Modern Regression

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Lecture 02.1.2 (v1.0.1)

Signposting

- ► The previous section 02.1.1 is about interpretation of Regression in general.
- ► This lecture contains the mathematical content for Modern Regression (Matrix representation).

Linear algebra view of covariance

- ► The covariance matrix of a random variable X
- ▶ Where X is an $n \times 1$ matrix, i.e. a column vector,
- ▶ has entries:

$$Cov(X)_{ij} = Cov(X_i, X_j) = \mathbb{E}[(X_i - \mu_i)(X_j - \mu_j)].$$

► The matrix form for this is:

$$\Sigma = \mathbb{E}[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T],$$

ightharpoonup Where $\mu=\mathbb{E}[X]$.

Linear algebra view of correlation

 Division by standard deviations is required to correctly generalise the scalar correlation:

$$Corr(X,Y) = \frac{\mathbb{E}[(X - \mu_X)(Y - \mu_Y)]}{\sigma_X \sigma_Y}.$$

► The matrix form for correlation is:

$$\operatorname{Corr}(X) = (\operatorname{diag}(\Sigma))^{-1/2} \Sigma (\operatorname{diag}(\Sigma))^{-1/2}$$

The matrix inversion is not computationally challenging because it is for a diagonal matrix.

Regression is analogous to linear algebra with noise

Most problems in Linear Algebra can be seen as solving a system of linear equations:

$$Ax + b = 0.$$

- ► However, data are not usually generated from a linear model.
- We therefore typically seek the least-bad fit that we can:

$$\min||Ax + b||_2^2 = \min \sum_{i=1}^{N} (Ax_i - b)^2$$

- i.e. we find A and b such that they minimise the distance (in the squared L_2 norm)
- Linear Algebra is therefore a very powerful way to view regression.

Matrix form of least squares

- \blacktriangleright Consider data X' with p' features (columns) and n observations.
- Given the regression problem:

$$\mathbf{y} = \mathbf{X}'\beta' + \mathbf{b} + \mathbf{e}$$

- to find β' (a matrix dimension $p' \times 1$))
- ▶ and b to minimise 'error':

Matrix form of least squares

We construct a simpler representation by adding a constant feature:

$$\mathbf{X} = \begin{bmatrix} 1 & \mathbf{X}_{11} & \cdots & \mathbf{X}_{1p'} \\ & & \cdots & \\ 1 & \mathbf{X}_{n1} & \cdots & \mathbf{X}_{np'} \end{bmatrix}$$

- which has p = p' + 1 features.
- ▶ We now solve the analogous equation:

$$y = X\beta + e$$

which has the same solution but is in a more convenient form.

Mean Squared Error (MSE)

► The prediction error is:

$$\mathbf{e}(\beta) = \mathbf{y} - \mathbf{X}\beta$$

And the estimation error can be written:

$$MSE(\beta) = \frac{1}{n} \mathbf{e}^T \mathbf{e}$$

Minimising MSE

▶ Taking (vector) derivatives with respect to β :

$$\nabla \text{MSE}(\beta) = \frac{1}{n} (\nabla \mathbf{y}^T \mathbf{y} - 2\nabla \beta^T x^T \mathbf{y} + \nabla \beta^T \mathbf{X}^T \mathbf{X} \beta) \quad \text{(I)}$$

$$= \frac{1}{n} (0 - 2x^T \mathbf{y} + 2\mathbf{X}^T \mathbf{X} \beta) \quad \text{(2)}$$

• which is zero at the optimum $\hat{\beta}$:

$$\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} - \mathbf{X}^T \mathbf{y} = 0$$

with the solution:

$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}.$$

Exercise: For the case p'=1, check that this solution is the same as you can find in regular linear algebra textbooks.

The Hat Matrix

➤ There is an important and response independent quantity hidden in the prediction:

$$\mathbf{H} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$$

► The fitted values are:

$$\hat{y} = X\hat{\beta} = X(X^TX)^{-1}X^Ty = Hy$$

- ightharpoonup H is dimension $N \times N$
- lackbox H "projects" y into the fitted value space \hat{y}

Properties of the Hat Matrix

- ▶ Influence: $\frac{\partial \hat{y}_i}{\partial y_j} = H_{ij}$. So H controls how much a change in one observation changes the estimates of each other point.
- **symmetry**: $H^T = H$. So influence is symmetric.
- ▶ **Idempotency**: $H^2 = H$. So the predicted value for any projected point is the predicted value itself.
- ► You should read up on these and other vector algebra properties.

Residuals and the Hat Matrix

► The residuals can be written:

$$e = y - Hy = (I - H)y$$

- ► I H is also symmetric and idempotent, and can also be interpreted in terms of Influence.
- ▶ Because of this,

$$MSE(\hat{\beta}) = \frac{1}{n} \mathbf{y}^T (1 - \mathbf{H})^T (1 - \mathbf{H}) \mathbf{y} = \frac{1}{n} \mathbf{y}^T (1 - \mathbf{H}) \mathbf{y}$$

Expectations

If the data were generated by our model(!) then they are described by a random variable Y:

$$\mathbf{Y} = \mathbf{x}\beta + \epsilon$$

- lacktriangle where ϵ is an n imes 1 matrix of RVs with mean $oldsymbol{0}$ and covariance $\sigma^s {
 m I.}$
- From this it is straightforward to show that the fitted values are unbiased:

$$\mathbb{E}[\hat{y}] = \mathbb{E}[HY] = \mathbf{x}\beta$$

using the properties of Expectations with the symmetry and idempotency of H.

Covariance

► Similarly, it is straightforward to show that

$$Var[\hat{y}] = \sigma^2 H$$

using the properties of Variances with the symmetry and idempotency of $\boldsymbol{H}.$

Reflection

- By the end of the course, you should:
 - Be able to define correlation and regression in multivariate context
 - ▶ Be able to perform basic calculations using these concepts
 - Be able to extend intuition about their application.
- ► This is something worth reading up on
 - ► You should really understand univariate regression

Signposting

- ► The mathematics behind Modern Regression is entirely analogous to the mathematics of a huge range of advanced, scalable Machine Learning tools.
- ▶ This is one of the places you should do your homework.

References

There is a lot more technical detail in Cosma Shalizi's course on **Modern** Regression:

Modern Regression, by Coasma Shalizi http://www.stat.cmu.edu/~cshalizi/mreg/15/lectures/