Homework 5 David Yang

- 1. Suppose  $X_1, \ldots, X_n$  are i.i.d. Normal with mean  $\mu$  and variance  $\sigma^2$ .
  - a) Show the MLE for  $\sigma$  is the square root of the MLE for  $\sigma^2$  (in general, if  $\hat{\theta}$  is the MLE of  $\theta$ , then  $\hat{\theta}$  is the MLE of  $\theta$ , then  $\hat{\varphi} = g(\hat{\theta})$  is the MLE for  $\varphi = g(\theta)$ ).

Solution. If  $X_1, \ldots, X_n$  are i.i.d. Normal with mean  $\mu$  and variance  $\sigma^2$ ,

$$L(\mu, \sigma^2) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2} \frac{(x_i - \mu)^2}{2\sigma^2}}$$
$$= (2\pi\sigma^2)^{-\frac{n}{2}} e^{-\frac{\sum (x_i - \mu)^2}{2\sigma^2}}.$$

The log-likelihood is

$$l(\mu, \sigma^2) = n \log \left(\frac{1}{\sqrt{2\pi\sigma^2}}\right) - \frac{\sum (x_i - \mu)^2}{2\sigma^2}$$
$$= -\frac{n}{2} \log \left(2\pi\sigma^2\right) - \frac{\sum (x_i - \mu)^2}{2\sigma^2}.$$

Setting the partials to 0 to solve for the MLEs, we get that

$$\begin{split} \frac{\partial l}{\partial \mu} &= 2 \frac{\sum (x_i - \mu)}{2\sigma^2} \\ \frac{\partial l}{\partial \sigma^2} &= -\frac{n}{2} \cdot \frac{1}{2\pi\sigma^2} \cdot 2\pi - \frac{1}{2} \cdot \left(\sum (x_i - \mu)^2\right) \cdot \left(-\frac{1}{\sigma^4}\right) \\ &= -\frac{n}{2\sigma^2} + \frac{1}{2\sigma^4} \sum (x_i - \mu)^2 \end{split}$$

The first equation is equivalent to  $\sum (x_i - \mu) = 0$ , so solving for  $\mu$  gives

$$\hat{\mu} = \frac{\sum x_i}{n} = \overline{x}.$$

The second equation is equivalent to  $\frac{n}{2\sigma^2} = \frac{1}{2\sigma^4} \sum (x_i - \mu)^2$ , and solving for  $\sigma^2$  gives

$$\sigma^2 = \frac{\sum (x_i - \mu)^2}{n}.$$

Plugging in the MLE for  $\mu$ , we get the MLE for  $\hat{\sigma}^2$ :

$$\hat{\sigma}^2 = \frac{\sum (x_i - \mu)^2}{n}$$
$$= \left[\frac{1}{n} \left(\sum (x_i - \overline{x})^2\right)\right]. \blacksquare$$

b) <b>Sho</b>	w that $\hat{\sigma}^2$	$= \sum \frac{(X_i - \bar{X})^2}{n}$	has	$\mathbf{smaller}$	mean	square	error	than	the	unbiase	d
esti	mate $s^2 = \sum_{i=1}^{n} a_i s_i$	$\sum \frac{(X_i - \overline{X})^2}{n-1}$ .									

Solution.

- c) Find an expression for the expected value of the sample standard deviation  $s=\sqrt{s^2}$ . Use this to construct an unbiased estimate for  $\sigma$ . You could check your answer by generating sample variances from the Gamma distribution implied by a given n and  $\sigma$ .
- 2. Suppose  $X_1, \ldots, X_n$  are i.i.d. Poisson( $\theta$ ).
  - a) Let  $Y = \sum X_i$ . If n = 3, find  $P(X_1 = 2, X_2 = 3, X_3 = 0 \mid Y = 5, \theta = 2)$ . How does this change if  $\theta = 3$ ?

Solution.

b) Show that  $\overline{X}$  is the MLE for  $\theta$ , and that the reciprocal Fisher information gives the exact variance of  $\hat{\theta}$ . Verify that  $\operatorname{Poisson}(\theta)$  is a 1-parameter exponential family, so you know you can use the second derivative formula for the Fisher information.

Solution.

c) If n=10 and  $\theta=10$ , check the Normal approximation to the distribution of  $\overline{X}$ . Use the continuity correction to approximate  $P(\overline{X}=10)$  and compare this to the exact probability based on the associated Poisson distribution.

Solution.

d) Show that  $\theta \sim \operatorname{Gamma}\left(\alpha, \frac{\alpha}{u}\right)$  defines a conjugate family of prior distributions for  $\theta$ , and that the posterior mean for  $\theta \mid \overline{x}$  is a weighted average of  $\overline{x}$  and the prior mean  $\mu$ , with more weight on  $\overline{x}$  as n increases.

Solution.