

Quantifying the Influence of Climate on Human Conflict

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A rapidly growing body of research examines whether human conflict can be affected by climatic changes. Drawing from archeology, criminology, economics, geography, history, political science, and psychology, we assemble and analyze the 60 most rigorous quantitative studies and document, for the first time, a remarkable convergence of results. We find strong causal evidence linking climatic events to human conflict across a range of spatial and temporal scales and across all major regions of the world. The magnitude of climate's influence is substantial: for each 1 standard deviation (1σ) change in climate toward warmer temperatures or more extreme rainfall, median estimates indicate that the frequency of interpersonal violence rises 4% and the frequency of intergroup conflict rises 14%. Because locations throughout the inhabited world are expected to warm 2-4°C by 2050, amplified rates of human conflict could represent a large and critical impact of anthropogenic climate change.

Human behavior is complex, and despite the existence of institutions designed to promote peace, interactions between individuals and groups sometimes lead to conflict. When such conflict becomes violent, it can have dramatic consequences on human wellbeing. Mortality alone from war and interpersonal violence amounts to 0.5-1 million deaths annually (1, 2), with non-lethal impacts including injury and lost economic opportunities affecting millions more. Because the stakes are so high, understanding the causes of human conflict has been a major project in the social sciences.

Researchers working across multiple disciplines including archeology, criminology, economics, geography, history, political science, and psychology have long debated the extent to which climatic changes are responsible for causing conflict, violence or political instability. Numerous pathways linking the climate to these outcomes have been proposed. For example, climatic changes may alter the supply of a resource and cause disagreement over its allocation, or climatic conditions may shape the relative appeal of using violence or cooperation to achieve some preconceived objective. Qualitative researchers have a well-developed history of studying these issues (3-7) dating back at least to the start of the twentieth century (8). Yet in recent years, growing recognition that the climate is changing, coupled with improvements in data quality and computing, have prompted an explosion of quantitative analyses seeking to test these theories and quantify the strength of these previously proposed linkages. Thus far, this work has remained scattered across multiple disciplines and has been difficult to synthesize given the disparate methodologies, data and interests of the various research teams.

Here we assemble the first comprehensive synthesis of this rapidly growing quantitative literature. We adopt a broad definition of "conflict," using the term to encompass a range of outcomes from individual-level violence and aggression to country-level political instability and

civil war. We then collect all available candidate studies and - guided by previous criticisms that not all correlations imply causation (9-11) - focus only on those quantitative studies that can reliably infer causal associations (9, 12) between climate variables and conflict outcomes. The studies we examine exploit either experimental or natural-experimental variation in climate, where the latter term refers to variation in climate over time that is plausibly independent of other variables that also affect conflict. To meet this standard, studies must account for unobservable confounding factors across populations, as well as for unobservable time-trending factors that could be correlated with both climate and conflict (13). In many cases we obtained data from studies that did not meet this criteria and reanalyzed it using a common statistical model that met this criteria (see supplementary materials). The importance of this rigorous approach is highlighted by an example where our standardized analysis generated findings consistent with other studies but at odds with the original conclusions of the study in question (14).

In total, we obtained 60 primary studies that either met this criteria or were reanalyzed with a method that met

this criteria (Table 1). Collectively, these studies analyze 45 different conflict data sets published across 26 different journals and represent the work of over 190 researchers from around the world. Our evaluation summarizes the recent explosion of research on this topic, with 78% of studies released since 2009 and the median study released in 2011. We collect findings across a wide range of conflict outcomes, across time periods spanning 10,000 B.C.E. to the present day, and across all major regions of the world (Fig. 1).

While various conflict outcomes differ in important ways, we find that the behavior of these outcomes relative to the climate system is remarkably similar. Put most simply, we find that large deviations from normal precipitation and mild temperatures systematically increase the risk of many types of conflict, often substantially, and that this relationship appears to hold over a variety of temporal and spatial scales. Our meta-analysis of studies that examine populations in the post-1950 era suggest that these relationships continue to be highly significant in the modern world - although there are important differences in the magnitude of the relationship when different variables are considered: the standardized effect of temperature is generally larger than the standardized effect of rainfall and the effect on intergroup violence (e.g., civil war) is larger than the effect on interpersonal violence (e.g., assault). We conclude that there is substantially more agreement and generality in the findings of this burgeoning literature than has been previously recognized. Given the large potential changes in precipitation and temperature regimes projected in coming decades, our findings have important implications for the social impact of anthropogenic climate change in both low- and high-income countries.

Estimation of Climate-Conflict Linkages

Reliably measuring an effect of climatic conditions on human conflict is complicated by the inherent complexity of social systems. In particular, a central concern is whether statistical relationships can be interpreted causally or if they are confounded by omitted variables. To address this concern, we restrict our attention to studies with research designs that are a scientific experiment or that approximate one (i.e., “natural experiments”). After describing how studies meet this criteria, we discuss how we interpret the precision of results, how we assess the “importance” of climatic factors, and how we address choices over functional form.

Research Design

In an ideal experiment, we would observe two identical populations, change the climate of one, and observe whether this “treatment” lead to more or less conflict relative to the “control” conditions. Because the climate cannot be experimentally manipulated, researchers primarily rely on natural experiments where a given population is compared to itself at different moments in time when it is exposed to different climatic conditions - conditions which are exogenously determined by the climate system (9, 15). In this research design, a single population serves as both the “control” population - e.g., just before a change in climatic conditions- and the “treatment” population - e.g., just after a change in climatic conditions. Inferences are thus based only on how a fixed population responds to different climatic conditions which vary over time, and time-series or longitudinal analysis is used to construct a credible estimate for the causal effect of climate on conflict (12, 15, 16).

To minimize statistical bias and to improve the comparability of studies, we focus on studies that use versions of the general model

$$\text{conflict_variable}_{it} = \beta \times \text{climate_variable}_{it} + \mu_i + \theta_t + \epsilon_{it} \quad (1)$$

where locations are indexed by i , observational periods are indexed by t , β is the parameter of interest and ϵ is the error. If different locations in a sample exhibit different average levels of conflict - perhaps because of cultural, historical, political, economic, geographic or institutional differences between the locations - this will be accounted for by the location-specific constants μ (commonly known as “fixed effects”). Time-specific constants θ (a dummy for each time period) flexibly account for other time-trending variables such as economic growth or gradual demographic changes that could be correlated with both climate and conflict. In some cases, such as in time series, the θ_t parameters may be replaced by a generic trend (e.g., $\theta \times t$) which is possibly nonlinear and is either common to all locations or may be location-specific (e.g., $\theta_i \times t$). Our conclusions from the literature are based only on those studies that implement Eq. 1 or one of the mentioned alternatives. In select cases, when studies did not meet this criteria but the data from these analyses were publicly available or supplied by the authors, we reanalyzed the data using this common method (see supplementary materials). Many estimates of Eq. 1 in the literature and in our reanalysis account for temporal and/or spatial autocorrelation in the error term ϵ , although this adjustment was not considered a requirement for inclusion here. In the case of some paleoclimatological/archeological studies, formal statistical analysis is not implemented because the outcome variables of interest are essentially singular cataclysmic events; however, we include these studies because they follow populations over time at a fixed location and are thus implicitly using the model in Eq. 1 (these cases are noted in Table 1).

We do not consider studies that are purely cross-sectional, i.e., studies that only compare rates of conflict across different locations and that attribute differences in average levels of conflict to average climatic conditions. There are many ways in which populations differ from one another (culture, history, etc.), many of them unobserved, and these

“omitted variables” are likely to confound these analyses. In the language of the natural experiment, the “treatment” and “control” populations in these analyses are not comparable units, so we cannot infer whether a climatic “treatment” has a causal effect or not (12, 13, 15–17). For example, a cross-sectional study might compare average rates of civil conflict in Norway and Nigeria, attributing observed differences to the different climate of these countries - despite the fact that there are clearly many other relevant ways in which these countries differ. Nonetheless, some studies use cross-sectional analyses and attempt to control for confounding variables in regression analyses, typically using a handful of covariates such as average income or political indices. However, because the full suite of determinants of conflict are unknown and unmeasured, it is likely impossible that any cross-sectional study can explicitly account for all important differences between populations. Rather than presuming that all confounders are accounted for, the studies we evaluate only compare Norway or Nigeria to themselves at different moments in time, thereby ensuring that the structure, history and geography of comparison populations are nearly identical.

Some studies implement versions of Eq. 1 that are expanded to explicitly “control” for potential confounding factors, such as average income. In many cases this approach is more harmful than helpful because it introduces bias in the coefficients describing the effect of climate on conflict. This problem occurs when researchers control for variables that are themselves affected by climate variation, causing either (i) the signal in the climate variable of interest to be inappropriately absorbed by the “control” variable, or (ii) the estimate to be biased because populations differ in unobserved ways that become artificially correlated with climate when the “control” variable is included. This methodological error is commonly termed “bad control” (12) and we exclude results obtained using this approach. The difficulty in this setting is that climatic variables affect many of the socioeconomic factors commonly included as control variables - things like crop production, infant mortality, population (via migration or mortality), and even political regime type. To the extent that these outcome variables are used as controls in Eq. 1, studies might draw mistaken conclusions about the relationship between climate and conflict. Because this error is so salient in the literature, we provide examples below. A full treatment can be found in (12, 18).

For an example of (i), consider whether variation in temperature increases conflict. In many studies of conflict, researchers often employ a “standard” set of controls which are correlates of conflict, such as per capita income. However, evidence suggests that income is itself affected by temperature (19–21), so if part of the effect of temperature on conflict is through income, then “controlling” for income in Eq. 1 will lead the researcher to underestimate the role of temperature in conflict. This occurs because much of the effect of temperature will be absorbed by the income variable, biasing the temperature coefficient toward zero. At the extreme, if temperature influences conflict only through income, then controlling for income would lead the researcher in this example to draw exactly the wrong conclusion about the relationship between temperature and conflict: that there is no effect of temperature on conflict.

For an example of (ii), imagine that a measure of “politics” and temperature both have a causal effect on conflict and that both politics and temperature have an effect on income, but that income has no effect on conflict. If politics and temperature are uncorrelated, estimates of Eq. 1 that do not control for politics will still recover the unbiased effect of temperature. However, If income is introduced to Eq. 1 as a “control” while politics is left out of the model, perhaps because it is more difficult to measure, then it will appear as if there is an association between income and conflict because income will be serving as a proxy measure for politics. In addition, this adjustment to Eq. 1 also biases the estimated effect of temperature. This bias occurs because the types of countries that have high income when temperature is high are different, in terms of their average politics, than those countries that have high income when

temperature is low. Thus, if income is “held fixed” as a control variable in a regression model, the comparison of conflict across temperatures is not an apples-to-apples comparison because politics will be systematically different across countries at different temperatures, generating a bias that can have either sign. In this example, the inclusion of income in the model leads to two incorrect conclusions: it biases the estimated relationship between climate and conflict and it implicates income as playing a role in conflict when it does not.

Statistical Precision

We consider each study’s estimated relationship between climate and conflict as well as the estimate’s precision. Because sampling variability and sample sizes differ across studies, some analyses present results that are more precise than other studies. Recognizing this fact is important when synthesizing a diverse literature, as some apparent differences between studies can be reconciled by evaluating the uncertainty in their findings. For example, some studies report associations that are very large or very small but with uncertainties that are also very large, leading us to place less confidence in these extreme findings. This intuition is formalized in our meta-analysis which aggregates results across studies by down-weighting results that are less precisely estimated.

The strength of a finding is sometimes summarized in a statement regarding its “statistical significance,” which describes the signal-to-noise ratio in an individual study. However, in principle the “signal” is a relationship that exists in the real world and cannot be affected by the researcher, whereas the level of “noise” in a given study’s finding (i.e., its uncertainty) is a feature specific to that study - a feature that can be affected by a researcher’s decisions, such as the size of the sample they choose to analyze. Thus, while it is useful to evaluate whether individual findings are statistically significant and it is important to down-weight highly imprecise findings, individual studies provide useful information even when they are not statistically significant.

To summarize the evidence that each statistical study provides while also taking into account its precision, we separately consider three questions for each study in Table 1: (1) Is the estimated average effect of climate on conflict quantitatively “large” in magnitude (discussed below), regardless of its uncertainty? (2) Is the reported effect large enough and estimated with sufficient precision that the study can reject the null hypothesis of “no relationship” at the 5% level? (3) If the study cannot reject the hypothesis of “no relationship,” can it reject the hypothesis that the relationship is quantitatively large? In the literature, often only question 2 is evaluated in any single analysis. Yet it is important to consider the magnitude of climate influence (question 1) separately from its statistical precision because the magnitude of these effects tell us something about the potential importance of climate as a factor that may influence conflict, so long as we are mindful that evidence is weaker if a study’s results are less certain. In cases where the estimated effect is smaller in magnitude and not statistically different than zero, it is important to consider whether a study provides strong evidence of zero association - i.e., the study rejects the hypothesis that an effect is large in magnitude (question 3) - or relatively weak evidence because the estimated confidence interval spans large effects as well as spanning zero effect.

Evaluating If an Effect Is “Important”

Evaluating whether an observed causal relationship is “important” is a subjective judgement that is not essential to our scientific understanding of whether there is a causal relationship. Nonetheless, because “importance” in this literature has sometimes been incorrectly conflated with statistical precision or inferred from incorrect interpretations of Eq. 1 and its variants, we explain our approach to evaluating importance.

Our preferred measure of importance is to ask a straightforward

question: Do changes in climate cause changes in conflict risk that an expert, policy-maker or citizen would consider large? To aid comparisons, we operationalize this question by considering an effect important if authors of a particular study state the size of the effect is substantive, or if the effect is greater than a 10% change in conflict risk for each 1 standard deviation (1σ) change in climate variables. This second criteria uses an admittedly arbitrary threshold, and other threshold selections would be justifiable. However, we contend this threshold is relatively conservative since most policy-makers or citizens would be concerned by effects well below $10\%/ \sigma$. For instance, since random variation in a normally distributed climate variable lies in a 4σ range for 95% of its realizations, even a $3\%/ \sigma$ effect size would generate variation in conflict of 12% of its mean, which is probably important to those individuals experiencing these shifts.

In some prior studies, authors have argued that a particular estimated effect is “unimportant” based on whether a climatic variable substantially changes goodness-of-fit measures (e.g., R^2) for a particular statistical model, sometimes in comparison to other predictor variables (14, 22–24). We do not use this criteria here for two reasons. First, goodness-of-fit measures are sensitive to the quantity of noise in a conflict variable: more noise reduces goodness-of-fit - thus, under this metric, irrelevant measurement errors that introduce noise into conflict data will reduce the apparent “importance” of climate as a cause of conflict, even if the effect of climate on conflict is quantitatively large. Second, comparing the goodness-of-fit across multiple predictor variables often makes little sense in many contexts since (i) longitudinal models typically compare variables that predict both *where* a conflict will occur and *when* a conflict occur and (ii) these models typically compare the causal effect of climatic variables with the non-causal effects of confounding variables, such as endogenous covariates. These are apples-to-oranges comparisons and the faulty logic of both types of comparison are made clear with examples.

For an example of (i), consider an analyst comparing violent crime over time in New York City and North Dakota who finds that the number of police on the street each day are important for predicting how much crime occurs on that day, but that a population variable describes more of the variation in crime since crime and population in North Dakota are both low. Clearly this comparison is not informative, since the reason that there is little crime in North Dakota has nothing to do with the reason why crime is lower in New York City on days when there are many police on the street. The argument that variations in climate are “not important” to predicting when conflict occurs because other variables are good predictors of where conflict occurs is analogous to the strange statement that the number of police in New York City are “not important” for predicting crime rates because North Dakota has lower crime that is attributable to its lower population.

For an example of (ii), suppose that both higher rainfall and higher household income lower the likelihood of civil conflict, but household income is not observed and instead a variable describing the average observable number of cars each household owns is included in the regression. Because wealthier households are better able to afford cars, the analyst finds that populations with more cars have a lower risk of conflict. This relationship clearly does not have a causal interpretation and comparing the “effect of car ownership on conflict” with the effect of rainfall on conflict does not help us better understand the importance of the rainfall variable. Published studies that make similar comparisons do so with variables that the authors suggest are more relevant than cars, but the uninformative nature of comparisons between causal effects and non-causal correlations is the same.

Functional Form and Evidence of Nonlinearity

Some studies assume a linear relationship between climatic factors

and conflict risk, while others assume a non-linear relationship. Taken as a whole, the evidence suggests that over a sufficiently large range of temperatures and rainfall levels, both temperature and precipitation appear to have a non-linear relationship with conflict, at least in some contexts. However, this curvature is not apparent in every study, probably because the range of temperatures or rainfall levels contained within a sample may be relatively limited. Thus, most studies report only linear relationships that should be interpreted as local linearizations of a more complex - and possibly curved - response function.

As we will show, all modern analyses that address temperature impacts find that higher temperatures lead to more conflict. However, a few historical studies that examine temperate locations during cold epochs do find that abrupt cooling from an already cold baseline temperature may lead to conflict. Taken together, this collection of locally linear relationships indicates a global relationship with temperature that is non-linear.

In studies of rainfall impacts, the distinction between linearity and curvature is made fuzzy by the multiple ways that rainfall changes have been parameterized in existing studies. Not all studies use the same independent variable, and because a simple transformation of an independent variable can change the response function from curved to linear and visa versa, this makes it difficult to determine whether results agree. In an attempt to make findings comparable, when replicating the studies that originally examine a non-linear relationship between rainfall and conflict we follow the approach of Hidalgo *et al.* (25) and use the absolute value of rainfall deviations from the mean as the independent variable; in studies that originally examined linear relationships we leave the independent variable unaltered. Because these two approaches in the literature (and our reanalysis) differ, we make the distinction clear in our figures through the use of two different colors.

Results from the Quantitative Literature

We divide our review topically, examining in turn the evidence on how climatic changes shape personal violence, group-level violence, and the breakdown of social order and political institutions. Results from twelve example studies of recent data (post-1950) are displayed in Fig. 2, which we replicated using the common statistical framework described above, and which were chosen to represent a broad cross section of outcomes, geographies, and time periods (see supplementary materials). Findings from several studies of historical data are collected in Fig. 3, where the different time scales of climatic events can be easily compared. A listing and description of all primary studies are in Table 1. For a detailed description and evaluation of each individual study, we refer readers to (26).

Personal Violence and Crime

Studies in psychology and economics have repeatedly found that subjects are more likely to exhibit aggressive or violent behavior toward others if ambient temperatures at the time of observation are higher (Fig. 2, A to C), a result that has been obtained in both experimental (27, 28) and natural-experimental (29–39) settings. Documented aggressive behaviors that respond to temperature range from somewhat less consequential - e.g., horn-honking while driving (27) and inter-player violence during sporting events (36) - to much more serious - e.g., the use of force during police training (28), domestic violence within households (29, 37), and violent crimes such as assault or rape (30–35, 38). Although the physiological mechanism linking temperature to aggression remains unknown, the causal association appears robust across a variety of contexts. Importantly, because aggression at high temperature increases the likelihood that intergroup conflicts escalate in some contexts (36) and the likelihood that police officers use force (28), it is possible that this mechanism could affect the prevalence of larger scale group-level conflicts.

In low-income settings, extreme rainfall events that adversely affect agricultural income are also associated with higher rates of personal violence (40–42) and property crime (43). High temperatures are also associated with increased property crime (34, 35, 38), but violent crimes appear to rise with temperature more quickly than property crimes (38).

Group-Level Violence and Political Instability

Some forms of intergroup violence, such as Hindu-Muslim riots (Fig. 2D), tend to be more likely following extreme rainfall conditions (44–47). This relationship between intergroup violence and rainfall is primarily documented in low-income settings, suggesting that reduced agricultural production may be an important mediating mechanism - although alternative explanations cannot be excluded.

Low water availability (23, 46, 48–57), very low temperatures (58–63) and very high temperatures (14, 21, 23, 51, 64–66) have been associated with organized political conflicts in a variety of low-income contexts (Fig. 2, E, F, H, I, K, and L). The structure of this relationship again seems to implicate a pathway through climate-induced changes in income, either agricultural (48, 67–69) or non-agricultural (20, 21), although this hypothesis remains speculative. Large deviations from normal precipitation have also been shown to lead to the forceful reallocation of wealth (25) (Fig. 2G) or the non-violent replacement of incumbent leaders (70, 71) (Fig. 2J).

Some authors recently suggested that contradictory evidence is widespread among quantitative studies of climate and human conflict (72–74), but the level of disagreement appears overstated. Two studies (22, 24) estimate that temperature and rainfall events have a limited impact on civil war in Africa, but the confidence intervals around these estimates are sufficiently wide that they do not reject a relatively large effect of climate on conflict that is consistent with 35 other studies of modern data and 28 other studies of inter-group conflict. Within the broader literature of primary statistical studies, these results represent 4% of all reported findings (Table 1). Isolated studies also suggest that windstorms and floods have limited observable effect on civil conflicts (75) and that anomalously high rainfall is associated with higher incidence of terrorist attacks (76).

Institutional Breakdown

Under sufficiently high levels of climatological stress, pre-existing social institutions may strain beyond recovery and lead to major changes in governing institutions (77–79) (Fig. 3C), a process that often involves the forcible removal of rulers. High levels of climatological stress have also led to major changes in settlement patterns and social organization (80, 81) (Fig. 3D). Finally, in extreme cases, entire communities, civilizations and empires collapse entirely following large changes in climatic conditions (62, 79, 80, 82–89) (Fig. 3, A to C, E, and F). These documented catastrophic failures all precede the twentieth century, yet the level of economic development in these communities at the time of their collapse was similar to the level of development in many poor countries of the modern world [see (26) for a comparison], an indicator that these historical cases may continue to have modern relevance.

Synthesis of Findings

Once attention is restricted to those studies able to make rigorous causal claims about the relationship between climate and conflict, some general patterns become clear. Here we identify, for the first time, commonalities across results that span diverse social systems, climatological stimuli and research disciplines.

Generality: Samples, Spatial Scales, and Rates of Climate Change

Social conflicts at all scales and levels of organization appear susceptible to climatic influence, and multiple dimensions of the climate

system are capable of influencing these various outcomes. Studies documenting this relationship can be found in data samples covering 10,000 BCE to the present and this relationship has been identified multiple times in each major region, as well as in multiple samples with global coverage (Fig. 1A).

Climatic influence on human conflict appears in both high and low income societies, although some types of conflict, such as civil war, are rare in high income populations do not exhibit a strong dependence on climate in those regions (51). Nonetheless, many other forms of conflict in high income countries such as violent crime (35, 38), police violence (28), or leadership changes (71), do respond to climatic changes. These forms of conflict are individually less extreme, but their total social cost may be large because they are widespread. For example, during 1979-2009 there were more than two million violent crimes (assault, murder and rape) per year on average in the United States alone (38), so small percentage changes can lead to substantial increases in the absolute number of these types of events.

Climatic perturbations at spatial scales ranging from a building (27, 28, 36) to the globe (51) have been found to influence human conflict or social stability (Fig. 1B). The finding that climate influences conflict across multiple scales suggests that coping or adaptation mechanisms are often limited. Interestingly, as shown in Fig. 1B, there is a positive association between the temporal and spatial scales of observational units in studies documenting a climate-conflict link. This might indicate that larger social systems are less vulnerable to high frequency climate events, or it may be that higher-frequency climate events are more difficult to detect in studies examining outcomes over wide spatial scales.

Finally, it is sometimes argued that societies are particularly resilient to climate perturbations of a specific temporal scale - perhaps they are capable of buffering themselves against short-lived climate events, or alternatively that they are able to adapt to conditions that are persistent. With respect to human conflict, the available evidence does not support either of these claims. Climatic anomalies of all temporal durations, from the anomalous hour (28) to the anomalous millennium (81), have been implicated in some form of human conflict (Fig. 1B).

The association between climatic events and human conflict is general in the sense that it has been observed almost everywhere: across types of conflict, across human history, across regions of the world, across income groups, across the various durations of climatic changes, and across all spatial scales. However, it is *not* true that all types of climatic events influence all forms of human conflict or that climatic conditions are the sole determinant of human conflict. The influence of climate is detectable across contexts, but we strongly emphasize that it is only one of many factors that contribute to conflict [see (90) for a review of these other factors].

The Direction and Magnitude of Climatic Influence on Human Conflict

We must consider the magnitude of the climate's influence in order to evaluate whether climatic events play an important role in the occurrence of conflict, and whether anthropogenic climate change has the potential to substantially alter future conflict outcomes. Quantifying the magnitude of climatic impact in archeological/paleo-climatological studies is difficult because outcomes of interest are often one-off cataclysmic events (e.g., societal collapse) and we typically do not observe how the universe of societies would have responded to similar sized shocks. Modern data samples, however, generally contain a large number of comparable social units (e.g., countries) that are repeatedly exposed to climatic variation, and this setting that is more amenable to statistical analyses that quantify how changes in climate affect the risk of conflict within an individual social unit.

To compare quantitative results across studies of modern data, we

computed standardized effect sizes for those studies where it was possible to do so, evaluating the effect of a 1 σ change in the explanatory climate variable and expressing the result as a percentage change in the outcome variable. Because we restrict our attention to studies that examine changes in climate variables over time, the relevant standard deviation is based only on inter-temporal changes at each specific location instead of comparing variation in climate across different geographic locations.

Our results are displayed in Figs. 4 and 5 (colors match Figs. 2 and 3). Nearly all studies suggest that warmer temperatures, lower or more extreme rainfall, or warmer El Niño-Southern Oscillation (ENSO) conditions lead to a 2-40% increase in the conflict outcome per 1 σ in the observed climate variable. The consistent direction of temperature's influence is particularly remarkable since all 27 modern estimates (including ENSO and temperature-based drought indices, 20 estimates are shown in Figs. 4 and 5) indicate that warmer conditions generate more conflict, a result that would be extremely unlikely to occur by chance alone if temperature had no effect on conflict. It is more difficult to interpret whether the sign of rainfall-related variables agree because these variables are parametrized several different ways, so Figs. 4 and 5 present likelihoods for different parameterizations separately. However, if all modern rainfall estimates are pooled (including ENSO and rainfall-based drought indices, 13 estimates are shown in Figs. 4 and 5) using signs shown in Figs. 4 and 5, then the sign of the effect in 16 out of 18 estimates agree.

Under the assumption that there is some underlying similarity across studies, we compute the average effect of climate variables across studies by weighting each estimate according to its precision (the inverse of the estimated variance), a common approach that penalizes uncertain estimates (91). We also calculate the confidence interval on this mean by assuming independence across studies, although this assumption is not critical to our central findings (in the supplementary materials we present results where we relax this assumption and show that it is not essential). The precision-weighted average effect on interpersonal conflict is a 2.3% increase for each 1 σ change in climatic variables (s.e.= 0.12%, $p < 0.001$, Fig. 4 and table S1) and the analogous estimate for intergroup conflict is 11.1% (s.e.= 1.3%, $p < 0.001$, Fig. 5 and table S1). These precision-weighted averages are relatively un-influenced by outliers since outlier estimates in our sample tend to have low precision and thus low weight in the meta-analysis. The corresponding medians, which are also insensitive to outliers, are comparable: 3.9% for personal conflict and 13.6% for group conflict. If we restrict our attention to only the effects of temperature, the precision-weighted average effect is similar for interpersonal conflict (2.3%), but for intergroup conflict rises to 13.2% per 1 σ in temperature (s.e.= 2.0, $p < 0.001$, Fig. 5). Regarding the interpretation of these effect sizes, we note that while the average effect for interpersonal violence is smaller than the average effect for intergroup conflict in percentage terms, the baseline number of incidents of interpersonal violence is dramatically higher, meaning a small percentage increase can represent a substantial increase in total incidents.

We estimate the precision-weighted probability distribution of study-level effect-sizes in Figs. 4 and 5 and in table S1. These distributions are centered at the precision-weighted averages described above and can be interpreted as the distribution of results from which studies' findings are drawn. The distribution for interpersonal conflict is narrow around its mean, likely because most interpersonal conflict studies focus on one country (the United States) and use very large samples and derive very precise estimates. The distribution for intergroup conflict is broader and covers values that are larger in magnitude, with an interquartile range 6 to 14% per 1 σ and the 5-95th percentiles spanning -5 to 32% per 1 σ (table S1). We estimate that for the intergroup and interpersonal conflict studies, respectively, 10% and 0% of the probability mass of the distributions of effect sizes lies below zero.

Figures 4 and 5 make it clear that even though there is substantial agreement across results, some heterogeneity across estimates remains. It is possible that some of this variation is meaningful, perhaps because different types of climate variables have different impacts or because the social, economic, political or geographic conditions of a society mediate its response to climatic events. For instance, poorer populations appear to have larger responses, consistent with prior findings that such populations are more vulnerable to climatic shifts (51). However, it is also possible that some of this variation is due to differences in how conflict outcomes are defined, to measurement error in climate variables, or to remaining differences in model specifications that we could not correct in our reanalysis.

To formally characterize the variation in estimated responses across studies, we use a Bayesian hierarchical model that does not require knowledge of the source of between-study variation (92) (see supplementary materials). Under this approach, estimates of the precision-weighted mean are essentially unchanged, and we recover estimates for the between-study standard deviation (a measure of the underlying dispersion of “true” effect sizes across studies) that are half of the precision-weighted mean for interpersonal conflict, and two-thirds of the precision-weighted mean for intergroup conflict (median estimates; see supplementary materials, fig. S3 and tables S2 and S3). By comparison, if variation in effect sizes across studies was driven by sampling variation alone, then this standard deviation in the underlying distribution of effect sizes would be zero. This suggests “true” effects likely differ across settings, and understanding this heterogeneity should be a primary goal of future research.

Publication Bias

Publication bias is a longstanding concern across the sciences, with a common form of bias arising from the research community’s perceived preference for positive rather than null results. Although it is always possible that publication bias played a role in the publication of a specific analysis, there are multiple reasons why publication bias is unlikely to be driving our findings about the literature on climate and conflict. First, we include working papers in our analysis (as is common practice in the social sciences), thereby eliminating editorial selection. Second, the central results presented here are replicated in multiple disciplines and across diverse samples. Third, the large number of positive findings present in the literature since 2009 could provide limited professional incentive for researchers to publish yet another positive finding, and benefits might be higher to those who publish results with alternative findings. Fourth, many analyses are not explicitly focused on the direct effect of climate on conflict but instead use climatic variations instrumentally (25, 35, 48, 71, 77) or account for it as an ancillary covariate in their analysis [e.g., (37)] while trying to study a different research question - indicating that these authors have little professional stake in the sign, magnitude or statistical significance of the climatic effects they are presenting. Fifth, we reanalyze the raw data from many studies using a common statistical framework, possibly “undoing” adjustments that authors might be making to their analysis (consciously or unconsciously) that make their findings appear stronger - partial support for this idea is provided by individual studies that present significant results, but whose results are only marginally significant or no longer significant after our reanalysis (see supplementary materials for details). Finally, we look for evidence of publication bias by examining whether the statistical strength of individual studies reflects their sample size (93) and do not find systematic evidence of strong bias in absolute terms or in comparison to other social science literatures (see fig. S4, table S4, and supplementary materials).

Implications for Future Climatic Changes

The above evidence makes a *prima facie* case that future anthropogenic climate change could worsen conflict outcomes across the globe in comparison to a future with no climatic changes, given the large expected increase in global surface temperatures and the likely increase in variability of precipitation across many regions over coming decades (94, 95). Recalling our finding that a 1σ change in a location’s temperature is associated with an average 2.3% increase in the rate of interpersonal conflict and a 13.2% increase in the rate of intergroup conflict, and assuming that future populations will respond to climatic shifts similarly to how current populations respond, one can consider the potential effect of anthropogenic warming by rescaling expected temperature changes according to each location’s historical variability. While not all conflict outcomes have been shown responsive to changes in temperature, many have, and the results uniformly indicate that increasing temperatures are harmful in regions that are temperate or warm initially. In Fig. 6 we plot expected warming by 2050, computed as the ensemble mean for 21 climate models running the A1B emissions scenario, in terms of location-specific standard deviations (96). Almost all inhabited locations warm by $> 2\sigma$, with the largest increases exceeding 4σ in tropical regions that are already warm and currently experience relatively low inter-annual temperature variability. These large climatological changes, combined with the quantitatively large effect of climate on conflict - particularly intergroup conflict - suggest that amplified rates of human conflict could represent a large and critical impact of anthropogenic climate change.

Two reasons are often given as to why climate change might not have a substantive impact on human conflict: future climate change will occur gradually and will thus allow societies to adapt, and the modern world today is less susceptible to climate variation than it has been in the past. However, if slower-moving climate shocks have smaller effects, or if the world has become less climate sensitive, it is unfortunately not obvious in the data. Gradual climatic changes appear to adversely affect conflict outcomes, and the majority of the studies we review use a sample period that extends into the 21st century (recall Fig. 1). Furthermore, some studies explicitly examine whether populations inhabiting hotter climates exhibit less conflict when hot events occur, but find little evidence that these areas are more adapted (31, 38). We also note that many of the modern linkages between high temperature anomalies and intergroup conflict have been characterized in Africa (14, 23, 52, 64, 66) or the global tropics and subtropics (21, 51), regions with hot climates where we would expect populations to be best adapted to high temperatures. Nevertheless, it is always possible that future populations will adapt in previously unobserved ways, but it is impossible to know if and to what extent these adaptations will make conflict more or less likely.

Studies of non-conflict outcomes do indicate that in some situations, historical adaptation to climate is observable, albeit costly (97–100), while in other cases there is limited evidence that any adaptation is occurring (19, 101). To our knowledge, no study has characterized the scale or scope for adaptation to climate in terms of conflict outcomes, and we believe this is an important area for future research. Given the quantitatively large effect of climate on conflict, future adaptations will need to be dramatic if they are to offset the potentially large amplification of conflict.

Future Research

Given the remarkable consistency of available quantitative evidence linking climate and conflict, in our view the top research priority in this field should be to narrow the number of competing explanatory hypotheses. Beyond efforts to mitigate future warming, limiting climate’s future influence on conflict requires that we understand the causal pathways that generate the observed association. This task is made difficult by the likely situation that multiple mechanisms contribute to the observed relationships and that different mechanisms dominate in different con-

texts. The rich qualitative literature (3–7) suggest that a multiplicity of mechanisms may be at work.

To date, no study has been able to conclusively pin down the full set of causal mechanisms, although some studies find suggestive evidence that a particular pathway contributes to the observed association in a particular context. In most cases, this is accomplished by “fingerprinting” the effect of climate on an intermediary variable, such as income, and showing that the same statistical fingerprint is visible in the climate's effect on conflict. This approach - typically called “instrumental variables” (12) in the social sciences - identifies a mechanism linking climate and conflict under the assumption that climate's only influence on conflict is through the particular intermediate variable in question. Because this assumption is often difficult or impossible to test, evidence from this approach is more suggestive than conclusive in uncovering mechanisms (51).

An alternate and promising research design that can help rule out certain hypotheses is to study situations where plausibly exogenous events block a proposed pathway in a “treated” subpopulation and then to compare whether the climate-conflict association persists or disappears in both the treatment and control subpopulations. An example of this approach, Sarsons (2011) examines whether rainfall shortages in India lead to riots because they depress local agricultural income (45). By showing that rainfall shortages and riots continue to occur together in districts with dams that supply irrigation, investments that partially decouple local agricultural income from temporary rain shortfalls, Sarsons argues that the rainfall effect on riots is unlikely to be operating solely through changes in local agricultural income.

Plausible Mechanisms

The following hypotheses have, in our judgement, received the strongest empirical support in existing analyses, although the evidence is still often inconclusive. A common hypothesis focuses on *local economic conditions and labor markets*, and argues that when climatic events cause economic productivity to decline (19–21, 68, 69, 102–104), the value of engaging in conflict is likely to rise relative to the value of participating in normal economic activities (48, 52, 105–110). A competing hypothesis on *state capacity* argues that these declines in economic productivity reduce the strength of governmental institutions (e.g., if tax revenues fall), curtailing their ability to suppress crime and rebellion or encouraging competitors to initiate conflict during these periods of relative state weakness (61, 70, 71, 77–79, 84, 85).

A second set of hypotheses focus on what has more generally been termed “grievances”. Hypotheses about *inequality* contend that when climatic events increase actual (or perceived) social and economic inequalities in a society (111, 112), this could increase conflict by motivating attempts to redistribute assets (25, 34, 35, 43). Evidence linking changes in *food prices* to conflict (61, 113–115) can be interpreted similarly - e.g., food riots due to a government's perceived inability to keep food affordable - particularly when some members of society can influence food markets (111, 116).

Climate-induced *migration and urbanization* might also be implicated in conflict. If climatic events cause large population displacements or rapid urbanization (97, 117, 118), this might lead to conflicts over geographically stationary resources that are unrelated to the climate (119) but become relatively scarce where populations concentrate. Changes in climate might also affect the *logistics* of human conflict (76, 120), for example by altering the physical environment (e.g., road quality) in which disputes or violence might occur (52, 120, 121). Finally, climate anomalies might result in conflict because they can make *cognition and attribution* more difficult or error-prone, or they may affect *aggression* through some physiological mechanism. For instance, climatic events may alter individuals' ability to reason and correctly interpret events (27,

28, 30, 31, 34–36), possibly leading to conflicts triggered by misunderstandings. Alternatively, if climatic changes and their economic consequences are inaccurately attributed to the actions of an individual or group (63, 122–125), for example an inept political leader (71), this may lead to violent actions that try to return economic conditions to normal by removing the “offending” population.

Selecting Climate Variables and Conflict Outcomes

Climate variables that have been previously analyzed, such as seasonal temperatures, precipitation, water availability indices, and climate indices, may be correlated with one another and autocorrelated across both time and space. For instance, temperature and precipitation time-series tend to be negatively correlated in much of the tropics and drought indices tend to be spatially correlated (51, 126). Unfortunately, only a few of the existing studies account for the correlations between different variables, so it may be that some studies mistakenly measure the influence of an omitted climate variable by proxy [see (126) for a complete discussion of this issue]. Except for the experiments linking temperature to aggression (27, 28), only a few studies demonstrate that a specific climate variable is more important for predicting conflict than other climate variables or that climatic changes during a specific season are more important than during other seasons. Furthermore, no study isolates a particular type of climatic change as the most influential and no study has identified whether temporal or spatial autocorrelations in climatic variables are mechanistically important. Identifying the climatic variables, timing of events and forms of autocorrelation that influence conflict will help us better understand the mechanisms linking climatic changes to conflict.

A similar situation exists with the choice of conflict outcomes. Most analyses simply document changes in the rate at which conflicts are reported in aggregate, but this approach provides only limited insight into how the evolution of conflict is impacted by climatic variables. A path for future investigation is to link climate data with richer conflict data that describes different stages of the conflict “lifecycle.” For example, future studies could examine how often non-violent group disputes become violent. Two studies in this review (28, 36) demonstrate the usefulness of selecting conflict-variables other than total conflict rates. By examining the probability that an initial confrontation escalates rather than just counting the total number of conflicts, these studies demonstrate that high temperatures lead to more violence by increasing the likelihood that a small conflict escalates into a larger conflict.

Conclusion

Findings from a growing corpus of rigorous quantitative research across multiple disciplines suggest that past climatic events have exerted significant influence on human conflict. This influence appears to extend across the world, throughout history, and at all scales of social organization. We do not conclude that climate is the sole - or even primary - driving force in conflict, but we do find that when large climate variations occur, they can have substantial effects on the incidence of conflict across a variety of contexts. The median effect of a 1 σ change in climate variables generates an 14% change in the risk of intergroup conflict and a 4% change in interpersonal violence, across the studies that we review where it is possible to calculate standardized effects. If future populations respond similarly to past populations, then anthropogenic climate change has the potential to substantially increase conflict around the world, relative to a world without climate change.

Although there is remarkable convergence of quantitative findings across disciplines, many open questions remain. Existing research has successfully established a causal relationship between climate and conflict but is unable to fully explain the mechanisms. This fact motivates our proposed research agenda and urges caution when applying statisti-

cal estimates to future warming scenarios. Importantly, however, it does not imply that we lack evidence of a causal association. The studies in this analysis were selected for their ability to provide reliable causal inferences and they consistently point toward the existence of at least one causal pathway. To place the state of this research in perspective, it is worth recalling that statistical analyses identified the smoking of tobacco as a proximate cause of lung cancer by the 1930's (127), although the research community was unable to provide a detailed account of the mechanisms explaining the linkage until many decades later. So although future research will be critical in pinpointing why climate affects human conflict, disregarding the potential effect of anthropogenic climate change on human conflict in the interim is, in our view, a dangerously misguided interpretation of the available evidence.

Numerous competing theories have been proposed to explain the linkages between the climate and human conflict, but none have been convincingly rejected and all appear to be consistent with at least some existing results. It seems likely that climatic changes influence conflict through multiple pathways that may differ between contexts and innovative research to identify these mechanisms is a top research priority. Achieving this research objective holds great promise, as the policies and institutions necessary for conflict resolution can only be built if we understand why conflicts arise. The success of such institutions will be increasingly important in the coming decades as changes in climatic conditions amplify the risk of human conflicts.

References and Notes

1. C. Mathers, T. Boerma, D. Ma Fat, *The Global Burden of Disease: 2004 Update* (World Health Organization, 2008).
2. R. Lozano, M. Naghavi, K. Foreman, S. Lim, K. Shibuya, V. Aboyans, J. Abraham, T. Adair, R. Aggarwal, S. Y. Ahn, M. Alvarado, H. R. Anderson, L. M. Anderson, K. G. Andrews, C. Atkinson, L. M. Baddour, S. Barker-Collo, D. H. Bartels, M. L. Bell, E. J. Benjamin, D. Bennett, K. Bhalla, B. Bikbov, A. Bin Abdulhak, G. Birbeck, F. Blyth, I. Bolliger, S. Boufous, C. Bucello, M. Burch, P. Burney, J. Carapetis, H. Chen, D. Chou, S. S. Chugh, L. E. Coffeng, S. D. Colan, S. Colquhoun, K. E. Colson, J. Condon, M. D. Connor, L. T. Cooper, M. Corriere, M. Cortinovis, K. C. de Vaccaro, W. Couser, B. C. Cowie, M. H. Criqui, M. Cross, K. C. Dabhadkar, N. Dahodwala, D. De Leo, L. Degenhardt, A. Delossantos, J. Denenberg, D. C. Des Jarlais, S. D. Dharmanatne, E. R. Dorsey, T. Driscoll, H. Duber, B. Ebel, P. J. Erwin, P. Espindola, M. Ezzati, V. Feigin, A. D. Flaxman, M. H. Forouzanfar, F. G. Fowkes, R. Franklin, M. Fransen, M. K. Freeman, S. E. Gabriel, E. Gakidou, F. Gaspari, R. F. Gillum, D. Gonzalez-Medina, Y. A. Halasa, D. Harding, J. E. Harrison, R. Havmoeller, R. J. Hay, B. Hoen, P. J. Hotez, D. Hoy, K. H. Jacobsen, S. L. James, R. Jasrasaria, S. Jayaraman, N. Johns, G. Karthikeyan, N. Kassebaum, A. Keren, J. P. Khoo, L. M. Knowlton, O. Kobusingye, A. Koranteng, R. Krishnamurthi, M. Lipnick, S. E. Lipshultz, S. L. Ohno, J. M. O'Donnell, M. F. MacIntyre, L. Mallinger, L. March, G. B. Marks, R. Marks, A. Matsumori, R. Matzopoulos, B. L. Mayosi, J. H. McAnulty, M. M. McDermott, J. McGrath, G. A. Mensah, T. R. Merriman, C. Michaud, M. Miller, T. R. Miller, C. Mock, A. O. Mocumbi, A. A. Mokdad, A. Moran, K. Mulholland, M. N. Nair, L. Naldi, K. M. Narayan, K. Nasser, P. Norman, M. O'Donnell, S. B. Omer, K. Ortblad, R. Osborne, D. Ozgediz, B. Pahari, J. D. Pandian, A. P. Rivero, R. P. Padilla, F. Perez-Ruiz, N. Perico, D. Phillips, K. Pierce, C. A. Pope III, E. Porrini, F. Pourmalek, M. Raju, D. Ranganathan, J. T. Rehm, D. B. Rein, G. Remuzzi, F. P. Rivara, T. Roberts, F. R. De León, L. C. Rosenfeld, L. Rushton, R. L. Sacco, J. A. Salomon, U. Sampson, E. Sanman, D. C. Schwebel, M. Segui-Gomez, D. S. Shepard, D. Singh, J. Singleton, K. Sliwa, E. Smith, A. Steer, J. A. Taylor, B. Thomas, I. M. Tleyjeh, J. A. Towbin, T. Truelsen, E. A. Undurraga, N. Venketasubramanian, L. Vijayakumar, T. Vos, G. R. Wagner, M. Wang, W. Wang, K. Watt, M. A. Weinstock, R. Weintraub, J. D. Wilkinson, A. D. Woolf, S. Wulf, P. H. Yeh, P. Yip, A. Zabetian, Z. J. Zheng, A. D. Lopez, C. J. Murray, M. A. AlMazroa, Z. A. Memish, Global and regional mortality from 235 causes of death for 20 age groups in 1990 and 2010: A systematic analysis for the Global Burden of Disease Study 2010. *Lancet* **380**, 2095–2128 (2012). [doi:10.1016/S0140-6736\(12\)61728-0](https://doi.org/10.1016/S0140-6736(12)61728-0) [Medline](#)
3. M. A. Levy, Is the environment a national security issue? *Int. Secur.* **20**, 35 (1995). [doi:10.2307/2539228](https://doi.org/10.2307/2539228)
4. T. F. Homer-Dixon, *Environment, Scarcity and Violence* (Princeton University Press, 1999).
5. J. Scheffran, M. Brzoska, H. G. Brauch, P. M. Link, J. Schilling, Eds., *Climate Change, Human Security and Violent Conflict: Challenges for Societal Stability*, vol. 8 (Springer Verlag, 2012).
6. T. Deligiannis, The evolution of environment-conflict research: Toward a livelihood framework. *Glob. Environ. Polit.* **12**, 78–100 (2012). [doi:10.1162/GLEP_a_00098](https://doi.org/10.1162/GLEP_a_00098)
7. K. W. Butzer, Collapse, environment, and society. *Proc. Natl. Acad. Sci. U.S.A.* **109**, 3632–3639 (2012). [doi:10.1073/pnas.1114845109](https://doi.org/10.1073/pnas.1114845109) [Medline](#)
8. E. Huntington, Climatic change and agricultural exhaustion as elements in the fall of Rome. *Q. J. Econ.* **31**, 173 (1917). [doi:10.2307/1883908](https://doi.org/10.2307/1883908)
9. P. W. Holland, Statistics and causal inference. *J. Am. Stat. Assoc.* **81**, 945–960 (1986). [doi:10.1080/01621459.1986.10478354](https://doi.org/10.1080/01621459.1986.10478354)
10. N. P. Gleditsch, Armed conflict and the environment: A critique of the literature. *J. Peace Res.* **35**, 381–400 (1998). [doi:10.1177/0022343398035003007](https://doi.org/10.1177/0022343398035003007)
11. J. D. Angrist, J.-S. Pischke, The credibility revolution in empirical economics: How better research design is taking the con out of econometrics. *J. Econ. Perspect.* **24**, 3–30 (2010). [doi:10.1257/jep.24.2.3](https://doi.org/10.1257/jep.24.2.3)
12. J. D. Angrist, J.-S. Pischke, *Mostly Harmless Econometrics: An Empiricist's Companion* (Princeton University Press, 2008).
13. J. Wooldridge, *Econometric Analysis of Cross Section and Panel Data* (The MIT press, 2002).
14. O. M. Theisen, Climate clashes? Weather variability, land pressure, and organized violence in Kenya, 1989–2004. *J. Peace Res.* **49**, 81–96 (2012). [doi:10.1177/0022343311425842](https://doi.org/10.1177/0022343311425842)
15. D. Freedman, Statistical models and shoe leather. *Sociol. Methodol.* **21**, 291 (1991). [doi:10.2307/270939](https://doi.org/10.2307/270939)
16. W. H. Greene, *Econometric Analysis, Fifth Edition* (Prentice Hall, 2003).
17. A. Gelman, J. Hill, Data Analysis using Regression and Multilevel/Hierarchical Models (Cambridge University Press, 2006).
18. J. D. Angrist, A. B. Krueger, *Empirical Strategies in Labor Economics* (Elsevier Science, 1999), vol. 3, chap. 23, pp. 1277–1366.
19. W. Schlenker, M. J. Roberts, Nonlinear temperature effects indicate severe damages to U.S. crop yields under climate change. *Proc. Natl. Acad. Sci. U.S.A.* **106**, 15594–15598 (2009). [doi:10.1073/pnas.0906865106](https://doi.org/10.1073/pnas.0906865106) [Medline](#)
20. S. M. Hsiang, Temperatures and cyclones strongly associated with economic production in the Caribbean and Central America. *Proc. Natl. Acad. Sci. U.S.A.* **107**, 15367–15372 (2010). [doi:10.1073/pnas.1009510107](https://doi.org/10.1073/pnas.1009510107) [Medline](#)
21. M. Dell, B. F. Jones, B. A. Olken, Climate change and economic growth: Evidence from the last half century. *Am. Econ. J. Macroecon.* **4**, 66–95 (2012). [doi:10.1257/mac.4.3.66](https://doi.org/10.1257/mac.4.3.66)
22. H. Buhaug, Reply to Burke *et al.*: Bias and climate war research. *Proc. Natl. Acad. Sci. U.S.A.* **107**, E186–E187 (2010). [doi:10.1073/pnas.1015796108](https://doi.org/10.1073/pnas.1015796108)
23. J. O'Loughlin, F. D. Witmer, A. M. Linke, A. Laing, A. Gettelman, J. Dudhia, Climate variability and conflict risk in East Africa, 1990–2009. *Proc. Natl. Acad. Sci. U.S.A.* **109**, 18344–18349 (2012). [doi:10.1073/pnas.1205130109](https://doi.org/10.1073/pnas.1205130109) [Medline](#)
24. O. Theisen, H. Holtermann, H. Buhaug, Climate wars? assessing the claim that drought breeds conflict. *Int. Secur.* **36**, 79–106 (2011). [doi:10.1162/ISEC_a_00065](https://doi.org/10.1162/ISEC_a_00065)
25. F. Hidalgo, S. Naidu, S. Nichter, N. Richardson, Economic determinants of land invasions. *Rev. Econ. Stat.* **92**, 505–523 (2010). [doi:10.1162/REST_a_00007](https://doi.org/10.1162/REST_a_00007)
26. S. M. Hsiang, M. Burke, Climate, conflict and social stability: What does the evidence say? *Clim. Change*, available at http://papers.ssrn.com/sol3/papers.cfm?abstract_id=2302245.
27. D. T. Kenrick, S. W. Macfarlane, Ambient temperature and horn honking: A field study of the heat/aggression relationship. *Environ. Behav.* **18**, 179–191 (1986). [doi:10.1177/0013916586182002](https://doi.org/10.1177/0013916586182002)
28. A. Vrij, J. Van Der Steen, L. Koppelaar, Aggression of police officers as a function of temperature: An experiment with the fire arms training system. *J. Community Appl. Soc.* **4**, 365–370 (1994). [doi:10.1002/casp.2450040505](https://doi.org/10.1002/casp.2450040505)
29. A. Auliciems, L. DiBartolo, Domestic violence in a subtropical environment: Police calls and weather in brisbane. *Int. J. Biometeorol.* **39**, 34–39 (1995). [doi:10.1007/BF01320891](https://doi.org/10.1007/BF01320891)
30. E. Cohn, J. Rotton, Assault as a function of time and temperature: A

- moderator-variable time-series analysis. *J. Pers. Soc. Psychol.* **72**, 1322–1334 (1997). doi:10.1037/0022-3514.72.6.1322
31. J. Rotton, E. G. Cohn, Violence is a curvilinear function of temperature in Dallas: A replication. *J. Pers. Soc. Psychol.* **78**, 1074–1081 (2000). doi:10.1037/0022-3514.78.6.1074 Medline
 32. B. J. Bushman, M. C. Wang, C. A. Anderson, Is the curve relating temperature to aggression linear or curvilinear? A response to Bell (2005) and to Cohn and Rotton (2005). *J. Pers. Soc. Psychol.* **89**, 74–77 (2005). doi:10.1037/0022-3514.89.1.74 Medline
 33. C. A. Anderson, B. J. Bushman, R. W. Groom, Hot years and serious and deadly assault: Empirical tests of the heat hypothesis. *J. Pers. Soc. Psychol.* **73**, 1213–1223 (1997). doi:10.1037/0022-3514.73.6.1213 Medline
 34. C. Anderson, K. Anderson, N. Dorr, K. DeNeve, M. Flanagan, Temperature and aggression. *Adv. Exp. Soc. Psychol.* **32**, 63–133 (2000). doi:10.1016/S0065-2601(00)80004-0
 35. B. Jacob, L. Lefgren, E. Moretti, The dynamics of criminal behavior: Evidence from weather shocks. *J. Hum. Resour.* **42**, 489 (2007).
 36. R. P. Larrick, T. A. Timmerman, A. M. Carton, J. Abrevaya, Temper, temperature, and temptation: Heat-related retaliation in baseball. *Psychol. Sci.* **22**, 423–428 (2011). doi:10.1177/0956797611399292 Medline
 37. D. Card, G. B. Dahl, Family violence and football: the effect of unexpected emotional cues on violent behavior. *Q. J. Econ.* **126**, 103–143 (2011). doi:10.1093/qje/qjr001 Medline
 38. M. Ranson, Crime, weather and climate change, *Harvard working paper* (2012). URL http://papers.ssrn.com/sol3/papers.cfm?abstract_id=2111377.
 39. D. Mares, Climate change and levels of violence in socially disadvantaged neighborhood groups. *J. Urban Health* **90**, 1 (2013). Medline
 40. E. Miguel, Poverty and witch killing. *Rev. Econ. Stud.* **72**, 1153–1172 (2005). doi:10.1111/0034-6527.00365
 41. S. Sekhri, A. Storeygard, Dowry deaths: Consumption smoothing in response to climate variability in india, *Working paper* (2012). URL <http://goo.gl/2pfZr>.
 42. D. Blakeslee, R. Fishman, Rainfall shocks and property crimes in agrarian societies: Evidence from india, *SSRN working paper* (2013). URL http://papers.ssrn.com/sol3/papers.cfm?abstract_id=2208292.
 43. H. Mehlum, E. Miguel, R. Torvik, Poverty and crime in 19th century Germany. *J. Urban Econ.* **59**, 370–388 (2006). doi:10.1016/j.jue.2005.09.007
 44. A. T. Bohlken, E. J. Sergenti, Economic growth and ethnic violence: An empirical investigation of Hindu-Muslim riots in India. *J. Peace Res.* **47**, 589–600 (2010). doi:10.1177/0022343310373032
 45. H. Sarsons, Rainfall and conflict, *Harvard working paper* (2011). URL www.econ.yale.edu/conference/neudc11/papers/paper_199.pdf.
 46. C. S. Hendrix, I. Salehyan, Climate change, rainfall, and social conflict in Africa. *J. Peace Res.* **49**, 35–50 (2012). doi:10.1177/0022343311426165
 47. J. K. Kung, C. Ma, Can cultural norms reduce conflicts? Confucianism and peasant rebellions in Qing China, *Working paper* (2012). URL http://ahcc2012.org/papers/S6B-2_Kai-singKung_Ma.pdf.
 48. E. Miguel, S. Satyanath, E. Sergenti, Economic shocks and civil conflict: An instrumental variables approach. *J. Polit. Econ.* **112**, 725–753 (2004). doi:10.1086/421174
 49. M. Levy *et al.*, Freshwater availability anomalies and outbreak of internal war: Results from a global spatial time series analysis, *International workshop on Human Security and Climate Change, Holmen, Norway* (2005). URL www.ciesin.columbia.edu/pdf/waterconflict.pdf.
 50. Y. Bai, J. Kung, Climate shocks and Sino-nomadic conflict. *Rev. Econ. Stat.* **93**, 970–981 (2011). doi:10.1162/REST_a_00106
 51. S. M. Hsiang, K. C. Meng, M. A. Cane, Civil conflicts are associated with the global climate. *Nature* **476**, 438–441 (2011). doi:10.1038/nature10311 Medline
 52. M. Harari, E. La Ferrara, Conflict, climate and cells: A disaggregated analysis, *Working paper* (2013). URL www-2.ies.su.se/Nobel2012/Papers/LaFerrera_Harari.pdf.
 53. M. Couttenier, R. Soubeyran, Drought and civil war in Sub-Saharan Africa. *Econ. J.* <http://onlinelibrary.wiley.com/doi/10.1111/eoj.12042/pdf> (2013). doi:10.1111/eoj.12042
 54. M. Cervellati, U. Sunde, S. Valmori, Disease environment and civil conflicts, *IZA DP No. 5614* (2011). URL <http://papers.ssrn.com/abstract=1806415>.
 55. H. Fjelde, N. von Uexkull, Climate triggers: Rainfall anomalies, vulnerability and communal conflict in sub-Saharan Africa. *Polit. Geogr.* **31**, 444–453 (2012). doi:10.1016/j.polgeo.2012.08.004
 56. R. Jia, Weather shocks, sweet potatoes and peasant revolts in historical China. *Econ. J.* <http://onlinelibrary.wiley.com/doi/10.1111/eoj.12037/pdf> (2013). doi:10.1111/eoj.12037
 57. H. F. Lee, D. D. Zhang, P. Brecke, J. Fei, Positive correlation between the north atlantic oscillation and violent conflicts in europe. *Clim. Res.* **56**, 1–10 (2013). doi:10.3354/cr01129
 58. D. Zhang, C. Y. Jim, G. C.-S. Lin, Y.-Q. He, J. J. Wang, H. F. Lee, Climatic change, wars and dynastic cycles in China over the last millennium. *Clim. Change* **76**, 459–477 (2006). doi:10.1007/s10584-005-9024-z
 59. D. D. Zhang, P. Brecke, H. F. Lee, Y. Q. He, J. Zhang, Global climate change, war, and population decline in recent human history. *Proc. Natl. Acad. Sci. U.S.A.* **104**, 19214–19219 (2007). doi:10.1073/pnas.0703073104 Medline
 60. R. Tol, S. Wagner, Climate change and violent conflict in Europe over the last millennium. *Clim. Change* **99**, 65–79 (2010). doi:10.1007/s10584-009-9659-2
 61. D. D. Zhang, H. F. Lee, C. Wang, B. Li, Q. Pei, J. Zhang, Y. An, The causality analysis of climate change and large-scale human crisis. *Proc. Natl. Acad. Sci. U.S.A.* **108**, 17296–17301 (2011). doi:10.1073/pnas.1104268108 Medline
 62. U. Buntgen, W. Tegel, K. Nicolussi, M. McCormick, D. Frank, V. Trouet, J. O. Kaplan, F. Herzig, K. U. Heussner, H. Wanner, J. Luterbacher, J. Esper, 2500 years of European climate variability and human susceptibility. *Science* **331**, 578–582 (2011). doi:10.1126/science.1197175
 63. R. W. Anderson, N. D. Johnson, M. Koyama, From the persecuting to the protective state? Jewish expulsions and weather shocks from 1100 to 1800, *SSRN working paper* (2013). URL <http://ssrn.com/abstract=2212323>.
 64. M. B. Burke, E. Miguel, S. Satyanath, J. A. Dykema, D. B. Lobell, Warming increases the risk of civil war in Africa. *Proc. Natl. Acad. Sci. U.S.A.* **106**, 20670–20674 (2009). doi:10.1073/pnas.0907998106 Medline
 65. C. Almer, S. Boes, Climate (change) and conflict: Resolving a puzzle of association and causation, *Working paper #dp1203*, Universitaet Bern, Departement Volkswirtschaft (2012); <http://ideas.repec.org/p/ube/dpvwib/dp1203.html>.
 66. J.-F. Maystadt, O. Ecker, A. Mabiso, Extreme weather and civil war in Somalia: Does drought fuel conflict through livestock price shocks? *IFPRI working paper* (2013). URL www.ifpri.org/publication/extreme-weather-and-civil-war-somalia.
 67. W. Schlenker, M. J. Roberts, Nonlinear effects of weather on corn yields. *Rev. Agric. Econ.* **28**, 391–398 (2006). doi:10.1111/j.1467-9353.2006.00304.x
 68. W. Schlenker, D. Lobell, Robust negative impacts of climate change on African agriculture. *Environ. Res. Lett.* **5**, 014010 (2010). doi:10.1088/1748-9326/5/1/014010
 69. D. Lobell, M. Burke, Climate Change and Food Security: Adapting Agriculture to a Warmer World (Springer, 2010).
 70. E. Chaney, Revolt on the Nile: Economic shocks, religion and political influence, *Econometrica* (in press); <http://scholar.harvard.edu/chaney/publications/revolt-nile-economic-shocks-religion-and-political-power-0>.
 71. P. J. Burke, Economic growth and political survival. *B. E. J. Macroecon.* **12** (2012).
 72. J. Scheffran, M. Brzoska, J. Kominek, P. M. Link, J. Schilling, Climate change and violent conflict. *Science* **336**, 869–871 (2012). doi:10.1126/science.1221339
 73. N. Gleditsch, Whither the weather? Climate change and conflict. *J. Peace Res.* **49**, 3–9 (2012). doi:10.1177/0022343311431288
 74. T. Bernauer, T. Böhmelt, V. Koubi, Environmental changes and violent conflict. *Environ. Res. Lett.* **7**, 015601 (2012). doi:10.1088/1748-9326/7/1/015601
 75. D. Bergholt, P. Lujala, Climate-related natural disasters, economic growth, and armed civil conflict. *J. Peace Res.* **49**, 147–162 (2012). doi:10.1177/0022343311426167
 76. I. Salehyan, C. Hendrix, Climate shocks and political violence, *Annual Convention of the International Studies Association* (2012). URL <http://goo.gl/7RGEZd>.
 77. P. J. Burke, A. Leigh, Do output contractions trigger democratic change? *Am. Econ. J. Macroecon.* **2**, 124–157 (2010). doi:10.1257/mac.2.4.124
 78. M. Brückner, A. Ciccone, Rain and the democratic window of opportunity. *Econometrica* **79**, 923–947 (2011). doi:10.3982/ECTA8183
 79. G. Yancheva, N. R. Nowaczyk, J. Mingram, P. Dulski, G. Schettler, J. F.

- Negeandank, J. Liu, D. M. Sigman, L. C. Peterson, G. H. Haug, Influence of the intertropical convergence zone on the East Asian monsoon. *Nature* **445**, 74–77 (2007). doi:10.1038/nature05431 [Medline](#)
80. P. B. deMenocal, Cultural responses to climate change during the late Holocene. *Science* **292**, 667–673 (2001). doi:10.1126/science.1059827
 81. R. Kuper, S. Kröpelin, Climate-controlled Holocene occupation in the Sahara: Motor of Africa's evolution. *Science* **313**, 803–807 (2006). doi:10.1126/science.1130989
 82. D. W. Stahle, M. K. Cleaveland, D. B. Blanton, M. D. Therrell, D. A. Gay, The lost colony and Jamestown droughts. *Science* **280**, 564–567 (1998). doi:10.1126/science.280.5363.564
 83. H. Cullen, P. B. deMenocal, S. Hemming, G. Hemming, F. H. Brown, T. Guilderson, F. Sirocko, Climate change and the collapse of the Akkadian empire: Evidence from the deep sea. *Geology* **28**, 379 (2000). doi:10.1130/0091-7613(2000)28<379:CCATCO>2.0.CO;2
 84. G. H. Haug, D. Günther, L. C. Peterson, D. M. Sigman, K. A. Hughen, B. Aeschlimann, Climate and the collapse of Maya civilization. *Science* **299**, 1731–1735 (2003). doi:10.1126/science.1080444
 85. B. M. Buckley, K. J. Anchukaitis, D. Penny, R. Fletcher, E. R. Cook, M. Sano, C. Nam, A. Wichienkeo, T. T. Minh, T. M. Hong, Climate as a contributing factor in the demise of Angkor, Cambodia. *Proc. Natl. Acad. Sci. U.S.A.* **107**, 6748–6752 (2010). doi:10.1073/pnas.0910827107 [Medline](#)
 86. W. P. Patterson, K. A. Dietrich, C. Holmden, J. T. Andrews, Two millennia of North Atlantic seasonality and implications for Norse colonies. *Proc. Natl. Acad. Sci. U.S.A.* **107**, 5306–5310 (2010). doi:10.1073/pnas.0902522107 [Medline](#)
 87. D. J. Kennett, S. F. Breitenbach, V. V. Aquino, Y. Asmerom, J. Awe, J. U. Baldini, P. Bartlein, B. J. Culleton, C. Ebert, C. Jazwa, M. J. Macri, N. Marwan, V. Polyak, K. M. Prufer, H. E. Ridley, H. Sodemann, B. Winterhalder, G. H. Haug, Development and disintegration of Maya political systems in response to climate change. *Science* **338**, 788–791 (2012). doi:10.1126/science.1226299
 88. R. L. Kelly, T. A. Surovell, B. N. Shuman, G. M. Smith, A continuous climatic impact on Holocene human population in the Rocky Mountains. *Proc. Natl. Acad. Sci. U.S.A.* **110**, 443–447 (2013). doi:10.1073/pnas.1201341110 [Medline](#)
 89. R. M. D'Anjou, R. S. Bradley, N. L. Balascio, D. B. Finkelstein, Climate impacts on human settlement and agricultural activities in northern Norway revealed through sediment biogeochemistry. *Proc. Natl. Acad. Sci. U.S.A.* **109**, 20332–20337 (2012). doi:10.1073/pnas.1212730109 [Medline](#)
 90. C. Blattman, E. Miguel, Civil war. *J. Econ. Lit.* **48**, 3–57 (2010). doi:10.1257/jel.48.1.3
 91. L. V. Hedges, I. Olkin, *Statistical Method for Meta-Analysis* (Academic Press, 1985).
 92. A. Gelman, J. B. Carlin, H. S. Stern, D. B. Rubin, *Bayesian Data Analysis* (Chapman & Hall/CRC, 2004).
 93. D. Card, A. B. Krueger, Time-series minimum-wage studies: A meta-analysis. *Am. Econ. Rev.* **85**, 238 (1995).
 94. G. Meehl *et al.*, Global Climate Projections. In: Climate Change 2007: The Physical Science Basis. Contribution of Working Group I to the Fourth Assessment Report of the Intergovernmental Panel on Climate Change (Cambridge University Press, 2007).
 95. IPCC, Managing the Risks of Extreme Events and Disasters to Advance Climate Change Adaptation (Cambridge University Press, 2012).
 96. G. A. Meehl, C. Covey, K. E. Taylor, T. Delworth, R. J. Stouffer, M. Latif, B. McAvaney, J. F. B. Mitchell, The WCRP CMIP3 Multimodel Dataset: A new era in climate change research. *Bull. Am. Meteorol. Soc.* **88**, 1383–1394 (2007). doi:10.1175/BAMS-88-9-1383
 97. R. Hornbeck, The enduring impact of the American Dust Bowl: Short and long-run adjustments to environmental catastrophe. *Am. Econ. Rev.* **102**, 1477–1507 (2012). doi:10.1257/aer.102.4.1477
 98. G. D. Libecap, R. H. Steckel, Eds., *The Economics of Climate Change: Adaptations Past and Present* (The University of Chicago Press, 2011).
 99. O. Deschênes, M. Greenstone, Climate change, mortality, and adaptation: Evidence from annual fluctuations in weather in the US. *Am. Econ. J. Appl. Econ.* **3**, 152–185 (2011). doi:10.1257/app.3.4.152
 100. S. M. Hsiang, D. Narita, Adaptation to cyclone risk: Evidence from the global cross-section. *Climate Change Econ.* **3** (2012).
 101. M. Burke, K. Emerick, Adaptation to climate change: Evidence from US agriculture, *Working paper* (2013). URL <http://ssrn.com/abstract=2144928>.
 102. S. Barrios, L. Bertinelli, E. Strobl, Trends in rainfall and economic growth in Africa: A neglected cause of the African growth tragedy. *Rev. Econ. Stat.* **92**, 350–366 (2010). doi:10.1162/rest.2010.11212
 103. J. Graff Zivin, M. Neidell, Temperature and the allocation of time: Implications for climate change, *J. Labor. Econ.* (in press). URL www.nber.org/papers/w15717.
 104. B. Jones, B. Olken, Climate shocks and exports. *Am. Econ. Rev. Pap. Proc.* **100**, 454–459 (2010). doi:10.1257/aer.100.2.454
 105. O. Dube, J. Vargas, Commodity price shocks and civil conflict: Evidence from Colombia. *Rev. Econ. Stud.* (2013). URL <http://restud.oxfordjournals.org/content/early/2013/02/15/restud.rdt009.abstract>
 106. J. Angrist, A. Kugler, Rural windfall or a new resource curse? Coca, income, and civil conflict in Colombia. *Rev. Econ. Stat.* **90**, 191–215 (2008). doi:10.1162/rest.90.2.191
 107. S. Chassang, G. Padro-i-Miquel, Economic shocks and civil war. *Q. J. Polit. Sci.* **4**, 211–228 (2009). doi:10.1561/100.00008072
 108. E. Dal Bó, P. Dal Bó, Workers, warriors, and criminals: Social conflict in general equilibrium. *J. Eur. Econ. Assoc.* **9**, 646–677 (2011). doi:10.1111/j.1542-4774.2011.01025.x
 109. E. Berman, J. Shapiro, J. Felter, Can hearts and minds be bought? The economics of counterinsurgency in Iraq. *J. Polit. Econ.* **119**, 766–819 (2011). doi:10.1086/661983
 110. Y. Lei, G. Michaels, Do giant oilfield discoveries fuel internal armed conflicts? *CEP Discussion Paper* (2011). URL <http://cep.lse.ac.uk/pubs/download/dp1089.pdf>.
 111. R. Grove, The Great El Niño of 1789–93 and its global consequences: Reconstructing an extreme climate event in world environmental history. *Mediev. Hist. J.* **10**, 75–98 (2007). doi:10.1177/097194580701000203
 112. J. K. Anttila-Hughes, S. M. Hsiang, Destruction, disinvestment, and death: Economic and human losses following environmental disaster, *SSRN working paper* (2012). URL http://papers.ssrn.com/abstract_id=2220501.
 113. M. Lagi, K. Bertrand, Y. Bar-Yam, The food crises and political instability in North Africa and the Middle East, *arXiv working paper* (2011). URL <http://arxiv.org/abs/1108.2455>.
 114. R. Arezki, M. Brückner, Food prices and political instability, *IMF working paper* (2011). URL www.imf.org/external/pubs/ft/wp/2011/wp1162.pdf.
 115. C. B. Barrett, Ed., Food or Consequences? Food Security and Its Implications for Global Sociopolitical Stability (Oxford University Press, 2013).
 116. B. Carter, R. Bates, Public policy, price shocks, and civil war in developing countries, *Harvard working paper* (2012). URL www.wcfia.harvard.edu/sites/default/files/bcarter_publicpolicycivilwar.pdf.
 117. S. Barrios, L. Bertinelli, E. Strobl, Climatic change and rural-urban migration: The case of Sub-Saharan Africa. *J. Urban Econ.* **60**, 357–371 (2006). doi:10.1016/j.jue.2006.04.005
 118. S. Feng, M. Oppenheimer, W. Schlenker, Climate change, crop yields, and internal migration in the United States, *NBER working paper 17734* (2012). URL www.nber.org/papers/w17734.
 119. P. S. Jensen, K. S. Gleditsch, Rain, growth, and civil war: The importance of location. *Defence Peace Econ.* **20**, 359–372 (2009). doi:10.1080/10242690902868277
 120. J. Fearon, D. Laitin, Ethnicity, insurgency, and civil war. *Am. Polit. Sci. Rev.* **97**, 75 (2003). doi:10.1017/S0003055403000534
 121. C. K. Butler, S. Gates, African range wars: Climate, conflict, and property rights. *J. Peace Res.* **49**, 23–34 (2012). doi:10.1177/0022343311426166
 122. C. Achen, L. Bartels, Blind retrospection. electoral responses to drought, flu, and shark attacks, *Estudios/Working Papers (Centro de Estudios Avanzados en Ciencias Sociales)* (2004). URL www.march.es/ceacs/publicaciones/working/archivos/2004_199.pdf.
 123. D. Hibbs Jr., Voting and the macroeconomy, *The Oxford Handbook of Political Economy* (2006).
 124. A. J. Healy, N. Malhotra, C. H. Mo, Irrelevant events affect voters' evaluations of government performance. *Proc. Natl. Acad. Sci. U.S.A.* **107**, 12804–12809 (2010). doi:10.1073/pnas.1007420107 [Medline](#)
 125. M. Manacorda, E. Miguel, A. Vigorito, Government transfers and political support. *Am. Econ. J. Appl. Econ.* **3**, 1–28 (2011). doi:10.1257/app.3.3.1 [Medline](#)

126. M. Auffhammer, S. Hsiang, W. Schlenker, A. Sobel, Using weather data and climate model output in economic analyses of climate change. *Rev. Environ. Econ. Policy* **7**, 181–198 (2013). doi:10.1093/reep/ret016
127. H. Witschi, A short history of lung cancer. *Toxicol. Sci.* **64**, 4–6 (2001). doi:10.1093/toxsci/64.1.4 Medline
128. S. M. Hsiang, Visually-weighted regression, *SSRN working paper* (2013). URL http://papers.ssrn.com/sol3/papers.cfm?abstract_id=2265501.
129. G. S. Watson, Smooth regression analysis. *Sankhya* **26**, 359 (1964).
130. E. A. Nadaraya, On estimating regression. *Theory Probab. Appl.* **9**, 141–142 (1964). doi:10.1137/1109020
131. D. R. Legates, C. J. Willmott, Mean seasonal and spatial variability global surface air temperature. *Theor. Appl. Climatol.* **41**, 11–21 (1990). doi:10.1007/BF00866198
132. A. E. Sutton, J. Dohrn, K. Loyd, A. Tredennick, G. Bucini, A. Solórzano, L. Prihodko, N. P. Hanan, Does warming increase the risk of civil war in Africa? *Proc. Natl. Acad. Sci. U.S.A.* **107**, E102 (2010). doi:10.1073/pnas.1005278107 Medline
133. H. Buhaug, H. Hegre, H. Strand, Sensitivity analysis of climate variability and civil war, *PRIO working paper* (2010); <http://goo.gl/Ar3xox>.
134. H. Buhaug, Climate not to blame for African civil wars. *Proc. Natl. Acad. Sci. U.S.A.* **107**, 16477–16482 (2010). doi:10.1073/pnas.1005739107 Medline
135. M. Burke, J. Dykema, D. Lobell, E. Miguel, S. Satyanath, Climate and civil war: Is the relationship robust? *NBER working paper 16440* (2010). URL www.nber.org/papers/w16440.
136. M. B. Burke, E. Miguel, S. Satyanath, J. A. Dykema, D. B. Lobell, Climate robustly linked to African civil war. *Proc. Natl. Acad. Sci. U.S.A.* **107**, E185 (2010). doi:10.1073/pnas.1014879107 Medline
137. M. B. Burke, E. Miguel, S. Satyanath, J. A. Dykema, D. B. Lobell, Reply to Sutton *et al.*: Relationship between temperature and conflict is robust. *Proc. Natl. Acad. Sci. U.S.A.* **107**, E103 (2010). doi:10.1073/pnas.1005748107
138. A. Ciccone, Economic shocks and civil conflict: A comment. *Am. Econ. J. Appl. Econ.* **3**, 215–227 (2011). doi:10.1257/app.3.4.215 Medline
139. E. Miguel, S. Satyanath, Re-examining economic shocks and civil conflict. *Am. Econ. J. Appl. Econ.* **3**, 228–232 (2011). doi:10.1257/app.3.4.228 Medline
140. A.-C. Disdier, K. Head, The puzzling persistence of the distance effect on bilateral trade. *Rev. Econ. Stat.* **90**, 37–48 (2008). doi:10.1162/rest.90.1.37
141. D. Card, J. Kluve, A. Weber, Active labour market policy evaluations: A meta-analysis. *Econ. J.* **120**, F452–F477 (2010). doi:10.1111/j.1468-0297.2010.02387.x

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Supplementary Materials

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Supplementary Text

Figs. S1 to S4

Tables S1 to S4

References (140, 141)

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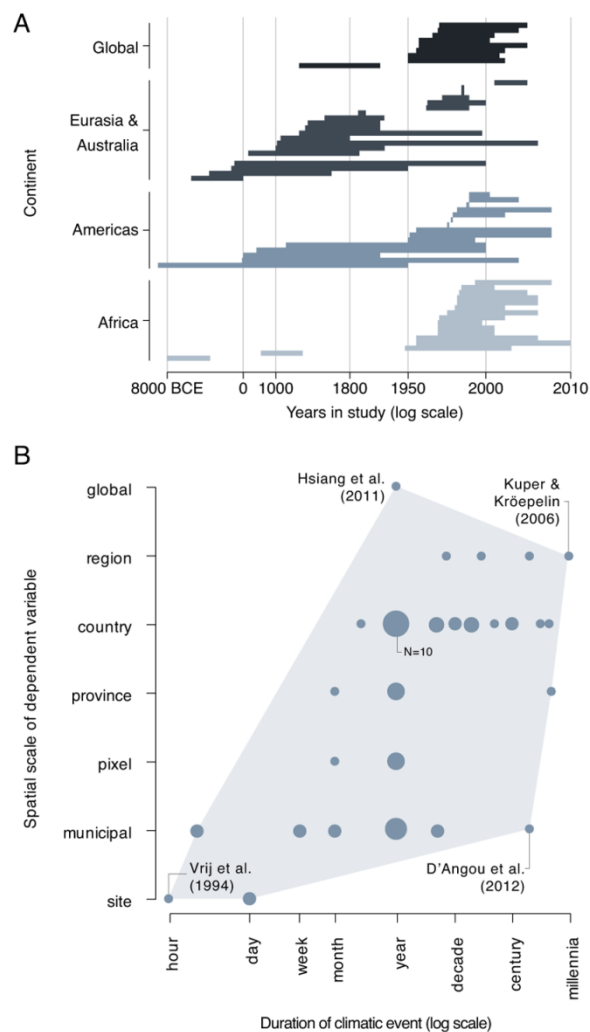


Fig. 1. Samples and spatiotemporal resolutions of 60 studies examining intertemporal associations between climatic variables and human conflict. (A) The location of each study region (y-axis) against the period of time included in the study (x-axis). The x-axis is scaled according to *log years before present* but labeled according to the year of the *Common Era*. **(B)** The level of aggregation in social outcomes (y-axis) against the timescale of climatic events (x-axis). The envelope of spatial and temporal scales where associations are documented is shaded, with studies at extreme vertices labeled for reference. Marker size denotes the number of studies at each location, with the smallest bubbles marking individual studies and the largest bubble marking 10 studies.

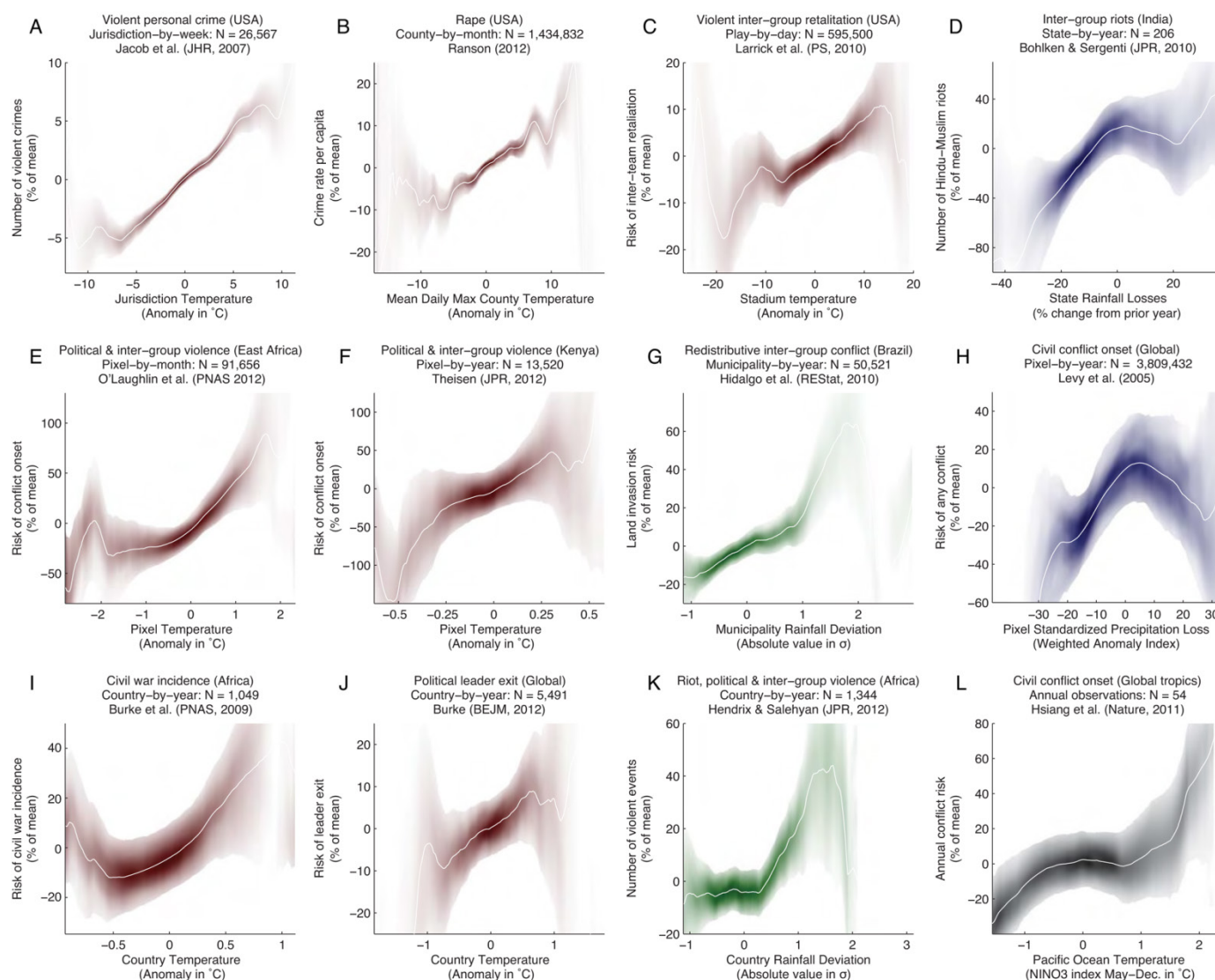


Fig. 2. Empirical studies indicate that climatological variables have a large effect on the risk of violence or instability in the modern world. Examples from studies of modern data that identify the causal effect of climate variables on human conflict. Both dependent and independent variables have had location-effects and trends removed, so all samples have a mean of zero. Relationships between climate and conflict outcomes are shown with non-parametric “watercolor regressions”, where the color intensity of 95% confidence intervals depicts the likelihood that the true regression line passes through a given value (darker is more likely) (128). White line is the conditional mean (129, 130). Climate variables are indicated by color: red = temperature, green = rainfall deviations from normal, blue = precipitation loss, black = ENSO. Panel titles describe the outcome variable, location, unit of analysis, sample size and study. Because the samples examined in each study differ, the units and scales change across each panel (see Figs. 4 and 5 for standardized effect-sizes). “Rainfall deviation” represents the absolute value of location-specific rainfall anomalies, with both abnormally high and abnormally low rainfall events described as having a large rainfall deviation. “Precipitation Loss” is an index describing how much lower precipitation is relative to the prior year or long-term mean.

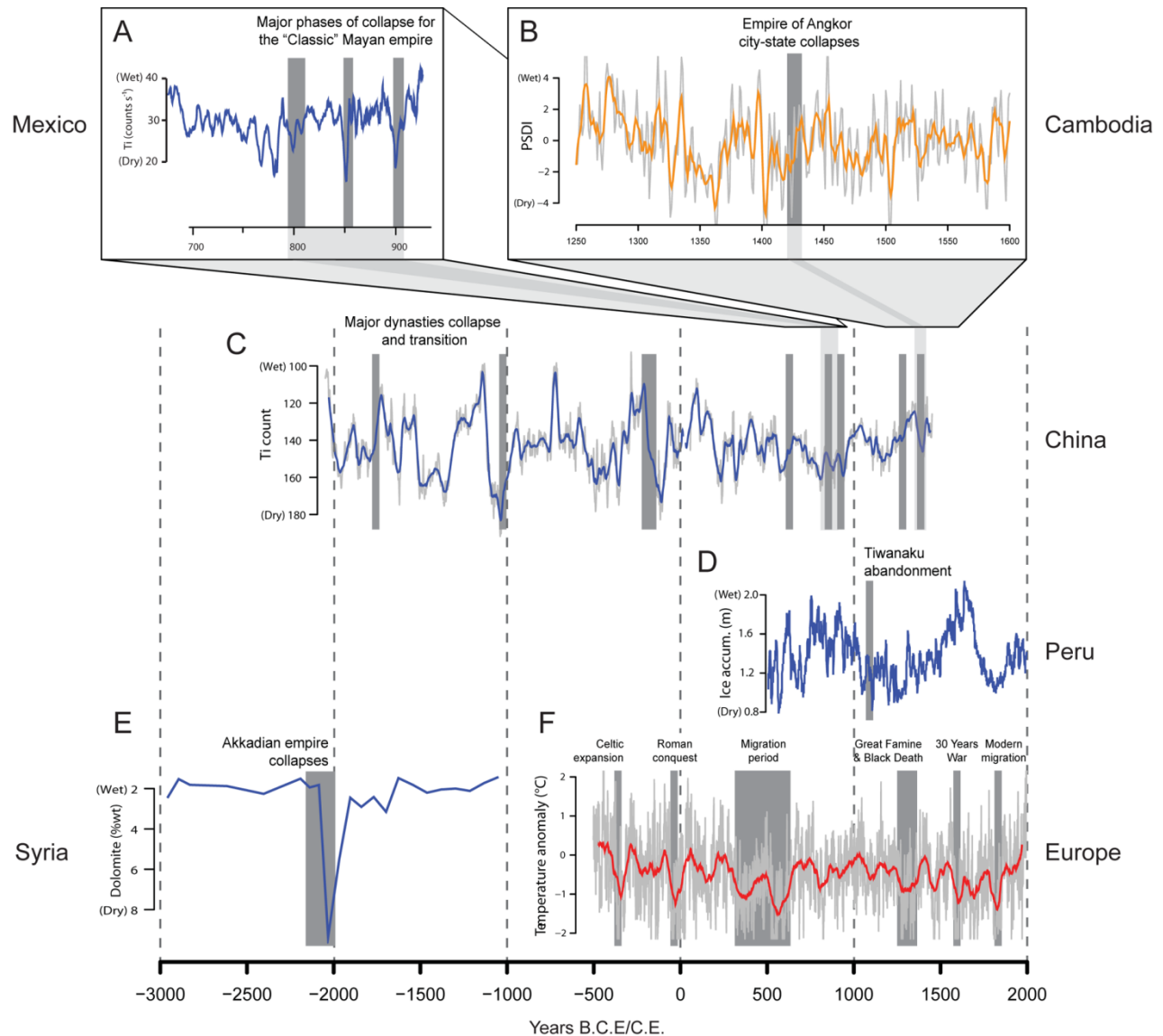


Fig. 3. Examples of paleoclimate reconstructions that find associations between climatic changes and human conflict. Lines are climate reconstructions (red = temperature, blue = precipitation, orange = drought; smoothed moving averages when light gray lines are shown), and dark gray bars indicate periods of substantial social instability, violent conflict, or the breakdown of political institutions. (A) Alluvial sediments from the Cariaco Basin indicate substantial multi-year droughts coinciding with the collapse of the Maya (84). (B) Reconstruction of a drought index from tree rings in Vietnam show sustained mega-droughts prior the collapse of the Angkor kingdom (85). (C) Sediments from Lake Huguang Maar in China indicate abrupt and sustained periods of reduced summertime precipitation that coincided with most major dynastic transitions (79). The collapse of the Tang Dynasty (907) coincided with the terminal collapse of the Maya (A), both of which occurred when the Pacific Ocean altered rainfall patterns in both hemispheres (79). Similarly, the collapse of the Yuan Dynasty (1368) coincided with collapse of Angkor (B) which shares the same regional climate. (D) Tiwanaku cultivation of the Lake Titicaca region ended abruptly following a drying of the region, as measured by ice accumulation in the Quelccaya Ice Cap, Peru (80). (E) Continental dust blown from Mesopotamia into the Gulf of Oman indicate terrestrial drying that is coincident with the collapse of the Akkadian empire (83). (F) European tree rings indicate that anomalously cold periods were associated with major periods of instability on the European continent (62).

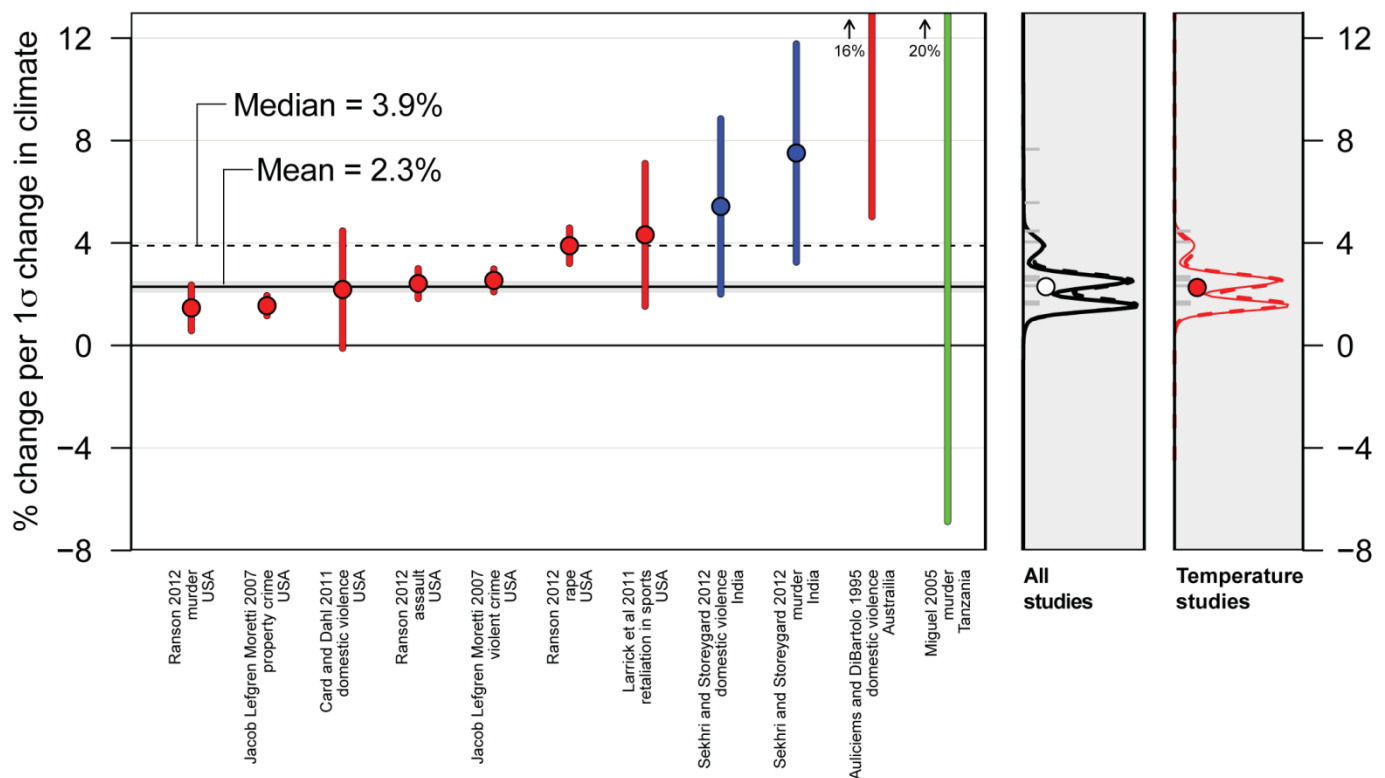


Fig. 4. Modern empirical estimates for the effect of climatic events on the risk of interpersonal violence. Each marker represents the estimated effect of a 1σ increase in a climate variable, expressed as a percentage change in the outcome variable relative to its mean. Whiskers represent the 95% confidence interval on this point estimate. Colors indicate the forcing climate variable. A coefficient is positive if conflict increases with higher temperature (red), greater rainfall loss (blue), or greater rainfall deviation from normal (green). Dashed line is the median estimate, and the solid black line the precision-weighted mean with its 95% confidence interval shown in gray. The panels on the right show the precision-weighted mean effect (circle) and the distribution of study results for all 11 results looking at individual conflict or for the subset of 8 results focusing on temperature effects; distributions of effect sizes are either precision-weighted (solid line) or derived from a Bayesian hierarchical model (dashed line). See supplementary materials for details on the individual studies and on the calculation of mean effects and their distribution.

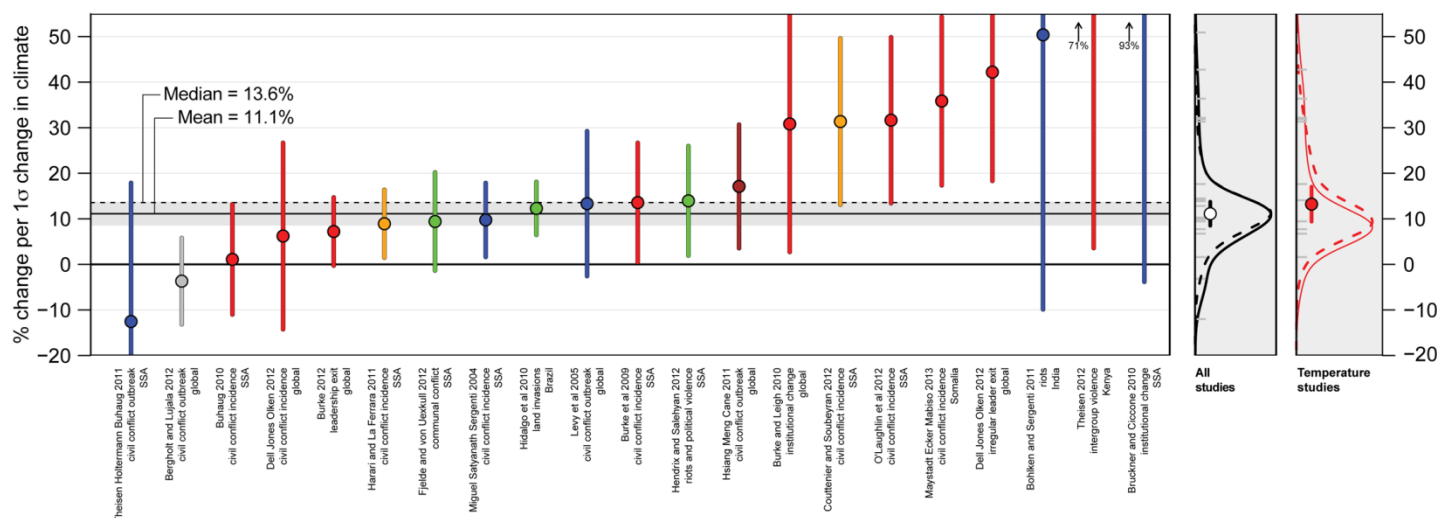


Fig. 5. Modern empirical estimates for the effect of climatic events on the risk of intergroup conflict. Each marker represents the estimated effect of a 1σ increase in a climate variable, expressed as a percentage change in the outcome variable relative to its mean. Whiskers represent the 95% confidence interval on this point estimate. Colors indicate the forcing climate variable. A coefficient is positive if conflict increases with higher temperature (red), greater rainfall loss (blue), greater rainfall deviation from normal (green), more floods and storms (gray), more El Niño-like conditions (brown), or more drought (orange) as captured by different drought indices. Dashed line is the median estimate, and the solid black line the precision-weighted mean with its 95% confidence interval shown in gray. The panels at right show the precision-weighted mean effect (circle) and the distribution of study results for all 21 results looking at intergroup conflict or for the subset of 12 results focusing on temperature effects (which includes the ENSO and drought studies); distributions of effect sizes are either precision-weighted (solid line) or derived from a Bayesian hierarchical model (dashed line). See supplementary materials for details on the individual studies and on the calculation of mean effects and their distribution.

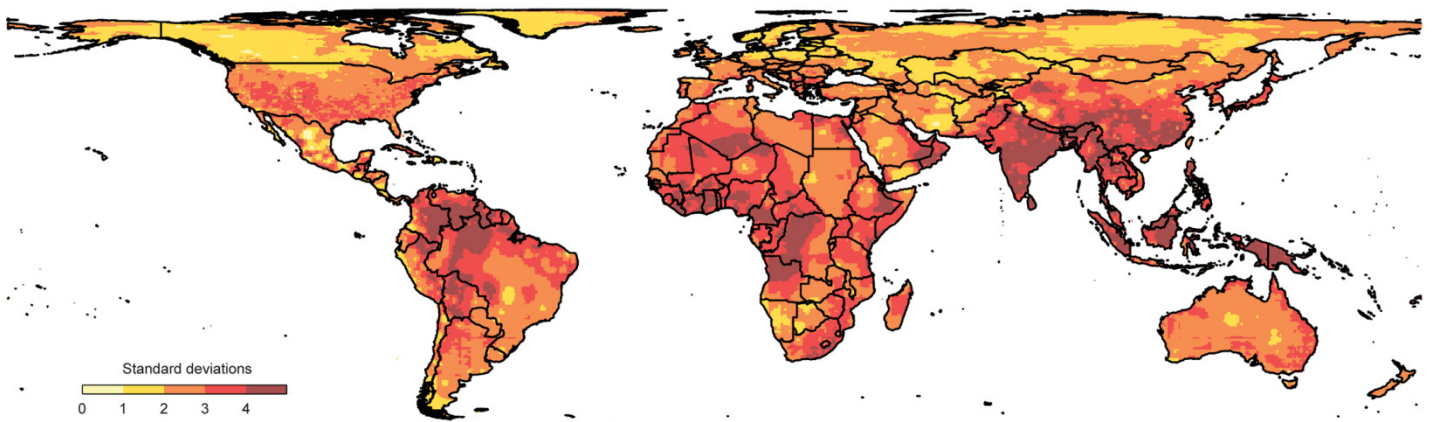


Fig. 6. Projected temperature change by 2050 as a multiple of the local historical standard deviation (σ) of temperature. Temperature projections are for the A1B scenario and are averaged across 21 global climate models reporting in the CMIP3 (96). Changes are the difference between projected annual average temperature in 2050 and average temperature in 2000. The historical standard deviation of temperature is calculated from annual average temperatures at each grid cell over the period 1950-2008, using University of Delaware data (131). The map is an equal-area projection.

Table 1. Unique quantitative studies testing for a relationship between climate and conflict, violence or political instability.

| Study | Sample Period | Sample Region | Time Unit | Spatial Unit | Independent Variable | Dependent Variable | Stat. Test | Large effect | Reject $\beta = 0$ | Reject $\beta = 10\%$ | Ref. |
|------------------------------------|---------------|---------------|-----------|--------------|----------------------|--|------------|--------------|--------------------|-----------------------|------|
| Interpersonal Conflict (15) | | | | | | | | | | | |
| Anderson et al. 2000 ^a | 1950–1997 | USA | annual | country | temp | violent crime | Y | | Y | - | (34) |
| Auliciems et al. 1995 ^x | 1992 | Australia | week | municipality | temp | domestic violence | Y | | Y | - | (29) |
| Blakeslee et al. 2013 | 1971–2000 | India | annual | municipality | rain | & property crime | Y | | Y | - | (42) |
| Card et al. 2011 ^{†x} | 1995–2006 | USA | day | municipality | temp | domestic violence | Y | | Y | - | (37) |
| Cohn et al. 1997 ^b | 1987–1988 | USA | hours | municipality | temp | violent crime | Y | | Y | - | (30) |
| Jacob et al. 2007 ^{*x} | 1995–2001 | USA | week | municipality | temp | violent & property crime | Y | | Y | - | (35) |
| Kenrick et al. 1986 ‡ | 1985 | USA | day | site | temp | hostility | Y | | Y | - | (27) |
| Larrick et al. 2011 ^{*x†} | 1952–2009 | USA | day | site | temp | violent retaliation | Y | Y | Y | - | (36) |
| Mares 2013 | 1990–2009 | USA | month | municipality | temp | violent crime | Y | Y | Y | - | (39) |
| Miguel 2005 ^{x†} | 1992–2002 | Tanzania | annual | municipality | rain | murder violent | Y | Y | No | No | (40) |
| Mehlum et al. 2006 | 1835–1861 | Germany | annual | province | rain | & property crime | Y | Y | Y | - | (43) |
| Ranson 2012 ^{*x} | 1960–2009 | USA | month | county | temp | personal violence | Y | Y | Y | - | (38) |
| Rotton et al. 2000 ^b | 1994–1995 | USA | hours | municipality | temp | violent crime murder & domestic violence | Y | Y | Y | - | (31) |
| Sekhri et al. 2013 ^x | 2002–2007 | India | annual | municipality | rain | violence | Y | Y | Y | - | (41) |
| Vrij et al. 1994‡ | 1993 | Netherlands | hours | site | temp | police use of force | Y | Y | Y | - | (28) |

| Study | Sample Period | Sample Region | Table 1: (continued) | | Independent Variable | Dependent Variable | Stat. Test | Large effect | Reject $\beta = 0$ | Reject $\beta = 10\%$ | Ref. | |
|---------------------------------------|---------------|---------------|-----------------------|---------------|----------------------|---|------------|---|--------------------|-----------------------|------|------|
| | | | Time Unit | Spatial Unit | | | | | | | | |
| Intergroup Conflict (31) | | | | | | | | | | | | |
| Almer et al. 2012 | 1985–2008 | SSA | annual | country | rain/temp | civil conflict | Y | Y | No | No | (65) | |
| Anderson et al. 2013 | 1100–1800 | Europe | decade | municipality | temp | minority expulsion | Y | Downloaded from http://science.sciencemag.org/ on August 2, 2017 | Y | - | (63) | |
| Bai et al. 2010 | 220–1839 | China | decade | country | rain | transboundary | Y | | Y | - | (50) | |
| Bergholt et al. 2012 ⁺⁺ | 1980–2007 | Global | annual | country | flood/storm | civil conflict | Y | | No | No | Y | (75) |
| Bohlken et al. 2011 ^{*+} | 1982–1995 | India | annual | province | rain | inter-group | Y | | No | No | No | (44) |
| Buhaug 2010 ⁺ | 1979–2002 | SSA | annual | country | temp | civil conflict | Y | | No | No | No | (22) |
| Burke 2012 ^{*++} | 1963–2001 | Global | annual | country | rain/temp | political instability | Y | | No | No [◇] | No | (71) |
| Burke et al. 2009 ^{*++c} | 1981–2002 | SSA | annual | country | temp | civil conflict | Y | | Y | Y | - | (64) |
| Cervellati et al. 2011 | 1960–2005 | Global | annual | country | drought | civil conflict | Y | | Y | Y | - | (54) |
| Chaney 2011 | 641–1438 | Egypt | annual | country | Nile floods | political instability | Y | | Y | Y | - | (70) |
| Couttenier et al. 2011 ⁺ | 1957–2005 | SSA | annual | country | PDSI [○] | civil conflict political instability | Y | | Y | Y | - | (53) |
| Dell et al. 2012 ⁺ | 1950–2003 | Global | annual | country | temp | & civil conflict | Y | Y | Y | - | (21) | |
| Fjelde et al. 2012 ⁺⁺ | 1990–2008 | SSA | annual | province | rain | inter-group | Y | No | No [◇] | No | (55) | |
| Harari et al. 2013 ⁺ | 1960–2010 | SSA | annual | pixel (1°) | drought | civil conflict | Y | Y | Y | - | (52) | |
| Hendrix et al. 2012 ^{*++} | 1991–2007 | SSA | annual | country | rain | inter-group | Y | Y | Y | - | (46) | |
| Hidalgo et al. 2010 ^{*++} | 1988–2004 | Brazil | annual | municipality | rain | inter-group | Y | Y | Y | - | (25) | |
| Hsiang et al. 2011 ^{*+} | 1950–2004 | Global | annual | world | ENSO [⊕] | civil conflict | Y | Y | Y | - | (51) | |
| Jia 2012 | 1470–1900 | China | annual | province | drought/flood | peasant rebellion | Y | Y | Y | - | (56) | |
| Kung et al. 2012 | 1651–1910 | China | annual | county | rain | peasant rebellion | Y | Y | Y | - | (47) | |
| Lee et al. 2013 | 1400–1999 | Europe | decade | region | NAO [⊗] | violent conflict | Y | Y | Y | - | (57) | |
| Levy et al. 2005 ^{*++} | 1975–2002 | Global | annual | pixel (2.5°) | rain | civil conflict | Y | Y | No [◇] | No | (49) | |
| Maystadt et al. 2013 ⁺ | 1997–2009 | Somalia | month | province | temp | civil conflict | Y | Y | Y | - | (66) | |
| Miguel et al. 2004 ^{+d} | 1979–1999 | SSA | annual | country | rain | civil war | Y | Y | Y | - | (48) | |
| O'Laughlin et al. 2012 ^{*++} | 1990–2009 | E. Africa | month | pixel (1°) | rain/temp | civil/inter-group | Y | Y | Y | - | (23) | |
| Salehyan et al. 2012 | 1979–2006 | Global | annual | country | PDSI [○] | civil/inter-group | Y | Y | Y | - | (76) | |
| Sarsons 2011 | 1970–1995 | India | annual | municipality | rain | inter-group | Y | Y | Y | - | (45) | |
| Theisen et al. 2011 ⁺⁺ | 1960–2004 | Africa | annual | pixel (0.5°) | rain | civil conflict | Y | No | No | No | (24) | |
| Theisen 2012 ^{*++} | 1989–2004 | Kenya | annual | pixel (0.25°) | rain/ temp | civil/inter-group | Y | Y | No [◇] | No | (14) | |
| Tol et al. 2009 | 1500–1900 | Europe | decade centur y | region | rain/temp | transboundary | Y | Y | Y | - | (60) | |
| Zhang et al. 2007 ^e | 1400–1900 | N. Hem. | y | region | temp | instability | Y | Y | Y | - | (59) | |

| Table 1: (continued) | | | | | | | | | | | |
|--|----------------|---------------|-----------|--------------|----------------------|---------------------------|------------|---|--------------------|-----------------------|------|
| Study | Sample Period | Sample Region | Time Unit | Spatial Unit | Independent Variable | Dependent Variable | Stat. Test | Large effect | Reject $\beta = 0$ | Reject $\beta = 10\%$ | Ref. |
| Institutional breakdown & population collapse (15) | | | | | | | | Downloaded from http://science.sciencemag.org/ on August 2, 2017 | | | |
| Brückner et al. 2011 [†] | 1980–2004 | SSA | annual | country | rain | inst. change | Y | | Y | - | (78) |
| Buckley et al. 2010 ^{**} | 1030–2008 | Cambodia | decade | country | drought | collapse | No | | - | - | (85) |
| Büntgen et al. 2011 ^{**} | 400 BCE–2000 | Europe | decade | region | rain/temp | instability | No | | - | - | (62) |
| Burke et al. 2010 ^{††} | 1963–2007 | Global | annual | country | rain/temp | inst. change | Y | | Y | - | (77) |
| Cullen et al. 2000 ^{**} | 4000 BCE–0 | Syria | century | country | drought | collapse | No | | - | - | (83) |
| D'Anjou et al. 2012 | 550 BCE–1950 | Norway | century | municipality | temp | collapse | Y | | Y | - | (89) |
| DeMenocal 2001 ^{**} | 500–2000 | Peru | century | country | drought | collapse | No | | - | - | (80) |
| Haug et al. 2003 ^{**} | 0–1900 | Mexico | century | country | drought | collapse | No | | - | - | (84) |
| Kelly et al. 2013 | 10050 BCE–1950 | USA | century | state | temp/rain | collapse | Y | Y | Y | - | (88) |
| Kennett et al. 2012 | 40 BCE–2006 | Belize | decade | country | rain | collapse | No | - | - | - | (87) |
| Kuper et al. 2006 | 8000–2000 BCE | N. Africa | millennia | region | rain | collapse | No | - | - | - | (81) |
| Patterson et al. 2010 | 200 BCE–1700 | Iceland | decade | country | temp | collapse | No | - | - | - | (86) |
| Stahle et al. 1998 | 1200–2000 | USA | multiyear | municipality | PDSI [⊙] | collapse | No | - | - | - | (82) |
| Yancheva et al. 2007 ^{**} | 2100 BCE–1700 | China | century | country | rain/temp | collapse | No | - | - | - | (79) |
| Zhang et al. 2006 | 1000–1911 | China | decade | country | temp | civil conflict & collapse | Y | Y | Y | - | (58) |
| Number of studies (60 total): | | | | | | | 50 | 47 | 37 | 1 | |
| Fraction of those using statistical tests: | | | | | | | 100% | 94% | 74% | 2% | |

“Stat. Test” is Y if the analysis uses formal statistical methods to quantify the influence of climate variables and uses hypothesis testing procedures. “Large Effect” is Y if the point estimate for the effect size is considered substantial by the authors or is greater in magnitude than 10% of the mean risk level for a 1σ change in climate variables. “Reject $\beta = 0$ ” is Y if the study rejects an effect size of zero at the 95% confidence level. \diamond the effect size in the study is statistically significant at the 10% level but not at the 5% level. “Reject $\beta = 10\%$ ” is Y if the study is able to reject the hypothesis that the effect size is larger than 10% of the mean risk level for a 1σ change in climate variables. * Shown in Fig. 2. ** Shown in Fig. 3. \times Shown in Fig. 4. $^+$ Shown in Fig. 5. † Reanalyzed using the common statistical model containing location fixed effects and trends (see supplementary materials). ‡ Actual experiment. a Also see (33). b Also see discussion in (32). c SSA is Sub-Saharan Africa. Also see discussion in (22, 132–137). d Also see discussion in (138, 139). e Also see (61). \odot PDSI is the Palmer Drought Severity Index. \oplus ENSO is the El Niño-Southern Oscillation. \otimes NAO is the North Atlantic Oscillation.

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