

Prediction and inference for epidemic models using likelihood-based methods

2nd New Zealand Workshop on Uncertainty Quantification and Inverse Problems

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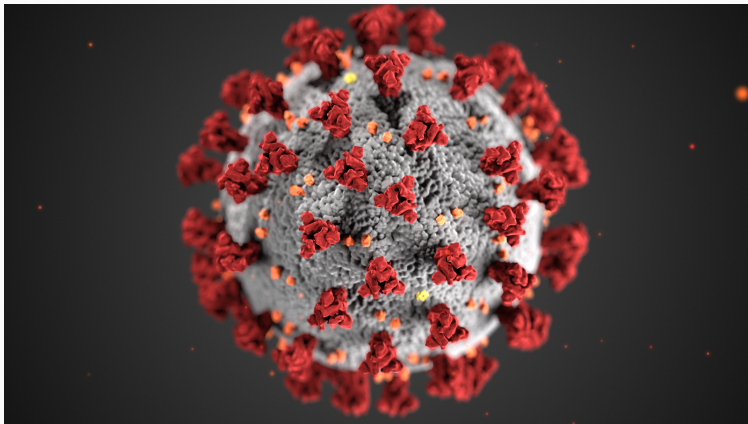
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4. Case Study

Epidemic modelling is a *severely neglected* field of applied mathematics, due to its *low economic and societal impacts*.

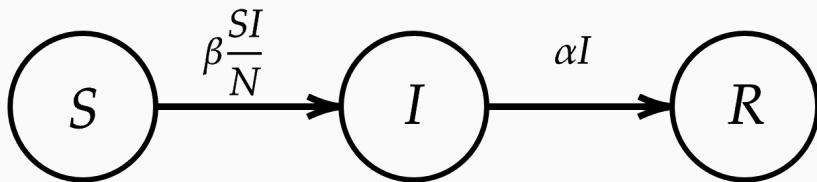


Various ways to model epidemic outbreaks:

- Phenomenological Models
- Deterministic ODEs
- Stochastic SDEs
- Branching Processes
- Agent-based Simulations

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$$\frac{dS}{dt} = -\beta \frac{SI}{N}$$

$$\frac{dI}{dt} = \beta \frac{SI}{N} - \alpha I$$

$$\frac{dR}{dt} = \alpha I$$

In general:

$$y = g(x) + \epsilon$$

$$\frac{dx}{dt} = f(x; \theta)$$

Two sources of information:

- the data, y
- the differential equation model, f

How do we use both?

Integrate the differential equation for proposal θ

$$\theta \mapsto \frac{dx}{dt} = f(x; \theta) \rightarrow x$$

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Form a likelihood function based on the observation function

$$x(\theta) \xrightarrow{y=g(x)+\epsilon} p(y|\theta)$$

Integrate the differential equation for proposal θ

$$\theta \xrightarrow[\frac{dx}{dt}=f(x;\theta)]{} x$$

Form a likelihood function based on the observation function

$$x(\theta) \xrightarrow{y=g(x)+\epsilon} p(y|\theta)$$

Then apply your chosen statistical philosophy to this likelihood

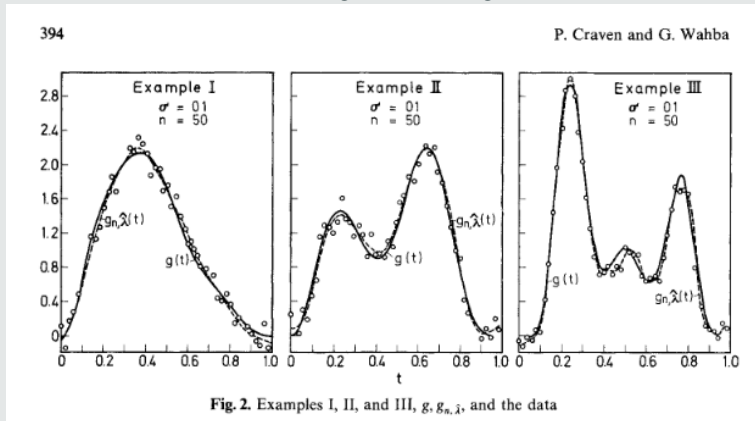
$$p(y|\theta) \xrightarrow{\text{magic}} \text{results}$$

If we had some estimate of x , we could find an estimate of θ pretty easily, without integration:

$$\frac{d}{dt}x_{\text{est}} = f(x_{\text{est}}; \theta)$$

If **An example of G**

We can do this, for example, using smoothing splines.



from Craven and Wahba (1979), Smoothing Noisy Data with Spline Functions, Numer. Math, 31,377-403

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If we can't find an appropriate G (e.g. when the x is only partially observed), then we cannot use this method.

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We can get around that restriction, if we estimate both x and θ simultaneously.

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Can project x onto some smooth basis ($x = \Phi c$), to reduce the dimension of the problem.

Two-level process:

Inner: for a given θ , compute a optimal smooth \hat{c}

$$\hat{c}(\theta) = \min_{c|\theta} \left\{ \|y - g(\Phi_c)\|^2 + \lambda \left\| \frac{d}{dt}(\Phi_c) - f(\Phi_c, \theta) \right\|^2 \right\}$$

Outer: Optimise over θ , using the inner to eliminate nuisance parameters (c)

$$\min_{\theta} \left\{ \|y - g(\Phi_{\hat{c}(\theta)})\|^2 \right\}$$

Given a likelihood:

$$\mathcal{L}(x_1, x_2)$$

Marginal likelihood:

$$\mathcal{L}_{X_1}(x_1) = \int \mathcal{L}(x_1, x_2) dx_2$$

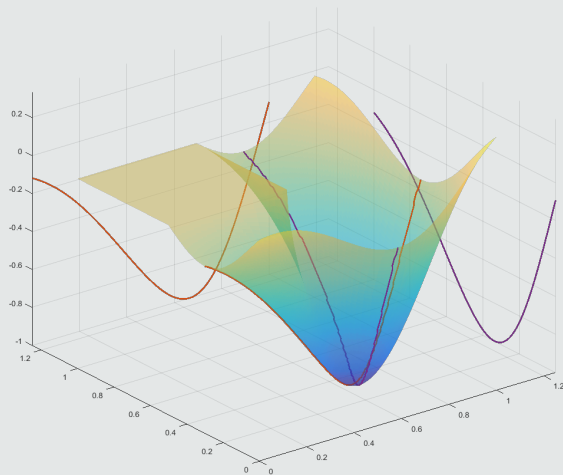
Profile likelihood:

$$\mathcal{L}_{P, X_1}(x_1) = \max_{x_2} \mathcal{L}(x_1, x_2)$$

Gi Profiling

Ma

Pro



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The profile likelihood is more natural under a frequentist lens, since the maximum likelihood estimator is left unchanged:

$$\max_{x_1, x_2} \mathcal{L}(x_1, x_2) = \max_{x_1} \mathcal{L}_{P, X_1}(x_1)$$

The generalised profiling method feels a lot like taking the profile likelihood of the inner objective.

Using ideas from mixed estimation (Theil & Golberger 1961), we can motivate the inner objective as a likelihood.

$$\mathcal{L}(\theta, c) = \|L(y - g(x))\|^2 + \left\| W \left(\frac{d}{dt}(\Phi_c) - f(\Phi_c; \theta) \right) \right\|^2$$

*Can also add regularisation to this in a justified manner

$$\mathcal{L}(\theta, c) = \|L(y - g(\Phi_c))\|^2 + \left\| W \left(\frac{d}{dt}(\Phi_c) - f(\Phi_c; \theta) \right) \right\|^2$$

Also estimating covariances $(LL^T)^{-1}$, $(WW^T)^{-1}$.

Can't do this all at once, so choose to use an iterative method (iteratively reweighted least squares).

1. Estimate (x, c) for a fixed (L, W) .
2. Estimate (L, W) for a fixed (x, c) .
3. Return to Step 1.

We can optimise the (log-)likelihood to get a point estimate (the MLE).

We really also want confidence sets to go with this.

Claim: The likelihood contains all the information about the uncertainty in the problem.

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All we need to do is to understand the behaviour of the likelihood with respect to quantities of interest.

Profile likelihood is a good way of doing this.

Also gives us information about identifiability.

If we have model $y = X(\theta) + \epsilon$, and we estimate MLE as $\hat{\theta}$ and the error as e , we can get an idea of the uncertainty in the estimate by sampling from fits of θ_{BS} for realisations of e .

$$X(\hat{\theta}) + e = X(\theta_{BS})$$

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$$X(\hat{\theta}) + e = X(\theta_{BS})$$

We can also use the randomised maximum likelihood interpretation of this, and sample from the perturbed data instead:

$$y + e = X(\theta_{RML})$$

Case Study

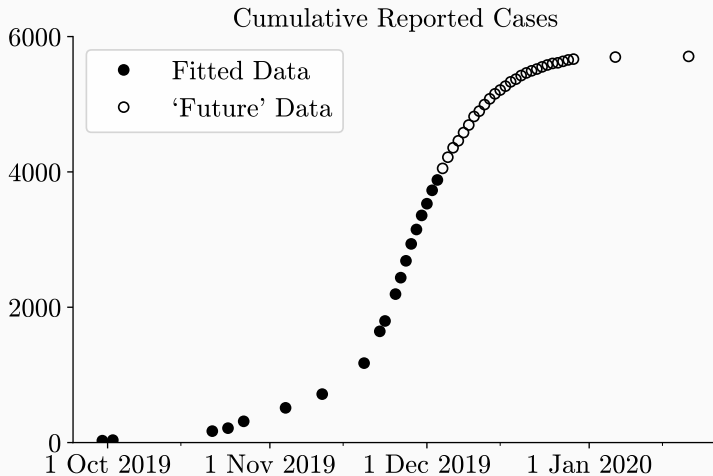
Case Study: 2019 Samoan Measles Outbreak



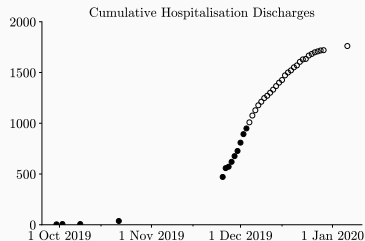
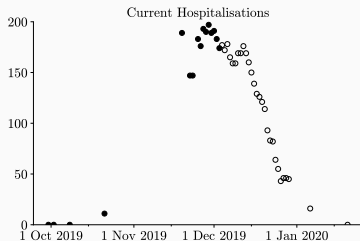
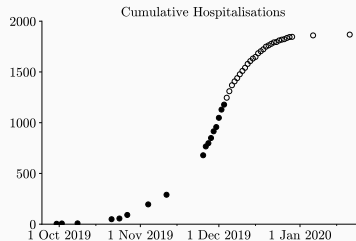
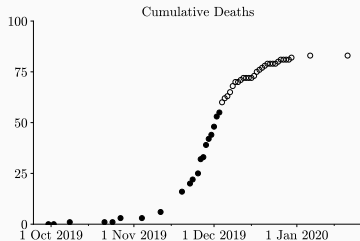
Measles outbreak in the small South Pacific island nation of Samoa

- September 2019 – January 2020
- Over 5700 cases (pop. 200 000)
- 83 deaths
- Anomalously low vaccination coverage in infants
 - 40% MCV1
 - 28% MCV2

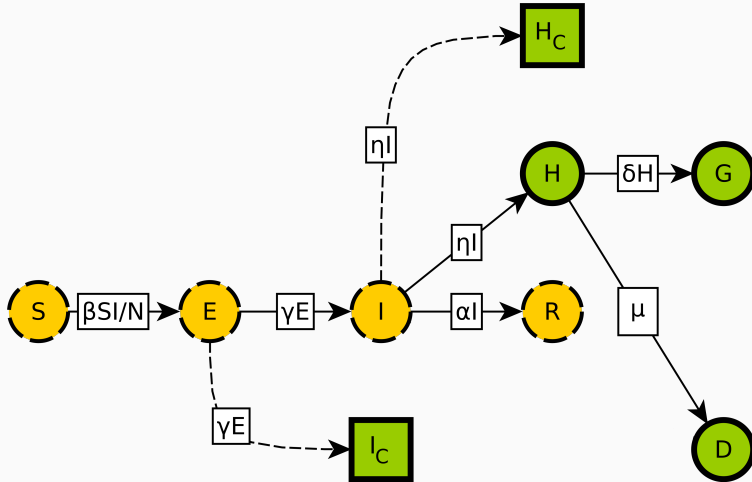
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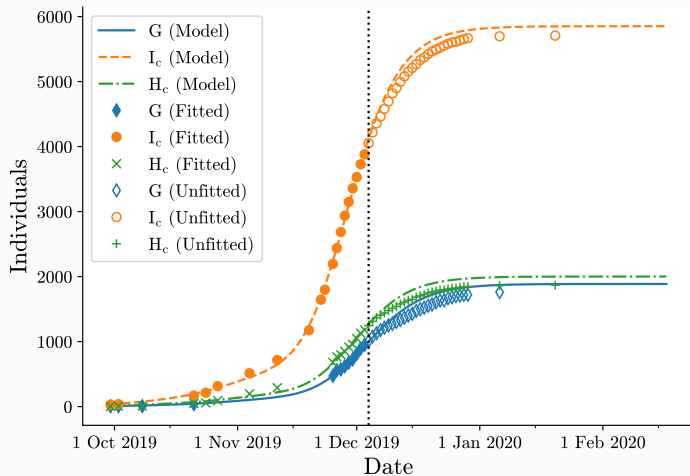
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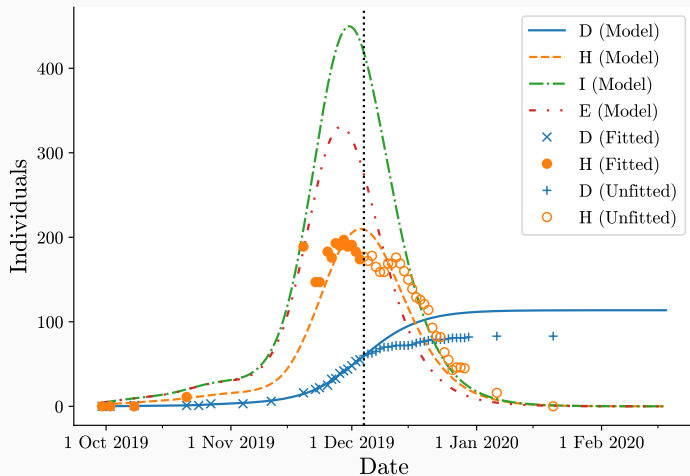
Case Study: Model



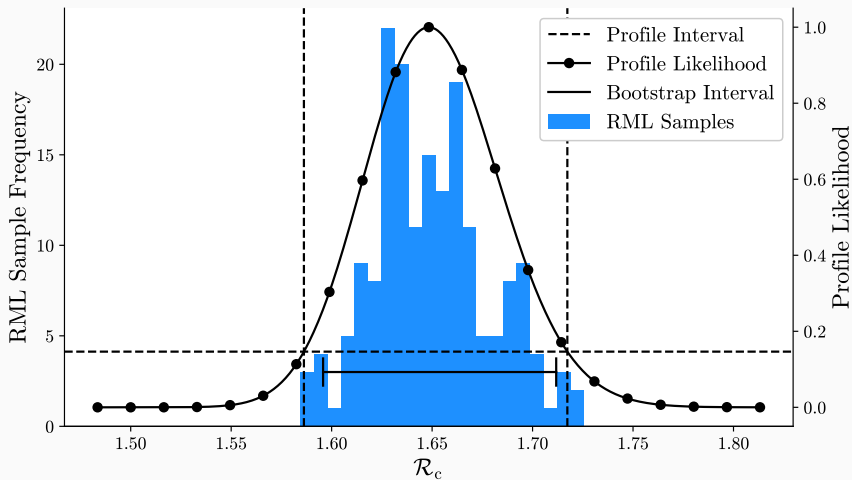
Case Study: Results (MLE)



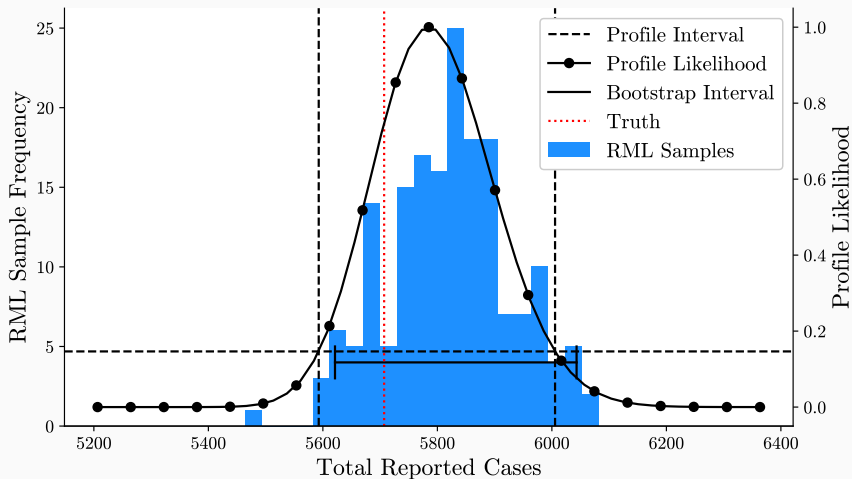
Case Study: Results (MLE)



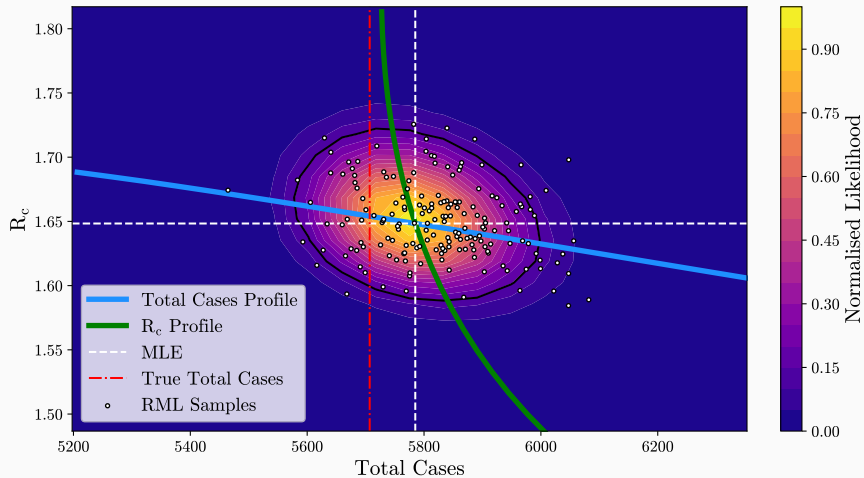
Case Study: Results (UQ)



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- Can avoid the integration/forward problem solve step by estimating the state and the parameters simultaneously
- There exist frequentist methods (profile likelihood, bootstrapping) to perform UQ on likelihood formulations
- This all seems to work on a case study of real outbreak data
- Though there is some tuning that needs to be done by hand

Thanks for coming!

Questions?

arXiv: 2103.16058

References

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