## CMPT 210: Probability and Computing

Lecture 6

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### Recap

**Sample (outcome) space** S: Nonempty (countable) set of possible outcomes. Example: When we threw one dice, the sample space is  $\{1, 2, 3, 4, 5, 6\}$ .

**Outcome**  $\omega \in \mathcal{S}$ : Possible "thing" that can happen. Example: When we threw one dice, a possible outcome is  $\omega = 1$ .

**Event** E: Any subset of the sample space. Example: When we threw one dice, a possible event is  $E = \{6\}$  (first example) or  $E = \{3,6\}$  (second example).

**Probability function** on a sample space S is a total function  $Pr: S \to [0,1]$ . For any  $\omega \in S$ ,

$$0 \leq \Pr[\omega] \leq 1 \quad ; \quad \sum_{\omega \in \mathcal{S}} \Pr[\omega] = 1 \quad ; \quad \Pr[E] = \sum_{\omega \in E} \Pr[\omega]$$

**Union**: For mutually exclusive events  $E_1, E_2, \dots, E_n$ ,  $Pr[E_1 \cup E_2 \cup \dots E_n] = Pr[E_1] + Pr[E_2] + \dots + Pr[E_n]$ .

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### **Probability Rules**

Complement rule:  $Pr[E] = 1 - Pr[E^c]$ .

*Proof*: Recall that  $E \cap E^c = \{\}$  and  $E \cup E^c = S$ . Since E and  $E^c$  are disjoint,

$$\Pr[E \cup E^c] = \Pr[E] + \Pr[E^c] \implies \Pr[S] = \Pr[E] + \Pr[E^c] \implies \Pr[E^c] = 1 - \Pr[E].$$

Inclusion-Exclusion when for containing containing the second  $E, F, \Pr[E \cup F] = \Pr[E] + \Pr[F] - \Pr[E \cap F]$ .

*Proof*: May be on midterm, but it is unlikely

$$\Pr[E \cup F] = \sum_{\omega \in \{E \cup F\}} \Pr[\omega] = \sum_{\omega \in \{E - F\}} \Pr[\omega] + \sum_{\omega \in \{F - E\}} \Pr[\omega] + \sum_{\omega \in \{E \cap F\}} \Pr[\omega]$$
(Since disjoint)

$$\begin{split} &= \left[ \sum_{\omega \in \{E - F\}} \Pr[\omega] + \sum_{\omega \in \{E \cap F\}} \Pr[\omega] \right] + \left[ \sum_{\omega \in \{F - E\}} \Pr[\omega] + \sum_{\omega \in \{E \cap F\}} \Pr[\omega] \right] - \sum_{\omega \in \{E \cap F\}} \Pr[\omega] \\ &= \sum_{\omega \in E} \Pr[\omega] + \sum_{\omega \in F} \Pr[\omega] - \sum_{\omega \in \{E \cap F\}} \Pr[\omega] = \Pr[E] + \Pr[F] - \Pr[E \cap F] \end{split}$$

**Union Bound**: For any two events E, F,  $Pr[E \cup F] \leq Pr[E] + Pr[F]$ .

*Proof*: By the inclusion-exclusion rule,  $\Pr[E \cup F] = \Pr[E] + \Pr[F] - \Pr[E \cap F]$ . Since probabilities are non-negative,  $\Pr[E \cap F] \ge 0$  and hence,  $\Pr[E \cup F] \le \Pr[E] + \Pr[F]$ .

**Union Bound**: For any events  $E_1, E_2, E_3, \dots E_n$ ,

$$\Pr[E_1 \cup E_2 \cup E_3 \ldots \cup E_n] \leq \sum_{i=1}^n \Pr[E_i]$$
 se inclusion of

The events contained in B that are not within A have a non-zero probability of occuring

**Monotonicity rule**: For events  $A \text{ and } B = A \cup (B - A)$ , then Pr[A] < Pr[B].

$$Proof$$
:  $pr(B) = pr(A) + pr(B - A)$ 

$$\Pr[A] = \sum_{\omega \in A} \Pr[\omega] = \sum_{\omega \in B} \Pr[\omega] - \sum_{\omega \in \{B-A\}} \Pr[\omega] \implies \Pr[A] < \Pr[B]$$

(Since probabilities are non-negative.)

### **Uniform Probability Spaces**

# Events can have different probabilities, but outcomes must have different probabilities

**Definition**: A probability space is uniform if  $Pr[\omega]$  is the same for every outcome  $\omega \in \mathcal{S}$ .

Since 
$$\sum_{\omega \in \mathcal{S}} \Pr[\omega] = 1 \implies \Pr[\omega] = \frac{1}{|\mathcal{S}|}$$
 for all  $\omega \in \mathcal{S}$ .

Example: For a standard dice,  $\mathcal{S} = \{1, 2, 3, 4, 5, 6\}$ ,  $\Pr[\{1\}] = \Pr[\{2\}] = \ldots = \Pr[\{6\}] = \frac{1}{6}$ .

$$\Pr[E] = \sum_{\omega \in E} \Pr[\omega] = |E| \Pr[\omega] = \frac{|E|}{|S|}.$$

Example: For a standard dice, if  $E = \{3,6\}$ , then,  $Pr[E] = \frac{|E|}{|S|} = \frac{2}{6} = \frac{1}{3}$ .

Hence, for uniform probability spaces, computing the probability is equivalent to counting the outcomes we "care" about.

For uniform probability space, calculating probabilities is equivalent to counting size of space.

## e tBackstouthrowing dicenumber.

$$p(1) = p(3) = p(5) = p$$

pability of getting an even number is 2p

$$pr(2) = pr(4) = pr(6) = 2p$$

p probabilites of all outcomes must equal one.

Q:  $Sup^{9}P_{0}\overline{s}e^{1}$  we have a loaded (not "standard") dice such that the probability of getting an even number p = 1/9. Since that of getting an odd number (all even numbers are equally likely, and so are the odd numbers). What is the probability of getting a 6?

Let p be the probability of getting an odd number. Probability of getting an even number = 2p.

$$\sum_{\omega \in \mathcal{S}} \Pr[\omega] = 1 \implies 3p + 3(2p) = 1 \implies p = \frac{1}{9}$$
. Hence, probability of getting an odd number  $= \frac{1}{9}$ . Probability of getting a 6 = Probability of getting an even number  $= \frac{2}{9}$ .

- Q: What is the probability that we get either a 3 or a 6? 3/9 = 1/3
- Q: What is the probability that we get a prime number

1 is not prime.

Q: Suppose we select a card at random from a standard deck of 52 cards. What is the probability of getting:

- A spade 13/52 = 1/4
- A spade facecard 3/52 =
- A black card 26/52 =
- The gueen of hearts 1/52
- An ace 4/52 = 1/13

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**Q**: A class consists of 6 men and 4 women. An exam is given and the students are ranked according to their performance. Assuming that no two students obtain the same scores and all rankings are considered equally likely, what is the probability that women receive the top 4

SCOURITORM probability space

Change it into rearranging string of 1111 000000 In general, let the number of men be m and let the number of women be w.

Number of possible rankings = Number of permutations = (m + w)!.

The event of interest is that where the women achieve the top scores. In a possible ranking, let's fix the top w slots for women. The w women can be arranged in w! ways. And the m men can be arranged in m! ways. Sense; where women receive the top scores m! w!. E -> set of the rankings where the four women get the top scores pr(E) = |E|/|S| since it is a uniform probability space

Since all rankings are equally likely, probability that women receive the top w scores  $=\frac{m!w!}{(m+w)!}$ . In this case, since m=6 and w=4, probability that women receive the top 4 scores  $=\frac{6!4!}{10!}$ .

S -> all rankings

E -> women receive the top t scores

$$ISI = (m + w)!$$

Q: A class consists of m men and w women. An exam is given and the students are ranked according to their performance. Assuming that no two students obtain the same scores and all rankings are conveniently and the same scores ar

Number of ways to select the t women that have top scores  $= {w \choose t}$ . The top t women can be arranged in t! ways. The number of remaining students is equal to m+w-t. These can be arranged in (m+w-t)! ways. Hence, total number of rankings where women receive the top t scores  $= {w \choose t} (m+w-t)!$  t!.

As before, the total number of rankings = (m+w)!. Since all rankings are equally likely, the probability that women receive the top t scores =  $\frac{\binom{w}{t}(m+w-t)!}{(m+w)!} = \frac{w!(m+w-t)!}{(w-t)!(m+w)!}$ 

If t = w, then we get the answer on the previous slide

#### selection made randomly -> uniform probability space

S = set of all dosenittees considered so of size five which has 3 cs and two math students Q: A committee of size 5 is to be selected from a group of 6 CS and 9 Math students (no double majors allowed). If the selection is made randomly, what is the probability that the committee consists of 3 CS and 2 Math students?

Number of possible ways of selecting the committee =  $|S| = {15 \choose 5}$ .

The event of interest (E) requires choosing 3 CS and 2 Math students. Number of ways we can select the CS students  $= \binom{6}{3}$ . Similarly, number of ways we can select the Math students  $= \binom{9}{2}$ .

Hence,  $|E|=\binom{6}{3}\binom{9}{2}\implies \Pr[E]=\frac{|E|}{\binom{|E|}{S!}}=\frac{\binom{6}{3}\binom{15}{15}}{\binom{15}{15}}$ . On exam, you hust define was S and E is.

cause for every selection of the math students, you have a selection of the cs students

# S: ways to select k items Each item is equally siderby ions of k items that has alpha.

pr(E) = IEI/ISI

**Q**: From a set of n items a random sample of size k is to be selected. What is the probability a given item  $(\alpha)$  will be among the k selected items?

Number of ways of choosing the sample =  $\binom{n}{k}$ .

If we want a particular item in the sample, number of ways of choosing the other items  $=\binom{n-1}{k-1}$ .

Hence, probability that a given item will be among the k selected  $=\frac{\binom{n-1}{k-1}}{\binom{n}{k}}=\frac{k}{n}$ .

> event that alpha is one of the k selected

#### F -> event that B is selected

**Q**: From a set of n items a random sample of size k is to be selected. Given two items of interest:  $\alpha$  and  $\beta$ , what is the probability that (i) both  $\alpha$  and  $\beta$  will be among the k selected (ii) at least one of  $\alpha$  or  $\beta$  will be among the k selected?

- (i) If we want both  $\alpha$  and  $\beta$  to be in the sample, number of ways of choosing the other items =  $\binom{n-2}{k-2}$ . Hence, probability that both  $\alpha$  and  $\beta$  will be in the sample =  $\frac{\binom{n-2}{k-2}}{\binom{n}{k}} = \frac{k(k-1)}{n(n-1)}$ .
- (ii) Let A be the event that item  $\alpha$  is in the selection.  $\Pr[A] = \frac{k}{n}$ . Similarly B be the event that item  $\beta$  is in the selection.  $\Pr[B] = \frac{k}{n}$ . We want to compute  $\Pr[A \cup B]$ . By the union-rule,  $\Pr[A \cup B] = \Pr[A] + \Pr[B] \Pr[A \cap B]$ . Hence, probability that either  $\alpha$  or  $\beta$  will be among the k selected items  $= \frac{2k}{n} \frac{k(k-1)}{n(n-1)}$ . pr(ii) + pr(iii) = 1

(iii) If we want neither  $\alpha$  nor  $\beta$  to be in the sample, number of ways of choosing the items =

 $\binom{n-2}{k}$ . Hence, probability that neither  $\alpha$  nor  $\beta$  will be in the sample  $=\frac{\binom{n-2}{k}}{\binom{n}{k}}=\frac{(n-k)(n-k-1)}{n(n-1)}$ .

Calculate pr(E U F) = pr(E) + pr(F) - pr(EF)

Another way is to take the complement of pr(alpha and beta not in selection)

All permutations are equally likely
S -> all possible permuations
E -> all possible permutations where the third letter is B

Q: Let us consider random permutations of the letters (i) ABBA (ii) ABBA'. What is the probability that the third letter is B?

A' is different from A.

a: 3!/2! ways of rearranging remaining letters. There is 4!/2!2! ways to rearrange all the litters. The probability is 3/6 = 1/2

Assuming a' is unique: there are 4!/2! ways to rearrange the letters here are 3! ways to rearrange the letters such that the third letter is B the probability is 3! / 4! 2!



### Birthday Paradox

When using pidgeon hole principle, define the function and say what is a and b

Q: There are students in the class. What is the probability that two students have their birthdays in the same week? 1 by the pidgeon hole principle

Q: In this class, what is the probability that two students share the same birthday? Assume that (i) each student is equally likely to be born on any day of the year, (ii) no leap years and (iii) student birthdays are independent of each other.

Let n be the number of state of state of the number of days in the year. Let's order the students according to their 10-34 of days sequence is (11 Feb, 23 April, 31 August, ...). First S -> set of all possible birthday sequences

E -> set of all possible birthday sequences

E -> set of all birthday sequences where two students have the same birthday

The first student's birtiday can be one of 23 March milarly, the second student's birthday can be one of d days, and so on. By the product rule, the total number of birthday sequences = $d \times d \times \ldots = d^n$ 

 $ISI = 365^{75}$ 

### Birthday Paradox

The event of interest is that two students share the same birthday. Let us compute the probability of the event that NO two students share the same birthday, and then use the complement rule.

The first birthday can be chosen in d ways, the second in d-1 ways, and so on. By the generalized product rule, the number of birthday sequences such that no birthday is shared =  $d \times (d-1) \times (d-2) \times \dots (d-(n-1))$ .

Hence, the probability that no two students share the same birthday

$$=\frac{\frac{\text{the number of birthday sequences such that no birthday is shared}}{\text{total number of birthday sequences}} = \frac{\frac{d \times (d-1) \times (d-2) \times \dots (d-(n-1))}{d^n}}{\text{Text}}$$

$$= \left(1 - \frac{0}{d}\right) \times \left(1 - \frac{1}{d}\right) \dots \left(1 - \frac{n-1}{d}\right) \leq \exp(-0/d) \times \exp(-1/d) \dots \exp(-(n-1)/d)$$

$$(\text{for } x > 0, \ 1 - x \leq \exp(-x))$$

$$= \exp\left(\frac{-0}{d} + \frac{-1}{d} + \dots \frac{-(n-1)}{d}\right) = \exp\left(-\frac{n(n-1)}{2d}\right)$$

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### Birthday Paradox

Probability that two students share a birthday  $\geq 1 - \exp\left(-\frac{n(n-1))}{2d}\right)$ . Let's plot for d=365.

As n increases, the value of the expression goes closer to 1.

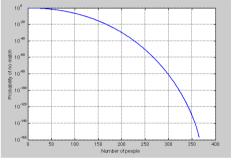


Figure 1: Plotting  $\exp\left(-\frac{n(n-1)}{2d}\right)$  for d=365

In our class, there is >96.4% that two students have the same birthday!

### Birthday Principle

If there are n pigeons and d pigeonholes, then the probability that two pigeons occupy the same hole is  $\geq 1 - \exp\left(-\frac{n(n-1)}{2d}\right)$ 

For  $n=\lceil \sqrt{2d} \rceil$ , probability that two pigeons occupy the same hole is about  $1-\frac{1}{e}\approx 0.632$ .

Example: If we are randomly throwing  $\lceil \sqrt{2d} \rceil$  balls into d bins, then the probability that two balls land in the same bin is around 0.632.

Later in the course, we will see applications of this principle to load balancing.

