

Model-based Distances

Recall that 1 weakness of the Hamming distance = proportion of sites that differ

is that it fails to account for back substitutions or multiple substitutions

on an edge $\xrightarrow{\quad}$ and unless sequences are

$A \rightarrow G \rightarrow A$

$A \rightarrow C \rightarrow G$

very closely related, the Hamming distance tends to under estimate the amount of evolutionary distance between 2 taxa.

To address this, we use our models JC, K2P, K3ST, GTR, GM to "correct" the Hamming distance and to account for unseen changes.

Example: Jukes-Cantor model and Jukes-Cantor distance

Parameters: single edge thought of as path from taxon a to b in tree

$$\vec{p}_0 = (.25 \ .25 \ .25 \ .25)$$

$$Q = \begin{pmatrix} -\alpha & \alpha/3 & \alpha/3 & \alpha/3 \\ & -\alpha & & \\ & & -\alpha & \\ & & & -\alpha \end{pmatrix}, \text{ branch length } t, \text{ and Markov matrix } M(t) =$$

$$M = \begin{pmatrix} 1-\alpha & \alpha/3 & \alpha/3 & \alpha/3 \\ & 1-\alpha & & \\ & & 1-\alpha & \\ & & & 1-\alpha \end{pmatrix}$$

$$\text{with } \alpha(t) = \frac{3}{4} \left(1 - \frac{4}{3} \alpha t \right)$$

Assuming the JC model, the expected pattern frequency array is

$$P = \text{diag}([.25 \ .25 \ .25 \ .25]) M(t) = \begin{pmatrix} \frac{1}{4}(1-\alpha) & \alpha/12 & \alpha/12 & \alpha/12 \\ \alpha/12 & \frac{1}{4}(1-\alpha) & \alpha/12 & \alpha/12 \\ \alpha/12 & & \ddots & \\ \alpha/12 & & & \end{pmatrix}$$

with a QS above

[No data here. Yet.]

If we interpret α = rate in units $\frac{\# \text{ substitutions}}{\text{time}}$, then

$$\alpha t = \left(\frac{\# \text{ substitutions per site}}{\text{time } t} \right) (\text{time } t) \quad \leftarrow (\text{model at single site})$$

$$\alpha t = \# \text{ of substitutions over time } t$$

including those hidden ones

This is (will be) the Jukes Center distance once we solve for it:

$$a(t) = a = \frac{3}{4} \left(1 - e^{-\frac{4}{3} \alpha t} \right)$$

$$\Rightarrow \boxed{\alpha t = -\frac{3}{4} \ln \left(1 - \frac{4}{3} a \right)} = \# \text{ of subst. per site over elapsed time } t$$

Since we don't have a in hand, we must estimate it from data,

i.e. from the empirical pattern freq. array

In theory, S_0 and S_i disagree with probability $12 \left(\frac{a}{12} \right) = a$

Sum of off-diagonal
entries

Thus, we estimate a with $\hat{a} = \frac{\# \text{ of sites with non-constant pattern}}{\# \text{ of sites } n}$

= Hamming distance!

Def: The Jukes-Center distance $d_{JC}(S_0, S_1)$ between aligned sequences is

$$d_{JC} = -\frac{3}{4} \ln \left(1 - \frac{4}{3} \hat{a} \right)$$

$$\hat{a} = d_{\text{Hamming}}(S_0, S_1)$$

Ex.

S_0 : A A A C G G C A T G
 S_1 : A A A T G G C T A G

$$d_{\text{Hamming}}(S_0, S_1) = 3/10 = .3$$

$$d_{\text{JC}}(S_0, S_1) = -\frac{3}{4} \ln\left(1 - \frac{4}{3}(.3)\right) \approx .38$$

Several Comments:

- $d_{\text{JC}} = .38 > .3 = d_{\text{Hamming}}$ to account for hidden mutations
- $d_{\text{JC}}(S_0, S_1) = -\frac{3}{4} \ln\left(1 - \frac{4}{3} \hat{a}\right) \Rightarrow 0 \leq \hat{a} < \frac{3}{4}$ for the log to make sense

This makes sense: if 2 sequences are generated at random

S_0 : - - -

S_1 : - - -

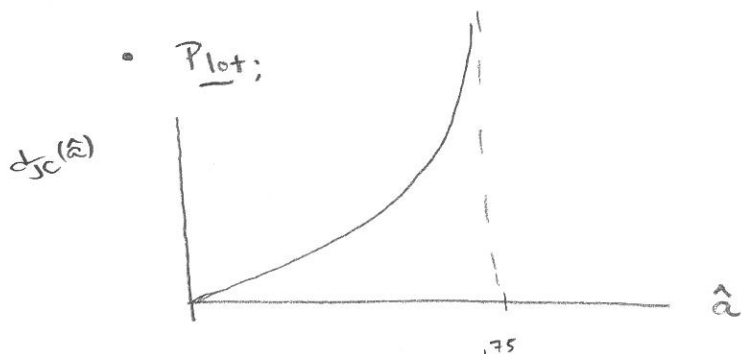
using the root distribution ($1/4, 1/4, 1/4, 1/4$)

then S_0 and S_1 will agree roughly

$1/4$ of the time and disagree roughly

$\hat{a} = 3/4$ of the sites $d(S_0, S_1) \gg 0$

• Plot;



Thus as \hat{a} gets close to .75,

$d_{\text{JC}}(S_0, S_1) \rightarrow \infty$, huge distances

"Saturated sequences"

informally, you can not differentiate them from 2 randomly selected sequences

• Since JC is time-reversible on a tree

$$\begin{array}{c}
 t_1 \quad t_2 \\
 \diagdown \quad \diagup \\
 S_0 \quad S_1
 \end{array} \equiv \frac{t_1 + t_2}{S_0 \quad S_1}$$

or more generally we can consider paths in trees.

Similar methods can be used to derive distance formulas for K2P, K3ST, GTR. 4.

$$d_{K2P}(S_1, S_2) = -\frac{1}{2} \ln(1 - 2\hat{b} - \hat{e}) - \frac{1}{4} \ln(1 - 2\hat{e})$$

\hat{b} = proportion of observed transitions
 \hat{e} = proportion of observed transv.

$$d_{K3}(S_1, S_2) = -\frac{1}{4} \left(\ln(1 - 2\hat{b} - 2\hat{e}) + \ln(1 - 2\hat{b} - \hat{d}) + \ln(1 - 2\hat{e} - \hat{d}) \right)$$

\hat{b}
 \hat{c}
 \hat{d} are best estimates for $M_{K3P} = \begin{pmatrix} * & b & c & d \end{pmatrix}$

OR $P = \text{diag}([1/4 \ 1/4 \ 1/4 \ 1/4]) M_{K3P}$

If $\hat{e} = \hat{d}$, then this simplifies to K2P.

There is also a more general GTR distance \rightarrow see book. Requires

- 1) normalization
- 2) Knowledge of Trace of a matrix.

Why normalize?

Eg. JC model

Probability: (Model)

$$d_{JC}(S_1, S_2) = -\frac{3}{4} \ln\left(1 - \frac{4}{3}a\right)$$

||
 αt
 {

Intertwined

$$\alpha t = 2\alpha\left(\frac{t}{2}\right) = 4\alpha\left(\frac{t}{4}\right) = c\alpha\left(\frac{t}{c}\right)$$

We can compute the product αt ,

but neither α or t separately

Estimate from data.

$$d_{JC}(S_0, S_1) = -\frac{3}{4} \ln\left(1 - \frac{4}{3}\hat{a}\right)$$

However, α represents a total mutation rate: in JC

$$\underbrace{\frac{1}{4}\alpha}_{\text{rate leaving state A}} + \underbrace{\frac{1}{4}\alpha}_G + \underbrace{\frac{1}{4}\alpha}_C + \underbrace{\frac{1}{4}\alpha}_T = \alpha$$

rate leaving state

Normalize so that $\alpha=1$

$$Q = \begin{pmatrix} -1 & 1/3 & 1/3 & 1/3 \\ \vdots & \ddots & \ddots & \ddots \end{pmatrix} \text{ etc.}$$

$$\alpha t = 1(t) = \frac{\# \text{ of subst.}}{t} \cdot t$$

$$= t \# \text{ of substitutions}$$

i.e. t measures the expected number of substitutions over the elapsed time.

MATLAB eq's