CMPT 210: Probability and Computing

Lecture 18

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If Range[X] =
$$\{x_1, x_2, ... x_n\}$$
, Range[Y] = $\{y_1, y_2, ... y_n\}$, then for $x \in \text{Range}(X)$, $[X = x] = [X = x \cap y = y_1] \cup [X = x \cap y = y_2] \cup ... \cup [X = x \cap y = y_n]$ $\implies \Pr[X = x] = \Pr[X = x \cap y = y_1] + \Pr[X = x \cap y = y_2] + ... + \Pr[X = x \cap y = y_n].$

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$$\implies \mathsf{PDF}_X[x] = \sum_i \mathsf{PDF}_{X,Y}[x,y_i].$$

Hence, we can obtain the distribution for each r.v. from the joint distribution by "marginalizing" over the other r.v's.

Q: Suppose that 3 batteries are randomly chosen from a group of 3 new, 4 used but still working, and 5 defective batteries. If the batteries are distinct and we let X and Y denote, respectively, the number of new and used but still working batteries that are chosen, completely specify $PDF_{X,Y}$.

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For
$$i \in [3], j \in [3]$$
, $\mathsf{PDF}_{X,Y}[i,j] = \mathsf{Pr}[X = i \cap Y = j | X + Y \le 3] = \frac{\binom{3}{i}\binom{4}{j}\binom{5}{3-i-j}}{\binom{12}{3}}$.

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 $PDF_{X,Y}[0,0] = \frac{\binom{5}{3}}{\binom{12}{3}} = 10/220$, $PDF_{X,Y}[1,2] = \frac{\binom{3}{1}\binom{4}{2}\binom{5}{2}}{\binom{12}{3}} = 18/220$.

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Table 4.1 $P\{X = i, Y = j\}$.						
i j	0	1	2	3	Row Sum $= P\{X = i\}$	
0	$\frac{10}{220}$	$\frac{40}{220}$	$\frac{30}{220}$	$\frac{4}{220}$	$\frac{84}{220}$	
1	$\frac{30}{220}$	$\frac{60}{220}$	$\frac{18}{220}$	0	$\frac{108}{220}$	
2	$\frac{15}{220}$	$\frac{12}{220}$	0	0	$\frac{27}{220}$	
3	$\frac{1}{220}$	0	0	0	$\frac{1}{220}$	
Column Sums = $P\{Y = j\}$	56 220	112 220	$\frac{48}{220}$	$\frac{4}{220}$		



Expectation - Examples

For a random variable $X: \mathcal{S} \to V$ and a function $g: V \to \mathbb{R}$, we define $\mathbb{E}[g(X)]$ as follows:

$$\mathbb{E}[g(X)] := \sum_{x \in \mathsf{Range}(X)} g(x) \Pr[X = x]$$

If
$$g(x) = x$$
 for all $x \in \text{Range}(X)$, then $\mathbb{E}[g(X)] = \mathbb{E}[X]$.

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Q: For a standard dice, if X is the r.v. corresponding to the number that comes up on the dice, compute $\mathbb{E}[X^2]$ and $(\mathbb{E}[X])^2$

For a standard dice, $X \sim \text{Uniform}(\{1, 2, 3, 4, 5, 6\})$ and hence,

$$\mathbb{E}[X^2] = \sum_{x \in \{1, 2, 3, 4, 5, 6\}} x^2 \Pr[X = x] = \frac{1}{6} \left[1^2 + 2^2 + \dots + 6^2 \right] = \frac{91}{6}$$

$$(\mathbb{E}[X])^2 = \left(\sum_{x \in \{1, 2, 3, 4, 5, 6\}} x \Pr[X = x]\right)^2 = \left(\frac{1}{6} \left[1 + 2 + \dots + 6\right]\right)^2 = \frac{49}{4}$$

Deviation from the Mean

We have developed tools to calculate the mean of random variables. Getting a handle on the expectation is useful because it tell us what would happen on average.

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Example: Consider three random variables W, Y and Z whose PDF's can be given as:

W = 0	(with $p=1$)
Y = -1	(with $p=1/2$)
= +1	(with $p=1/2$)
Z = -1000	(with $p=1/2$)
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Though $\mathbb{E}[W] = \mathbb{E}[Y] = \mathbb{E}[Z] = 0$, these distributions are quite different. Z can take values really far away from its expected value, while W can take only one value equal to the mean. Hence, we want to understand how much does a random variable "deviate" from its mean.

Standard way to measure the deviation from the mean is to calculate the variance. For r.v. X,

$$\operatorname{Var}[X] = \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_{x \in \operatorname{Range}(X)} (x - \mu)^2 \operatorname{Pr}[X = x] \qquad \text{(where } \mu := \mathbb{E}[X])$$

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Since X is a Bernoulli random variable, X=1 with probability p and X=0 with probability 1-p. Recall that $\mathbb{E}[X]=\mu=(0)(1-p)+(1)(p)=p$.

$$Var[X] = \sum_{x \in \{0,1\}} (x-p)^2 \Pr[X = x] = (0-p)^2 \Pr[X = 0] + (1-p)^2 \Pr[X = 1]$$
$$= p^2 (1-p) + (1-p)^2 p = p(1-p)[p+1-p] = p(1-p).$$

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For a Bernoulli r.v. X, $Var[X] = p(1-p) \le \frac{1}{4}$. Hence, the variance is maximum when p = 1/2 (equal probability of getting heads/tails).

Alternate definition of variance: $Var[X] = \mathbb{E}[X^2] - \mu^2 = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$.

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$$\begin{aligned} \textit{Proof} : \mathsf{Var}[X] &= \mathbb{E}[(X - \mathbb{E}[X])^2] = \sum_{x \in \mathsf{Range}(X)} (x - \mu)^2 \; \mathsf{Pr}[X = x] \\ &= \sum_{x \in \mathsf{Range}(X)} (x^2 - 2\mu x + \mu^2) \; \mathsf{Pr}[X = x] \\ &= \sum_{x \in \mathsf{Range}(X)} (x^2 \, \mathsf{Pr}[X = x]) - (2\mu x \, \mathsf{Pr}[X = x]) + (\mu^2) \, \mathsf{Pr}[X = x] \end{aligned}$$

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(Since μ is a constant does not depend on the x in the sum.)

$$= \mathbb{E}[X^2] - 2\mu \mathbb{E}[X] + \mu^2 \sum_{x \in \mathsf{Range}(X)} \mathsf{Pr}[X = x] \quad (\mathsf{Definition of } \mathbb{E}[X] \; \mathsf{and} \; \mathbb{E}[X^2])$$

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$$\mathbb{E}[X^2] = \sum_{x \in \{1, 2, 3, 4, 5, 6\}} x^2 \Pr[X = x] = \frac{1}{6} \left[1^2 + 2^2 + \dots + 6^2 \right] = \frac{91}{6}$$

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$$\implies \text{Var}[X] = \frac{91}{6} - \frac{49}{4} \approx 2.917$$

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$$\mathbb{E}[X] = \sum_{i=1}^{n} v_i \Pr[X = v_i] = \frac{1}{n} [v_1 + v_2 + \dots v_n] \quad ; \quad \mathbb{E}[X^2] = \frac{1}{n} [v_1^2 + v_2^2 + \dots v_n^2].$$

$$\implies \mathsf{Var}[X] = \frac{[v_1^2 + v_2^2 + \dots v_n^2]}{n} - \left(\frac{[v_1 + v_2 + \dots v_n]}{n}\right)^2$$

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Recall that $\mathbb{E}[W] = \mathbb{E}[Y] = \mathbb{E}[Z] = 0$.

 $Var[W] = \mathbb{E}[W^2] - (\mathbb{E}[W])^2 = \mathbb{E}[W^2] = \sum_{w \in \mathsf{Range}(W)} w^2 \Pr[W = w] = 0^2(1) = 0$. The variance of W is zero because it can only take one value and the r.v. does not "vary".

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$$Var[Z] = \mathbb{E}[Z^2] = \sum_{z \in Range(Z)} z^2 \Pr[Z = z] = (-1000)^2 (1/2) + (1000)^2 (1/2) = 10^6.$$

Hence, the variance can be used to distinguish between r.v.'s that have the same mean.

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Recall that for a coin s.t. Pr[heads] = p, R is the r.v. equal to the number of coin tosses we need to get the first heads. Let A be the event that we get a heads in the first toss. Using the law of total expectation,

$$\mathbb{E}[R^2] = \mathbb{E}[R^2|A] \Pr[A] + \mathbb{E}[R^2|A^c] \Pr[A^c]$$

 $\mathbb{E}[R^2|A] = 1$ ($R^2 = 1$ if we get a heads in the first coin toss) and $\Pr[A] = p$. Hence,

$$\mathbb{E}[R^2] = (1)(p) + \mathbb{E}[R^2|A^c](1-p) \quad ; \quad \mathbb{E}[R^2|A^c] = \sum_{k=1} k^2 \Pr[R = k|A^c]$$

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Note that
$$\Pr[R = k | A^c] = \Pr[R = k | \text{ first toss is a tails}] = (1 - p)^{k-2} p = \Pr[R = k - 1]$$

$$\implies \mathbb{E}[R^2 | A^c] = \sum_{k=1} k^2 \Pr[R = k - 1] = \sum_{t=0} (t+1)^2 \Pr[R = t] \qquad (t := k-1)$$

Continuing from the previous slide,

$$\mathbb{E}[R^2|A^c] = \sum_{t=0}^{\infty} (t+1)^2 \Pr[R=t] = \sum_{t=0}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=0}^{\infty} t \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + \sum_{t=1}^{\infty} \Pr[R=t] + 2 \mathbb{E}[R] + 1$$

Continuing from the previous slide,

$$\mathbb{E}[R^2|A^c] = \sum_{t=0}^{\infty} (t+1)^2 \Pr[R=t] = \sum_{t=0}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=0}^{\infty} t \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t]$$

$$= \sum_{t=1}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + \sum_{t=1}^{\infty} \Pr[R=t] = \mathbb{E}[R^2] + 2\mathbb{E}[R] + 1$$

Putting everything together,

$$\mathbb{E}[R^2] = (1)(p) + (\mathbb{E}[R^2] + 2\mathbb{E}[R] + 1])(1-p) \implies p \,\mathbb{E}[R^2] = p + 2(1-p)\mathbb{E}[R] + (1-p)\mathbb{E}[1]$$

$$\implies p \,\mathbb{E}[R^2] = p + \frac{2(1-p)}{p} + (1-p) \qquad (\mathbb{E}[R] = \frac{1}{p}, \,\mathbb{E}[1] = 1)$$

Continuing from the previous slide,

$$\mathbb{E}[R^2|A^c] = \sum_{t=0}^{\infty} (t+1)^2 \Pr[R=t] = \sum_{t=0}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=0}^{\infty} t \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + \sum_{t=1}^{\infty} \Pr[R=t] + 2 \mathbb{E}[R^2] + 2 \mathbb{E}[R] + 1$$

Putting everything together,

$$\mathbb{E}[R^{2}] = (1)(p) + (\mathbb{E}[R^{2}] + 2\mathbb{E}[R] + 1])(1 - p) \implies p \,\mathbb{E}[R^{2}] = p + 2(1 - p)\mathbb{E}[R] + (1 - p)\mathbb{E}[1]$$

$$\implies p \,\mathbb{E}[R^{2}] = p + \frac{2(1 - p)}{p} + (1 - p) \qquad (\mathbb{E}[R] = \frac{1}{p}, \,\mathbb{E}[1] = 1)$$

$$\implies \mathbb{E}[R^{2}] = \frac{2(1 - p)}{p^{2}} + \frac{1}{p} \implies \mathbb{E}[R^{2}] = \frac{2 - p}{p^{2}}$$

$$\implies \text{Var}[R] = \mathbb{E}[R^{2}] - (\mathbb{E}[R])^{2} = \frac{2 - p}{p^{2}} - \frac{1}{p^{2}} = \frac{1 - p}{p^{2}}$$

Continuing from the previous slide,

$$\mathbb{E}[R^2|A^c] = \sum_{t=0}^{\infty} (t+1)^2 \Pr[R=t] = \sum_{t=0}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=0}^{\infty} t \Pr[R=t] + \sum_{t=0}^{\infty} \Pr[R=t]$$

$$= \sum_{t=1}^{\infty} t^2 \Pr[R=t] + 2 \sum_{t=1}^{\infty} t \Pr[R=t] + \sum_{t=1}^{\infty} \Pr[R=t] = \mathbb{E}[R^2] + 2\mathbb{E}[R] + 1$$

Putting everything together,

$$\mathbb{E}[R^{2}] = (1)(p) + (\mathbb{E}[R^{2}] + 2\mathbb{E}[R] + 1])(1 - p) \implies p \,\mathbb{E}[R^{2}] = p + 2(1 - p)\mathbb{E}[R] + (1 - p)\mathbb{E}[1]$$

$$\implies p \,\mathbb{E}[R^{2}] = p + \frac{2(1 - p)}{p} + (1 - p) \qquad (\mathbb{E}[R] = \frac{1}{p}, \,\mathbb{E}[1] = 1)$$

$$\implies \mathbb{E}[R^{2}] = \frac{2(1 - p)}{p^{2}} + \frac{1}{p} \implies \mathbb{E}[R^{2}] = \frac{2 - p}{p^{2}}$$

$$\implies \text{Var}[R] = \mathbb{E}[R^{2}] - (\mathbb{E}[R])^{2} = \frac{2 - p}{p^{2}} - \frac{1}{p^{2}} = \frac{1 - p}{p^{2}}$$