

HW 6

1. y_t is not stationary. It follows a random walk, so

$$\begin{aligned} E[y_t] &= E[c_t] + E[s_t] \\ &= E[c_t] \\ &= \mu + \dots \end{aligned}$$

So since the expected value of y_t depends on time, it's not stationary

$$\begin{aligned} 2. \Delta y_t &= \Delta c_t + \Delta s_t \\ &= \mu + v_t + y_t - y_{t-1} \end{aligned}$$

$$\begin{aligned} E[\Delta y_t] &= \mu + E[y_t] - E[y_{t-1}] \\ &= \mu \end{aligned}$$

$$\begin{aligned} \gamma_0 &= \text{var}(\Delta y_t) \\ &= \text{var}(v_t + y_t - y_{t-1}) \\ &= \sigma_v^2 + 2\sigma_u^2 \end{aligned}$$

$$\begin{aligned}
 \gamma_j &= \text{cov}(\Delta y_t, \Delta y_{t-j}) \\
 &= E[(u_t + u_t - u_{t-1})(u_{t-j} + u_{t-j} - u_{t-j-1})] \\
 &= \begin{cases} -\sigma_u^2 & \text{if } j=1 \\ 0 & j>1 \end{cases}
 \end{aligned}$$

$$\rho_j = \frac{\gamma_j}{\gamma_0} = \begin{cases} -\frac{\sigma_u^2}{\sigma_y^2 + 2\sigma_y^2} & \text{if } j=1 \\ 0 & j>1 \end{cases}$$

3. Since $\rho_j = 0$ for $j > 1$

$$\Rightarrow \Delta y_t = \mu + \varepsilon_t + \theta \varepsilon_{t-1}$$

$$\text{where } \varepsilon_t \sim \text{i.i.d. } N(0, \sigma_\varepsilon^2)$$

So it implies the autocorrelation is $MA(1)$

4) We have

$$\gamma_0 = (1 + \theta^2) \sigma_\varepsilon^2$$

$$\gamma_1 = \theta \sigma_\varepsilon^2$$

Which implies

$$\sigma_v^2 + 2\sigma_u^2 = (1+\theta^2)\sigma_v^2$$

$$- \sigma_u^2 = \theta\sigma_v^2$$

$$\Rightarrow \sigma_v^2 = -\sigma_u^2/\theta$$

$$\text{and } \sigma_u^2 \theta^2 + (\sigma_v^2 + 2\sigma_u^2) \theta + \sigma_u^2 = 0$$

$$\Rightarrow \theta = \frac{-(\sigma_v^2 + 2\sigma_u^2) \pm \sqrt{(\sigma_v^2 + 2\sigma_u^2)^2 - 4\sigma_u^2}}{2\sigma_u^2}$$

ADL (1, 1)

$$1. (1 - \phi L) y_t = \alpha + (\beta_0 + \beta_1 L) x_t + \varepsilon_t$$

$$\phi L y_t = \alpha + \beta_1 L x_t + \varepsilon_t$$

Assume $|\phi| < 1$, so y_t is stationary

$$E[y_t] = \alpha + \phi E[y_{t-1}] + \beta_0 E[x_t] + \beta_1 E[x_{t-1}] + E[\varepsilon_t]$$

y is stationary, so $E[y_{t-1}] = E[y_t]$

$$\Rightarrow (1 - \phi) E[y_t] = \alpha + (\beta_0 + \beta_1) E[x_t]$$

$$\Rightarrow E[y_t] = \frac{\alpha + (\beta_0 + \beta_1) E[x_t]}{1 - \phi}$$

$$E[y_t] = \frac{\alpha}{1 - \phi} + \frac{(\beta_0 + \beta_1)}{1 - \phi} E[x_t]$$

$$2. y_t - y_{t-1} = \alpha + \phi y_{t-1} + \beta_0 x_t + \beta_1 x_{t-1} + \varepsilon_t - y_{t-1}$$

$$\Delta y_t = \alpha + (\phi - 1) y_{t-1} + \beta_0 x_t + \beta_1 x_{t-1} + \varepsilon_t$$

$$\Delta y_t = \alpha + (\phi - 1) y_{t-1} + \beta_0 x_t - \beta_0 x_{t-1} + \beta_1 x_{t-1} + \varepsilon_t + \beta_1 x_t$$

$$\Delta y_t = \beta_0 \Delta x_t + (\phi - 1) y_{t-1} + \alpha + \beta_0 x_{t-1} + \beta_1 x_{t-1} + \varepsilon_t$$

$$\Delta y_t = \beta_0 \Delta x_t - (1 - \phi) \left(\frac{\alpha}{1 - \phi} + \frac{(\beta_0 + \beta_1)}{1 - \phi} x_{t-1} \right) + \varepsilon_t$$

$$\begin{aligned}
 3. \quad E[\Delta y_t] &= \beta_0 E[\Delta x_t] - (1-\phi) E\left[y_{t-1} - \frac{\alpha}{1-\phi} + \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1}\right] \\
 &= \beta_0 E[\Delta x_t] - (1-\phi) \left[E[y_{t-1}] - E\left[\frac{\alpha}{1-\phi} + \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1}\right] \right] \\
 &= \beta_0 E[\Delta x_t] \quad \text{from part 1}
 \end{aligned}$$

4,

$$\begin{aligned}
 \Delta y_t &= \beta_0 \Delta x_t - (1-\phi) \left(y_{t-1} - \frac{\alpha}{1-\phi} - \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1} \right) + \varepsilon_t \\
 &= \beta_0 \Delta x_t - (1-\phi) (y_{t-1} - y_{t-1}) + 0 \\
 &= \beta_0 \Delta x_t
 \end{aligned}$$

$$5. \quad \varepsilon_t = 0 \quad y_{t-1} > \frac{\alpha}{1-\phi} + \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1}$$

$$\begin{aligned}
 \Delta y_t &= \beta_0 \Delta x_t - (1-\phi) \left(y_{t-1} - \frac{\alpha}{1-\phi} - \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1} \right) + \varepsilon_t \\
 &= \beta_0 \Delta x_t - (1-\phi) \left(y_{t-1} - \frac{\alpha}{1-\phi} - \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1} \right) \\
 &= \beta_0 \Delta x_t - (1-\phi) (+) \\
 &\quad +
 \end{aligned}$$

$$\Rightarrow \Delta y_t < \beta_0 \Delta x_t$$

$$y_{t-1} > (\cdot)$$

$$\Delta y_t = \beta_0 \Delta x_t - (1-\phi) \left(y_{t-1} - \frac{\alpha}{1-\phi} - \frac{\rho_0 + \beta_1}{1-\phi} x_{t-1} \right)$$

$$= B_0 \Delta x_t - \underbrace{(1-\phi) C_-}_{C_-}$$

$$=) \Delta y_t > \Delta \Delta x_t$$

Problem 4

$$1) \quad LRV = r_0 + 2 \sum_{j=1}^{\infty} r_j$$

$$\text{we have } (a/v/v|1) \quad r_0 = \frac{\sigma^2}{1-\phi^2}$$

$$r_j = \phi r_{j-1}$$

$$\begin{aligned} \Rightarrow LRV &= r_0 + 2 \sum_{j=1}^{\infty} \phi^j r_0 \\ &= r_0 + \frac{2\phi}{1-\phi} r_0 \\ &= \frac{1+\phi}{1-\phi} r_0 \\ &= \frac{\sigma^2}{(1-\phi)^2} \end{aligned}$$

2) Sec cond

Hass 14.20)

SoR cod;

HW6

2023-05-23

```
# Load packages
pacman::p_load(dynlm, sandwich, lmtest, here)
```

2)

```
# Set parameters
mu = 1
phi = 0.75
sigma = 1
n = 500

# Set seed and simulate process
set.seed(101)
y = mu + arima.sim(model = list(ar = phi), n = n)

# Using formula derived in handwritten notes, calculate long run variance
(LRV = sigma^2/(1-phi)^2)
```

```
## [1] 16
```

```
# Calculate standard error
(se = sqrt(LRV/n))
```

```
## [1] 0.1788854
```

3)

```
# Fit AR(1) model
ar1 = dynlm(y ~ L(y,1))
summary(ar1)
```

```
##
## Time series regression with "ts" data:
## Start = 2, End = 500
##
## Call:
## dynlm(formula = y ~ L(y, 1))
##
## Residuals:
```

```
##      Min      1Q   Median      3Q      Max
## -3.11444 -0.65392 -0.00975  0.61632  2.65135
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.18392    0.04817   3.818 0.000152 ***
## L(y, 1)      0.75143    0.02958  25.406 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.9599 on 497 degrees of freedom
## Multiple R-squared:  0.565, Adjusted R-squared:  0.5641
## F-statistic: 645.5 on 1 and 497 DF, p-value: < 2.2e-16
```

```
# Calculate fitted parameters
phi_hat = ar1$coefficients[2]
sigma_hat = mean(ar1$residuals^2)
```

```
# Use these to calculate LRV
(LRV_hat = sigma_hat^2/(1-phi)^2)
```

```
## [1] 13.47752
```

```
# Calculate standard error
(se_hat = sqrt(LRV/n))
```

```
## [1] 0.1788854
```

5)

```
# Regress y on a constant to calculate newey-west standard errors
new = dynlm(y ~ 1)
```

```
# Calculate newey-west se
(newey = coeftest(new, vcov=NeweyWest(new, prewhite=FALSE)))
```

```
##
## t test of coefficients:
##
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)  0.73516    0.14888   4.9379 1.079e-06 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
(se_new = newey[2])
```

```
## [1] 0.1488822
```

```
# Calculate LRV using these ses
(LRV_new = 500*se_new^2)
```

```
## [1] 11.08296
```

Hansen 14.20

a

```
# Load data
fred_df = haven::read_dta(here("HW6", "FRED-MD.dta"))
# Make unemployment rate convenient time series
unrate = ts(fred_df$unrate, frequency=12, start=1959)

# Function for calculating various lengths of AR models and their AICs
ar_sim = function(p){
  mod = dynlm(unrate ~ L(unrate, 1:p), start=c(1960,1)) # p is number of lags
  aic = AIC(mod)
  return(list(mod, aic))
}

# Calculate for 1 through 8 lags
(models = lapply(1:8, ar_sim))
```

```
## [[1]]
## [[1]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)
##           0.03113           0.99456
##
##
## [[1]][[2]]
## [1] -420.3042
##
##
## [[2]]
## [[2]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
```

```

## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2
##           0.03672           1.11070           -0.11702
##
##
## [[2]][[2]]
## [1] -427.9882
##
##
## [[3]]
## [[3]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2  L(unrate, 1:p)3
##           0.05035           1.08020           0.17781           -0.26649
##
##
## [[3]][[2]]
## [1] -477.8072
##
##
## [[4]]
## [[4]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2  L(unrate, 1:p)3
##           0.06106           1.03292           0.20937           -0.07172
## L(unrate, 1:p)4
##           -0.18079
##
##
## [[4]][[2]]
## [1] -499.298
##
##
## [[5]]
## [[5]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:

```

```

## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2  L(unrate, 1:p)3
##      0.072216      1.003344      0.196562      -0.036887
## L(unrate, 1:p)4  L(unrate, 1:p)5
##      -0.004773      -0.170310
##
##
## [[5]][[2]]
## [1] -518.2109
##
##
## [[6]]
## [[6]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2  L(unrate, 1:p)3
##      0.07933      0.98759      0.19671      -0.04076
## L(unrate, 1:p)4  L(unrate, 1:p)5  L(unrate, 1:p)6
##      0.01298      -0.07677      -0.09300
##
##
## [[6]][[2]]
## [1] -522.3805
##
##
## [[7]]
## [[7]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2  L(unrate, 1:p)3
##      0.08413      0.98224      0.19235      -0.03967
## L(unrate, 1:p)4  L(unrate, 1:p)5  L(unrate, 1:p)6  L(unrate, 1:p)7
##      0.01036      -0.06572      -0.03616      -0.05745
##
##
## [[7]][[2]]
## [1] -522.7277
##
##
## [[8]]

```

```
## [[8]][[1]]
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:p), start = c(1960, 1))
##
## Coefficients:
##      (Intercept)  L(unrate, 1:p)1  L(unrate, 1:p)2  L(unrate, 1:p)3
##      0.0841171      0.9822538      0.1923565      -0.0396568
## L(unrate, 1:p)4  L(unrate, 1:p)5  L(unrate, 1:p)6  L(unrate, 1:p)7
##      0.0103606      -0.0657103      -0.0361923      -0.0576041
## L(unrate, 1:p)8
##      0.0001603
##
##
## [[8]][[2]]
## [1] -520.7277
```

b

```
# Print AICs
(aic = sapply(1:8, function(x){models[[x]][[2]]}))
```

```
## [1] -420.3042 -427.9882 -477.8072 -499.2980 -518.2109 -522.3805 -522.7277
## [8] -520.7277
```

c

```
# Find minimum AIC
which.min(aic) # the AR(7) model has the lowest AIC
```

```
## [1] 7
```

d

```
summary(dynlm(unrate ~ L(unrate, 1:7), start=c(1960,1)))
```

```
##
## Time series regression with "ts" data:
## Start = 1960(1), End = 2017(12)
##
## Call:
## dynlm(formula = unrate ~ L(unrate, 1:7), start = c(1960, 1))
##
## Residuals:
```

```

##      Min      1Q   Median      3Q      Max
## -0.56971 -0.10061 -0.00689  0.09910  0.72160
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    0.08413    0.02541   3.311 0.000977 ***
## L(unrate, 1:7)1  0.98224    0.03769  26.064 < 2e-16 ***
## L(unrate, 1:7)2  0.19235    0.05285   3.640 0.000294 ***
## L(unrate, 1:7)3 -0.03967    0.05327  -0.745 0.456728
## L(unrate, 1:7)4  0.01036    0.05329   0.194 0.845852
## L(unrate, 1:7)5 -0.06572    0.05331  -1.233 0.218051
## L(unrate, 1:7)6 -0.03616    0.05289  -0.684 0.494375
## L(unrate, 1:7)7 -0.05745    0.03768  -1.524 0.127845
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.165 on 688 degrees of freedom
## Multiple R-squared:  0.9892, Adjusted R-squared:  0.9891
## F-statistic: 9043 on 7 and 688 DF, p-value: < 2.2e-16

```