

# Stats 150, Fall 2019

## Lecture 5, Thurs, 9/12/2019

### Topics Today:

- §1.4 - 1.8 of text
- The remaining of readings are assigned outside of lecture.

## 1 Key Points for Homework

Pitman gives a few key pointers (which are from the textbook) that may help with finishing the homework due tonight.

(1) Recall the definition of an irreducible chain. That is,

$$\forall x, y \in S, \exists_n : P^n(x, y) > 0.$$

This forbids a random walk on a graph with 2 or more components (closed classes). Most of the chains we commonly deal with (and in our homework) are irreducible.

(2) Fact: (See text). If  $P$  is irreducible and if there is a stationary probability vector  $\pi$  for  $P$  (that is, we can solve  $\pi P = \pi$  where  $\sum_x \pi_x = 1, \pi_x \geq 0$ ), then all the states are recurrent (the chain is recurrent). (Theorem 1.7 in Durrett).

### Definition: Positive Recurrent -

We say that the chain is **positive recurrent** when, for some or for all  $x$ :

$$\mathbb{E}_x T_x < \infty$$

which is:

$$\mathbb{E}_x T_x = \sum_{n=1}^{\infty} \mathbb{P}_x(T_x \geq n).$$

We should check that if  $E_x T_x < \infty$  for some  $x$  and  $P$  is irreducible, then

$$E_x T_x < \infty, \quad \forall x.$$

### Definition: Null recurrent -

If a state is recurrent but not positive recurrent (for example  $P_x(T_x < \infty) = 1$  but  $E_x T_x = \infty$ ), then we say that  $x$  is **null recurrent**.

## 2 Review

Pitman reminds us that there is a formula relating the mean return time and the stationary probability (Theorem 1.21 Durrett):

$$\pi_x = \frac{1}{\mathbb{E}_x T_x}$$

As a simple corollary, this formula directly implies that  $\pi$  is unique. There is no doubt about this for a stationary measure in terms of the mean recurrence time. If we discuss a system of countably infinite space, our traditional linear algebra may fail. This result provides an interpretation beyond a system of finitely many equations and unknowns.

**Remark:** Conversely, if  $P$  is irreducible and positive recurrent, then there exists this  $\pi$ . This is almost trivial, but of course we have to check that  $\pi$  is a stationary probability.

### 3 Example of Null Recurrence

**Example:** Consider a simple (symmetric) random walk with equal probability of going either direction on  $\mathbb{Z} = \{\dots, -2, -1, 0, 1, 2, \dots\}$ . We take the usual notation  $S_n$  for the walk.

Start at  $x = 0$ , so that  $S_n := \Delta_1 + \Delta_2 + \dots + \Delta_n$ , where  $\Delta_k$  has the value  $+1$  with probability  $\frac{1}{2}$  and  $-1$  with probability  $\frac{1}{2}$ . Now this gives:

$$P^n(0, 0) = \begin{cases} 0, & \text{if } n \text{ is odd} \\ \binom{2m}{m} \left(\frac{1}{2}\right)^{2m}, & \text{if } n = 2m \text{ is even} \end{cases}.$$

Now Pitman notes we can tell recurrence or transience by looking at the fact that the total number of visits to 0 follows a geometric distribution with  $(1 - \rho_0)$ :

$$\mathbb{E}_0(\text{total \# visits to } 0) = \sum_{n=1}^{\infty} P^n(0, 0)$$

But we know that  $\binom{2m}{m} \left(\frac{1}{2}\right)^{2m}$  is the same as the probability of  $m$  heads and  $m$  tails in  $2m$  tosses. Increasing tosses gives a very ‘flat’ normal curve because the mean of  $\mathbb{E}_0 S_{2m} = 0$ . Now because the variance of each summed term is 1, the mean square is:

$$\mathbb{E}_0 S_{2m}^2 = \underbrace{1 + 1 + \dots + 1}_{2m} = 2m.$$

We call this “diffusion”, in that on average the center of our distribution goes nowhere, but the distribution spreads out and flattens.

Using Stirling’s formula (or the Normal Approximation), that is,

$$n! \sim \left(\frac{n}{e}\right)^n \sqrt{2\pi n},$$

and apply this to our earlier expression to show that:

$$P^{2m}(0, 0) \sim \frac{C}{\sqrt{m}},$$

where  $C$  is some constant.

To see recurrence versus transience, we look at, from earlier,

$$\sum_{n=1}^{\infty} P^n(0, 0) = \sum_{m=1}^{\infty} P^{2m}(0, 0) \sim \sum_{m=1}^{\infty} \frac{C}{\sqrt{m}} = \infty$$

A rather paradoxical fact: This implies that the expected return time to 0 is infinite:

$$\mathbb{E}_0 T_0 = \infty,$$

although we are sure to (eventually) return with probability 1.  
Recall that the definition of recurrent gives:

$$\mathbb{P}_x(T_x < \infty) = 1 \iff \mathbb{P}_x(T_x \geq n) \downarrow 0 \text{ as } n \uparrow \infty.$$

Also, we should know that positive recurrence implies recurrence, but not necessarily conversely.

**Remark:** Pitman summarizes that on our homework, we can quote the result that:

If we have a stationary measure, then the chain is **positive recurrent**.

## 4 Notion of $x$ -blocks of a Markov chain

Start at  $x$  (for simplicity) or wait until we hit  $x$ . Then look the successive return times  $T_x^{(i)}$  which is the  $i$ th copy of  $T_x$ . Now recall this has the Strong Markov Property, which gives us two things:

- (1) Every  $T_x^{(i)}$  has the same distribution as  $T_x$ .
- (2) Further, they are independent copies. That is,  $T_x^{(1)}, T_x^{(2)}, \dots$  are independent.

Now Pitman mentions a variation on this theme of  $x$ -blocks, which explains many things:

**Example:** Let  $N_{xy}^{(2)} :=$  the # of visits to  $y$  in the  $i$ th block of length  $T_x$ . In our previous in-class example, this gives a sequence:

$$2, 0, 6, 0, 4, 2, \dots$$

Now for some book keeping, consider what happens if we sum over all states  $y$ . Of course, this just gives the length of  $T_x^{(i)}$  by ‘Accounting 101’.

$$\sum_{y \in S} N_{xy}^{(i)} = T_x^{(i)}.$$

Now this implies that there is a formula involving expectations. Take  $\mathbb{E}_x$ , the expectation starting at  $x$ :

$$\sum_{y \in S} \mathbb{E}_x N_{xy}^{(i)} = \mathbb{E}_x T_x^{(i)},$$

where this is really the same equation for all  $i$  by the Strong Markov Property. Fix  $x, y$  and look at  $N_{xy}^{(1)}, N_{xy}^{(2)}, \dots$ , each of which:

- (1)  $N_{xy}^{(i)}$  has the same distribution as  $N_{xy} := N_{xy}^{(1)}$ .
- (2) Further, the  $N_{xy}^{(i)}$  are independent and identically distributed (iid).

Pitman reminds us that as we return to  $x$ , via the Strong Markov Property, nothing of the past changes our expectations or distributions going forward.

Break time.

## 5 Positive Recurrent Chains ( $P$ irreducible)

Notice that if  $\mathbb{E}_x T_x < \infty$ , and we define  $N_{xy}$  as we have earlier, then we can let:

$$\begin{aligned}\mu(x, y) &:= \mathbb{E}_x(N_{xy}) \\ \mu(x) &:= \mathbb{E}_x T_x = \text{mean length of } x\text{-block}\end{aligned}$$

Correspondingly to our Accounting 101 from earlier, we write:

$$\sum_{y \in S} \mu(x, y) = \mu(x) < \infty.$$

Further, we can show (see text for details) that if we sum:

$$\sum_y \mu(x, y) P(y, z) = \mu(x, z),$$

or in other words,  $\mu(x, \cdot)$  is a stationary measure (not a stationary probability, as it is an unnormalized measure). That is,

$$\mu(x, \cdot) P = \mu(x, \cdot).$$

This is important because it gives us a simple explicit construction of a stationary measure  $\mu(x, \cdot)$  for every state  $x$  in state space  $S$  of a positive recurrent (PR) irreducible chain with matrix  $P$ . Notice that this is not just any measure. By convention, we say that the number of times we visit  $x$  in the duration of  $T_x$  is 1 (this is necessary to satisfy our constructions today). That is, we must not count a visit twice, and we must set:

$$\mu(x, x) := 1,$$

in order to get:

$$\sum_{y \in S} \mu(x, y) = \mu(x) < \infty.$$

Now to get a stationary probability measure, we take:

$$\pi(y) = \frac{\mu(x, y)}{\sum_z \mu(x, z)} = \frac{\mu(x, y)}{\mu(x)},$$

and this does NOT depend on  $x$ . We can take any reference state and we get the same thing when we look at these ratios.

### 5.1 Explanation of the Key Formula

We may ask why we have:

$$\sum_y \mu(x, y) P(y, z) = \mu(x, z). \quad (1)$$

Recall that  $\mu(x, y)$  is the expected number of hits on  $y$  before  $T_x$ . That is,

$$\mu(x, y) = \mathbb{E}_x[\# \text{ of hits on } y \text{ before } T_x]$$

Now, every time we hit  $y$ , then  $P(y, z)$  is the probability that the next step is to state  $z$ . Therefore (at least intuitively),  $\mu(x, y)P(y, z)$  has a particular meaning. That is,

$$\mu(x, y)P(y, z) = \mathbb{E}_x ( \# \text{ of transitions } y \rightarrow z \text{ before } (\leq) T_x )$$

The distribution of a single  $x$ -block gives the following formulas for the invariant probability measure  $\pi$ :

$$\pi(x) = \frac{1}{\mathbb{E}_x T_x}, \quad \frac{\pi(y)}{\pi(x)} = \mu(x, y)$$

## 6 Limit Theorems

If we let  $N_n(y) := \sum_{k=1}^n 1(X_k = y) = \#$  of hits on  $y$  in first  $n$  steps, then:

$$\begin{aligned} \mathbb{E}_x \frac{N_n(y)}{n} &= \text{mean } \# \text{ hits on } y \text{ per unit time up to } n \\ &= \frac{1}{n} \sum_{k=1}^n P^k(x, y) \rightarrow \pi(y) \end{aligned}$$

We have this Cesàro mean convergence always for irreducible positive recurrent chains (these themselves do not converge, but their average converges). Now if we additionally impose aperiodicity, we have:

$$P^n(x, y) \rightarrow \pi(y),$$

always for irreducible and positive recurrent and aperiodic.

### 6.1 Review and Audience Questions:

A null recurrent chain has a stationary measure with reference state  $x$  assigned as measure 1. If we do this on a simple random walk, we find that the expected probability of every state is 1, which explains why we expect to spend so much time to return back to  $x$ .

If we have a stationary probability measure:

$$\mathbb{E}_\pi \frac{1}{n} \sum_{k=1}^n P^k(x, y) = \pi(y),$$

we can argue that the stationary measure must be approached in the limit (and hence is unique as a limit must be unique).

Lecture ends here.