

Module 4: Factorial Treatment Structure

Analyzing a Factorial: Model, ANOVA, and Decision Flow

Example 4.2: Bakery

Treatment Structure: 3×2 Full Factorial

- Shelf Height (Bottom, Middle, Top)
- Shelf Width (Regular, Wide)

Design Structure: CRD with $r = 2$ stores per treatment combination.

Response: Bread Sales



What changes when we add a factor?

- We add:
 - Another main effect
 - An interaction term
- We must decide what effects are present *before* interpreting the treatment means.

The goal of the analysis is to decide which effects matter.

The Treatment Effects Model (Two-way Factorial)

$$y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \epsilon_{ijk} \text{ with } \epsilon_{ijk} \sim \text{iid } N(0, \sigma^2)$$

for $i = 1, 2, 3, \dots, a; j = 1, 2, \dots, b; k = 1, 2, \dots, r$

ANOVA for Two-way Factorials

Just like one-way ANOVA: $SST = SSTrt + SSE$

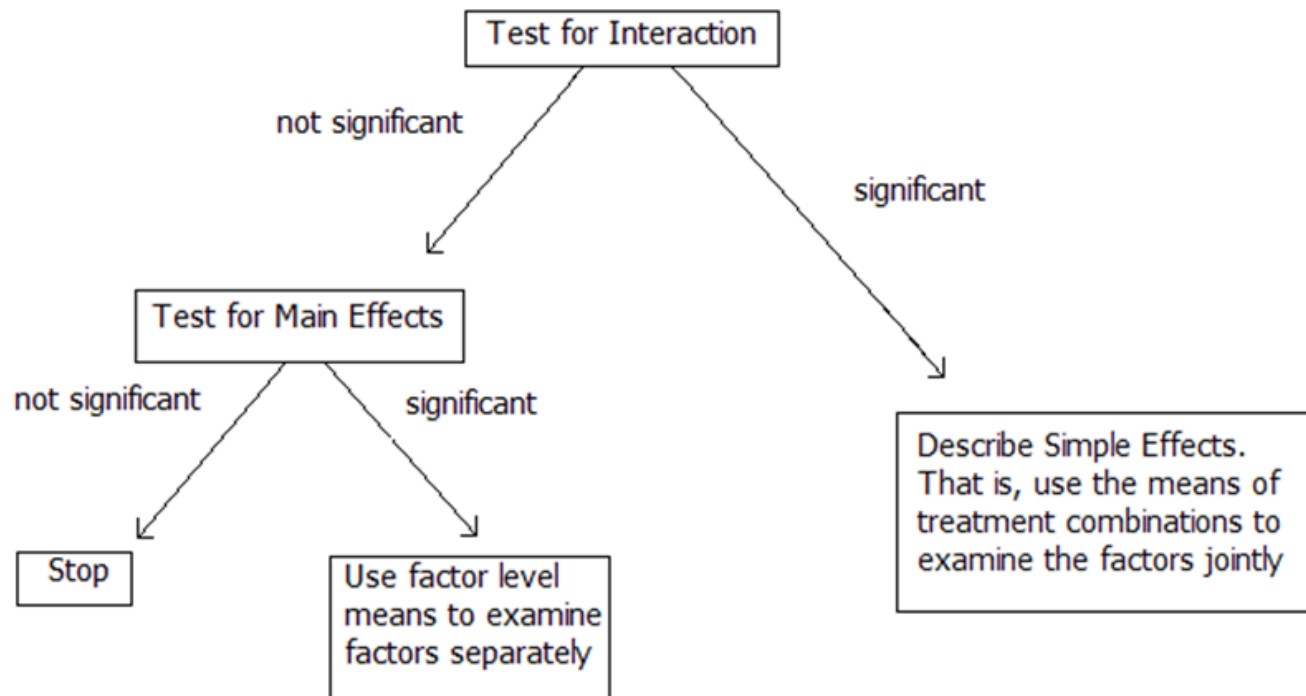
For a factorial: $SSTrt = SSA + SSB + SSAB$

- Main effect of factor A (Height): $SSA = rb \sum_i (\bar{y}_{i..} - \bar{y}_{..})^2$
- Main effect of factor B (Width): $SSB = ra \sum_j (\bar{y}_{.j.} - \bar{y}_{..})^2$
- AB Interaction Effect (Height x Width): $SSAB = SST - SSE - SSA - SSB$

Full ANOVA Table (Two-way)

Source (SV)	df	SS	MS	F	
A	$a - 1$	SSA	$MSA = \frac{SSA}{a-1}$	$\frac{MSA}{MSE}$	Do avg sales differ across shelf heights?
B	$b - 1$	SSB	$MSB = \frac{SSB}{b-1}$	$\frac{MSB}{MSE}$	Do avg sales differ across shelf widths?
AB	$(a - 1)(b - 1)$	SSAB	$MSAB = \frac{SSAB}{(a-1)(b-1)}$	$\frac{MSAB}{MSE}$	Does the effect of width <i>depend</i> on height?
Error: e.u.(AB)	$(r - 1)ab$	SSE	$MSE = \frac{SSE}{(r-1)ab}$		
Total	$N - 1$	SST			

Decision Flowchart



Example 4.2: Skeleton ANOVA

Source of Variation $DF = 12 \text{ stores} - 1 = 11 \text{ total df}$

R: Fitting the Model

Let's prep the data...

```
1 bakery_data <- read_csv("data/04_bakery_data.csv") %>%  
2   mutate(height = factor(height, levels = c("bottom", "middle", "top")),  
3         width = factor(width, levels = c("regular", "wide"))  
4 head(bakery_data)
```

```
# A tibble: 6 × 4  
  height width  sales placement  
  <fct> <fct>  <dbl> <chr>  
1 bottom regular    47 bottomregular  
2 bottom regular    43 bottomregular  
3 bottom wide      46 bottomwide  
4 bottom wide      40 bottomwide  
5 middle regular    62 middleregular  
6 middle regular    68 middleregular
```

```
1 levels(bakery_data$height)
```

```
[1] "bottom" "middle" "top"
```

```
1 levels(bakery_data$width)
```

```
[1] "regular" "wide"
```

R: Fitting the Model

```
1 options(contrasts = c("contr.sum", "contr.poly"))
2 bakery_mod <- lm(sales ~ height + width + height:width, data = bakery_data)
```

```
1 anova(bakery_mod)
```

Analysis of Variance Table

Response: sales

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
height	2	1544	772.00	74.7097	5.754e-05 ***
width	1	12	12.00	1.1613	0.3226
height:width	2	24	12.00	1.1613	0.3747
Residuals	6	62	10.33		

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```
1 summary(bakery_mod)
```

Call:

lm(formula = sales ~ height + width + height:width, data = bakery_data)

Residuals:

Min	1Q	Median	3Q	Max
-3	-2	0	2	3

Coefficients:

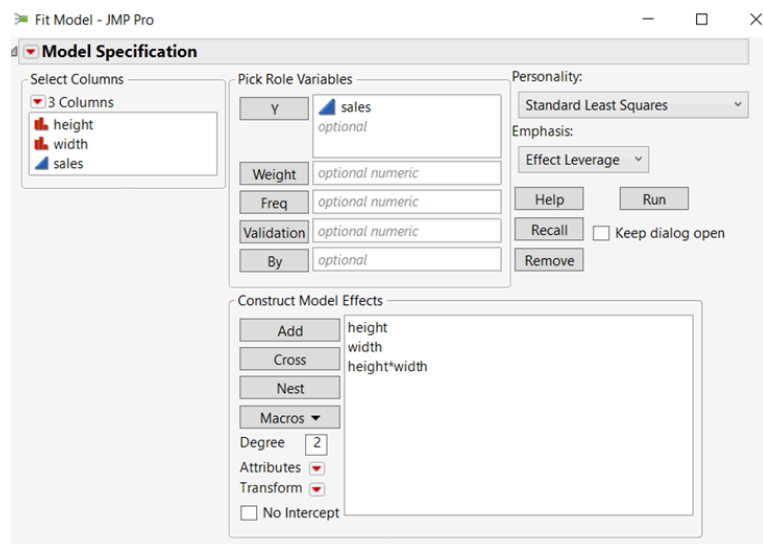
	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	51.000	0.928	54.959	2.44e-09 ***
height1	-7.000	1.312	-5.334	0.00177 **
height2	16.000	1.312	12.192	1.85e-05 ***
width1	-1.000	0.928	-1.078	0.32261
height1:width1	2.000	1.312	1.524	0.17835
height2:width1	-1.000	1.312	-0.762	0.47494

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.215 on 6 degrees of freedom
Multiple R-squared: 0.9622, Adjusted R-squared: 0.9308
F-statistic: 30.58 on 5 and 6 DF, p-value: 0.0003384

JMP: Fitting the Model

▼ Analyze > Fit Model > Assign Y = Response + Highlight both treatment factors and click Macros > Full Factorial



Analysis of Variance

Source	DF	Sum of Squares	Mean Square	F Ratio
Model	5	1580.0000	316.000	30.5806
Error	6	62.0000	10.333	Prob > F
C. Total	11	1642.0000		0.0003*

Effect Tests

Source	Nparm	DF	Sum of Squares	F Ratio	Prob > F
height	2	2	1544.0000	74.7097	<.0001*
width	1	1	12.0000	1.1613	0.3226
height*width	2	2	24.0000	1.1613	0.3747

JMP: Fitting the Model

▼ Response > Expanded Estimates

Expanded Estimates				
Nominal factors expanded to all levels				
Term	Estimate	Std Error	t Ratio	Prob> t
Intercept	51	0.927961	54.96	<.0001*
height[bottom]	-7	1.312335	-5.33	0.0018*
height[middle]	16	1.312335	12.19	<.0001*
height[top]	-9	1.312335	-6.86	0.0005*
width[regular]	-1	0.927961	-1.08	0.3226
width[wide]	1	0.927961	1.08	0.3226
height[bottom]*width[regular]	2	1.312335	1.52	0.1783
height[bottom]*width[wide]	-2	1.312335	-1.52	0.1783
height[middle]*width[regular]	-1	1.312335	-0.76	0.4749
height[middle]*width[wide]	1	1.312335	0.76	0.4749
height[top]*width[regular]	-1	1.312335	-0.76	0.4749
height[top]*width[wide]	1	1.312335	0.76	0.4749

Recall, “sum to zero”

$\mu^{\wedge} =$

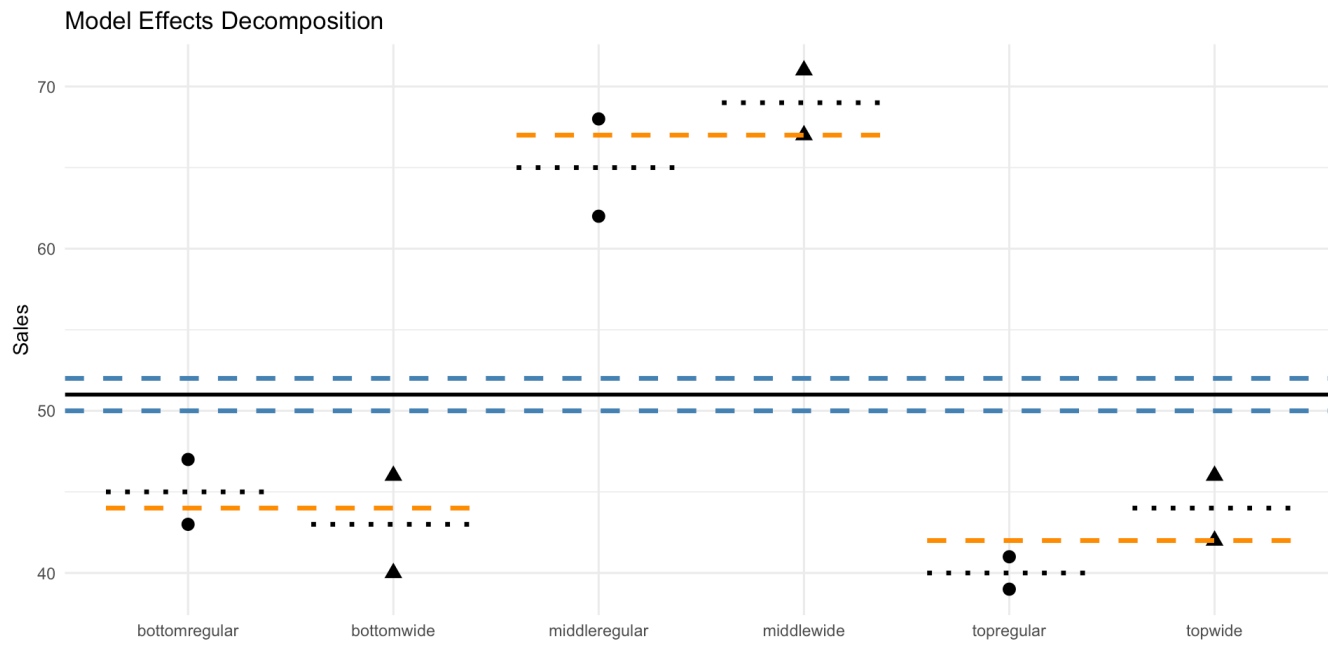
Height	
Bottom	$\alpha_1^{\wedge} = 7$
Middle	$\alpha_2^{\wedge} = 16$
Top	$\alpha_3^{\wedge} =$

Width	
Regular	$\beta_1^{\wedge} = -1$
Wide	$\beta_2^{\wedge} =$

	Regular	Wide
Bottom	$\widehat{\alpha\beta}_{11} = 2$	$\widehat{\alpha\beta}_{12} =$
Middle	$\widehat{\alpha\beta}_{21} = -1$	$\widehat{\alpha\beta}_{22} =$
Top	$\widehat{\alpha\beta}_{31} =$	$\widehat{\alpha\beta}_{32} =$

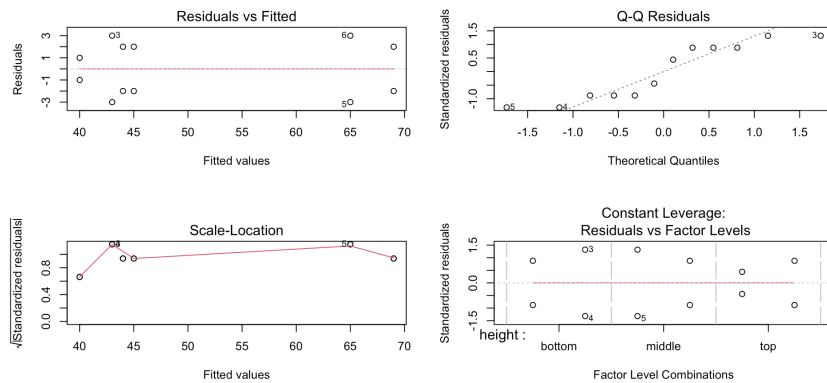
Decomposition of Model Effects

Recall our statistical effects model: $y_{ijk} = \mu + \alpha_i + \beta_j + \alpha\beta_{ij} + \epsilon_{ijk}$



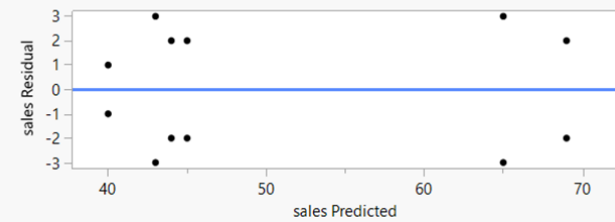
Model Diagnostics

```
1 par(mfrow = c(2,2))
2 plot(bakery_mod)
```



```
1 par(mfrow = c(1,1))
```

Residual by Predicted Plot



Residual Normal Quantile Plot

