

Regression discontinuity I & II

April 1, 2020

PMAP 8521: Program Evaluation for Public Service
Andrew Young School of Policy Studies
Spring 2020

Plan for today

Arbitrary cutoffs & causal inference

Drawing lines & measuring gaps

Main RDD concerns

RDD with R

Arbitrary cutoffs & causal inference

Rules to access programs

**Lots of policies and programs are based
on arbitrary rules and thresholds**

**If you're above the threshold, you're in
the program; if you're below, you're not**

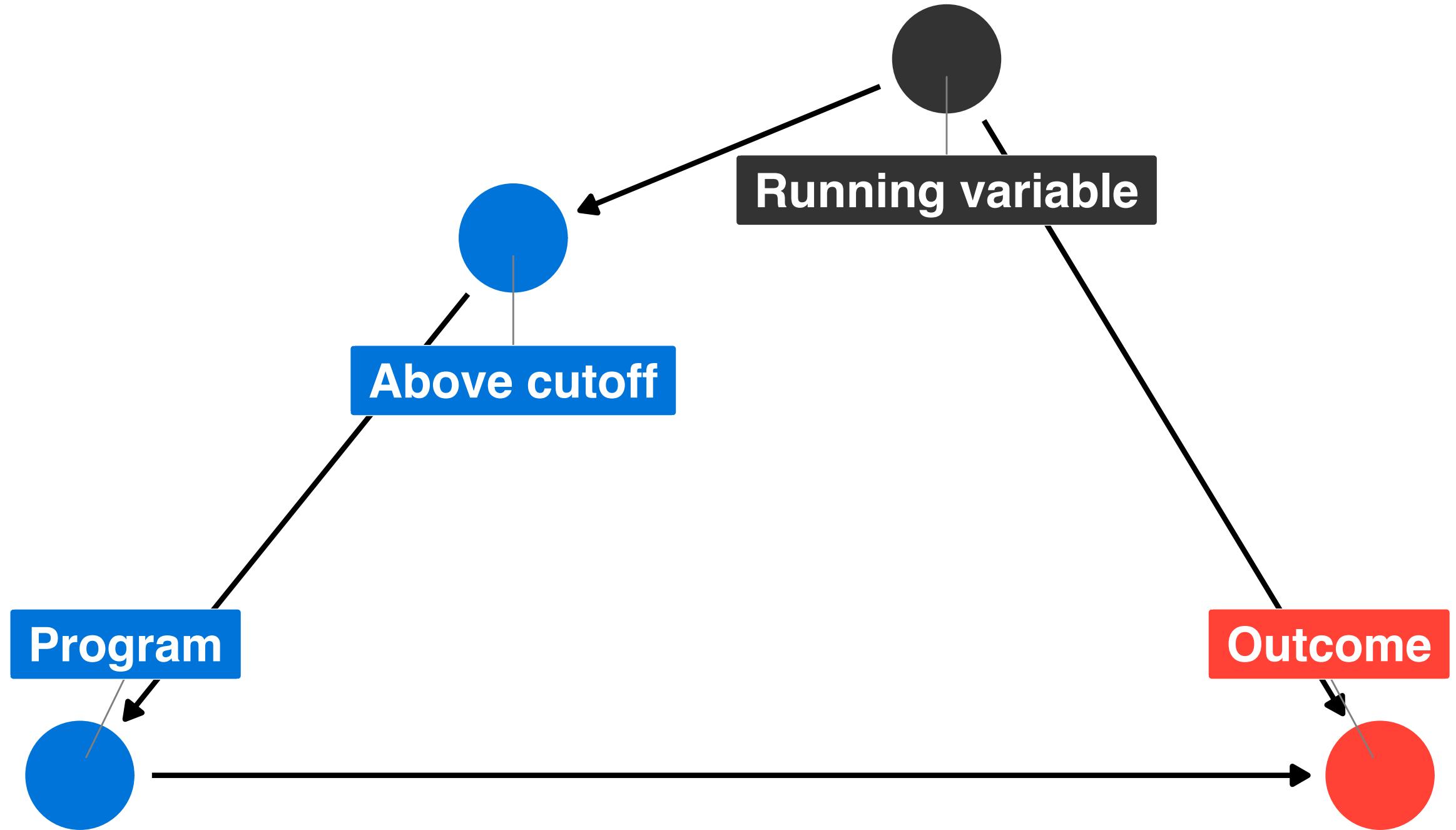
Key terms

Running/forcing variable

Index or measure that determines eligibility

Cutoff/cutpoint/threshold

Number that formally assigns access to program



Discontinuities everywhere!

Size	Annual	Monthly	138%	150%	200%	Medicaid 138%
1	\$12,760	\$1,063	\$17,609	\$19,140	\$25,520	ACA subsidies 100*-400%
2	\$17,240	\$1,437	\$23,791	\$25,860	\$34,480	CHIP 200%
3	\$21,720	\$1,810	\$29,974	\$32,580	\$43,440	SNAP/Free lunch 130%
4	\$26,200	\$2,183	\$36,156	\$39,300	\$52,400	Reduced lunch 130-185%
5	\$30,680	\$2,557	\$42,338	\$46,020	\$61,360	
6	\$35,160	\$2,930	\$48,521	\$52,740	\$70,320	
7	\$39,640	\$3,303	\$54,703	\$59,460	\$79,280	
8	\$44,120	\$3,677	\$60,886	\$66,180	\$88,240	

Hypothetical AIG program

If you score 75+ on a test, you get into
an academically and intellectually
gifted (AIG) during-school program

Participated in AIG program

TRUE
FALSE

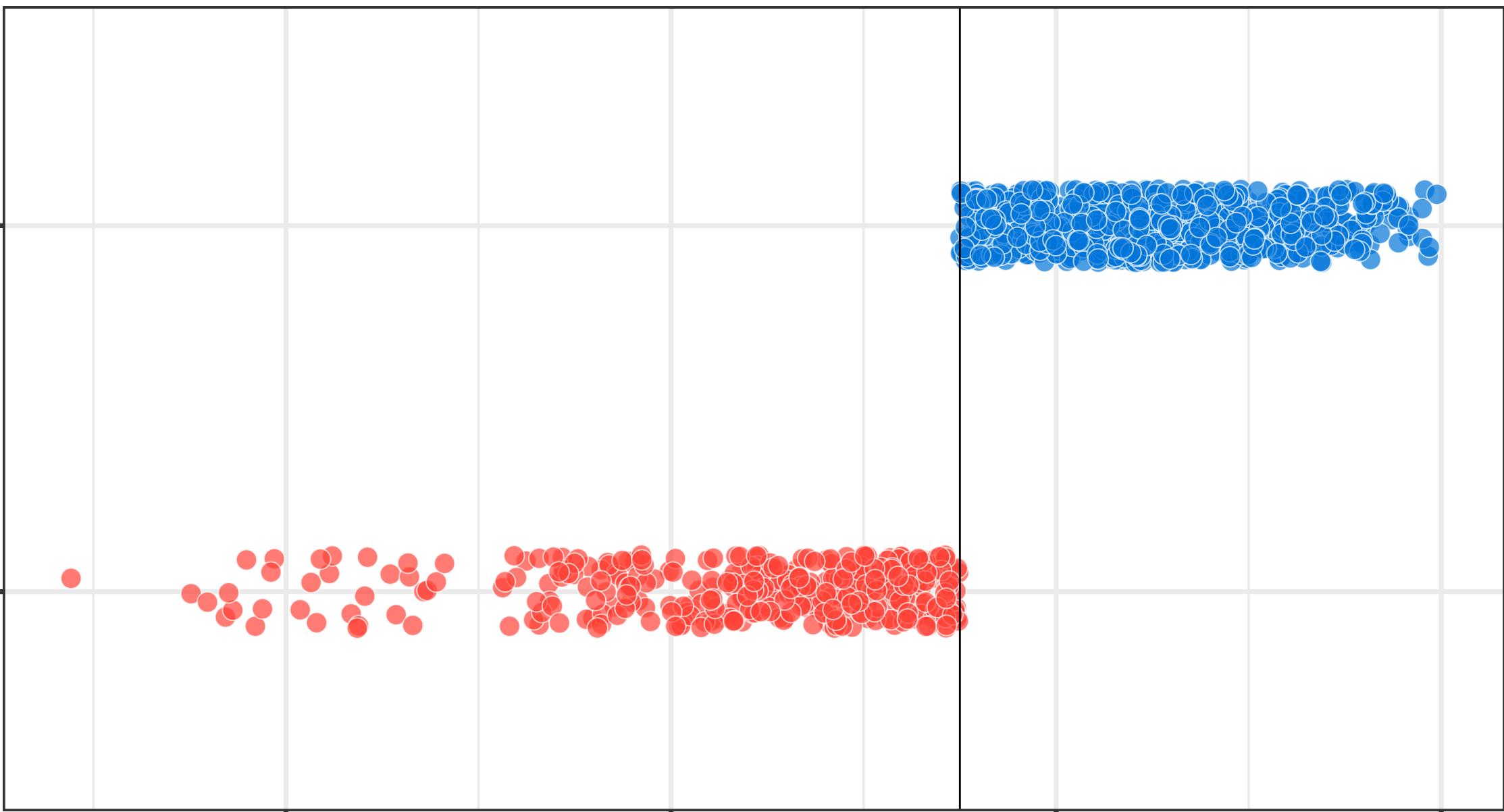
40

60

80

100

AIG test score



Causal inference intuition

People right before and right after the threshold are essentially the same

Participated in AIG program

TRUE
FALSE

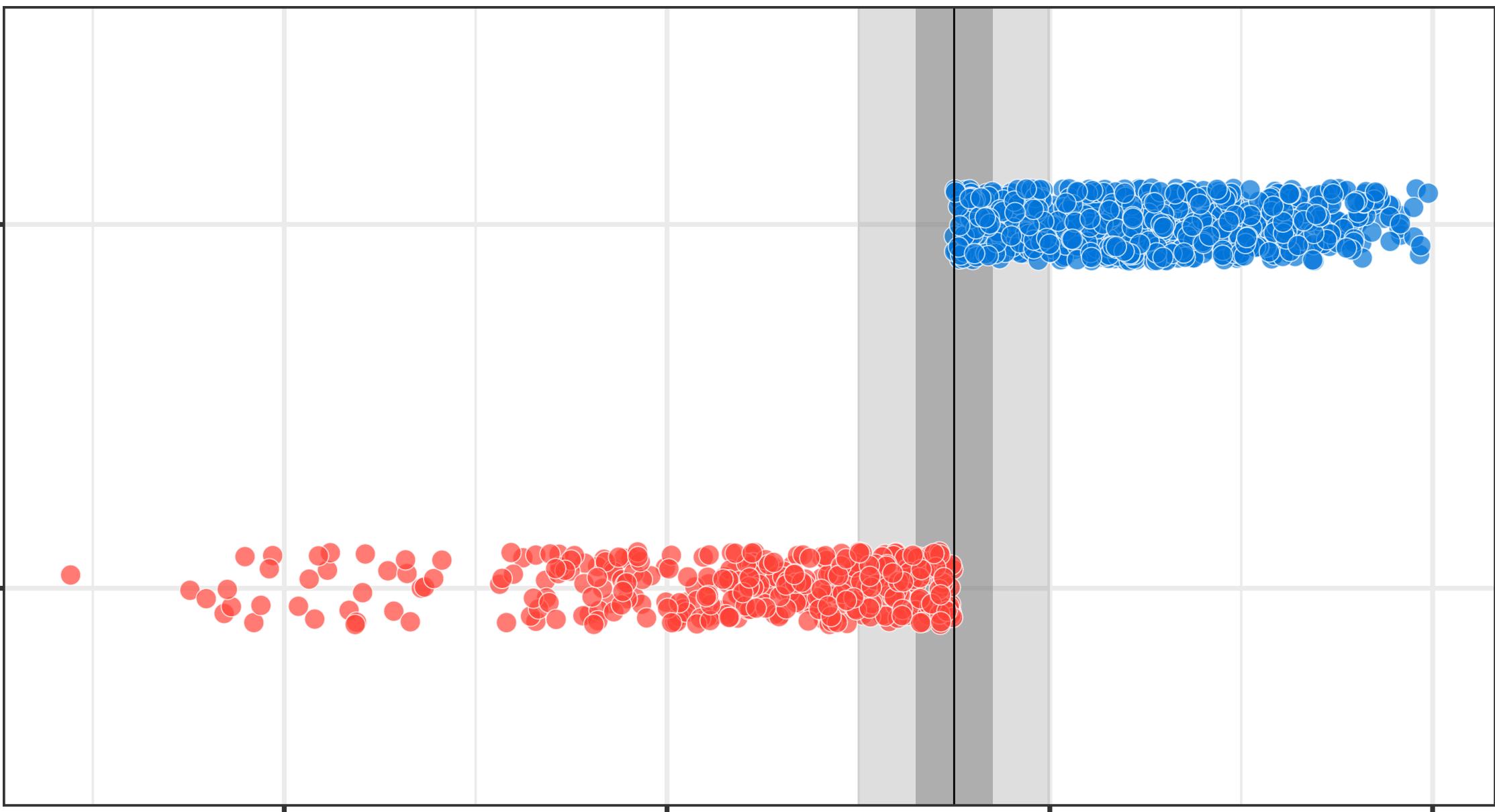
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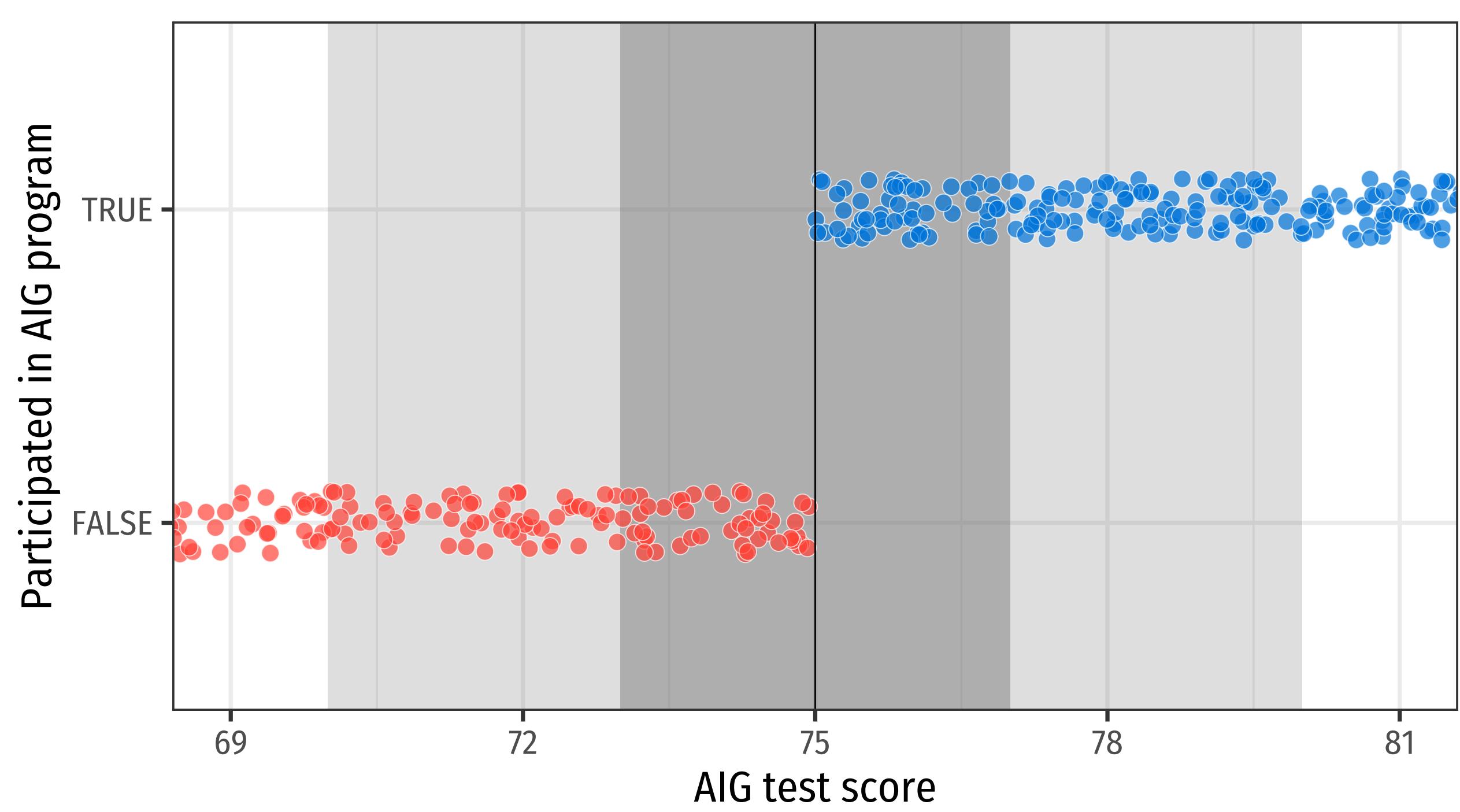
60

80

100

AIG test score



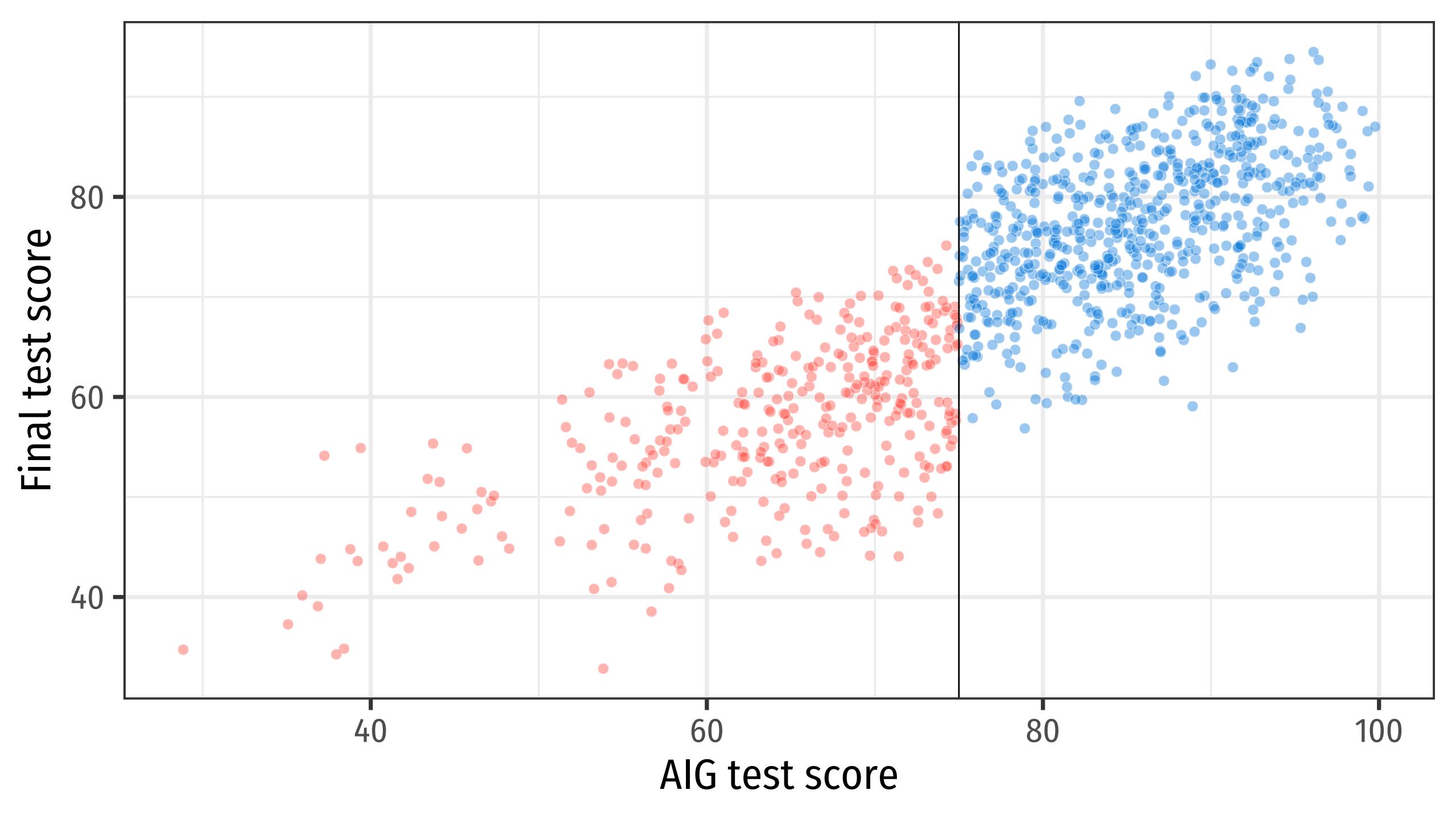


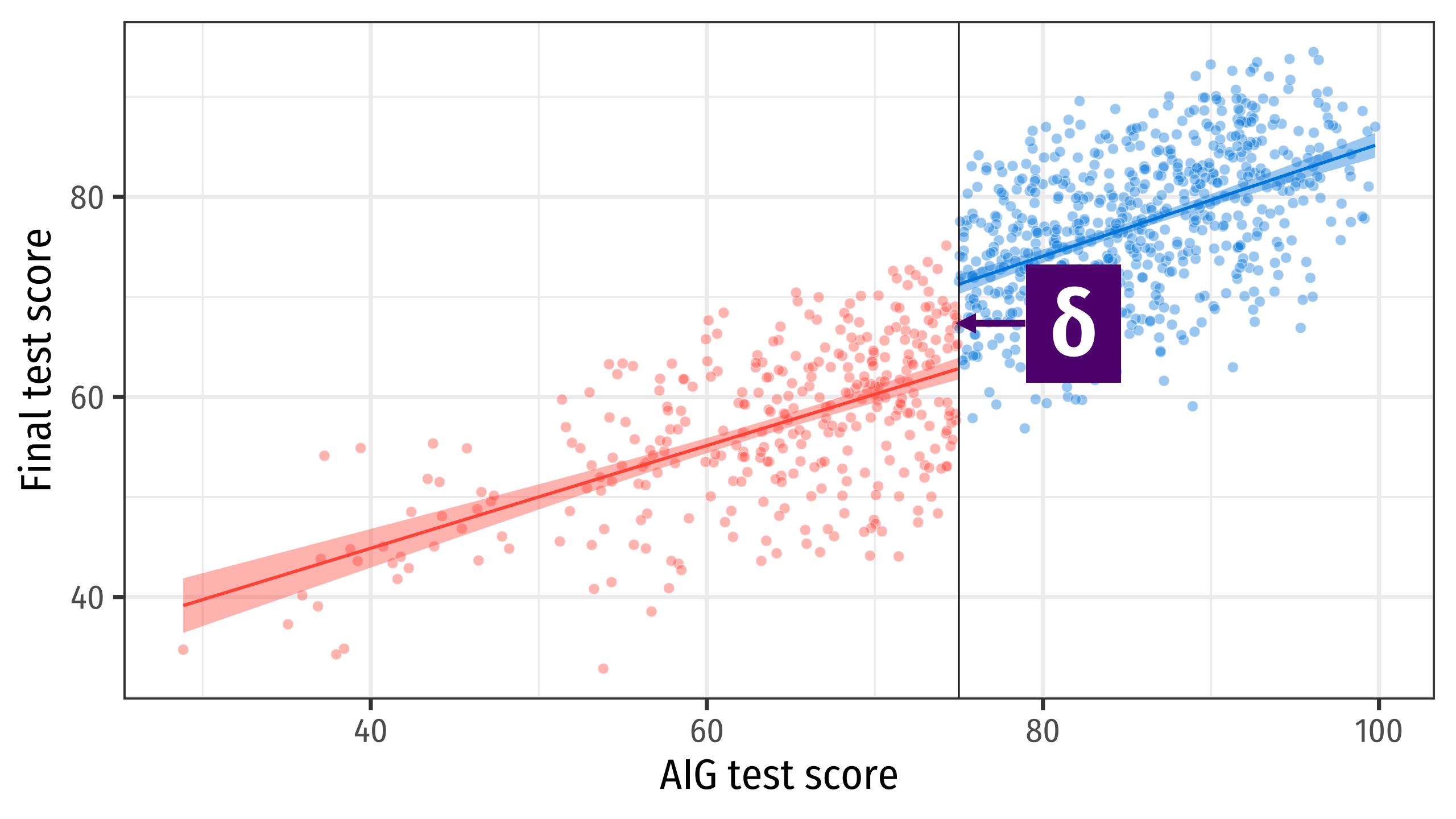
Causal inference intuition

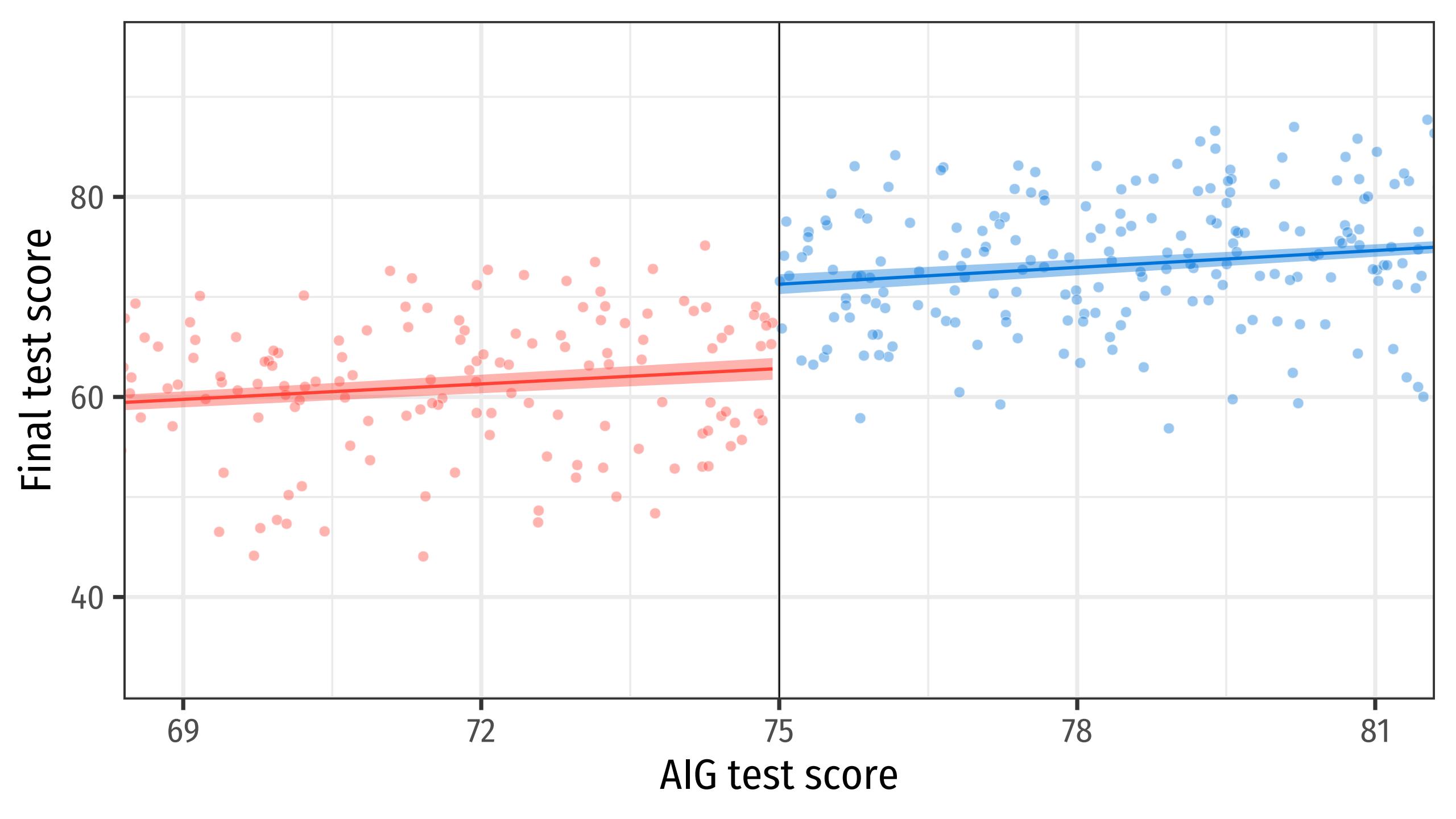
People right before and right after the threshold are essentially the same

Pseudo treatment and control groups!

Compare outcomes for those right before/after, calculate difference







Geographic discontinuities

When Time Is of the Essence: A Natural Experiment on How Time Constraints Influence Elections

Jerome Schafer, Ludwig Maximilian University of Munich
John B. Holbein, University of Virginia

Foundational theories of voter turnout suggest that time is a key input in the voting decision, but we possess little causal evidence about how this resource affects electoral behavior. In this article, we use over two decades of elections data and a novel geographic regression discontinuity design that leverages US time zone boundaries. Our results show that exogenous shifts in time allocations have significant political consequences. Namely, we find that citizens are less likely to vote if they live on the eastern side of a time zone border. Time zones also exacerbate participatory inequality and push election results toward Republicans. Exploring potential mechanisms, we find suggestive evidence that these effects are the consequence of insufficient sleep and moderated by the convenience of voting. Regardless of the exact mechanisms, our results indicate that local differences in daily schedules affect how difficult it is to vote and shape the composition of the electorate.

Although in recent years the administrative barriers to voting have declined in many democracies (Blais 2010), many eligible citizens still fail to vote. In the United States, about 40% of registered voters do not partic-

vote, many nonvoters report “not having enough time”—or a close derivative (e.g., “I’m too busy” or “[Voting] takes too long”; Pew Research Center 2006). Moreover, recent studies suggest that levels of turnout may be shaped by time costs such

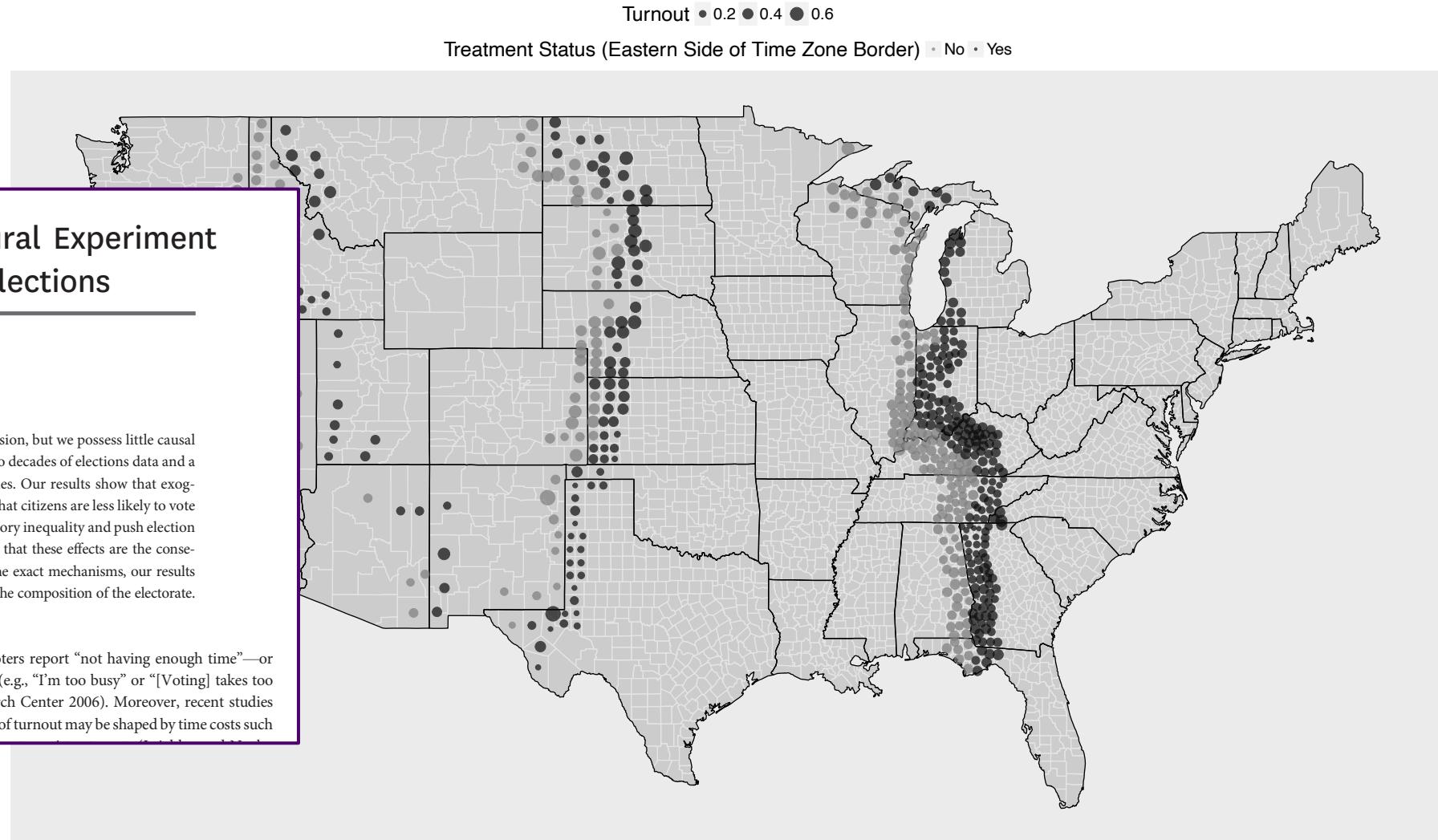
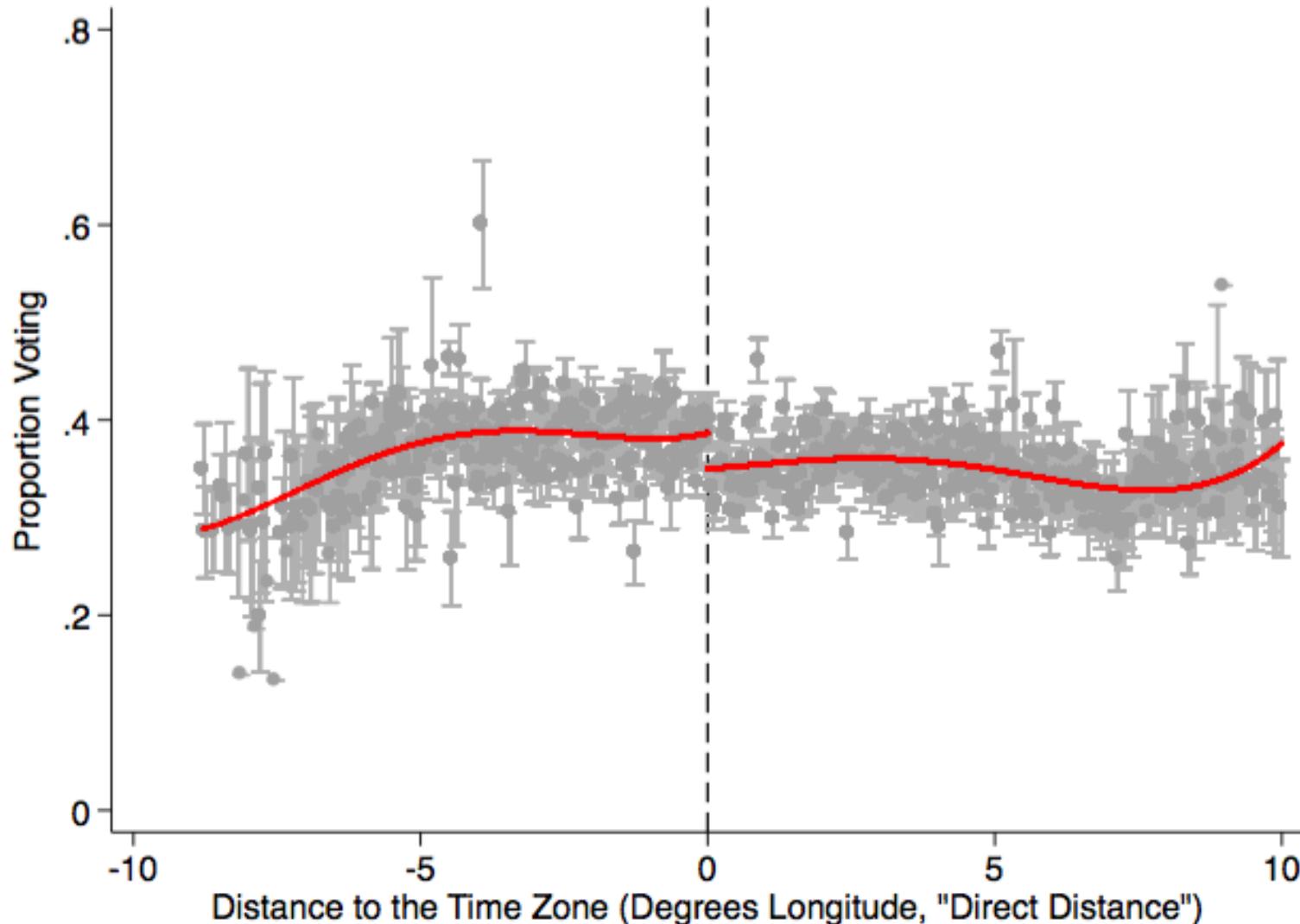


Figure 1 shows counties (with their geographic centroids marked) on either side of the time zones in the continental United States as of Election Day on 2010. The map shows counties within 1 degree (latitude and longitude) of the time zone boundaries.

Geographic discontinuities



**Lower turnout in
counties on the
eastern side of the
boundary**

**Election schedules
cause fluctuations
in turnout**

Time discontinuities

After Midnight: A Regression Discontinuity Design in Length of Postpartum Hospital Stays[†]

By DOUGLAS ALMOND AND JOSEPH J. DOYLE JR.*

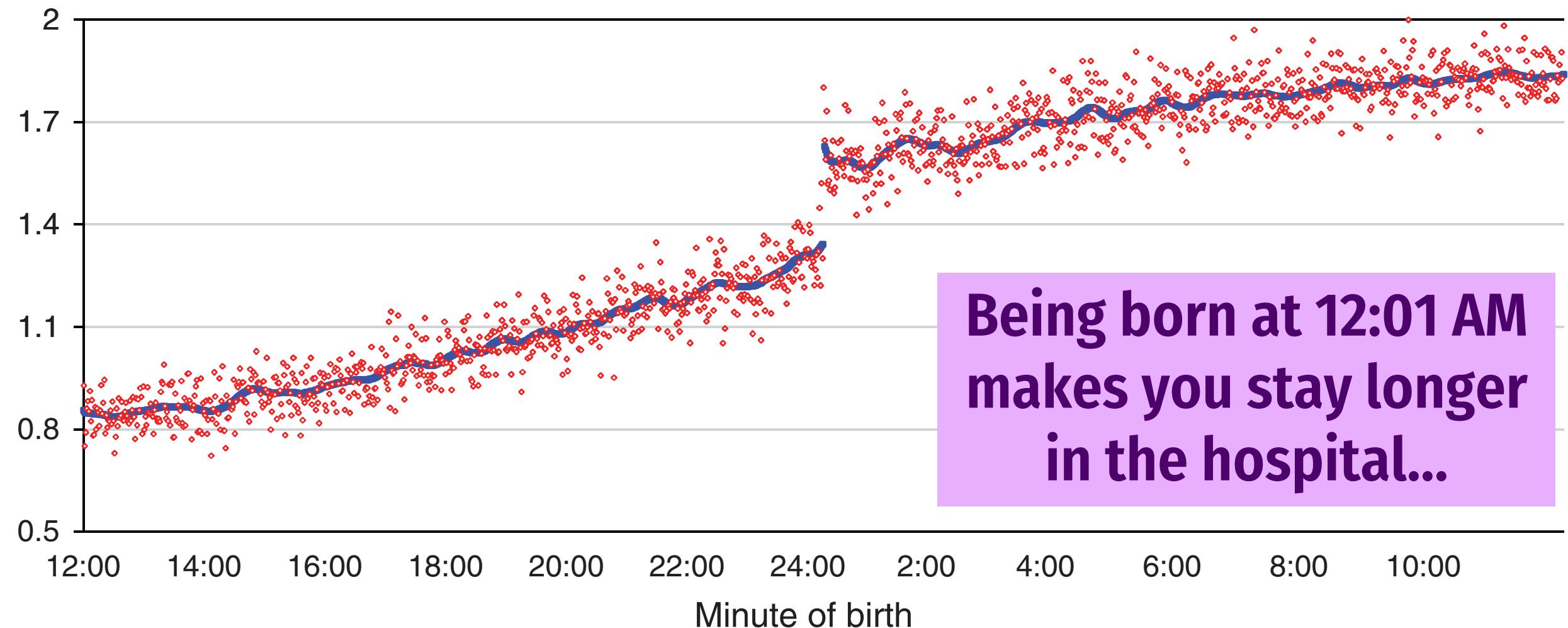
Estimates of moral hazard in health insurance markets can be confounded by adverse selection. This paper considers a plausibly exogenous source of variation in insurance coverage for childbirth in California. We find that additional health insurance coverage induces substantial extensions in length of hospital stay for mother and newborn. However, remaining in the hospital longer has no effect on readmissions or mortality, and the estimates are precise. Our results suggest that for uncomplicated births, minimum insurance mandates incur substantial costs without detectable health benefits. (JEL D82, G22, I12, I18, J13)

**California requires that
insurance cover two
days of post-partum
hospitalization**

**Does extra time in the
hospital improve
health outcomes?**

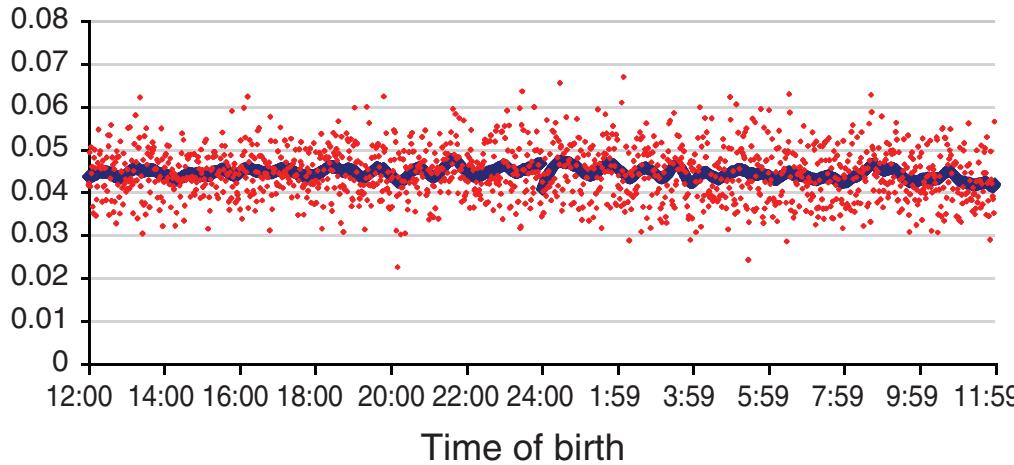
Time discontinuities

Panel B. Additional midnights: after law change

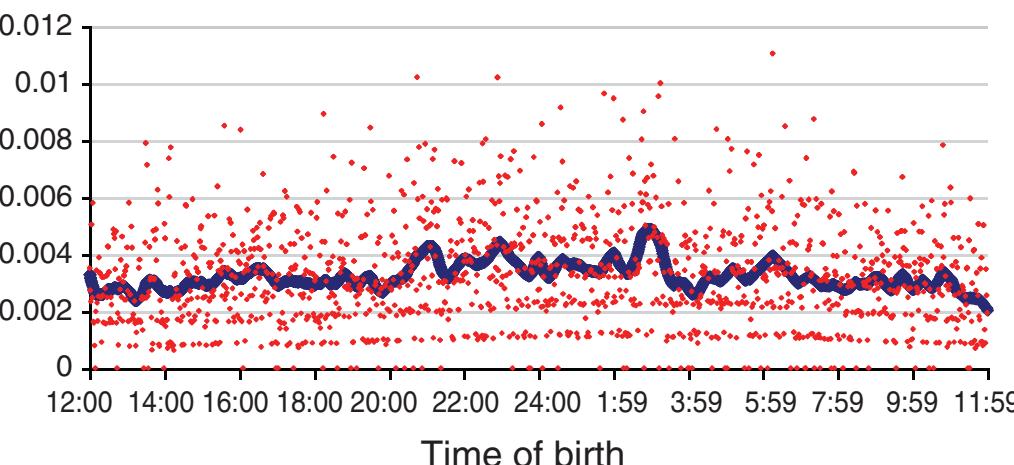


Time discontinuities

Panel B. Twenty-eight day readmission rate: after law change



Panel D. Twenty-eight day mortality rate: after law change



...but being born at
12:01 AM has no effect
on readmission rates
or mortality rates

Test score discontinuities

THE EFFECT OF ATTENDING THE FLAGSHIP STATE UNIVERSITY ON EARNINGS: A DISCONTINUITY-BASED APPROACH

Mark Hoekstra*

Abstract—This paper examines the effect of attending the flagship state university on the earnings of 28 to 33 year olds by combining confidential admissions records from a large state university with earnings data collected through the state's unemployment insurance program. To distinguish the effect of attending the flagship state university from the effects of confounding factors correlated with the university's admission decision or the applicant's enrollment decision, I exploit a large discontinuity in the probability of enrollment at the admission cutoff. The results indicate that attending the most selective state university causes earnings to be approximately 20% higher for white men.

I. Introduction

WHILE there has been considerable study of the effect of educational attainment on earnings, less is known regarding the economic returns to college quality. This paper examines the economic returns to college quality in the context of attending the most selective public state university. It does so using an intuitive regression discontinuity design that compares the earnings of 28 to 33 year olds who were barely admitted to the flagship to those of individuals who were barely rejected.

Convincingly estimating the economic returns to college quality requires overcoming the selection bias arising from the fact that attendance at more selective universities is likely correlated with unobserved characteristics that them-

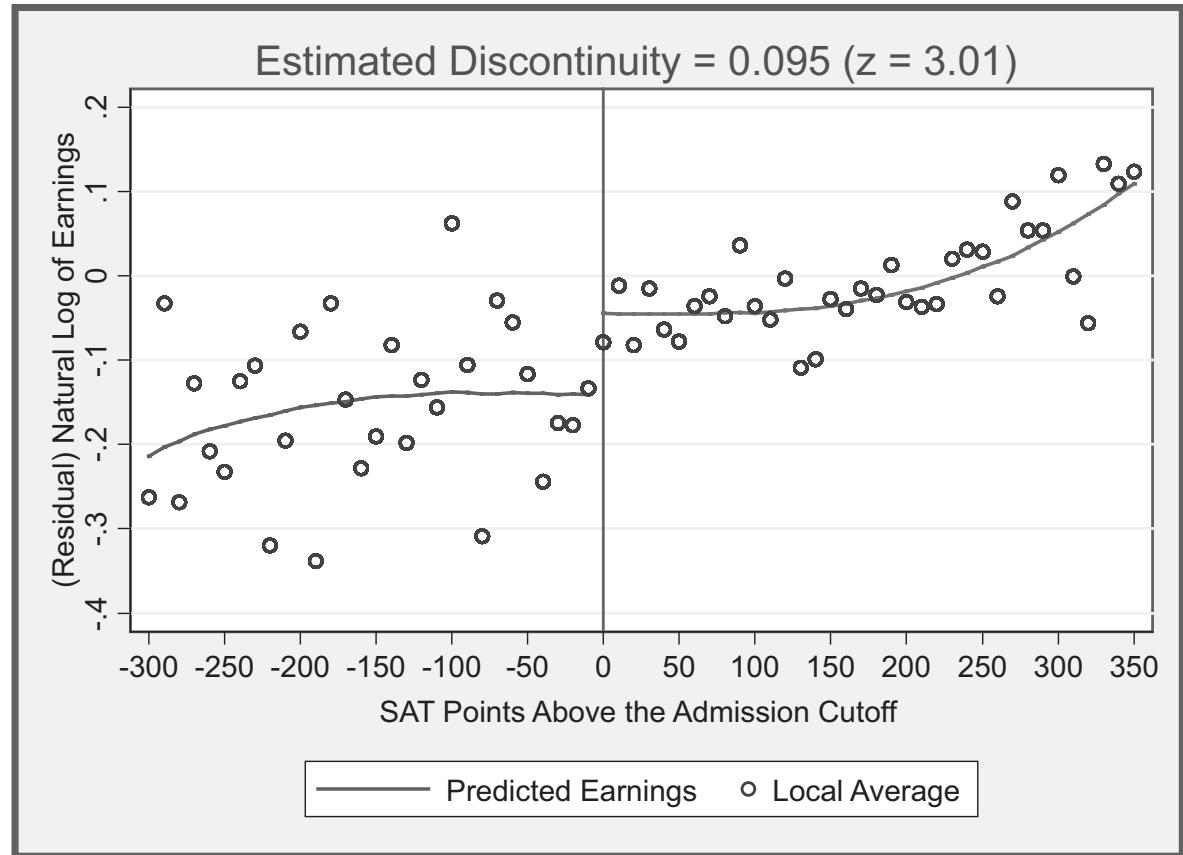
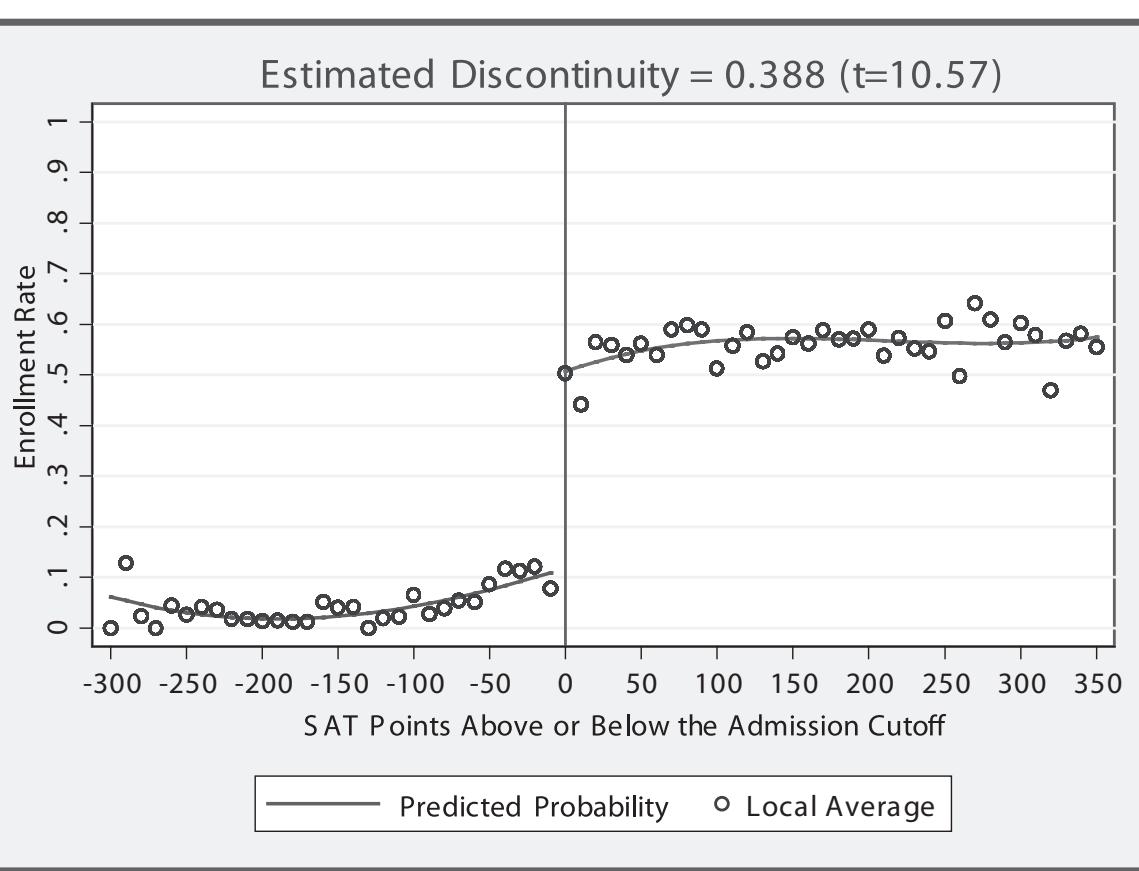
leges but chose to attend less selective institutions. They find that attending more selective colleges has a positive effect on earnings only for students from low-income families. Brewer, Eide, and Ehrenberg (1999) estimate the payoff by explicitly modeling high school students' choice of college type and find significant returns to attending an elite private institution for all students. Behrman, Rosenzweig, and Taubman (1996) identify the effect by comparing female twin pairs and find evidence of a positive payoff from attending Ph.D.-granting private universities with well-paid senior faculty. Using a similar approach, Lindahl and Regner (2005) use Swedish sibling data and show that cross-sectional estimates of the selective college wage premium are twice the within-family estimates.

This paper uses a different strategy in that it identifies the effect of school selectivity on earnings by comparing the earnings of those just below the cutoff for admission to the flagship state university to those of applicants who were barely above the cutoff for admission. To do so, I combined confidential administrative records from a large flagship state university with earnings records collected by the state through the unemployment insurance program. To put the selectivity of the flagship in context, the average SAT scores

Does going to the main state university (i.e. UGA) make you earn more money?

SAT scores are an arbitrary cutoff for accessing the university

Test score discontinuities



Cutoff seems rule-based

Earnings slightly higher

RDDs are all the rage

People love these things!

They're intuitive, compelling, and highly graphical

ABSTRACT

Methods Matter: P-Hacking and Causal Inference in Economics*

The economics 'credibility revolution' has promoted the identification of causal relationships using difference-in-differences (DID), instrumental variables (IV), randomized control trials (RCT) and regression discontinuity design (RDD) methods. The extent to which a reader should trust claims about the statistical significance of results proves very sensitive to method. Applying multiple methods to 13,440 hypothesis tests reported in 25 top economics journals in 2015, we show that selective publication and p-hacking is a substantial problem in research employing DID and (in particular) IV. RCT and RDD are much less problematic. Almost 25% of claims of marginally significant results in IV papers are misleading.

JEL Classification: A11, B41, C13, C44

Keywords: research methods, causal inference, p-curves, p-hacking, publication bias

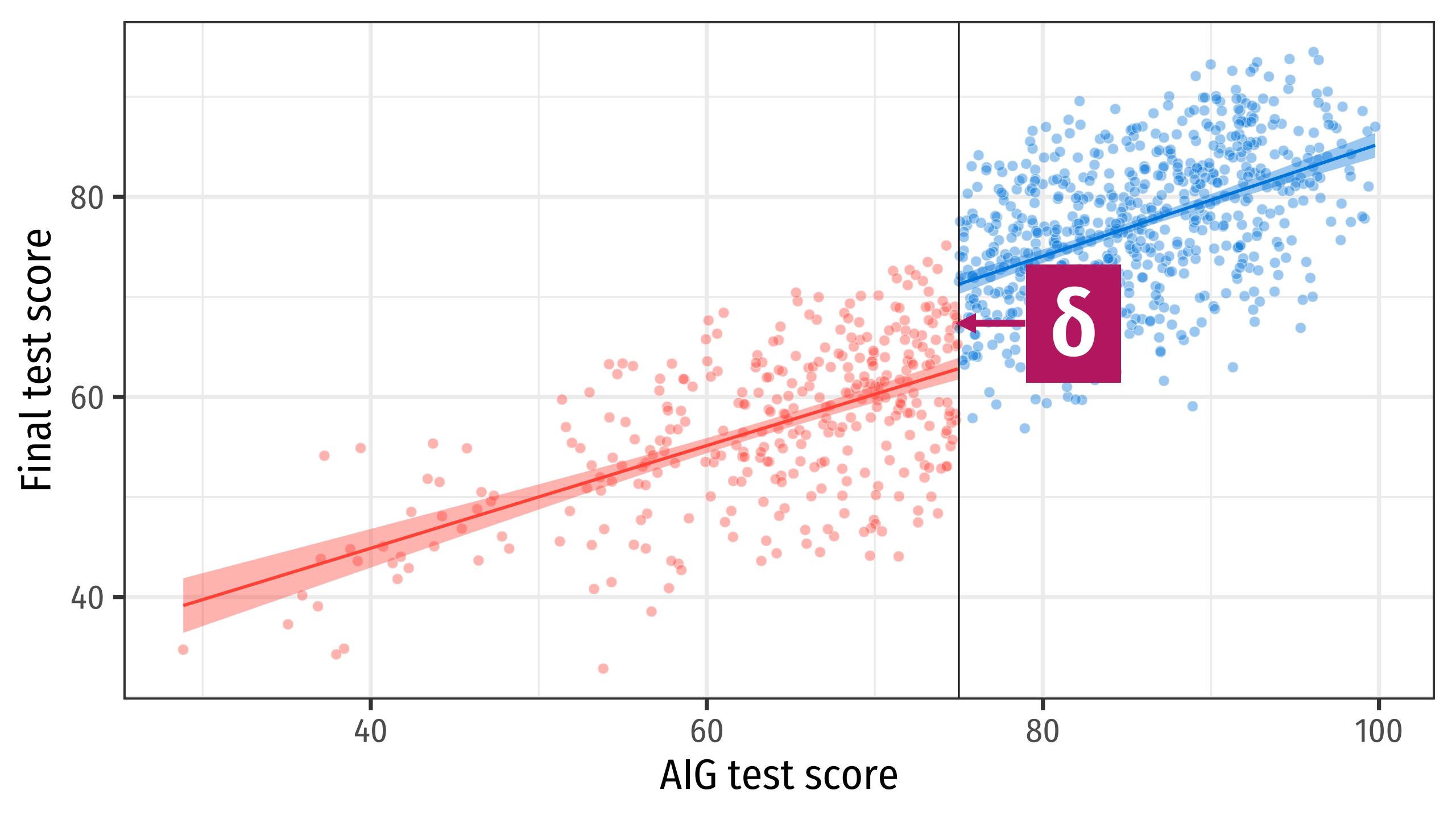
Less susceptible to p-hacking and selective publication than DID or IV

Drawing lines &
measuring gaps

Main goal of RD

**Measure the gap in the outcome for
people on both sides of the cutpoint**

**Gap = δ =
local average treatment effect
(LATE)**



Drawing lines

The size of the gap depends on how you draw the lines on each side of the cutoff

The type of lines you choose can change the estimate of δ —sometimes by a lot!

There's no one right way to draw lines!

Line-drawing considerations

Parametric vs. nonparametric lines

Measuring the gap

Bandwidths

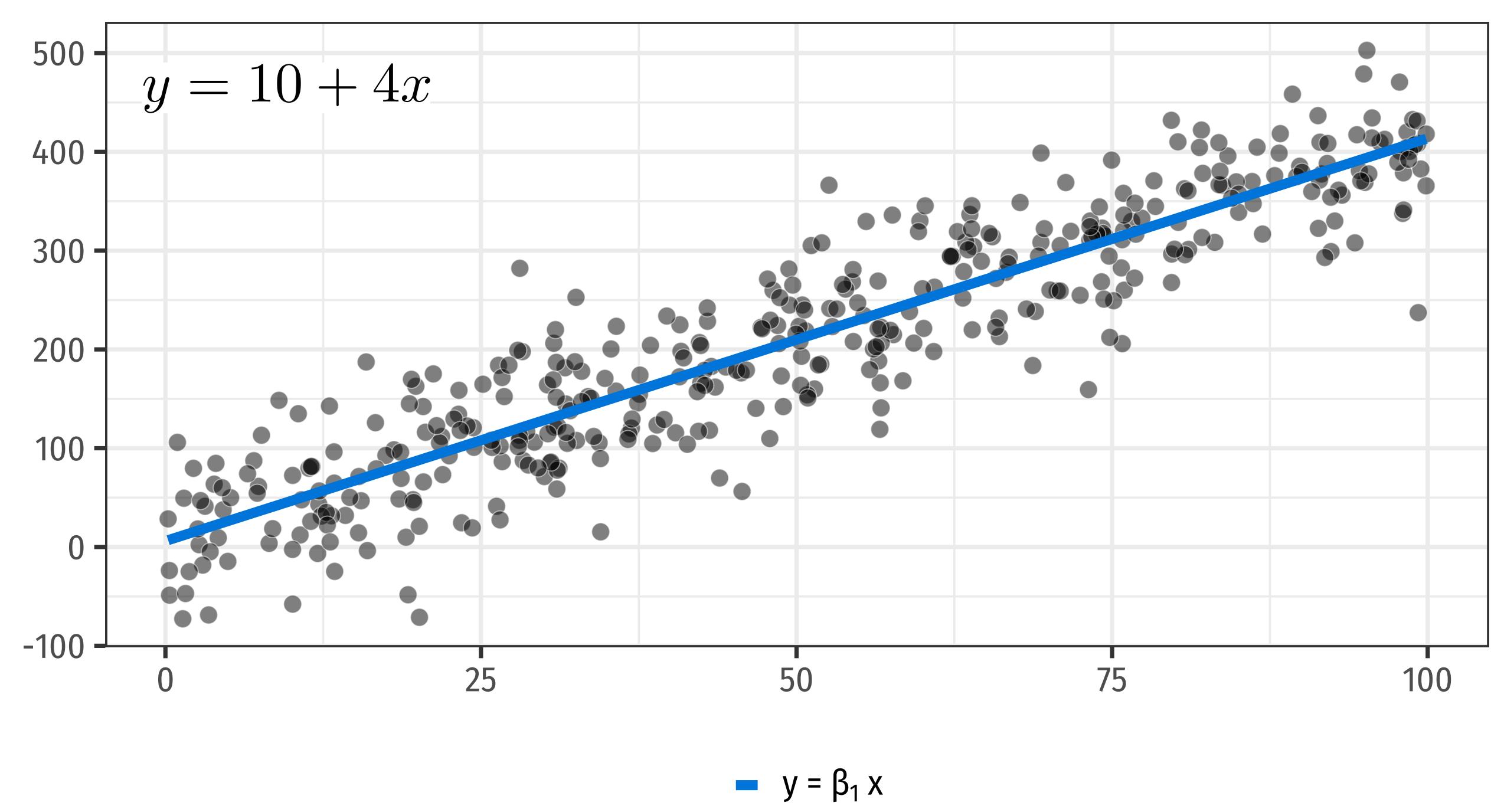
Kernels

Parametric lines

Formulas with *parameters*

$$y = mx + b$$

$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$



Parametric lines

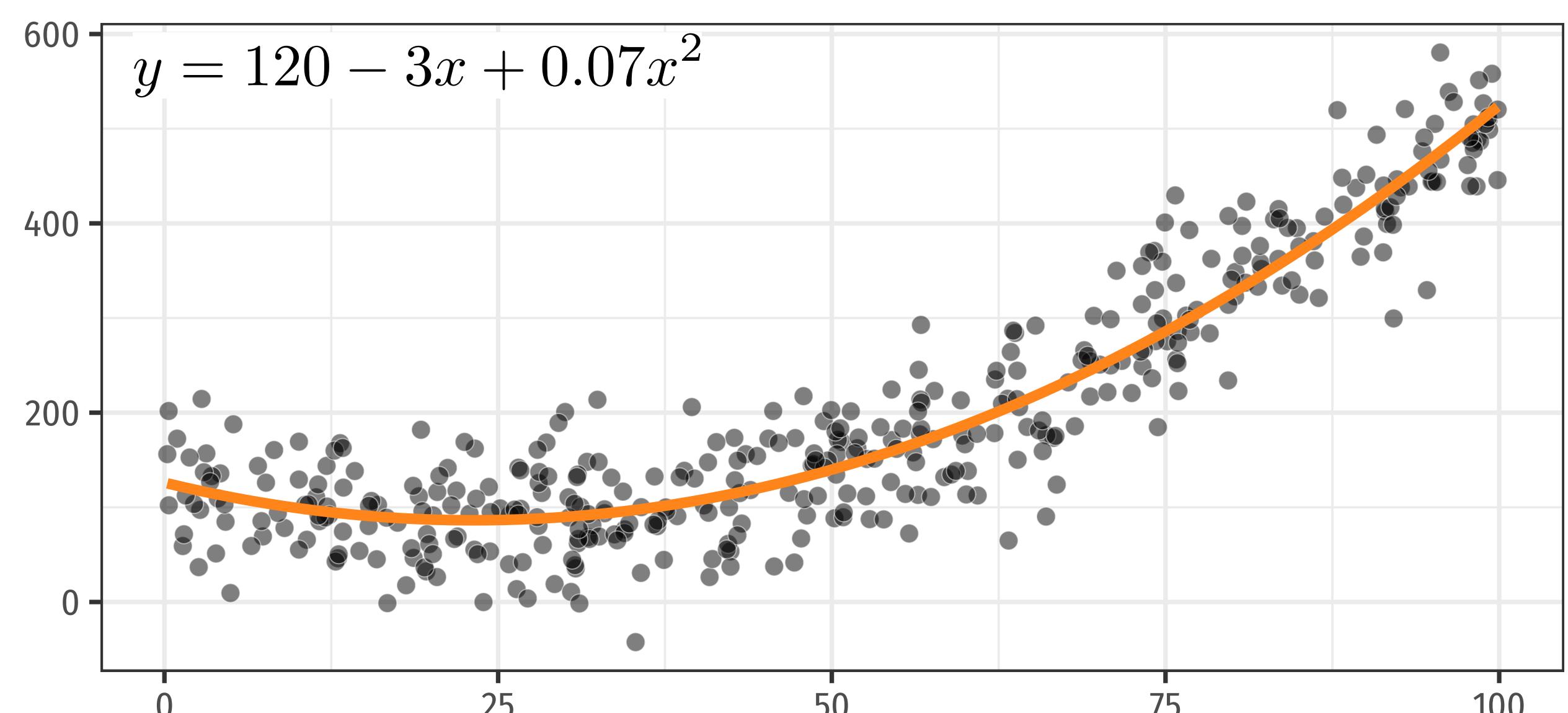
Not just for straight lines!

Make curvy with exponents or trigonometry

$$y = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3$$

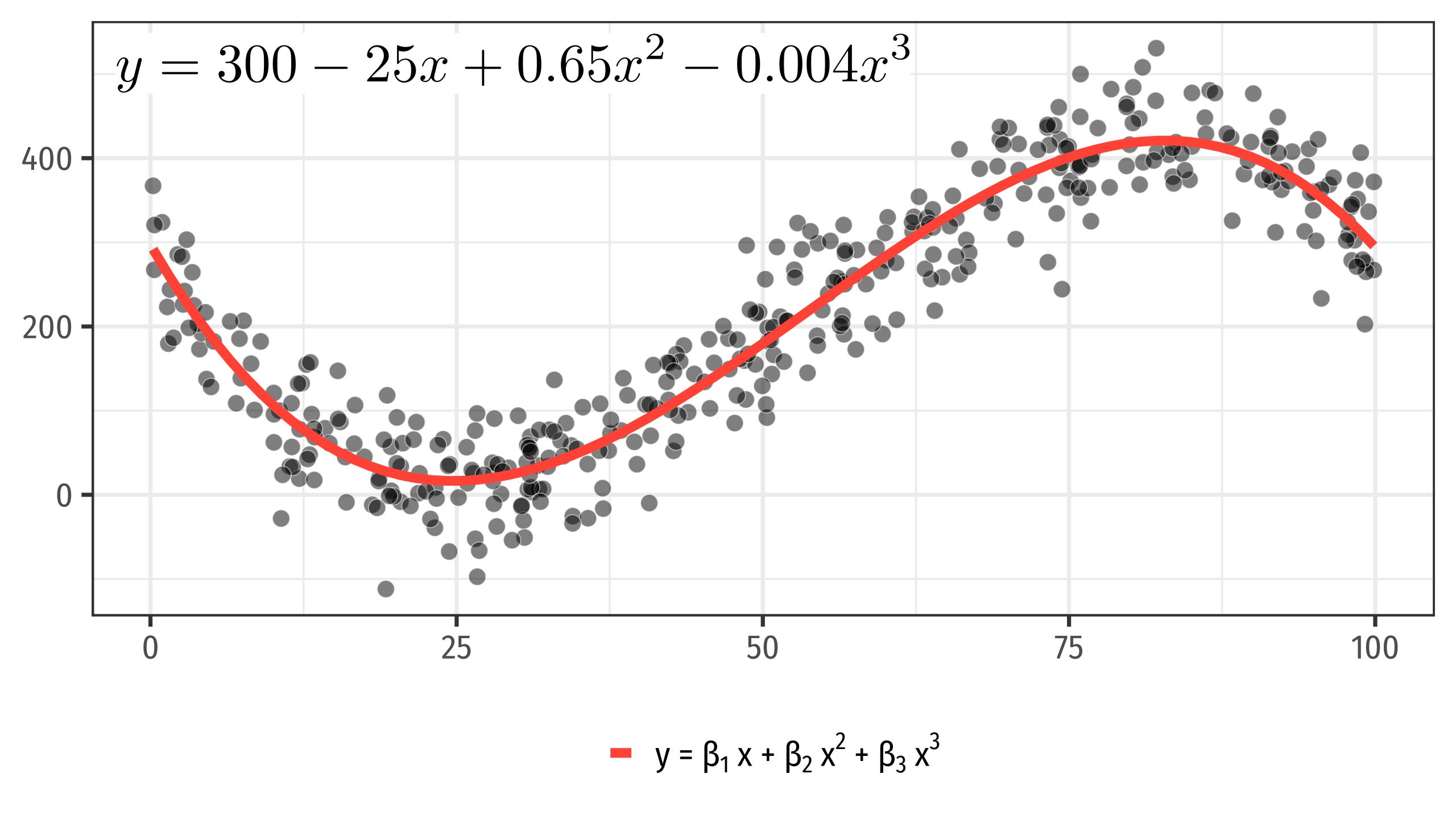
$$y = \beta_0 + \beta_1 x + \beta_2 \sin(x)$$

$$y = 120 - 3x + 0.07x^2$$

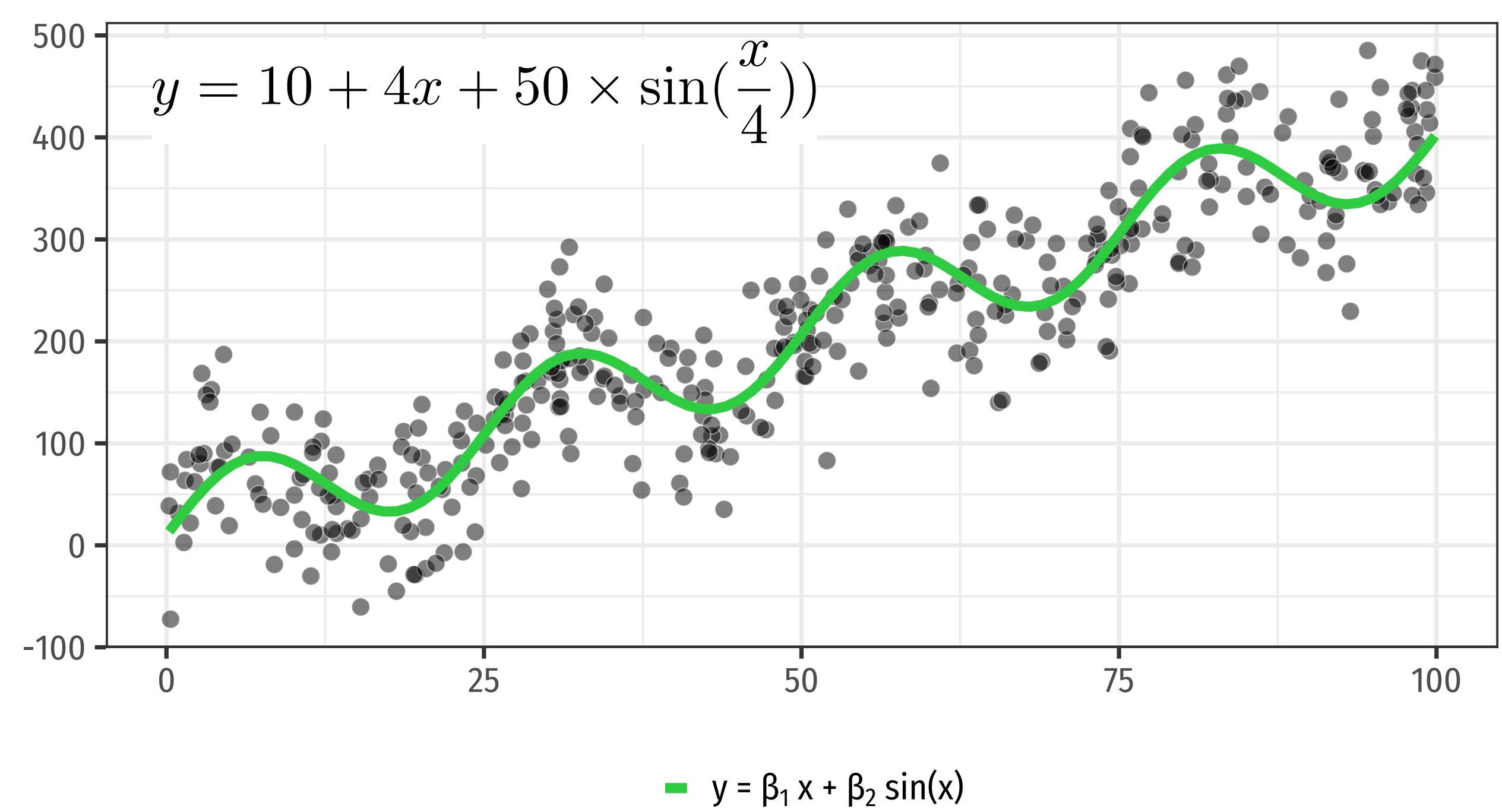


— $y = \beta_1 x + \beta_2 x^2$

$$y = 300 - 25x + 0.65x^2 - 0.004x^3$$



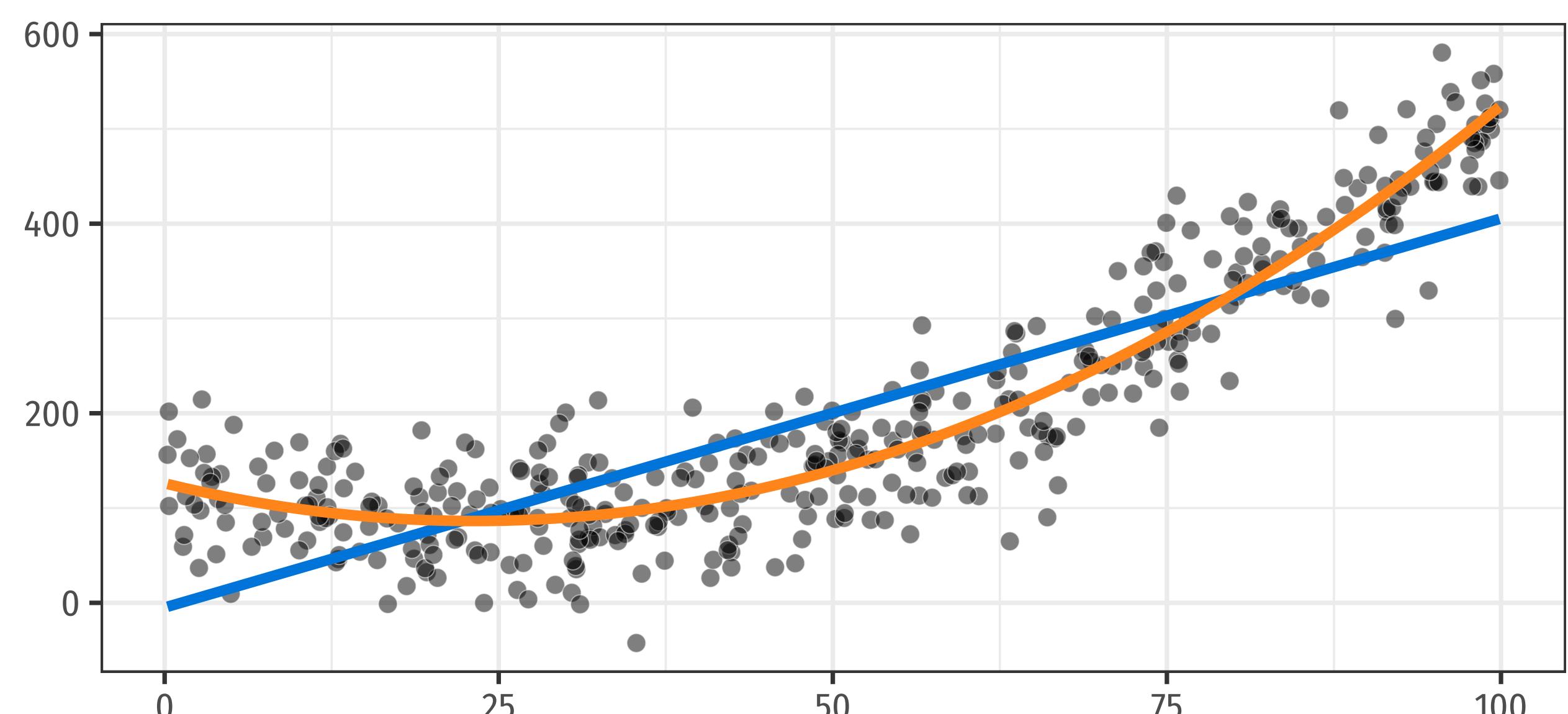
$$y = 10 + 4x + 50 \times \sin\left(\frac{x}{4}\right)$$



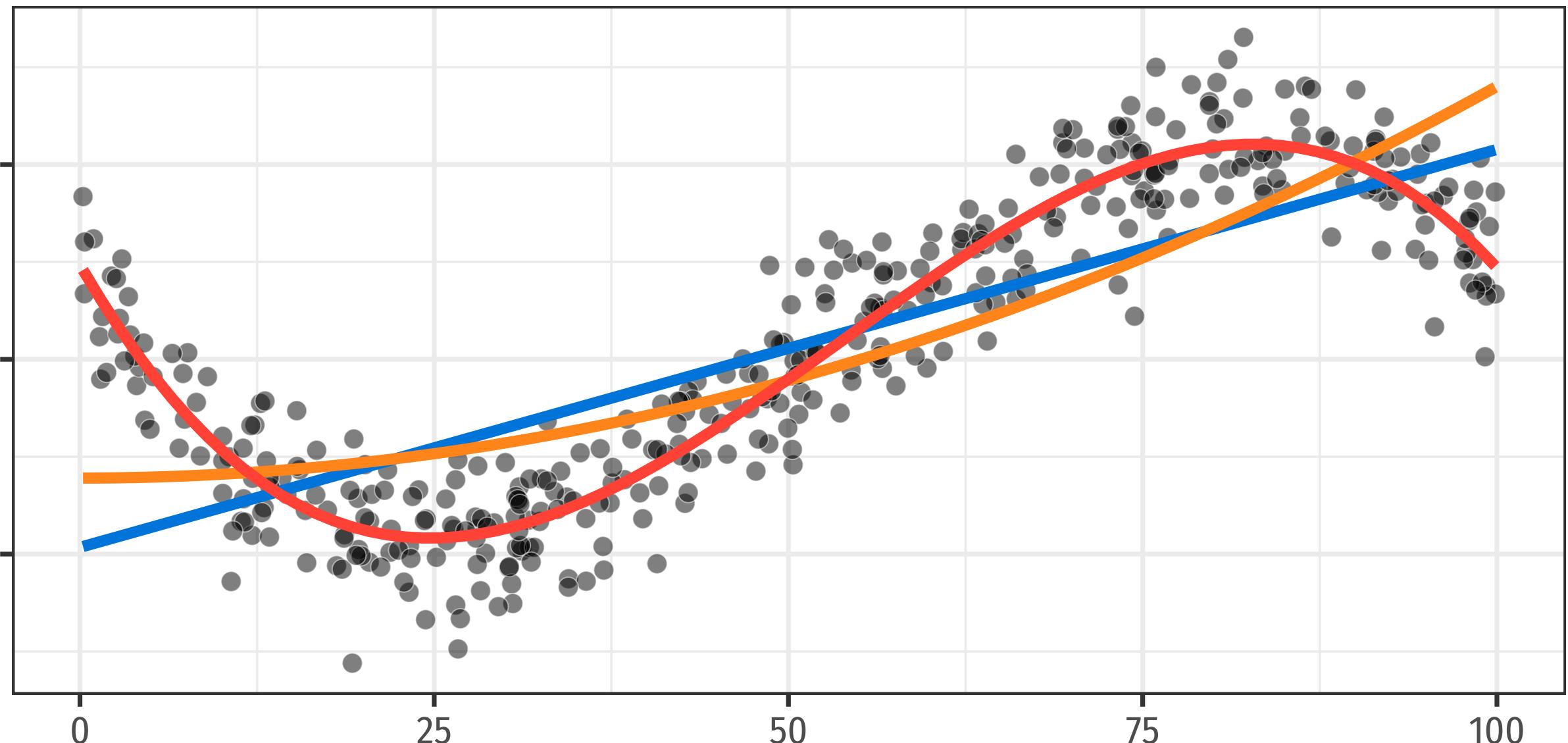
Parametric lines

It's important to get the parameters right!

Line should fit the data pretty well



— $y = \beta_1 x$ — $y = \beta_1 x + \beta_2 x^2$



— $y = \beta_1 x$ — $y = \beta_1 x + \beta_2 x^2$ — $y = \beta_1 x + \beta_2 x^2 + \beta_3 x^3$

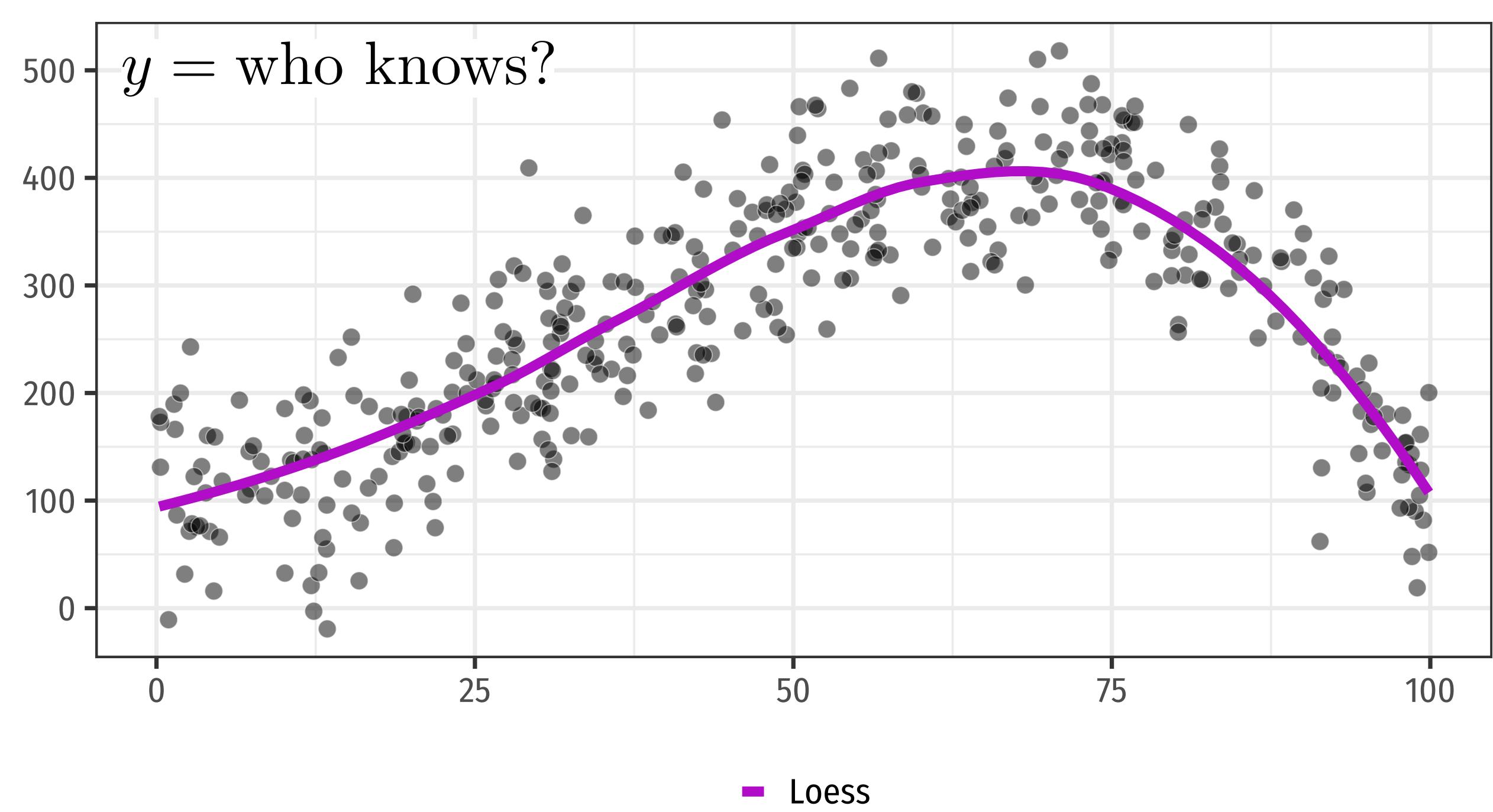
Nonparametric lines

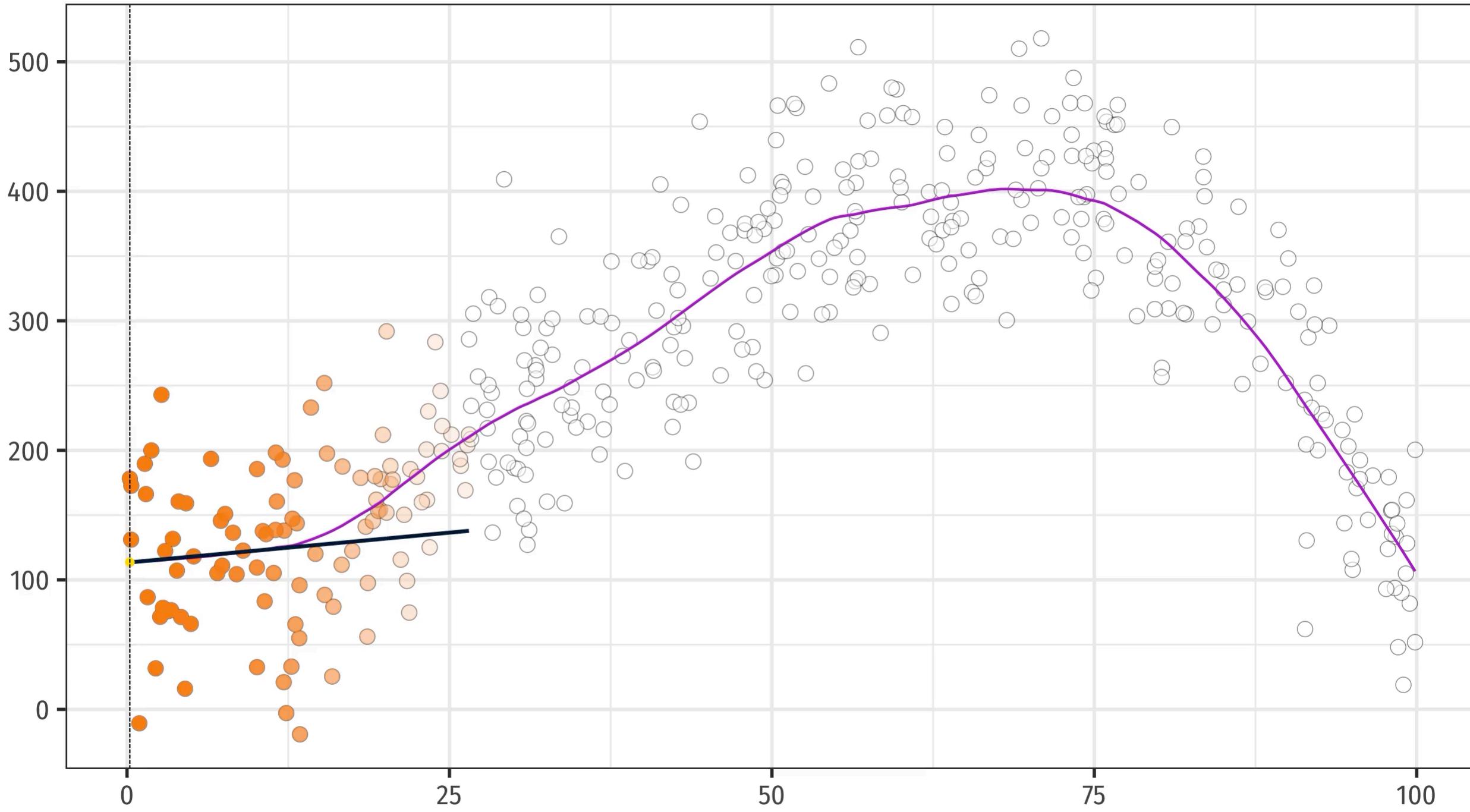
Lines without parameters

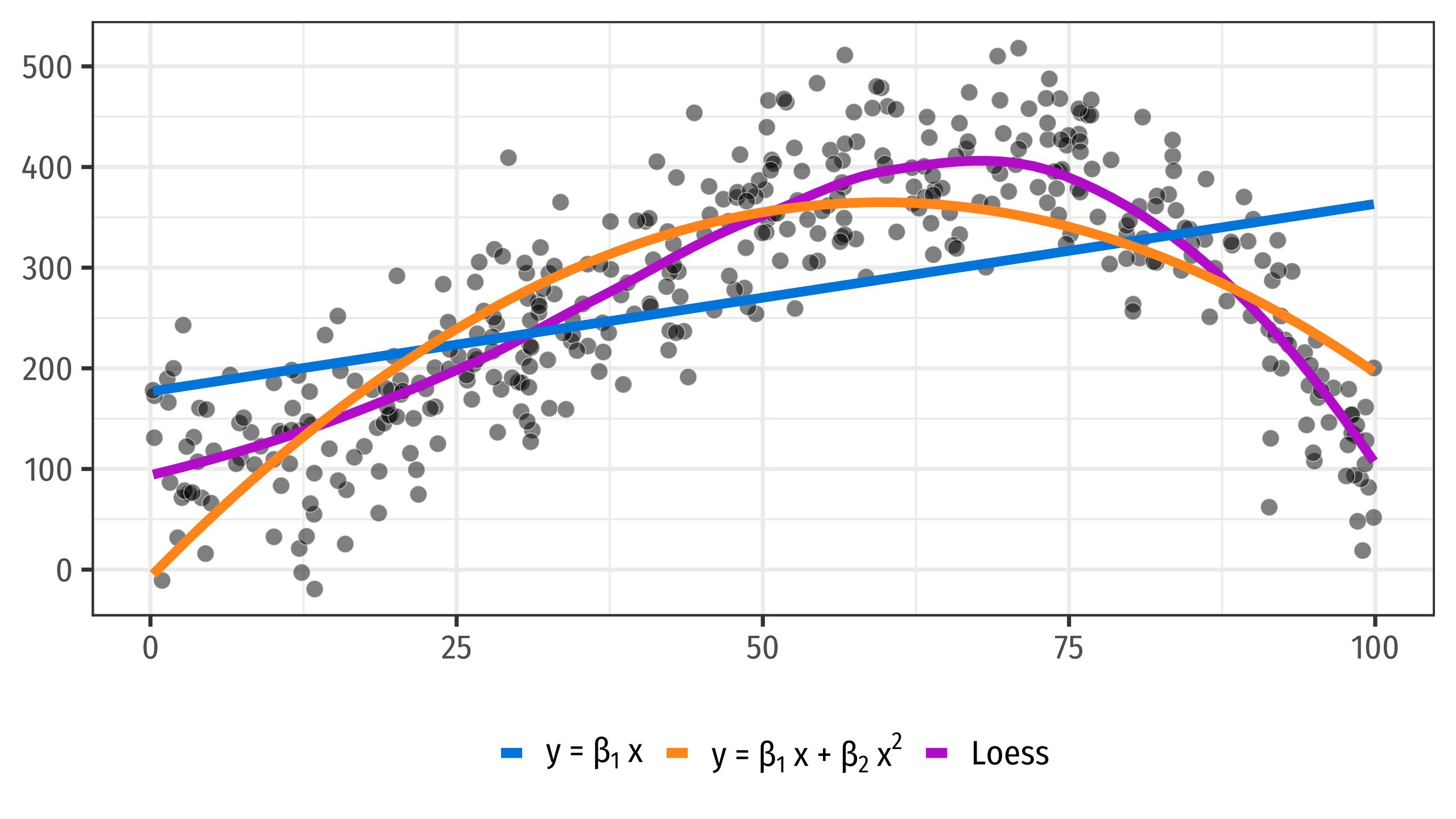
**Use the data to find the best line,
often with windows and moving averages**

**Locally estimated/weighted scatterplot smoothing
(LOESS/LOWESS)
is a common method**

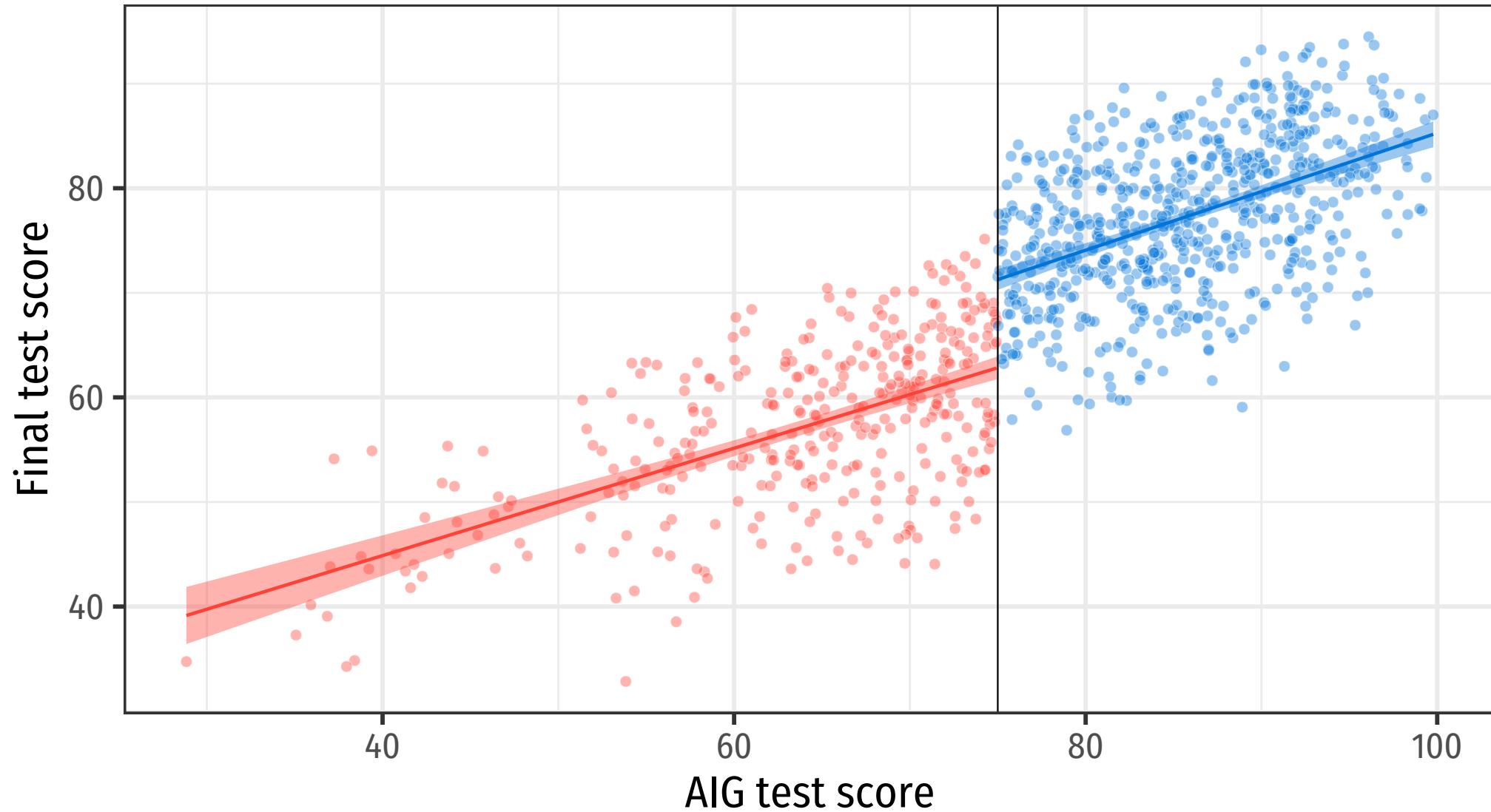
$y = \text{who knows?}$







Measuring gap with parametric lines



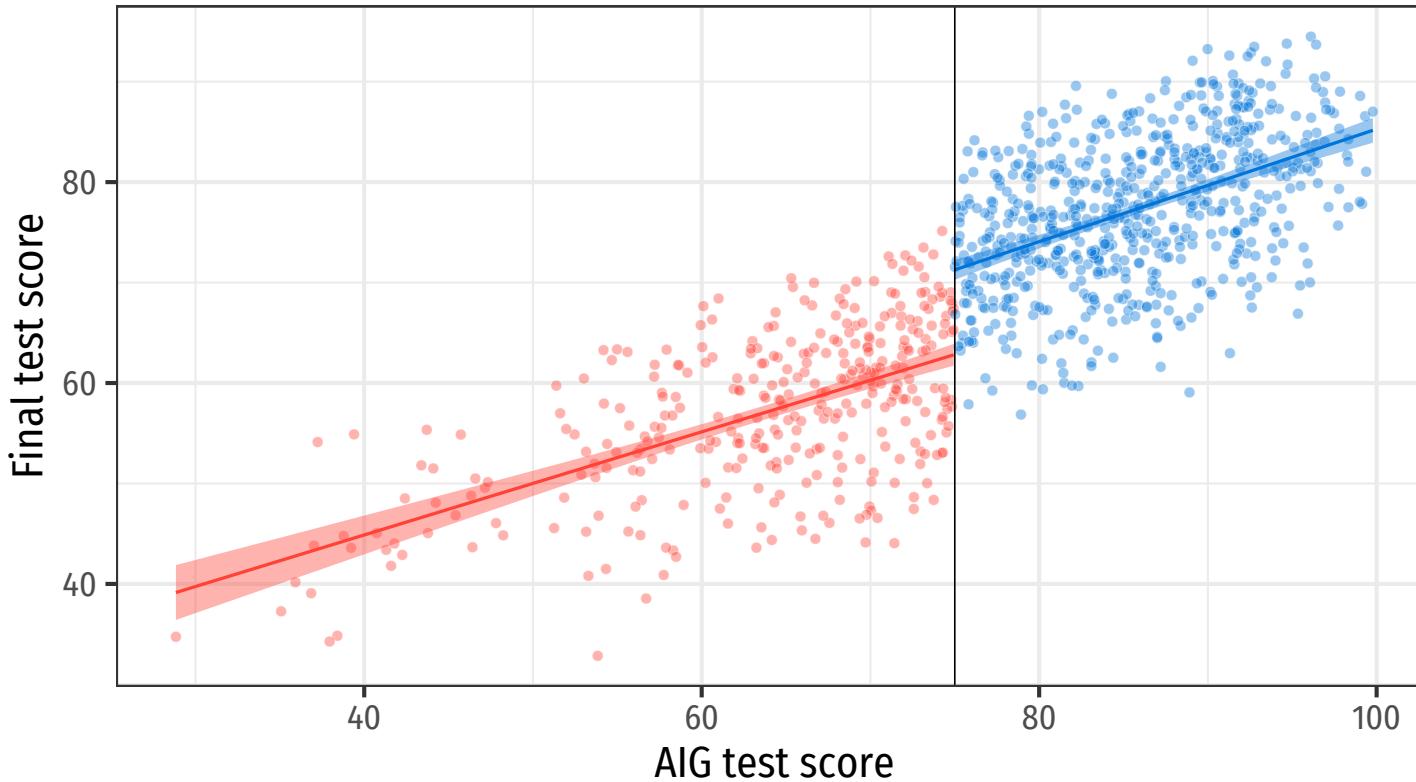
Measuring gap with parametric lines

Easiest way: center the running variable

ID	outcome	running_var	running_var_centered	treatment
1	90.0	69	-6	FALSE
2	85.7	75	0	TRUE
3	85.8	78	3	TRUE
4	85.7	65	-10	FALSE
5	84.4	76	1	TRUE

$$y = \beta_0 + \beta_1 \text{Running variable (centered)} + \beta_2 \text{Indicator for treatment}$$

Measuring gap with parametric lines



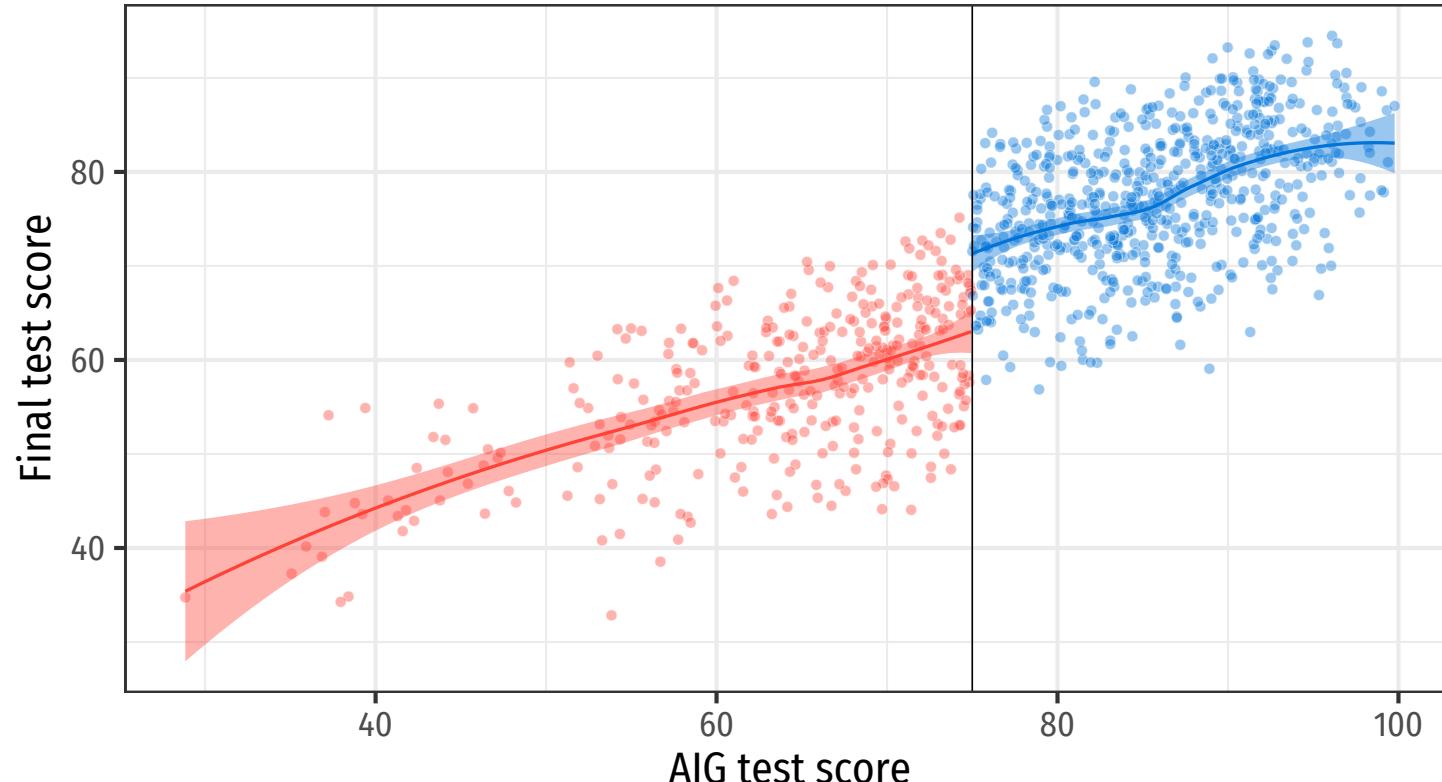
```
program_data <- aig_program %>%  
  mutate(aig_score_centered = test_score - 75)  
  
model1 <- lm(final_score ~ aig_score_centered + aig,  
              data = program_data)
```

```
> tidy(model1)  
# A tibble: 3 x 5  
  term                estimate std.error statistic p.value  
  <chr>                 <dbl>     <dbl>      <dbl>    <dbl>  
  1 (Intercept)            63.1      0.466     135.    0.  
  2 aig_score_centered      0.535     0.0278    19.2  2.09e-70  
  3 aigTRUE                  8.47      0.743     11.4  2.27e-28
```

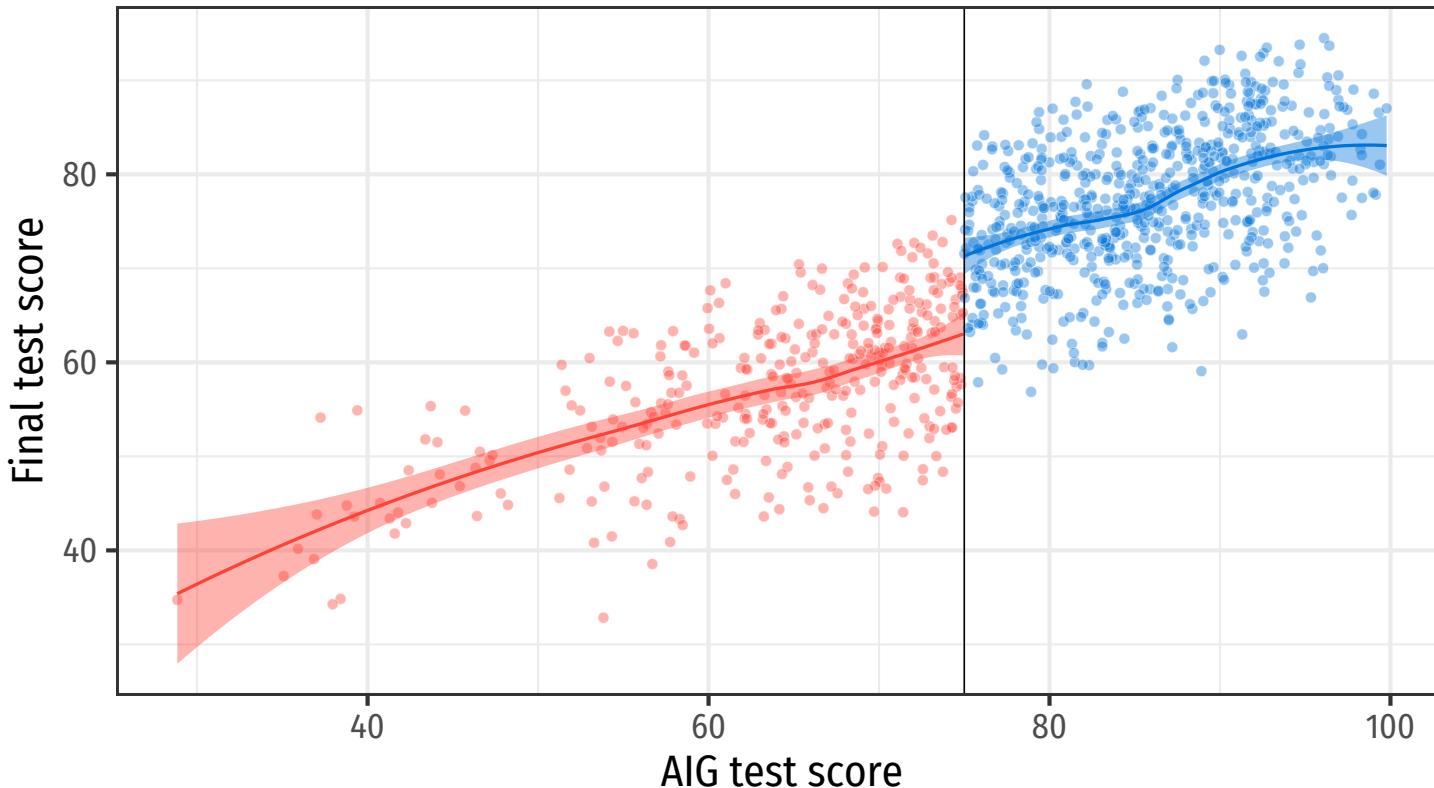


Measuring gap with nonparametric lines

Can't use regression; use `rdrobust` R package



Measuring gap with nonparametric lines



```
rdrobust(y = aig_program$final_score,  
         x = aig_program$test_score,  
         c = 75) %>%  
  summary()
```

Method	Coef.	Std. Err.	z	P> z	[95% C.I.]
Conventional	8.011	1.494	5.360	0.000	[5.082 , 10.940]
Robust	-	-	4.437	0.000	[4.315 , 11.144]

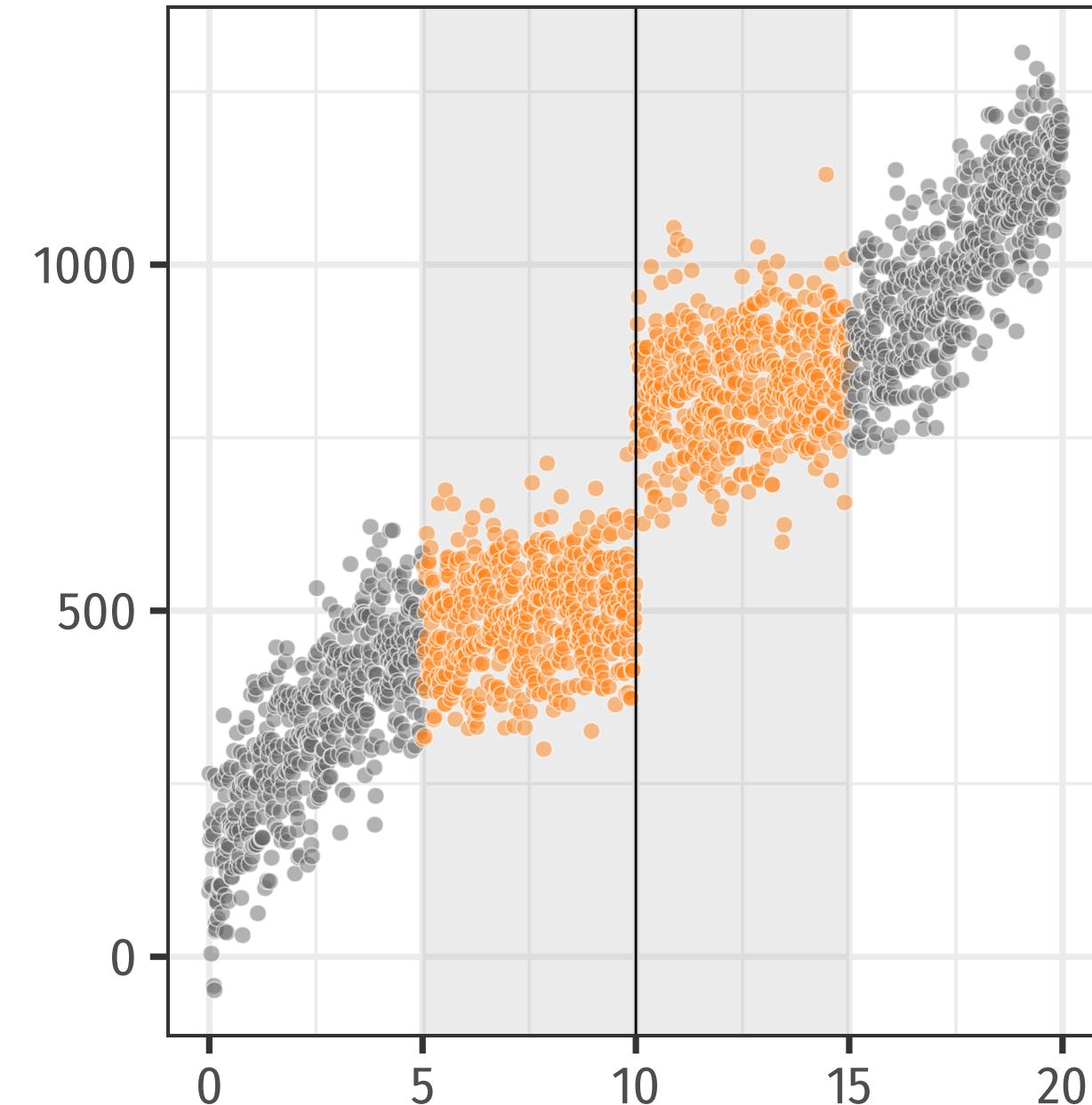
Bandwidths

All you really care about is the area right around the cutoff

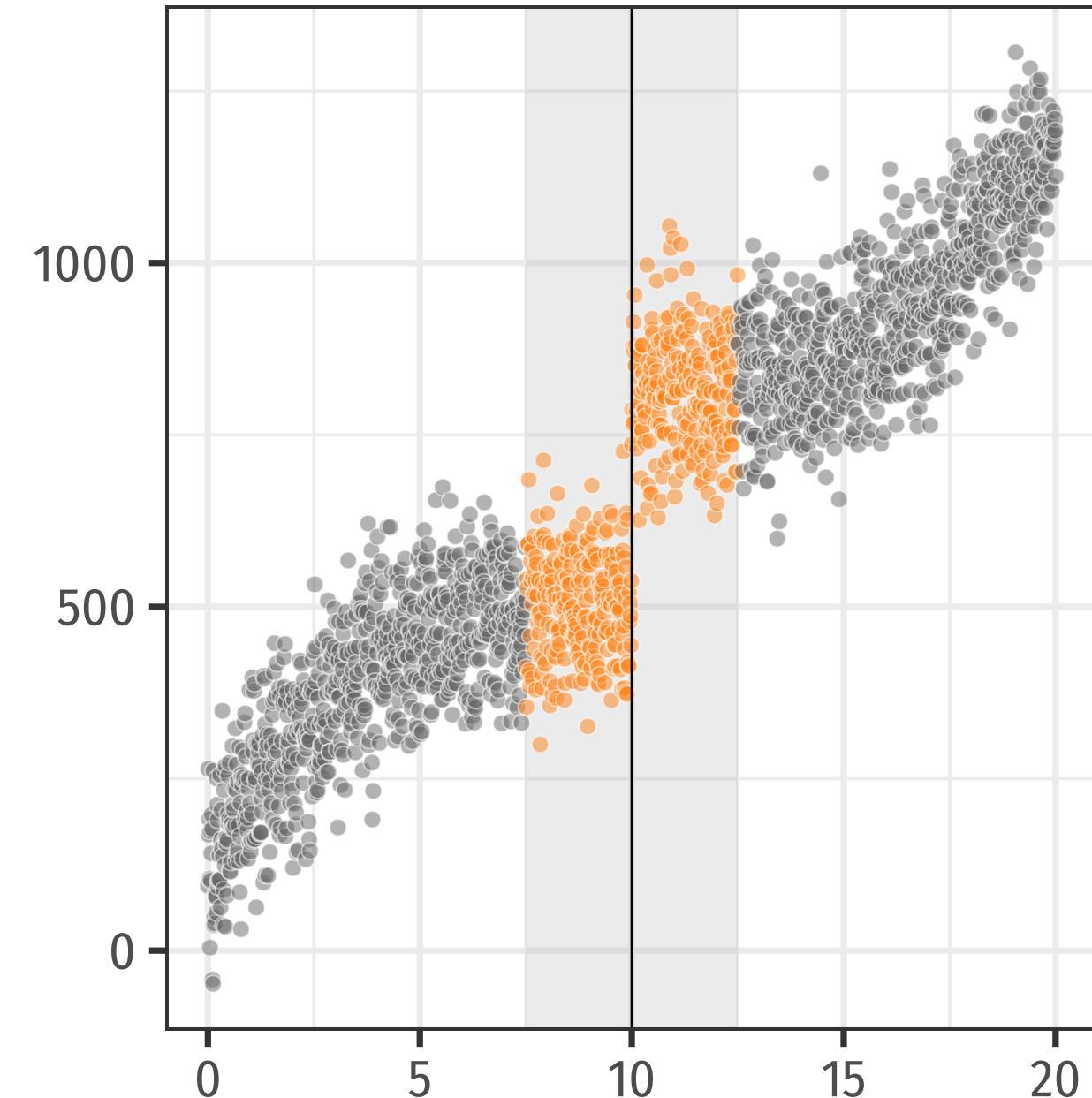
Observations far away don't matter because they're not comparable

Bandwidth = window around cutoff

Bandwidth = 5



Bandwidth = 2.5



Bandwidths

Algorithms exist to choose optimal width

Also use common sense

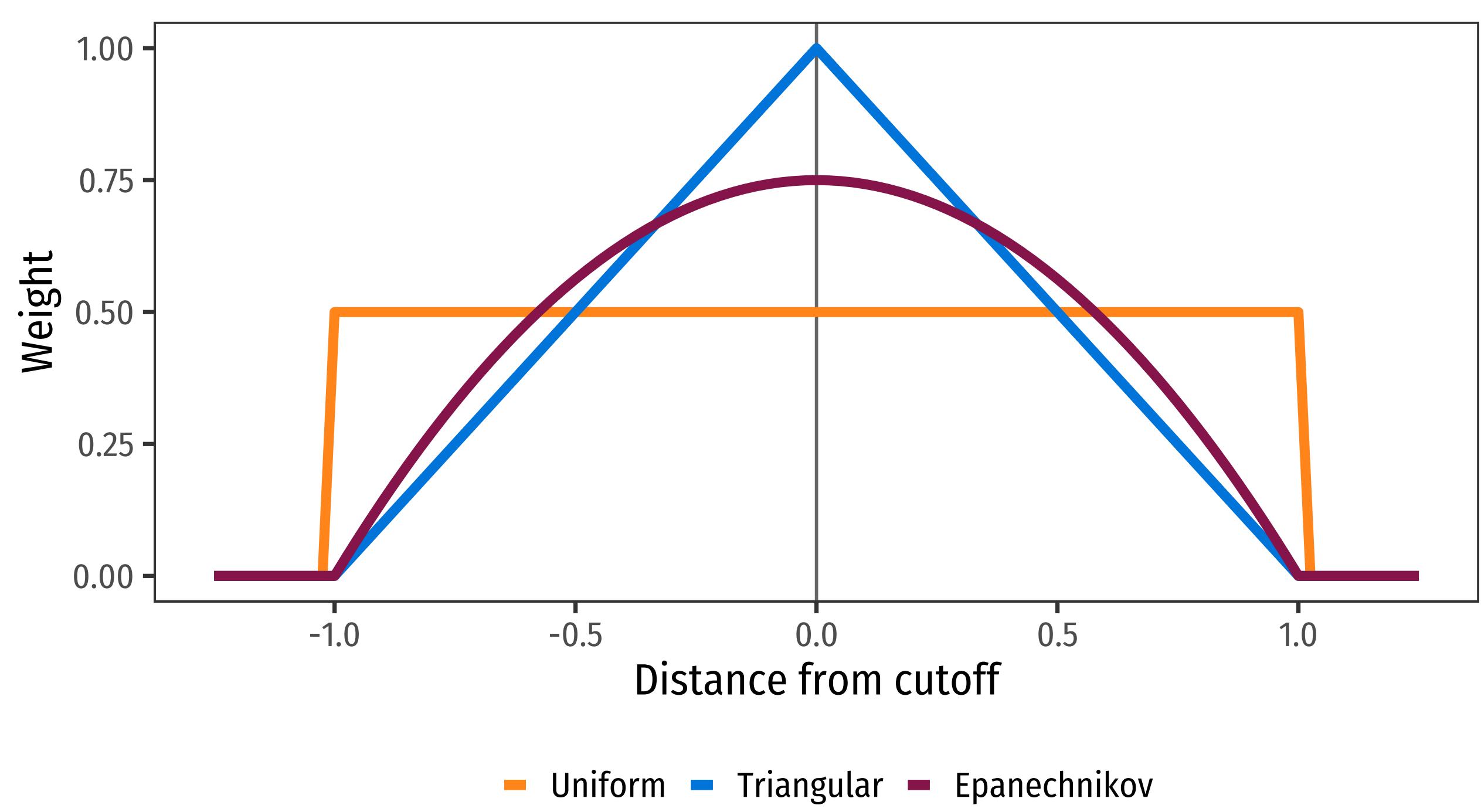
Maybe ± 5 for the AIG test?

For robustness, check what happens if
you double and halve the bandwidth

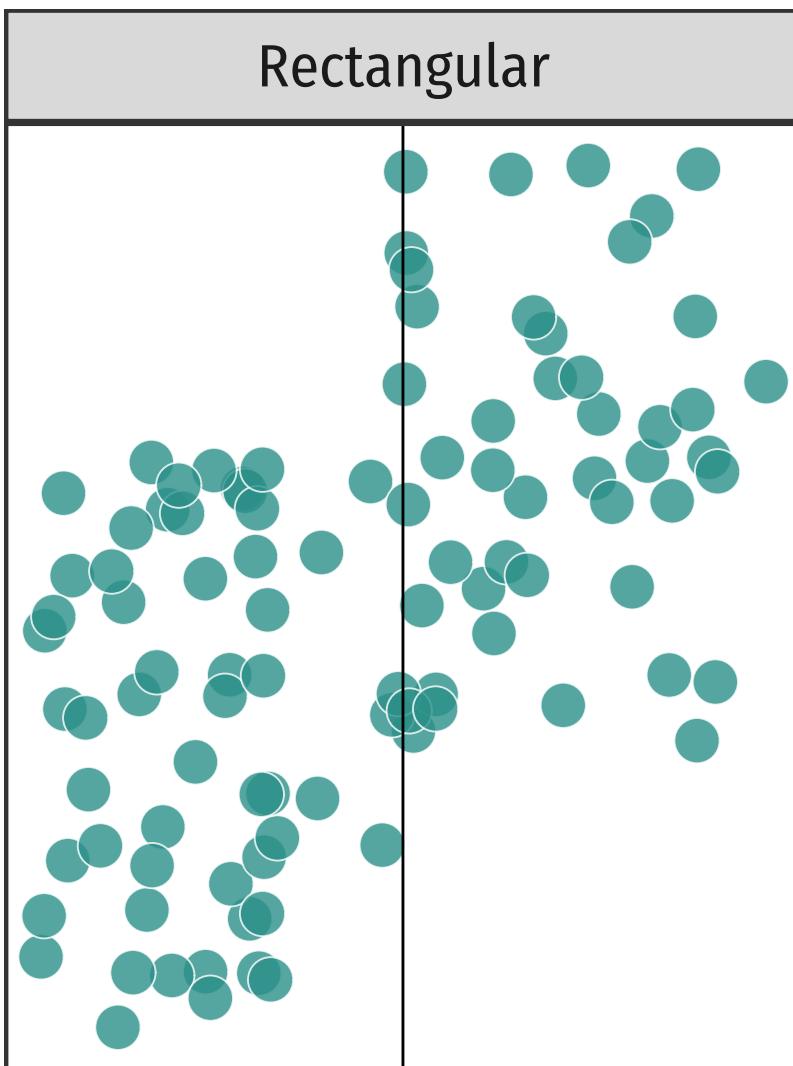
Kernels

Because we care the most about observations right by the cutoff, give more distant ones less weight

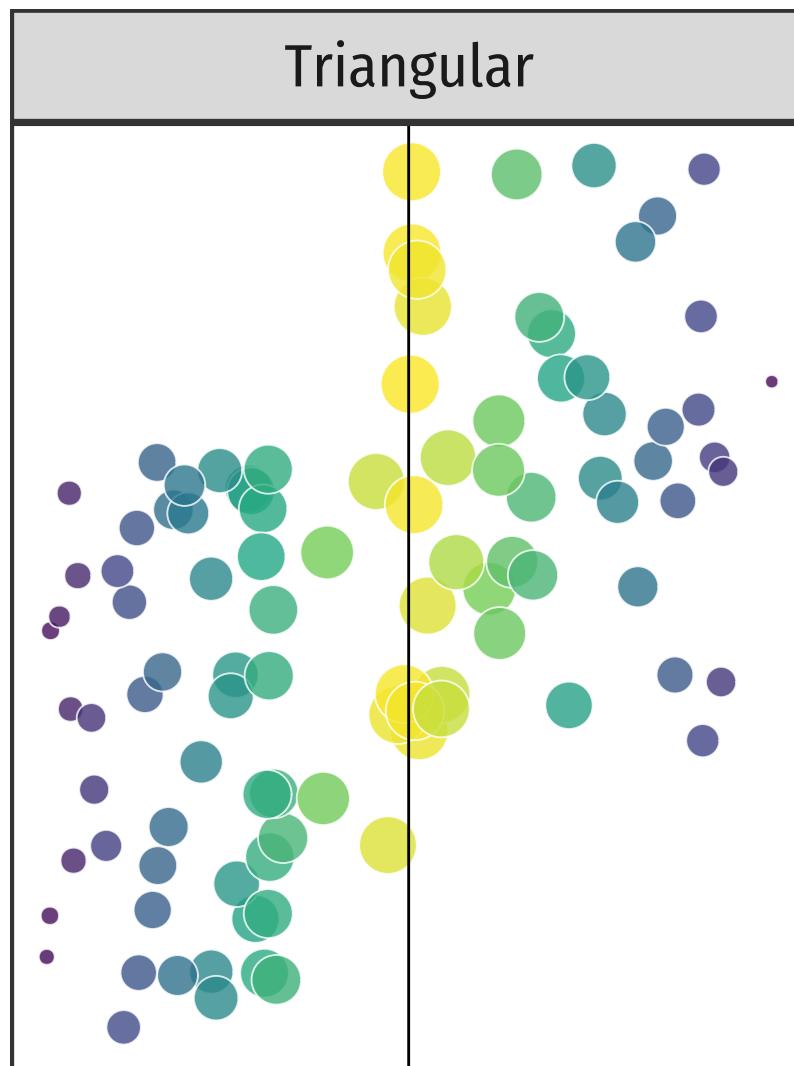
Kernel = method for assigning importance to values based on their distance to the cutoff



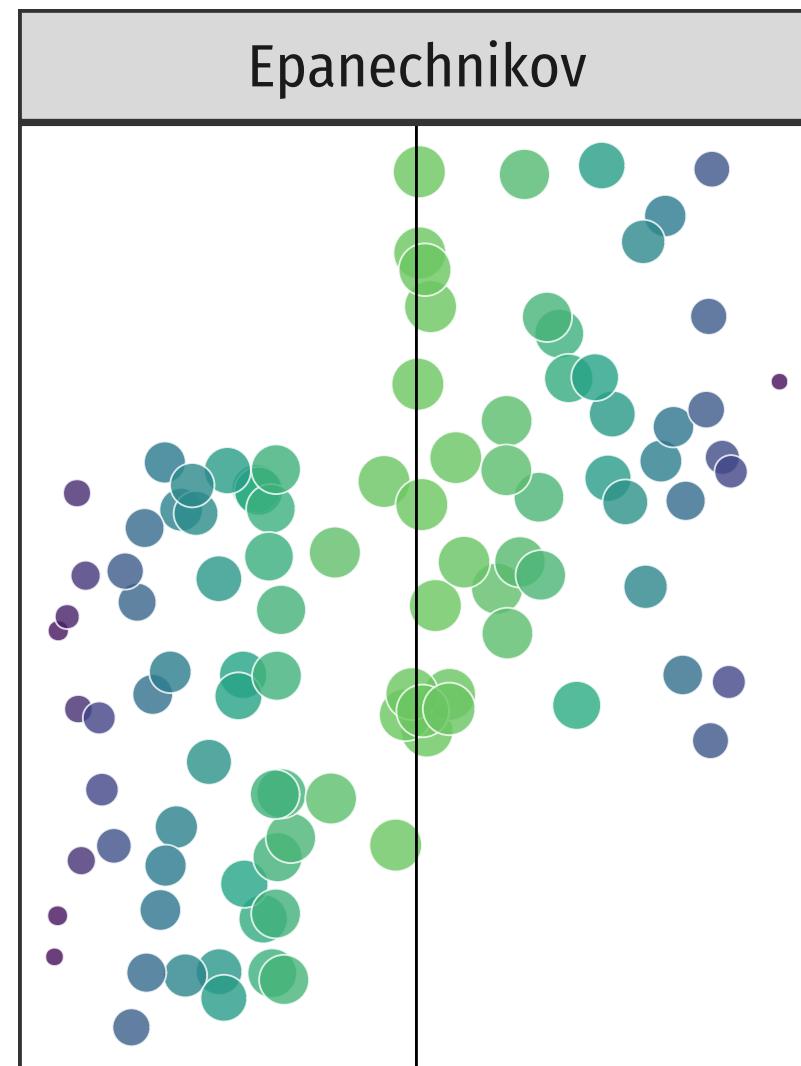
Rectangular



Triangular



Epanechnikov



Weight



Try everything!

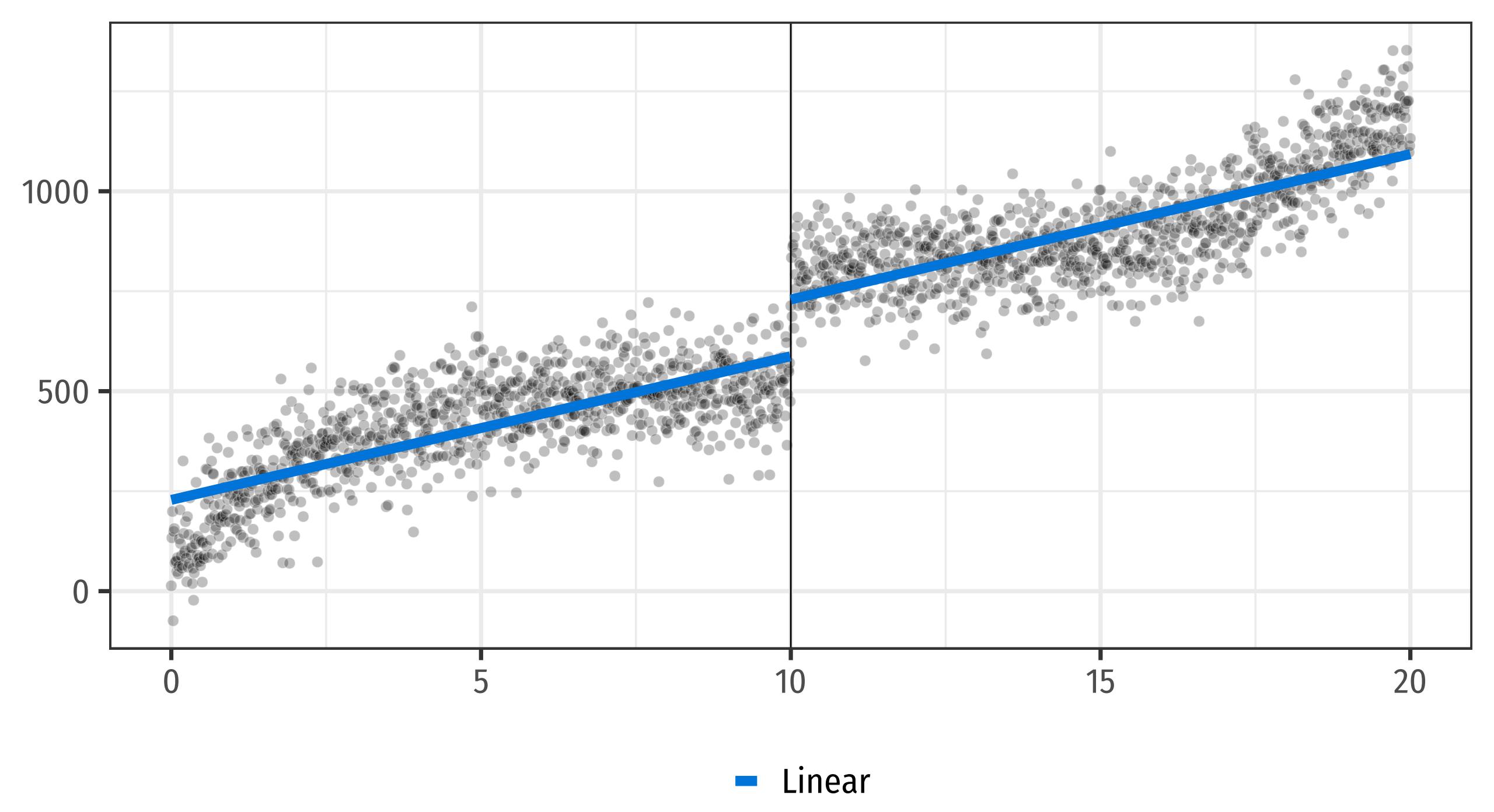
Your estimate of δ depends on all these:

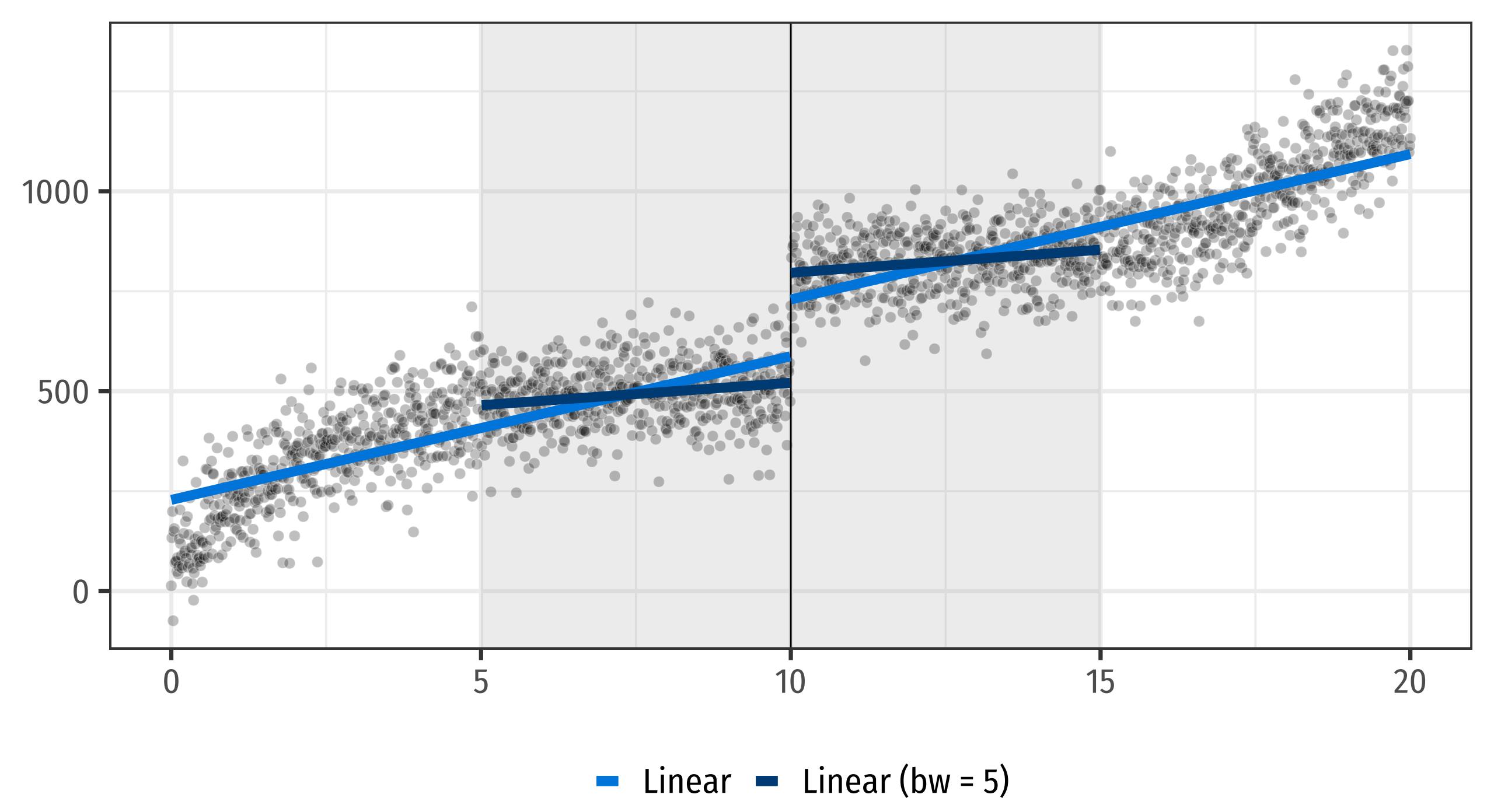
Line type (parametric vs. nonparametric)

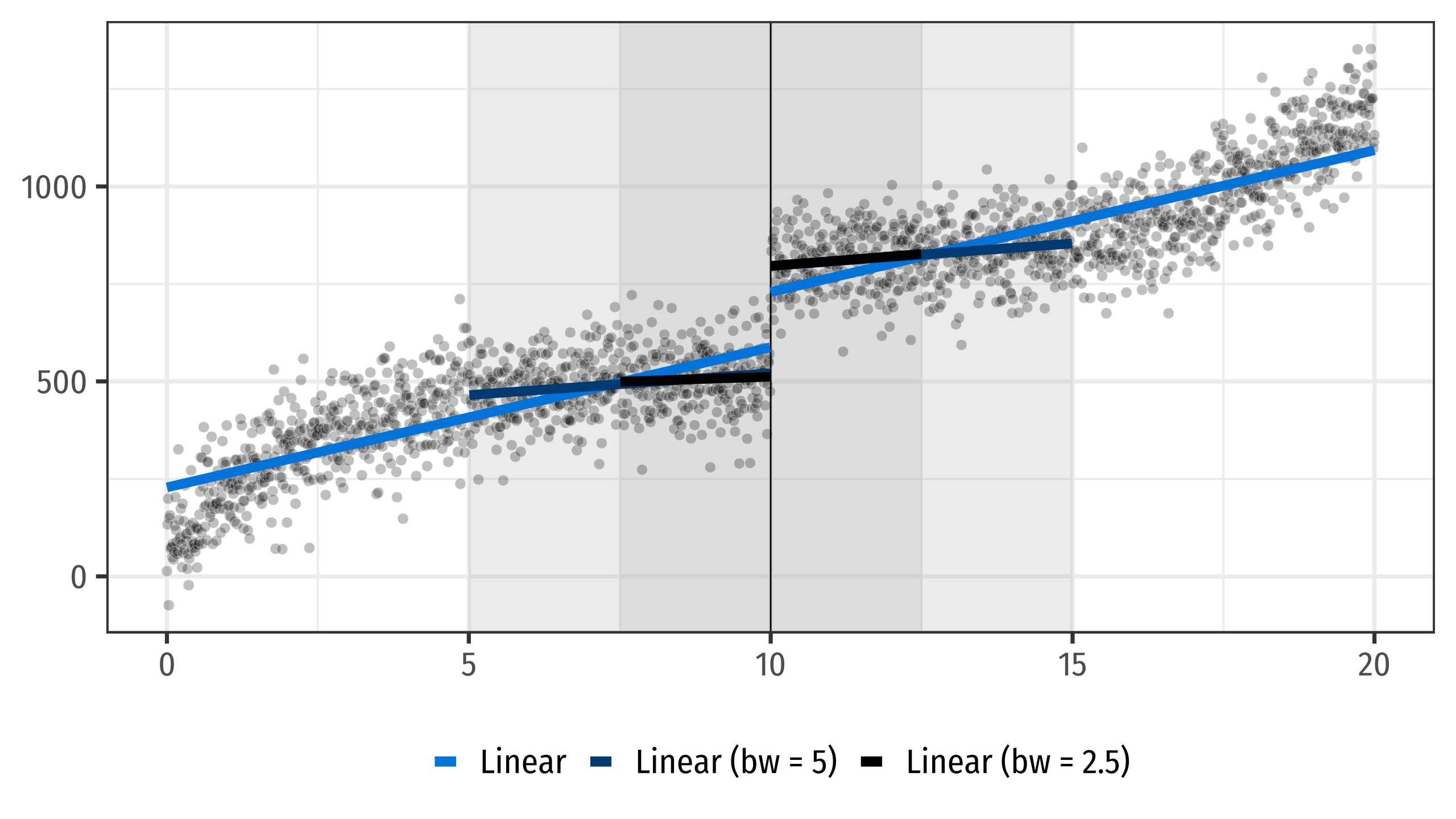
Bandwidth (wide vs. narrow)

Kernel weighting

Try lots of different combinations!





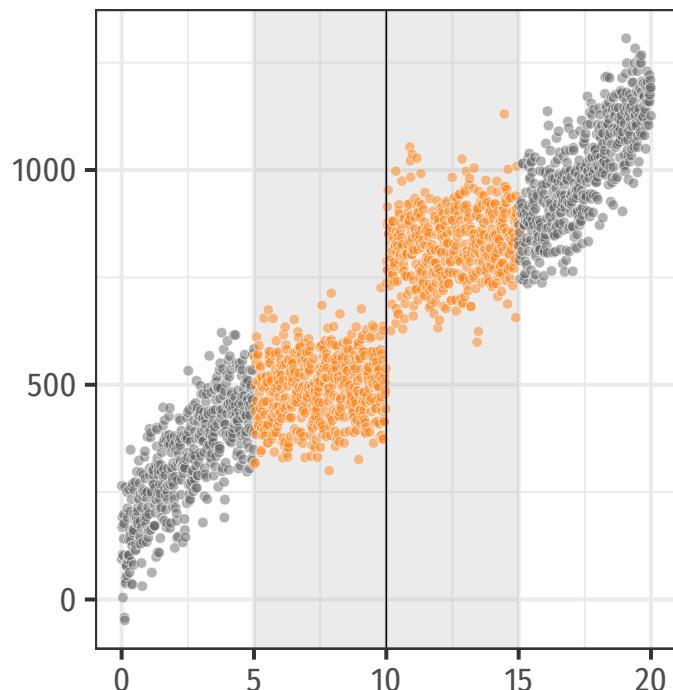


Main RDD concerns

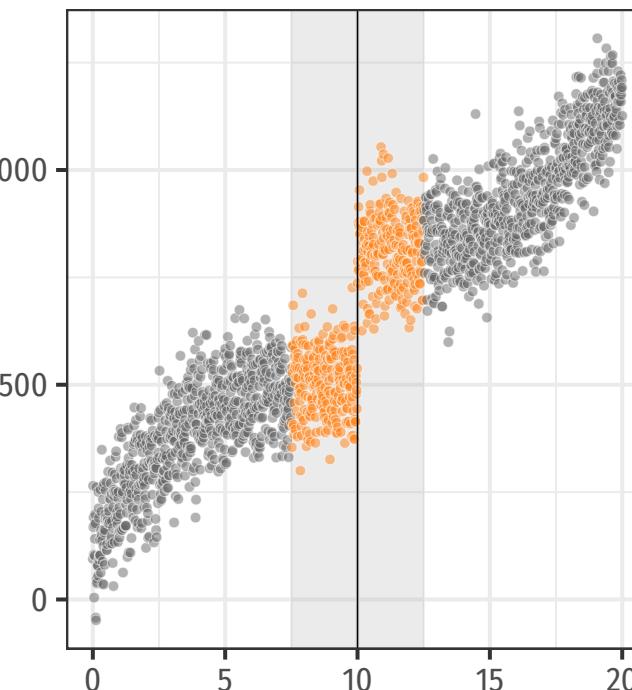
It's greedy!

You need *lots* of data, since
you're throwing most of it away

Bandwidth = 5



Bandwidth = 2.5



It's limited in scope!

You're only measuring the ATE for people in the bandwidth

Local Average Treatment Effect (LATE)

It's limited in scope!

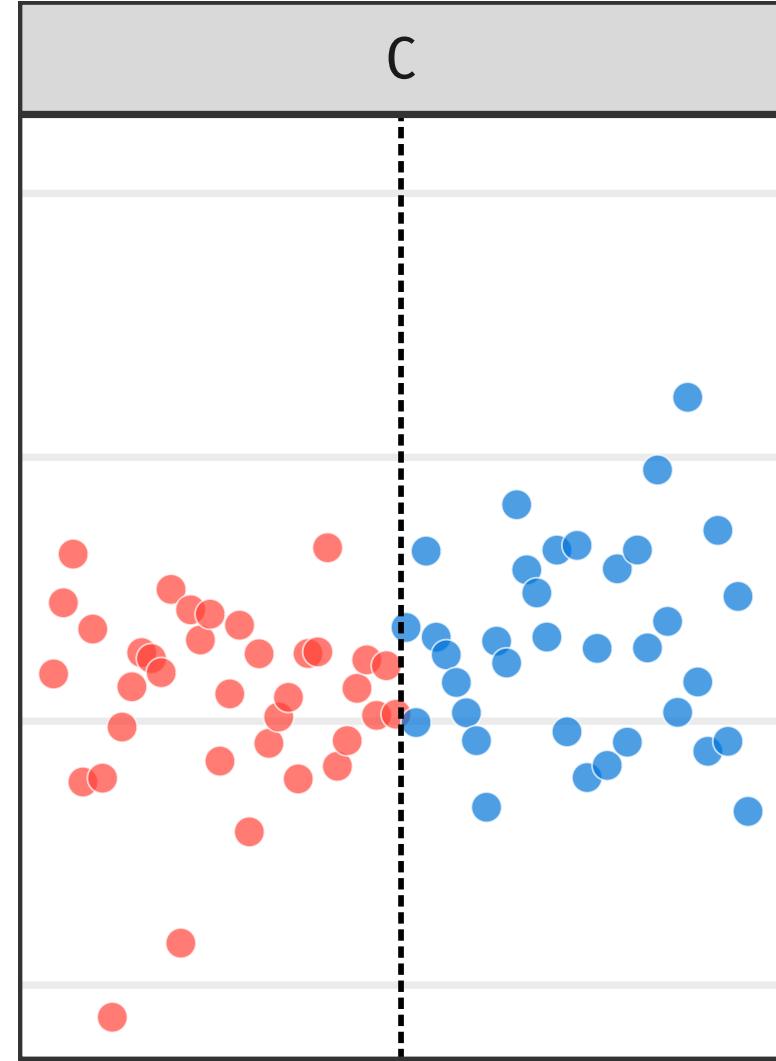
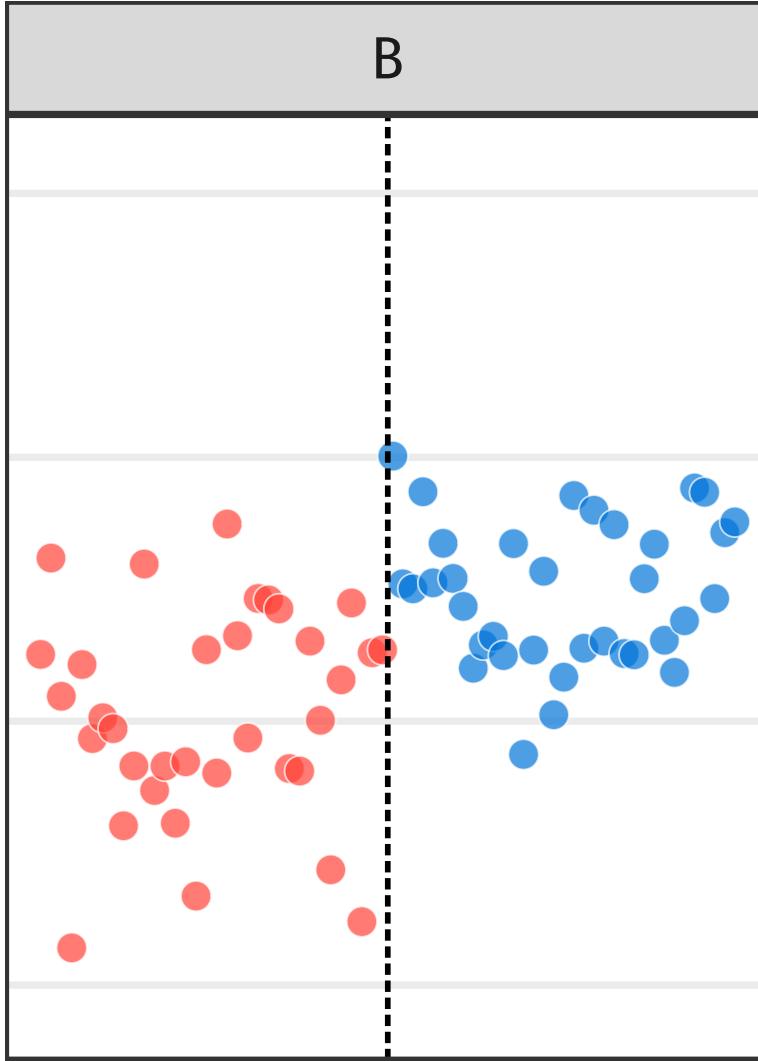
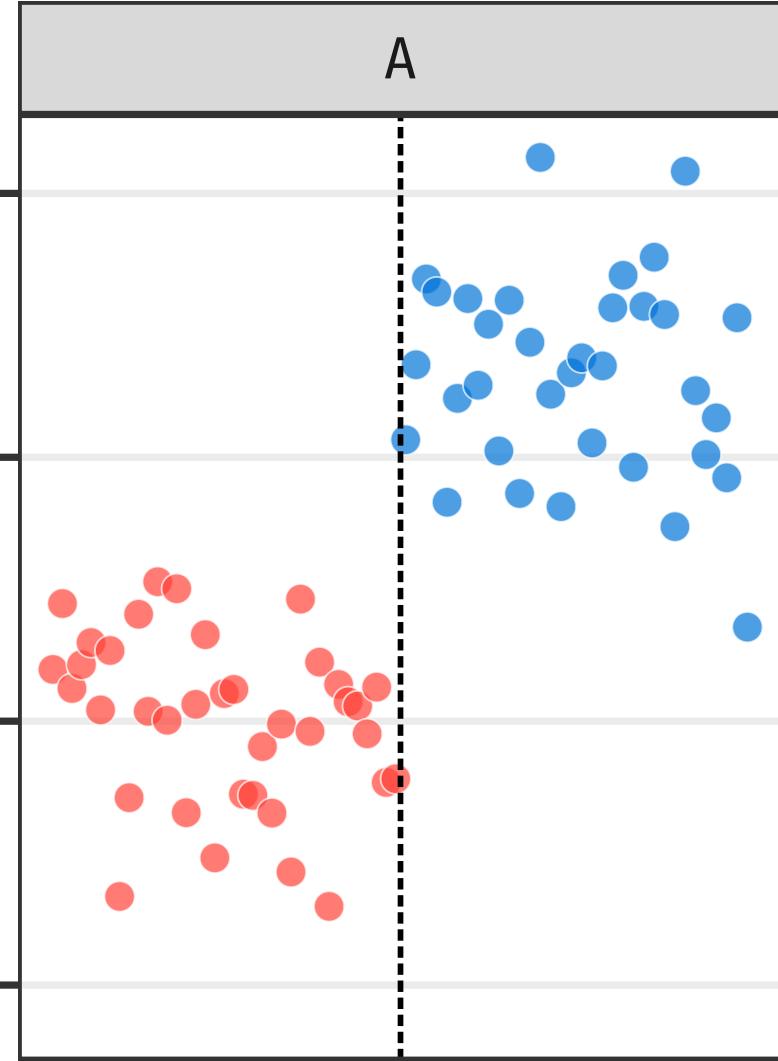
You can't make population-level claims with a LATE

(But can you really do that with RCTs and diff-in-diff anyway?)

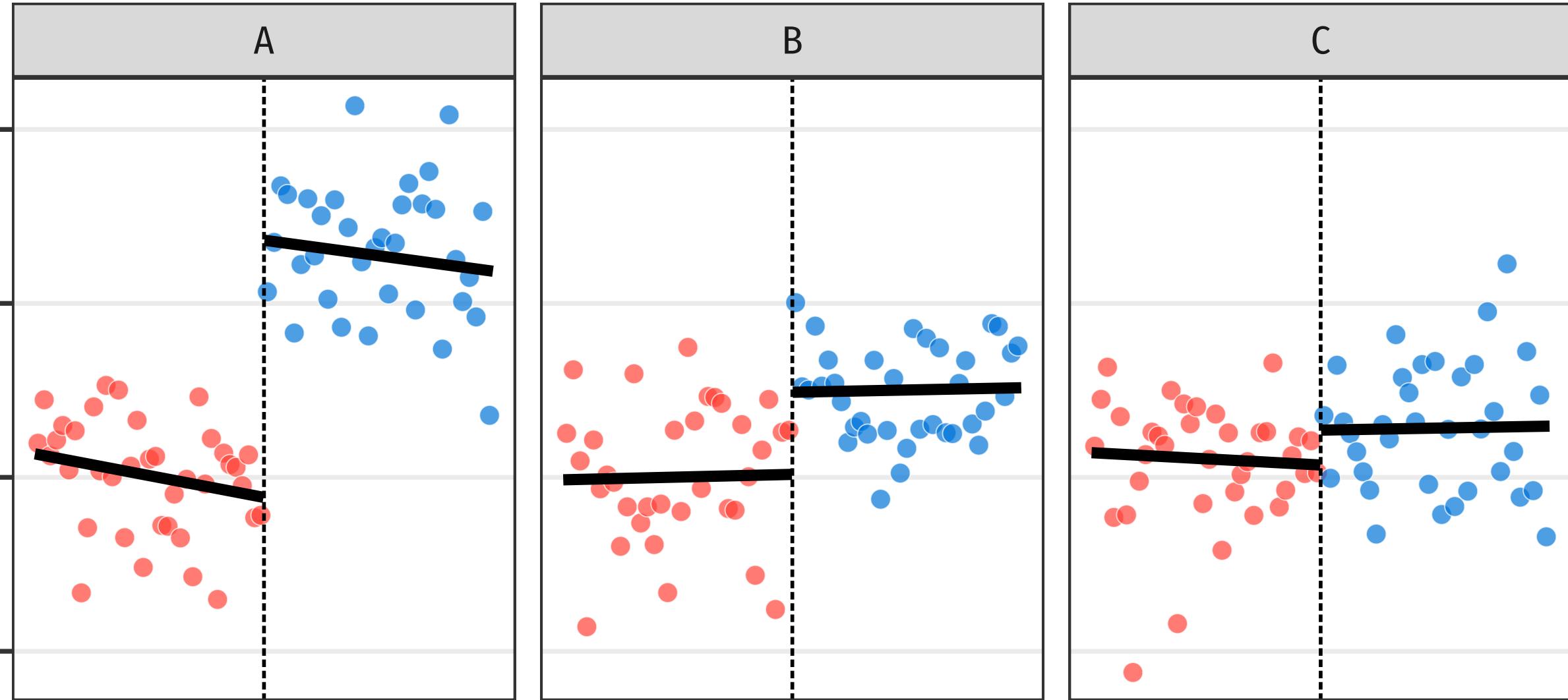
“The realistic conclusion to draw is that all quantitative empirical results that we encounter are ‘local’”

Angrist and Pischke, *Mostly Harmless Econometrics*, p. 23–24

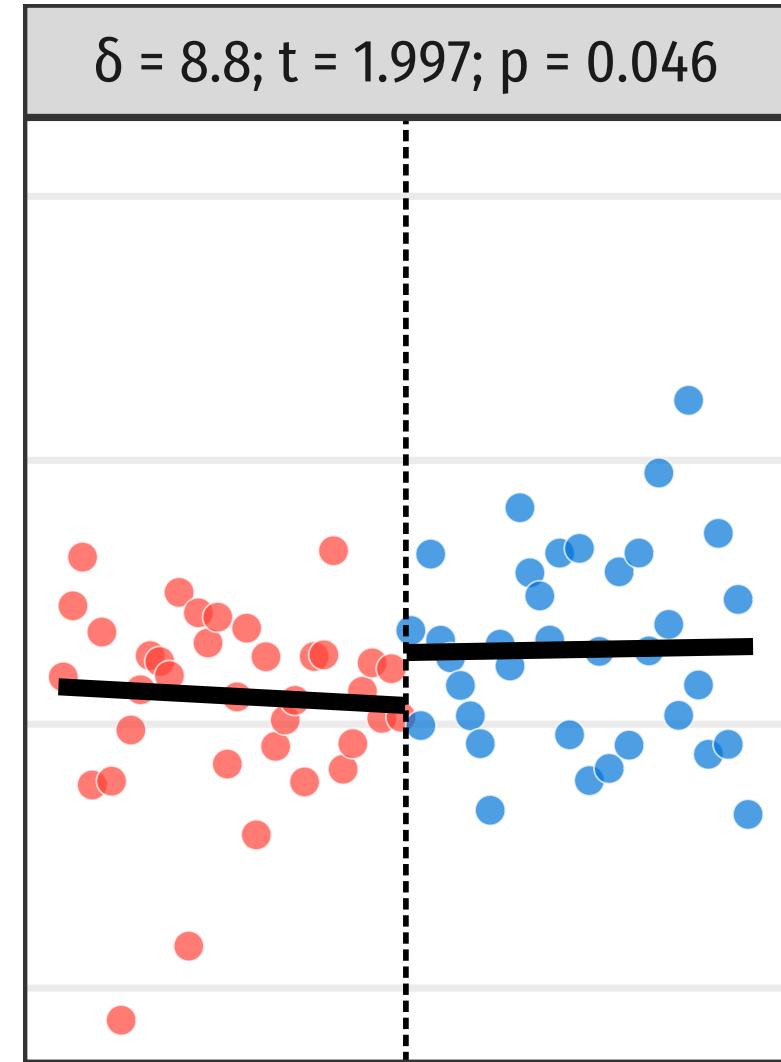
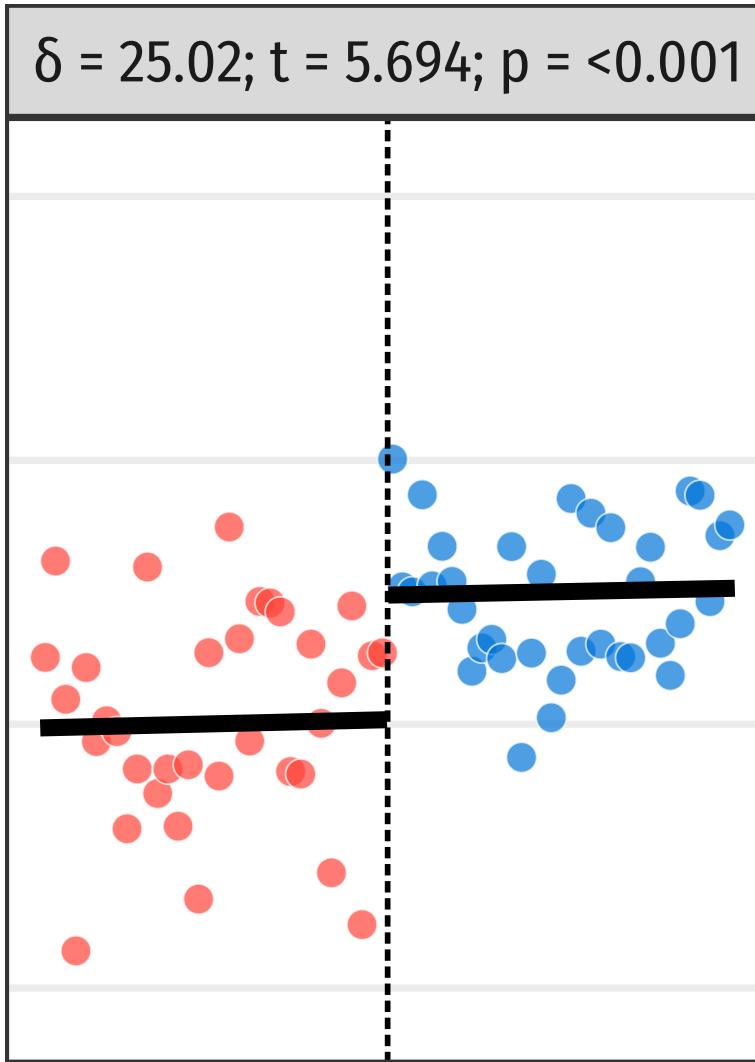
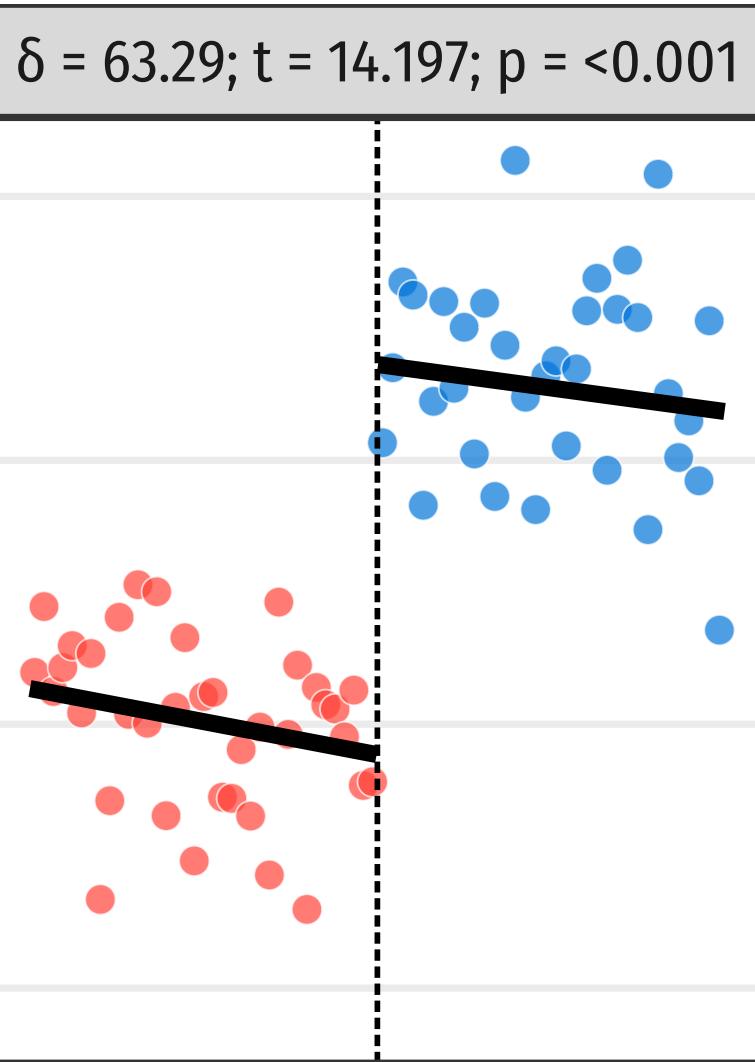
Graphics are neat!



Which ones are significant?



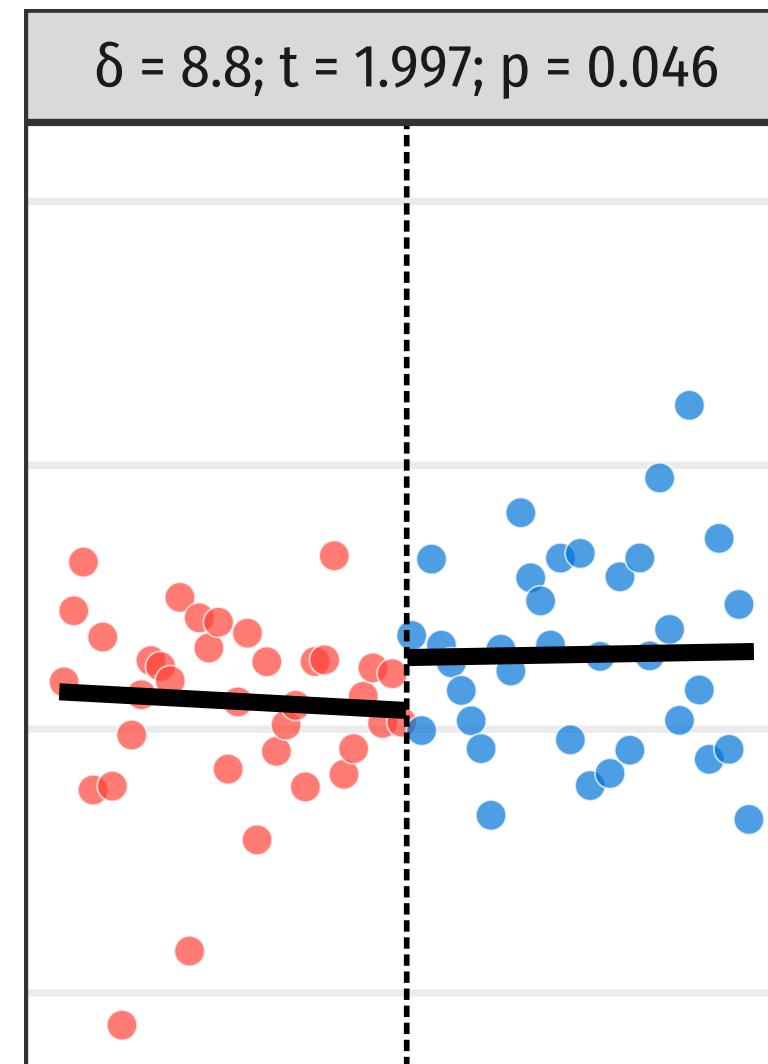
All of them!



Don't rely only on graphics!

Super clear breaks
are uncommon

Make graphs, but
also find the
actual δ value



Manipulation!

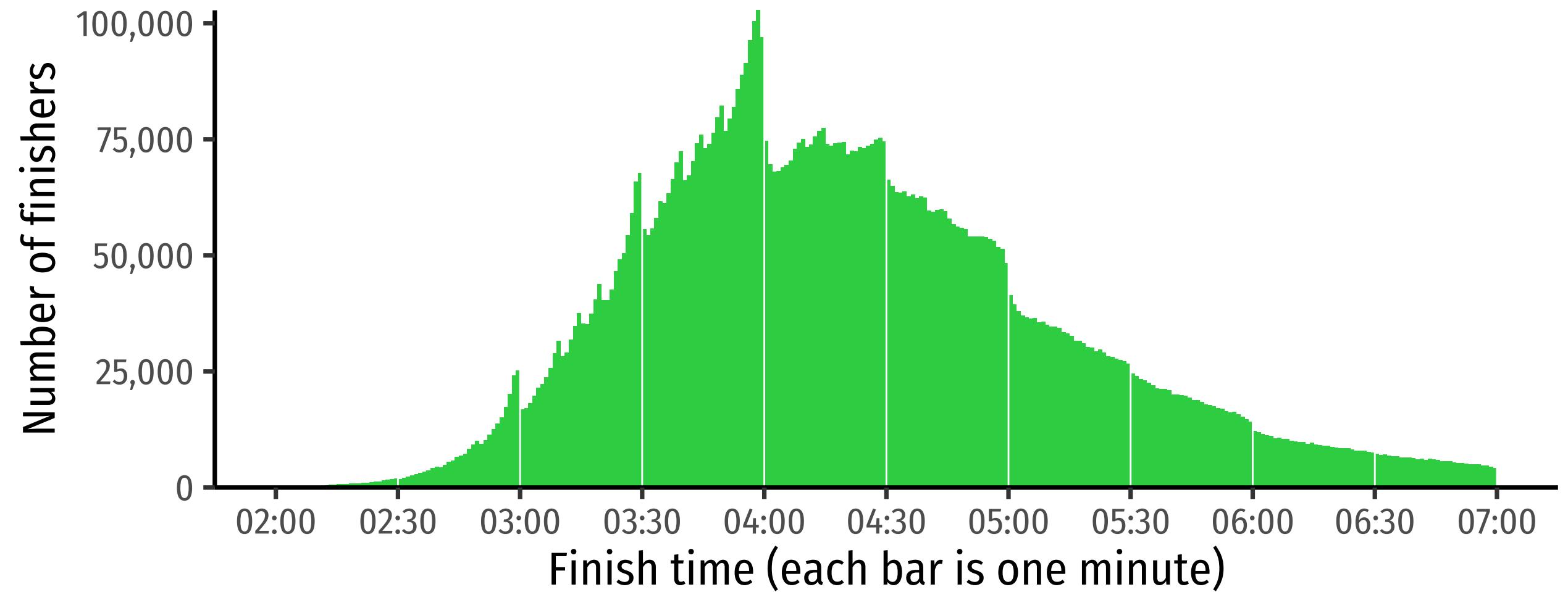
(if people know about the cutoff)

People might fudge numbers or work to hit the threshold to get in/out of program

If so, those right next to the cutoff are no longer comparable treatment/control groups

Distribution of marathon finishing times

N = 9,589,053



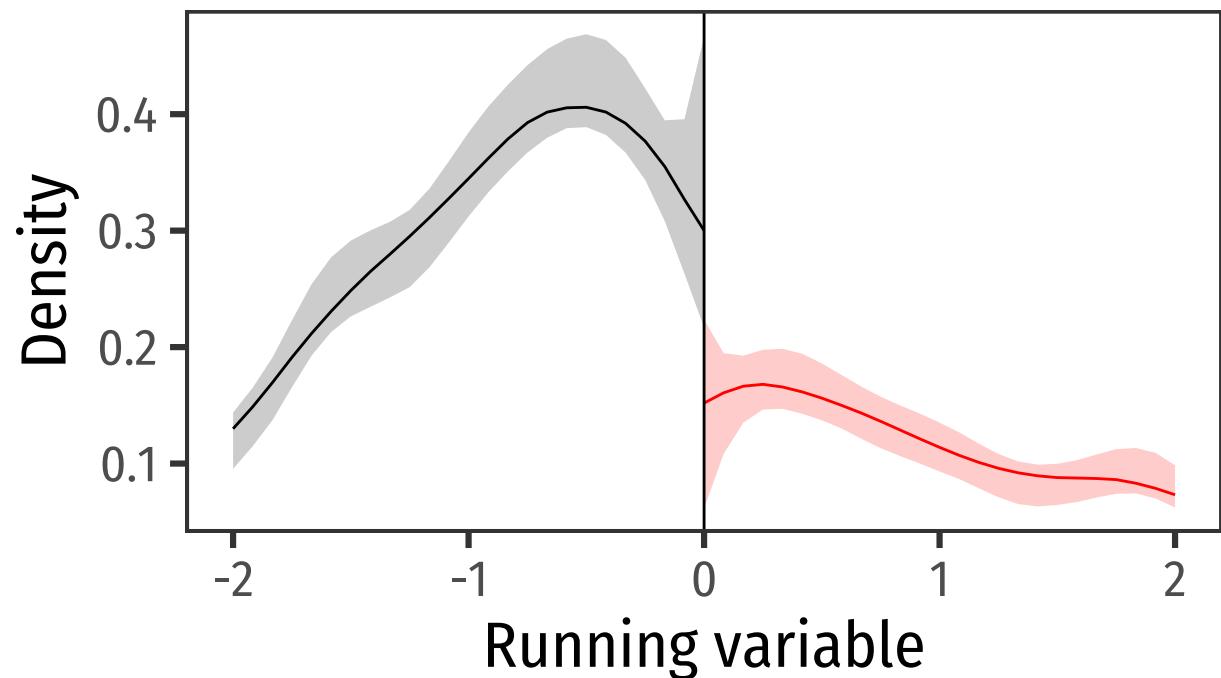
Eric J. Allen, Patricia M. Dechow, Devin G. Pope, George Wu (2017)
Reference-Dependent Preferences: Evidence from Marathon Runners.
Management Science 63(6):1657-1672. <https://doi.org/10.1287/mnsc.2015.2417>

Manipulation

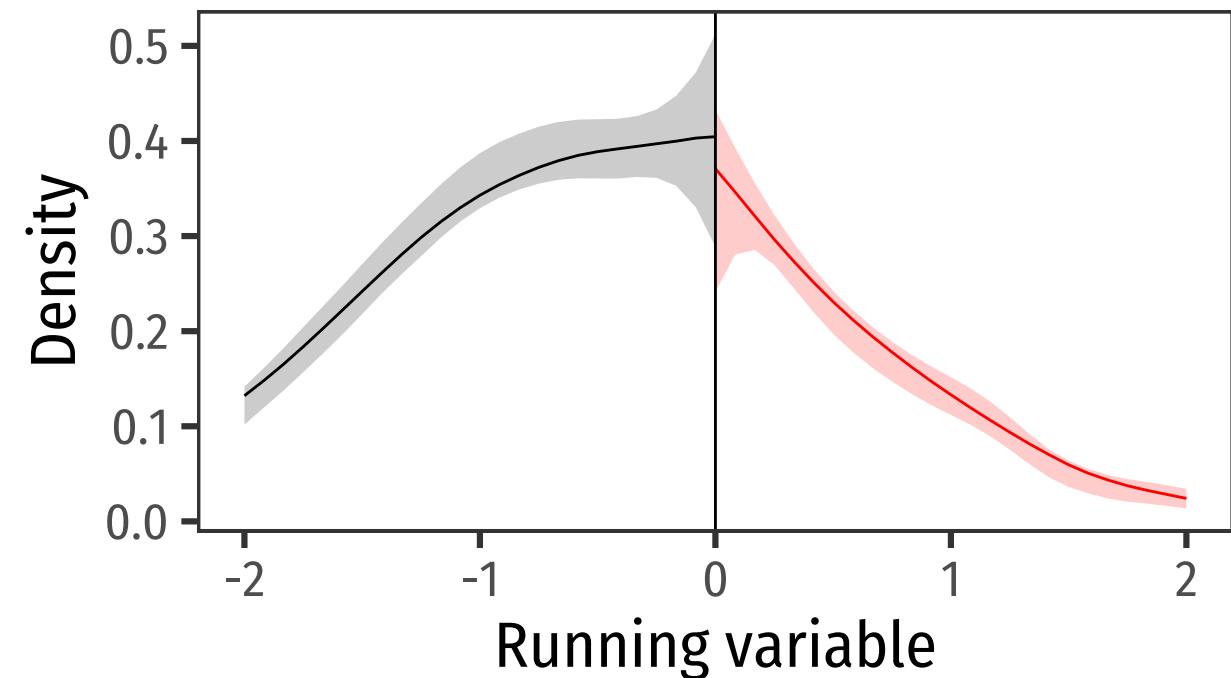
Check with a McCrary density test

`rddensity::rdplotdensity() in R`

Manipulation



No manipulation



Noncompliance!

**People on the margin of the cutoff
might end up in/out of the program**

**The ACA, subsidies, Medicaid,
and 138% of the poverty line**

Sharp vs. fuzzy discontinuities

Sharp discontinuity

TRUE

Perfect compliance

FALSE

40

60

80

100

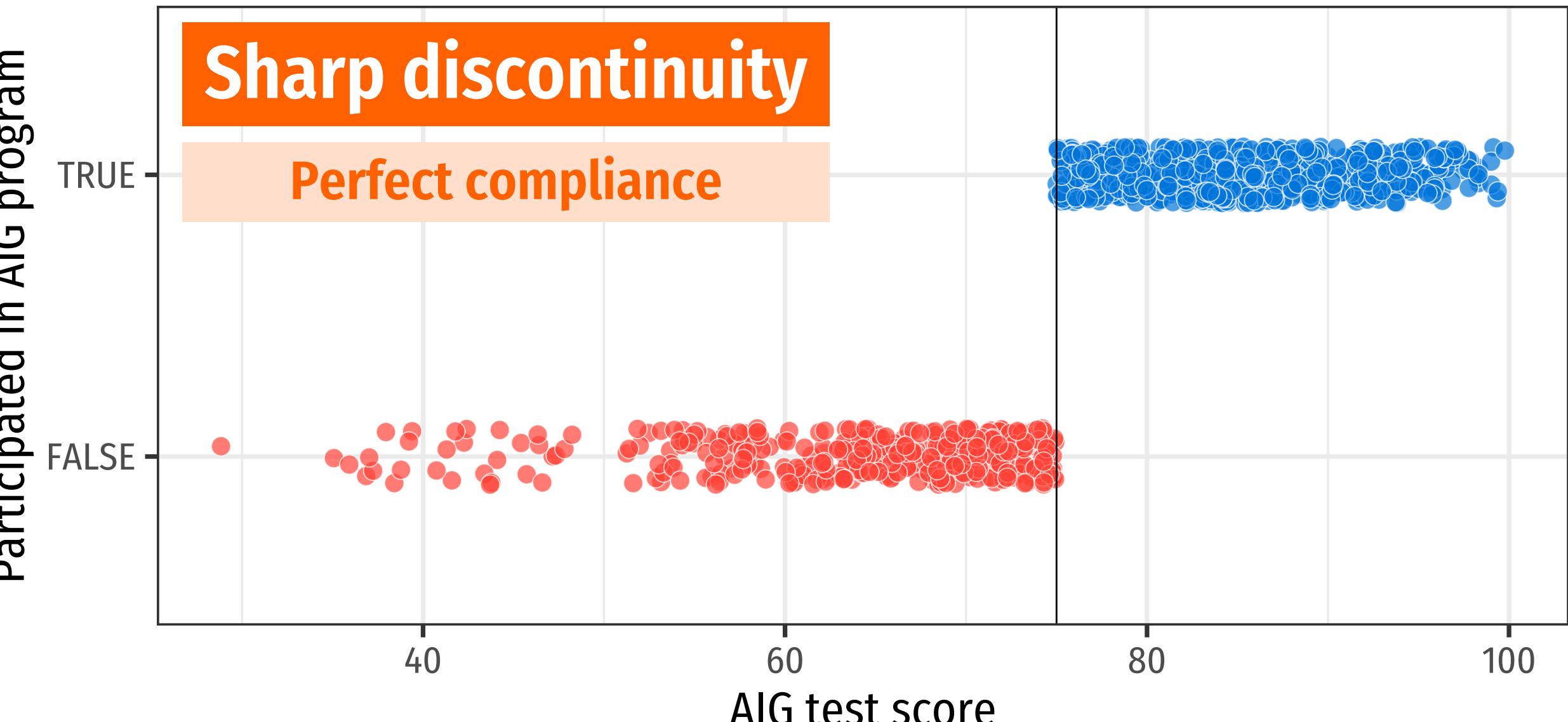
AIG test score

Test score ≥ 75

• FALSE

• TRUE

Participated in AIG program



Fuzzy discontinuity

TRUE

Imperfect compliance

Participated in AIG program

FALSE

40

60

80

100

AIG test score

Test score ≥ 75 • FALSE • TRUE



Fuzzy discontinuities

Address noncompliance with instrumental variables (more on those next time!)

Use an instrument for which side of the cutoff people should be on

**Effect is only for compliers near the cutoff
(complier LATE; doubly local)**

RDD with R

1: Is assignment to treatment rule-based?

If not, stop!

2: Is design fuzzy or sharp?

Either is fine; sharp is easier.

3: Is there a discontinuity in running variable at cutpoint?

Hopefully not.

**4: Is there a discontinuity in outcome
variable at cutpoint in running variable?**

Hopefully.

5: How big is the gap?

Measure parametrically and nonparametrically.

R time!