## Chapter 7

# Probability.

#### 7.1 Probability Spaces.

**Definition 7.1.1.** A probability space defined by a tuple  $(\Omega, P)$  such that:

- 1.  $\Omega$  is a set, called the sample space. Any element  $\omega \in \Omega$  is named an atomic event. Conceptually, we think of atomic events as possible outcomes of our experiment. Any subset  $A \subset \Omega$  is an event.
- 2. P, called the probability function, is a function that assigns a number in [0,1] to any event, denoted as  $P:2^{\Omega} \to [0,1]$ , and satisfies:
  - (a) For any event  $A \subset \Omega$ ,  $P(A) = \sum_{w \in A} P(w)$ .
  - (b) Normalization, over the atomic events, to 1, which means  $\sum_{\omega \in \Omega} P(\omega) = 1$ .

#### **Example 7.1.1.** [COMMENT] Add dice roll, as an example.

**Claim 7.1.1.** *Probability function satisfies the following properties:* 

- 1.  $P(\emptyset) = 0$ .
- 2. Monotonic, If  $A \subset B \subset \Omega$  then  $P(A) \leq P(B)$ .
- 3. Union Bound,  $P(A \cup B) \leq P(A) + P(B)$ .
- 4. Additivity for disjointness events. If  $A \cap B = \emptyset$  then  $P(A \cup B) = P(A) + P(B)$ .
- 5. Denote by  $\bar{A}$  the complementary event of A, which means  $A \cup \bar{A} = \Omega$ . Then,  $P(\bar{A}) = 1 P(A)$ .

**Example 7.1.2.** Let's proof the additivity of disjointness property. Let A, B disjointness events, so  $A \cap B = \emptyset$  then

$$\begin{split} P(A \cup B) &= \sum_{w \in A \cup B} P(w) \\ &= \underbrace{\sum_{w \in A, w \notin B} P(w)}_{P(A)} + \underbrace{\sum_{w \in B, w \notin A} P(w)}_{P(B)} + \underbrace{\sum_{w \in A, w \in B} P(w)}_{Q(A)} \\ &= P(A) + P(B) \end{split}$$

**Definition 7.1.2.** Let  $(\Omega, P)$  be a probability space. A random variable X on  $(\Omega, P)$  is a function  $X: \Omega \to \mathbb{R}$ . An indicator, is a random variable defined by an event  $A \subset \Omega$  as follows

$$X(\omega) = \begin{cases} 1 & \omega \in A \\ 0 & \omega \notin A \end{cases}$$

Sometimes, we will use the notation  $\{X = x\}$  to denote the event A such:

$$A = \{\omega : X(\omega) = x\} := \{X = x\}$$

**Example 7.1.3.** [COMMENT] Add dice roll, as an example.

**Definition 7.1.3.** We will say that two random variable  $X,Y:\Omega\to\mathbb{R}$  are independent if for any  $x\in\operatorname{Im} X$  and  $y\in\operatorname{Im} Y$ :

$$P(X = x \cap Y = y) = P(X = x) \cdot P(Y = y)$$

### 7.2 [COMMENT] Throw Keys to Cells.

**[COMMENT]** Add the description of throwing keys to cells. Define the random variable  $X_i^j$ .

**Definition 7.2.1.** Let  $X : \Omega \to \mathbb{R}$  be a random variable, the expectation of X is

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega) P(\omega) = \sum_{x \in \operatorname{Im} X} x P(X = x)$$

Observes that if P is distributed uniformly, then the expectation of X is just the arithmetic mean:

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega) P(\omega) = \frac{1}{|\Omega|} \sum_{\omega \in \Omega} X(\omega)$$

**Claim 7.2.1.** *The expectation satisfies the following properties:* 

- 1. Monotonic, If  $X \leq Y$  (for any  $\omega \in \Omega$ ) then  $\mathbf{E}[X] \leq \mathbf{E}[Y]$ .
- 2. Linearity, for  $a, b \in \mathbb{R}$  it holds that  $\mathbf{E}[aX + by] = a\mathbf{E}[X] + b\mathbf{E}[Y]$ .
- 3. Independently, if X, Y are independent, then  $\mathbf{E}[X \cdot Y] = \mathbf{E}[X] \cdot \mathbf{E}[Y]$ .
- 4. For any constant  $a \in \mathbb{R}$  we have that  $\mathbf{E}[a] = a$ .

*Proof.* 1. Monotonic, if  $X \leq Y$  then:

$$\mathbf{E}\left[X\right] = \sum_{\omega \in \Omega} X(\omega) P(\omega) \le \sum_{\omega \in \Omega} Y(\omega) P(\omega) = \mathbf{E}\left[Y\right]$$

2. Linearity,

$$\mathbf{E}[aX + bY] = \sum_{\omega \in \Omega} (aX(\omega) + bY(\omega)) P(\omega)$$
$$= a \sum_{\omega \in \Omega} X(\omega) P(\omega) + b \sum_{\omega \in \Omega} Y(\omega) P(\omega)$$

3. Independently,

$$\begin{split} \mathbf{E}\left[XY\right] &= \sum_{x,y \in \operatorname{Im} X \times \operatorname{Im} Y} xy P(X = x \cap Y = y) \\ &= \sum_{x,y \in \operatorname{Im} X \times \operatorname{Im} Y} xy P(X = x) P(Y = y) \\ &= \sum_{x \in \operatorname{Im} X} \sum_{y \in \operatorname{Im} Y} xy P(X = x) P(Y = y) \\ &= \sum_{x \in \operatorname{Im} X} x P(X = x) \sum_{y \in \operatorname{Im} Y} y P(Y = y) \\ &= \sum_{x \in \operatorname{Im} X} x P(X = x) \mathbf{E}\left[Y\right] \\ &= \mathbf{E}\left[X\right] \mathbf{E}\left[Y\right] \end{split}$$

4. Let X be the random variable which is also the constant function  $X(\omega)=a$  for any  $\omega\in\Omega$ . Then we have that

$$\begin{split} \mathbf{E}\left[X\right] &= \sum_{\omega \in \Omega} X(\omega) P(\omega) \\ &= \sum_{\omega \in \Omega} a P(\omega) = a \cdot 1 = a \end{split}$$

**Example 7.2.1.** [COMMENT] Expectation of indicators and their multiplication.

**Example 7.2.2.** [COMMENT] How many keys trowed into the same cell as the first key thrown to?

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1 let B[0:n-1] be a new array 2 for i \leftarrow [0,n-1] do 3 | make B_i an empty list 4 end 5 for i \leftarrow [1,n] do 6 | insert A_i into list B_{\lfloor nA_i \rfloor}] 7 end 8 for i \leftarrow [0,n-1] do 9 | sort list B_i 10 end 11 concatenate the lists B_0,B_1,..,B_{n-1} together and 12 return the concatenated lists Algorithm 1: bucket-sort(A,n)
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Denote by  $X_i:[n] \to [n]$  then random variable that counts the number of elements fallen in the *i*th bucket. The Expectation of the sorting running time is:

$$\begin{split} \mathbf{E}\left[T\right] &= \mathbf{E}\left[\text{Inserting into buckets } + \sum_{i} \text{Sorting } i \text{th bucket} \right] \\ &= \mathbf{E}\left[\Theta(n) + \sum_{i} X_{i}^{2}\right] = \Theta(n) + \sum_{i} \mathbf{E}\left[X_{i}^{2}\right] \\ \mathbf{E}\left[X_{i}^{2}\right] &= \mathbf{E}\left[\left(\sum_{i} X_{i}^{j}\right)^{2}\right] = \mathbf{E}\left[\sum_{j,j'} X_{i}^{j} X_{i}^{j'}\right] \\ &= \sum_{j,j'} \mathbf{E}\left[X_{i}^{j} X_{i}^{j'}\right] = \sum_{j \neq j'} \mathbf{E}\left[X_{i}^{j} X_{i}^{j'}\right] + \sum_{j} \mathbf{E}\left[X_{i}^{j} X_{i}^{j}\right] \\ &= \sum_{j \neq j'} \mathbf{E}\left[X_{i}^{j} X_{i}^{j'}\right] + \sum_{j} \mathbf{E}\left[X_{i}^{j}\right] \\ &= 2\binom{n}{2}\left(\frac{1}{n}\right)^{2} + n \cdot \frac{1}{n} \\ &= \frac{n-1}{n} + 1 \end{split}$$