## **Chapter 7**

## Probability.

## 7.1 Probability Spaces.

**Definition 7.1.1.** A probability space is defined by a tuple  $(\Omega, P)$ , where:

- 1.  $\Omega$  is a set, called the sample space. Any element  $\omega \in \Omega$  is an atomic event. Conceptually, we think of atomic events as possible outcomes of our experiment. Any subset  $A \subset \Omega$  is an event.
- 2. P, called the probability function, is a function that assigns a number in [0,1] to any event, denoted as  $P: 2^{\Omega} \to [0,1]$ , and satisfies:
  - (a) For any event  $A \subset \Omega$ ,  $P(A) = \sum_{\omega \in A} P(\omega)$ .
  - (b) Normalization over the atomic events to 1, which means  $\sum_{\omega \in \Omega} P(\omega) = 1$ .

**Example 7.1.1.** Consider a dice rolling, where each of the faces is indexed by 1, 2, 3, 4, 5, 6 and has an equal chance of being rolled. Therefore, our atomic events are associated with the rolling result, and P is defined as  $P(\omega) = \frac{1}{6}$  for any such atomic event. An example of an event can be A = "the dice falls on an even number". The probability of this outcome is:

$$P(A) = \sum_{\omega \in A} P(\omega) = P(\{2\}) + P(\{4\}) + P(\{6\}) = 3 \cdot \frac{1}{6} = \frac{1}{2}$$

**Claim 7.1.1.** The probability function satisfies the following properties:

- 1.  $P(\emptyset) = 0$ .
- 2. Monotonicity: If  $A \subset B \subset \Omega$ , then  $P(A) \leq P(B)$ .
- 3. Union Bound:  $P(A \cup B) \leq P(A) + P(B)$ .
- 4. Additivity for disjoint events: If  $A \cap B = \emptyset$ , then  $P(A \cup B) = P(A) + P(B)$ .
- 5. Complementarity: Denote by  $\bar{A}$  the complementary event of A, which means  $A \cup \bar{A} = \Omega$ . Then,  $P(\bar{A}) = 1 P(A)$ .

**Example 7.1.2.** Let's proof the additivity of disjointness property. Let A, B disjointness events, so  $A \cap B = \emptyset$  then

$$P(A \cup B) = \sum_{w \in A \cup B} P(w)$$

$$= \underbrace{\sum_{w \in A, w \notin B} P(w) + \sum_{w \in B, w \notin A} P(w) + \sum_{w \in A, w \in B} 0}_{P(w)} P(w)$$

$$= P(A) + P(B)$$

**Definition 7.1.2.** Let  $(\Omega, P)$  be a probability space. A random variable X on  $(\Omega, P)$  is a function  $X: \Omega \to \mathbb{R}$ . An indicator, is a random variable defined by an event  $A \subset \Omega$  as follows

$$X(\omega) = \begin{cases} 1 & \omega \in A \\ 0 & \omega \notin A \end{cases}$$

Sometimes, we will use the notation  $\{X = x\}$  to denote the event A such:

$$A = \{\omega : X(\omega) = x\} := \{X = x\}$$

**Example 7.1.3.** Consider rolling a pair of dice. Denote by  $X:[6] \times [6] \to [6]$  the random variable that is set to be the result of the first roll. Let Y be defined in almost the same way, but setting the result of the second die. Namely, if we denote by  $\{(i,j)\}$  the atomic event associated with sample i on the first die and j on the second die, then:

$$X(\{i,j\}) = i$$
$$Y(\{i,j\}) = j$$

In addition, one can define the random variable z as the sum, Z = X + Y. Since the sum is also a function from  $\Omega$  to  $\mathbb{R}$ , Z is also a random variable. An example of an indicator could be W, which gets 1 if  $Z \in \{2,7,8\}$ .

**Definition 7.1.3.** We will say that two random variable  $X,Y:\Omega\to\mathbb{R}$  are independent if for any  $x\in\operatorname{Im} X$  and  $y\in\operatorname{Im} Y$ :

$$P(X = x \cap Y = y) = P(X = x) \cdot P(Y = y)$$

**Example 7.1.4.** X, Y defined in Example 7.1.3 are independent.

$$\begin{split} P(\{X=i\} \cap \{Y=j\}) &= \sum_{i'=i \text{ and } j'=j} P(\{(i',j')\}) = P(\{(i,j)\}) \\ &= \frac{1}{36} = \frac{1}{6} \cdot \frac{1}{6} = P(X=i)P(Y=j) \end{split}$$

## 7.2 Throwing Keys to Cells.

Imagines that following experiment, we have m cells and n keys (balls, numbers, your favorite object type). We throw each of the keys independently to cells. We would like to analyze how capacity of the cells is distributed.

**Example 7.2.1.** 1. What is the probability that the first and the second keys will be thrown to the first cell?

- 2. What is the probability that the first and the second keys will be thrown to the same cell?
- 3. What is the probability that in each cell there is at most one key? Let us define the indicator  $X_i^j$  which indicate that the jth key fallen into the ith cell.
  - 1. So first we been asked whether  $X_1^1 \cdot X_1^2 = 1$
  - 2. Now we been asked whether there exists i such that  $X_j^1 \cdot X_j^2 = 1$
  - 3. Now we been asked whether there exists i such that  $X_i^1 \cdot X_i^2 = 1$

**Definition 7.2.1.** Let  $X : \Omega \to \mathbb{R}$  be a random variable, the expectation of X is

$$\mathbf{E}\left[X\right] = \sum_{\omega \in \Omega} X(\omega) P(\omega) = \sum_{x \in \operatorname{Im} X} x P(X = x)$$

Observes that if P is distributed uniformly, then the expectation of X is just the arithmetic mean:

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega) P(\omega) = \frac{1}{|\Omega|} \sum_{\omega \in \Omega} X(\omega)$$

**Claim 7.2.1.** *The expectation satisfies the following properties:* 

- 1. Monotonic, If  $X \leq Y$  (for any  $\omega \in \Omega$ ) then  $\mathbf{E}[X] \leq \mathbf{E}[Y]$ .
- 2. Linearity, for  $a, b \in \mathbb{R}$  it holds that  $\mathbf{E}[aX + by] = a\mathbf{E}[X] + b\mathbf{E}[Y]$ .
- 3. Independently, if X, Y are independent, then  $\mathbf{E}[X \cdot Y] = \mathbf{E}[X] \cdot \mathbf{E}[Y]$ .
- 4. For any constant  $a \in \mathbb{R}$  we have that  $\mathbf{E}[a] = a$ .

*Proof.* 1. Monotonic, if  $X \leq Y$  then:

$$\mathbf{E}\left[X\right] = \sum_{\omega \in \Omega} X(\omega) P(\omega) \le \sum_{\omega \in \Omega} Y(\omega) P(\omega) = \mathbf{E}\left[Y\right]$$

2. Linearity,

$$\begin{split} \mathbf{E}\left[aX + bY\right] &= \sum_{\omega \in \Omega} \left(aX(\omega) + bY(\omega)\right) P(\omega) \\ &= a \sum_{\omega \in \Omega} X(\omega) P(\omega) + b \sum_{\omega \in \Omega} Y(\omega) P(\omega) \end{split}$$

3. Independently,

$$\begin{split} \mathbf{E}\left[XY\right] &= \sum_{x,y \in \operatorname{Im} X \times \operatorname{Im} Y} xy P(X = x \cap Y = y) \\ &= \sum_{x,y \in \operatorname{Im} X \times \operatorname{Im} Y} xy P(X = x) P(Y = y) \\ &= \sum_{x \in \operatorname{Im} X} \sum_{y \in \operatorname{Im} Y} xy P(X = x) P(Y = y) \\ &= \sum_{x \in \operatorname{Im} X} x P(X = x) \sum_{y \in \operatorname{Im} Y} y P(Y = y) \\ &= \sum_{x \in \operatorname{Im} X} x P(X = x) \mathbf{E}\left[Y\right] \\ &= \mathbf{E}\left[X\right] \mathbf{E}\left[Y\right] \end{split}$$

4. Let X be the random variable which is also the constant function  $X(\omega)=a$  for any  $\omega\in\Omega$ . Then we have that

$$\mathbf{E}[X] = \sum_{\omega \in \Omega} X(\omega) P(\omega)$$
$$= \sum_{\omega \in \Omega} a P(\omega) = a \cdot 1 = a$$

**Example 7.2.2.** Let X be an indicator of event A, what are  $\mathbf{E}[X]$  and  $\mathbf{E}[X^2]$ ? Recall that  $X(\omega) = 1$  only if  $\omega \in A$  and 0 otherwise, thus:

$$X^{k}(\omega) = \begin{cases} 1^{k} = 1 & \omega \in A \\ 0^{k} = 0 & \textit{else} \end{cases} \Rightarrow X^{k}(\omega) = X(\omega)$$

Therefore,

$$\mathbf{E}\left[X^{k}\right] = \sum_{\omega \in \Omega} X^{k}(\omega) P(\omega) = \sum_{\omega \in \Omega} X(\omega) P(\omega) = \mathbf{E}\left[X\right]$$

**Example 7.2.3.** [COMMENT] How many keys trowed into the same cell as the first key thrown to?

Denote by  $X_i : [n] \to [n]$  then random variable that counts the number of elements fallen in the *i*th bucket. The Expectation of the sorting running time is:

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1 let B[0:n-1] be a new array

2 for i \leftarrow [0,n-1] do

3 | make B_i an empty list

4 end

5 for i \leftarrow [1,n] do

6 | insert A_i into list B_{\lfloor nA_i \rfloor}]

7 end

8 for i \leftarrow [0,n-1] do

9 | sort list B_i

10 end

11 concatenate the lists B_0, B_1, ..., B_{n-1} together and

12 return the concatenated lists

Algorithm 1: bucket-sort(A, n)
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$$\begin{split} \mathbf{E}\left[T\right] &= \mathbf{E}\left[\text{ Inserting into buckets } + \sum_{i} \text{Sorting } i \text{th bucket}\right] \\ &= \mathbf{E}\left[\Theta(n) + \sum_{i} X_{i}^{2}\right] = \Theta(n) + \sum_{i} \mathbf{E}\left[X_{i}^{2}\right] \\ \mathbf{E}\left[X_{i}^{2}\right] &= \mathbf{E}\left[\left(\sum_{i} X_{i}^{j}\right)^{2}\right] = \mathbf{E}\left[\sum_{j,j'} X_{i}^{j} X_{i}^{j'}\right] = \sum_{j,j'} \mathbf{E}\left[X_{i}^{j} X_{i}^{j'}\right] \\ &= \sum_{j \neq j'} \mathbf{E}\left[X_{i}^{j} X_{i}^{j'}\right] + \sum_{j} \mathbf{E}\left[X_{i}^{j} X_{i}^{j}\right] \\ &= \sum_{j \neq j'} \mathbf{E}\left[X_{i}^{j} X_{i}^{j'}\right] + \sum_{j} \mathbf{E}\left[X_{i}^{j}\right] \\ &= 2\binom{n}{2}\left(\frac{1}{n}\right)^{2} + n \cdot \frac{1}{n} \\ &= \frac{n-1}{n} + 1 = 2 - \frac{1}{n} \Rightarrow \mathbf{E}\left[T\right] = \Theta(n) + n\left(2 - \frac{1}{n}\right) = \Theta(n) \end{split}$$