

# Chapter 10: Comparing multiple independent populations

(Ott & Longnecker Sections: 14.2 and 14.5)

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Part 2

<https://dzwang91.github.io/stat371/>



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- The variances are the same for all treatments
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How do we check the equal variance?



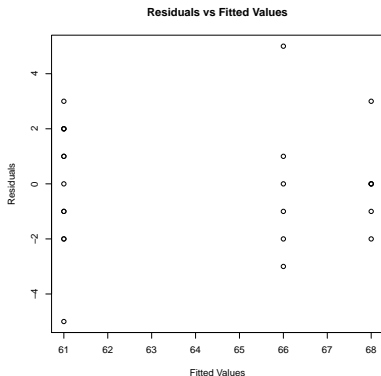
- The **fitted values** are the treatment means  $\bar{y}_{i.}$ .
- The **residuals** are the differences between the observed data ( $y_{ij}$ ) and the treatment means  $y_{ij} - \bar{y}_{i.}$ .
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- In fact, the sum of squares error (SSE) is the sum of the squares of the residuals.
- **The approach to check equal variance:** If the spreads of residuals are about the same for each treatment, then we are safe to assume equal variance.
- So we need the **Residuals vs Fitted values plot**.

# Residuals vs Fitted values plot



Assume there are 3 treatments. The y-axis is called 'Residuals' and the x-axis is called 'Fitted Values'.



There might be slightly less spread in the last group, but they're close enough.



# Residuals vs Fitted plot continued





# Residuals vs Fitted plot continued



The SD of weights of 5 elephants will tend to be larger than the SD of weights of 5 hamsters.



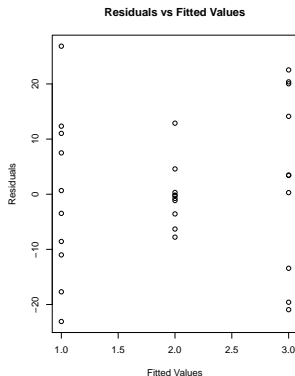
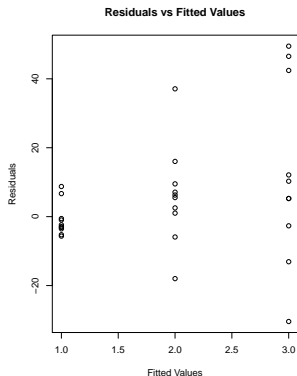


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# Residuals vs Fitted plot continued



- The reason we plot the residuals vs fitted values is that it is often the case that variability will increase with increasing fitted values.
- Some bad examples: In the first, we see the classic funnel pattern where the variability increases with the fitted value. In the second, we still see non-constant variance even though there is no pattern relative to the fitted value.





- Another option for checking equal variance is to use the ratio of SDs guideline first mentioned when comparing two populations.
- Since there are now more than two groups, it is typical to **take the ratio of the largest and smallest sample SDs** - if this ratio passes the test, then every other pair will as well.
- For the example in Part 1 of Chapter 10, the sample SDs for the four groups are 1.83, 2.83, 1.67, and 2.62. The ratio of the smallest to the largest is  $1.67/2.83 = 0.59$ , which falls between 0.5 and 2.0, so assuming the variances equal should be safe.



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- One way to figure these things out is with post-hoc pairwise tests.

$$t = \frac{\bar{y}_{i.} - \bar{y}_{j' .}}{\sqrt{MSE(1/n_i + 1/n_{j'})}}$$

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- Or make a  $100(1 - \alpha)\%$  CI on the difference as follows. **If the CI contains 0, then we don't reject the null hypothesis**, which means the treatment has the same effect.

$$\bar{y}_{i.} - \bar{y}_{i' .} \pm t_{df_E, \alpha/2} \sqrt{MSE(1/n_i + 1/n_{i'})}$$

- We will primarily work with CIs. Suppose we want 95% CIs,  $t_{20,0.025} = 2.086$ , and  $MSE = 5.6$ , so
  - Trt 1 vs Trt 2:  
 $61 - 66 \pm 2.086\sqrt{5.6(1/4 + 1/6)} = -5 \pm 3.19 = (-8.19, -1.81)$
  - Trt 1 vs Trt 3:  
 $61 - 68 \pm 2.086\sqrt{5.6(1/4 + 1/6)} = -7 \pm 3.19 = (-10.19, -3.81)$
  - Trt 1 vs Trt 4:  
 $61 - 61 \pm 2.086\sqrt{5.6(1/4 + 1/8)} = 0 \pm 3.02 = (-3.02, 3.02)$
  - Trt 2 vs Trt 3:  
 $66 - 68 \pm 2.086\sqrt{5.6(1/6 + 1/6)} = -2 \pm 2.85 = (-4.85, 0.85)$
  - Trt 2 vs Trt 4:  
 $66 - 61 \pm 2.086\sqrt{5.6(1/6 + 1/8)} = 5 \pm 2.67 = (2.33, 7.67)$
  - Trt 3 vs Trt 4:  
 $68 - 61 \pm 2.086\sqrt{5.6(1/6 + 1/8)} = 7 \pm 2.67 = (4.33, 9.67)$
- The conclusion is that treatments 2 and 3 are the same, and 1 and 4 are the same, but 2 and 3 differ from 1 and 4.



- The information is summarized by sorting the treatment means from largest to smallest, and then adding letter codes. Two treatments share a letter if they do not differ significantly:

Treatment	Sample Mean	Letter Code
3	68	A
2	66	A
1	61	B
4	61	B



If we end up doing a bunch of pairwise tests, why do we use ANOVA?

- Theoretically, the F-test is the most powerful test for the hypotheses we specified, provided all of our assumptions are met.
- Practically, if there are a lot of treatments, if there are really no differences between the treatments, doing one ANOVA could save time over doing many pairwise tests.