

STAT350 Tutorial 6

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This week's tutorial will cover diagnostics for leverage and influence. This is an extension to what we covered in previous tutorials with regards to outliers (observations with unusual response values) and residual analysis, in that we will be looking at the data we have for points that effect our model.

We will be using a different dataset this week; the savings data from the **faraway** package which includes per capita financial information for 50 countries.

```
library(faraway)
data(savings)
head(savings)
```

```
##           sr pop15 pop75      dpi ddpi
## Australia 11.43 29.35  2.87 2329.68 2.87
## Austria   12.07 23.32  4.41 1507.99 3.93
## Belgium   13.17 23.80  4.43 2108.47 3.82
## Bolivia    5.75 41.89  1.67  189.13 0.22
## Brazil    12.88 42.19  0.83  728.47 4.56
## Canada     8.79 31.72  2.85 2982.88 2.43
```

In this dataset **sr** is the response and represents the ratio of savings to disposable income, **dpi** is the disposable income, **ddpi** is the rate of change for disposable income, and **pop15** and **pop75** represent the percentage of the population which are below the age of 15 and above the age of 75 respectively.

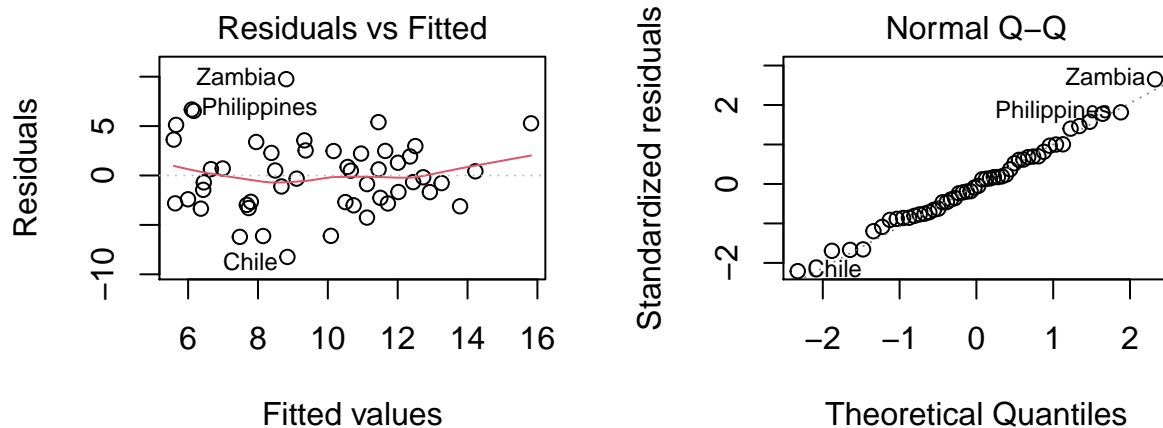
I'm now going to fit the multiple linear regression model with all four predictors included. Here using `.` to the right of the `~` tells R to use all other variables in the model.

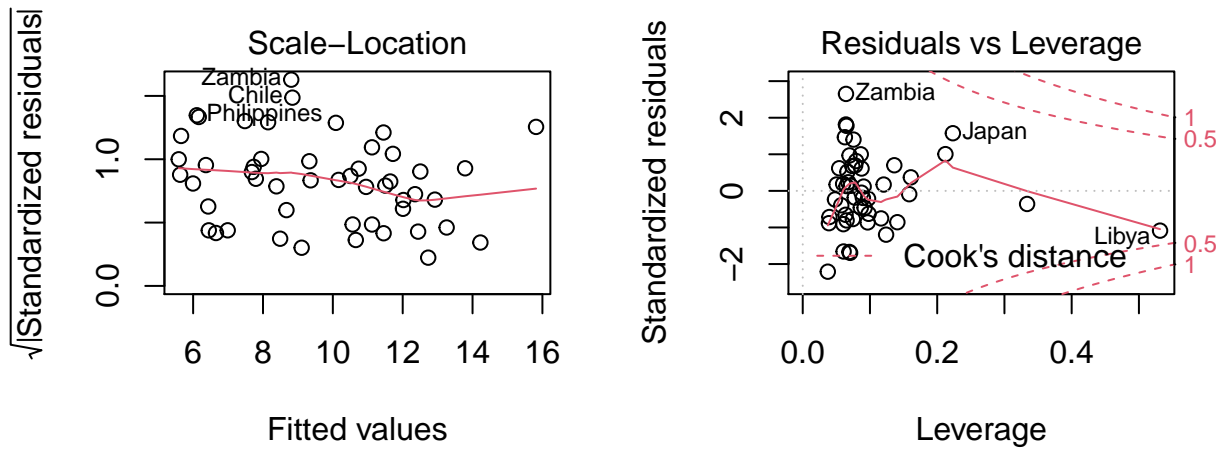
```
mdl <- lm(sr ~ ., data = savings)
(mdl_sum <- summary(mdl))
```

```
##
```

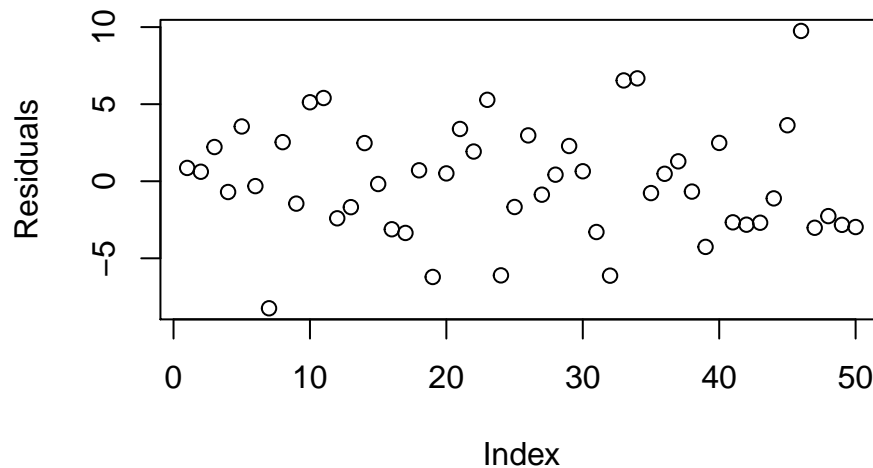
```
## Call:
## lm(formula = sr ~ ., data = savings)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -8.2422 -2.6857 -0.2488  2.4280  9.7509
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 28.5660865   7.3545161   3.884 0.000334 ***
## pop15       -0.4611931   0.1446422  -3.189 0.002603 **
## pop75       -1.6914977   1.0835989  -1.561 0.125530
## dpi         -0.0003369   0.0009311  -0.362 0.719173
## ddpi         0.4096949   0.1961971   2.088 0.042471 *
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.803 on 45 degrees of freedom
## Multiple R-squared:  0.3385, Adjusted R-squared:  0.2797
## F-statistic: 5.756 on 4 and 45 DF,  p-value: 0.0007904
```

So, everything looks good in the model summary – now we'll check assumptions





From the first plot we can see that the linear assumption is met. The second plot shows us that the normality of errors assumption is met. The third along with the first plot allow us to say that the constant variance assumption is met. Lastly, from the below residuals versus index plot we see that the uncorrelated errors assumption is met.



As for unusual points, Zambia and Chile should be checked due to their relatively large residuals. And, Libya should be checked due to its high leverage (we'll go more into this in the next section).

Leverage

The leverage of a point is defined by its place in the design space; a point far from all others can have a great effect on the properties of a given regression model. Leverages, h_{ii} , are given by the diagonal elements of the hat matrix:

$$H = X(X'X)^{-1}X.$$

A common rule given is that points whose leverage exceeds $2p/n$ should be looked at more closely, where n is the number of data points and p is the number of coefficients in the model. For our dataset that value is:

```
2*ncol(savings)/nrow(savings)
```

```
## [1] 0.2
```

A quick way to get the influences of points in R is with the `influence()` function, using the model object as its only argument. The `hat` component of the resulting object gives us the diagonal of the hat matrix.

```
mdl_inf <- influence(mdl)
sort(mdl_inf$hat, decreasing = TRUE)
```

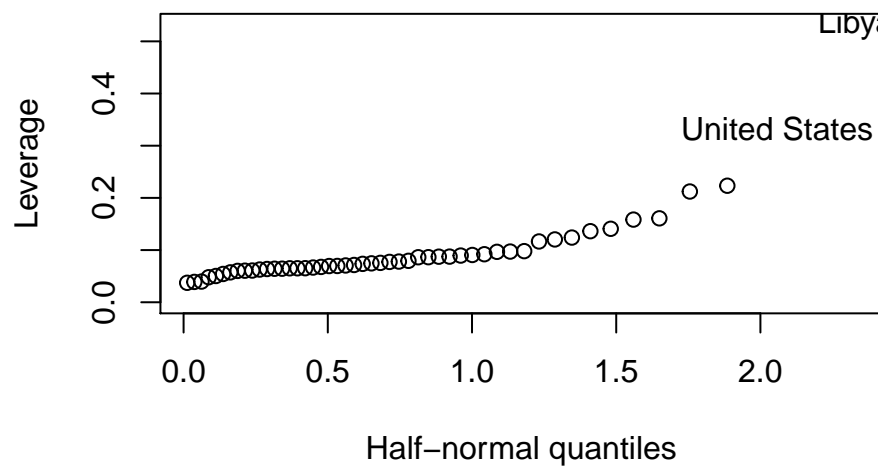
##	Libya	United States	Japan	Ireland	South Rhodesia
##	0.53145676	0.33368800	0.22330989	0.21223634	0.16080923
##	Canada	Jamaica	France	Sweden	Austria
##	0.15840239	0.14076016	0.13620478	0.12398898	0.12038393
##	United Kingdom	Uruguay	Portugal	Greece	Finland
##	0.11651375	0.09794717	0.09714946	0.09662073	0.09204246
##	Netherlands	Bolivia	Belgium	Germany	Luxembourg
##	0.09061400	0.08947114	0.08748248	0.08735739	0.08634787
##	Venezuela	Malta	China	Spain	Costa Rica
##	0.08628365	0.07940290	0.07795899	0.07732854	0.07546780
##	Tunisia	Switzerland	India	Iceland	Brazil
##	0.07456729	0.07359423	0.07145213	0.07049590	0.06955944
##	Paraguay	Australia	Italy	Malaysia	South Africa
##	0.06937188	0.06771343	0.06651170	0.06523300	0.06510405
##	Peru	Zambia	Philippines	Ecuador	Denmark
##	0.06504891	0.06433163	0.06425415	0.06372651	0.06271782
##	Korea	Guatemala	Honduras	Colombia	New Zealand

##	0.06079915	0.06049212	0.06008079	0.05730171	0.05421789
##	Nicaragua	Norway	Turkey	Panama	Chile
##	0.05035056	0.04793213	0.03964224	0.03897459	0.03729796

So, Libya, the US, Japan, and Ireland have leverages higher than 0.2 and should be looked at closer as they have the potential to be influential. Typically a point which has high leverage paired with a large residual it is likely to be influential. None of these countries had large residuals, so we'll have to dig a little more.

The `halfnorm()` function of the `faraway` package provides a useful way to visualize these leverages.

```
halfnorm mdl_inf$hat, labs = names(mdl_inf$hat), ylab = 'Leverage')
```



Influence

Influential points in the data are points that if removed would cause noteworthy changes in the estimates of the model coefficients. They typically have large leverages paired with unusual response values, therefore giving it more say in the fitted model values compared to other points.

Cook's Distance

Cook's distance is a measure of the distance between predicted values between models fit with and without point i . Your textbook says points with $D_i > 1$ are considered to be influential.

$$D_i = \frac{r_i^2}{p} \frac{h_{ii}}{1 - h_{ii}}$$

Now, I'll calculate the cook's distance for our model using the `cooks.distance()` function with the model object as it's only argument.

```
mdl_cook <- cooks.distance(mdl)
sort(mdl_cook, decreasing = TRUE)
```

##	Libya	Japan	Zambia	Ireland	Philippines
##	2.680704e-01	1.428162e-01	9.663275e-02	5.439637e-02	4.522120e-02
##	Peru	Iceland	Paraguay	Sweden	Chile
##	4.401457e-02	4.352902e-02	4.157229e-02	4.055963e-02	3.781324e-02
##	Korea	Costa Rica	Denmark	Jamaica	Venezuela
##	3.555386e-02	3.207537e-02	2.879580e-02	2.402677e-02	1.886141e-02
##	Greece	France	United Kingdom	Brazil	United States
##	1.590102e-02	1.547176e-02	1.496628e-02	1.402735e-02	1.284481e-02
##	Malta	Guatamala	Tunisia	Malaysia	Uruguay
##	1.146827e-02	1.067111e-02	9.562447e-03	9.113404e-03	8.532329e-03
##	China	Switzerland	Belgium	Panama	Ecuador
##	8.156984e-03	7.334746e-03	7.154674e-03	6.333674e-03	5.818699e-03
##	South Rhodesia	New Zealand	Finland	Turkey	Luxembourg
##	5.267290e-03	4.379219e-03	4.364051e-03	4.224370e-03	3.993882e-03
##	Italy	Colombia	Portugal	Austria	Australia
##	3.919100e-03	1.879460e-03	9.733900e-04	8.175997e-04	8.035888e-04
##	Bolivia	Spain	Norway	Honduras	Nicaragua
##	7.278744e-04	5.659085e-04	5.558570e-04	4.741920e-04	3.226479e-04
##	Canada	India	Netherlands	South Africa	Germany
##	3.106199e-04	2.965778e-04	2.744377e-04	2.405063e-04	4.736572e-05

So, we see that for our data and model we don't get any D_i larger than 1, with Libya having the greatest value at 0.268.

How does the model change if we delete Libya from the data set?

```
mdl_lib <- lm(sr ~ ., savings, subset = (row.names(savings) != 'Libya'))
summary(mdl_lib)
```

```
##
## Call:
## lm(formula = sr ~ ., data = savings, subset = (row.names(savings) !=
##      "Libya"))
##
## Residuals:
```

	Min	1Q	Median	3Q	Max
	-8.0699	-2.5408	-0.1584	2.0934	9.3732

```
##
## Coefficients:
```

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	24.5240460	8.2240263	2.982	0.00465 **
pop15	-0.3914401	0.1579095	-2.479	0.01708 *
pop75	-1.2808669	1.1451821	-1.118	0.26943
dpi	-0.0003189	0.0009293	-0.343	0.73312
ddpi	0.6102790	0.2687784	2.271	0.02812 *

```
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 3.795 on 44 degrees of freedom
## Multiple R-squared:  0.3554, Adjusted R-squared:  0.2968
## F-statistic: 6.065 on 4 and 44 DF,  p-value: 0.0005617
```

Here's the original model summary:

```
##
## Call:
## lm(formula = sr ~ ., data = savings)
##
## Residuals:
```

	Min	1Q	Median	3Q	Max
	-8.2422	-2.6857	-0.2488	2.4280	9.7509

```
##
```

```
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 28.5660865  7.3545161   3.884 0.000334 ***
## pop15       -0.4611931  0.1446422  -3.189 0.002603 **
## pop75       -1.6914977  1.0835989  -1.561 0.125530
## dpi         -0.0003369  0.0009311  -0.362 0.719173
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##
## Residual standard error: 3.803 on 45 degrees of freedom
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```

So, the biggest difference I see is in the coefficient for `ddpi`, changing from 0.41 to 0.61 – quite the change considering the magnitude of the coefficient.

So, what do we do with these points?

The answer to this question is heavily dependent on context and similar to how we treat outliers. That is, if we know the point is influential due to experimental error it is safe to delete the point. Otherwise we should leave it in. There are methods of regression that are less susceptible to changes caused by the presence of influential points.