

Midterm: Theory of Linear Models

Fall 2024

Due: Monday, 10/21 at 10:00am, hard-copy handed-in

Directions:

- Type your solutions in RMarkdown or latex
- Everyone must work independently: no consulting with other students, professors, or posting questions online
- You may use any textbook and your class notes for reference.
- **No ChatGPT**
- Along with your exam, provide a list of references that you used to complete the problems

1. Consider the variance components model $Y = g + \varepsilon$, where Y is an $n \times 1$ response vector, g is a $n \times 1$ random effects vector with distribution $g \sim \mathcal{N}(0, \sigma_g^2 K)$, where K is an $n \times n$ symmetric positive definite matrix; assume $\varepsilon \sim \mathcal{N}(0, \sigma^2 I_n)$ and that g and ε are independent.
 - a. What is $\text{Var}(Y)$?
 - b. Compute the maximum likelihood estimate for σ_g^2 .
 - c. Suppose X is an $n \times p$ matrix of covariates such that $K = XX^T$. Provide an interpretation of the elements of K with respect to the covariates measured in X .
 - d. Provide an interpretation of the quantity $\frac{\sigma_g^2}{\sigma_g^2 + \sigma^2}$.
 - e. Compute the score test statistic for hypothesis $H_0 : \sigma_g^2 = 0$.
 - f. Consider now the model $Y = X\beta + \eta$, $\eta \sim \mathcal{N}(0, \sigma^2 I_n)$. Compute the score test statistic for $H_0 : \beta = 0$.
 - g. Are the two score test null hypotheses always equivalent? Why or why not?
 - h. Assuming the null hypotheses are equivalent, identify a situation for which the two tests will have different power.
2. A random variable $X = \sum_{i=1}^k Z_i^2$, for Z_i independent standard normal random variables, follows a *Chi-Squared* distribution with k degrees of freedom. Show that $Y^T \Sigma^{-1} Y \sim \chi_k^2$ for $Y \sim \mathcal{N}(0, \Sigma)$. Show your work and carefully explain your reasoning for full credit.
3. Let P_X and P_W be the projection matrices corresponding respectively to the column spaces of matrices X and W , respectively. Assume X is $n \times r$ with $\text{rank}(X) = r$, and W is $n \times q$ with $\text{rank}(W) = q$. Under what conditions will the projections commute? I.e., when is $P_X P_W = P_W P_X$ true?
4. Consider the regression model $Y = X\beta + \varepsilon$ with $\varepsilon \sim \mathcal{N}(0, \Sigma)$, where Σ is a positive definite covariance matrix that is otherwise unspecified.
 - a. Verify that the generalized least squares estimator $\hat{\beta}_G = (X^T \Sigma^{-1} X)^{-1} X^T \Sigma^{-1/2} Y$ is unbiased.
 - b. Compute the variance of $\hat{\beta}_G$.
 - c. In this setting, calculate the bias of the usual OLS estimator $\hat{\beta} = (X^T X)^{-1} X^T Y$.
 - d. Compute the variance of $\hat{\beta}$.
 - e. What is the statistical advantage of using $\hat{\beta}_G$ instead of $\hat{\beta}$ when we know that $\Sigma \neq \sigma^2 I_n$? Explain your answer in rigorous statistical terms referring to concepts such as power, variance, mean squared error, bias, etc.
5. Consider the regression model $Y = X\beta + \varepsilon$, with $\mathbb{E}[\varepsilon] = 0$. Assume X is full column rank so that

$$\hat{Y} = X(X^T X)^{-1} X^T Y.$$

- a. Show that $\text{Cov}(\hat{Y}, (Y - \hat{Y})) = 0$.
 - b. Show that if $\varepsilon \sim \mathcal{N}(0, \sigma^2 I_n)$, then \hat{Y} and $Y - \hat{Y}$ are independent.
 - c. Prove or Disprove: if $\varepsilon \sim \mathcal{N}(0, \Sigma)$, then \hat{Y} and $Y - \hat{Y}$ are independent.
6. Let $Y_i = \beta \cdot i + \varepsilon_i, i = 1, \dots, n, \mathbb{E}\varepsilon_i = 0, \text{Var}(\varepsilon_i) = i^2 \sigma^2$.
- a. Compute the usual Least Squares Estimate for β .
 - b. Compute the best linear unbiased estimator for β .
 - c. What is the difference in variance for the estimators from parts (a) and (b)?
7. Consider the model $Y_{ij} = \alpha_i + \beta_j + \varepsilon_{ij}$, where $i = 1, \dots, a; j = 1, \dots, b$, and $\varepsilon_{ij} \stackrel{iid}{\sim} (0, \sigma^2)$.
- a. Clearly describe the structure of the design matrix for this model. That is, we wish to write the model in matrix form as $Y = X\gamma + \varepsilon$ for some matrix X .
 - b. Recall that a function $\sum c_i \alpha_i + \sum d_j \beta_j$ is estimable iff the linear contrast vector $u^T = (c_1, \dots, c_a, d_1, \dots, d_b)$ is in $\mathcal{C}(X^T)$. Identify a condition of $c_1, \dots, c_a, d_1, \dots, d_b$ that is necessary and sufficient for the estimability of $u^T \gamma$.