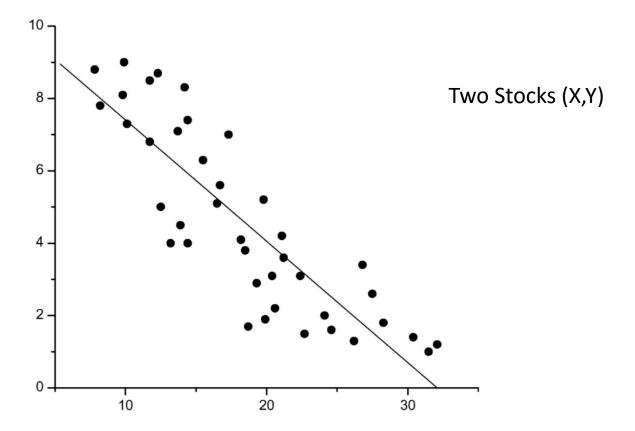
Not always!

• Ture for Normal distribution, Bernoulli distribution...

False for uniform distribution...

Linear regression

Find the relationship between X and Y

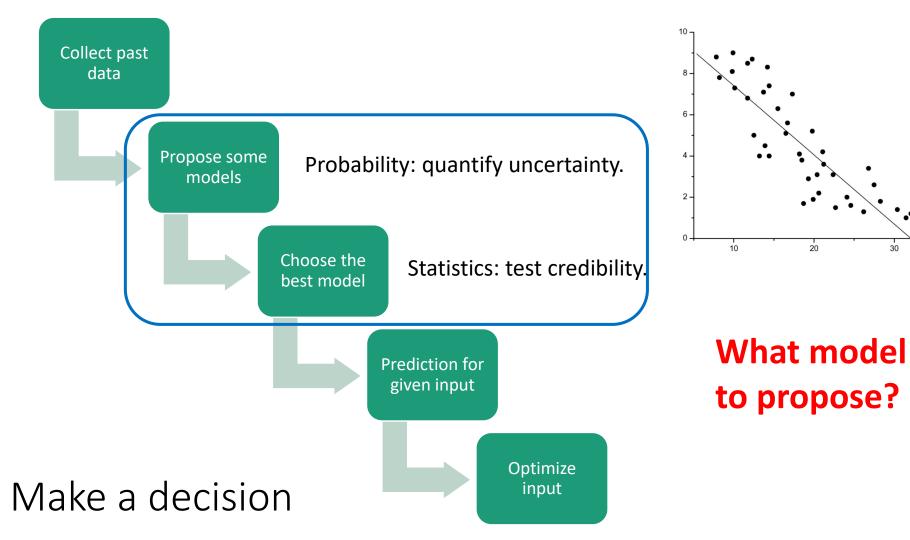


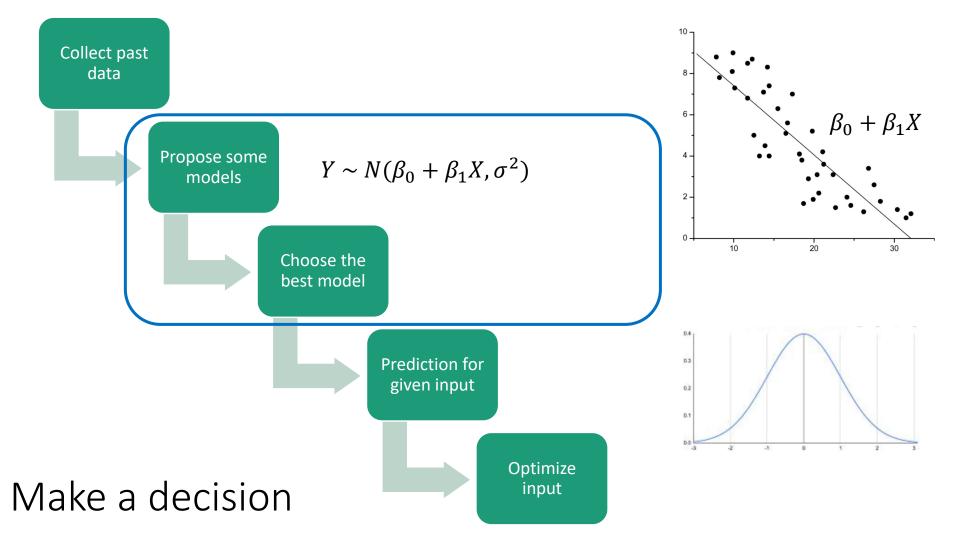
#### Correlation

- With a positive correlation (positively correlated),
  - Larger x implies larger y (vice versa)
- With zero correlation (uncorrelated),
  - No clear relationship between x and y
- With a negative correlation (negatively correlated),
  - Larger x implies smaller y (vice versa)

Negative: larger x implies smaller y.

- Question: when x increases by a certain quantity, what's the reduction in y?
- Use a line to approximate the relationship:
  - Regression analysis.

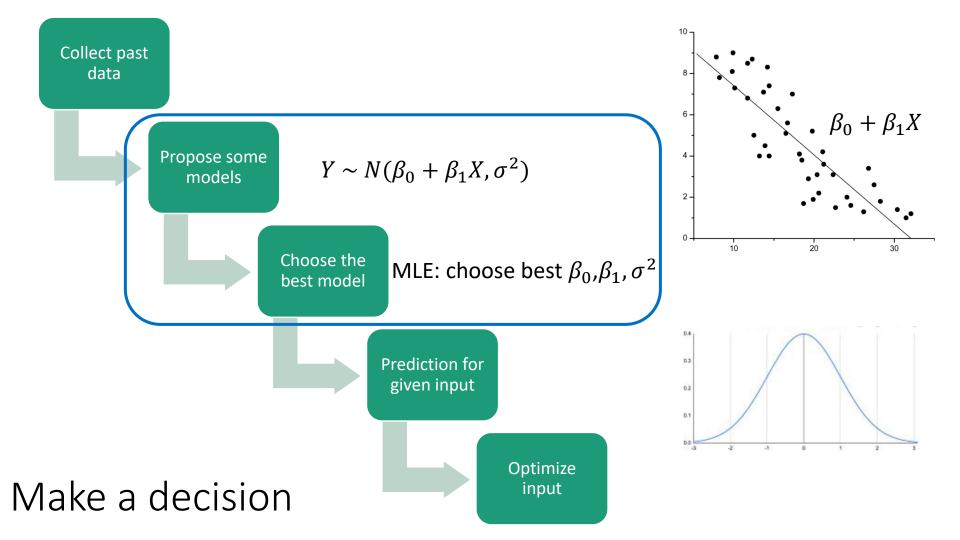




• Larger x implies smaller y.

 But when x increases by a certain quantity, what's the reduction in y?

• Regression analysis: knowing  $\beta_0 + \beta_1 X$ , you can answer.



PDF for normal

$$\frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2} \left(\frac{x-\mu}{\sigma}\right)^2\right].$$

- Samples:  $(X_1, Y_1), ..., (X_N, Y_N)$
- For the model with  $eta_0,eta_1,\sigma^2$  , the likelihood is

$$(X_N, Y_N)$$
 $S_0, \beta_1, \sigma^2$ , the likelihood is
$$\frac{1}{(\sqrt{2\pi})^n \sigma^n} \exp\left[-\frac{1}{2} \frac{\Sigma_i (Y_i) - \beta_1 X_i - \beta_0}{\sigma^2}\right]$$

 $Y \sim N(\beta_0 + \beta_1 X, \sigma^2)$ 

$$\frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right].$$

 $\Sigma_i (Y_i - \beta_1 X_i - \beta_0)^2$ 

 $Y \sim N(\beta_0 + \beta_1 X, \sigma^2)$ 

• Samples:  $(X_1, Y_1), ..., (X_N, Y_N)$ 

• For the model with  $\beta_0, \beta_1, \sigma^2$  , the likelihood is

$$\frac{1}{(\sqrt{2\pi})^n \sigma^n} \exp \left[ -\frac{1}{2} \frac{\Sigma_i (Y_i - \beta_1 X_i - \beta_0)^2}{\sigma^2} \right]$$

• Given  $\sigma^2$ , to maximize the likelihood, we only need to minimize

• Taking derivative over  $\beta_0$  and  $\beta_1$ , we have

$$\Sigma_i (Y_i - \beta_1 X_i - \beta_0) = 0$$

$$\Sigma_i (Y_i - \beta_1 X_i - \beta_0) X_i = 0$$

$$\frac{1}{\sqrt{2\pi}\sigma} \exp\left[-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right].$$

• Samples:  $(X_1, Y_1), ..., (X_N, Y_N)$ 

• For the model with 
$$\beta_0$$
,  $\beta_1$ ,  $\sigma^2$ , the likelihood is

$$\frac{1}{(\sqrt{2\pi})^n \sigma^n} \exp \left[ -\frac{1}{2} \frac{\sum_i (Y_i - \beta_1 X_i - \beta_0)^2}{\sigma^2} \right]$$

• Given  $\sigma^2$ , to maximize the likelihood, we only need to minimize

$$\Sigma_i(Y_i - \beta_1 X_i - \beta_0)^2$$

• Taking derivative over  $\beta_0$  and  $\beta_1$ , we have

$$\Sigma_i(Y_i - \beta_1 X_i - \beta_0) = 0$$

$$\Sigma_i(Y_i - \beta_1 X_i - \beta_0) X_i = 0$$

From High School: Least square regression 最小二乘法

 $Y \sim N(\beta_0 + \beta_1 X, \sigma^2)$ 

$$\Sigma_i(Y_i - \beta_1 X_i - \beta_0)X_i = 0$$
 AND  $\Sigma_i(Y_i - \beta_1 X_i - \beta_0) = 0$ 

Eliminate  $\beta_0$  first:

$$\Sigma_i(Y_i - \beta_1 X_i - \beta_0) = 0 \rightarrow \beta_0 = \frac{1}{N} \Sigma_i(Y_i - \beta_1 X_i) = \bar{Y} - \beta_1 \bar{X}$$

MLE: 
$$\widehat{\beta_1} = \frac{\sum_i (X_i - X) (Y_i - Y)}{\sum_i (X_i - \overline{X})^2}$$

$$\widehat{\beta_0} = \overline{Y} - \widehat{\beta_1} \overline{X}$$

### When simple regression is invalid?

• The model we propose is not correct.

$$Y \sim N(\beta_0 + \beta_1 X, \sigma^2) \text{ or } Y - \beta_0 - \beta_1 X \sim N(0, \sigma^2)$$

Linear regression assumes that...

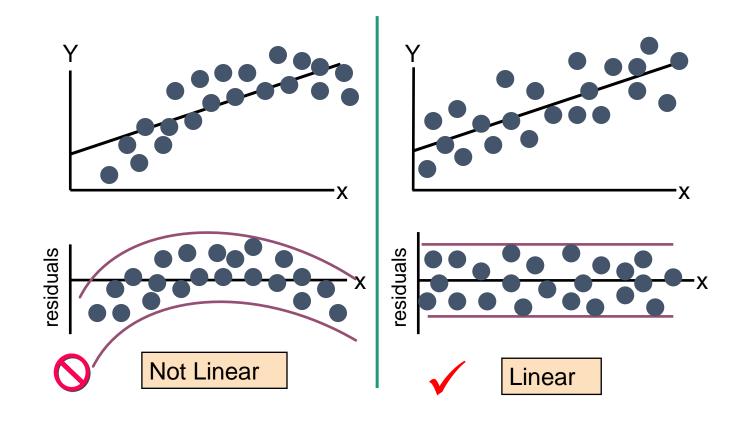
- The relationship between X and Y is linear
- 2. The variance of  $Y \beta_0 \beta_1 X$  at every value of X is the **same** (homogeneity of variances)

### Residual Analysis: check assumptions

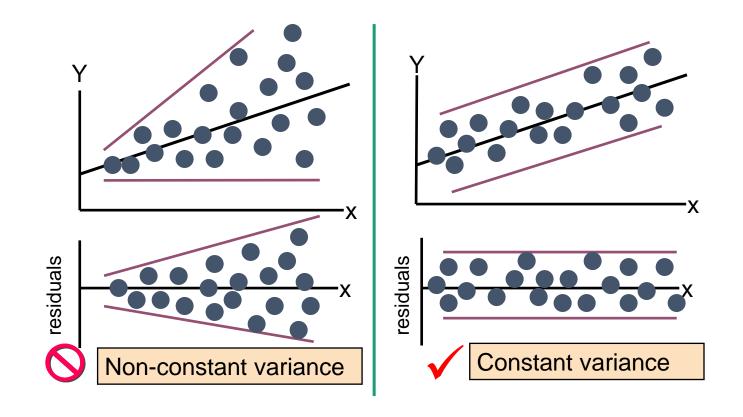
Residual: 
$$e_i := Y_i - \widehat{\beta_0} - \widehat{\beta_1} X_i$$

- Check the assumptions by examining the residuals
  - Examine for linearity assumption:
    - $e_i$  does not depend on  $X_i$
  - Evaluate constant-variance assumption:
    - variance of  $e_i$  does not depend on  $X_i$
- Graphical Analysis of Residuals: Can plot residuals vs. X

### Residual Analysis for Linearity



### Residual Analysis for constant-variance



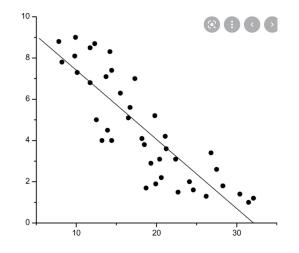
#### **Summary of Regression Analysis:**

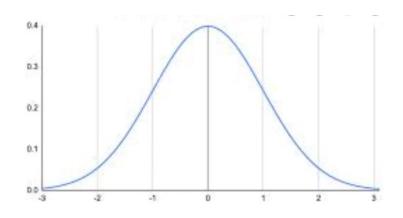
- Target: find a relationship between X and Y.
  - Proposed models:  $Y \sim N(\beta_0 + \beta_1 X, \sigma^2)$ .
  - Choose the best parameters: MLE.
  - Check whether the model is acceptable: Residual analysis.

## 1. Why we propose $Y \sim N(\beta_0 + \beta_1 X, \sigma^2)$ ?

From data, we observe that

- They are more likely to be linearly dependent with each other.
- Y is centralized at some value  $\beta_0 + \beta_1 X$ .





### 2. Why we use MLE to estimate parameters?

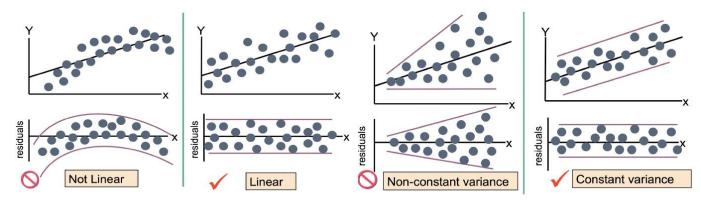
• Among the proposed models, choose the one with the **largest probability** that the samples are generated.

#### Recall example

- Drug 1: p = 0.8
- Drug 2: p = 0.9
- Samples: M = 81 successes from N=100 experiments
- Which drug do you think is experimented?

#### 3. Whether the model is acceptable?

- The model assumption may be incorrect.
  - The relationship between X and Y is linear
  - The variance of  $Y \beta_0 \beta_1 X$  at every value of X is the same (homogeneity of variances)
- Residual analysis



a)