

Econometrics: Multiple Regression and Applications

ECON4004: LAB 4

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Intro

- ◇ Duong Trinh
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 - ◇ Email: Duong.Trinh@glasgow.ac.uk
- ◇ ECON4004-LB01
 - ◇ Wednesday 10am -12 pm
 - ◇ 5 sessions (7-Feb, 14-Feb, 21-Feb, 28-Feb, 6-March)
 - ◇ ST ANDREWS:357
- ◇ ECON4004-LB02
 - ◇ Wednesday 12-2 pm
 - ◇ 5 sessions (7-Feb, 14-Feb, 21-Feb, 28-Feb, 6-March)
 - ◇ ST ANDREWS:357

Record Attendance

Plan for LAB 3

- ◇ Exercise 1: based on Stock & Watson, E10.1
- ◇ Exercise 2: based on Stock & Watson, E10.2

- ◇ We will focus on *“Panel Data - Fixed Effects Regressions”*

BRIEF REVIEW

Panel Data - What it looks like...

Panel data is a dataset in which the behavior of entities (i) are observed across time (t).

$(X_{it}, Y_{it}),$

$i = 1, \dots, n; t = 1, \dots, T$

These entities could be states, companies, families, individuals, countries, etc.

| Entity | Year | Y | X1 | X2 | X3 | |
|--------|------|---|----|----|----|-------|
| 1 | 1 | # | # | # | # | |
| 1 | 2 | # | # | # | # | |
| 1 | 3 | # | # | # | # | |
| : | : | : | : | : | : | : |
| 2 | 1 | # | # | # | # | |
| 2 | 2 | # | # | # | # | |
| 2 | 3 | # | # | # | # | |
| : | : | : | : | : | : | : |
| 3 | 1 | # | # | # | # | |
| 3 | 2 | # | # | # | # | |
| 3 | 3 | # | # | # | # | |

Panel Data - The Long Form

Preparing data into
panel data format:

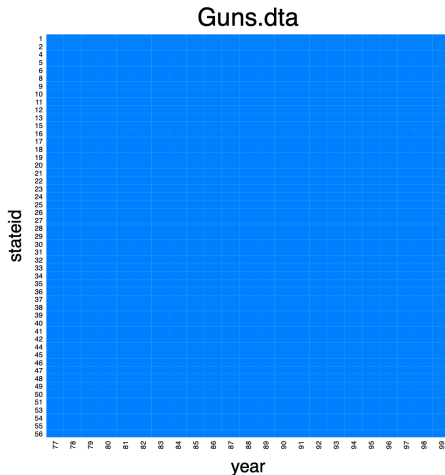
Entity and Time in rows.
&
Variables in columns.

| Entity | Year | Y | X1 | X2 | X3 | |
|--------|------|---|----|----|----|------|
| 1 | 1 | # | # | # | # | |
| 1 | 2 | # | # | # | # | |
| 1 | 3 | # | # | # | # | |
| : | : | : | : | : | : | : |
| 2 | 1 | # | # | # | # | |
| 2 | 2 | # | # | # | # | |
| 2 | 3 | # | # | # | # | |
| : | : | : | : | : | : | : |
| 3 | 1 | # | # | # | # | |
| 3 | 2 | # | # | # | # | |
| 3 | 3 | # | # | # | # | |

Balanced Panel

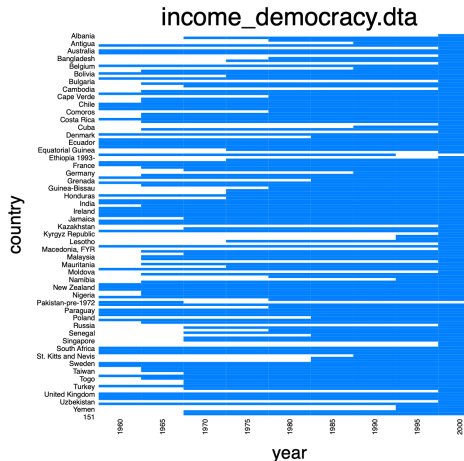
All entities are observed
across all times.

51 states \times 23 years
= 1173 observations



Unbalanced Panel

Some entities are not observed in some time periods.



[SN] STATA command for Setting Data as Panel

Once the data is in long form, we need to set it as panel so we can use Stata's panel data xt commands.

```
*xtset entityid timeid  
//'entityid' and 'timeid' have to be in numeric format
```

```
. use "Guns.dta", clear
```

```
. xtset stateid year
```

Panel variable: **stateid** (strongly balanced)

Time variable: **year**, 77 to 99

Delta: 1 unit

```
. use "income_democracy.dta", clear
```

```
. xtset code year
```

Panel variable: **code** (unbalanced)

Time variable: **year**, 1960 to 2000, but with gaps

Delta: 1 unit

Note: id means unique identifiers for entity (entityid) or for time period (timeid)

[SN] STATA command for Visualizing Panel Data

Once the data is set as panel, we can use a series of xt commands to analyze it.

For visualization

```
*xtline varname_of_intetereast  
//graphs by entities
```

```
*xtline varname_of_intetereast, overlay legend(off)  
//all entities in one graph
```

Fixed Effects Regressions - General Form

For $i = 1, \dots, n$ and $t = 1, \dots, T$

$$Y_{it} = \overbrace{\beta_1 X_{1,it} + \beta_2 X_{2,it} + \dots + \beta_k X_{k,it}}^{\text{k regressor}} + \underbrace{\alpha_i}_{\substack{\text{entity} \\ \text{fixed effects}}} + \underbrace{\lambda_t}_{\substack{\text{time} \\ \text{fixed effects}}} + u_{it}$$

(3) Both E&T Fixed Effects: $Y_{it} = \beta_1 X_{it} + \underbrace{\alpha_i}_{\substack{\text{entity} \\ \text{fixed effects}}} + \underbrace{\lambda_t}_{\substack{\text{time} \\ \text{fixed effects}}} + u_{it}$

(1) Entity Fixed Effects: $Y_{it} = \beta_1 X_{it} + \alpha_i + u_{it}$

(2) Time Fixed Effects: $Y_{it} = \beta_1 X_{it} + \lambda_t + u_{it}$

(1) Entity Fixed Effects (Time-invariant)

$$Y_{it} = \beta_1 X_{it} + \underbrace{\alpha_i}_{\text{n entity-specific intercepts}} + u_{it}$$

\Downarrow

$$Y_{it} = \underbrace{\beta_0}_{\text{an intercept}} + \beta_1 X_{it} + \underbrace{\gamma_2 D2_i + \gamma_3 D3_i + \dots + \gamma_n Dn_i}_{\text{with n-1 entity binary indicators}} + u_{it}$$

Entity 1: $\alpha_1 = \beta_0$

Entity 2: $\alpha_2 = \beta_0 + \gamma_2$

Entity 3: $\alpha_3 = \beta_0 + \gamma_3$

\vdots

Entity n: $\alpha_n = \beta_0 + \gamma_n$

Note: This form suggests estimating regression model with $n - 1$ binary indicators by OLS, but we can also use “entity-demeaned” OLS algorithm.

(2) Time Fixed Effects (Entity-invariant)

$$Y_{it} = \beta_1 X_{it} + \overbrace{\lambda_t}^{\text{T time-specific intercepts}} + u_{it}$$

\Downarrow

$$Y_{it} = \underbrace{\beta_0}_{\text{an intercept}} + \beta_1 X_{it} + \underbrace{\delta_2 B2_t + \delta_3 B3_t + \dots + \delta_T B T_t}_{\text{with n-1 time binary indicators}} + u_{it}$$

Time period **1**: $\lambda_1 = \beta_0$

Time period **2**: $\lambda_2 = \beta_0 + \delta_2$

Time period **3**: $\lambda_3 = \beta_0 + \delta_3$

\vdots

Time period **T**: $\lambda_T = \beta_0 + \delta_T$

Note: This form suggests estimating regression model with $T - 1$ binary indicators by OLS, but we can also use “time-demeaned” OLS algorithm.

(3) Both Entity and Time (Two-way) Fixed Effects

$$Y_{it} = \beta_1 X_{it} + \overbrace{\alpha_i}^{\text{entity fixed effects}} + \overbrace{\lambda_t}^{\text{time fixed effects}} + u_{it}$$

\Downarrow

$$Y_{it} = \underbrace{\beta_0}_{\text{an intercept}} + \beta_1 X_{it} + \underbrace{\gamma_2 D2_i + \gamma_3 D3_i + \dots + \gamma_n Dn_i}_{\text{with } n-1 \text{ entity binary indicators}} + \underbrace{\delta_2 B2_t + \delta_3 B3_t + \dots + \delta_T BT_t}_{\text{with } n-1 \text{ time binary indicators}} + u_{it}$$

Note: This form suggests estimating regression model with $n - 1$ entity binary indicators and $T - 1$ time binary indicators by OLS, but we can also combine with “demeaned” OLS algorithm.

[SN] STATA syntax for Fixed Effects Regressions

(1) Entity Fixed Effects

```
*xtreg yvar xvar, fe vce(cluster entityid)
```

(2) Time Fixed Effects

```
*regress yvar xvar i.timeid, fe vce(cluster entityid)
```

(3) Both Entity & Time Fixed Effects

```
*xtreg yvar xvar i.timeid, fe vce(cluster entityid)
```

Note: id means unique identifiers for entity (entityid) or for time period (timeid)

Clustered Standard Errors

| | $t = 1$ | $t = 2$ | \dots | $t = T$ |
|----------|----------|----------|---------|----------|
| $i = 1$ | u_{11} | u_{12} | \dots | u_{1T} |
| $i = 2$ | u_{21} | u_{22} | \dots | u_{2T} |
| \vdots | \vdots | \vdots | \dots | \vdots |
| $i = n$ | u_{n1} | u_{n2} | \dots | u_{nT} |

- ◇ Sampling are *i.i.d* across entities.
- ◇ But if the omitted factors comprising the error term u_{it} are serially correlated, then u_{it} is serially correlated aka *autocorrelated* - that is, correlated over time within an entity.

Standard errors need to allow both for this autocorrelation and for potential heteroskedasticity → use **clustered standard errors**.

```
*[, vce(cluster entityid)]  
//add this option at the end of STATA regressions
```

Exercise 1: based on Stock & Watson, E10.1

Excercise 1: based on Stock & Watson, E10.1

- ◇ **Objective:** Analyze effects of *concealed weapons laws* on *violent crimes*.
 - ◇ [Proponents:] More people carry concealed weapons, crime will decline because criminals will be deterred from attacking other people.
 - ◇ [Opponents:] Crime will increase because of accidental or spontaneous use of the weapons.

- ◇ **Dataset:** `Guns.dta`
 - ◇ A balanced panel of data from the 50 U.S. states plus the District of Columbia for 23 years (1977-1999).

- ◇ **Key variables:**
 - ◇ `stateid`: ID number of states (Alabama = 1, Alaska = 2, etc.)
 - ◇ `year`: year (1977-1999)
 - ◇ `shall`: = 1 if the state has a shall-carry law in effect in that year = 0 otherwise
 - ◇ `vio`: violent crime rate.

Variable Description

| Variable | Definition |
|--------------------|---|
| <i>vio</i> | violent crime rate (incidents per 100,000 members of the population) |
| <i>rob</i> | robbery rate (incidents per 100,000) |
| <i>mur</i> | murder rate (incidents per 100,000) |
| <i>shall</i> | = 1 if the state has a shall-carry law in effect in that year = 0 otherwise |
| <i>incarc rate</i> | incarceration rate in the state in the previous year (sentenced prisoners per 100,000 residents; value for the previous year) |
| <i>density</i> | population per square mile of land area, divided by 1000 |
| <i>avginc</i> | real per capita personal income in the state, in thousands of dollars |
| <i>pop</i> | state population, in millions of people |
| <i>pm1029</i> | percent of state population that is male, ages 10 to 29 |
| <i>pw1064</i> | percent of state population that is white, ages 10 to 64 |
| <i>pb1064</i> | percent of state population that is black, ages 10 to 64 |
| <i>stateid</i> | ID number of states (Alabama = 1, Alaska = 2, etc.) |
| <i>year</i> | Year (1977-1999) |

Source: Ayres, I. and Donohue, J.J., 2002. Shooting down the more guns, less crime hypothesis.

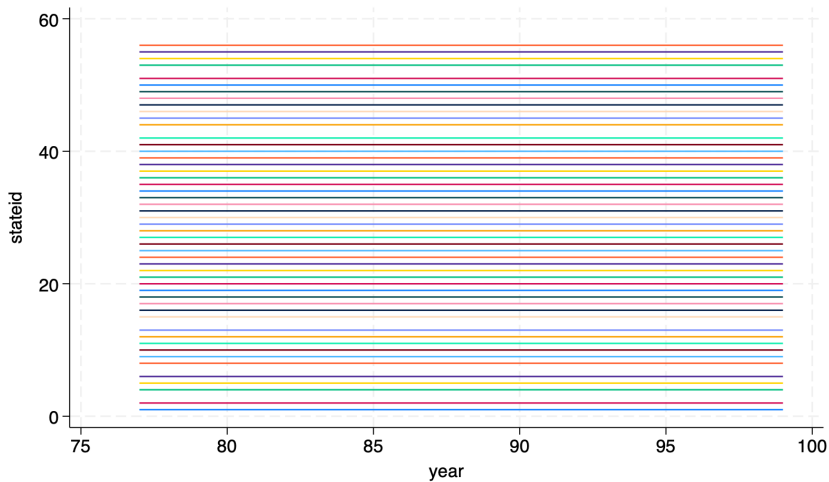
```
. xtset stateid year
```

Panel variable: **stateid** (strongly balanced)

Time variable: **year**, 77 to 99

Delta: 1 unit

```
. xtline stateid, overlay legend(off)
```



Questions

(a) Estimate a linear regression model of:

- ◇ (M1) $\ln(\text{vio})$ against shall.
- ◇ (M2) $\ln(\text{vio})$ against shall, `incarc_rate`, `density`, `avginc`, `pop`, `pb1064`, `pw1064` and `pm1029`.
 - i. Interpret the coefficient on shall in (M2).
Is this estimate large or small in a real-world sense?
 - ii. Does adding the control variables in (M2) change the estimated effect of a shall-issue law in (M1) as measured by statistical significance? As measured by the real-world significance of the estimated coefficient?
 - iii. Suggest a variable that varies across states but plausibly varies little or not at all over time and that could cause omitted variable bias in (M2).

Questions

- (b) Do the results change when you add *state fixed effects*?
If so, which set of regression results is more credible, and why?
- (c) Do the results change when you add *time fixed effects*?
If so, which set of regression results is more credible, and why?
- (d) Repeat the analysis using $\ln(\text{rob})$ and $\ln(\text{mur})$ in place of $\ln(\text{vio})$.

(a-i,ii) Estimate linear regression models (M1) and (M2)

From OLS estimation results for (M1) (»stata) and (M2) (»stata)

- ◇ The coefficient is -0.368 , which suggests that shall-issue laws reduce violent crime by 36%. This is a large effect.
- ◇ The coefficient in (1) is -0.443 ; in (2) it is -0.368 . Both are highly statistically significant. Adding the control variables results in a small drop in the coefficient.

(a-iii) Suggest a variable that varies across states but plausibly varies little or not at all over time and that could cause omitted variable bias in (M2).

There are several examples:

- ◇ Residents' attitudes towards guns and crime, these are typically slow to change.
- ◇ Quality of police and other crime-prevention programs.
- ◇ For historical reasons, cities can have very different crime rates.
- ◇ Geographic features.
- ◇ etc.

⇒ These factors constitute an *unobserved state effect*/ *state fixed effect*/ *unobserved state heterogeneity*. If they and the variable of interest (shall) are correlated, omitting such variables results in OVB (**heterogeneity bias**).

(b) Do the results change when you add *state fixed effects*?

Good news: These time-invariant factors can be effectively captured by the state fixed effect a_i ! (»review)

(b) Do the results change when you add *state fixed effects*?

```
*xtreg yvar xvar, fe vce(cluster entityid)
```

```
. xtreg lvio shall $basevars, fe vce(cluster stateid)
```

```
Fixed-effects (within) regression      Number of obs   =    1,173
Group variable: stateid              Number of groups =     51
```

```
R-squared:                            Obs per group:
    Within = 0.2178                      min =      23
    Between = 0.0033                     avg =     23.0
    Overall = 0.0001                     max =      23
```

```
corr(u_i, Xb) = -0.3687                F(8, 50)        =    34.10
                                          Prob > F        =    0.0000
```

(Std. err. adjusted for 51 clusters in **stateid**)

| lvio | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-------------|-------------|---------------------|-------|-------|----------------------|-----------|
| shall | -.0461415 | .0417616 | -1.10 | 0.275 | -.1300223 | .0377392 |
| incarc_rate | -.000071 | .0002504 | -0.28 | 0.778 | -.0005739 | .0004318 |
| density | -.1722901 | .1376129 | -1.25 | 0.216 | -.4486936 | .1041135 |
| avginc | -.0092037 | .0129649 | -0.71 | 0.481 | -.0352445 | .016837 |
| pop | .0115247 | .014224 | 0.81 | 0.422 | -.0170452 | .0400945 |
| pb1064 | .1042804 | .0326849 | 3.19 | 0.002 | .0386308 | .1699301 |
| pw1064 | .0408611 | .0134585 | 3.04 | 0.004 | .0138289 | .0678932 |
| pm1029 | -.0502725 | .0206949 | -2.43 | 0.019 | -.0918394 | -.0087057 |
| _cons | 3.866017 | .7701057 | 5.02 | 0.000 | 2.319214 | 5.412819 |

(b) Do the results change when you add *state fixed effects*?

- ◇ In (M3) the coefficient on shall falls to -0.046 , a large reduction in the coefficient from (M2). The estimate is not statistically significantly different from zero.
- ◇ Evidently there was important omitted variable bias leading to a spurious effect in (M2).
- ◇ The constant reported in `xtreg, fe` results implies the estimated average fixed effect.

Like experiments → Let's compare:

```
reg lvio shall i.stateid
```

```
xtreg lvio shall, fe
```

(c) Do the results change when you add *time fixed effects*?

```
*xtreg yvar xvar i.timeid, fe vce(cluster entityid)
```

```
. xtreg lvio shall $basevars i.year, fe vce(cluster stateid)
```

Fixed-effects (within) regression Number of obs = 1,173
Group variable: **stateid** Number of groups = 51

R-squared: Obs per group:

| | |
|--------------------|------------------------|
| Within = 0.4180 | min = 23 |
| Between = 0.0419 | avg = 23.0 |
| Overall = 0.0009 | max = 23 |

corr(u_i, Xb) = -0.2929 F(30, 50) = 56.86
 Prob > F = 0.0000

(Std. err. adjusted for 51 clusters in **stateid**)

| lvio | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-------------|-------------|---------------------|-------|-------|----------------------|----------|
| shall | -.0279935 | .0407168 | -0.69 | 0.495 | -.1097757 | .0537886 |
| incarc_rate | .000076 | .0002079 | 0.37 | 0.716 | -.0003416 | .0004935 |
| density | -.091555 | .1238622 | -0.74 | 0.463 | -.3403396 | .1572296 |
| avginc | .0009587 | .0164931 | 0.06 | 0.954 | -.0321688 | .0340861 |
| pop | -.0047544 | .0152294 | -0.31 | 0.756 | -.0353436 | .0258347 |
| pb1064 | .0291862 | .0495407 | 0.59 | 0.558 | -.0703192 | .1286916 |
| pw1064 | .0092501 | .0237564 | 0.39 | 0.699 | -.0384659 | .0569662 |
| pm1029 | .0733254 | .0524733 | 1.40 | 0.168 | -.0320704 | .1787211 |
| year | | | | | | |
| 78 | .0585261 | .0161556 | 3.62 | 0.001 | .0260767 | .0909755 |
| 79 | .1639486 | .0244579 | 6.70 | 0.000 | .1148233 | .2130738 |

(c) Do the results change when you add *time fixed effects*?

```
. quiet xtreg lvio shall $basevars i.year, fe vce(cluster stateid)
```

```
. _ms_extract_varlist i.year
```

```
. test `r(varlist)'
```

```
( 1) 77b.year = 0
( 2) 78.year = 0
( 3) 79.year = 0
( 4) 80.year = 0
( 5) 81.year = 0
( 6) 82.year = 0
( 7) 83.year = 0
( 8) 84.year = 0
( 9) 85.year = 0
(10) 86.year = 0
(11) 87.year = 0
(12) 88.year = 0
(13) 89.year = 0
(14) 90.year = 0
(15) 91.year = 0
(16) 92.year = 0
(17) 93.year = 0
(18) 94.year = 0
(19) 95.year = 0
(20) 96.year = 0
(21) 97.year = 0
(22) 98.year = 0
(23) 99.year = 0
```

```
Constraint 1 dropped
```

```
F( 22, 50) = 21.62
```

```
Prob > F = 0.0000
```

(c) Do the results change when you add *time fixed effects*?

- ◇ The coefficient falls further to -0.028 . The coefficient is insignificantly different from zero.
- ◇ The time effects are jointly statistically significant, so this regression seems better specified than (M3).

Table of Results - Exercise 1

| Regressor | Models | | | |
|----------------------|--|--|-------------------------------------|-------------------------------------|
| | (M1) | (M2) | (M3) | (M4) |
| <i>shall</i> | -0.443*** (0.157) [-0.76, -0.13] | -0.368*** (0.114) [-0.60, -0.14] | -0.0461 (0.042) [-0.13, 0.04] | -0.0280 (0.041) [-0.11, 0.05] |
| <i>Controls</i> | No | Yes | Yes | Yes |
| <i>State effects</i> | No | No | Yes | Yes |
| <i>Time effects</i> | No | No | No | Yes |

Notes: Clustered standard errors shown in parentheses and 95% confidence intervals are shown in brackets; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

Excercise 2: based on Stock & Watson, E10.2

Picture the Scenario

- ◇ **Objective:** Do citizens demand more democracy and political freedom as their incomes grow? That is, is democracy a normal good?
- ◇ **Dataset:** Income-Democracy.dta
 - ◇ a panel data set from 195 countries for the years 1960, 1965, ..., 2000.
- ◇ **Key variables:** For each country in each year
 - ◇ Dem_ind: an index of political freedom/democracy.
 - ◇ Log_GDPPC: per capita income.
 - ◇ various demographic controls.

Variable Description

| Variable Name | Description |
|-------------------|--|
| <i>country</i> | country name |
| <i>year</i> | year |
| <i>dem_ind</i> | index of democracy |
| <i>log_gdppc</i> | logarithm of real GDP per capita |
| <i>log_pop</i> | logarithm of population |
| <i>age_1</i> | fraction of the population age 0-14 |
| <i>age_2</i> | fraction of the population age 15-29 |
| <i>age_3</i> | fraction of the population age 30-44 |
| <i>age_4</i> | fraction of the population age 45-59 |
| <i>age_5</i> | fraction of the population age 60 and older |
| <i>educ</i> | average years of education for adults (25 years and older) |
| <i>age_median</i> | median age |
| <i>code</i> | country code |

Source: Acemoglu, Daron, Simon Johnson, James A. Robinson, and Pierre Yared. 2008. "Income and Democracy." *American Economic Review*, 98 (3): 808-42.

Notes: The income and demographic variable are lagged five years.

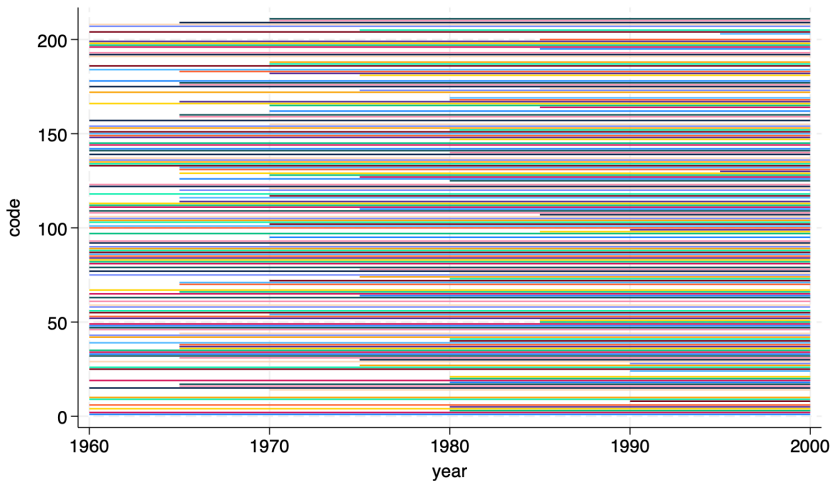
```
. xtset code year
```

Panel variable: **code** (unbalanced)

Time variable: **year**, 1960 to 2000, but with gaps

Delta: 1 unit

```
. xtline code, overlay legend(off)
```



Questions

- (a) Is the data set a balanced panel? Explain.
- (b) The logarithm of per capita income is labeled `Log_GDPPC`. Regress `Dem_ind` on `Log_GDPPC`. Use standard errors that are clustered by country.
 - i. How large is the estimated coefficient on `Log_GDPPC`?
Is the coefficient statistically significant?
 - ii. If per capita income in a country increases by 20%, by how much is `Dem_ind` predicted to increase?
What is a 95% confidence interval for the prediction?
Is the predicted increase in `Dem_ind` large or small? (Explain what you mean by large or small.)
 - iii. Why is it important to use clustered standard errors for the regression?
Do the results change if you do not use clustered standard errors?

Questions

- (c) Estimate the regression in (b), allowing for country fixed effects. How do your answers to (b-i) change?
- (d) Estimate the regression in (b), allowing for time and country fixed effects. How do your answers to (b-i) change?
- (e) There are additional demographic controls in the data set. Should these variables be included in the regression? If so, how do the results change when they are included?

(a) Is the data set a balanced panel? Explain.

- ◇ The dataset is unbalanced because data are available over different years for different countries. For example, data on `Dem_ind` are available
 - ◇ for Andorra during 1970, 1995, and 2000;
 - ◇ for Afghanistan during 1960, 1965, ..., 2000.

(b) Regress Dem_ind on Log_GDPPC using standard errors clustered by country.

```
. reg dem_ind log_gdppc, vce(cluster code) //using clustered standard errors
```

```
Linear regression               Number of obs   =       958
                               F(1, 149)         =     396.40
                               Prob > F           =     0.0000
                               R-squared          =     0.4385
                               Root MSE       =     .2719
```

(Std. err. adjusted for 150 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|--------|-------|----------------------|-----------|
| log_gdppc | .2356731 | .011837 | 19.91 | 0.000 | .212283 | .2590632 |
| _cons | -1.354828 | .1004215 | -13.49 | 0.000 | -1.553262 | -1.156394 |

(b-i) How large is the estimated coefficient on Log_GDPPC?

The coefficient is 0.236 with a standard error of 0.012. The 95% confidence interval is 0.212 to 0.259. The coefficient is statistically significant.

(b-ii) If per capita income in a country increases by 20%, by how much is Dem_ind predicted to increase?

Hint: Review Regression models with functional forms involving logarithms.

- ◇ A 20% increase in GDP per capita implies that \log_gdp increases by approximately 0.20, so that Dem_ind is predicted to increase by approximately $0.20 \times 0.236 = 0.0472$, or about 1/8 of the standard deviation in the dataset.
- ◇ The 95% confidence for the effect is (approximately) 0.20×0.212 to $0.20 \times 0.259 = 0.0424$ to 0.0518.

(b-iii) Why is it important to use clustered standard errors for the regression?

```
. reg dem_ind log_gdppc, vce(cluster code) //using clustered standard errors
```

```
Linear regression               Number of obs   =      958
                                F(1, 149)      =     396.40
                                Prob > F         =     0.0000
                                R-squared        =     0.4385
                                Root MSE     =     .2719
```

(Std. err. adjusted for 150 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|--------|-------|----------------------|-----------|
| log_gdppc | .2356731 | .011837 | 19.91 | 0.000 | .212283 | .2590632 |
| _cons | -1.354828 | .1004215 | -13.49 | 0.000 | -1.553262 | -1.156394 |

```
. reg dem_ind log_gdppc, robust //without clustering
```

```
Linear regression               Number of obs   =      958
                                F(1, 956)      =     1100.57
                                Prob > F         =     0.0000
                                R-squared        =     0.4385
                                Root MSE     =     .2719
```

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|--------|-------|----------------------|----------|
| log_gdppc | .2356731 | .007104 | 33.17 | 0.000 | .221732 | .2496143 |
| _cons | -1.354828 | .0607549 | -22.30 | 0.000 | -1.474056 | -1.2356 |

(b-iii) Why is it important to use clustered standard errors for the regression?

- ◇ Clustered standard errors are needed because of country-specific omitted factors in the regressions.
- ◇ The clustered standard error for `Dem_ind` is 0.012; the unclustered standard error is smaller (0.007) because it ignores the positive within-country autocorrelation of the errors.

(c) Estimate the regression in (b), allowing for country fixed effects.

Estimation result ([»stata](#))

The estimated coefficient falls by a factor of 3, to 0.083 with a standard error of 0.032. The estimated effect, while significantly smaller is still statistically significant at the 1% significance level.

(d) Estimate the regression in (b), allowing for time and country fixed effects.

Estimation and test results ([»stata](#))

The estimated coefficient falls further to 0.054, approximately 1/5 of the value that omits time and country fixed effects. The estimate is not statistically significant at the 10% level.

(e) Include additional demographic controls in the regression

Estimation and test results ([»stata](#))

When age, population, and education are included, the estimated coefficient on `log_gdppc` falls further to 0.025 with a standard error of 0.054. Jointly, these variables are not statistically significant in the regression, although the age variables are significant at the 10% level.

Table of Results - Exercise 2

| Regressor | Models | | | | |
|---|---------------------------------------|---------------------------------------|---------------------------------------|-------------------------------------|-------------------------------------|
| | (M1) | (M2) | (M3) | (M4) | (M5) |
| <i>log_GDPPC</i> | 0.236*** (0.012) [0.212, 0.259] | 0.235*** (0.012) [0.211, 0.259] | 0.083*** (0.031) [0.021, 0.146] | 0.054 (0.042) [-0.030, 0.137] | 0.025 (0.054) [-0.057, 0.120] |
| <i>Controls</i> | No | No | No | No | Yes |
| <i>Country effects</i> | No | No | Yes | Yes | Yes |
| <i>Time effects</i> | No | Yes | No | Yes | Yes |
| F-statistics and p-values testing exclusion of groups of variables | | | | | |
| <i>Time effects</i> | | 9.31 (0.000) | | 5.73 (0.00) | 4.61 (0.000) |
| <i>Age variables</i> | | | | | 2.12 (0.08) |
| <i>Age, educ, pop variables</i> | | | | | 1.44 (0.21) |

Notes: Clustered standard errors shown in parentheses and 95% confidence intervals are shown in brackets; *** $p < 0.01$, ** $p < 0.05$, * $p < 0.1$.

STATA CODES & RESULTS

Exercise 1(a-i) ([»back\(1a\)](#))

```
. regress lvio shall, vce(cluster stateid)
```

Linear regression

| | | |
|---------------|---|--------|
| Number of obs | = | 1,173 |
| F(1, 50) | = | 7.96 |
| Prob > F | = | 0.0068 |
| R-squared | = | 0.0866 |
| Root MSE | = | .61735 |

(Std. err. adjusted for 51 clusters in stateid)

| lvio | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-------|-------------|---------------------|-------|-------|----------------------|-----------|
| shall | -.4429646 | .1570184 | -2.82 | 0.007 | -.7583452 | -.1275839 |
| _cons | 6.134919 | .0790269 | 77.63 | 0.000 | 5.976189 | 6.293649 |

Exercise 1(a-ii) ([»back\(1a\)](#))

```
. global basevars "incarc_rate density avginc pop pb1064 pw1064 pm1029"  
  
. reg lvio shall $basevars, vce(cluster stateid)
```

| | | | |
|-------------------|---------------|---|--------|
| Linear regression | Number of obs | = | 1,173 |
| | F(8, 50) | = | 62.13 |
| | Prob > F | = | 0.0000 |
| | R-squared | = | 0.5643 |
| | Root MSE | = | .42769 |

(Std. err. adjusted for 51 clusters in stateid)

| lvio | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-------------|-------------|---------------------|-------|-------|----------------------|-----------|
| shall | -.3683869 | .113937 | -3.23 | 0.002 | -.5972361 | -.1395378 |
| incarc_rate | .0016126 | .0005999 | 2.69 | 0.010 | .0004076 | .0028177 |
| density | .0266885 | .0414909 | 0.64 | 0.523 | -.0566485 | .1100255 |
| avginc | .0012051 | .0240808 | 0.05 | 0.960 | -.0471626 | .0495728 |
| pop | .0427098 | .011729 | 3.64 | 0.001 | .0191515 | .0662681 |
| pb1064 | .0808526 | .0713875 | 1.13 | 0.263 | -.0625334 | .2242386 |
| pw1064 | .0312005 | .03409 | 0.92 | 0.364 | -.0372713 | .0996723 |
| pm1029 | .0088709 | .0340964 | 0.26 | 0.796 | -.0596137 | .0773554 |
| _cons | 2.981738 | 2.166513 | 1.38 | 0.175 | -1.369831 | 7.333307 |

Exercise 1(b-i)

```
*xtreg yvar xvar, fe vce(cluster entityid)
```

```
. xtreg lvio shall $basevars, fe vce(cluster stateid)
```

Fixed-effects (within) regression Number of obs = 1,173
Group variable: **stateid** Number of groups = 51

R-squared: Obs per group:

| | |
|--------------------|------------------------|
| Within = 0.2178 | min = 23 |
| Between = 0.0033 | avg = 23.0 |
| Overall = 0.0001 | max = 23 |

| | | | |
|-------------------------|----------|---|--------|
| | F(8, 50) | = | 34.10 |
| corr(u_i, Xb) = -0.3687 | Prob > F | = | 0.0000 |

(Std. err. adjusted for 51 clusters in **stateid**)

| lvio | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-------------|-------------|---------------------|-------|-------|----------------------|-----------|
| shall | -.0461415 | .0417616 | -1.10 | 0.275 | -.1300223 | .0377392 |
| incarc_rate | -.000071 | .0002504 | -0.28 | 0.778 | -.0005739 | .0004318 |
| density | -.1722901 | .1376129 | -1.25 | 0.216 | -.4486936 | .1041135 |
| avginc | -.0092037 | .0129649 | -0.71 | 0.481 | -.0352445 | .016837 |
| pop | .0115247 | .014224 | 0.81 | 0.422 | -.0170452 | .0400945 |
| pb1064 | .1042804 | .0326849 | 3.19 | 0.002 | .0386308 | .1699301 |
| pw1064 | .0408611 | .0134585 | 3.04 | 0.004 | .0138289 | .0678932 |
| pm1029 | -.0502725 | .0206949 | -2.43 | 0.019 | -.0918394 | -.0087057 |
| _cons | 3.866017 | .7701057 | 5.02 | 0.000 | 2.319214 | 5.412819 |

Exercise 1(b-ii-I)

```
*xtreg yvar xvar i.timeid, fe vce(cluster entityid)
```

```
. xtreg lvio shall $basevars i.year, fe vce(cluster stateid)
```

| | | | |
|-----------------------------------|------------------|---|-------|
| Fixed-effects (within) regression | Number of obs | = | 1,173 |
| Group variable: stateid | Number of groups | = | 51 |

Alternative,

```
. quiet tabulate year, generate(yr)
```

```
. xtreg lvio shall $basevars yr*, fe vce(cluster stateid)
```

note: yr23 omitted because of collinearity.

| | | | |
|-----------------------------------|------------------|---|-------|
| Fixed-effects (within) regression | Number of obs | = | 1,173 |
| Group variable: stateid | Number of groups | = | 51 |

Alternative,

```
. quiet tabulate year, generate(yr)
```

```
. global yr_vars "yr2 yr3 yr4 yr5 yr6 yr7 yr8 yr9 yr10 yr11 yr12 yr13 yr14 yr15 yr16 yr17 yr18 yr19 yr20 yr21 yr22 yr23"
```

```
. xtreg lvio shall $basevars $yr_vars, fe vce(cluster stateid)
```

| | | | |
|-----------------------------------|------------------|---|-------|
| Fixed-effects (within) regression | Number of obs | = | 1,173 |
| Group variable: stateid | Number of groups | = | 51 |

Exercise 1(b-ii-II)

```
. quiet tabulate year, generate(yr)

. global yr_vars "yr2 yr3 yr4 yr5 yr6 yr7 yr8 yr9 yr10 yr11 yr12 yr13 yr14 yr15 yr16 yr17 yr18 yr19 yr20 yr21 yr22 yr23"

. xtreg lvio shall $basevars $yr_vars, fe vce(cluster stateid)
```

```
Fixed-effects (within) regression      Number of obs   =      1,173
Group variable: stateid                Number of groups =       51
```

```
R-squared:                               Obs per group:
    Within = 0.4180                        min =          23
    Between = 0.0419                       avg =         23.0
    Overall = 0.0009                       max =          23
```

```
                                F(30, 50)      =      56.86
corr(u_i, Xb) = -0.2929                Prob > F       =      0.0000
```

(Std. err. adjusted for 51 clusters in **stateid**)

| lvio | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-------------|-------------|---------------------|-------|-------|----------------------|----------|
| shall | -.0279935 | .0407168 | -0.69 | 0.495 | -.1097757 | .0537886 |
| incarc_rate | .000076 | .0002079 | 0.37 | 0.716 | -.0003416 | .0004935 |
| density | -.091555 | .1238622 | -0.74 | 0.463 | -.3403396 | .1572296 |
| avginc | .0009587 | .0164931 | 0.06 | 0.954 | -.0321688 | .0340861 |
| pop | -.0047544 | .0152294 | -0.31 | 0.756 | -.0353436 | .0258347 |
| pbl064 | .0291862 | .0495407 | 0.59 | 0.558 | -.0703192 | .1286916 |
| pw1064 | .0092501 | .0237564 | 0.39 | 0.699 | -.0384659 | .0569662 |
| pm1029 | .0733254 | .0524733 | 1.40 | 0.168 | -.0320704 | .1787211 |

Exercise 1(b-iii)

```
. quiet xtreg lvio shall $basevars $yr_vars, fe vce(cluster stateid)
```

```
. test $yr_vars
```

```
( 1)  yr2 = 0
( 2)  yr3 = 0
( 3)  yr4 = 0
( 4)  yr5 = 0
( 5)  yr6 = 0
( 6)  yr7 = 0
( 7)  yr8 = 0
( 8)  yr9 = 0
( 9)  yr10 = 0
(10)  yr11 = 0
(11)  yr12 = 0
(12)  yr13 = 0
(13)  yr14 = 0
(14)  yr15 = 0
(15)  yr16 = 0
(16)  yr17 = 0
(17)  yr18 = 0
(18)  yr19 = 0
(19)  yr20 = 0
(20)  yr21 = 0
(21)  yr22 = 0
(22)  yr23 = 0
```

| |
|--------------------|
| F(22, 50) = 21.62 |
| Prob > F = 0.0000 |

Exercise 2(b-I)

```
. reg dem_ind log_gdppc, vce(cluster code) //using clustered standard errors
```

Linear regression

| | | |
|---------------|---|--------|
| Number of obs | = | 958 |
| F(1, 149) | = | 396.40 |
| Prob > F | = | 0.0000 |
| R-squared | = | 0.4385 |
| Root MSE | = | .2719 |

(Std. err. adjusted for 150 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|--------|-------|----------------------|-----------|
| log_gdppc | .2356731 | .011837 | 19.91 | 0.000 | .212283 | .2590632 |
| _cons | -1.354828 | .1004215 | -13.49 | 0.000 | -1.553262 | -1.156394 |

Exercise 2(b-II)

```
. reg dem_ind log_gdppc, robust //without clustering
```

| | | | |
|-------------------|---------------|---|---------|
| Linear regression | Number of obs | = | 958 |
| | F(1, 956) | = | 1100.57 |
| | Prob > F | = | 0.0000 |
| | R-squared | = | 0.4385 |
| | Root MSE | = | .2719 |

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|--------|-------|----------------------|----------|
| log_gdppc | .2356731 | .007104 | 33.17 | 0.000 | .221732 | .2496143 |
| _cons | -1.354828 | .0607549 | -22.30 | 0.000 | -1.474056 | -1.2356 |

```
. //time fixed effects only
. reg dem_ind log_gdppc y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000, vce(cluster code)
```

| | | | |
|-------------------|---------------|---|--------|
| Linear regression | Number of obs | = | 958 |
| | F(9, 149) | = | 56.80 |
| | Prob > F | = | 0.0000 |
| | R-squared | = | 0.4767 |
| | Root MSE | = | .26357 |

(Std. err. adjusted for 150 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|--------|-------|----------------------|-----------|
| log_gdppc | .2351066 | .0122358 | 19.21 | 0.000 | .2109285 | .2592847 |
| y1965 | -.0756674 | .0205124 | -3.69 | 0.000 | -.1162003 | -.0351346 |
| y1970 | -.2387064 | .0343905 | -6.94 | 0.000 | -.3066625 | -.1707504 |
| y1975 | -.2801793 | .0341178 | -8.21 | 0.000 | -.3475965 | -.2127621 |
| y1980 | -.2445214 | .0331917 | -7.37 | 0.000 | -.3101086 | -.1789341 |
| y1985 | -.2415764 | .035402 | -6.82 | 0.000 | -.3115311 | -.1716216 |
| y1990 | -.2064564 | .0328683 | -6.28 | 0.000 | -.2714045 | -.1415083 |
| y1995 | -.1720611 | .0351667 | -4.89 | 0.000 | -.2415511 | -.1025712 |
| y2000 | -.1687362 | .0344812 | -4.89 | 0.000 | -.2368716 | -.1006009 |
| _cons | -1.156693 | .1062762 | -10.88 | 0.000 | -1.366696 | -.9466898 |

```
. quiet reg dem_ind log_gdppc y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000, vce(cluster code)
```

```
. test y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000
```

```
( 1) y1965 = 0
```

```
( 2) y1970 = 0
```

```
( 3) y1975 = 0
```

```
( 4) y1980 = 0
```

```
( 5) y1985 = 0
```

```
( 6) y1990 = 0
```

```
( 7) y1995 = 0
```

```
( 8) y2000 = 0
```

| | |
|--------------|--------|
| F(8, 149) = | 9.31 |
| Prob > F = | 0.0000 |

Exercise 2(c) (»back(2c))

```
. //country fixed effects only
. xtreg dem_ind log_gdppc, fe vce(cluster code)
```

```
Fixed-effects (within) regression      Number of obs   =      958
Group variable: code                  Number of groups =      150
```

```
R-squared:                               Obs per group:
    Within = 0.0197                      min =      1
    Between = 0.5365                     avg =     6.4
    Overall = 0.4385                     max =      9
```

```
corr(u_i, Xb) = 0.6173                  F(1, 149)       =      7.06
                                          Prob > F        =     0.0088
```

(Std. err. adjusted for 150 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|-----------------------------------|-------|-------|----------------------|----------|
| log_gdppc | .083741 | .0315258 | 2.66 | 0.009 | .0214456 | .1460364 |
| _cons | -.115316 | .257198 | -0.45 | 0.655 | -.6235425 | .3929106 |
| sigma_u | .26651952 | | | | | |
| sigma_e | .20351058 | | | | | |
| rho | .63168655 | (fraction of variance due to u_i) | | | | |

Exercise 2(d-I) (»back(2d))

```
. // both country and time fixed effects
. xtreg dem_ind log_gdppc y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000, fe vce(cluster code)
```

```
Fixed-effects (within) regression               Number of obs   =       958
Group variable: code                          Number of groups =       150
```

```
R-squared:                                     Obs per group:
    Within = 0.1182                             min =          1
    Between = 0.3832                             avg =          6.4
    Overall = 0.3190                             max =          9
```

```
corr(u_i, Xb) = 0.4393                        F(9, 149)        =       5.65
                                                Prob > F         =     0.0000
```

(Std. err. adjusted for 150 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|---------------------|-------|-------|----------------------|-----------|
| log_gdppc | .0535878 | .042432 | 1.26 | 0.209 | -.0302585 | .137434 |
| y1965 | .0002347 | .0209199 | 0.01 | 0.991 | -.0411033 | .0415727 |
| y1970 | -.1268076 | .0340453 | -3.72 | 0.000 | -.1940816 | -.0595337 |
| y1975 | -.1477264 | .0370153 | -3.99 | 0.000 | -.2208692 | -.0745836 |
| y1980 | -.097822 | .0355399 | -2.75 | 0.007 | -.1680494 | -.0275947 |
| y1985 | -.0871025 | .0391062 | -2.23 | 0.027 | -.1643769 | -.009828 |
| y1990 | -.0421216 | .0353035 | -1.19 | 0.235 | -.1118818 | .0276385 |
| y1995 | .0095646 | .0426094 | 0.22 | 0.823 | -.0746322 | .0937613 |
| y2000 | .0323636 | .0432037 | 0.75 | 0.455 | -.0530075 | .1177348 |
| _cons | .1802954 | .327202 | 0.55 | 0.582 | -.4662601 | .8268508 |

sigma_u .2835593

sigma_e .19397224

rho .68122712 (fraction of variance due to u_i)

Exercise 2(d-II) (»back(2d))

```
. quiet xtreg dem_ind log_gdppc y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000, fe vce(cluster code)

. test y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000

( 1)  y1965 = 0
( 2)  y1970 = 0
( 3)  y1975 = 0
( 4)  y1980 = 0
( 5)  y1985 = 0
( 6)  y1990 = 0
( 7)  y1995 = 0
( 8)  y2000 = 0
```

| | |
|--------------|--------|
| F(8, 149) = | 5.73 |
| Prob > F = | 0.0000 |

Exercise 2(e-I) (»back(2e))

```
. xtreg dem_ind log_gdppc log_pop educ age_2 age_3 age_4 age_5 y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000, fe vce(cluster code)
note: y2000 omitted because of collinearity.
```

```
Fixed-effects (within) regression      Number of obs   =    680
Group variable: code                  Number of groups =    96
```

```
R-squared:                            Obs per group:
    Within = 0.1327                      min =      1
    Between = 0.0165                     avg  =    7.1
    Overall = 0.0359                     max  =     8
```

```
F(14, 95) = 3.22
corr(u_i, Xb) = -0.2385                 Prob > F         = 0.0004
```

(Std. err. adjusted for 96 clusters in code)

| dem_ind | Coefficient | Robust std. err. | t | P> t | [95% conf. interval] | |
|-----------|-------------|-----------------------------------|-------|-------|----------------------|-----------|
| log_gdppc | .0252013 | .0539626 | 0.47 | 0.642 | -.0819281 | .1323306 |
| log_pop | -.0692295 | .1245471 | -0.56 | 0.580 | -.3164868 | .1780278 |
| educ | -.0004013 | .0232475 | -0.02 | 0.986 | -.0465534 | .0457509 |
| age_2 | -.5255157 | .609811 | -0.86 | 0.391 | -1.736144 | .6851122 |
| age_3 | -2.481235 | .8941413 | -2.77 | 0.007 | -4.25633 | -.7061401 |
| age_4 | .2978116 | 1.297809 | 0.23 | 0.819 | -2.278666 | 2.874289 |
| age_5 | .6054059 | 1.296464 | 0.47 | 0.642 | -1.9684 | 3.179212 |
| y1965 | -.1582882 | .119523 | -1.32 | 0.189 | -.3955713 | .0789949 |
| y1970 | -.2599235 | .1858436 | -2.46 | 0.016 | -.4700497 | -.0497974 |
| y1975 | -.2815462 | .0923211 | -3.05 | 0.003 | -.4648267 | -.0982656 |
| y1980 | -.2052376 | .0784034 | -2.62 | 0.010 | -.360888 | -.0495871 |
| y1985 | -.1537799 | .0596167 | -2.58 | 0.011 | -.272134 | -.0354257 |
| y1990 | -.1008704 | .0443634 | -2.27 | 0.025 | -.1889428 | -.0127979 |
| y1995 | -.0397771 | .0220018 | -1.81 | 0.074 | -.0834561 | .0039019 |
| y2000 | 0 | (omitted) | | | | |
| _cons | 1.649546 | 1.406281 | 1.17 | 0.244 | -1.142275 | 4.441368 |
| sigma_u | .30898253 | | | | | |
| sigma_e | .19831249 | | | | | |
| rho | .70824619 | (fraction of variance due to u_i) | | | | |

Exercise 2(e-II) (»back(2e))

```
. quiet xtreg dem_ind log_gdppc log_pop educ age_2 age_3 age_4 age_5 y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000, fe vce(cluster code)
```

```
. test y1965 y1970 y1975 y1980 y1985 y1990 y1995 y2000
```

```
( 1) y1965 = 0
( 2) y1970 = 0
( 3) y1975 = 0
( 4) y1980 = 0
( 5) y1985 = 0
( 6) y1990 = 0
( 7) y1995 = 0
( 8) o.y2000 = 0
      Constraint 8 dropped
```

| |
|-------------------|
| F(7, 95) = 4.61 |
| Prob > F = 0.0002 |

```
. test age_2 age_3 age_4 age_5
```

```
( 1) age_2 = 0
( 2) age_3 = 0
( 3) age_4 = 0
( 4) age_5 = 0
```

| |
|-------------------|
| F(4, 95) = 2.12 |
| Prob > F = 0.0837 |

```
. test age_2 age_3 age_4 age_5 educ log_pop
```

```
( 1) age_2 = 0
( 2) age_3 = 0
( 3) age_4 = 0
( 4) age_5 = 0
( 5) educ = 0
( 6) log_pop = 0
```

| |
|-------------------|
| F(6, 95) = 1.44 |
| Prob > F = 0.2070 |