

# Notes on causal inference

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## 1 Why linear Gaussian is non-identifiable?

Theorem quoted in Theorem 4.2 [PJS17]

**Theorem 1.1** (Linear Gaussian). *Given that  $Y = \alpha X + N_Y$ , where  $X \perp\!\!\!\perp N_Y$ . Then there exists  $N_X$  and  $\beta$  such that  $X = \beta Y + N_X$  such that  $Y \perp\!\!\!\perp N_X$  if and only if both  $X$  and  $N_Y$  are normally distributed.*

*Proof.* We need to prove both directions.

**Step 1.** The  $\Rightarrow$  direction: assume  $X, Y_N \sim \mathcal{N}$ .

So we need to find some  $\beta$  and  $N_X$  such that they makes  $Y$  and  $N_X$  to be independent.

Since  $Y \sim \mathcal{N}$ , it would be easier if  $N_X$  is also normal. In that case, we only need to design  $N_X$  so that

$$\text{cov}(Y, N_X) = 0.$$

So the goal is to find  $N_X$  that satisfy:

- $N_X$  is normal.
- $\text{cov}(Y, N_X) = 0$ .

Let start from

$$N_X = X - \beta Y = X - \beta(\alpha X + N_Y) = (1 - \alpha\beta)X - \beta N_Y$$

haha, this is already normal little enforcement to ensure the whole thing is not 0. So the other thing should be

$$\begin{aligned} & \text{cov}(N_X, Y) = 0 \\ \Leftrightarrow & \text{cov}((1 - \alpha\beta)X - \beta N_Y, \alpha X + N_Y) = 0 \\ \Leftrightarrow & (1 - \alpha\beta)\alpha \text{var}(X) + (1 - \alpha\beta)\text{cov}(X, N_Y) - \alpha\beta \text{cov}(N_Y, X) - \beta \text{var}(N_Y) = 0 \\ \Leftrightarrow & (1 - \alpha\beta)\alpha \text{var}(X) - \beta \text{var}(N_Y) = 0 \\ \Leftrightarrow & \alpha \text{var}(X) = \alpha^2 \beta \text{var}(X) + \beta \text{var}(N_Y) \\ \Leftrightarrow & \beta = \frac{\alpha \text{var}(X)}{\alpha^2 \text{var}(X) + \text{var}(N_Y)} \end{aligned}$$

**Step 2.** The  $\Leftarrow$  direction: assume the existence of  $N_X, \beta$  and ... It turns out the reverse direction is harder to prove. We need to invoke Darmois-Skitovich theorem.

**Theorem 1.2** (Darmois-Skitovich). *Let  $X_1, \dots, X_N$  be mutually independent. If there are exists nonzero coefficients  $\alpha_i, \beta_i$  such that*

$$Y_1 = \sum_{i=1}^N \alpha_i X_i, \quad Y_2 = \sum_{i=1}^N \beta_i X_i$$

*where  $Y_1 \perp\!\!\!\perp Y_2$ , then all  $X_i$ 's are normally distributed.*

Now we have two RV  $X \perp\!\!\!\perp N_Y$ , and there are some nonzero coefficients that make

$$\begin{cases} Y = \alpha X + N_Y \\ N_X = (1 - \alpha\beta)X - \beta N_Y \end{cases}$$

$Y \perp\!\!\!\perp N_X$ . Hence by Darmois-Skitovich theorem,  $X, N_Y$  are normally distributed.

Of course, we need to handle some corner cases where some of the coefficients are 0. But since I'm lazy ...  $\square$

Whenever two RVs are not independent, either one RV causes the other or either there is a common cause to both of them. In the former case, we want to determine the direction of the causality given joint pdf of 2 RVs. For example, if  $X$  causes  $Y$ , our basic assumption is  $Y = f(X, N_Y)$  where the noise  $N_Y \perp\!\!\!\perp X$ . In general, if one tries to model  $X = g(Y, N_X)$ , the estimated noise  $N_X$  would satisfy  $N_X \not\perp\!\!\!\perp Y$ . So the difference in  $N_Y \perp\!\!\!\perp X$  and  $N_X \not\perp\!\!\!\perp Y$  are the key to determine direction of causality. However, in case that  $X, N_Y$  are independent and normal, and the model are linear, then the reverse model  $N_X, Y$  are also independent, as pointed out in Theorem 1.1. Hence, in that case, although the causal effect is unidirectional in reality, one cannot distinguish between 2 directions using only joint pdf of  $X, Y$ .

## References

- [PJS17] Jonas Peters, Dominik Janzing, and Bernhard Schölkopf. *Elements of causal inference: foundations and learning algorithms*. The MIT Press, 2017.