

ABSOLUTELY CONTINUOUS RANDOM VARIABLES

X discrete $\rightarrow F_x$ piecewise constant

We say that X is a **CONTINUOUS R.V.**

if $F_x : \mathbb{R} \rightarrow [0, 1]$ is continuous ($X : \Omega \rightarrow \mathbb{R}$)

Def $X : \Omega \rightarrow \mathbb{R}$ r.v. We say that X is
A**BSOLUTELY CONTINUOUS** if there exists a function

$f_x : \mathbb{R} \rightarrow [0, +\infty)$ s.t.

$$F_x(x) = P(X \leq x) = \int_{-\infty}^x f_x(\epsilon) d\epsilon \quad \forall x \in \mathbb{R}$$

f_x is called the **DENSITY** of the r.v.

Note X absolutely continuous $\Rightarrow X$ continuous

- In general f_x not continuous, just integrable.

Properties of density

$$1) \int_{-\infty}^{\infty} f_x(\epsilon) d\epsilon = 1$$

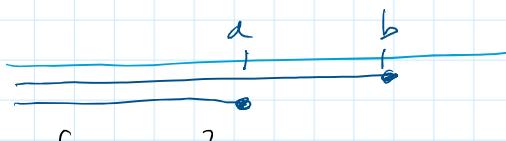
$$P(X \leq +\infty) = P(X \in \mathbb{R}) = 1 = \mu_X(\mathbb{R})$$

$$2) P(a < X \leq b) = F_x(b) - F_x(a) \quad \forall a < b, a, b \in \mathbb{R}$$

↪ for any r.v.

$$F_x(b) = P(X \leq b) \quad F_x(a) = P(X \leq a)$$

$$\{a < X \leq b\} = \{X \leq b\} \setminus \{X \leq a\}$$



and $\{X \leq a\} \subset \{X \leq b\}$

$$\Rightarrow P(a < X \leq b) = P(X \leq b) - P(X \leq a)$$

if X is absolutely continuous then

$$P(a \leq X \leq b) = \int_a^b f_X(\epsilon) d\epsilon$$

$$L = \int_{-\infty}^b f_X(\epsilon) d\epsilon - \int_{-\infty}^a f_X(\epsilon) d\epsilon = \int_a^b f_X(\epsilon) d\epsilon$$

3) for a general set $B \in \mathcal{B}(\mathbb{R})$

$$P(X \in B) = \int_B f_X(\epsilon) d\epsilon$$

4) $P(X = a) = 0 \quad \forall a \in \mathbb{R}$

$$L = P(\{\omega : X(\omega) = a\}) = \int_a^a f_X(\epsilon) d\epsilon = 0$$

Note

$$1 = P(X \in \mathbb{R}) = P\left(\bigcup_{x \in \mathbb{R}} \{X = x\}\right)$$

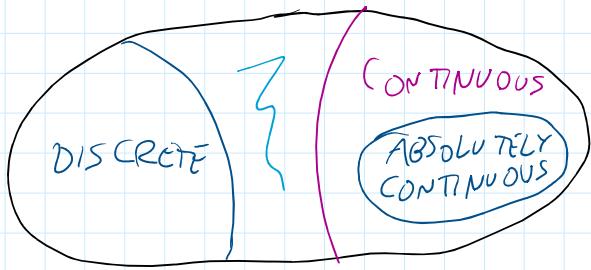
$$\bigcup_{x \in \mathbb{R}} \{X = x\} \quad \begin{matrix} \text{disjoint union} \\ \text{NOT COUNTABLE} \end{matrix}$$

$$\neq \sum_{x \in \mathbb{R}} P(X = x) = \sum_{x \in \mathbb{R}} 0 = 0$$

$$\text{We see that } P\left(\bigcup_{n \in \mathbb{N}} A_n\right) = \sum_{n \in \mathbb{N}} P(A_n)$$

ONLY for COUNTABLE UNIONS of sets

NOT for uncountable unions



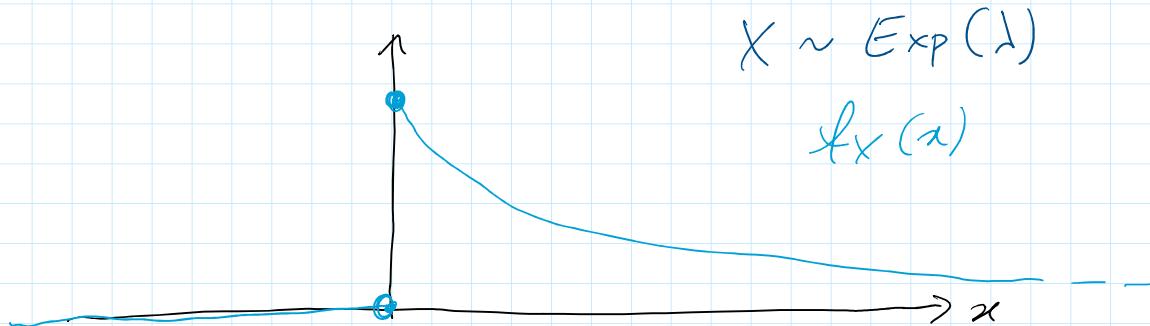
Facts

- there exist continuous r.v. which are not absolutely continuous
- there exist r.v. neither discrete nor continuous

Example of abs. cont. r.v.

$\lambda > 0$, X is EXPONENTIAL r.v. with parameter λ

$$f_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ \lambda e^{-\lambda x} & \text{if } x \geq 0 \end{cases}$$



Compute $\int_{-\infty}^{\infty} f_X(t) dt$ has to be 1

$$\begin{aligned} \int_{-\infty}^{+\infty} f_X(t) dt &= \int_{-\infty}^0 0 dt + \int_0^{+\infty} \lambda e^{-\lambda x} dx \\ &= 0 + \cancel{\lambda} \left[-\frac{1}{\lambda} e^{-\lambda x} \right]_{x=0}^{x \rightarrow +\infty} = -\cancel{e^{-\lambda \infty}} + e^0 = 1 \end{aligned}$$

Compute distribution function

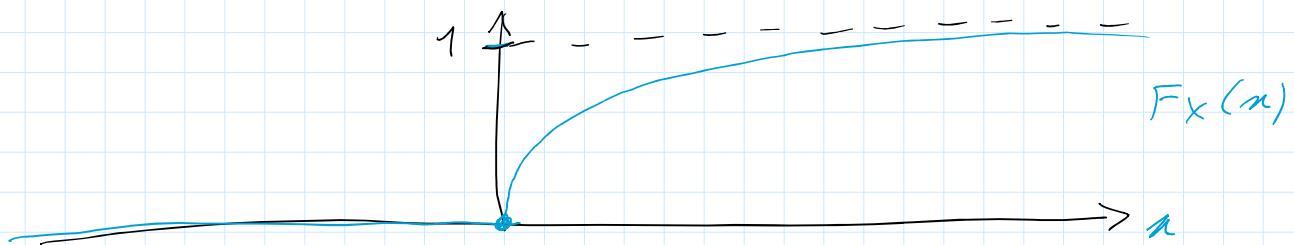
Compute distribution function

$$F_X(x) = \int_{-\infty}^x f_X(t) dt$$

$$\bullet = \int_{-\infty}^x 0 dt = 0 \quad \text{if } x < 0$$

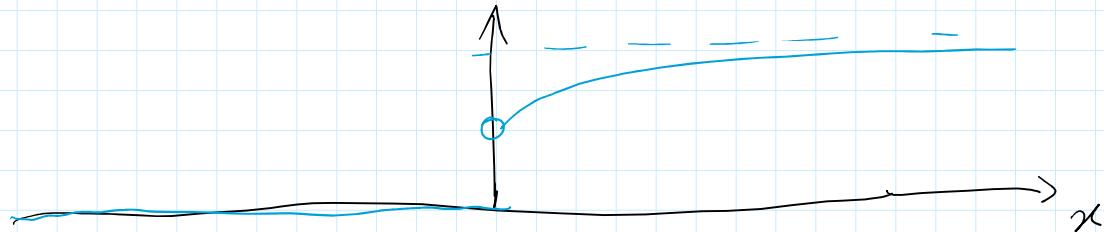
$$\bullet \text{ if } x > 0 \quad F_X(x) = \int_{-\infty}^0 0 dt + \int_0^x \lambda e^{-\lambda t} dt \\ = 0 + \lambda \left[-\frac{1}{\lambda} e^{-\lambda t} \right]_{t=0}^{t=x} = -e^{-\lambda x} + 1$$

$$F_X(x) = \begin{cases} 0 & \text{if } x \leq 0 \\ 1 - e^{-\lambda x} & \text{if } x > 0 \end{cases}$$



F_X is continuous

Example of a distribution function of a r.v.
neither discrete nor continuous



has a jump \rightarrow NOT continuous

not piecewise constant for $x > 0 \rightarrow$ NOT discrete.

Remark

When X is absolutely continuous r.v., it means that

When X is absolutely continuous r.v., it means that
 $F_X : \mathbb{R} \rightarrow [0, 1]$ is absolutely continuous function
 μ_X is a measure absolutely continuous
 (with respect to Lebesgue measure on \mathbb{R})

EXPECTATION (MEAN) OF RANDOM VARIABLES

$X : \Omega \rightarrow \mathbb{R}$ "expected value"

$X \sim \text{Ber}(p)$ $X(\omega) \in \{0, 1\}$

$$P(X=0) = 1-p \quad P(X=1) = p$$

"The value of X is 1 p -times
 0 $(1-p)$ -times

$$\text{expected value} = 0 \cdot (1-p) + 1 \cdot p$$

$$\begin{aligned} E[X] &= 1 P(X=1) + 0 P(X=0) \\ &= p \end{aligned}$$

- Consider X s.c. $P(X=a) = p$, $P(X=b) = 1-p$
 $a, b \in \mathbb{R}$, $a \neq b$

$$\begin{aligned} E[X] &= a P(X=a) + b P(X=b) \\ &= a p + b(1-p) \end{aligned}$$

$$p = \frac{1}{2} \rightarrow E[X] = \frac{a+b}{2}$$

Let X be a FINITE r.v. with density p_X

$$\exists N \subseteq \mathbb{R}, N \text{ finite } P(X \in N) = 1$$

$\dots \cup r_{N-1} \cup r_N \cup \dots$

$\exists N \subseteq \mathbb{R}$, N finite $P(X \in N) = 1$

$$\forall k \in N \quad P(X = k) = p_X(k)$$

$$E[X] = \sum_{k \in N} k P(X = k) = \sum_{k \in N} k p_X(k)$$

Example $X \sim \text{Bin}(n, p)$

$$k \in \{0, 1, \dots, n\} \quad p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$$

$$E[X] = \sum_{k=0}^n k p_X(k) = \sum_{k=0}^n k \binom{n}{k} p^k (1-p)^{n-k}$$

$$k \binom{n}{k} = k \frac{n!}{k! (n-k)!} = \frac{n \cdot (n-1)!}{(k-1)! ((n-1)-(k-1))!}$$

$$k! = k \cdot (k-1)(k-2) \cdots 1 = k (k-1)!$$

$$= n \sum_{k=1}^n \binom{n-1}{k-1} p^k (1-p)^{n-k}$$

$$= n \sum_{k=1}^n \binom{n-1}{k-1} p \cdot p^{k-1} (1-p)^{(n-1)-(k-1)}$$

$$= np \sum_{k=1}^n \binom{n-1}{k-1} p^{k-1} (1-p)^{(n-1)-(k-1)}$$

$$= 1$$

$$\text{because } (p + (1-p))^{n-1} = \sum_{l=0}^{n-1} \binom{n-1}{l} p^l (1-p)^{(n-1)-l}$$

$$(l=k-1) \sum_{k=1}^n \binom{n-1}{k-1} p^{k-1} (1-p)^{(n-1)-(k-1)}$$

$$\Rightarrow E[X] = np \quad \text{if } X \sim \text{Bin}(n, p)$$

if X is discrete we could define

$$E[X] = \sum_{x \in N} x P(X = x) \quad N \text{ countable}$$

sum of series \rightarrow can be finite or infinite
 can be absolutely convergent or not

BETTER to consider case when

$$\sum_{x \in N} |x| P(X=x) \text{ is finite}$$

and thus to define expectation in such case

X discrete, $N \subseteq \mathbb{R}$ countable $P(X \in N) = 1$

$$S_+ = \sum_{\substack{x \in N \\ x > 0}} x p_X(x)$$

$$S_- = \sum_{\substack{x \in N \\ x < 0}} (-x) p_X(x)$$

S_+ and S_- are always defined, might be infinite

$$S_+, S_- \in [0, +\infty]$$

Def If at least one of S_+ and S_- are finite
 then we can define

$$E[X] = S_+ - S_-$$

Note $E[X] = \begin{cases} +\infty & \text{if } S_+ = +\infty, S_- < +\infty \\ +\infty & \text{if } S_+ = +\infty, S_- \geq +\infty \\ -\infty & \text{if } S_+ < +\infty, S_- = +\infty \end{cases}$

If both S_+ and S_- are FINITE, we say that
 the r.v. X is INTEGRABLE and we have

$$E[X] = S_+ - S_- \text{ is finite}$$

Remark With this definition

X is integrable $\Leftrightarrow |X|$ is integrable
and $E[|X|] = S_+ + S_-$

Proof if $S_+, S_- < \infty$ then $S_+ + S_-$ is finite

$$\begin{aligned} E[|X|] &= \sum_{x \in \mathbb{R}} |x| p_X(x) \\ &= \sum_{x \geq 0} x p_X(x) + \sum_{x < 0} (-x) p_X(x) \\ &= S_+ + S_- \quad \blacksquare \end{aligned}$$

Instead $E[X] = \sum_{x \in \mathbb{R}} x p_X(x)$ (if series is absolutely convergent)

$$\begin{aligned} &= \sum_{x \geq 0} x p_X(x) + \sum_{x \leq 0} x p_X(x) \\ &= \sum_{x \geq 0} x p_X(x) - \sum_{x \geq 0} (-x) p_X(x) \\ &= S_+ - S_- \end{aligned}$$

$L^1(\Omega)$ = space of integrable random variables on Ω

$$X \in L^1(\Omega) \Leftrightarrow \sum_{x \in \mathbb{R}} |x| p_X(x) < +\infty$$

Note $Y \geq 0$ discrete

$E[Y]$ is always defined, $E[Y] \in [0, +\infty]$

Example $X \sim \text{Geo}(p)$

GEOMETRIC r.v. of parameter $p \in [0, 1]$

$$P(X=k) = (1-p)^k p \quad k \in \mathbb{N}, k=0, 1, \dots$$

$X \geq 0 \rightarrow E[X] \text{ exists}$

$$E[X] = \sum_{k=0}^{\infty} k (1-p)^k p = \frac{1}{p}$$