

# BACS HW (Week 10)

108070023

**Question 1)** Download `demo_simple_regression_rsqr.R` from Canvas – it has a function that runs a regression simulation. This week, the simulation also reports  $R^2$  along with the other metrics from last week.

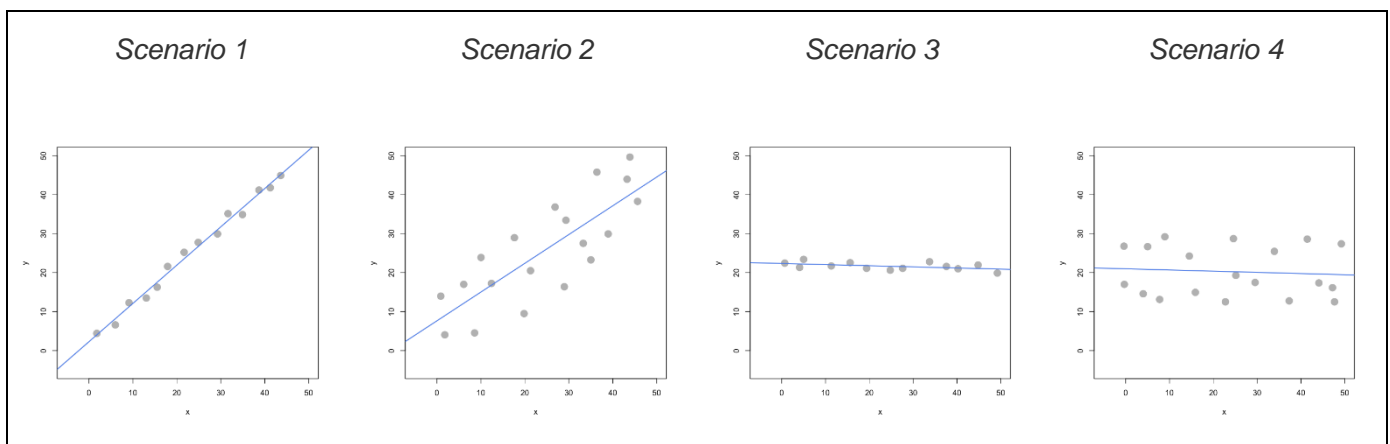
To answer the questions below, understand each of these four scenarios by simulating them:

Scenario 1: Consider a very narrowly dispersed set of points that have a negative or positive steep slope

Scenario 2: Consider a widely dispersed set of points that have a negative or positive steep slope

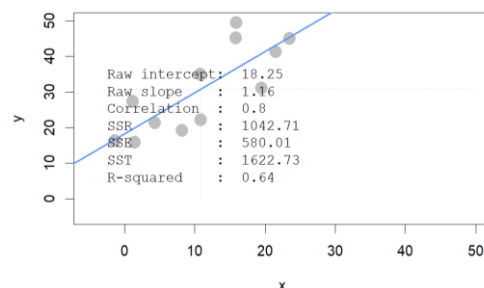
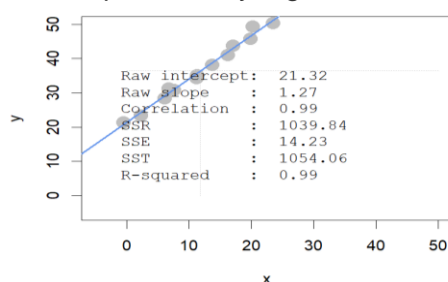
Scenario 3: Consider a very narrowly dispersed set of points that have a negative or positive shallow slope

Scenario 4: Consider a widely dispersed set of points that have a negative or positive shallow slope



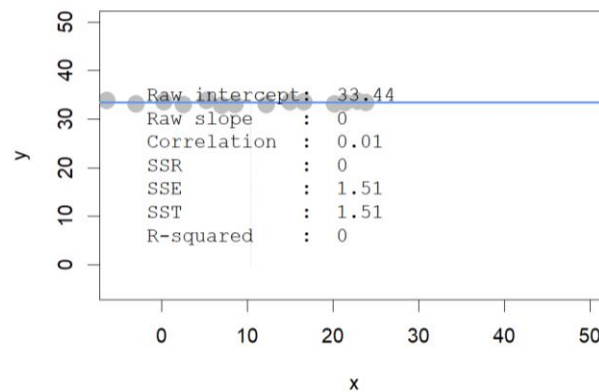
a. Comparing scenarios 1 and 2, which do we expect to have a stronger  $R^2$ ?

Ans: We'll expect scenario 1 to have a stronger  $R$ . Since  $R$  is the portion that SSR take part in SST (how much SST is explained by regression), so if we explain it intuitively, the narrower the set of points lie, the closer the set of points to the line; therefore, SST is more explainable by regression line.

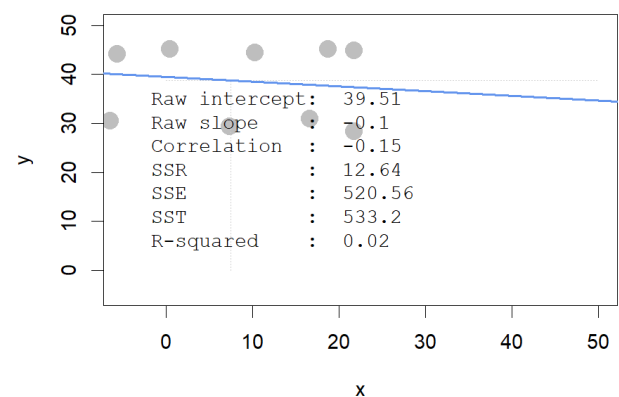
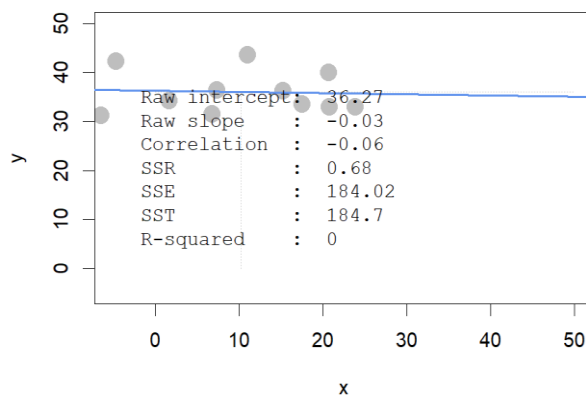


b. Comparing scenarios 3 and 4, which do we expect to have a stronger  $R^2$  ?

Ans: I'll say that they both have similar  $R^2$  close to 0, but if points of scenario 4 are widely dispersed enough,  $R^2$  of scenario 4 will be stronger. Let's look at figures below carefully, SSR of scenario 2 is a little bit bigger than scenario 1, while the amount of it is minor compared to SST. Therefore, we'll observe the two scenarios to have similar  $R^2$  that is close to 0.



< scenario 1 >



< two kinds of scenario 2 >

c. Comparing scenarios 1 and 2, which do we expect has bigger/smaller SSE, SSR, and SST? (intuitively)

Ans: We'll expect scenario 2 has the bigger SST and SSE, since points in scenario 2 are more dispersed than scenario 1, and they may result in a bigger SST. Also, points in scenario 2 lie widely compared to scenario 1 (lie farther to the line), which comes up with a bigger SSE (minimum sum of square of distance between  $\hat{y}$  and  $y$ ). However, we'll expect SSR of scenario 2 to be stronger. As we discussed above, SSR means the how much SST is explained by the regression line, if the points lie narrowly, the trend of them will look more similar to the regression line, and result in stronger SSR.

d. Comparing scenarios 3 and 4, which do we expect has bigger/smaller SSE, SSR, and SST? (intuitively)

Ans: We'll expect scenario 4 to have higher SSE and SST, since its points disperse more widely. As for SSR, if the points are widely dispersed enough, it may result in the bigger SSR compared to scenario 3.

**Question 2)** Let's perform regression ourselves on the `programmer_salaries.txt` dataset we saw in class

You can read the file using `read.csv("programmer_salaries.txt", sep="\t")`

a. First, use the `lm()` function to estimate the model `Salary ~ Experience + Score + Degree`

(show the beta coefficients,  $R^2$  and the first 5 values of `y` (`$fitted.values`) and `$residuals`)

```
> #2-a
> sal<-read.csv("programmer_salaries.txt", sep="\t")
> reg<-lm(Salary~Experience+Score+Degree,data = sal)
> head(sal,5)
  Experience Score Degree Salary
1          4    78     0   24.0
2          7   100     1   43.0
3          1    86     0   23.7
4          5    82     1   34.3
5          8    86     1   35.8
6         10    84     1   38.0
> summary(reg)
```

Call:

```
lm(formula = Salary ~ Experience + Score + Degree, data = sal)
```

Residuals:

Min	1Q	Median	3Q	Max
-3.8963	-1.7290	-0.3375	1.9699	5.0480

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	7.9448	7.3808	1.076	0.2977
Experience	1.1476	0.2976	3.856	0.0014 **

```

Score          0.1969      0.0899      2.191      0.0436 *
Degree         2.2804      1.9866      1.148      0.2679
---
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 2.396 on 16 degrees of freedom
Multiple R-squared:  0.8468,    Adjusted R-squared:  0.8181
F-statistic: 29.48 on 3 and 16 DF,  p-value: 9.417e-07

> head(reg$residual,n=5)
      1      2      3      4      5
-3.8962605  5.0479568 -2.3290112  2.1879860 -0.5425072
> head(reg$fitted.values,n=5)
      1      2      3      4      5
27.89626 37.95204 26.02901 32.11201 36.34251

```

b. Use only linear algebra (and the geometric view of regression) to estimate the regression yourself:

i. Create an  $X$  matrix that has a first column of 1s followed by columns of the independent variables

```

> #2-b-i
> salx<-data.matrix(sal[,1:3])
> salx<-cbind(c(1),salx)
> colnames(salx)[1]<-"intercept"

```

ii. Create a  $y$  vector with the Salary values (*only show the code*)

```

> #2-b-ii
> saly<-sal$Salary

```

iii. Compute the  $\beta_{\text{hat}}$  vector of estimated regression coefficients (*show the code and values*)

```

> #2-b-iii
> beta_h<-(solve(t(salx)%*%salx))%*%t(salx)%*%saly
> beta_h
      [,1]

```

```
intercept  7.944849
Experience  1.147582
Score      0.196937
Degree     2.280424
```

iv. Compute a `y_hat` vector of estimated y values, and a `res` vector of residuals (show the code and the first 5 values of `y_hat` and `res`)

```
> #2-b-iv
> saly_h<-salyx%%beta_h
> head(saly_h,n=5)
      [,1]
[1,] 27.89626
[2,] 37.95204
[3,] 26.02901
[4,] 32.11201
[5,] 36.34251
> res<-saly-saly_h
> head(res,n=5)
      [,1]
[1,] -3.8962605
[2,]  5.0479568
[3,] -2.3290112
[4,]  2.1879860
[5,] -0.5425072
```

v. Using only the results from (i) – (iv), compute SSR, SSE and SST (show the code and values)

```
> #2-b-v
> SST<-sum((saly-mean(saly))^2)
> SSR<-sum((saly_h-mean(saly))^2)
> SSE<-sum((saly-saly_h)^2)
> data.frame(SST,SSR,SSE)
      SST      SSR      SSE
1 599.7855 507.896 91.88949
```

c. Compute  $R^2$  for in two ways, and confirm you get the same results (*show code and values*):

i. Use any combination of SSR, SSE, and SST

```
> #2-c-i
> R_square_1<-SSR/SST
> R_square_1
[1] 0.8467961
```

ii. Use the squared correlation of vectors y and y

```
> #2-c-ii
> R_square_2<-cor(saly,saly_h)^2
> R_square_2
      [,1]
[1,] 0.8467961
```

(see question 3 on next page)

**Question 3)** We're going to take a look back at the early heady days of global car manufacturing, when American, Japanese, and European cars competed to rule the world. Take a look at the data set in file `auto-data.txt`. We are interested in explaining **what kind of cars have higher fuel efficiency (mpg).**

a. Let's first try exploring this data and problem:

i. Visualize the data in any way you feel relevant (report only relevant/interesting ones)

```
> #3-a-i
> library(dplyr)
> library(ggplot2)
> #install.packages("cowplot")
> library(cowplot)
> auto <- read.table("auto-data.txt", header=FALSE, na.strings = "?")
> names(auto) <- c("mpg", "cylinders", "displacement", "horsepower",
"weight",
+               "acceleration", "model_year", "origin", "car_name")
> auto$brand <-
as.character(lapply(auto$car_name,function(x){unlist(strsplit(x," "))[1]}))
> my_summary <- auto %>%
+   count(brand, sort = TRUE)
```

```
> my_summary #too many categories of car_name and brand, not valuable for  
linear regression model
```

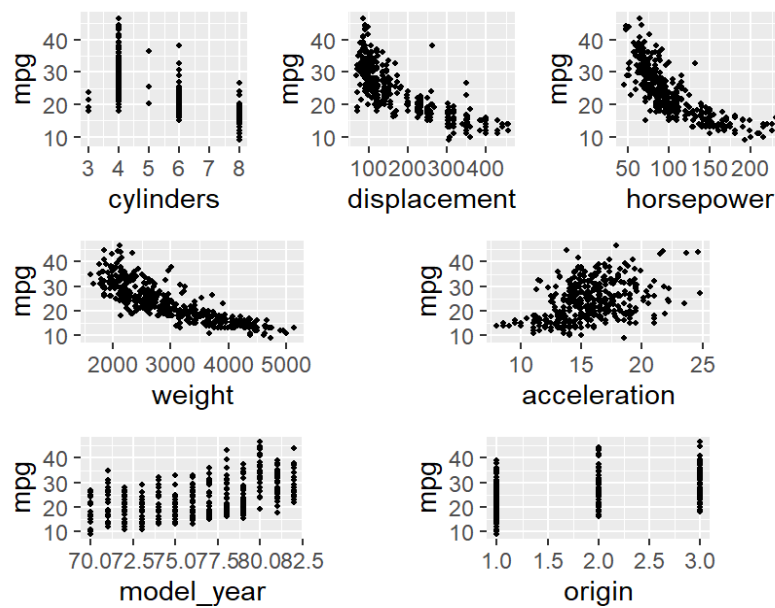
	brand	n
1	ford	51
2	chevrolet	43
3	plymouth	31
4	amc	28
5	dodge	28
6	toyota	25
7	datsum	23
8	buick	17
9	pontiac	16
10	volkswagen	15
11	honda	13
12	mercury	11
13	mazda	10
14	oldsmobile	10
15	fiat	8
16	peugeot	8
17	audi	7
18	chrysler	6
19	volvo	6
20	vw	6
21	renault	5
22	opel	4
23	saab	4
24	subaru	4
25	chevy	3
26	bmw	2
27	cadillac	2
28	maxda	2
29	mercedes-benz	2
30	capri	1
31	chevroelt	1
32	hi	1
33	mercedes	1
34	nissan	1
35	toyouta	1

```

36     triumph 1
37     vokswagen 1
> auto['car_name']<-NULL #abandon car_name column
> auto['brand']<-NULL
> # Scatter plot
> cylinders <- ggplot(auto, aes(x = cylinders, y = mpg))+
+   geom_point(size=0.8)
> displacement <- ggplot(auto, aes(x = displacement, y = mpg))+
+   geom_point(size=0.8)
> horsepower <- ggplot(auto, aes(x = horsepower, y = mpg))+
+   geom_point(size=0.8)
> weight <- ggplot(auto, aes(x = weight, y = mpg))+
+   geom_point(size=0.8)
> acceleration <- ggplot(auto, aes(x = acceleration, y = mpg))+
+   geom_point(size=0.8)
> model_year <- ggplot(auto, aes(x = model_year, y = mpg))+
+   geom_point(size=0.8)
> origin <- ggplot(auto, aes(x = origin, y = mpg))+
+   geom_point(size=0.8)
> ggdraw() +
+   draw_plot(cylinders, 0, .6, .33, .35) +
+   draw_plot(displacement, .33, .6, .33, .35) +
+   draw_plot(horsepower, .66, .6, .33, .35) +
+   draw_plot(weight, 0, .3, .4, .3) +
+   draw_plot(acceleration, .5, .3, .4, .3) +
+   draw_plot(model_year, 0, 0, .4, .3) +
+   draw_plot(origin, .5, 0, .4, .3)

```





<relationship between mpg and each variable>

- ii. Report a correlation table of all variables, rounding to two decimal places (in the `cor()` function, set `use="pairwise.complete.obs"` to handle missing values)

```
> #3-a-ii
> round(cor(auto[,c("mpg", "cylinders", "displacement", "horsepower",
"weight", "acceleration", "model_year", "origin")]), use =
"pairwise.complete.obs"), digits=2)
```

	mpg	cylinders	displacement	horsepower
mpg	1.00	-0.78	-0.80	-0.78
cylinders	-0.78	1.00	0.95	0.84
displacement	-0.80	0.95	1.00	0.90
horsepower	-0.78	0.84	0.90	1.00
weight	-0.83	0.90	0.93	0.86
acceleration	0.42	-0.51	-0.54	-0.69
model_year	0.58	-0.35	-0.37	-0.42
origin	0.56	-0.56	-0.61	-0.46

	weight	acceleration	model_year	origin
mpg	-0.83	0.42	0.58	0.56
cylinders	0.90	-0.51	-0.35	-0.56
displacement	0.93	-0.54	-0.37	-0.61
horsepower	0.86	-0.69	-0.42	-0.46
weight	1.00	-0.42	-0.31	-0.58
acceleration	-0.42	1.00	0.29	0.21

model_year	-0.31	0.29	1.00	0.18
origin	-0.58	0.21	0.18	1.00

iii. From the visualizations and correlations, which variables seem to relate to mpg?

A: 'cylinders', 'displacement', 'horsepower', 'weight' have correlation higher than 0.75 or smaller than -0.75, so I'll say they seem to relate to mpg.

iv. Which relationships might not be linear? (*don't worry about linearity for rest of this HW*)

Ans: 'acceleration', 'model year', 'origin' may not have linear relationship with mpg since they have lower correlation, and the distribution of origin and model year seem like several vertical lines, I can't observe linear relationship in these graphs.

v. Are there any pairs of independent variables that are highly correlated ( $r > 0.7$ )?

```
> #3-a-v
> library(reshape2)
> diag(cor_table) <- 0
> cor_melt <- melt(cor_table)
> hight_cor <- cor_melt[order(abs(cor_melt$value), decreasing = T) &
abs(cor_melt$value) > 0.7, ]
>
> ##### eliminate the same combination of variable
> #sort two variable by first character order
> hight_cor[1:2] <- t( apply(hight_cor[1:2], 1, sort) )
> #eliminate the same variable combination
> hight_cor[!duplicated(hight_cor[1:2]), ]
      Var1      Var2 value
2  cylinders      mpg -0.78
3 displacement      mpg -0.80
4  horsepower      mpg -0.78
5         mpg      weight -0.83
11 cylinders displacement  0.95
12 cylinders  horsepower  0.84
13 cylinders      weight  0.90
20 displacement horsepower  0.90
21 displacement      weight  0.93
29  horsepower      weight  0.86
```

Ans: The above result shows pairs of independent variables that are highly correlated ( $r > 0.7$ )

- b. Let's create a linear regression model where mpg is dependent upon all other suitable variables (*Note: origin is categorical with three levels, so use `factor(origin)` in `lm(...)` to split it into two dummy variables*)
- i. Which independent variables have a 'significant' relationship with mpg at 1% significance?

```
> #3-b
> mpg_reg<-
lm(mpg~cylinders+displacement+horsepower+weight+acceleration+model_year+fac
tor(origin),data = auto)
> summary(mpg_reg)
```

Call:

```
lm(formula = mpg ~ cylinders + displacement + horsepower + weight +
    acceleration + model_year + factor(origin), data = auto)
```

Residuals:

Min	1Q	Median	3Q	Max
-9.0095	-2.0785	-0.0982	1.9856	13.3608

Coefficients:

	Estimate	Std. Error	t value
(Intercept)	-1.795e+01	4.677e+00	-3.839
cylinders	-4.897e-01	3.212e-01	-1.524
displacement	2.398e-02	7.653e-03	3.133
horsepower	-1.818e-02	1.371e-02	-1.326
weight	-6.710e-03	6.551e-04	-10.243
acceleration	7.910e-02	9.822e-02	0.805
model_year	7.770e-01	5.178e-02	15.005
factor(origin)2	2.630e+00	5.664e-01	4.643
factor(origin)3	2.853e+00	5.527e-01	5.162

Pr(>|t|)

(Intercept)	0.000145 ***
cylinders	0.128215
displacement	0.001863 **
horsepower	0.185488
weight	< 2e-16 ***
acceleration	0.421101
model_year	< 2e-16 ***

```

factor(origin)2 4.72e-06 ***
factor(origin)3 3.93e-07 ***
---
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 3.307 on 383 degrees of freedom
Multiple R-squared: 0.8242, Adjusted R-squared: 0.8205
F-statistic: 224.5 on 8 and 383 DF, p-value: < 2.2e-16

```

Ans: 'Intercept', 'displacement', 'weight', 'model\_year', 'origin' have significant relationship with mpg at 1% significance?

- ii. Looking at the coefficients, is it possible to determine which independent variables are the *most effective* at increasing mpg? If so, which ones, and if not, why not? (hint: units!)

Ans: No, because those variable are in different scales, it's not possible to compare their magnitude directly. We should standardize them and compare between them.

- c. Let's try to resolve some of the issues with our regression model above.

- i. Create fully standardized regression results: are these slopes easier to compare?  
(note: consider if you should standardize origin)

```

> #3-c-i
> auto_std <- cbind(scale(auto[1:7]),auto$origin)#origin is categorical
variable
> colnames(auto_std) <- colnames(auto[1:8])
> auto_std <- data.frame(auto_std)
> mpg_stdreg<-
lm(mpg~cylinders+displacement+horsepower+weight+acceleration+model_year+fac
tor(origin),data = auto_std)
> summary(mpg_stdreg)

```

Call:

```
lm(formula = mpg ~ cylinders + displacement + horsepower + weight +
    acceleration + model_year + factor(origin), data = auto_std)
```

Residuals:

Min	1Q	Median	3Q	Max
-1.15270	-0.26593	-0.01257	0.25404	1.70942

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	-0.13323	0.03174	-4.198	3.35e-05
cylinders	-0.10658	0.06991	-1.524	0.12821
displacement	0.31989	0.10210	3.133	0.00186
horsepower	-0.08955	0.06751	-1.326	0.18549
weight	-0.72705	0.07098	-10.243	< 2e-16
acceleration	0.02791	0.03465	0.805	0.42110
model_year	0.36760	0.02450	15.005	< 2e-16
factor(origin)2	0.33649	0.07247	4.643	4.72e-06
factor(origin)3	0.36505	0.07072	5.162	3.93e-07

(Intercept)	***
cylinders	
displacement	**
horsepower	
weight	***
acceleration	
model_year	***
factor(origin)2	***
factor(origin)3	***

---

Signif. codes:

0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.423 on 383 degrees of freedom

Multiple R-squared: 0.8242, Adjusted R-squared: 0.8205

F-statistic: 224.5 on 8 and 383 DF, p-value: < 2.2e-16

Ans: Yes, it is easier to interpret the coefficient after standardization. According to the report above, we can find out that weight is the most effective at increasing mpg since it has the highest correlation with mpg.

- ii. Regress mpg over each *nonsignificant* independent variable, individually.  
Which ones become significant when we regress mpg over them individually?

```
> #3-c-ii
> #unsignificant var:cylinders,horsepower,acceleration
> cy<-lm(mpg~cylinders,data=auto)
> summary(cy)
```

Call:

```
lm(formula = mpg ~ cylinders, data = auto)
```

Residuals:

Min	1Q	Median	3Q	Max
-14.2607	-3.3841	-0.6478	2.5538	17.9022

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	42.9493	0.8330	51.56	<2e-16 ***
cylinders	-3.5629	0.1458	-24.43	<2e-16 ***

---

Signif. codes:

0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 4.942 on 396 degrees of freedom

Multiple R-squared: 0.6012, Adjusted R-squared: 0.6002

F-statistic: 597.1 on 1 and 396 DF, p-value: < 2.2e-16

```
> hp<-lm(mpg~horsepower,data=auto)
```

```
> summary(hp)
```

Call:

```
lm(formula = mpg ~ horsepower, data = auto)
```

Residuals:

Min	1Q	Median	3Q	Max
-13.5710	-3.2592	-0.3435	2.7630	16.9240

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	39.935861	0.717499	55.66	<2e-16 ***
horsepower	-0.157845	0.006446	-24.49	<2e-16 ***

---

Signif. codes:

0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```

Residual standard error: 4.906 on 390 degrees of freedom
(因為不存在，6 個觀察量被刪除了)
Multiple R-squared: 0.6059, Adjusted R-squared: 0.6049
F-statistic: 599.7 on 1 and 390 DF, p-value: < 2.2e-16

> ac<-lm(mpg~acceleration,data=auto)
> summary(ac)

Call:
lm(formula = mpg ~ acceleration, data = auto)

Residuals:
    Min       1Q   Median       3Q      Max
-18.007  -5.636  -1.242   4.758  23.192

Coefficients:
              Estimate Std. Error t value Pr(>|t|)
(Intercept)    4.9698     2.0432   2.432  0.0154 *
acceleration    1.1912     0.1292   9.217 <2e-16 ***
---
Signif. codes:
0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

```

```

Residual standard error: 7.101 on 396 degrees of freedom
Multiple R-squared: 0.1766, Adjusted R-squared: 0.1746
F-statistic: 84.96 on 1 and 396 DF, p-value: < 2.2e-16

```

Ans: The above figures show the regression line for three nonsignificant independent variable: **cylinders**, **horsepower**, **acceleration**. According to the p-value of each regression, all of them are significant if we regress mpg over them individually.

iii. Plot the density of the *residuals*: are they normally distributed and centered around zero?

(get the residuals of a fitted linear model, e.g. `regr <- lm(...)`, using `regr$residuals`)

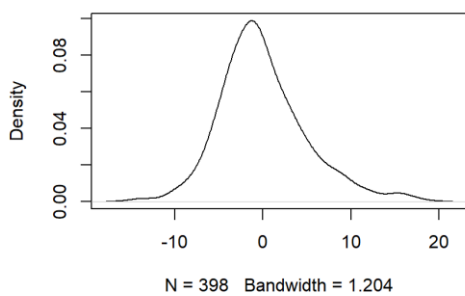
```

> #3-c-iii
> plot(density(cy$residuals),main = "cylinders$residuals")
> plot(density(hp$residuals),main = "horsepower$residuals")
> plot(density(ac$residuals),main = "acceleration$residuals")

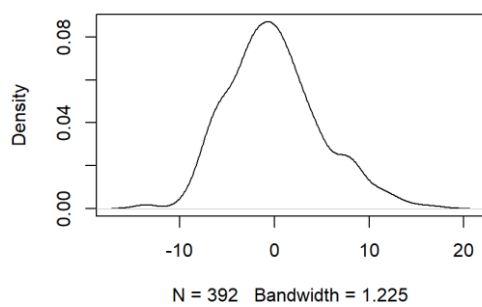
```

Ans: Yes, they all center around 0 and their distribution look similar to normal distribution.

**cylinders\$residuals**



**horsepower\$residuals**



**acceleration\$residuals**

