# ME41105 - IV Assignment 2 **Bayesian filtering and state estimation**

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# About the assignment

Make the assignments in student pairs, you receive both one grade. Please read through this whole document first such that you have an overview of what you need to do.

This assignment contains *Questions* and *Exercises*:

- You should address all of the questions in a 2 or 3 page report (excluding plots and figures). Please provide separate answers for each Question in your report, using the same Question number as in this document. Your answer should address all the issues raised in the Question, but typically should not be longer than a few lines.
- The Exercises are tasks for you to do, typically implementing a function or performing an experiment. Threrefore, first study the relevant provided code before working on an exercise, as the code may always contains comments refering to each specific exercise. If you do not fully understand the exercise, it may become more clear after reading the relevant code comments! Do not directly address the exercises in your report. Instead, you should submit your solution code together with the report. Experimental results may be requested in accompanying questions.

### You will be graded on:

- 1. Quality of your answers in the report: Did you answer the Questions correctly, and demonstrate understanding of the issue at hand? All Questions are weighted equally.
- 2. Quality of your code: Does the code work as required?
- 3. Quality of presentation: Is your report readable (sentences easy to understand, no grammar mistakes, clear figures)? Is the code you wrote clear and commented?

# **Submitting**

- Before you start, go to the course's **Brightspace** page, and enroll with your partner in a lab group (found under the 'Collaboration' page).
- To submit, upload **two** items on the Brightspace 'Assignments' page:

pdf attachment A pdf with your report.

Do not forget to add your student names and ids on the report.

zip attachment A zip archive with your Matlab code for this assignment

Do NOT add the data files! They are large and we have them already ...

deadline Saturday, December 1 2018, 23:59

- Note that only a single submission is required for your group, and only your last group submission is kept by Brightspace. You are responsible for submitting on time. So, do not wait till the last moment to submit your work, and verify that your files were uploaded correctly. Connection problems and forgotten attachments are not a valid excuses. The due deadline is automatically enforced by Brightspace. If your submission is not on time, you receive an automatic '1'.
- You may only hand in work by you and your lab partner, done this year. You are responsible for not sharing your work neither publicly, nor privately with other students outside your group. Do not put your code or report on public servers. If we believe that you have

used material from other groups, or that you have submitted material that is not yours, it will be reported to the exam committee. This may ultimately result in a failing grade or an expulsion.

• If code is submitted that was written with ill intent, e.g. to manipulate files in the users home directory that were not specified by the task, you will immediately fail the course.

# **Getting assistance**

The primary occasion to obtain help with this assignment is during the lab practicum contact hours of the Intelligent Vehicles course. An instructor and student assistants will be present at the practicum to give you feedback and support. If you find errors, ambiguously phrased exercises, or have another question about this lab assignment, please use the Brightspace lab support forum. This way, all students can benefit from the questions and answers equally. If you cannot discuss your issue on the forum, please contact Julian Kooij (J.F.P.Kooij@tudelft.nl) directly.

Remember that for help on what a specific Matlab command somefunction does or how to use it, use can type from the Matlab command line help somefunction, or doc somefunction.

# Bayesian filtering and state estimation

The previous lab assignment considered object detection from sensor measurements, in particular classifying pedestrians from video frames. This lab assignment we first focus on the next stage in the processing pipeline, namely combining object detections of multiple moving targets into a track representation. The vehicle in this exercise is equipped with a range sensor that obtains noisy distance measurements of the vehicle surroundings from which the true target tracks must be recovered. We shall use probabilistic filtering with Kalman Filters to predict and update estimated target positions, and to assign measurements to tracks. The second part of the lab then considers the related problem of vehicle self-localization, i.e. determine where is the vehicle within the (known) world. This too will be done using a probabilistic filter, but this time we will deal with non-linear vehicle dynamics and measurements using a particle filter.

# Kalman filters and multi-object tracking

Exercises refer to code sections in assignment\_mot.m.

We will start by defining a simple dynamical model for a moving object, such as a person or vehicle. In the context of tracking, such an object of interest is also called a target.

In our setup, a target's coordinates are expressed as 2D (x, y) coordinates. The target's state  $\mathbf{x}_t = [x_t, y_t, \dot{x}_t, \dot{y}_t]^{\mathsf{T}}$  at time t describes its position and velocity  $(\dot{x}_t, \dot{y}_t)$ . We will use a fixed time interval of whole units, i.e.  $t=1,2,3,\cdots$  thus  $\Delta t=1$ . Let the vector  $\mathbf{z}_t=[z_t^x,z_t^y]^{\top}$  represent a positional measurement at that time instance t. Note that we do not directly measure the target's velocity.

We assume that targets generally move at constant velocity, but some deviations on this strict assumption are allowed by including process noise acting on both its position and velocity. This results in an additional random offset  $\epsilon_t = [\epsilon_t^x, \epsilon_t^y, \epsilon_t^x, \epsilon_t^y]$  at each time instance, and there is noise on the measurements as well, given by offset  $\eta_t = [\eta_t^x, \eta_t^y]$ . Written as equations,

$$x_t = x_{t-1} + \dot{x}_{t-1}\Delta t + \epsilon_t^x \qquad y_t = y_{t-1} + \dot{y}_{t-1}\Delta t + \epsilon_t^y \tag{1}$$

$$\dot{x}_t = \dot{x}_{t-1}\Delta t + \epsilon_t^{\dot{x}} \qquad \qquad \dot{y}_t = \dot{y}_{t-1}\Delta t + \epsilon_t^{\dot{y}}$$
(2)

$$z_t^x = x_t + \eta_t^x z_t^y = y_t + \eta_t^y. (3)$$

The noise offsets are assumed to be *independent* and Normally distributed, namely,

$$\epsilon_{t}^{x} \sim \mathcal{N}(0, \sigma_{\text{pos}}^{2}) \qquad \qquad \epsilon_{t}^{y} \sim \mathcal{N}(0, \sigma_{\text{pos}}^{2}) \qquad \qquad (4)$$

$$\epsilon_{t}^{\dot{x}} \sim \mathcal{N}(0, \sigma_{\text{vel}}^{2}) \qquad \qquad \epsilon_{t}^{\dot{y}} \sim \mathcal{N}(0, \sigma_{\text{vel}}^{2}) \qquad \qquad (5)$$

$$\eta_{t}^{x} \sim \mathcal{N}(0, \sigma_{\text{obs}}^{2}) \qquad \qquad \eta_{t}^{y} \sim \mathcal{N}(0, \sigma_{\text{obs}}^{2}) \qquad \qquad (6)$$

$$\epsilon_t^{\dot{x}} \sim \mathcal{N}(0, \sigma_{\text{vel}}^2)$$
  $\epsilon_t^{\dot{y}} \sim \mathcal{N}(0, \sigma_{\text{vel}}^2)$  (5)

$$\eta_t^x \sim \mathcal{N}(0, \sigma_{\text{obs}}^2)$$

$$\eta_t^y \sim \mathcal{N}(0, \sigma_{\text{obs}}^2)$$
(6)

where  $\sigma^2_{\rm pos}$ ,  $\sigma^2_{\rm vel}$  and  $\sigma^2_{\rm obs}$  are the variances of the noise distributions (noise for x and y are equally distributed). To track the target, we first reformulate these assumptions about the dynamics as a single Linear Dynamical System, see Appendix A.

Question 1.1. Formulate the above dynamics using Equations (17) and (18) as a single LDS with 4D state  $x_t$  and 2D measurements  $z_t$ . This entails that you have to determine matrices F, H,  $\Sigma_{\mathbf{x}}$ , and  $\Sigma_{\mathbf{z}}$ . Express  $\Sigma_{\mathbf{x}}$ , and  $\Sigma_{\mathbf{z}}$  in terms of  $\sigma_{\text{pos}}^2$ ,  $\sigma_{\text{vel}}^2$  and  $\sigma_{\text{obs}}^2$ , and note the tip in the Appendix on covariance matrices for independently distributed random variables. What four matrices did you find?

Once we have formulated our problem as a LDS, the (linear) Kalman Filter (KF) can be used to estimate the target's true position, and maintain an uncertainty estimate. We shall implement the Kalman Filter using this LDS.

Exercise 1.1. Add the missing LDS matrices that you found in the previous question to the KF setup in kf\_init.m. Values for the noise parameters  $\sigma_{\rm pos}^2$ ,  $\sigma_{\rm vel}^2$  and  $\sigma_{\rm obs}^2$  are given in that file too. The code block in the assignment script performs some checks on your solution.

Before moving to a situation with measurements, we focus on studying the effect of the Kalman predict step on a state distribution. Let's assume a simple initial state distribution, namely  $P(\mathbf{x}_0) = \mathcal{N}(\mathbf{x}_0 | \boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0)$  with  $\boldsymbol{\mu}_0 = [0, 0, 0, 0]^{\top}$  and  $\boldsymbol{\Sigma}_0 = I$  (where I means the identity matrix).

Exercise 1.2. Implement the predict step in  $kf_predict_step.m$ . The equations of the Kalman Filter predict step can be found in Appendix B. The code in the assignment script should show you that repeated application of the predict step will increase the initial uncertainty in both x and y direction.

The function that visualizes the distribution takes as argument the two dimensions from the state vector for which the distribution is plotted.

- *Exercise 1.3.* Instead of dimensions 1 and 2, now visualize how the predict step affected the uncertainty on state dimensions 1 and 3, being  $x_t$  and  $\dot{x}_t$ . Then, also study the uncertainty for state dimensions 1 and 4, being  $x_t$  and  $\dot{y}_t$ .
- **Question 1.2.** Include the final plot for dimensions 1 & 3, and for 1 & 4 in your report. Explain the direction of the shapes that you see: Are  $x_t$  and  $\dot{x}_t$  (not?) correlated? How would you intuitively explain this? Are  $x_t$  and  $\dot{y}_t$  (not?) correlated? How would you intuitively explain this?

Single target Kalman Filter We now turn towards a simulation where there is a single pedestrian moving in front of the vehicle. In our test setup demonstrated in Figure 1, the vehicle is not moving, and has a  $90^{\circ}$  front-facing distance sensor which measures proximity to nearby objects, such as pedestrians. Our goal is to use the sensor to monitor the presence of any pedestrians, and if detected, follow their path in front of the vehicle, and express their estimated positions as a sequence of groundplane coordinates.

- Exercise 1.4. Run the code block with simulation code, which generates simulated measurements. A figure shows a matrix with all measurement over the T time steps. New random measurements can be made by altering the seed in <code>jk\_srand</code>. You can also change the variable <code>scenario\_version</code> from 1 to 2 to make the simulated pedestrian stop half way.
- Exercise 1.5. For both scenario versions ('crossing' and 'stopping'), also run the next code block which animates the simulated situation. Try to see if you can visualize a single frame from the animation by changing the time steps in the animation for-loop. Also try to speed-up or slow-down the animation with the pause command. Understanding how the script works will help you debug and study the results in the following exercises.

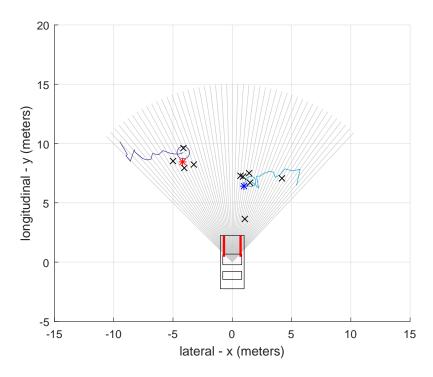


Figure 1: Multi-object tracking in the groundplane, seen in a top-down view. The vehicle's sensor is located at the origin of the coordinate system. The sensor measures distances up to 15 meters along R rays over a  $90^\circ$  angle. In the shown time instance, a pedestrian is crossing in front of the vehicle (true target positions shown as blue star). The  $\times$  marks indicate measured distances along the sensor rays. The solid blue circle shows the filtered 2D mean position of the target, and its variance at 1 standard deviation. The blue line shows the estimated position from past frames. The dotted blue circle shows the 1 standard deviation of where the filter expects the noisy observations to be.

Now let's see what your Kalman predict implementation does if we initialize it on top of the pedestrian's position, and predict its movements.

Exercise 1.6. Run the Kalman predict step on the simulated scenario.

Clearly, the uncertainty distribution only grows as the filter never receives new information on the target. We must utilize the measurements to reduce its uncertainty. The run\_single\_kf implements the minimal Kalman Filter algorithm. For each time step, it (1) converts the measured distances along the rays to 2D positions, (2) executes the KF predict step, and finally (3) executes the KF update step on usable measurements.

Exercise 1.7. Implement the update step in the kf\_update\_step.m function. The equations of the Kalman Filter update step can also be found in Appendix B.

Your implementation should now be able to track the pedestrian reasonably well!

Exercise 1.8. Experiment with increasing and decreasing process noise  $\Sigma_x$  and observation noise  $\Sigma_y$ , for instance by scaling all its elements by 0.01 or 100 (see the commented line in the assignment script just after calling kf\_init()).

**Question 1.3.** When does the track become more smooth, when the process noise and/or observation noise covariance is large or small? What for effect does this have on the pedestrian 'stopping' scenario?

**Single target with outliers** Now we consider that the sensor measurements also include false positives, which are random and therefore do not provide any evidence on the target's position. We refer to such measurements as *outliers*.

*Exercise 1.9.* Run your current KF implementation on the same scenario but with the new measurements that contain outliers.

We can try to remove outliers by pre-processing the range measurements. For instance, we can compare the distance measurement of each sensor ray to that of the adjacent rays, and remove large differences. Hence, this is a form of *spatial filtering* (not to be confused with the *temporal filtering* of the KF). For this example, we shall use a median filter for spatial filtering. For instance, a 3-rd order median filter compares each distance measurement to its two surrounding measurements, and takes the median value of all three.

Exercise 1.10. A spatial median filter implementation is provided in the preprocess\_measurements function. Try out different filter orders in the function, namely 1-st order, 3-rd order, 7-th order, and see what happens to the measurements in the measurement matrix visualization.

**Question 1.4.** Which order works best for the current setup? Add a matrix visualization with all the pre-processed sensor measurement in your report.

A different approach is not to alter the measurements themselves, but to try to determine which measurements are from the tracked target, and which should be discarded. This is done through comparing the measurements to the filtered position of the KF. If the distance between the two positions is below a certain "gating" threshold, the measurement can be used in the update step.

Exercise 1.11. Implement this gating strategy in test\_gating\_score: compute the Euclidean distance between the measurement, and the expected position given by the Kalman Filter. Accept the measurement if the distance is below some threshold (in meters). Try to get a feeling of what threshold works good by looking at the tracking results.

**Question 1.5.** What happens if the gating threshold is too large (e.g. 10 meters)? What happens if the gating threshold is too small (e.g. 1 meters)? Which gating threshold did you select?

**Two known targets** Now we move to the next scenario where there are two moving targets, for which there are again two scenario version: 'crossing' and 'stopping'.

*Exercise 1.12.* Generate measurements for the two target scenario, and plot the measurements over all time steps in a single figure. Do this for both scenario versions.

We assume that we know where the targets start, and initialize a KF for each. The function run\_multiple\_kfs executes the predict and update steps for both KFs, using *data association* at each time step to decide which measurements belong to which KF.

Exercise 1.13. In associate\_measurements\_to\_tracks, finish the code to assign measurements to tracks. To do this, we compare a measurement's gating scores for all Kalman Filters, and assign it to the best scoring KF (if it exceeds the gating threshold).

**Multi-object tracking** In general, we do not know how many targets there are (or if there are any targets at all). Our tracker therefore needs to initialize and remove KFs when it is appropriate. The provided function run\_multi\_object\_tracker extends run\_multiple\_kfs with track management code. It starts a new KF for unused observations, and terminates it when it has had no updates for too many iterations.

Exercise 1.14. Run your multi-object tracker on the scenario. Try to create an optimal multi-object tracker by pre-processing and/or adjusting your gating threshold, and/or the track termination threshold. You probably won't get it to work perfectly, this is inherently a hard problem!

**Question 1.6.** What settings did you find to work best? How many tracks were created?

**Question 1.7.** To answer the previous question, did you try out several variations of the simulation by altering the seed of the random number generator <code>jk\_srand</code>? Motivate your choice: why is it important to try/not try variations of the scenario when optimizing an intelligent vehicle's tracker?

### 2 Vehicle self-localization

Exercises refer to code sections in assignment\_self\_localization.m.

After tracking other objects in the first part of the assignment, we now turn to localizing the vehicle itself. Consider the situation where we have a description of the static environment. One such description is an *occupancy map*, also known as *occupancy grid*. In our example, we use a simple binary occupancy grid which divides the 2D groundplane into square cells. Each cell can either be occupied (i.e. contain an obstacle), or not occupied (i.e. free space). The grid can be used in the sensor model to determine what distance we would measure at a particular location in a particular direction.

The vehicle is equipped with the same range sensor as in the previous exercise, but now has R=32 rays and full  $360^\circ$  coverage. Thus, the sensor can reliably measure the distance to the closest obstacle in a large number of directions around the vehicle, see Figure 2. Apart from the measurements, we also know the vehicle's *control input* at each time step, which is given as a 2D vector  $\mathbf{u}_t = [v_t, \omega_t]^\top$ . Here  $v_t$  is the vehicle's velocity (in meters per second), and  $\omega_t$  its turning rate (in radians per second).

Exercise 2.1. Run the self-localization scenario setup and visualization code. You will see the vehicle drive around some street blocks, while measuring its environment.

In the following exercise, you will implement a Particle Filter to track the vehicle itself through its environment using its measurements only.

The particle filter allows you to deal with the non-linear dynamics of the vehicle and sensor model. Additionally, it can represent more complex state distributions than a the single Gaussian of the Kalman Filter. It does this by keeping a set of N particles  $S_t = \{\mathbf{x}_t^{(1)}, \cdots \mathbf{x}_t^{(N)}\}$  which

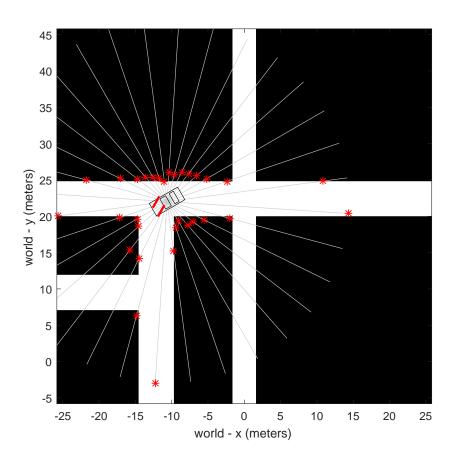


Figure 2: Vehicle self-localization problem. The vehicle moves through a binary grid world, represented by an occupancy map (here black regions define occupied space, white regions show free space). The vehicle has a range sensor with R=32 rays and  $360^\circ$  coverage. The red stars show the R measured distances at the current time step. Note the red lines showing the vehicle steering angle.

represent possible vehicle states at time t. Note that in this notation  $\mathbf{x}^{(i)}$  represents the i-th particle. You can think of the particles as hypotheses on the true vehicle state. In our case, each particle  $\mathbf{x} = [x, y, \theta]^{\top}$  will consist of a 2D position (x, y) plus orientation angle  $\theta$  to describe the forward direction of the vehicle. Please read Appendix C.

**PF predict step** We will use simple non-linear dynamics f to propagate the particles,

$$f(\mathbf{x}_t^{(i)}, \mathbf{u}_t, \Delta t) = \mathbf{x}_t^{(i)} + \begin{pmatrix} v_t \, \Delta t \, \sin(\theta_t^{(i)}) \\ v_t \, \Delta t \, \cos(\theta_t^{(i)}) \\ \omega_t \Delta_t \end{pmatrix}. \tag{7}$$

These dynamics define the how the particle would move in an 'ideal' deterministic world, i.e.  $f(\mathbf{x}_t^{(i)}, \mathbf{u}_t, \Delta t)$  is the *expected* particle state at  $\mathbf{x}_{t+1}^{(i)}$ . Then, Gaussian noise is added to this predicted state of each particle to account for the uncertainty in the motion, see Equation (32).

**Question 2.1.** Using Equation (7), write out the expected state  $\mathbf{x}_2 = [x_2, y_2, \theta_2]^{\top}$  as a function of the initial state  $\mathbf{x}_0 = [x_0, y_0, \theta_0]^{\top}$ , the control inputs  $\mathbf{u}_0$  and  $\mathbf{u}_1$  and  $\Delta t$  (assumed to be constant over all frames). Note that  $\mathbf{x}_2 = f(f(\mathbf{x}_0, \mathbf{u}_0, \Delta t), \mathbf{u}_1, \Delta t)$ , and that we do not consider any Gaussian noise in this question. Give your answer as the three equations, namely for  $x_2, y_2$  and  $\theta_2$ . Which of these equations is/are linear, and which non-linear?

Exercise 2.2. Complete the PF predict step in pf\_predict\_step by propagating the particles with the dynamics in Equation (7), and then adding the Gaussian noise sampled from  $\mathcal{N}(0,\Sigma_x)$ . The covariance  $\Sigma_x$  is given in the pf\_predict\_step function.

The assignment code block initializes N=1000 samples at the initial vehicle state (position+orientation), and applies your prediction step several times. It visualizes the position (not orientation) of the particles with a scatter plot. You should see that the particles create a nonlinear curved distribution. You can try changing the velocity and steering angle parameters used in the prediction. Notice that the used particle dynamics do *not* take the obstacles from the map into account.

**PF update step** Next, we will use the measurement to evaluate how likely each particle is. Before implementing code, we empirically study what the measurements would be at a few interesting predefined particle states. In the assignment code, you can select one of several interesting particles  ${\bf x}$ . The code block also shows how to create a virtual sensor located at the particle's position and orientation, and use it to create an ideal 'expected' distance measurement vector  $\hat{{\bf z}}=h({\bf x})$  of its environment.

Exercise 2.3. Run the code block for the 7 different particle positions, and visually compare the expected distance measurements of the particle to the actual measurements obtained at the true vehicle position. Notice that several particles result in very similar measurements.

**Question 2.2.** Which of the 7 provided particle locations results in a very similar measurement pattern to the actual measurement pattern? Why do the measurement patterns of those particles look like the actual measurements, and why are those of the other particles so different (look at the visualized vehicle state, its surroundings, and sensor measurements)?

The measurement model computes the probability of observing the given measurements at a particular particle state. But instead of normal probabilities, we will use log-probabilities, read Appendix D. Adapting Equation 33, the log measurement likelihood becomes

$$\log P(\mathbf{z}_t|\mathbf{x}_t^{(i)}) = \alpha + \sum_{r=1}^{R} -(z_r - \hat{z}_r^{(i)})^2 / 2\sigma^2$$
(8)

where  $\alpha$  is a certain constant.

**Question 2.3.** Prove that this equation follows from Equation (33) (you should use the definition of the Gaussian distribution p.d.f.). How does this show that  $\alpha$  is constant for all particles  $\mathbf{x}_{t}^{(i)}$ ?

The likelihood defines the *particle's weight*, but since the weights are eventually renormalized to probabilities that sum to 1, we only need a weight up to a certain constant factor. In the log space, this means that we can drop the  $\alpha$  term, i.e.

$$\log w^{(i)} = \sum_{r=1}^{R} -(z_r - \hat{z}_r^{(i)})^2 / 2\sigma^2.$$
(9)

Exercise 2.4. Complete the map\_measurement\_loglik function using Equation (9). Note that the particles in Exercise 2.3 that you found to have similar measurements to the true position should have high probability.

Exercise 2.5. Now implement the PF update step in pf\_update\_step. The update step should evaluate the measurement log-likelihood for each particle, normalize these a probability distribution, and use these probabilities to resample the particles (with replacement).

**Complete Particle Filter** Next we will combine the predict and update step to run the full particle filtering algorithm on the test scenario. We will start with the situation where the vehicle's initial position and orientation are known.

*Exercise 2.6.* Run your particle filter on the full scenario. Try your filter with at least N=5, N=10, and N=100 particles, but feel free to try other values. Repeat each experiment several times to reduce the effect of chance in your evaluation. Note that the magenta dots represent the predicted particles, and the greed dots the resampled ones. As a summary of all particles, the mean and covariance are computed, and shown as a blue 2D Gaussian distribution. See Figure 3.

**Question 2.4.** What benefit(s) do you notice of having few particles? What benefit(s) do you notice of having many particles? How many particles do you (approximately) need to reliably keep track of the vehicle?

Now let us consider that we do not know where the vehicle is (or maybe we lost track and need to reinitialize).

Exercise 2.7. This time, initialize all particles randomly at the unoccupied cells. You can do this by setting INITIAL\_POSITION\_KNOWN = false at the start of the code block for Exercise 2.6. Run the Particle Filter on the whole sequence multiple times as before, and try out various amount of N. You may want to try out values larger then N=100.

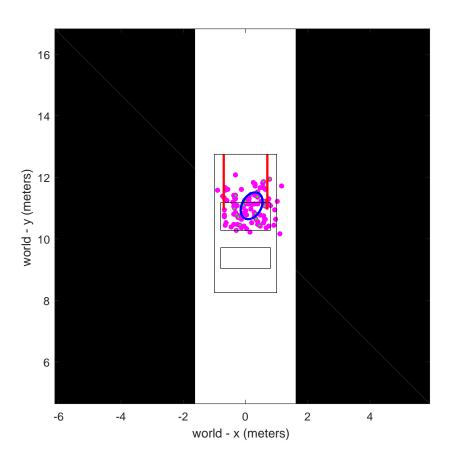


Figure 3: Particle filter with accurate estimate of the vehicle position (closeup on vehicle). The magenta dots show the location of the prediction particles, the green dots which particle state have been resampled (maybe multiple times). The blue 2D gauss shows the overall mean and covariance among the resampled particles.

**Question 2.5.** How does the number of particles correlate with the risk that the filter does not recover the vehicle position by the end of the sequence? How many particles are approximately needed to reliably track the vehicle up to the end of the sequence? Describe your experiments on which you base your conclusions (e.g. which values of N you tried, how many times you repeated each run, etc.).

You may have noticed that if all the particles in the filter have lost track of the actual vehicle state, it will not recover anymore. One strategy to tackle this situation is to randomly reinitialize part of the particles.

Exercise 2.8. Adjust the PF code such that before each update step a percentage of the N particles are replaced by randomly reinitialized particles, see <code>frac\_reinit</code> in the assignment code block. You can use <code>pf\_init\_freespace</code> to create new initial samples. Try out reinitialization percentages of 5%, 25%, 50%, and 75%, on the situation where the initial position is not known. Do this again with N=20 and N=100 particles multiple times.

**Question 2.6.** What happens if you reinitialize a very low fraction of the particles? What happens if you reinitialize a very large fraction of the particles? For the same reinitialization factor, is there a benefit or downside for N=20 versus N=100? Motivate your answer.

# 3 Mathematical proofs with Linear Algebra

One may not find the Kalman Update equations (Equations (27) to (31) in Appendix B) particularly intuitive, or trivial at fist glance. Would it not have been possible to come up with some other set of equations? What are the assumptions underlying these equations? It turns out that these equations follow from the using linear models, Gaussian distributions for the prior, process and observation noise, and consistent application of Bayes' rule of probability. In this assignment, you will not do any coding, but use your knowledge on linear algebra to work out mathematical proofs which you will include in your report.

In Appendix E you can read about computing the posterior of a Multivariate normal distribution in closed form. We can proof that the Kalman filter update equations basically compute this same posterior, though this proof requires the use of some known matrix identities. Here are two matrix identities that you can use in your proofs, namely the so-called Woodbury matrix identity,

$$(A + XBX^{\top})^{-1} = A^{-1} - A^{-1}X(B^{-1} + X^{\top}A^{-1}X)^{-1}X^{\top}A^{-1}$$
(10)

and also

$$BX^{\top}(A + XBX^{\top})^{-1} = (B^{-1} + X^{\top}A^{-1}X)^{-1}X^{\top}A^{-1}.$$
 (11)

### 3.1 Covariance update

Each Appendix gives a different definition on how to compute the covariance update:

$$\Sigma_t = (I - KH) \widehat{\Sigma}_t$$
 (KF covariance update, (31) in Appendix B)  
 $\Sigma_t = (\widehat{\Sigma}_t^{-1} + H^{\top} \Sigma_{\mathbf{z}}^{-1} H)^{-1}$  (MVN posterior covariance, (38) in Appendix E)

Question 3.1. Give a derivation to prove that these two definitions are the same, i.e. that

$$(I - KH)\widehat{\boldsymbol{\Sigma}}_t = (\widehat{\boldsymbol{\Sigma}}_t^{-1} + H^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} H)^{-1}.$$
 (12)

### 3.2 Mean update

Likewise, each Appendix gives a different definition on how to compute the updated mean:

$$\mu_t = \widehat{\mu}_t + Ke$$
 (KF mean update, (30) in Appendix B) 
$$\mu_t = (\widehat{\boldsymbol{\Sigma}}_t^{-1} + \boldsymbol{H}^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \boldsymbol{H})^{-1} (\widehat{\boldsymbol{\Sigma}}_t^{-1} \widehat{\boldsymbol{\mu}}_t + \boldsymbol{H}^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \mathbf{z})$$
 (MVN posterior mean, (39) in Appendix E)

Proving that these mean update rules are the same is a bit more complicated. Here we will make a start, which still requires you to prove two parts:

$$\mu_{t} = \widehat{\mu}_{t} + Ke$$

$$= \underbrace{(I - KH)\widehat{\mu}_{t}}_{\text{Part A}} + \underbrace{K\mathbf{z}}_{\text{Part B}}$$

$$= \underbrace{(\widehat{\boldsymbol{\Sigma}}_{t}^{-1} + H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}H)^{-1}\widehat{\boldsymbol{\Sigma}}_{t}^{-1}\widehat{\mu}_{t}}_{\text{Part A}} + \underbrace{(\widehat{\boldsymbol{\Sigma}}_{t}^{-1} + H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}H)^{-1}H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}\mathbf{z}}_{\text{Part B}}$$

$$= (\widehat{\boldsymbol{\Sigma}}_{t}^{-1} + H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}H)^{-1}(\widehat{\boldsymbol{\Sigma}}_{t}^{-1}\widehat{\mu}_{t} + H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}\mathbf{z})$$

$$= (\widehat{\boldsymbol{\Sigma}}_{t}^{-1} + H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}H)^{-1}(\widehat{\boldsymbol{\Sigma}}_{t}^{-1}\widehat{\mu}_{t} + H^{\top}\boldsymbol{\Sigma}_{\mathbf{z}}^{-1}\mathbf{z})$$
(MVN posterior mean)

Question 3.2. Prove Part A,

$$(I - KH)\widehat{\boldsymbol{\mu}}_t = (\widehat{\boldsymbol{\Sigma}}_t^{-1} + H^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} H)^{-1} \widehat{\boldsymbol{\Sigma}}_t^{-1} \widehat{\boldsymbol{\mu}}_t$$
 (15)

**Question 3.3.** Finally, prove Part B:

$$K\mathbf{z} = (\widehat{\boldsymbol{\Sigma}}_{t}^{-1} + H^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} H)^{-1} H^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \mathbf{z}$$
(16)

So, what did we learn from proving these equalities? Well, we already could easily see that the Kalman Filter *prediction* equations consisted of applying the linear state transition on the mean Equation 23, and by *increasing the variance* of our estimate by summing the state and process covariance in Equation 24.

With our proofs, and the intuition gained at the end of Appendix E, we can now also intuitively understand that the Kalman Filter *update* equations 27-31 do the 'inverse', namely they just sum the state and observation *precision* matrices, and take a weighted average of the predicted mean and observation.

# **Appendix**

# Linear Dynamical System

A Linear Dynamical System (LDS) describes a sequence of states  $x_1, x_2, \cdots$ , where  $x_t$  is the state at time t, and each state is an N-dimensional vector. The transition from one state to the next can be written as a linear transformation of the previous state  $x_{t-1}$ , and some additional Normally distributed noise  $\epsilon_t$ . The state is called *latent* in the sense that it cannot be observed directly. In an LDS, a measurement is an M-dimensional vector  $\mathbf{z}_t$ . A measurement is related to the true latent state  $x_t$  trough a linear transformation, but is again some Normally distributed measurement 'noise'.

The full system can therefore be written as

$$\mathbf{x}_t = F\mathbf{x}_{t-1} + \boldsymbol{\epsilon}_t$$
 where  $\boldsymbol{\epsilon}_t \sim \mathcal{N}(0, \Sigma_{\mathbf{x}})$  (17)  
 $\mathbf{z}_t = H\mathbf{x}_t + \boldsymbol{\eta}_t$  where  $\boldsymbol{\eta}_t \sim \mathcal{N}(0, \Sigma_{\mathbf{z}})$  (18)

$$\mathbf{z}_t = H\mathbf{x}_t + \boldsymbol{\eta}_t$$
 where  $\boldsymbol{\eta}_t \sim \mathcal{N}(0, \Sigma_{\mathbf{z}})$  (18)

where F is the  $N \times N$  transition matrix, and, H is the  $M \times N$  observation matrix. Also,  $\Sigma_{\mathbf{x}}$  is the  $N \times N$  covariance matrix of the process noise, and  $\Sigma_{\mathbf{z}}$  the  $M \times M$  covariance matrix of the observation noise.  $\mathcal{N}(\mu, \Sigma)$  indicates a Multivariate Normal (i.e. Gaussian) distribution with mean  $\mu$  and covariance  $\Sigma$ .

Note that if all N elements  $\epsilon = [\epsilon_1, \dots, \epsilon_N]$  are *independently* distributed, and element  $\epsilon_i$  have variance  $\sigma_i^2$ , i.e. each  $\epsilon_i \sim \mathcal{N}(0, \sigma_i^2)$ , then  $\Sigma_{\mathbf{x}}$  is a diagonal,

$$\Sigma_{\mathbf{x}} = \begin{bmatrix} \sigma_1^2, & 0, & \cdots, & 0 \\ 0, & \sigma_2^2, & \cdots, & 0 \\ 0, & 0, & \ddots, & 0 \\ 0, & 0, & \cdots, & \sigma_N^2 \end{bmatrix}.$$
 (19)

### Kalman Filter

The Kalman Filter assumes that a process can be described as a Linear Dynamical System (see Appendix A). It maintains a posterior distribution over the process state, given all observations until the current time step. This state distribution is always represented by a Multivariate Normal distribution. Just as other Bayesian filters, it consists of the following steps:

**Initialize** To initiate the Kalman Filter at t = 0, we need to define the initial state distribution:

$$P(\mathbf{x}_0) = \mathcal{N}(\mathbf{x}_0 \mid \boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0)$$
 initial distribution (20)

This step can be followed with an update step if there are any initial observations. For subsequent time instances (t > 0), the initialization is replaced by a *predict* step (followed by an update).

**Predict step** The predict step uses the filtered state distribution  $P(\mathbf{x}_{t-1}|\mathbf{z}_{0:t-1})$  of the previous time instance to create a predicted state distribution  $P(\mathbf{x}_t|\mathbf{z}_{0:t-1})$  for the new time instance.

These Multivariate Normal distributions are parameterized as

$$P(\mathbf{x}_{t-1}|\mathbf{z}_{0:t-1}) = \mathcal{N}(\mathbf{x}_{t-1} \mid \boldsymbol{\mu}_{t-1}, \boldsymbol{\Sigma}_{t-1})$$
 posterior from previous time step (21)

$$P(\mathbf{x}_t|\mathbf{z}_{0:t-1}) = \mathcal{N}(\mathbf{x}_t \mid \widehat{\boldsymbol{\mu}}_t, \widehat{\boldsymbol{\Sigma}}_t)$$
 prediction for new time step (22)

The predict step computes the new parameters as

$$\hat{\mu}_t = F \mu_{t-1}$$
 predicted mean (23)

$$\widehat{\mathbf{\Sigma}}_t = F \mathbf{\Sigma}_{t-1} F^{\mathsf{T}} + \Sigma_{\mathbf{x}} \qquad \text{predicted covariance}$$
 (24)

**Update step** The update step incorporates evidence from measurements into the state distribution. On each time step, it can be called *zero or more times*, namely once for each available independent measurement  $\mathbf{z}_t$ .

For a single update with  $z_t$ , we consider the distribution before the update the Bayesian *prior*, and the resulting distribution the *posterior*, which are parameterized as

$$P(\mathbf{x}_t|\mathbf{z}_{0:t-1}) = \mathcal{N}(\mathbf{x}_t \mid \widehat{\boldsymbol{\mu}}_t, \widehat{\boldsymbol{\Sigma}}_t) \qquad \text{prior of } \mathbf{z}_t \text{ (from prediction or earlier updates)}$$
 (25)

$$P(\mathbf{x}_t|\mathbf{z}_{0:t}) = \mathcal{N}(\mathbf{x}_t \mid \boldsymbol{\mu}_t, \boldsymbol{\Sigma}_t)$$
 updated (posterior of  $\mathbf{z}_t$ ) (26)

The parameters of the updated distribution are computed as

$$e_t = \mathbf{z}_t - H\hat{\boldsymbol{\mu}}_t$$
 innovation (residual error) (27)

$$\mathbf{S}_t = H \widehat{\mathbf{\Sigma}}_t H^{\top} + \Sigma_{\mathbf{z}}$$
 innovation covariance (28)

$$K_t = \widehat{\mathbf{\Sigma}}_t H^{\mathsf{T}} \mathbf{S}_t^{-1}$$
 Kalman gain (29)

$$\mu_t = \widehat{\mu}_t + K_t e_t$$
 updated mean (30)

$$\Sigma_t = (I - K_t H) \hat{\Sigma}_t$$
 updated covariance (31)

The updated state distribution after all measurements can then serve as the input to the predict step at the next time instance t+1.

# **C** Particle Filtering

While Kalman Filtering represents the state distribution using multivariate Gaussian distributions, a Particle Filter does not assume a specific parametric form of the distribution. Instead, the distribution is represented by a set of *samples*. The prediction and update step manipulate the set of samples, such that it remains representative for the state distribution as time progresses and more measurements are obtained. Benefits include the unconstrained form of the distribution, support of non-linear motion and observation models, and the simple algorithmic implementation. A downside is that it has inherent randomness, and the number of particles may be too limited to represent the shape of the true distribution.

Let  $S_t = \{\mathbf{x}_1^1, \cdots, \mathbf{x}_t^N\}$  be the set of N particles at time t. The particle filter consists of the following steps:

**Initialize** Create the initial set  $S_0 = \{\mathbf{x}_0^1, \cdots, \mathbf{x}_0^N\}$  with N particles by sampling N times from the initial distribution  $P(\mathbf{x}_0)$ . Depending on the setup, the initial distribution may be very specific (i.e. we know where the object is), or evenly spread out over the space (we have to recover where it is).

**Predict** The predict step creates a set  $\widehat{S}_{t+1}$  from the current particle set  $S_t$ . It does this by applying a possibly non-linear dynamical model to move each particles forward, given the control input, and time difference  $\Delta t$  between current time step t and t+1. For the 2D space + orientation representation, a simple non-linear dynamical model  $f(\mathbf{x}_t^{(i)}, \mathbf{u}_t, \Delta t)$  is

$$\mathbf{x}_{t+1}^{(i)} = f(\mathbf{x}_t^{(i)}, \mathbf{u}_t, \Delta t) + \epsilon_t^{(i)} \quad \text{where } \epsilon_t^{(i)} \sim \mathcal{N}(0, \Sigma_x)$$
 (32)

which is an alternative way to formulate  $P(\mathbf{x}_{t+1}|\mathbf{x}_t,\mathbf{u}_t) = \mathcal{N}(\mathbf{x}_{t+1}|f(\mathbf{x}_t,\mathbf{u}_t,\Delta t),\Sigma_x)$ . Similar to the Kalman Filter,  $\epsilon_t^{(i)}$  is additional zero-mean Gaussian noise on the state. The predict step can therefore be split into two steps: first, apply the deterministic non-linear dynamics to move the particle forward using control input  $\mathbf{u}_t$ . Second, add some randomly sampled Gaussian noise to each particle.

**Update** In the update step, a new set  $S_{t+1} = \{\mathbf{x}_{t+1}^{(1)}, \cdots, \mathbf{x}_{t+1}^{(N)}\}$  of N particles is constructed by sampling N times (with replacement) particles from the predicted set  $\widehat{S}_{t+1}$ . The probability  $p^{(i)}$  that a particle i in  $\widehat{S}_{t+1}$  is sampled is proportional to the particle weight  $w^{(i)}$ , i.e.  $p^{(i)} = w^{(i)} / \sum_i w^{(i)}$ . To determine this weight we compare the observed R-dimensional measurement vector to the expected R-dimensional measurement of each particle, i.e. the observation likelihood  $P(\mathbf{z}|\mathbf{x}^{(i)})$  is used as the particle weight.

Let  $h(\mathbf{x}_t^{(i)}) = \hat{\mathbf{z}}^{(i)} = [\hat{z}_1^{(i)}, \cdots, \hat{z}_R^{(i)}]^{\top}$  be the expected measurement at the i-th particle  $\mathbf{x}_t^{(i)}$  in  $\widehat{S}_{t+1}$ . Here h() is the non-linear observation model that maps the particle's state to the R-dimensional measurement vector (e.g. the expected distances along R sensor rays using an occupancy map). We place a zero-mean Gaussians with variance  $\sigma_x^2$  on the R errors between the expected and observed values. In other words, the measurement likelihood is large when the error over all dimensions is small. The complete likelihood can thus be written as,

$$P(\mathbf{z}|\mathbf{x}^{(i)}) = P(\mathbf{z}|h(\mathbf{x}^{(i)})) = \prod_{r=1}^{R} P(z_r|\hat{z}_r^{(i)}) = \prod_{r=1}^{R} \mathcal{N}(z_r - \hat{z}_r^{(i)}|0, \sigma^2).$$
(33)

We use these likelihood as the non-zero particle weights. In other words, particles whose expected measurement is very similar to the actual measurement will have a high weight, which in return increases their chance of being present in the new resampled set  $S_{t+1}$ .

# D Using log-probabilities in computations

Often in numerical computations concerning many small probabilities it is advantageous to use log-probabilities instead.

First, log probabilities are much less prone to suffer from the rounding errors that the computer makes to internally represent the numbers. In some cases, normal probabilities can get so small that they become indistinguishable after the rounding errors.

Another benefit is that we can avoid many calls to  $\exp()$  and  $\log()$  that occur in the pdfs of many distributions, e.g. the Gaussian, which are relatively slow to compute. Note that multiplying terms of normal probabilities results in adding terms of log-probabilities.

Special care has to be taken however when eventually converting the log-probabilities back to normal probabilities, e.g. when normalizing.

# E Bayes' rule and the Multivariate Normal distribution

Bayes' rule describes the relation between the prior  $P(\mathbf{x})$ , likelihood term  $P(\mathbf{z}|\mathbf{x})$ , and posterior  $P(\mathbf{x}|\mathbf{z})$ . If the state  $\mathbf{x}$  is a continuous variable, it can be written as

$$P(\mathbf{x}|\mathbf{z}) = \frac{P(\mathbf{z}|\mathbf{x})P(\mathbf{x})}{\int P(\mathbf{z}|\mathbf{x})P(\mathbf{x}) d\mathbf{x}}.$$
 (Bayes' rule)

This rule is the foundation of many probabilistic filters, and can also be used to analytically obtain the posterior, given only the parametric forms of the prior and likelihood terms.

For instance, consider the case where the prior and likelihood terms are multivariate normal distributions, which are found in Linear Dynamical Systems. A D-dimensional Multivariate Normal distribution is a distribution over D-dimensional vectors in  $\mathbb{R}^D$ . At has a D-dimensional mean vector  $\mu$ , and  $D \times D$  positive-definite covariance matrix  $\Sigma$ . Its probability density function (p.d.f.) is given by the formula

$$\mathcal{N}(\mathbf{x}|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} |\Sigma|^{1/2}} \exp\left[\frac{1}{2} (\mathbf{x} - \mu)^{\top} \Sigma^{-1} (\mathbf{x} - \mu)\right]. \tag{35}$$

If both the prior and likelihood term are multivariate normal distributions, then the posterior turns out to also be a multivariate normal distribution. The parameters of this posterior can be computed in closed form given the parameters of the likelihood and prior distributions. Let the prior on  $\mathbf{x}$  be parameterized by mean  $\widehat{\mu}_t$  and covariance  $\widehat{\Sigma}_t$ , and the observation likelihood term for observation  $\mathbf{z}$  express a linear transformation H of  $\mathbf{x}$  with additional white noise with variance  $\Sigma_{\mathbf{z}}$ , i.e.

$$P(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\widehat{\boldsymbol{\mu}}_t, \widehat{\boldsymbol{\Sigma}}_t)$$
 (36)

$$P(\mathbf{z}|\mathbf{x}) = \mathcal{N}(\mathbf{z}|H\mathbf{x}, \Sigma_{\mathbf{z}}). \tag{37}$$

Using Bayes' rule, and a bit of calculus, it can be found that the posterior is also a multivariate Normal distribution with parameters  $\mu_t$  and  $\Sigma_t$ ,

$$P(\mathbf{x}|\mathbf{z}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_t, \boldsymbol{\Sigma}_t)$$
$$\boldsymbol{\Sigma}_t = (\hat{\boldsymbol{\Sigma}}_t^{-1} + \boldsymbol{H}^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \boldsymbol{H})^{-1}$$
(38)

$$\boldsymbol{\mu}_t = (\widehat{\boldsymbol{\Sigma}}_t^{-1} + \boldsymbol{H}^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \boldsymbol{H})^{-1} (\widehat{\boldsymbol{\Sigma}}_t^{-1} \widehat{\boldsymbol{\mu}}_t + \boldsymbol{H}^{\top} \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \mathbf{z}).$$
(39)

These equations may look confusing at first, by can be quite easily understood if we talk not about *covariance* matrices  $\Sigma$ , but instead in terms of *precision* matrices  $\Lambda = \Sigma^{-1}$ , which are the inverse of the covariance. In other words, when the variance is low the precision is high, and vice versa. Also, we can note that H and  $H^{\top}$  just project back and forth between the space of  $\mathbf{x}$  and  $\mathbf{z}$ .

Let's just focus on the case where  ${\bf x}$  and  ${\bf z}$  are in the same space such that H=I. Equations 38 and 39 can now be expressed instead as,

$$\mathbf{\Lambda}_t = \widehat{\mathbf{\Lambda}}_t + \Sigma_{\mathbf{z}}^{-1} \tag{40}$$

$$\boldsymbol{\mu}_t = (\widehat{\boldsymbol{\Lambda}}_t + \boldsymbol{\Sigma}_{\mathbf{z}}^{-1})^{-1} (\widehat{\boldsymbol{\Lambda}}_t \, \widehat{\boldsymbol{\mu}}_t + \boldsymbol{\Sigma}_{\mathbf{z}}^{-1} \mathbf{z}). \tag{41}$$

We can now have an intuitive understanding of what the posterior update does: it is simply summing the precisions (i.e. inverse covariances) of the prior and likelihood term. Also, the posterior mean is simply a weighted average of prior mean and observed  $\mathbf{z}$ , where each is weighted by their precision.