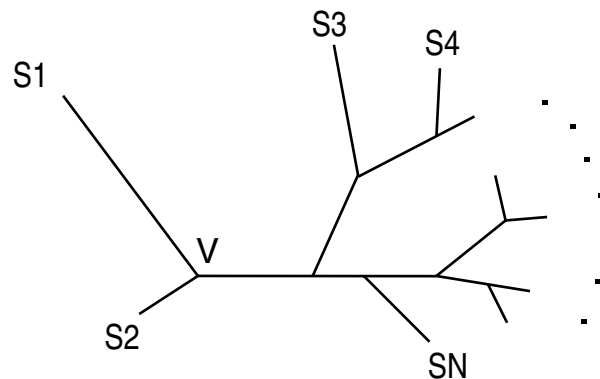


Models of Molecular Evolution: An Introduction

Elizabeth S. Allman
University of Southern Maine

MAT 500 Computational Methods in Genomics
University of Maine, Orono
November 11, 2003

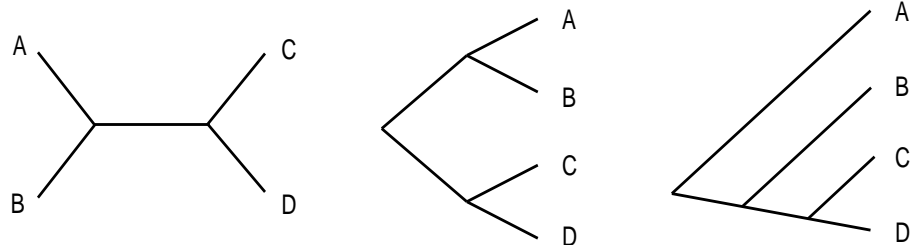


Which tree best relates N taxa?

Problem:

Given aligned biological sequences, presumed to have arisen from a common ancestral sequence, infer their evolutionary history.

A: AATCGCTGCTCGACC...
B: AAATGCTACTGGACC...
C: AAACGTTACTGGAGC...
D: AATCGTGGCTCGATC...



Do we care about root location? edge lengths?
description of mutation process along edges?
sequences at internal nodes?

In addition to the intrinsic interest of wanting to understand evolutionary history, there are *many* less obvious applications:

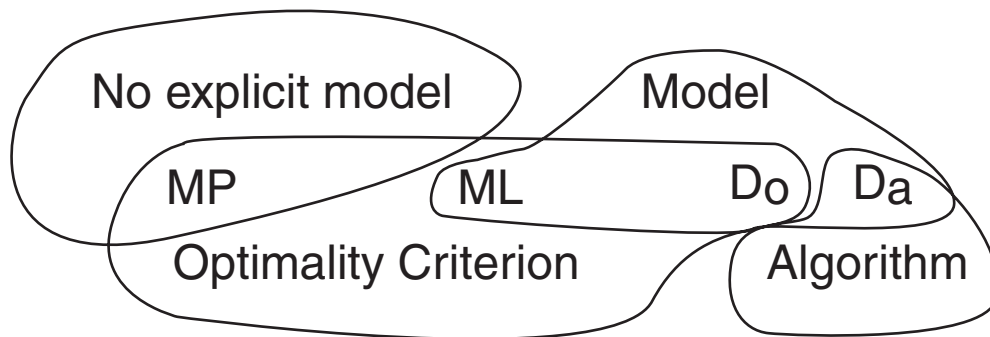
- Epidemiology – Florida dentist AIDS cluster
- Ecology – co-evolution of species and parasites; assessing diversity
- Conservation – Whales
- History – Dead Sea scrolls

Classical applications:

- inferring evolutionary relationships between primates

Major approaches in current use:

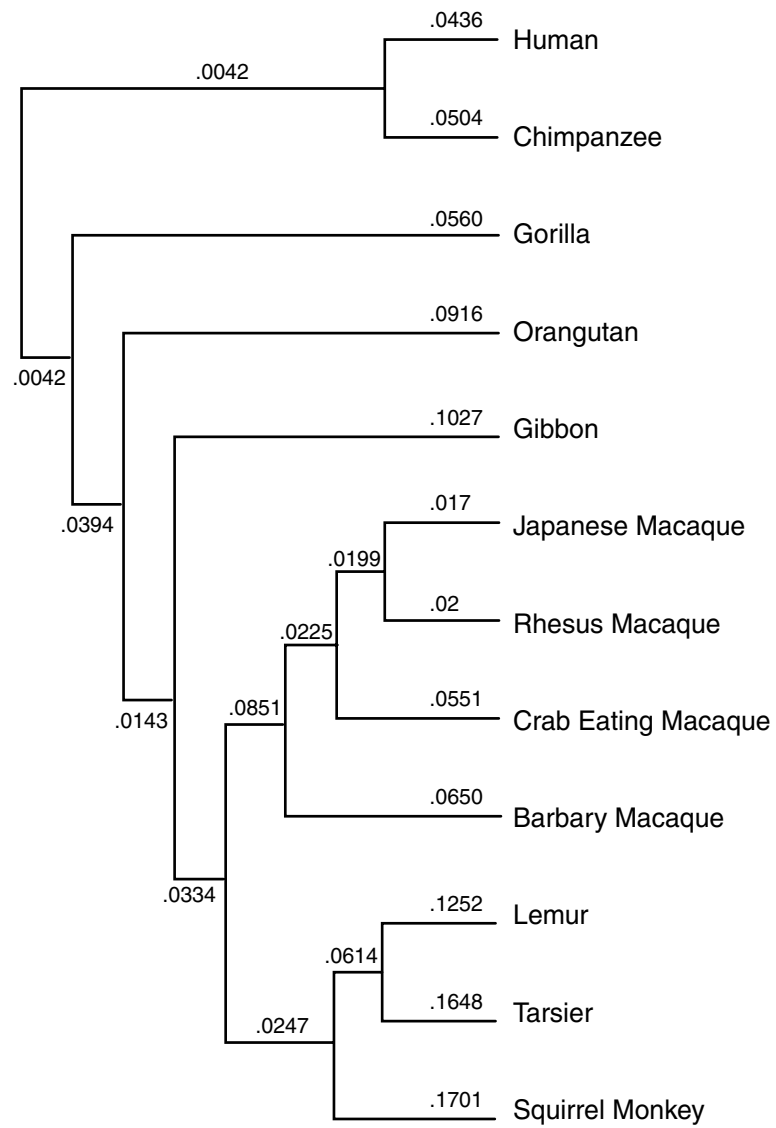
- Maximum Parsimony (MP)
- Distance Methods (D_o , D_a)
- Maximum Likelihood (ML)



Distance methods begin with distance matrix,
pairwise distances between species

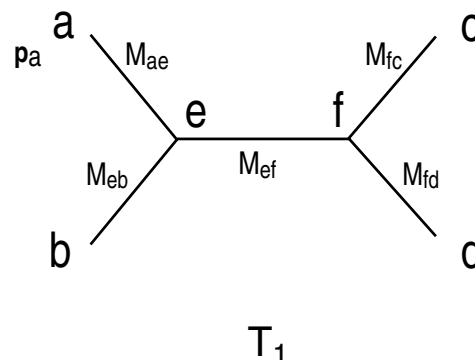
Gor	Orangu	Human	Chimp	Gibbon	CEMac	Lemur	BMacaq	JMacaq	SqMonk	RhMac	...
0	.1890	.1100	.1130	.2150	.3150	.3470	.2850	.2740	.3290	.2710	...
	0	.1790	.1920	.2110	.3170	.3440	.2790	.2890	.3390	.2920	...
		0	.0940	.2050	.2920	.3720	.3040	.2680	.3290	.2710	...
			0	.2140	.3240	.3720	.2920	.2850	.3480	.2980	...
				0	.3080	.3540	.2860	.2930	.3220	.2800	...
					0	.3610	.1350	.0880	.3540	.0990	...
						0	.3430	.3360	.3440	.3470	...
							0	.1370	.3570	.1300	...
										.	
										.	
										.	

Neighbor Joining leads to....



How to find distances between two DNA sequences?

Simplest distance/model of molecular evolution:
Jukes Cantor model



Quick review of elementary Probability:

\mathcal{P}_G = probability that base G occurs at root a .
 $\mathcal{P}_{A|G}$ = probability that a base G mutates to become an A .

Conditional probability

$$\mathcal{P}_{A|G} = \frac{\mathcal{P}(G, A)}{\mathcal{P}_G} = \frac{\mathcal{P}(G \text{ at } a \text{ and } A \text{ at } e)}{\mathcal{P}_G}$$

Explicit Models:

Model base substitutions at a single site

Assume

- i.i.d. – each site is an independent trial of the same probabilistic process
- Markov – probabilities of each substitution along an edge depend only on immediate ancestor base



Model parameters: the tree T

$$\text{root distribution } \mathbf{p}_r = \begin{pmatrix} \mathcal{P}_A \\ \mathcal{P}_G \\ \mathcal{P}_C \\ \mathcal{P}_T \end{pmatrix}$$

Markov matrix $M =$

$$\begin{pmatrix} \mathcal{P}(A|A) & \mathcal{P}(A|G) & \mathcal{P}(A|C) & \mathcal{P}(A|T) \\ \mathcal{P}(G|A) & \mathcal{P}(G|G) & \mathcal{P}(G|C) & \mathcal{P}(G|T) \\ \mathcal{P}(C|A) & \mathcal{P}(C|G) & \mathcal{P}(C|C) & \mathcal{P}(C|T) \\ \mathcal{P}(T|A) & \mathcal{P}(T|G) & \mathcal{P}(T|C) & \mathcal{P}(T|T) \end{pmatrix}$$

Jukes-Cantor Model

Additional assumptions:

- All bases occur with equal probability in the root distribution $\mathbf{p}_r = (.25, .25, .25, .25)$

- All possible base substitutions are equally likely, $A \leftrightarrow G$, $A \leftrightarrow C$, $A \leftrightarrow T$, $G \leftrightarrow T$, etc.

Markov matrix $M_{JC} =$

$$\begin{pmatrix} 1 - \alpha & \frac{\alpha}{3} & \frac{\alpha}{3} & \frac{\alpha}{3} \\ \frac{\alpha}{3} & 1 - \alpha & \frac{\alpha}{3} & \frac{\alpha}{3} \\ \frac{\alpha}{3} & \frac{\alpha}{3} & 1 - \alpha & \frac{\alpha}{3} \\ \frac{\alpha}{3} & \frac{\alpha}{3} & \frac{\alpha}{3} & 1 - \alpha \end{pmatrix}$$

Parameter α is a probability, but also may be interpreted as a rate.

Rate at which observable base substitutions occur over one time step and is measured in units (substitutions per site)/(time step)

Powers M_{JC}^t of the Jukes Cantor Markov matrix give the conditional probabilities after t time steps. Markov matrix $M_{JC}^t =$

$$\begin{pmatrix} \frac{1}{4} + \frac{3}{4}(1 - \frac{4}{3}\alpha)^t & \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & \dots \\ \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & \frac{1}{4} + \frac{3}{4}(1 - \frac{4}{3}\alpha)^t & & \\ \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & & \\ \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & \frac{1}{4} - \frac{1}{4}(1 - \frac{4}{3}\alpha)^t & & \end{pmatrix}$$

(Think: the Jukes-Cantor matrix modeling the evolutionary process (probability of base substitutions) at a single site between two sequences over t time steps.)

(The product of two JC matrices is again JC.)

Let $p(t)$ = the fraction of sites that differ between two sequences S_0 and S_1 .

Then $p(t)$ can be estimated from data (proportion of sites that differ). Using modeling principles,

$$p(t) = \frac{3}{4} - \frac{3}{4}(1 - \frac{4}{3}\alpha)^t$$

Jukes-Cantor distance:

$$p = \frac{3}{4} - \frac{3}{4}\left(1 - \frac{4}{3}\alpha\right)^t$$

Solving for t , gives

$$t = \frac{\ln\left(1 - \frac{4}{3}p\right)}{\ln\left(1 - \frac{4}{3}\alpha\right)}$$

Unrealistic to find either t or α . However,

$$\begin{aligned}\alpha t &= (\text{mutation rate})(\text{no. of time steps}) \\ &= (\text{no. of subst per site/time step}) \\ &\quad (\text{no. of time steps}) \\ &= \text{expected no. substitutions per site} \\ &\quad \text{during the elapsed time}\end{aligned}$$

Approximating $\ln\left(1 - \frac{4}{3}\alpha\right) \approx -\frac{4}{3}\alpha$, then

$$\alpha t \approx -\frac{3}{4} \ln\left(1 - \frac{4}{3}p\right) \equiv d_{JC},$$

where p is the fraction of the sites that disagree in S_0 and S_1 .

Eg. Consider the two aligned sequences

S_0 : ACTTGTCGGATGATCAGCGGTCCATGCACCTGACAACGGT

S_1 : ACATGTTGCTTGACGACAGGTCCATGCGCCTGAGAACGGC

Compute the JC distance between them.

$$d_{JC} \equiv -\frac{3}{4} \ln \left(1 - \frac{4}{3}p \right)$$