

Markov Chain Monte Carlo

Dustin Lang
Perimeter Institute for Theoretical Physics

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Borrowing heavily from Dan Foreman-Mackey's slides
<https://speakerdeck.com/dfm/data-analysis-with-mcmc>

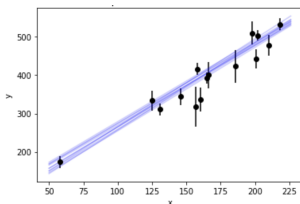
These slides are available at
<https://github.com/dstndstn/MCMC-talk>

Context - Generative data analysis

- ▶ In Astrophysics, often have a *generative model* of a phenomenon of interest:
- ▶ We can *simulate* or *predict* what we will observe with a *model*
- ▶ The model has *parameters*, and we want to put *constraints* on these parameters:
- ▶ *What range of parameters are consistent with our observations?*

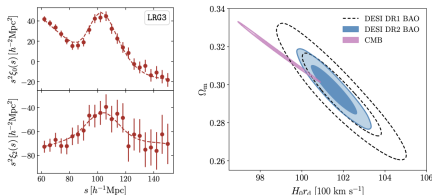
Context - Generative data analysis

- ▶ We can *simulate* or *predict* what we will observe with a *parameterized model*: $\hat{y}_i = f(\theta, x_i)$
- ▶ We can write down a *likelihood* of getting the observations we got: $\mathcal{L}(y_i|x_i, \theta) = g(y_i, x_i, \theta)$
- ▶ Eg, Gaussian likelihoods:
$$\mathcal{L}(y_i|x_i, \theta) = \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(y_i - f(\theta, x_i))^2}{2\sigma_i^2}\right)$$
- ▶ *Note*: after we have made our observations, y_i are fixed, as are x_i , σ_i (and f and g); only parameters θ are free
- ▶ We want to know about constraints on our *parameters* θ !
- ▶ *Bayes' theorem* lets us do this: $p(\theta|y_i, x_i) = \frac{\mathcal{L}(y_i|x_i, \theta)p(\theta)}{p(y_i)}$



Context - Generative data analysis

- ▶ *Bayes' theorem* lets us do this:
$$p(\theta|\{y_i\}, \{x_i\}) = \frac{\mathcal{L}(\{y_i\}|\{x_i\}, \theta)p(\theta)}{p(\{y_i\})}$$
- ▶ *Usually* we don't care about the $p(\{y_i\})$ (“evidence”) term
- ▶ *Often* we want the $p(\theta)$ (“prior”) term to be “uninformative” (that's not trivial)
- ▶ So, we usually want to *draw samples* from the *posterior distribution* $p(\theta|\{y_i\}, \{x_i\})$; that's what we report in our paper
- ▶ (given samples, we can draw *contours* of the probability distribution):



MCMC

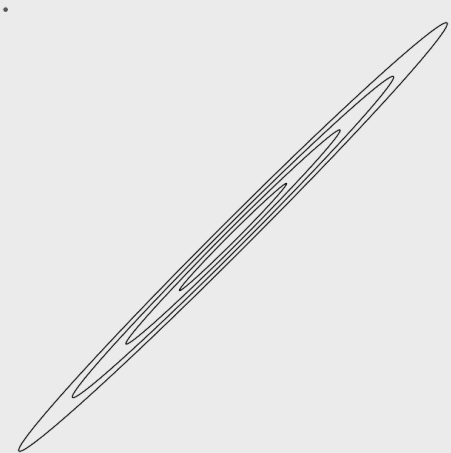
draws samples from a probability function

and all you need to be able to do is

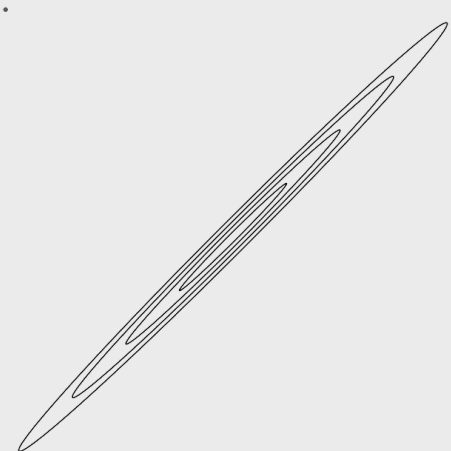
evaluate

the function

(up to a constant)

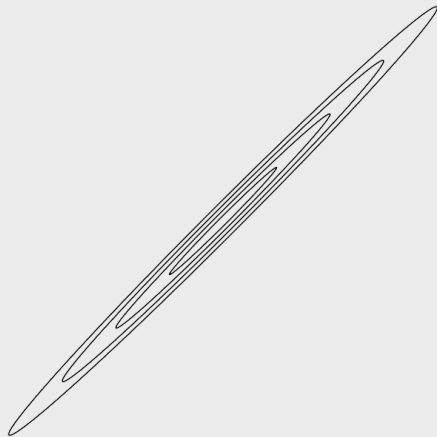


Metropolis-Hastings

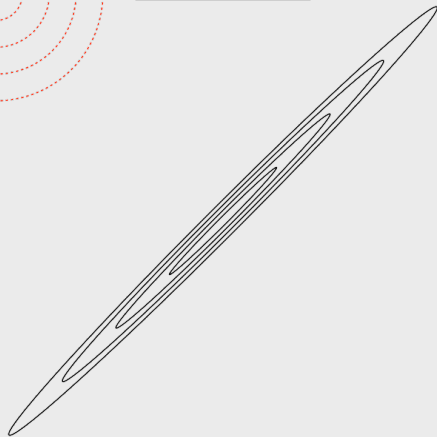
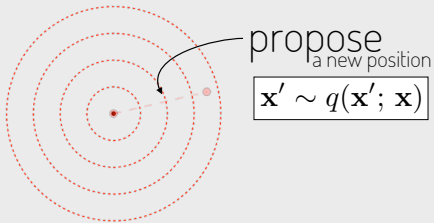


Metropolis-Hastings
in an ideal world

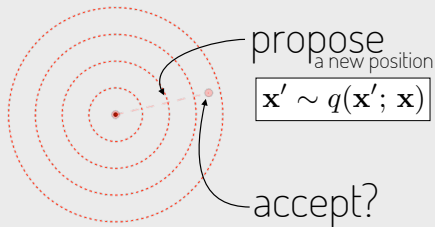
start here
perhaps



Metropolis-Hastings
in an ideal world

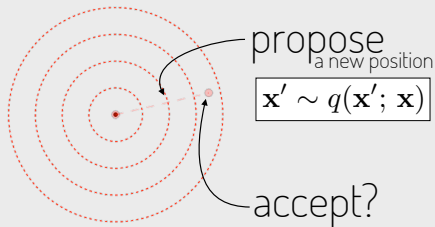


Metropolis-Hastings
in an ideal world



$$p(\text{accept}) = \min \left(1, \frac{p(\mathbf{x}')}{p(\mathbf{x})} \frac{q(\mathbf{x}; \mathbf{x}')}{q(\mathbf{x}'; \mathbf{x})} \right)$$

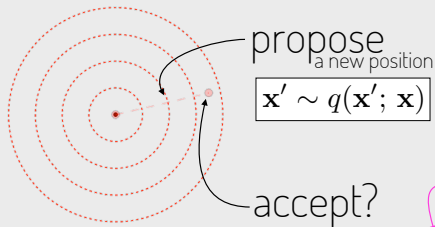
Metropolis-Hastings
in an ideal world



$$p(\text{accept}) = \min \left(1, \frac{p(\mathbf{x}')}{p(\mathbf{x})} \frac{q(\mathbf{x}; \mathbf{x}')}{q(\mathbf{x}'; \mathbf{x})} \right)$$

definitely.

Metropolis–Hastings
in an ideal world



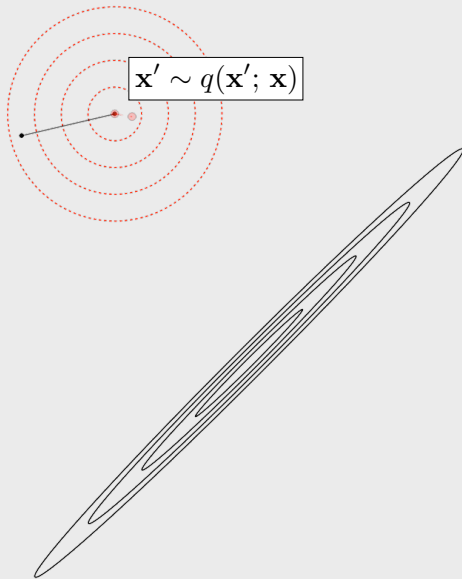
$$\mathbf{x}' \sim q(\mathbf{x}'; \mathbf{x})$$

only **relative** probabilities

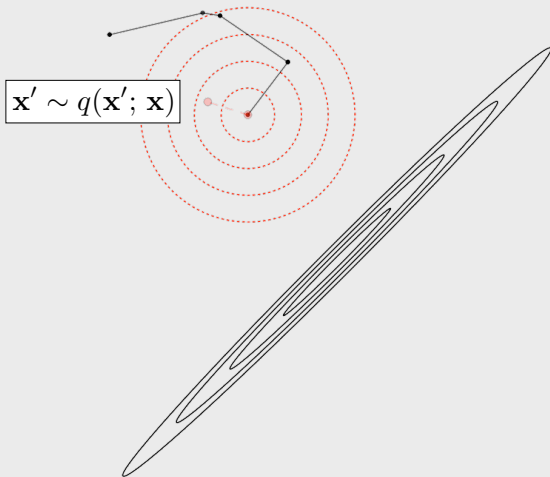
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Metropolis-Hastings
in an ideal world

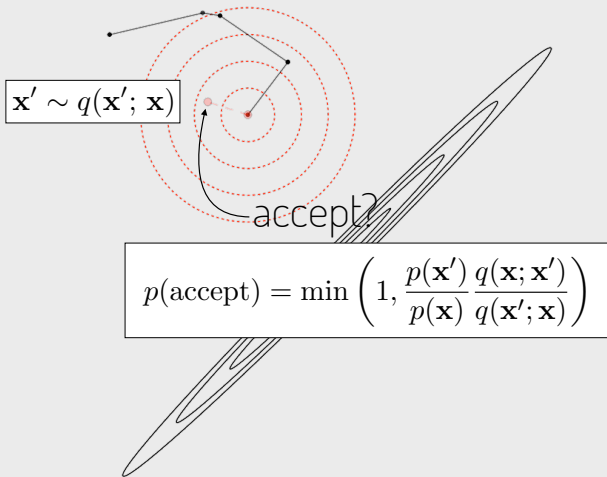


Metropolis–Hastings
in an ideal world



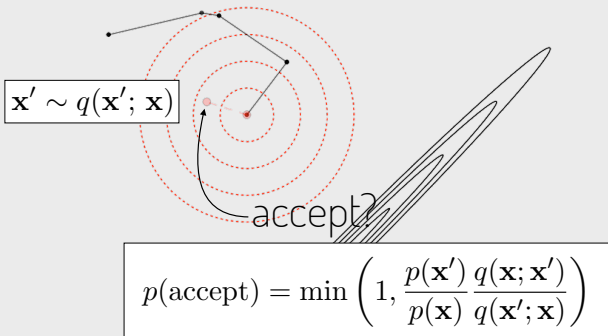
Metropolis–Hastings

in an ideal world



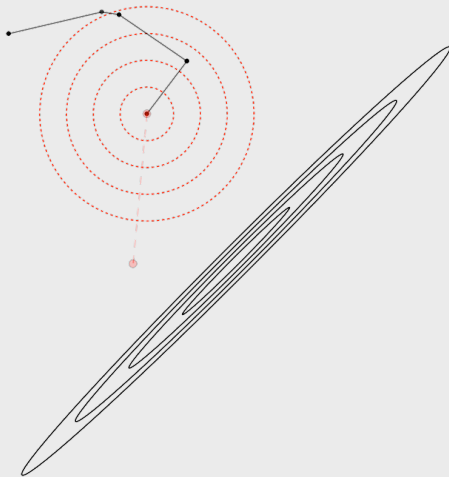
Metropolis-Hastings

in an ideal world

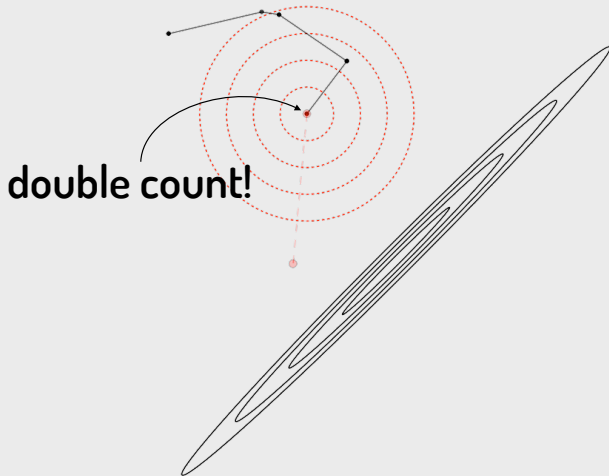


not this time.

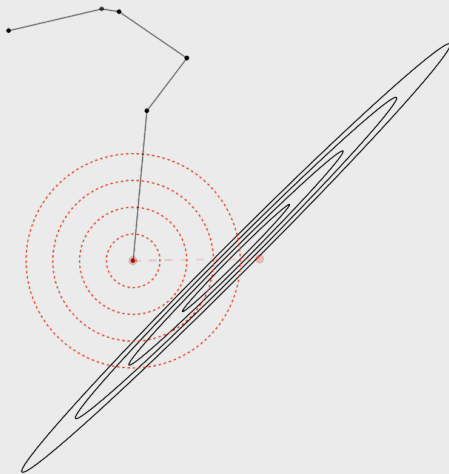
Metropolis-Hastings
in an ideal world



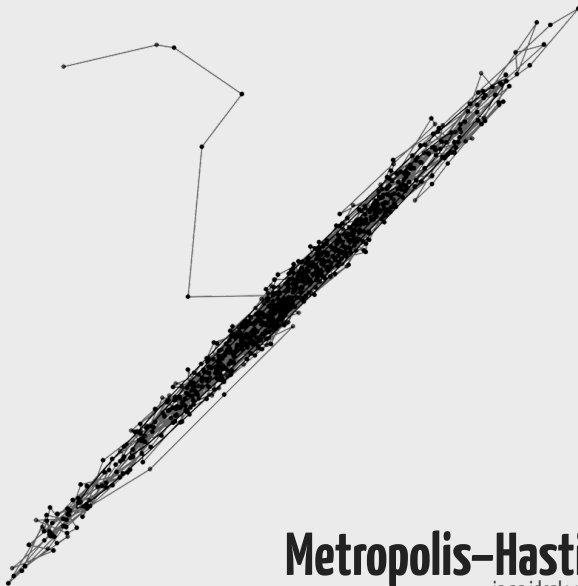
Metropolis–Hastings
in an ideal world



Metropolis-Hastings
in an ideal world



Metropolis-Hastings
in an ideal world



Metropolis–Hastings
in an ideal world

About the name

- ▶ **Monte Carlo**: a reference to the famous Monte Carlo Casino in Monaco, alluding to the randomness used in the algorithm
- ▶ **Markov Chain**: a list of samples, where each one is generated by a process that only looks at the previous one.
- ▶ **Markov**: a 19th-century Russian mathematician and impressive-moustache-haver with an [extensive list of things named after him](#)
- ▶ **Metropolis–Hastings**: lead authors of 1953 and 1970 papers (resp.) giving the algorithm with symmetric and general proposal distributions (resp.)

The Algorithm (1)

```
function mcmc(prob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    prob = prob_func(p)
    chain = []
    for i in 1:nsteps
        # propose a new position in parameter space
        # ...
        # compute probability at new position
        # ...
        # decide whether to jump to the new position
        # ...
        if # ...
            # ...
            # ...
        end
        # save the position
        append!(chain, p)
    end
    return chain
end
```

The Algorithm (2)

```
function mcmc(prob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    prob = prob_func(p)
    chain = []
    for i in 1:nsteps
        # propose a new position in parameter space
        p_new = propose_func(p)
        # compute probability at new position
        prob_new = prob_func(p_new)
        # decide whether to jump to the new position
        ratio = prob_new / prob
        if ratio > 1 or ratio > uniform_random()
            p = p_new
            prob = prob_new
        end
        # save the position
        append!(chain, p)
    end
    return chain
end
```

The Algorithm (3)

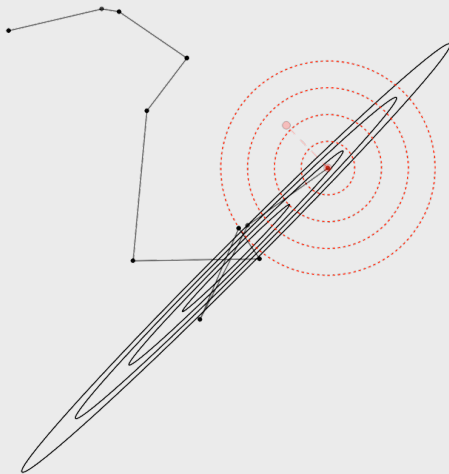
```
function mcmc(logprob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    logprob = logprob_func(p)
    chain = []
    for i in 1:nsteps
        # propose a new position in parameter space
        p_new = propose_func(p)
        # compute probability at new position
        logprob_new = logprob_func(p_new)
        # decide whether to jump to the new position
        ratio = exp(logprob_new - logprob)
        if ratio > 1 or ratio > uniform_random()
            p = p_new
            logprob = logprob_new
        end
        # save the position
        append!(chain, p)
    end
    return chain
end
```


The Algorithm (4)

```
function mcmc(logprob_func, propose_func, initial_pos, nsteps)
    p = initial_pos
    logprob = logprob_func(p)
    chain = []
    naccept = 0
    for i in 1:nsteps
        # propose a new position in parameter space
        p_new = propose_func(p)
        # compute probability at new position
        logprob_new = logprob_func(p_new)
        # decide whether to jump to the new position
        if exp(prob_new - prob) > uniform_random()
            p = p_new
            logprob = logprob_new
            naccept += 1
        end
        # save the position
        append!(chain, p)
    end
    return chain, naccept/nsteps
end
```

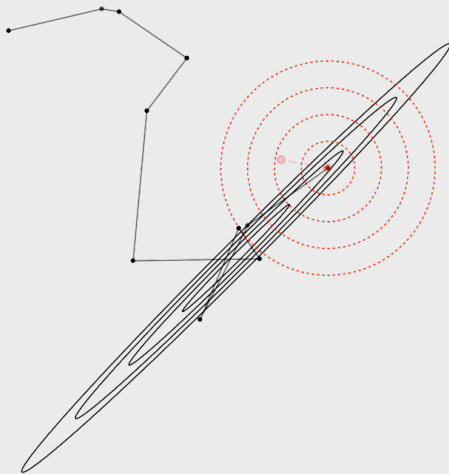
Practicalities

- ▶ How do I choose a proposal distribution?
- ▶ How many steps do I have to take?



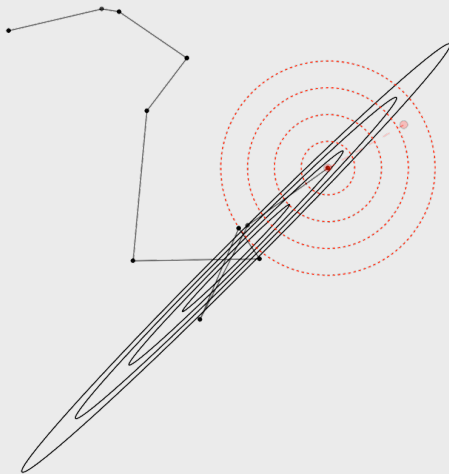
Metropolis-Hastings

in the real world

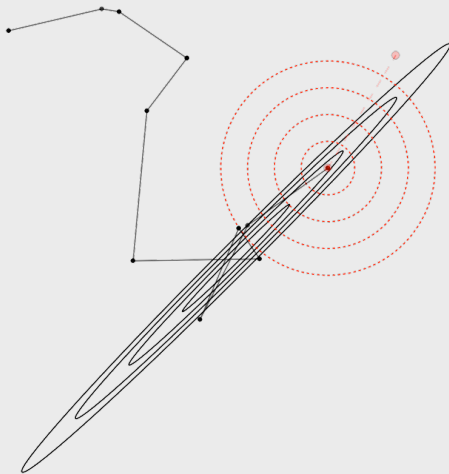


Metropolis-Hastings

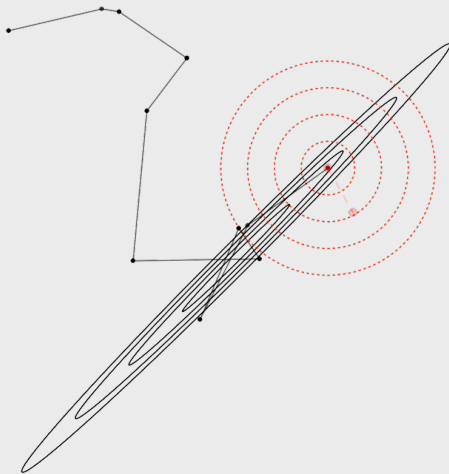
in the real world



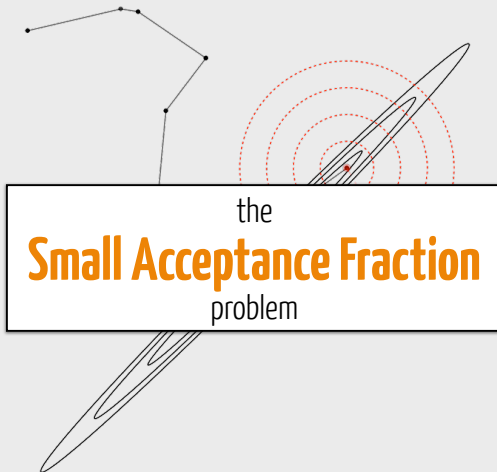
Metropolis-Hastings
in the real world



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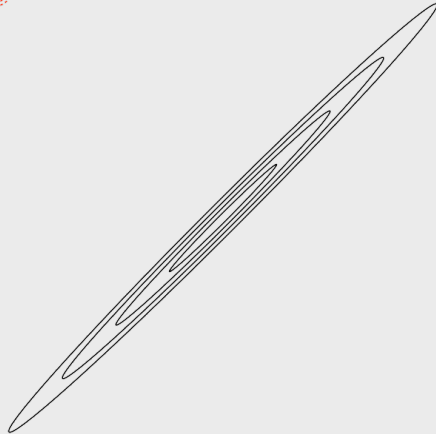


Metropolis–Hastings
in the real world



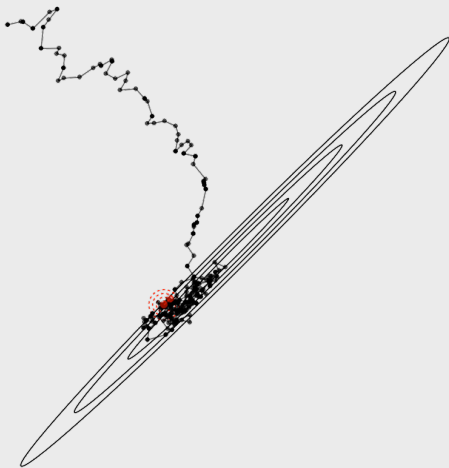
the
Small Acceptance Fraction
problem

Metropolis–Hastings
in the real world



Metropolis-Hastings

in the real world

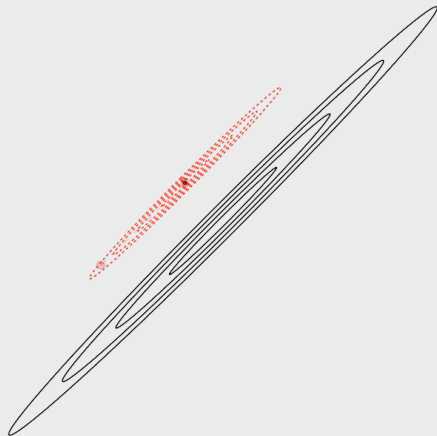


Metropolis-Hastings

in the real world

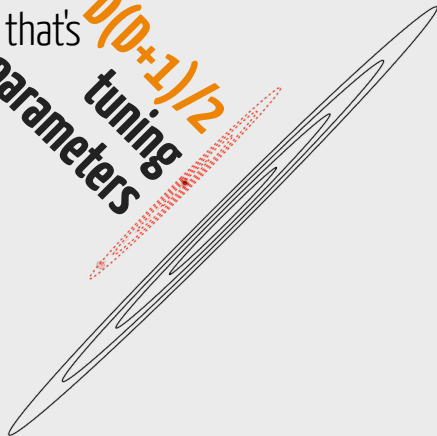


Metropolis–Hastings
in the real world



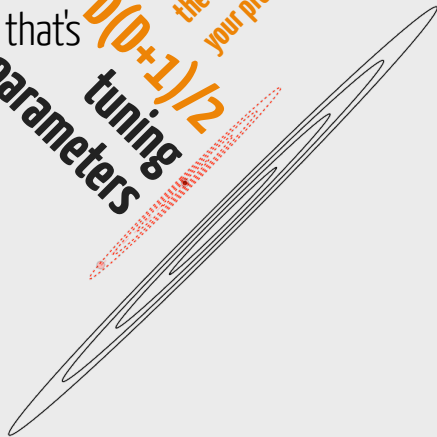
Metropolis–Hastings
in the real world

that's $O(D+1)/2$
tuning
parameters

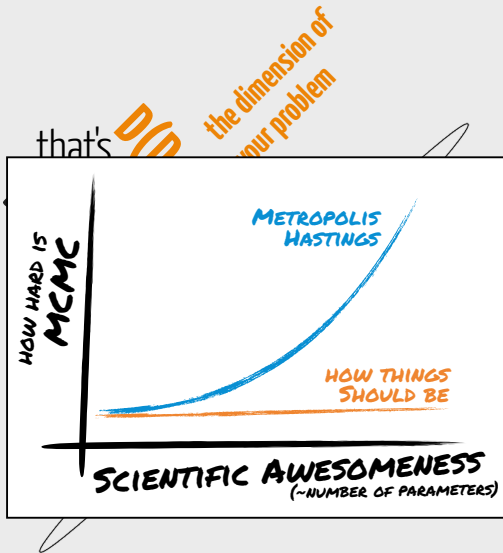


Metropolis–Hastings
in the real world

that's $D(D+1)/2$ tuning parameters
the dimension of your problem



Metropolis–Hastings
in the real world



Metropolis-Hastings

in the real world

How many samples do I need?

- ▶ Burn-in — skip the first N samples
- ▶ *Has my chain converged?*
- ▶ MCMC produces **correlated** samples, so
 - ▶ How correlated are my samples?
 - ▶ Can measure the *autocorrelation time* τ
 - ▶ Keep $1/\tau$ of the MCMC samples
 - ▶ eg <https://github.com/dfm/acor>
 - ▶ How many uncorrelated samples do I need?
 - ▶ No easy general answer to this question!
 - ▶ “How many can you afford?”

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Conclusions

- ▶ MCMC remains an essential tool for probabilistic inference
- ▶ For science: lets us constrain model parameters based on data (Bayesian inference)
- ▶ Beguilingly simple algorithm, but difficult practicalities
- ▶ MCMC has beautiful theoretical guarantees... as compute time $\rightarrow \infty$