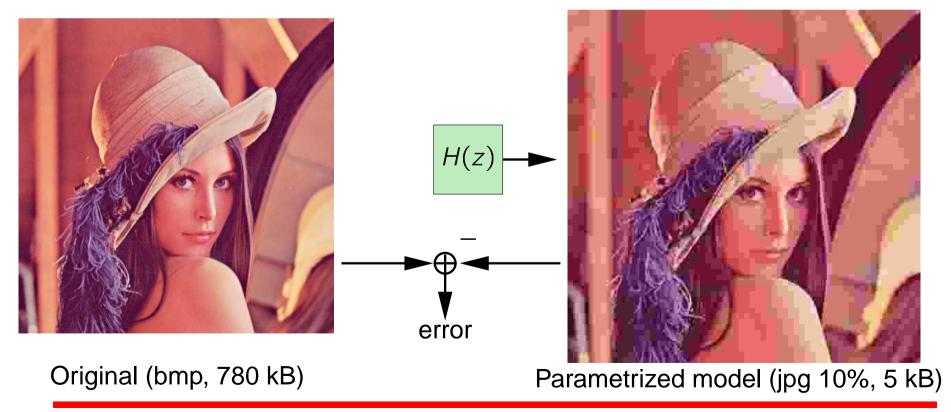
Digital Signal Processing (ET 4235)

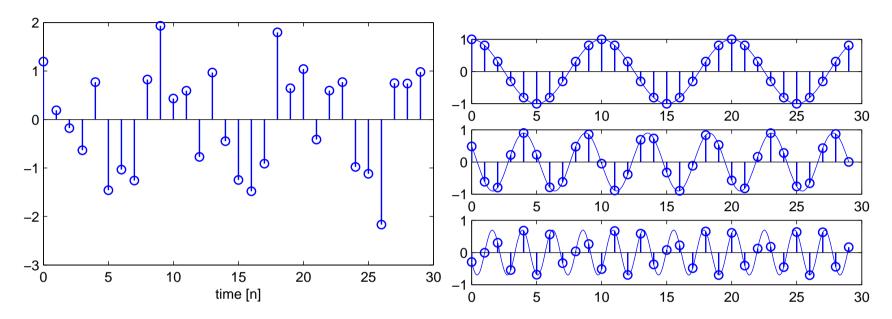
Chapter 4: Signal Modeling

Given a signal (set of samples), how can it be modeled using a filter?



Signal Modeling—Motivation

Motivation 1: Efficient transmission/storage

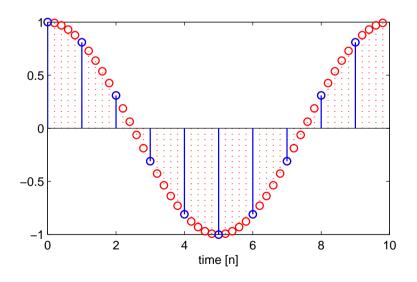


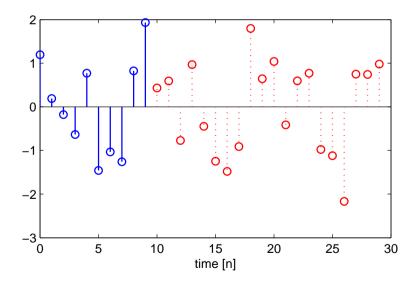
- Direct: store all samples
- Coded:
 - Model the signal, e.g., sum of sinusoids: $\sum \alpha_k \cos(n\omega_k + \phi_k)$
 - Estimate the parameters of the signal, $\{\alpha_k, \omega_k, \phi_k\}$
 - Store the parameters instead of the original samples.

Example: GSM speech coding, MP3 audio coding, JPEG image coding, · · ·

Motivation 2: Interpolation/extrapolation

Interpolation/extrapolation requires a model, e.g., bandlimited/lowpass, sum of sinusoids, etc.



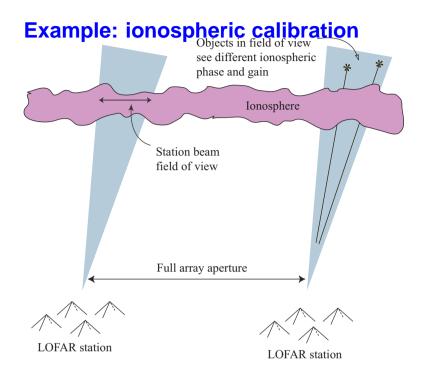


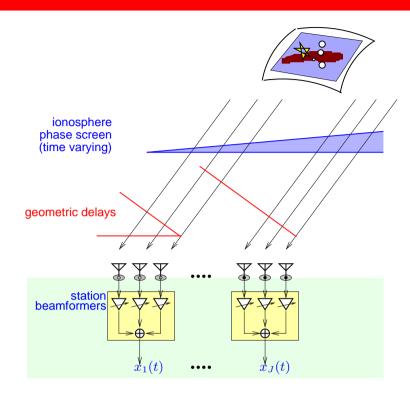
Extrapolation of Lena



Horizontal extrapolation:

(In fact, there are stationarity requirements that would prohibit this...)

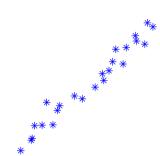




- The ionosphere causes small delays in the reception of signals that modify the apparent direction of astronomical sources.
- Low frequency radio telescopes can 'sample' the ionosphere in the direction of calibration sources. For other directions, we rely on interpolation.
- A simple ionospheric model specifies correlations in phase delay τ as function of distance between points: $\mathbf{C}_{\tau}(\mathbf{x}_1, \mathbf{x}_2) = 1 \alpha ||\mathbf{x}_1 \mathbf{x}_2||^{\beta}$

Example: Interpolation of correlated variables

Suppose we have samples of random variables (x, y) that are correlated. Given a new sample x, can we predict the corresponding y?



Pose the model: $y = \alpha x$, and estimate α by minimizing $E|y - \alpha x|^2$:

$$E\{xy\} = \alpha E\{xx\} \quad \Leftrightarrow \quad r_{xy} = \alpha r_{xx} \quad \Leftrightarrow \quad \alpha = \frac{r_{xy}}{r_{xx}}$$

If we stack the measured samples in vectors **x** and **y**, then we obtain the estimate

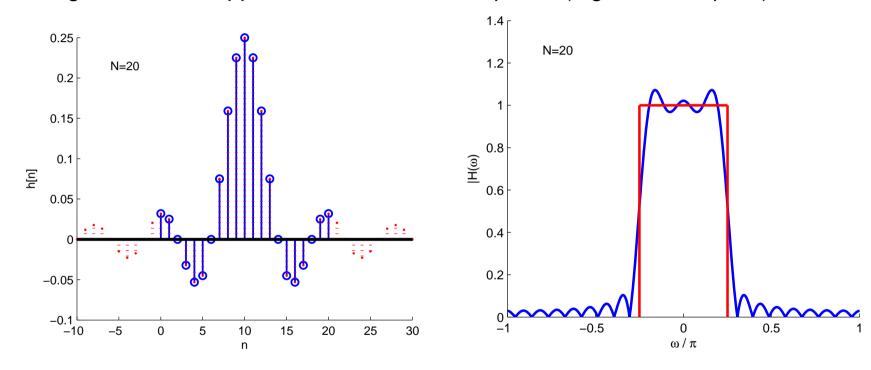
$$\hat{r}_{xx} = \frac{1}{N} \mathbf{x}^H \mathbf{x}$$
, $\hat{r}_{xy} = \frac{1}{N} \mathbf{x}^H \mathbf{y}$ \Rightarrow $\hat{\alpha} = \frac{\mathbf{x}^H \mathbf{y}}{\mathbf{x}^H \mathbf{x}}$

(cf. the Wiener filter in Ch. 7.) This is the solution of the Least Squares problem

$$\min_{\alpha} \|\mathbf{y} - \alpha \mathbf{x}\|^2$$

Motivation 3: Filter design

Design a filter that approximates a desired response (e.g., ideal lowpass)



The design depends on the filter model (FIR, IIR), filter order, error criteria, etc.

Signal models

Deterministic:

Signal models

standard input signal model signal signal observations impulse $\delta[n]$ \longrightarrow H(z) \longrightarrow $\hat{x}[n]$ (or known signal)

Stochastic: white noise $v[n] \longrightarrow H(z)$ $\hat{x}[n]$

Models for H(z)

ARMA(p, q):
$$H(z) = \frac{B(z)}{A(z)} = \frac{\sum_{k=0}^{q} b(k)z^{-k}}{1 + \sum_{k=1}^{p} a(k)z^{-k}}$$

Special cases: AR(p), MA(q)

Signal models

Model identification

Given observations x[n], $n=0,\cdots,N-1$ and filter order p,q, find the parameters of H(z) such that the modeled output signal $\hat{x}[n]=h[n]$ best matches the observations.

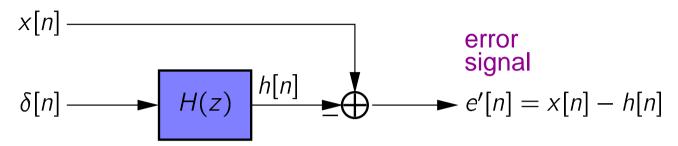
For stochastic signals, we will try to match the correlation sequences: $r_{\times}[k] = r_h[k]$.

Issues:

- stability of the filter (in case p > 0)
- computational complexity for parameter estimation
- model order selection
- error criterion for the approximation

Model identification via Least Squares

In the next slides, we consider deterministic input signals (impulse $\delta[n]$). The model is a filter H(z): LTI, causal, rational. The desired signal has x[n] = 0, n < 0.



- "Minimize the error" depends on the definition of error, and the norm.
 - Least squares: $\mathcal{E}_{LS} = \sum_{n=0}^{\infty} |e'[n]|^2$
- The minimization of $\mathcal{E}_{LS} = \sum |x[n] h[n]|^2$ with $H(z) = \frac{B(z)}{A(z)} = \frac{\sum_{k=0}^q b[k]z^{-k}}{1 + \sum_{k=1}^p a[k]z^{-k}}$ requires

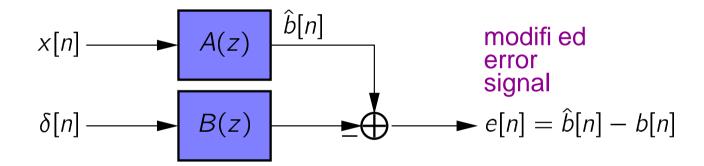
$$\begin{cases} \frac{\partial \mathcal{E}_{LS}}{\partial a^*[k]} = 0, & k = 1, \dots, p \\ \frac{\partial \mathcal{E}_{LS}}{\partial b^*[k]} = 0, & k = 0, \dots, q \end{cases}$$

The resulting p + q + 1 equations are nonlinear because of the division by A(z).

Model identification via Least Squares

Alternative that leads to tractable results: consider "weighted" error

$$E(z) = A(z)E'(z) = A(z)X(z) - B(z)$$



- Now A(z) is in the nominator, the error equation is linear.
- Techniques that are based on this: Pade Approximation, Prony's Method, Shank's Method.

Model identification via Pade Approximation

Pade approximation

We have p + q + 1 model parameters: Can match p + q + 1 signal samples exactly:

$$h[n] = x[n], \quad n = 0, \dots, p + q$$

■ How to find h[n] in terms of the parameters:

$$H(z) = \frac{B(z)}{A(z)} \Rightarrow H(z)A(z) = B(z) \Rightarrow h[n] * a[n] = b[n]$$

$$\Rightarrow h[n] + \sum_{k=1}^{p} a[k]h[n-k] = b[n]$$

where h[n] = 0, n < 0, and b[n] = 0, n < 0 or n > q

■ Match exactly with x[n] for $n = 0, \dots, p + q$:

$$x[n] + \sum_{k=1}^{p} a[k]x[n-k] = \begin{cases} b[n], & n = 0, \dots, q \\ 0, & n = q+1, \dots, p+q \end{cases}$$

Model identification via Pade Approximation

Write these equations in matrix form:

$$\begin{bmatrix}
x[0] & 0 & \cdots & 0 \\
x[1] & x[0] & \cdots & 0 \\
x[2] & x[1] & \cdots & x[0] \\
x[q] & x[q-1] & \cdots & x[q-p] \\
\hline
x[q+1]x[q] & \cdots & x[q-p] \\
x[q+p]x[q+p-1] & x[q]
\end{bmatrix} = \begin{bmatrix}
b[0] \\
b[1] \\
b[2] \\
\vdots \\
a[p]
\end{bmatrix}$$

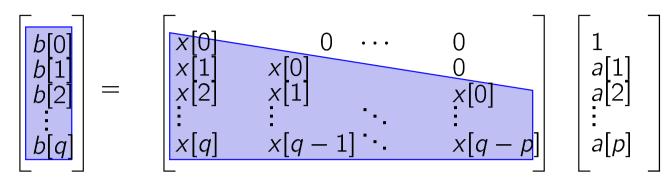
■ Take the submatrix that does not involve the b[k], and first solve for the a[k]:

This is a square $p \times p$ matrix equation: $\mathbf{X}_q \mathbf{\bar{a}} = -\mathbf{x}_{q+1}$.

The solution is $\bar{\mathbf{a}} = -\mathbf{X}_q^{-1}\mathbf{x}_{q+1}$.

Model identification via Pade Approximation

Plug the solution back to find the b[k]:



$$\mathbf{b} = \mathbf{X}_0 \mathbf{a}$$

Disadvantages of the Pade approximation

- The resulting model doesn't have to be stable
- Outside the used interval [0, p+q], the approximation may be very bad
- If X_q is singular, then a solution does not always exist that matches all signal values. (Sometimes a solution exists of lower order for A(z), B(z), i.e., smaller p, q, but then no guarantee on matching for samples beyond p + q.)

Derivation of Prony's Method

As before, E(z) = X(z)A(z) - B(z), i.e.,

$$e[n] = \begin{cases} x[n] + \sum_{k=1}^{p} a[k]x[n-k] - b[n], & n = 0, \dots, q \\ x[n] + \sum_{k=1}^{p} a[k]x[n-k], & n > q \end{cases}$$

- For Pade, we first solved e[n] = 0, $n = q + 1, \dots, q + p$ to find A(z).
- Prony: solve $\min_{\{a[k]\}} \sum_{n=q+1}^{\infty} |e[n]|^2 = \min_{\{a[k]\}} \sum_{n=q+1}^{\infty} |x[n] + \sum_{k=1}^{p} a[k]x[n-k]|^2$

In matrix form:

$$\min \left\| \begin{bmatrix} x[q] & x[q-1] & \dots & x[q-p+1] \\ x[q+1] & x[q] & \dots & x[q-p+2] \\ \vdots & \vdots & \vdots & \vdots \\ x[q+p-1]x[q+p-2] & \dots & x[q] \\ \hline x[q+p] & \dots & x[q+1] \\ x[q+p+1] & \dots & x[q+2] \\ \vdots & \vdots & \vdots & \vdots \\ a[p] \end{bmatrix} + \begin{bmatrix} x[q+1] \\ x[q+2] \\ x[q+p] \\ x[q+p+1] \\ x[q+p+2] \end{bmatrix} \right\|^{2}$$

$$\min \|\mathbf{X}_q \bar{\mathbf{a}} - (-\mathbf{x}_{q+1})\|^2$$

(Now, X_q refers to the infinite-dimensional matrix.)

- This is a Least-Squares problem of an overdetermined system of equations.
- The solution is $\bar{\mathbf{a}} = -\mathbf{X}_q^{\dagger}\mathbf{x}_{q+1} = -(\mathbf{X}_q^H\mathbf{X}_q)^{-1}\mathbf{X}_q^H\mathbf{x}_{q+1} = -\mathbf{R}_{\chi}^{-1}\mathbf{r}_{\chi}$. This assumes that $\mathbf{R}_{\chi} := \mathbf{X}_q^H\mathbf{X}_q$ is not singular.
- In practice, \mathbf{X}_q and \mathbf{x}_{q+1} are of finite size as determined by the available data.

$$\mathbf{R}_{\mathsf{x}} := \mathbf{X}_{q}^{\mathsf{H}} \mathbf{X}_{q} \qquad (p \times p \; \mathsf{matrix})$$

- **R**_x is positive (semi)definite by construction. We will see later that this makes A(z) (marginally) stable.
- If \mathbf{R}_{\times} is singular, then this indicates that the filter order can be reduced.
- After **a** is known, we find **b** (i.e., B(z)) precisely as in Pade's method: **b** = X_0 **a**. This makes e[n] = 0, $n = 0, \dots, q$.

Alternatively, we can find the numerator by minimizing e[n] over the entire data record. For the error criterion based on e[n], this does not make a difference. Shank's method switches back to e'[n] and minimizes over the entire data.

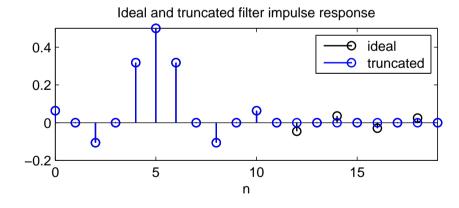
Suppose we want to design an ideal linear phase lowpass filter:

$$G(e^{j\omega}) = \left\{ egin{array}{ll} e^{-jn_d\omega}, & |\omega| < \pi/2 \ 0, & ext{otherwise} \end{array}
ight.$$

 n_d is the filter delay. The corresponding impulse response is

$$g[n] = \frac{\sin[(n-n_d)\pi/2]}{(n-n_d)\pi}$$

■ We will match p + q + 1 = 11 values, and choose $n_d = 5$.



■ Two cases: FIR filter (p = 0, q = 10), ARMA filter (p = 5, q = 5).

Pade approximation

Filter coefficients to match:

$$\mathbf{g} = \begin{bmatrix} 0.064 & -0.000 & -0.106 & 0.000 & 0.318 & 0.5 & 0.318 & 0.000 & -0.106 & -0.000 & 0.064 \end{bmatrix}$$

■ FIR filter (p = 0, q = 10):

$$h[n] = \begin{cases} g[n], & 0 \le n \le 10 \\ 0, & \text{otherwise} \end{cases}$$

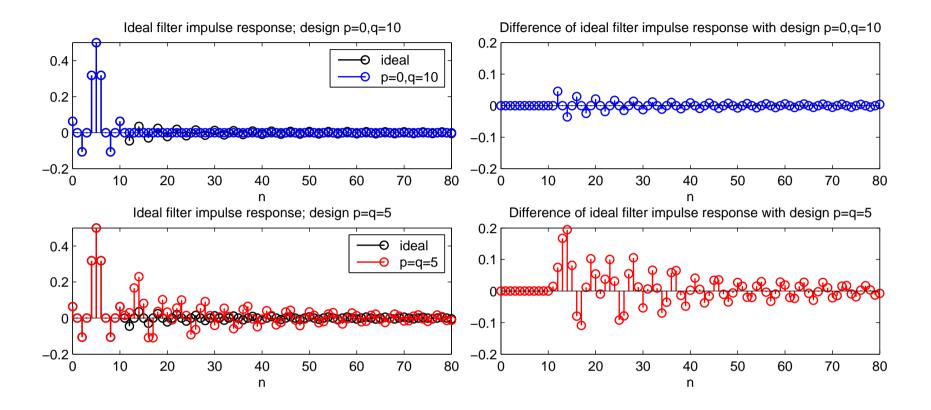
■ ARMA filter (p = q = 5): solving the Pade equations for a gives

$$\mathbf{a} = \begin{bmatrix} 1.000 & -2.526 & 3.677 & -3.485 & 2.131 & -0.703 \end{bmatrix}$$

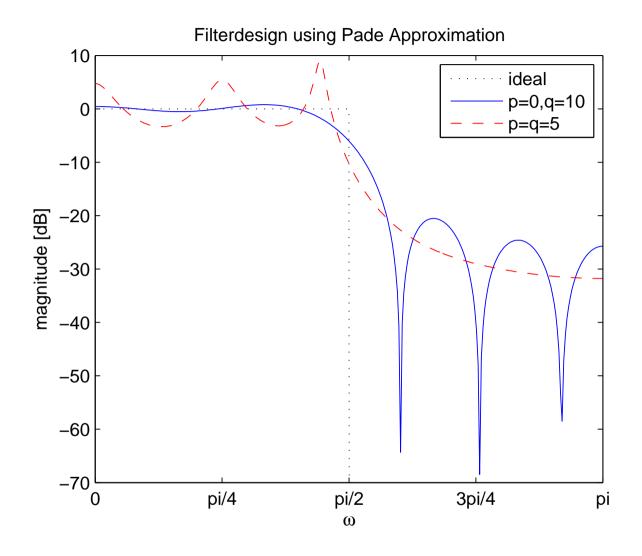
and subsequently the numerator coefficients are found as

$$\mathbf{b} = \begin{bmatrix} 0.064 & -0.161 & 0.128 & 0.046 & 0.064 & 0.021 \end{bmatrix}$$

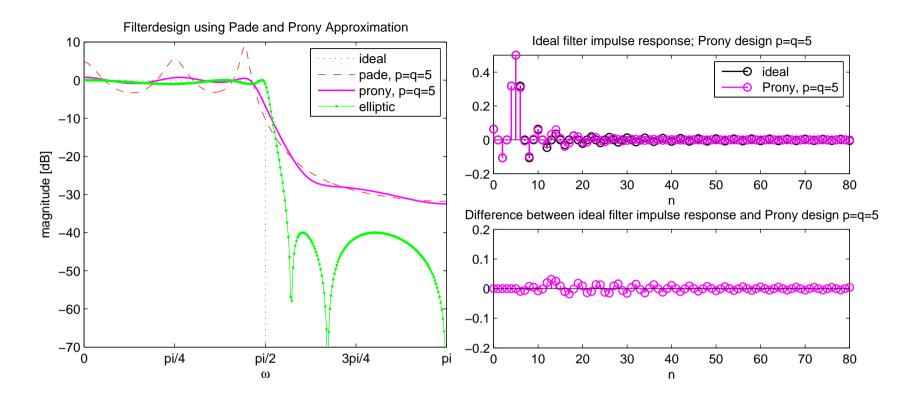
(Use Matlab [b,a] = prony(g,q,p) to find the design)



■ The ARMA(5,5) filter gives a larger error outside the specified interval. Also, this filter is not linear phase (no symmetry).



■ The ARMA(5,5) filter does not have a good frequency response in the passband.



- The ARMA(5,5) design using Prony's method is much better than the Pade approximation, since A(z) is designed to minimize the error over the entire domain
- A specialistic design (elliptic filter of 5th order) can still be better, with full control over the passband error and stopband attenuation.

Alternative writing of the equations

Recall: The solution is $\bar{\mathbf{a}} = -\mathbf{R}_{x}^{-1}\mathbf{r}_{x}$ with

$$\mathbf{R}_{\scriptscriptstyle X} = \mathbf{X}_q^H \mathbf{X}_q$$
 , $\mathbf{r}_{\scriptscriptstyle X} = \mathbf{X}_q^H \mathbf{x}_{q+1}$

Thus, the entries of \mathbf{R}_{x} and \mathbf{r}_{x} are

$$r_{X}(k,\ell) := [\mathbf{R}_{X}]_{k,\ell} = \sum_{n=q+1}^{\infty} x^{*}[n-k]x[n-\ell], \qquad r_{X}(k,0) := [\mathbf{r}_{X}]_{k} = \sum_{n=q+1}^{\infty} x^{*}[n-k]x[n]$$

The equation $\mathbf{R}_{\mathbf{x}}\mathbf{\bar{a}}=-\mathbf{r}_{\mathbf{x}}$ can be written as

$$\begin{bmatrix} r_{X}(1,1) & r_{X}(1,2) & \cdots & r_{X}(1,p) \\ r_{X}(2,1) & r_{X}(2,2) & \cdots & r_{X}(2,p) \\ \vdots & \ddots & \vdots & \vdots \\ r_{X}(p,1) & r_{X}(p,2) & \cdots & r_{X}(p,p) \end{bmatrix} \begin{bmatrix} a[1] \\ a[2] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} r_{X}(1,0) \\ r_{X}(2,0) \\ \vdots \\ r_{X}(p,0) \end{bmatrix}$$

The entries of \mathbf{R}_{\times} are considered *autocorrelations*. In a stochastic context (discussed later), we arrive at very similar equations.

■ These equations can also be summarized as

$$\sum_{\ell=1}^{p} a[\ell] r_{X}(k,\ell) = -r_{X}(k,0), \quad k = 1, \dots, p$$

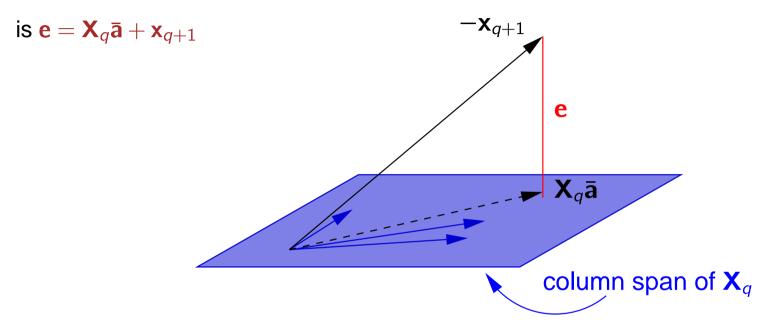
These are called the *Prony normal equations*.

Yet another writing form: $r_x(k,0) + \sum_{\ell=1}^p a[\ell] r_x(k,\ell) = 0$, $k = 1, \dots, p$

$$\begin{bmatrix} r_{X}(1,0) & r_{X}(1,1) & r_{X}(1,2) & \cdots & r_{X}(1,p) \\ r_{X}(2,0) & r_{X}(2,1) & r_{X}(2,2) & \cdots & r_{X}(2,p) \\ \vdots & \vdots & \ddots & \vdots \\ r_{X}(p,0) & r_{X}(p,1) & r_{X}(p,2) & \cdots & r_{X}(p,p) \end{bmatrix} \begin{bmatrix} 1 \\ a[1] \\ a[2] \\ \vdots \\ a[p] \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ a[p] \end{bmatrix}$$

Orthogonality principle

Consider the Least Squares problem $X_q\bar{a}\approx -x_{q+1}$. After solving, the error vector



The orthogonality principle states, in general, that the error vector for the optimal $\bar{\mathbf{a}}$ is orthogonal to all the columns of \mathbf{X}_q , i.e., $\mathbf{e} \perp \operatorname{colspan}(\mathbf{X}_q)$. If it was not the case, we could take a linear combination of these columns to reduce the error!

$$\mathbf{X}_{q}^{H}\mathbf{e} = 0 \quad \Leftrightarrow \quad \mathbf{X}_{q}^{H}(\mathbf{X}_{q}\mathbf{\bar{a}} + \mathbf{x}_{q+1}) = 0 \quad \Leftrightarrow \quad \mathbf{R}_{x}\mathbf{\bar{a}} = -\mathbf{r}_{x}$$

The minimum error

At the minimum, the total error is (due to the orthogonality principle)

$$\epsilon_{p,q} = \|\mathbf{e}\|^2 = \mathbf{e}^H \mathbf{e} = (\mathbf{X}_q \bar{\mathbf{a}} + \mathbf{x}_{q+1})^H \mathbf{e} = \mathbf{x}_{q+1}^H \mathbf{e}$$

This can further be written as

$$\epsilon_{p,q} = \mathbf{x}_{q+1}^H(\mathbf{X}_q\mathbf{ar{a}} + \mathbf{x}_{q+1}) = (\mathbf{x}_{q+1}^H\mathbf{X}_q)\mathbf{ar{a}} + \mathbf{x}_{q+1}^H\mathbf{x}_{q+1}$$

■ In terms of the autocorrelation sequence,

$$r_X(0,0) = \sum_{n=q+1}^{\infty} x^*[n]x[n] = \mathbf{x}_{q+1}^H \mathbf{x}_{q+1}, \qquad r_X(0,k) = \sum_{n=q+1}^{\infty} x^*[n]x[n-k] = [\mathbf{x}_{q+1}^H \mathbf{X}_q]_k$$

this can be written as

$$\epsilon_{p,q} = r_{x}(0,0) + \sum_{k=1}^{p} a[k]r_{x}(0,k)$$

This can be combined with the equation $\mathbf{R}_{x}\mathbf{\bar{a}}=-\mathbf{r}_{x}$ as

$$\begin{bmatrix}
r_X(0,0) & r_X(0,1) & r_X(0,2) & \cdots & r_X(0,p) \\
r_X(1,0) & r_X(1,1) & r_X(1,2) & \cdots & r_X(1,p) \\
r_X(2,0) & r_X(2,1) & r_X(2,2) & \cdots & r_X(2,p) \\
\vdots & \vdots & \vdots & \vdots & \vdots \\
r_X(p,0) & r_X(p,1) & r_X(p,2) & \cdots & r_X(p,p)
\end{bmatrix}
\begin{bmatrix}
1 \\
a[1] \\
a[2] \\
\vdots \\
a[p]
\end{bmatrix} = \begin{bmatrix}
\epsilon_{p,q} \\
0 \\
0 \\
\vdots \\
a[p]
\end{bmatrix}$$

or also as

$$\mathbf{R}_{\mathsf{X}}\mathbf{a}_{\mathsf{p}}=\epsilon_{\mathsf{p},\mathsf{q}}\mathbf{u}_{1}$$

where $\mathbf{u}_1 = [1, 0 \cdots 0]^T$ is a unit vector. These are the *augmented normal equations*.

All-pole modeling

Special case: all-pole modeling

If q = 0, then we have an all-pole model:

$$H(z) = \frac{b[0]}{1 + \sum_{k=1}^{p} a[k]z^{-k}}$$

- In some cases, this is an accurate physical model (e.g., speech)
- Even if it is not a valid physical model, it is attractive because it leads to fast computational algorithms to find a[k] (the Levinson algorithm)

Recall that the solution is given by solving the overdetermined system (q = 0)

$$\begin{bmatrix} x[0] & 0 & 0 & \cdots & 0 \\ x[1] & x[0] & 0 & \cdots & 0 \\ x[2] & x[1] & x[0] & \ddots & 0 \\ x[3] & x[2] & x[1] & \ddots & x[0] \\ \vdots & \vdots & \vdots & \vdots & \vdots \end{bmatrix} \begin{bmatrix} a[1] \\ a[2] \\ a[3] \end{bmatrix} = - \begin{bmatrix} x[1] \\ x[2] \\ x[3] \end{bmatrix}$$

All-pole modeling

The normal equations become (premultiply with \mathbf{X}_0^H)

$$\begin{bmatrix} x^*[0] & x^*[1] & x^*[2] & x^*[3] \cdots \\ 0 & \ddots & \ddots & \ddots & \ddots \\ 0 & 0 & x^*[0] & x^*[1] \cdots \end{bmatrix} \begin{bmatrix} x[0] & 0 & 0 \\ x[1] & \ddots & 0 \\ x[2] & \ddots & x[0] \\ x[3] & \ddots & x[1] \\ \vdots & \vdots & \vdots \end{bmatrix} \begin{bmatrix} a[1] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} x^*[0] & x^*[1] & x^*[2] & x^*[3] \cdots \\ 0 & \ddots & \ddots & \ddots & \ddots \\ 0 & 0 & x^*[0] & x^*[1] \cdots \end{bmatrix} \begin{bmatrix} x[1] \\ x[2] \\ x[3] \\ \vdots \end{bmatrix}$$
From this structure it is seen that the normal equations become

From this structure it is seen that the normal equations become

$$\begin{bmatrix} r_{X}(0) & r_{X}^{*}(1) & \cdots & r_{X}^{*}(p-1) \\ r_{X}(1) & r_{X}(0) & \cdots & r_{X}^{*}(p-2) \\ \vdots & \ddots & \vdots & \vdots \\ r_{X}(p-1) & r_{X}(p-2) & \cdots & r_{X}(0) \end{bmatrix} \begin{bmatrix} a[1] \\ a[2] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} r_{X}(1) \\ r_{X}(2) \\ \vdots \\ r_{X}(p) \end{bmatrix}$$

where the autocorrelation sequence is in fact "stationary":

$$r_X(k-\ell) = r_X(k,\ell) = \sum_{n=0}^{\infty} x[n-k]^* x[n-\ell]$$

All-pole modeling

For an all-pole model, the matrix \mathbf{R}_{\times} has a Toeplitz structure (constant along diagonals); it can be efficiently inverted using the Levinson algorithm (later in Ch.5).

■ The numerator will be chosen as $b[0] = \sqrt{\epsilon_p}$. It can be shown that then

$$r_X(k) = r_h(k), \qquad |k| \leq p$$

Meaning: the autocorrelation sequence of the filter matches the first p lags of the autocorrelation sequence of the specified data ("moment matching").

All-pole modeling with finite data

Suppose we have only N+1 samples, $x[0], \dots, x[N]$. What changes?

Autocorrelation method

Here, the entire data set is used, and extended with zeros:

$$\begin{bmatrix} x[0] & 0 & 0 \\ x[1] & \ddots & 0 \\ \vdots & \ddots & x[0] \\ x[N] & \ddots & x[1] \\ 0 & 0 & \vdots \\ 0 & 0 & x[N] \end{bmatrix} = \begin{bmatrix} x[1] \\ \vdots \\ x[N] \\ 0 \\ \vdots \\ 0 \end{bmatrix}$$

We will have the same normal equations as before, but with a different $r_x(k)$:

$$r_X(k) = \sum_{n=k}^{N} x[n]x^*[n-k], \ k = 0, \dots, p.$$

$$\begin{bmatrix} r_{X}(0) & \cdots & r_{X}^{*}(p-1) \\ \vdots & & \vdots \\ r_{X}(p-1) & \cdots & r_{X}(0) \end{bmatrix} \begin{bmatrix} a[1] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} r_{X}(1) \\ \vdots \\ r_{X}(p) \end{bmatrix}$$

All-pole modeling with finite data

Covariance method

If we know that x[n] is not zero for n < 0 and n > N, it is more accurate to omit those equations that contain an extension with zeros:

$$\begin{bmatrix} x[p-1] & \ddots & x[0] \\ x[p] & \ddots & x[1] \\ \vdots & & \ddots & \vdots \\ x[N-1] & \ddots & x[N-p] \end{bmatrix} \begin{bmatrix} a[1] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} x[p] \\ x[p+1] \\ \vdots \\ x[N] \end{bmatrix}$$

We will have slightly different normal equations as before, with $r_x(k, \ell)$ defined as

$$r_X(k,\ell) = \sum_{n=p}^{N} x[n-\ell]x^*[n-k], \ k,\ell = 0,\cdots, p.$$

$$\begin{bmatrix} r_{\mathsf{X}}(1,1) & \cdots & r_{\mathsf{X}}(1,p) \\ \vdots & & \vdots & \\ r_{\mathsf{X}}(p,1) & \cdots & r_{\mathsf{X}}(p,p) \end{bmatrix} \begin{bmatrix} a[1] \\ \vdots \\ a[p] \end{bmatrix} = - \begin{bmatrix} r_{\mathsf{X}}(1,0) \\ \vdots \\ r_{\mathsf{X}}(p,0) \end{bmatrix}$$

The Toeplitz property is lost.

Example: pole estimation

Consider the finite data sequence $\mathbf{x} = [1, \boldsymbol{\beta}, \boldsymbol{\beta}^2, \cdots, \boldsymbol{\beta}^N]^T$. Estimate an AR(1) model for this signal (p = 1):

$$H(z) = \frac{b(0)}{1 + a[1]z^{-1}}$$

Autocorrelation method

The normal equations collapse to $r_{\chi}(0)a[1] = -r_{\chi}(1)$, where

$$r_{X}(0) = [1, \beta, \beta^{2}, \cdots, \beta^{N}] \begin{bmatrix} 1 \\ \beta \\ \beta^{2} \\ \vdots \\ \beta^{N} \end{bmatrix} = \sum_{n=0}^{N} |\beta|^{2n} = \frac{1 - |\beta|^{2N+2}}{1 - |\beta|^{2}}$$

$$r_{X}(1) = [0, 1, \beta, \cdots, \beta^{N}] \begin{bmatrix} 1 \\ \beta \\ \vdots \\ \beta^{N} \\ 0 \end{bmatrix} = \sum_{n=0}^{N-1} \beta |\beta|^{2n} = \beta \frac{1 - |\beta|^{2N}}{1 - |\beta|^{2}}$$

$$\Rightarrow a[1] = -\beta \frac{1 - |\beta|^{2N}}{1 - |\beta|^{2N+2}}$$

Example: pole estimation

Covariance method

The normal equations collapse to $r_{\chi}(1,1)a[1] = -r_{\chi}(1,0)$, where

$$r_{x}(1,1) = [1, \beta, \beta^{2}, \cdots, \beta^{N-1}] \begin{bmatrix} 1 \\ \beta \\ \beta^{2} \\ \vdots \\ \beta^{N-1} \end{bmatrix} = \sum_{n=0}^{N-1} |\beta|^{2n} = \frac{1 - |\beta|^{2N}}{1 - |\beta|^{2}}$$

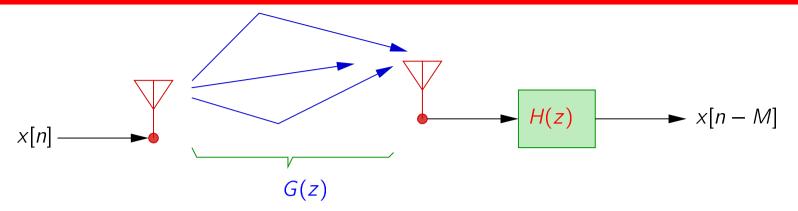
$$r_{\mathsf{X}}(1,0) = [1,eta,\cdots,eta^{N-1}] \left[egin{array}{c} eta \ eta^2 \ dots \ eta^N \end{array}
ight] = \sum_{n=0}^{N-1} eta |eta|^{2n} = eta rac{1-|eta|^{2N}}{1-|eta|^2}$$

Thus, $a[1] = -\beta$. Set b(0) = 1, then

$$H(z) = \frac{1}{1 - \beta z^{-1}}$$

The pole location is found exactly.

Example: channel inversion



Consider a communication channel with a known (estimated) transfer function G(z).

 \blacksquare At the receiver, we wish to equalize (invert) the channel using an equalizer H(z):

$$G(z)H(z) = 1 \Leftrightarrow g[n] * h[n] = \delta[n]$$

Further H(z) must be causal and stable, typically FIR.

If G(z) is not minimum-phase (has zeros outside the unit circle), causal inversion is not possible, and we allow for a delay:

$$G(z)H(z) = z^{-M}$$
 \Leftrightarrow $g[n] * h[n] = \delta[n - M] =: d[n]$

Example: channel inversion

Design of an FIR equalizer H(z) of length N:

$$e[n] = d[n] - h[n] * g[n] = d[n] - \sum_{\ell=0}^{N-1} h[\ell]g[n-\ell]$$

■ Minimize $\mathcal{E} = \sum_{n=0}^{\infty} |e[n]|^2 = \sum_{n=0}^{\infty} |d[n] - \sum_{\ell=0}^{N-1} h[\ell]g[n-\ell]|^2$:

$$\min_{\substack{\{h[\ell]\}}} \left\| \begin{bmatrix} g[0] & 0 & 0 \\ g[1] & \ddots & 0 \\ \vdots & \ddots & g[0] \\ g[N] & \ddots & g[1] \\ \vdots & \ddots & \vdots \end{bmatrix} - \begin{bmatrix} d[0] \\ \vdots \\ d[N-1] \\ d[N] \\ \vdots \end{bmatrix} \right\|^{2}$$

The LS solution of this overdetermined system $\mathbf{G}\mathbf{h} = \mathbf{d}$ is $\mathbf{h} = \mathbf{G}^{\dagger}\mathbf{d} = (\mathbf{G}^{H}\mathbf{G})^{-1}\mathbf{G}^{H}\mathbf{d}$.

■ Also, the corresponding normal equations are $(\mathbf{G}^H\mathbf{G})\mathbf{h} = \mathbf{G}^H\mathbf{d}$, or

$$\begin{bmatrix} r_g(0) & \cdots & r_g^*(N-1) \\ \vdots & & \vdots & \\ r_g(N-1) & \cdots & r_g(0) \end{bmatrix} \begin{bmatrix} h[0] \\ \vdots \\ h[N-1] \end{bmatrix} = - \begin{bmatrix} r_{dg}(0) \\ \vdots \\ r_{dg}(N-1) \end{bmatrix}$$

where $r_g(k) = \sum_{n=0}^{\infty} g[n]g^*[n-k]$ and $r_{dg}(k) = \sum_{n=0}^{\infty} d[n]g^*[n-k]$. Because the matrix $\mathbf{R}_g = \mathbf{G}^H \mathbf{G}$ is Toeplitz, it can be inverted efficiently.