

# Chapter 7

## Probability.

### 7.1 Probability Spaces.

**Definition 7.1.1.** A probability space is defined by a tuple  $(\Omega, P)$ , where:

1.  $\Omega$  is a set, called the sample space. Any element  $\omega \in \Omega$  is an atomic event. Conceptually, we think of atomic events as possible outcomes of our experiment. Any subset  $A \subset \Omega$  is an event.
2.  $P$ , called the probability function, is a function that assigns a number in  $[0, 1]$  to any event, denoted as  $P : 2^\Omega \rightarrow [0, 1]$ , and satisfies:

(a) For any event  $A \subset \Omega$ ,  $P(A) = \sum_{\omega \in A} P(\omega)$ .

(b) Normalization over the atomic events to 1, which means  $\sum_{\omega \in \Omega} P(\omega) = 1$ .

**Example 7.1.1.** Consider a dice rolling, where each of the faces is indexed by 1, 2, 3, 4, 5, 6 and has an equal chance of being rolled. Therefore, our atomic events are associated with the rolling result, and  $P$  is defined as  $P(\omega) = \frac{1}{6}$  for any such atomic event. An example of an event can be  $A = \text{"the dice falls on an even number"}$ . The probability of this outcome is:

$$P(A) = \sum_{\omega \in A} P(\omega) = P(\{2\}) + P(\{4\}) + P(\{6\}) = 3 \cdot \frac{1}{6} = \frac{1}{2}$$

**Claim 7.1.1.** The probability function satisfies the following properties:

1.  $P(\emptyset) = 0$ .
2. Monotonicity: If  $A \subset B \subset \Omega$ , then  $P(A) \leq P(B)$ .
3. Union Bound:  $P(A \cup B) \leq P(A) + P(B)$ .
4. Additivity for disjoint events: If  $A \cap B = \emptyset$ , then  $P(A \cup B) = P(A) + P(B)$ .
5. Complementarity: Denote by  $\bar{A}$  the complementary event of  $A$ , which means  $A \cup \bar{A} = \Omega$ . Then,  $P(\bar{A}) = 1 - P(A)$ .

**Example 7.1.2.** Let's proof the additivity of disjointness property. Let  $A, B$  disjoint events, so  $A \cap B = \emptyset$  then

$$\begin{aligned} P(A \cup B) &= \sum_{w \in A \cup B} P(w) \\ &= \overbrace{\sum_{w \in A, w \notin B} P(w)}^{P(A)} + \overbrace{\sum_{w \in B, w \notin A} P(w)}^{P(B)} + \overbrace{\sum_{w \in A, w \in B} P(w)}^0 \\ &= P(A) + P(B) \end{aligned}$$

**Definition 7.1.2.** Let  $(\Omega, P)$  be a probability space. A random variable  $X$  on  $(\Omega, P)$  is a function  $X : \Omega \rightarrow \mathbb{R}$ . An indicator, is a random variable defined by an event  $A \subset \Omega$  as follows

$$X(\omega) = \begin{cases} 1 & \omega \in A \\ 0 & \omega \notin A \end{cases}$$

Sometimes, we will use the notation  $\{X = x\}$  to denote the event  $A$  such:

$$A = \{\omega : X(\omega) = x\} := \{X = x\}$$

**Example 7.1.3.** Consider rolling a pair of dice. Denote by  $X : [6] \times [6] \rightarrow [6]$  the random variable that is set to be the result of the first roll. Let  $Y$  be defined in almost the same way, but setting the result of the second die. Namely, if we denote by  $\{(i, j)\}$  the atomic event associated with sample  $i$  on the first die and  $j$  on the second die, then:

$$\begin{aligned} X(\{i, j\}) &= i \\ Y(\{i, j\}) &= j \end{aligned}$$

In addition, one can define the random variable  $z$  as the sum,  $Z = X + Y$ . Since the sum is also a function from  $\Omega$  to  $\mathbb{R}$ ,  $Z$  is also a random variable. An example of an indicator could be  $W$ , which gets 1 if  $Z \in \{2, 7, 8\}$ .

**Definition 7.1.3.** We will say that two random variable  $X, Y : \Omega \rightarrow \mathbb{R}$  are independent if for any  $x \in \text{Im } X$  and  $y \in \text{Im } Y$ :

$$P(X = x \cap Y = y) = P(X = x) \cdot P(Y = y)$$

**Example 7.1.4.**  $X, Y$  defined in Example 7.1.3 are independent.

$$\begin{aligned} P(\{X = i\} \cap \{Y = j\}) &= \sum_{i'=i \text{ and } j'=j} P(\{(i', j')\}) = P(\{(i, j)\}) \\ &= \frac{1}{36} = \frac{1}{6} \cdot \frac{1}{6} = P(X = i)P(Y = j) \end{aligned}$$

## 7.2 Throwing Keys to Cells.

**Example 7.2.1.** Imagine that following experiment, we have  $m$  cells and  $n$  keys (balls, numbers, or your favorite object type). We throw each of the keys independently into the cells. The cells are identical, so the probability of hitting any of them is the same,  $1/m$ . We would like to analyze how the capacity of the cells is distributed.

1. What is the probability that the first and the second keys will be thrown to the first cell? What is the probability that the first and the second keys will be thrown to the same cell?
2. What is the probability that in each cell there is at most one key?

Let us define the indicator  $X_i^j$  which indicate that the  $j$ th key fallen into the  $i$ th cell.

1. So first we been asked whether  $X_1^1 \cdot X_1^2 = 1$ , Since this happens only if both  $X_1^1 = 1, X_1^2 = 1$  then by independently we have that:

$$\begin{aligned} P(X_1^1 \cdot X_1^2 = 1) &= P(X_1^1 = 1 \cap X_1^2 = 1) \\ &= P(X_1^1 = 1) \cdot P(X_1^2 = 1) = \frac{1}{m^2} \end{aligned}$$

Now, if we would like to answer whether the first and the second fall into the same cell then we have to ask if there is an  $i$  such that  $X_i^1 \cdot X_i^2 = 1$ , also notice that for any different  $i, i'$  the  $X_i^j$  and  $X_{i'}^j$  are indicators of disjointness events, that because  $j$  cannot be fallen to both  $i, i'$  cells. Therefore  $X_i^1 \cdot X_i^2$  and  $X_{i'}^1 \cdot X_{i'}^2$  are also indicators of disjointness events. Thus:

$$\begin{aligned} P(\exists i : X_i^1 \cdot X_i^2 = 1) &= P\left(\bigcup_i X_i^1 \cdot X_i^2 = 1\right) = \sum_i P(X_i^1 \cdot X_i^2 = 1) \\ &= m \cdot \frac{1}{m^2} = \frac{1}{m} \end{aligned}$$

We basically done. Yet we want to present the same calculation in a different notation that later on will be useful for computing expectations. Observe that the random variable that counts to "how many" cells both the first and the second fall is  $\sum_i X_i^1 \cdot X_i^2$ .

2. Now we been asked whether there exists  $i$  such that  $X_j^1 \cdot X_j^2 = 1$
3. Now we been asked whether there exists  $i$  such that  $X_j^1 \cdot X_j^2 = 1$

**Definition 7.2.1.** Let  $X : \Omega \rightarrow \mathbb{R}$  be a random variable, the expectation of  $X$  is

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega)P(\omega) = \sum_{x \in \text{Im } X} xP(X = x)$$

Observe that if  $P$  is distributed uniformly, then the expectation of  $X$  is just the arithmetic mean:

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega)P(\omega) = \frac{1}{|\Omega|} \sum_{\omega \in \Omega} X(\omega)$$

**Claim 7.2.1.** The expectation satisfies the following properties:

1. Monotonic, If  $X \leq Y$  (for any  $\omega \in \Omega$ ) then  $\mathbf{E}[X] \leq \mathbf{E}[Y]$ .
2. Linearity, for  $a, b \in \mathbb{R}$  it holds that  $\mathbf{E}[aX + bY] = a\mathbf{E}[X] + b\mathbf{E}[Y]$ .
3. Independently, if  $X, Y$  are independent, then  $\mathbf{E}[X \cdot Y] = \mathbf{E}[X] \cdot \mathbf{E}[Y]$ .

4. For any constant  $a \in \mathbb{R}$  we have that  $\mathbf{E}[a] = a$ .

*Proof.* 1. Monotonic, if  $X \leq Y$  then :

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega)P(\omega) \leq \sum_{\omega \in \Omega} Y(\omega)P(\omega) = \mathbf{E}[Y]$$

2. Linearity,

$$\begin{aligned} \mathbf{E}[aX + bY] &= \sum_{\omega \in \Omega} (aX(\omega) + bY(\omega)) P(\omega) \\ &= a \sum_{\omega \in \Omega} X(\omega)P(\omega) + b \sum_{\omega \in \Omega} Y(\omega)P(\omega) \end{aligned}$$

3. Independently,

$$\begin{aligned} \mathbf{E}[XY] &= \sum_{x,y \in \text{Im } X \times \text{Im } Y} xyP(X = x \cap Y = y) \\ &= \sum_{x,y \in \text{Im } X \times \text{Im } Y} xyP(X = x)P(Y = y) \\ &= \sum_{x \in \text{Im } X} \sum_{y \in \text{Im } Y} xyP(X = x)P(Y = y) \\ &= \sum_{x \in \text{Im } X} xP(X = x) \sum_{y \in \text{Im } Y} yP(Y = y) \\ &= \sum_{x \in \text{Im } X} xP(X = x)\mathbf{E}[Y] \\ &= \mathbf{E}[X]\mathbf{E}[Y] \end{aligned}$$

4. Let  $X$  be the random variable which is also the constant function  $X(\omega) = a$  for any  $\omega \in \Omega$ . Then we have that

$$\begin{aligned} \mathbf{E}[X] &= \sum_{\omega \in \Omega} X(\omega)P(\omega) \\ &= \sum_{\omega \in \Omega} aP(\omega) = a \cdot 1 = a \end{aligned}$$

□

**Example 7.2.2.** Let  $X$  be an indicator of event  $A$ , what are  $\mathbf{E}[X]$  and  $\mathbf{E}[X^2]$ ? Recall that  $X(\omega) = 1$  only if  $\omega \in A$  and 0 otherwise, thus:

$$X^k(\omega) = \begin{cases} 1^k = 1 & \omega \in A \\ 0^k = 0 & \text{else} \end{cases} \Rightarrow X^k(\omega) = X(\omega)$$

Therefore,

$$\mathbf{E}[X^k] = \sum_{\omega \in \Omega} X^k(\omega)P(\omega) = \sum_{\omega \in \Omega} X(\omega)P(\omega) = \mathbf{E}[X]$$

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1 let B[0 : n - 1] be a new array
2 for i ← [0, n-1] do
3   | make Bi an empty list
4 end
5 for i ← [1, n] do
6   | insert Ai into list B[nAi]
7 end
8 for i ← [0, n-1] do
9   | sort list Bi
10 end
11 concatenate the lists B0, B1, ..., Bn-1 together and
12 return the concatenated lists

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**Algorithm 1:** bucket-sort( $A, n$ )

**Example 7.2.3.** *[COMMENT] How many keys trowed into the same cell as the first key thrown to?*

Denote by  $X_i : [n] \rightarrow [n]$  then random variable that counts the number of elements fallen in the  $i$ th bucket. The Expectation of the sorting running time is:

$$\begin{aligned}
\mathbf{E}[T] &= \mathbf{E} \left[ \text{Inserting into buckets} + \sum_i \text{Sorting } i\text{th bucket} \right] \\
&= \mathbf{E} \left[ \Theta(n) + \sum_i X_i^2 \right] = \Theta(n) + \sum_i \mathbf{E}[X_i^2] \\
\mathbf{E}[X_i^2] &= \mathbf{E} \left[ \left( \sum_j X_i^j \right)^2 \right] = \mathbf{E} \left[ \sum_{j,j'} X_i^j X_i^{j'} \right] = \sum_{j,j'} \mathbf{E} [X_i^j X_i^{j'}] \\
&= \sum_{j \neq j'} \mathbf{E} [X_i^j X_i^{j'}] + \sum_j \mathbf{E} [X_i^j X_i^j] \\
&= \sum_{j \neq j'} \mathbf{E} [X_i^j X_i^{j'}] + \sum_j \mathbf{E} [X_i^j] \\
&= 2 \binom{n}{2} \left( \frac{1}{n} \right)^2 + n \cdot \frac{1}{n} \\
&= \frac{n-1}{n} + 1 = 2 - \frac{1}{n} \Rightarrow \mathbf{E}[T] = \Theta(n) + n \left( 2 - \frac{1}{n} \right) = \Theta(n)
\end{aligned}$$