

CMPT 210: Probability and Computing

Lecture 12

Sharan Vaswani

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Recap - (Basic) Freivald's Algorithm

- **Q:** For $n \times n$ matrices A , B and D , is $D = AB$? $O(n^2)$ time using Freivalds algorithm
- Last class, we proved that:

	Yes	No
$D = AB$	1	0
$D \neq AB$	$< \frac{1}{2}$	$\geq \frac{1}{2}$

Ask Sharan to go over the proof

Table 1: Probabilities for Basic Freivalds Algorithm

Probability amplification:
Want to amplify the probability of success.

Text

Frievald's Algorithm

By repeating the *Basic Frievald's Algorithm* m times, we will amplify the probability of success.
The resulting complete Frievald's Algorithm is given by:

- 1 Run the Basic Frievald's Algorithm for m independent runs.

Generate a new vector x .

Generation of the random vector is done independently.

Frievald's Algorithm

By repeating the *Basic Frievald's Algorithm* m times, we will amplify the probability of success. The resulting complete Frievald's Algorithm is given by:

- 1 Run the Basic Frievald's Algorithm for m independent runs.
- 2 If *any* run of the Basic Frievald's Algorithm outputs "no", output "no".
- 3 If *all* runs of the Basic Frievald's Algorithm output "yes", output "yes".

$$Dx = ABx$$

Every run is testing whether $D = AB$

Frievald's Algorithm

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- 3 If *all* runs of the Basic Frievald's Algorithm output "yes", output "yes".

If		Yes	No	As m increases, the probability of increasing the right answer increases.
	$D = AB$	1	0	
	$D \neq AB$	$< \frac{1}{2^m}$	$\geq 1 - \frac{1}{2^m}$	

Table 2: Probabilities for Frievald's Algorithm

Frievald's Algorithm

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- 2 If *any* run of the Basic Frievald's Algorithm outputs "no", output "no".
- 3 If *all* runs of the Basic Frievald's Algorithm output "yes", output "yes".

	Yes	No
$D = AB$	1	0
$D \neq AB$	$< \frac{1}{2^m}$	$\geq 1 - \frac{1}{2^m}$

Table 2: Probabilities for Frievald's Algorithm

If $m = 20$, then Frievald's algorithm will make mistake with probability $1/2^{20} \approx 10^{-6}$.

Computational Complexity: $O(mn^2)$
You run the algorithm m times,

Probability Amplification

Consider a randomized algorithm \mathcal{A} that is supposed to solve a binary decision problem i.e. it is supposed to answer either Yes or No. It has a one-sided error – (i) if the true answer is Yes, then the algorithm \mathcal{A} correctly outputs Yes with probability 1, but (ii) if the true answer is No, the algorithm \mathcal{A} incorrectly outputs Yes with probability $\leq \frac{1}{2}$.



No mistake if $D = AB$

Probability Amplification

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Let us define a new algorithm \mathcal{B} that runs algorithm \mathcal{A} m times, and if *any* run of \mathcal{A} outputs No, algorithm \mathcal{B} outputs No. If *all* runs of \mathcal{A} output Yes, algorithm \mathcal{B} outputs Yes.

Probability Amplification

Consider a randomized algorithm \mathcal{A} that is supposed to solve a binary decision problem i.e. it is supposed to answer either Yes or No. It has a one-sided error – (i) if the true answer is Yes, then the algorithm \mathcal{A} correctly outputs Yes with probability 1, but (ii) if the true answer is No, the algorithm \mathcal{A} incorrectly outputs Yes with probability $\leq \frac{1}{2}$.

Let us define a new algorithm \mathcal{B} that runs algorithm \mathcal{A} m times, and if *any* run of \mathcal{A} outputs No, algorithm \mathcal{B} outputs No. If *all* runs of \mathcal{A} output Yes, algorithm \mathcal{B} outputs Yes.

Q: What is the probability that algorithm \mathcal{B} correctly outputs Yes if the true answer is Yes, and correctly outputs No if the true answer is No?

Probability Amplification - Analysis

If A_i denotes run i of Algorithm \mathcal{A} , then

$$\Pr[\mathcal{B} \text{ outputs Yes} \mid \text{true answer is Yes}]$$

Each run is independent since each run's output is independent of the other's result.

$$= \Pr[\mathcal{A}_1 \text{ outputs Yes} \cap \mathcal{A}_2 \text{ outputs Yes} \cap \dots \cap \mathcal{A}_m \text{ outputs Yes} \mid \text{true answer is Yes}]$$

All runs must say Yes.

$$= \prod_{i=1}^m \Pr[\mathcal{A}_i \text{ outputs Yes} \mid \text{true answer is Yes}] = 1 \quad (\text{Independence of runs})$$

Probability Amplification - Analysis

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$$= \prod_{i=1}^m \Pr[\mathcal{A}_i \text{ outputs Yes} \mid \text{true answer is Yes}] = 1 \quad (\text{Independence of runs})$$

$$\Pr[\mathcal{B} \text{ outputs No} \mid \text{true answer is No}]$$

$$= 1 - \Pr[\mathcal{B} \text{ outputs Yes} \mid \text{true answer is No}]$$

$$= 1 - \Pr[\mathcal{A}_1 \text{ outputs Yes} \cap \mathcal{A}_2 \text{ outputs Yes} \cap \dots \cap \mathcal{A}_m \text{ outputs Yes} \mid \text{true answer is No}]$$

$$= 1 - \prod_{i=1}^m \Pr[\mathcal{A}_i \text{ outputs Yes} \mid \text{true answer is No}] \geq 1 - \frac{1}{2^m}.$$

Bounded by 1/2

Negation of the above statement

Probability Amplification - Analysis

If A_i denotes run i of Algorithm \mathcal{A} , then

$$\begin{aligned} & \Pr[\mathcal{B} \text{ outputs Yes} \mid \text{true answer is Yes}] \\ &= \Pr[\mathcal{A}_1 \text{ outputs Yes} \cap \mathcal{A}_2 \text{ outputs Yes} \cap \dots \cap \mathcal{A}_m \text{ outputs Yes} \mid \text{true answer is Yes}] \\ &= \prod_{i=1}^m \Pr[\mathcal{A}_i \text{ outputs Yes} \mid \text{true answer is Yes}] = 1 \end{aligned} \quad \text{(Independence of runs)}$$

$$\begin{aligned} & \Pr[\mathcal{B} \text{ outputs No} \mid \text{true answer is No}] \\ &= 1 - \Pr[\mathcal{B} \text{ outputs Yes} \mid \text{true answer is No}] \\ &= 1 - \Pr[\mathcal{A}_1 \text{ outputs Yes} \cap \mathcal{A}_2 \text{ outputs Yes} \cap \dots \cap \mathcal{A}_m \text{ outputs Yes} \mid \text{true answer is No}] \\ &= 1 - \prod_{i=1}^m \Pr[\mathcal{A}_i \text{ outputs Yes} \mid \text{true answer is No}] \geq 1 - \frac{1}{2^m}. \end{aligned}$$

When the true answer is Yes, both \mathcal{B} and \mathcal{A} correctly output Yes. When the true answer is No, \mathcal{A} incorrectly outputs Yes with probability $< \frac{1}{2}$, but \mathcal{B} incorrectly outputs Yes with probability $< \frac{1}{2^m} \ll \frac{1}{2}$. By repeating the experiment, we have “amplified” the probability of success.

\ll : means very small

Questions?

Concentration inequalities

Random Variables

Definition: A random “variable” R on a probability space is a total function whose domain is the sample space \mathcal{S} . The codomain is usually a subset of the real numbers.

$$F: \mathcal{S} \rightarrow V$$

Where V is a subset of \mathbb{R} :

Random Variables

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Example: Suppose we toss three independent, unbiased coins. Let C be the number of heads that appear.

Unbiased: $\Pr(\text{head}) = \Pr(\text{tails})$

$$\mathcal{S} = \{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\}$$

C is a total function that maps each outcome in \mathcal{S} to a number as follows: $C(HHH) = 3$, $C(HHT) = C(HTH) = C(THH) = 2$, $C(HTT) = C(THT) = C(TTH) = 1$, $C(TTT) = 0$.

C is a random variable that counts the number of heads in 3 tosses of the coin.

Everything up to the 12th lecture will be included

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C is a random variable that counts the number of heads in 3 tosses of the coin.

Example: I toss a coin, and define the random variable R which is equal to 1 when I get a heads, and equal to 0 when I get a tails.

Bernoulli random variables: Random variables with the codomain $\{0, 1\}$ are called Bernoulli random variables. E.g. R is a Bernoulli r.v.

Back to throwing dice

Q: Suppose we throw two standard dice one after the other. Let us define R to be the random variable equal to the sum of the dice. What is the domain, range of R ?

the sample space, which is $(1, 2, 3, 4, 5, 6)^2$

$$\text{Range}(R) = \{2, \dots, 12\}$$

Range is the values R can take on

Domain are the values that R can have as input.

Back to throwing dice

Q: Suppose we throw two standard dice one after the other. Let us define R to be the random variable equal to the sum of the dice. What is the domain, range of R ?

With replacement

Q: Three balls are randomly selected from an urn containing 20 balls numbered 1 through 20. The random variable M is the maximal value on the selected balls. What is the domain, range of

M ?
You have $\text{Dom}(M) = \{1, 2, 3, 4, 5, \dots, 20\}^3$
 $\text{Range}(M) = \{1, 2, 3, \dots, 20\}$

Back to throwing dice

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Q: In the above example, what is $2 \times M((1, 4, 6))$? Is M an invertible function?

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No, since you can have different tuples map to the same value.

Random Variables and Events

Indicates whether an event has happened.

Indicator Random Variable: An indicator random variable maps every outcome to either 0 or 1.

Example: Suppose we throw two standard dice, and define M to be the random variable that is 1 iff both throws of the dice produce a prime number, else it is 0.

$M : \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\} \rightarrow \{0, 1\}$. $M((2, 3)) = 1$, $M((3, 6)) = 0$.

An indicator random variable partitions the sample space into those outcomes mapped to 1 and those outcomes mapped to 0.

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The indicator random variable corresponding to an event E is denoted as \mathcal{I}_E , meaning that for $\omega \in E$, $\mathcal{I}_E[\omega] = 1$ and for $\omega \notin E$, $\mathcal{I}_E[\omega] = 0$. In the above example, $M = \mathcal{I}_E$ and since $(2, 4) \notin E$, $M((2, 4)) = 0$ and since $(3, 5) \in E$, $M((3, 5)) = 1$.

Indicates whether an event has happened or not.

Random Variables and Events

In general, a random variable that takes on several values partitions \mathcal{S} into several blocks.

Example: When we toss a coin three times, and define C to be the r.v. that counts the number of heads, C partitions \mathcal{S} as follows: $\mathcal{S} = \{\underbrace{HHH}_{C=3}, \underbrace{HHT, HTH, THH}_{C=2}, \underbrace{HTT, THT, TTH}_{C=1}, \underbrace{TTT}_{C=0}\}$.

Each block is a subset of the sample space and is therefore an event. For example, $[C = 2]$ is the event that the number of heads is two and consists of the outcomes $\{HHT, HTH, THH\}$.

Event

Event that the random variable takes on the value of 2

$$[C = i] = \{w \text{ in } \mathcal{S} \mid c[W] = i\}$$

$$\Pr(C = 2) = \Pr(\{HHT, HTH, THH\}) = 3/8.$$

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Since it is an event, we can compute its probability i.e.

$\Pr[C = 2] = \Pr[\{HHT, HTH, THH\}] = \Pr[\{HHT\}] + \Pr[\{HTH\}] + \Pr[\{THH\}]$. Since this is a uniform probability space, $\Pr[\omega] = \frac{1}{8}$ for $\omega \in \mathcal{S}$ and hence $\Pr[C = 2] = \frac{3}{8}$.

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Q: What is $\Pr[C = 0]$, $\Pr[C = 1]$ and $\Pr[C = 3]$?

Q: What is $\sum_{i=0}^3 \Pr[C = i]$?

Summing over all the partitions
is equivalent to summing the probability of
every event, which is equivalent to $\Pr(\mathcal{S})$

$\frac{3}{8}$

$\frac{1}{8}$

Random Variables and Events

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Q: What is $\Pr[C = 0]$, $\Pr[C = 1]$ and $\Pr[C = 3]$?

Q: What is $\sum_{i=0}^3 \Pr[C = i]$?

Since a random variable R is a total function that maps every outcome in \mathcal{S} to some value in the codomain, $\sum_{i \in \text{Range of } R} \Pr[R = i] = \sum_{i \in \text{Range of } R} \sum_{\omega \text{ s.t. } R(\omega)=i} \Pr[\omega] = \sum_{\omega \in \mathcal{S}} \Pr[\omega] = 1$.

Q: Suppose we throw two standard dice one after the other. Let us define R to be the random variable equal to the sum of the dice. What are the outcomes in the event $[R = 2]$?

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Q: What is $\Pr[R = 4]$, $\Pr[R = 9]$?

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Q: What is $\Pr[R = 4]$, $\Pr[R = 9]$?

Q: If M is the indicator random variable equal to 1 iff both throws of the dice produces a prime number, what is $\Pr[M = 1]$?

9/36

$\{2, 3\}, \{3, 5\}, \{5, 2\}, \{2, 5\}, \{5, 3\}, \{3, 3\}, \{2, 2\}, \{5, 5\}$

Distribution Functions

Probability density function (PDF): Let R be a random variable with codomain V . The probability density function of R is the function $\text{PDF}_R : V \rightarrow [0, 1]$, such that $\text{PDF}_R[x] = \Pr[R = x]$ if $x \in \text{Range}(R)$ and equal to zero if $x \notin \text{Range}(R)$.

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$$\sum_{x \in V} \text{PDF}_R[x] = \sum_{x \in \text{Range}(R)} \Pr[R = x] = 1.$$

Sum of a pdf is 1.

If there is something that cannot happen, the pdf of that event is 0.

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$$\sum_{x \in V} \text{PDF}_R[x] = \sum_{x \in \text{Range}(R)} \Pr[R = x] = 1.$$

Cumulative distribution function (CDF): If the codomain is a subset of the real numbers, then the cumulative distribution function is the function $\text{CDF}_R : \mathbb{R} \rightarrow [0, 1]$, such that $\text{CDF}_R[x] = \Pr[R \leq x]$. Equality is important

Importantly, neither PDF_R nor CDF_R involves the sample space of an experiment.

Abstract from specific examples

PDF and CDF correspond to a specific random variable.

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$$\sum_{x \in V} \text{PDF}_R[x] = \sum_{x \in \text{Range}(R)} \Pr[R = x] = 1. \quad \text{Remove use of sample space}$$

Cumulative distribution function (CDF): If the codomain is a subset of the real numbers, then the cumulative distribution function is the function $\text{CDF}_R : \mathbb{R} \rightarrow [0, 1]$, such that $\text{CDF}_R[x] = \Pr[R \leq x]$.

Remove use of Codomain

Importantly, neither PDF_R nor CDF_R involves the sample space of an experiment.

Example: If we flip three coins, and C counts the number of heads, then

$\text{PDF}_C[0] = \Pr[C = 0] = \frac{1}{8}$, and

$$\text{CDF}_C[2.3] = \Pr[C \leq 2.3] = \Pr[C = 0] + \Pr[C = 1] + \Pr[C = 2] = \frac{7}{8}.$$

variable C takes on values less than or equal to 2.3

$\text{Pdf}(4) = 0$ since we can never see 4 heads in three coin flips.

Distribution Functions

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Cumulative distribution function (CDF): If the codomain is a subset of the real numbers, then the cumulative distribution function is the function $\text{CDF}_R : \mathbb{R} \rightarrow [0, 1]$, such that $\text{CDF}_R[x] = \Pr[R \leq x]$.

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$$\text{CDF}_C[2.3] = \Pr[C \leq 2.3] = \Pr[C = 0] + \Pr[C = 1] + \Pr[C = 2] = \frac{7}{8}.$$

Q: What is $\text{CDF}_C[5.8]$? .

1 since you have summed across the range.

Distribution Functions

Probability density function (PDF): Let R be a random variable with codomain V . The probability density function of R is the function $\text{PDF}_R : V \rightarrow [0, 1]$, such that $\text{PDF}_R[x] = \Pr[R = x]$ if $x \in \text{Range}(R)$ and equal to zero if $x \notin \text{Range}(R)$.

$$\sum_{x \in V} \text{PDF}_R[x] = \sum_{x \in \text{Range}(R)} \Pr[R = x] = 1.$$

Cumulative distribution function (CDF): If the codomain is a subset of the real numbers, then the cumulative distribution function is the function $\text{CDF}_R : \mathbb{R} \rightarrow [0, 1]$, such that $\text{CDF}_R[x] = \Pr[R \leq x]$.

Importantly, neither PDF_R nor CDF_R involves the sample space of an experiment.

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$$\text{PDF}_C[0] = \Pr[C = 0] = \frac{1}{8}, \text{ and}$$

$$\text{CDF}_C[2.3] = \Pr[C \leq 2.3] = \Pr[C = 0] + \Pr[C = 1] + \Pr[C = 2] = \frac{7}{8}.$$

Q: Since most things are smaller than infinity, you will approach one.
Since most things are larger than -infinity, you will approach zero as you obtain less events.

For a general random variable R , as $x \rightarrow \infty$, $\text{CDF}_R[x] \rightarrow 1$ and $x \rightarrow -\infty$, $\text{CDF}_R[x] \rightarrow 0$.

Back to throwing dice

Q: Suppose we throw two standard dice one after the other. Let us define T to be the random variable equal to the sum of the dice. Plot PDF_T and CDF_T

$S: (1, 2, 3, 4, 5, 6)^2$

$T(2, 3) = 5$

$T: S \rightarrow V \text{ st } T(i, j) = i + j$

$\text{PDF}_T: V \rightarrow [0, 1]$

$\text{PDF}(2) = \Pr(T = 2) = \Pr(\{1, 1\}) = 1/36$

$\text{PDF}(3, 2) = 0$ since it is not part of the codomain

$\text{DF}[3] = \Pr(T = 3) = \Pr(\{1, 2\}, \{2, 1\}) = 2/36$

$\text{CDF}: R \rightarrow [0, 1]$

$\text{CDF}(2.3) = \Pr(c \leq 2.3) = \Pr(c = 2) = 1/36$

$\text{CDF}(3) = \Pr(C \leq 3) = \Pr(C = 2) + \Pr(C = 3) = 3/36$

$\text{CDF}(2.9999) = \Pr(C \leq 2.999) = \Pr(C = 2)$

$\text{CDF}(3.00001) = \Pr(C \leq 3.00001) = \Pr(C = 3) + \Pr(C = 2)$

Back to throwing dice

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Recall that $T : \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\} \rightarrow V$ where $V = \{2, 3, 4, \dots, 12\}$.

$\text{PDF}_T : V \rightarrow [0, 1]$ and $\text{CDF}_T : \mathbb{R} \rightarrow [0, 1]$.

For example, $\text{PDF}_T[4] = \Pr[T = 4] = \frac{3}{36}$ and $\text{PDF}_T[12] = \Pr[T = 12] = \frac{1}{36}$.

Back to throwing dice

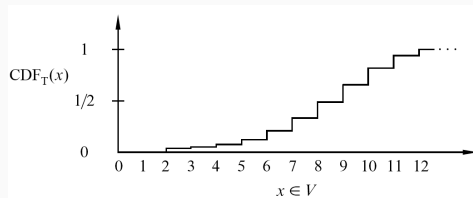
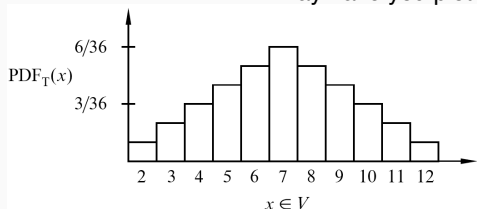
Q: Suppose we throw two standard dice one after the other. Let us define T to be the random variable equal to the sum of the dice. Plot PDF_T and CDF_T

Recall that $T : \{1, 2, 3, 4, 5, 6\} \times \{1, 2, 3, 4, 5, 6\} \rightarrow V$ where $V = \{2, 3, 4, \dots, 12\}$.

$\text{PDF}_T : V \rightarrow [0, 1]$ and $\text{CDF}_T : \mathbb{R} \rightarrow [0, 1]$.

For example, $\text{PDF}_T[4] = \Pr[T = 4] = \frac{3}{36}$ and $\text{PDF}_T[12] = \Pr[T = 12] = \frac{1}{36}$.

May have you plot out CDF or PDF in exam



CDF for discrete distributions is like a step function

Questions?

Distributions

Many random variables turn out to have the same PDF and CDF. In other words, even though R and T might be different random variables on different probability spaces, it is often the case that $\text{PDF}_R = \text{PDF}_T$. Hence, by studying the properties of such PDFs, we can study different random variables and experiments.

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Distribution over a random variable can be fully specified using the cumulative distribution function (CDF) (usually denoted by F). The corresponding probability density function (PDF) is denoted by f .

Distribution is interchangeable with telling you the CDF or the PDF

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Common Discrete Distributions in Computer Science:

- Bernoulli Distribution
- Next lecture will talk about the PDF and CDF of each of these distributions
- Uniform Distribution
 - Binomial Distribution
 - Geometric Distribution

Bernoulli Distribution

Canonical Example: We toss a biased coin such that the probability of getting a heads is p . Let R be the random variable such that $R = 1$ when the coin comes up heads and $R = 0$ if the coin comes up tails. R follows the Bernoulli distribution.

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PDF _{R} for Bernoulli distribution: $f: \{0, 1\} \rightarrow [0, 1]$ meaning that Bernoulli random variables take values in $\{0, 1\}$. It can be fully specified by the “probability of success” (of an experiment) p (probability of getting a heads in the example). Formally, PDF _{R} is given by:

$$f(1) = p \quad ; \quad f(0) = q := 1 - p.$$

In the example, $\Pr[R = 1] = f(1) = p = \Pr[\text{event that we get a heads}]$.

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Q: If X has a Bernoulli distribution, when is X also uniform?

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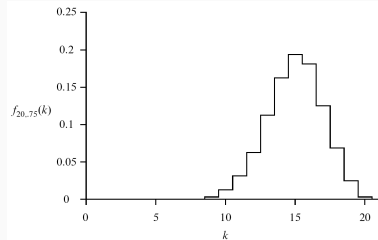
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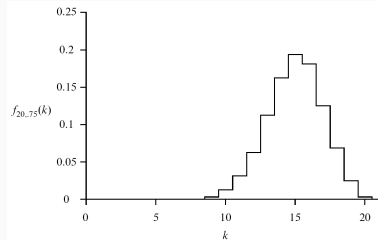
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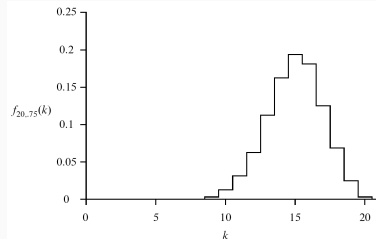
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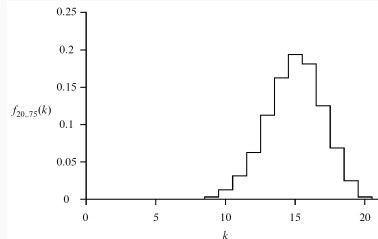


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By the Binomial Theorem, $\sum_{k \in \text{Range}(R)} \text{PDF}_R[k] = \sum_{k=0}^n \binom{n}{k} p^k (1-p)^{n-k} = (p + 1 - p)^n = 1$.

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$$= 1. \quad (\text{for } x \geq n)$$

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Canonical Example: We toss a biased coin independently multiple times. The probability of getting a heads is p . Let R be the random variable equal to the number of tosses needed to get the first heads. R follows the geometric distribution.

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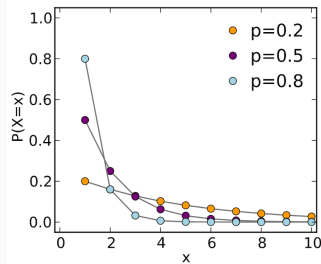
By the sum of geometric series, $\sum_{k \in \text{Range}(R)} \text{PDF}_R[k] = \sum_{k=1}^{\infty} (1 - p)^{k-1} p = \frac{p}{1 - (1 - p)} = 1$.

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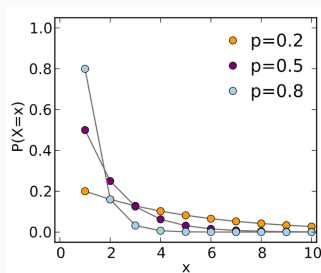
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