

# The Bayesian Lasso

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## MODIFIED BAYES' THEOREM:

$$P(H|X) = P(H) * \left( 1 + P(C) * \left( \frac{P(x|H)}{P(x)} - 1 \right) \right)$$

H: HYPOTHESIS

X: OBSERVATION

P(H): PRIOR PROBABILITY THAT H IS TRUE

P(x): PRIOR PROBABILITY OF OBSERVING X

P(C): PROBABILITY THAT YOU'RE USING  
BAYESIAN STATISTICS CORRECTLY

# The Bayesian Lasso<sup>1</sup>

## 1. Formulation

- ▶ Classical Regression
- ▶ Classical Lasso
- ▶ Bayesian Lasso

## 2. Selecting $\lambda$

- ▶ Classical Regression
- ▶ Classical Lasso

## 3. Comparison

## 4. Extensions

## 5. Synthetic Example

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<sup>1</sup>Park and Casella (2008)

# Classical Regression

$$y = \mu 1_n + X\beta + \epsilon$$

- ▶  $y$  is an  $n \times 1$  vector of responses
- ▶  $\mu$  is the overall mean
- ▶  $X$  is the  $n \times p$  matrix of **standardized** regressors
- ▶  $\beta = (\beta_1, \dots, \beta_p)^T$
- ▶  $\epsilon$  is an  $n \times 1$  vector of  $\stackrel{iid}{\sim} N(0, \sigma^2)$

Satisfies  $\min_{\beta} (\tilde{y} - X\beta)^T (\tilde{y} - X\beta)$

# Classical Lasso

## Formulation

$$\min_{\beta} (\tilde{y} - X\beta)^T (\tilde{y} - X\beta) + \lambda \sum_{j=1}^p |\beta_j|$$

$$\lambda \geq 0$$

# Classical Lasso

$$\min_{\beta} (\tilde{y} - X\beta)^T (\tilde{y} - X\beta) + \lambda \sum_{j=1}^p |\beta_j|$$

## Notes

1. Often called “penalized regression”
2. L1 penalty
3. “Shrinkage”  $\beta$  values are shrunk towards 0
4. Tune  $\lambda$  through cross validation

## Motivation

1. Model selection - often as a precursor to other models
2. Reduce overfitting
3. Easily extendable to generalized linear models

# Classical Lasso

$$\min_{\beta} (\tilde{y} - X\beta)^T (\tilde{y} - X\beta) + \lambda \sum_{j=1}^p |\beta_j|$$

## Drawbacks

1. Biases  $\beta$
2. Unreliable standard errors (issues with statistical tests)
3. Correlated features
4. Tuning issues / time

# Bayesian Lasso<sup>2</sup>

## Hierarchical Specification 1 (1 of 2)

$$y|\mu, X, \beta, \sigma^2 \sim N_n(\mu 1_n + X\beta, \sigma^2 I_n)$$

$$\beta|\sigma^2, \tau_1^2, \dots, \tau_p^2 \sim N_p(0_p, \sigma^2 D_t)$$

$$D_t = \text{diag}(\tau_1^2, \dots, \tau_p^2)$$

$$\sigma^2, \tau_1^2, \dots, \tau_p^2 \sim \pi(\sigma^2) d\sigma^2 \prod_{j=1}^P \frac{\lambda^2}{2} e^{-\lambda^2 \frac{\tau_j^2}{2}} d\tau_j^2$$

$$\sigma^2, \tau_1^2, \dots, \tau_p^2 > 0$$

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<sup>2</sup>Andrews and Mallows (1974)



# Bayesian Lasso

## Hierarchical Specification 1 (2 of 2)

The authors integrate out  $\tau_1^2, \dots, \tau_p^2$  which yields the conditional prior for  $\beta$  as a Laplace (double-exponential) distribution:

$$\pi(\beta|\sigma^2) = \prod_{j=1}^P \frac{\lambda}{2\sqrt{\sigma^2}} e^{\frac{-\lambda|\beta_j|}{\sqrt{\sigma^2}}}$$

$$\pi(\sigma^2) = IG(\alpha, \beta)$$

$$p(\mu) = U(a, b)$$

# Bayesian Lasso<sup>3</sup>

## Hierarchical Specification 2 (1 of 2)

$$y|\mu, X, \beta, \sigma^2 \sim N_n(\mu 1_n + X\beta, \sigma^2 I_n)$$

Authors integrate out  $\mu$

$$p(\beta) = N(A^{-1}X^T\tilde{y}, \sigma^2 A^{-1})$$

where

$$A = X^T X + D_\tau^{-1}$$

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<sup>3</sup>Bae and Mallick (2004)

# Bayesian Lasso

## Hierarchical Specification 2 (2 of 2)

$$p(\sigma^2) = IG(\frac{n-1}{2} + \frac{p}{2}, (\tilde{y} - X\beta)^T \frac{(\tilde{y} - X\beta)}{2} + \beta^T D_\tau^{-1} \frac{\beta}{2})$$

$$p(\tau_1^2, \dots, \tau_p^2) = \sqrt{\frac{\lambda'}{2\pi}} x^{-\frac{3}{2}} \exp\left\{-\frac{\lambda'(x - \mu')^2}{2(\mu')^2 x}\right\}$$

where

$$\mu' = \sqrt{\frac{\lambda^2 \sigma^2}{\beta_j^2}}$$

$$\lambda' = \lambda^2$$

# Choosing the Lasso Parameter: Classical Lasso

## Cross Validation

1. Cross validate over a grid of  $\lambda$  where  $\lambda \geq 0$ .
2. For each  $\lambda$  value find the error metric of interest.
3. Select the  $\lambda$  value that minimizes the metric of interest.

# Choosing the Lasso Parameter: Bayesian Lasso

Technique 1: Empirical Bayes

# Choosing the Lasso Parameter: Bayesian Lasso

## Technique 1: Empirical Bayes



# Choosing the Lasso Parameter: Bayesian Lasso

## Technique 1: Empirical Bayes<sup>4</sup>

1. Solve for a marginal maximum likelihood for  $\lambda$  using estimates of the hyperparameters.
2.  $\lambda$  is updated for each iteration using the estimates from the sample of the previous iteration

$$\lambda^{(k)} = \sqrt{\frac{2p}{\sum_{j=1}^p E_{\lambda^{(k-1)}}[\tau_j^2 | \tilde{y}]}}$$

3. Recommended initial value of:

$$\lambda^{(0)} = \frac{p \sqrt{\sigma_{LS}^2}}{\sum_{j=1}^p |\beta_j^{LS}|}$$

4.  $\beta_j^{LS}$  and  $\sigma_{LS}^2$  are estimated from least squares.

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<sup>4</sup>Casella (2001)

# Choosing the Lasso Parameter: Bayesian Lasso

## Technique 2: Hyperpriors

Authors recommend the diffuse Gamma hyperprior of  $\lambda^2$  in the following form

$$\pi(\lambda^2) = \frac{\delta^r}{\Gamma(r)} (\lambda^2)^{r-1} e^{-\delta \lambda^2}$$

$$\lambda^2 > 0, r > 0, \delta > 0$$

- ▶ Select  $r$  and  $\delta$  such that there is high probability near the maximum likelihood estimate to avoid mixing problems.
- ▶  $r = 0$  and  $\delta = 0$  are tempting but lead to an improper posterior.
- ▶ This formulation allows easy integration into a Gibbs sampler.



## Comparison

Consider the following data:<sup>5</sup> <sup>6</sup>

obs	age	sex	bmi	bp	s1	s2	s3	s4	s5
-0.01	0.80	1.06	1.30	0.46	-0.93	-0.73	-0.91	-0.05	0.42
-1.00	-0.04	-0.94	-1.08	-0.55	-0.18	-0.40	1.56	-0.83	-1.43
-0.14	1.79	1.06	0.93	-0.12	-0.96	-0.72	-0.68	-0.05	0.06
0.70	-1.87	-0.94	-0.24	-0.77	0.26	0.52	-0.76	0.72	0.48
-0.22	0.11	-0.94	-0.76	0.46	0.08	0.33	0.17	-0.05	-0.67
-0.72	-1.95	-0.94	-0.85	-0.41	-1.45	-1.67	0.87	-1.60	-0.86

- ▶ Measurements of 440 diabetic patients
- ▶ 10 baseline variables (centered and scaled)
- ▶ Response variable is a measure of disease progression one year after baseline

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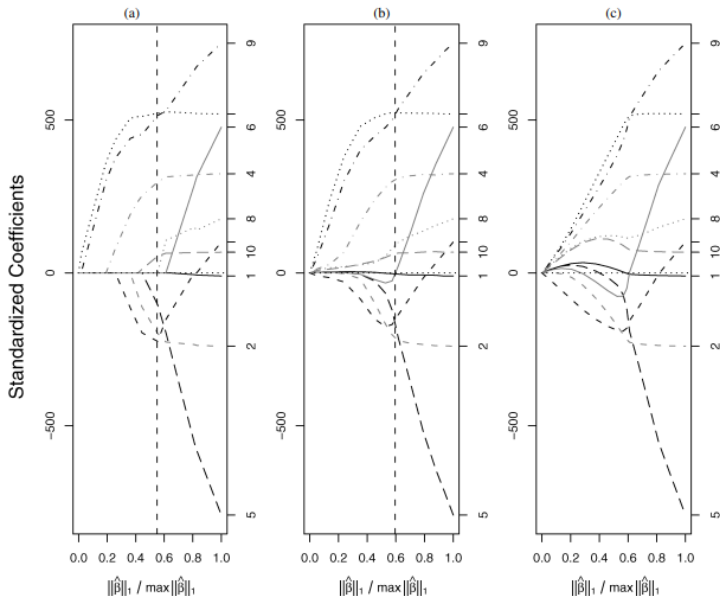
<sup>5</sup>{care} Efron et al. (2004)

<sup>6</sup>Zuber and Strimmer. (2021)

# Trace Plot of Coefficients by Lasso

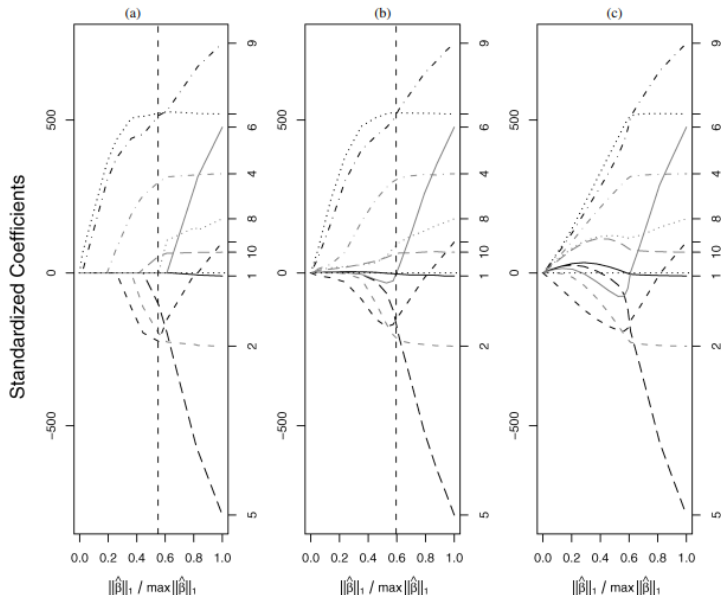
- a) Lasso
- b) Bayesian Lasso
- c) Ridge Regression

Vertical lines for the Lasso and Bayesian Lasso indicating the estimates chosen by n-fold cross-validation and marginal maximum likelihood.



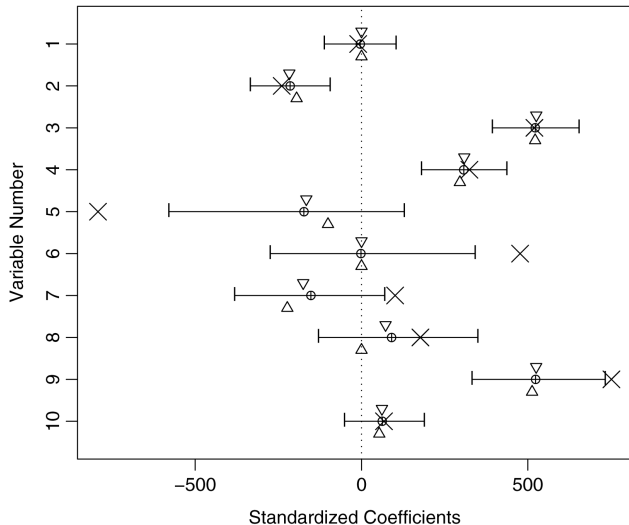
## Comparison

- ▶ The Bayesian Lasso estimates appear to be a compromise between the Lasso and ridge regression estimates
- ▶ The Bayesian Lasso appears to pull the more weakly related parameters to 0 faster than ridge regression



## Comparison

- ▶ Least squares estimates ( $\times$ ),
- ▶ Lasso estimates based on n-fold cross-validation ( $\triangle$ ),
- ▶ Lasso estimates chosen to match the L1 norm of the Bayes estimates ( )
- ▶ Posterior median Bayesian Lasso estimates ( $\otimes$ ) and corresponding 95% credible intervals (equal-tailed) with  $\lambda$  selected according to marginal maximum likelihood



# Extensions

## “Bridge” Regression<sup>7</sup>

$$\min_{\beta} (\tilde{y} - X\beta)^T (\tilde{y} - X\beta) + \lambda \sum_{j=1}^p |\beta_j|^2$$

$$\pi(\beta|\sigma^2) \propto \prod_{j=1}^P e^{-\lambda(\frac{|\beta_j|}{\sqrt{\sigma^2}})^2}$$

## Huberized Lasso<sup>8</sup>

$$\min_{\beta} \sum_{i=1}^n L(\tilde{y}_i - x_i^T \beta) + \lambda \sum_{j=1}^p |\beta_j|$$

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<sup>7</sup>Knight and Fu (2000)

<sup>8</sup>Rosset and Zhu (2007)

## Example

To compare the results, we will generate synthetic data in the form:

$$\text{logit}(p) = \log \frac{p}{1-p} = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$

```
num <- 100
```

```
x1 <- rnorm(num)
```

```
x2 <- rnorm(num)
```

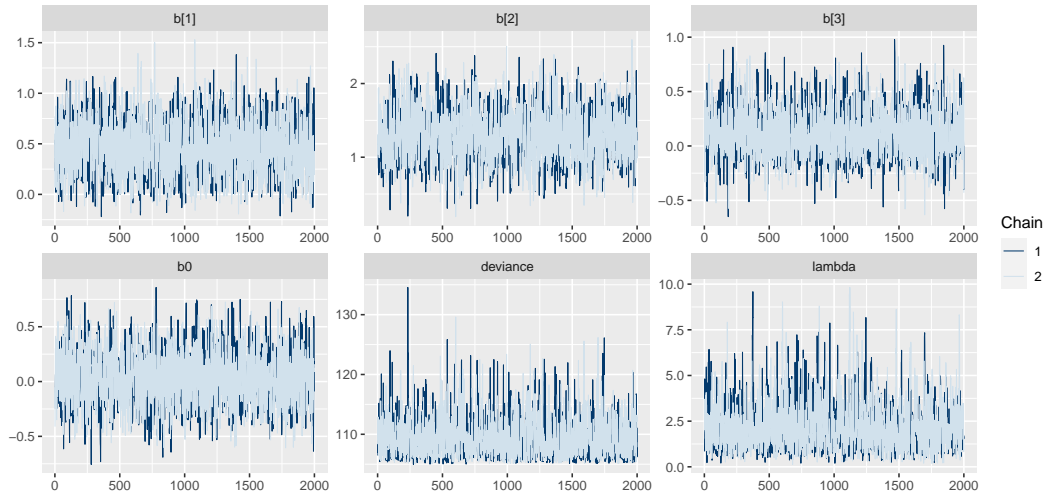
```
x3 <- rnorm(num)
```

```
prob <- exp(x1+2*x2) / (1+exp(x1+2*x2))
```

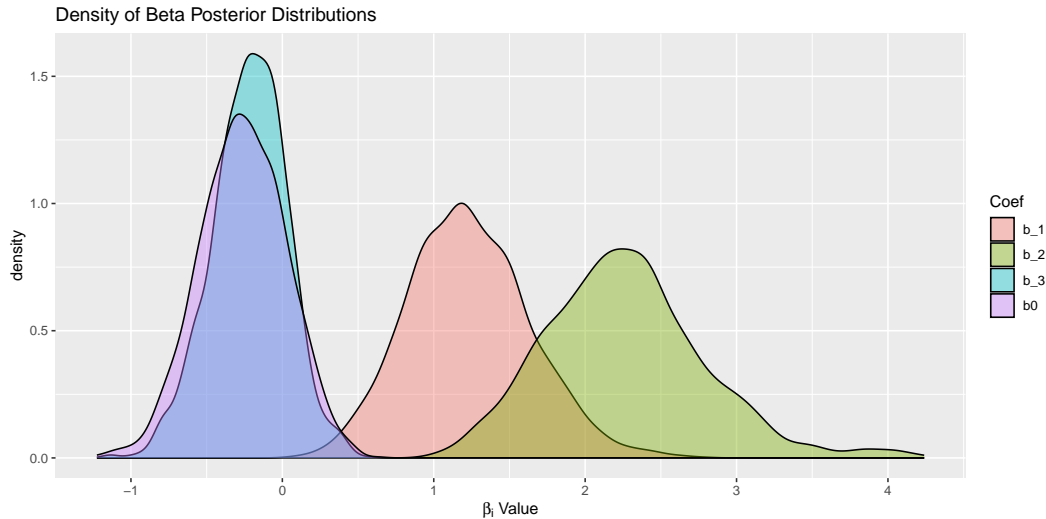
```
y <- rbinom(num, 1, prob)
```

# Example

Trace Plots



# Example





## Example

model	b0	b_1	b_2	b_3
Logistic Regression	-0.29	1.41	2.45	-0.21
Lasso	-0.21	0.87	1.73	-0.11
Bayes Lasso	-0.26	1.23	2.27	-0.22
Truth	0.00	1.00	1.00	0.00

## Bibliography

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