A New Statistical Test for Tobit Model and ZMP Model

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目录

L score test推导

模型设定

• ZIP统计模型:

$$\begin{cases}
P(Y_i = 0) = \omega + (1 - \omega)e^{-\lambda_i} \\
P(Y_i = y_i) = (1 - \omega)\frac{e^{-\lambda_i}\lambda_i^{y_i}}{y_i!} & (y_i > 0)!
\end{cases}$$
(1.1)

- 令:
 - $\theta = \frac{\omega}{1-\omega} =$ 常数 $(-f(0) \le \theta < \inf)$
 - $ln(\lambda_i) = X_i\beta$

• 示性函数:

$$I_{(condition)} = \begin{cases} 1 & \text{if condition is True} \\ 0 & \text{if condition is False} \end{cases}$$
 (1.2)

• 对数似然:

$$l(\lambda,\theta;y) = \sum_{i} \left\{ -log(1+\theta) + I_{y_i=0}log(\theta+e^{-\lambda_i}) + +I_{y_i>0} \left[-\lambda_i + y_ilog(\lambda_i) - log(y_i!) \right] \right\} \tag{1.3}$$

● 偏导方程:

$$\begin{split} \frac{dl}{d\beta_r} &= \sum_i \left\{ I_{(y_i=0)} \left(\frac{-e^{-\lambda_i}}{\theta + e^{-\lambda_i}} \right) \lambda_i x_{ir} + I_{y_i>0} (y_i - \lambda_i) x_{ir} \right\} \ r = 1, 2, ..., p \\ \frac{dl}{d\theta} &= \sum_i \left\{ \frac{-1}{(1+\theta)} + I_{y_i=0} \left(\frac{1}{\theta + e^{-\lambda_i}} \right) \right\} \end{split}$$

• ZIP的对数似然函数基础上构建得分向量:

$$U(\beta, \theta) = \begin{bmatrix} \frac{\partial l}{\partial \beta} \\ \frac{\partial l}{\partial \theta} \end{bmatrix}$$

Fisher信息矩阵:

$$J(\beta,\theta) = \begin{bmatrix} \frac{\partial^2 l}{\partial \beta \partial \beta'} & \frac{\partial^2 l}{\partial \beta \partial \theta} \\ \frac{\partial^2 l}{\partial \beta' \partial \theta} & \frac{\partial^2 l}{\partial \theta \partial \theta} \end{bmatrix}$$

在原假设成立下,根据Fisher信息矩阵构建Score检验统计量:

$$S(\hat{\beta}) = S(\hat{\beta}, 0) = U^{T}(\hat{\beta}, 0)[J(\hat{\beta}, 0)]^{-1}U(\hat{\beta}, 0) \sim \chi^{2}_{(1)}$$

Lscore test推导

• 所以在原假设条件下, $\theta = 0$:

$$\frac{dl}{d\theta}|_{\theta=0} = \sum_{i} \left\{ \frac{I_{(y_i=0)}}{e^{-\hat{\lambda}_i}} - 1 \right\}$$

• 计算得分向量得:

$$U^{T}(eta,0) = \left(0,...,0,\sum_{i}\left\{rac{I_{(y_{i}=0)}}{e^{-\hat{\lambda}_{i}}}-1
ight\}
ight)$$

• 计算二阶偏导:

$$\begin{split} \frac{d^2 I(.)}{d\beta r d\beta s} &= \sum_i \left\{ I_{y_i=0} \left(\frac{-e^{-\lambda_i} [(1-\lambda_i)\theta + e^{-\lambda_i}]}{[\theta + e^{-\lambda_i}]^2} \right) \lambda_i x_{ir} x_{is} - I_{(y_i>0)} \lambda_i x_i r x_i s \right\} \ r = 1...p, s = 1...p \\ \frac{d^2 I(.)}{d\beta r d\theta} &= \sum_i \left\{ I_{y_i=0} \left(\frac{-e^{-\lambda_i}}{[\theta + e^{-\lambda_i}]^2} \right) \lambda_i x_{ir} \right\} \ r = 1...p \\ \frac{d^2 I(.)}{d\theta^2} &= \sum_i \left\{ \frac{1}{(1+\theta)^2} - I_{y_i=0} \left(\frac{1}{[\theta + e^{-\lambda_i}]^2} \right) \right\} \end{split} \tag{1.4}$$

• 可以得到 $J(\hat{\beta},0)$:

•
$$\hat{J}_{r,s} = \sum_{i} \hat{\lambda}_{i} x_{ir} x_{is} \ r = 1...p, s = 1...p$$

•
$$\hat{J}_{r,p+1} = -\sum_{i} \hat{\lambda}_{i} x_{i} r \ r = 1...p$$

•
$$\hat{J}_{p+1,p+1} = \sum_{i} \frac{1 - e^{-\hat{\lambda}_i}}{e^{-\hat{\lambda}_i}}$$

• $J(\hat{\beta},0)$ 矩阵分块得:

$$J(\hat{\beta},0) = \left[\begin{array}{cc} \hat{J_{11}} & \hat{J_{12}} \\ \hat{J_{21}} & \hat{J_{22}} \end{array} \right]$$

- $\bullet \ \hat{J_{11}} = X^T diag(\hat{\lambda}X)$
- $\bullet \hat{J}_{12} = -X^T \hat{\lambda}$
- $\hat{J}_{21} = -\hat{\lambda^T} X^T$
- $\bullet \ \hat{J}_{22} = \sum_{i} \left(\frac{1}{e^{-\lambda_i}} 1 \right)$

• $J(\hat{\beta},0)^{-1}$ 定义为C:

$$J(\hat{eta},0)^{-1} = C(\hat{eta},0) = \left[egin{array}{cc} C_{11} & C_{12} \ C_{21} & C_{22} \end{array}
ight]$$

• 由于得分向量 $U(\hat{\beta},0)$ 的结构只需要 C_{22} 即可:

$$C_{22}^{-1} = \hat{J}_{22} - \hat{J}_{21}\hat{J}_{11}^{-1}\hat{J}_{12}$$

$$C_{22}^{-1} = \sum_{i} \left(\frac{1}{e^{-\hat{\lambda}_{i}}} - 1\right) - \hat{\lambda}^{T}X[X^{T}diag(\hat{\lambda})X]^{-1}X^{T}\hat{\lambda}$$

• Score统计量最终形式为:

$$S(\hat{\beta}) = S(\hat{\beta}, 0) = U^{T}(\hat{\beta}, 0)[J(\hat{\beta}, 0)]^{-1}U(\hat{\beta}, 0) \sim \chi^{2}_{(1)}$$

$$S(\hat{\beta}, 0) = \frac{\left\{\sum_{i=1}^{n} \left(\frac{I_{y_i=0}}{e^{-\hat{\lambda}_i}} - 1\right)\right\}^2}{\sum_{i} \left(\frac{1}{e^{-\hat{\lambda}_i}} - 1\right) - \hat{\lambda}^T X [X^T diag(\hat{\lambda})X]^{-1} X^T \hat{\lambda}}$$

● 可得:

$$S_{score} = \frac{\left\{\sum_{i=1}^{n} \left(\frac{I_{y_i=0}}{e^{-\hat{\lambda_i}}} - 1\right)\right\}}{\left\{\sum_{i} \left(\frac{1}{e^{-\hat{\lambda_i}}} - 1\right) - \hat{\lambda}^T X [X^T diag(\hat{\lambda})X]^{-1} X^T \hat{\lambda}\right\}^{1/2}}$$
$$S_{score} \sim N(0, 1)$$