

## Math 66 - Fall 2023 - Goldwyn

## HW 8 - due Wednesday 11/22

## Please do the following:

1. Consider a Bernoulli process of length 5:

$$\{X_1, X_2, X_3, X_4, X_5\}$$

with  $P(X_i = 1) = p$  for each  $i$ . What is the probability of observing a realization of this process in which there are exactly three “successes” (occurrences of  $X_i = 1$ ) and these occur on consecutive time steps?

2. \* Let  $T_1$  be the time of first success for a Bernoulli process. Recall we derived the mass function for this random variable in class, it is a geometric random variable. Derive the fact that  $E[T_1] = 1/p$

*Hints:*

- Begin from the definition  $E[T_1] = \sum_{n=1}^{\infty} nP\{T_1 = n\}$
- Plug in the formula for  $P\{T_1 = n\}$  (an expression in terms of  $p$  and  $q$ , derived in class)
- To complete the problem: relating the series  $\sum_{n=1}^{\infty} nP\{T_1 = n\}$  to the derivative of a geometric series.
- *Reminder:* For  $|x| < 1$ , the geometric series is absolutely convergent with  $\sum_{n=0}^{\infty} x^n = 1/(1-x)$ .

**Challenge (optional):** If you can't get enough of working with geometric series, try a similar analysis to derive  $Var(T_1) = q/p^2$

## For the following two problems:

Use the properties of *independent* and *stationary* increments.

Please write down an exact expression (do not simplify more than you want to, do not use a calculator to find numerical values).

3. Recall that the probability that there are  $k$  successes in  $n$  time steps of a Bernoulli process is the Binomial distribution

$$P\{N_n = k\} = \binom{n}{k} p^k q^{n-k}$$

Find the probability of the event  $P\{N_3 = 1, N_{\bullet} = 4, N_{10} = 7\}$

4. \* Recall that the probability that the  $k^{th}$  event of a Bernoulli process occurs at time  $n$  is

$$P\{T_k = n\} = \binom{n-1}{k-1} p^k q^{n-k}$$

Find the probability of the event  $P\{T_1 = 3, T_5 = 9, T_7 = 17\}$

**This problem is about the Poisson process, a continuous time stochastic process**

5. *Poisson process.* The counting process  $N_n$  for the Bernoulli process counts the number of success in  $n$  steps of the process. The Poisson process can be thought of as a version of this in continuous time, call it  $N_t$ , which counts the number of “successes” in time  $t$ , where  $t$  takes is real-valued and non-negative. The defining properties of the Poisson process are:

- A Poisson process is defined by the parameter  $\lambda$  that is the rate of successes (also called “arrivals”)
- The increments of  $N_t$  are independent and stationary. As a consequence,

$$P(N_{t+s} - N_s = k) = P(N_t = k) \quad \text{for all } t > 0$$

- The number of arrivals in an interval of length  $t$  has a Poisson distribution:

$$P(N_t = k) = \frac{(\lambda t)^k}{k!} e^{-\lambda t}$$

- (a) Find  $P(T_1 \leq t)$ , the cumulative distribution function for  $T_1$ .

**Tip:** How can  $P(T_1 > t)$  be expressed using the Poisson distribution for  $N_t$ , the number of events by time  $t$ ?

- (b) Take a derivative with respect to  $t$  of the function in you found in (a) to find the probability density function for  $T_1$ .

**Comment:** The distribution and density functions in (a) and (b) define the **exponential distribution**.

- (c) Find  $E[T_1]$ .

**Tip:** Notice that  $T_1$  takes only non-negative values. Applying the definition of expected value for continuous random variables, you must evaluate:

$$E[T_1] = \int_0^{\infty} t f(t) dt$$

where  $f(t)$  is the probability density function you found in (b).

6. A random variable  $X$  is called **memoryless** with respect to  $t$  if it satisfies the following:

$$P(X > s + t \mid X > s) = P(X > t) \quad \text{for all } s \text{ with } t \neq 0$$

- (a) Show that  $T_1$  is memoryless for the case that  $T_1$  represents the first success of a Bernoulli process
- (b) Show that  $T_1$  is memoryless for the case that  $T_1$  represents the first success of a Poisson process

**Remark:** Recall that  $T_1$  for the Bernoulli process is a geometric random variables (discrete), and  $T_1$  for the Poisson process is an exponential random variable (continuous). It is a fact that geometric and exponential random variables are the only memoryless random variables.

1. Consider a Bernoulli process of length 5:

$$\{X_1, X_2, X_3, X_4, X_5\}$$

with  $P(X_i = 1) = p$  for each  $i$ . What is the probability of observing a realization of this process in which there are exactly three "successes" (occurrences of  $X_i = 1$ ) and these occur on consecutive time steps?

We can get exactly three "successes" which occur at consecutive time steps in 3 ways:

$$X_1, X_2, X_3 \quad ; \quad X_2, X_3, X_4 \quad ; \quad X_3, X_4, X_5$$

Since each  $X_i$  is independent, each of these realizations occur with probability

$$\left(P(X_i = 1)\right)^3 \left(P(X_i = 0)\right)^2 = p^3 (1-p)^2 \quad \text{as we need 3 successes and 2 failures.}$$

Thus, the total probability is  $\boxed{3p^3(1-p)^2}$

2. \* Let  $T_1$  be the time of first success for a Bernoulli process. Recall we derived the mass function for this random variable in class, it is a geometric random variable. Derive the fact that  $E[T_1] = 1/p$

Hints:

- Begin from the definition  $E[T_1] = \sum_{n=1}^{\infty} nP\{T_1 = n\}$
- Plug in the formula for  $P\{T_1 = n\}$  (an expression in terms of  $p$  and  $q$ , derived in class)
- To complete the problem: relating the series  $\sum_{n=1}^{\infty} nP\{T_1 = n\}$  to the derivative of a geometric series.
- *Reminder:* For  $|x| < 1$ , the geometric series is absolutely convergent with  $\sum_{n=0}^{\infty} x^n = 1/(1-x)$ .

**Challenge (optional):** If you can't get enough of working with geometric series, try a similar analysis to derive  $\text{Var}(T_1) = q/p^2$

$T_1$  = time of first success in Bernoulli process

$$P\{T_1 = n\} = q^{n-1} p \quad (\text{"failures" in first } n-1 \text{ occurrences, success in } n^{\text{th}})$$

$$E(T_1) = \sum_{n=1}^{\infty} n p \{T_1 = n\}$$

$$= \sum_{n=1}^{\infty} n q^{n-1} p = p \sum_{n=1}^{\infty} n q^{n-1} = p \sum_{n=1}^{\infty} n (1-p)^{n-1} \quad \text{as } q = 1-p.$$

Note that  $\sum_{n=1}^{\infty} n (1-p)^{n-1} = -\frac{d}{dp} \sum_{n=1}^{\infty} (1-p)^n$  (negative due to chain rule)

$$= -\frac{d}{dp} \left( \frac{1-p}{1-(1-p)} \right)$$

$$= -\frac{d}{dp} \left( \frac{1-p}{p} \right) = -\frac{d}{dp} \left( \frac{1}{p} - 1 \right) = -\frac{d}{dp} \left( \frac{1}{p} \right) + \frac{d}{dp} (-1) = \frac{1}{p^2}$$

Thus,  $p \sum_{n=1}^{\infty} n (1-p)^{n-1} = p \left( \frac{1}{p^2} \right) = \frac{1}{p}$ , so  $\boxed{E[T_1] = \frac{1}{p}}$  as desired.

For the following two problems:

Use the properties of *independent* and *stationary* increments.

Please write down an exact expression (do not simplify more than you want to, do not use a calculator to find numerical values).

3. Recall that the probability that there are  $k$  successes in  $n$  time steps of a Bernoulli process is the Binomial distribution

$$P\{N_n = k\} = \binom{n}{k} p^k q^{n-k}$$

Find the probability of the event  $P\{N_3 = 1, N_6 = 4, N_{10} = 7\}$

4. \* Recall that the probability that the  $k^{\text{th}}$  event of a Bernoulli process occurs at time  $n$  is

$$P\{T_k = n\} = \binom{n-1}{k-1} p^k q^{n-k}$$

Find the probability of the event  $P\{T_1 = 3, T_5 = 9, T_7 = 17\}$

3. Using increments, we know

$$\begin{aligned} & P\{N_3 = 1, N_6 = 4, N_{10} = 7\} \\ &= P\{N_3 = 1, N_6 - N_3 = 3, N_{10} - N_5 = 3\} \\ &= P\{N_3 = 1\} P\{N_6 - N_3 = 3\} P\{N_{10} - N_5 = 3\} \\ &= P\{N_3 = 1\} P\{N_3 = 3\} P\{N_5 = 3\} \end{aligned}$$

independent property of increments

Stationary property of increments

We can calculate these individually using the given formula.

$$\begin{aligned} P\{N_3 = 1\} &= \binom{3}{1} p^1 q^2 \\ P\{N_3 = 3\} &= \binom{3}{3} p^3 \\ P\{N_5 = 3\} &= \binom{5}{3} p^3 q^2 \end{aligned} \quad \Rightarrow \quad P\{N_3 = 1, N_6 = 4, N_{10} = 7\} = \binom{3}{1} \binom{3}{3} \binom{5}{3} p^{1+3+3} q^{2+0+2}$$

$$= \boxed{\binom{3}{1} \binom{3}{3} \binom{5}{3} p^8 q^3}$$

4. Using increments, we know

$$\begin{aligned} & P\{T_1 = 3, T_5 = 9, T_7 = 17\} \\ &= P\{T_1 = 3, T_5 - T_1 = 6, T_7 - T_5 = 12\} \\ &= P\{T_1 = 3\} P\{T_5 - T_1 = 6\} P\{T_7 - T_5 = 12\} \\ &= P\{T_1 = 3\} P\{T_4 = 6\} P\{T_2 = 12\} \end{aligned}$$

independent property of increments

Stationary property of increments

We can calculate these individually using the given formula.

$$\begin{aligned} P\{T_1 = 3\} &= \binom{2-1}{1-1} p^1 q^2 = \binom{1}{0} p^1 q^2 \\ P\{T_4 = 6\} &= \binom{6-1}{4-1} p^4 q^2 = \binom{5}{3} p^4 q^2 \\ P\{T_2 = 12\} &= \binom{12-1}{2-1} p^2 q^{10} = \binom{11}{1} p^2 q^{10} \end{aligned} \quad \Rightarrow \quad P\{T_1 = 3, T_5 = 9, T_7 = 17\} = \binom{1}{0} \binom{5}{3} \binom{11}{1} p^{1+4+2} q^{2+2+10}$$

$$= \boxed{\binom{1}{0} \binom{5}{3} \binom{11}{1} p^7 q^{14}}$$

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**Tip:** How can  $P(T_1 > t)$  be expressed using the Poisson distribution for  $N_t$ , the number of events by time  $t$ ?

- (b) Take a derivative with respect to  $t$  of the function in you found in (a) to find the probability density function for  $T_1$ .

**Comment:** The distribution and density functions in (a) and (b) define the **exponential distribution**.

- (c) Find  $E[T_1]$ .

**Tip:** Notice that  $T_1$  takes only non-negative values. Applying the definition of expected value for continuous random variables, you must evaluate:

$$E[T_1] = \int_0^{\infty} t f(t) dt$$

where  $f(t)$  is the probability density function you found in (b).

- (a)  $T_1$  is the time of the first success in Poisson process

$$\begin{aligned} P(T_1 \leq t) &= 1 - P(T_1 > t) \\ &= 1 - P(N_t = 0) \quad \downarrow \begin{array}{l} T_1 > t \text{ means there are no successes} \\ \text{yet at time } t, \text{ so } N_t = 0. \end{array} \\ &= 1 - \frac{(\lambda t)^0}{0!} e^{-\lambda t} \\ &= 1 - e^{-\lambda t} \end{aligned}$$

Thus,  $P(T_1 \leq t) = 1 - e^{-\lambda t}$  for  $t > 0$  is the CDF for  $T_1$ .

$$\begin{aligned} (b) \quad P(T_1 = t) &= \frac{d}{dt} (1 - e^{-\lambda t}) \\ &= \lambda e^{-\lambda t} \end{aligned}$$

so the PDF of  $T_1$  (the exponential distribution) is

$$P(T_1 = t) = \lambda e^{-\lambda t} \quad \text{for } t > 0$$

(i)

$$E[T_1] = \int_0^{\infty} t f(t) dt$$
$$= \int_0^{\infty} t \lambda e^{-\lambda t} dt$$

Using integration by parts, let  $u = t$ ,  $dv = \lambda e^{-\lambda t} dt$  so  
 $du = dt$ ,  $v = -e^{-\lambda t}$

$$\int_0^{\infty} \underbrace{t}_{u} \underbrace{\lambda e^{-\lambda t}}_{dv} dt = \left[ \underbrace{-te^{-\lambda t}}_{uv} \right]_0^{\infty} - \int_0^{\infty} \underbrace{-e^{-\lambda t}}_v \underbrace{dt}_{du}$$

(Note:  $\lambda = \text{rate of success} > 0$ )

$$= 0 - \left[ \frac{1}{\lambda} e^{-\lambda t} \right]_0^{\infty}$$

$$= 0 - \left( 0 - \frac{1}{\lambda} \right) = \frac{1}{\lambda}$$

Thus,  $E[T_1] = \frac{1}{\lambda}$

6. A random variable  $X$  is called **memoryless** with respect to  $t$  if it satisfies the following:

$$P(X > s + t \mid X > s) = P(X > t) \quad \text{for all } s \text{ with } t \neq 0$$

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**Remark:** Recall that  $T_1$  for the Bernoulli process is a geometric random variables (discrete), and  $T_1$  for the Poisson process is an exponential random variable (continuous). It is a fact that geometric and exponential random variables are the only memoryless random variables.

(a) Let  $T_1$  be the first success of a Bernoulli process.

$$\begin{aligned} &P(T_1 > s + t \mid T_1 > s) \\ &= \frac{P(T_1 > s, T_1 > s + t)}{P(T_1 > s)} \quad \text{by definition of conditional probability} \\ &= \frac{P(T_1 > s + t)}{P(T_1 > s)} \quad \text{since if } T_1 > s + t, T_1 > s. \\ &= \frac{P(N_{s+t} = 0)}{P(N_s = 0)} = \frac{\binom{s+t}{0} q^{s+t}}{\binom{s}{0} q^s} = q^t. \end{aligned}$$

But note that  $q^t = P(N_t = 0) = P(T_1 > t)$  as well (using similar logic as above).

Thus,

$$P(T_1 > s + t \mid T_1 > s) = P(T_1 > t), \text{ so } T_1 \text{ (first success in Bernoulli process) is memoryless. } \blacksquare$$

(b) Let  $T_1$  be the first success of a Poisson process.

$$\begin{aligned} &P(T_1 > s + t \mid T_1 > s) \\ &= \frac{P(T_1 > s, T_1 > s + t)}{P(T_1 > s)} \quad \text{by definition of conditional probability} \\ &= \frac{P(T_1 > s + t)}{P(T_1 > s)} \quad \text{since if } T_1 > s + t, T_1 > s. \\ &= \frac{1 - P(T_1 \leq s + t)}{1 - P(T_1 \leq s)}. \end{aligned}$$

Using CDF derived in 5(a), we have

$$P(T_1 \leq s + t) = 1 - e^{-\lambda(s+t)}, \quad P(T_1 \leq s) = 1 - e^{-\lambda s}.$$

$$\begin{aligned}
 \text{So } P(T_1 > s+t \mid T_1 > s) &= \frac{1 - P(T_1 \leq s+t)}{1 - P(T_1 \leq s)} \\
 &= \frac{1 - (1 - e^{-\lambda(s+t)})}{1 - (1 - e^{-\lambda s})} \quad \text{by plugging in above expressions for } P(T_1 \leq s+t), P(T_1 \leq s) \\
 &= \frac{e^{-\lambda(s+t)}}{e^{-\lambda s}} = e^{-\lambda t}
 \end{aligned}$$

$$\begin{aligned}
 \text{Note that } e^{-\lambda t} &= 1 - (1 - e^{-\lambda t}) \\
 &= 1 - P(T_1 \leq t) \\
 &= P(T_1 > t) \quad \text{by the same logic as above.}
 \end{aligned}$$

Thus,

$$P(T_1 > s+t \mid T_1 > s) = P(T_1 > t), \text{ so } T_1 \text{ (first success in Poisson process) is memoryless. } \blacksquare$$