

distribution	pmf and domain	$E(X)$	$Var(X)$	mgf $M(\theta)$
<i>Bernoulli</i> (p)	$p(x) = p^x(1-p)^{1-x}, x = 0, 1$	p	$p(1-p)$	$1 - p + pe^\theta$
<i>Binomial</i> (n, p)	$p(x) = \binom{n}{x}p^x(1-p)^{n-x}, x = 0, 1, \dots, n$	np	$np(1-p)$	$(1 - p + pe^\theta)^n$
<i>Poisson</i> (λ)	$p(x) = \frac{e^{-\lambda}\lambda^x}{x!}, x = 0, 1, 2, \dots$	λ	λ	$e^{\lambda(e^\theta - 1)}$
<i>Geometric</i> (p)	$p(x) = p(1-p)^{x-1}, x = 1, 2, 3, \dots$	$\frac{1}{p}$	$\frac{1-p}{p^2}$	$\frac{pe^\theta}{1-(1-p)e^\theta}$ $\theta < -\ln(1-p)$
<i>Neg.bin</i> (r, p)	$p(x) = \binom{x-1}{r-1}p^r(1-p)^{x-r},$ $x = r, r+1, r+2, \dots$	$\frac{r}{p}$	$\frac{r(1-p)}{p^2}$	$\left(\frac{pe^\theta}{1-(1-p)e^\theta}\right)^r$ $\theta < -\ln(1-p)$
<i>Hyp.geom</i> (n, M, N)	$p(x) = \frac{\binom{M}{x}\binom{N-M}{n-x}}{\binom{N}{n}}, x = 0, 1, \dots, n$ $x \leq M, n-x \leq N-M$	$\frac{nM}{N}$	$\frac{nM}{N}(1 - \frac{M}{N})(\frac{N-n}{N-1})$	
<i>Discrete uniform</i>	$p(x) = \frac{1}{n}, x \in \{x_1, x_2, \dots, x_n\}$	$\frac{\sum_{i=1}^n x_i}{n}$	$\frac{\sum_{i=1}^n x_i^2 - \frac{(\sum_{i=1}^n x_i)^2}{n}}{n}$	$\frac{1}{n} \sum_{i=1}^n e^{\theta x_i}$
<i>logarithmic</i> (p)	$p(k) = -\frac{p^k}{k \ln(1-p)}, k = 1, 2, 3, \dots$	$-\frac{p}{(1-p) \ln(1-p)}$	$-\frac{p^2 + p \ln(p)}{(1-p)^2 (\ln(1-p))^2}$	$\frac{\ln(1-pe^\theta)}{\ln(1-p)}$ $\theta < -\ln(p)$
<i>multinomial</i>	$p(x_1, x_2, \dots, x_r) = \frac{n!}{x_1!x_2!\dots x_r!}p_1^{x_1}p_2^{x_2}\dots p_r^{x_r}$	$E(X_i) = np_i$	$\sigma_i^2 = np_i(1-p_i)$ $\sigma_{i,j} = -np_i p_j$ for $i \neq j$	

Euler Gamma function: for $\alpha > 0$, $\Gamma(\alpha) = \int_0^\infty u^{\alpha-1}e^{-u} du$.

Some properties of the Euler Gamma function:

1. If $n > 0$ is an integer, then $\Gamma(n) = (n-1)!$, **2.** $\Gamma(\frac{1}{2}) = \sqrt{\pi}$, **3.** If $x > 1$, then $\Gamma(x) = (x-1)\Gamma(x-1)$

Binomial theorem: $(a+b)^n = \sum_{j=1}^n \binom{n}{j} a^j b^{n-j}$. Geometric series: $\sum_{j=m}^\infty r^j = \frac{r^m}{1-r}$.

Liebniz rule: $\frac{d}{dx} \int_{a(x)}^{b(x)} g(x, y) dy = g(x, b(x))b'(x) - g(x, a(x))a'(x) + \int_{a(x)}^{b(x)} \frac{\partial}{\partial x} f(x, y) dy$.

distribution	pdf and domain	$E(X)$	$Var(X)$	mgf $M(\theta)$
$uniform(a, b)$	$f(x) = \frac{1}{b-a}, a < x < b$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	$\frac{e^{\theta b} - e^{\theta a}}{(b-a)\theta}$
$Exp(\lambda)$	$f(x) = \lambda e^{-\lambda x}, x > 0$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	$(1 - \frac{\theta}{\lambda})^{-1}$ $\theta < \lambda$
$Gamma(\alpha, \beta)$	$f(x) = \frac{x^{\alpha-1} e^{-x/\beta}}{\beta^\alpha \Gamma(\alpha)}, x > 0$	$\alpha\beta$	$\alpha\beta^2$	$(1 - \beta\theta)^{-\alpha}$ $\theta < 1/\beta$
$Normal(\mu, \sigma^2)$	$f(x) = \frac{e^{-\frac{1}{2}(\frac{x-\mu}{\sigma})^2}}{\sqrt{2\pi}\sigma^2}, -\infty < x < \infty$	μ	σ^2	$e^{\mu\theta + \frac{\sigma^2\theta^2}{2}}$
$log - normal(\mu, \sigma^2)$	$f(x) = \frac{e^{-\frac{(\ln(x)-\mu)^2}{2\sigma^2}}}{x\sqrt{2\pi}\sigma^2}, 0 < x < \infty$	$e^{\mu + \frac{\sigma^2}{2}}$	$[e^{\sigma^2} - 1]e^{2\mu + \sigma^2}$	
$chi - square \text{ with } \nu \text{ d.f. } \chi_\nu^2$	$f(x) = \frac{x^{\frac{\nu}{2}-1} e^{-x/2}}{2^{\frac{\nu}{2}} \Gamma(\frac{\nu}{2})}, x > 0$	ν	2ν	$(1 - 2\theta)^{-\frac{\nu}{2}}$ $\theta < \frac{1}{2}$
$Beta(\alpha, \beta)$ $\alpha > 0, \beta > 0$	$f(x) = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1}(1-x)^{\beta-1}, 0 < x < 1$	$\frac{\alpha}{\alpha+\beta}$	$\frac{\alpha\beta}{(\alpha+\beta)^2(\alpha+\beta+1)}$	
$F_{n,m} - \text{distribution}$ $n \text{ numer}, m \text{ denom df}$	$f(x) = \frac{\Gamma(\frac{n+m}{2})}{\Gamma(\frac{n}{2})\Gamma(\frac{m}{2})} (\frac{n}{m})^{\frac{n}{2}} x^{\frac{n}{2}-1} (1 + \frac{nx}{m})^{-\frac{n+m}{2}}$ $\text{for } x > 0$	$\frac{m}{m-2}$ $m > 2$	$\frac{2m^2(n+m-2)}{n(m-2)^2(m-4)}$ $m > 4$	
$t_m - \text{distribution}$ $m \text{ df}$	$f(x) = \frac{\Gamma(\frac{m+1}{2})}{\sqrt{\pi m} \Gamma(\frac{m}{2})} (1 + \frac{t^2}{m})^{-(\frac{m+1}{2})}$ $\text{for } -\infty < x < \infty$	0 $m > 1$	$\frac{m}{m-2}$ $m > 2$	
$Laplace \text{ dist./double exp. } \lambda > 0, c \in \mathbb{R}$	$f(x) = \frac{\lambda}{2} e^{-\lambda x-c }, x \in \mathbb{R}$	c	$\frac{2}{\lambda^2}$	$\frac{e^{c\theta}}{1 - \frac{\theta^2}{\lambda^2}}$ $-\lambda < \theta < \lambda$
$bivariate \text{ normal}$	$f(x, y) = \frac{e^{-\frac{Q(x, y)}{2(1-\rho^2)}}}{2\pi\sigma_x\sigma_y\sqrt{1-\rho^2}}, (x, y) \in \mathbb{R}^2$ $Q(x, y) = (\frac{x-\mu_x}{\sigma_x})^2 - 2\rho(\frac{x-\mu_x}{\sigma_x})(\frac{y-\mu_y}{\sigma_y}) + (\frac{y-\mu_y}{\sigma_y})^2$	$E(X) = \mu_x$ $E(Y) = \mu_y$	$Var(X) = \sigma_x^2$ $Var(Y) = \sigma_y^2$ $corr(X, Y) = \rho$	
$Pareto$ $\alpha > 0, \lambda > 0$	$f(x) = \frac{\alpha\lambda^\alpha}{x^{\alpha+1}}, x > \lambda$	$\frac{\alpha\lambda}{\alpha-1}$ $\alpha > 1$	$\frac{\alpha\lambda^2}{(\alpha-1)^2(\alpha-2)}$ $\alpha > 2$	