

Varying impacts of letters of recommendation on college admissions

Approximate balancing weights for subgroup effects in observational studies

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(joint work with Avi Feller and Jesse Rothstein)

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- 2016: UC Berkeley pilot program requests **letters of recommendation** (LORs) for undergrad admission
 - No LOR requirement at other UCs/CSUs
- *Goal:* Identify students from non-traditional backgrounds who might be overlooked
 - "Holistic review" of applicants
- *Concern:* Adverse impact on disadvantaged applicants, especially under-represented minority (URM) students



Chair of the UC Berkeley Academic Senate

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~> **LORs discontinued** before study results released

LOR Pilot Program: Subgroup Effects

Our question: Impact of submitting LORs on admissions

- Variation across pre-defined subgroups, especially URM status
- Design an observational study; one of many potential cuts at this problem

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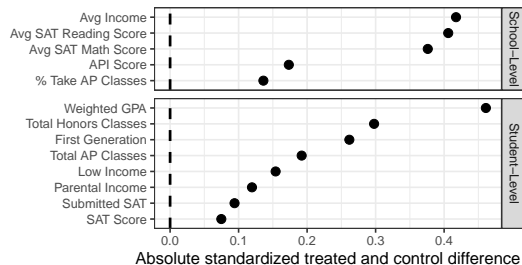
- Partially pooled balancing weights, control both **local balance** and **global balance**
- *Dual:* IPW with hierarchical propensity score

No evidence of differential impacts on URM applicants

Letters of Recommendation: Pilot Study

LOR Pilot Study: Overview

- Total $N = 40,541$ applicants in 2016
[exclude athletes, other groups]
 - 14,596 invited to submit LORs
 - 11,143 submitted LoRs
- Two admissions readers
 - Scores of {No, Possible, Yes}
 - Admitted with 1-2 Yes votes
- Invitation to submit LORs:
[+ funkiness due to timing]
 - First reader score of "possible"
 - Predicted possible score of >50%
[nearly all URM]



LoR Pilot Study: Subgroups

URM: Under-Represented Minority

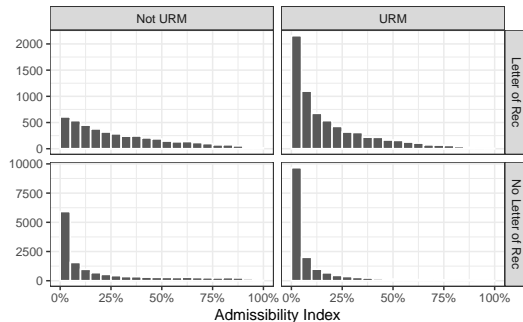
- Low-income or first-gen college
 - Underrepresented racial/ethnic group
 - Low-performing high school
- [~ 55% of all applicants]

AI: Admissibility Index

- Predicted prob. of admissions using 2015 data

Define subgroups by URM \times AI bin

- + First reader score; college applied to



Setup and Background

Setup

For applicant $i = 1, \dots, N$ observe

- Outcome $Y_i \in \mathbb{R}$ (admission)
- Treatment status $W_i \in \{0, 1\}$ (submit LoRs)
- Covariates $X_i \in \mathcal{X}$
- Group indicator $G_i \in \{1, \dots, K\}$

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Estimands: Overall ATT and subgroup CATTs

$$\tau = \mathbb{E}[Y(1) - \underbrace{Y(0)}_{\hat{\mu}_0 = \sum \hat{\gamma}_i Y_i} \mid W = 1] \quad \text{and} \quad \tau_g = \mathbb{E}[Y(1) - \underbrace{Y(0)}_{\hat{\mu}_{0g} = \sum_{G=g} \hat{\gamma}_i Y_i} \mid W = 1, G = g]$$

Setup

Strong ignorability (sensitivity analysis in paper)

$$Y(1), Y(0) \perp\!\!\!\perp W \mid X, G \quad \text{and} \quad e(X, G) \equiv P(W = 1 \mid X, G) < 1$$

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Many methods for subgroup effects under ignorability

- **Outcome model** and **design-based** approaches
- *Review: 2018 ACIC data challenge* [Carvalho et al., 2020]

Inverse Propensity Score Weighting identities, with known $e(x, g)$

$$\mu_0 = \mathbb{E}[\text{missing } Y(0) \mid \text{treated}] = \mathbb{E}\left[\underbrace{\frac{e(x, g)}{1 - e(x, g)}}_{\text{weights}} Y^{\text{obs}} \mid \text{control}\right]$$

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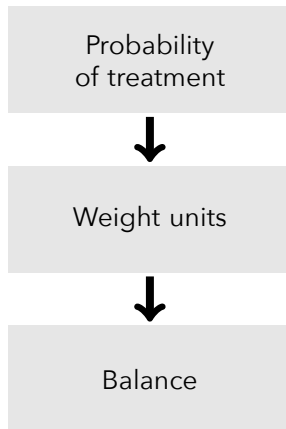
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~> How to estimate weights?

Background: Traditional IPW workflow

Goal: $\hat{e}(x, g)$ close to $e(x, g)$

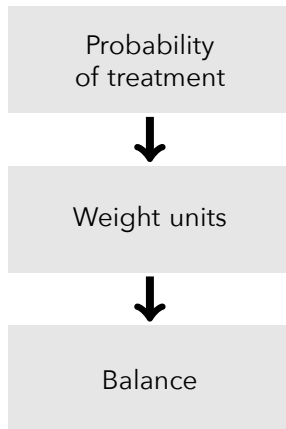
1. **Directly** estimate $\hat{e}(x, g)$, via MLE, ML, etc.
2. Calculate weights $\hat{\gamma} = \frac{\hat{e}(x, g)}{1 - \hat{e}(x, g)}$
3. **Indirectly** balance covariates



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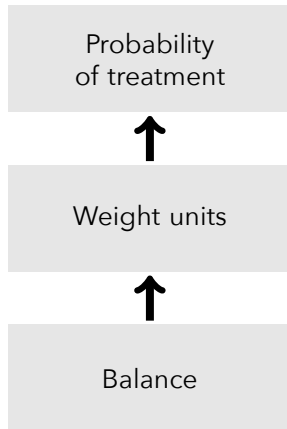
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3. **Indirectly** balance covariates
 - Poor finite sample performance, esp with many covariates



Background: Balancing weights workflow

Goal: $\hat{\gamma}$ close to $\frac{e(x,g)}{1-e(x,g)}$

1. **Directly** estimate $\hat{\gamma}$ to balance covariates
2. **Indirectly** estimate $\hat{e}(x, g) = \frac{\hat{\gamma}}{1+\hat{\gamma}}$



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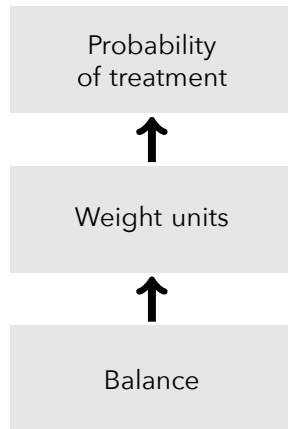
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- Old history as raking and calibration in survey sampling with non-response

[Deming and Stephan, 1940; Deville et al., 1993]

- New causal inference literature

[Hainmueller, 2011; Zubizarreta, 2015; Athey et al., 2018; Chernozhukov et al., 2018]



Balancing weights to estimate
subgroup effects

Balancing weights for local balance only

Error for effect in subgroup g

For outcome model $m_0 = \eta_g \cdot \phi(x)$; weighting estimator $\hat{\mu}_{0g} = \sum \gamma Y_i$

$$\text{error}_g \leq \|\eta_g\|_2 \|\text{Local Balance}_g\|_2 + \|\gamma\|_2$$

Can generalize to flexible outcome models [Hirshberg et al., 2019; Hazlett, 2020]

Balancing local balance only

Balancing weights for subgroup g

$$\begin{aligned} \min_{\gamma} \quad & \|\text{Local Balance}_g\|_2^2 + \frac{\lambda_g}{2} \|\gamma\|_2^2 \\ \text{s.t.} \quad & \sum \gamma_i = 1, \quad \gamma_i \geq 0 \end{aligned}$$

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Challenge:

- Small subgroups can be hard to balance well
- Balancing subgroups separately \rightarrow poor global balance

Balancing global balance and local balance

Partially Pooled Balancing Weights

$$\min_{\gamma} \quad \sum_g \|\text{Local Balance}_g\|_2^2 + \frac{\lambda_g}{2} \|\gamma\|_2^2$$

$$\text{s.t.} \quad \sum_{G_i=g} \gamma_i = n_{1g}, \quad \gamma_i \geq 0$$

$$\text{Global Balance} = 0$$

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$$\text{s.t.} \quad \sum_{G_i=g} \gamma_i = n_{1g}, \quad \gamma_i \geq 0$$

$$\text{Global Balance} = 0$$

- Overall errors depends on both global balance and local balance
- Further expand to control *differences* in local balance

~> *Tuning parameter*: Global parameter $\lambda \Rightarrow \lambda_g = \lambda/n_g$

Dual perspective: M estimation of treatment odds

Dual when optimizing for for **local balance** only

$$\text{Population: } \underbrace{\frac{e(x, g)}{1 - e(x, g)}}_{\text{inverse prop. score weights}} \sim \underbrace{\alpha_g + \beta'_g \phi(x)}_{\text{balancing weights}}$$

Dual perspective: M estimation of treatment odds

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Population: $\underbrace{\frac{e(x, g)}{1 - e(x, g)}}_{\text{inverse prop. score weights}} \sim \underbrace{\alpha_g + \beta_g' \phi(x)}_{\text{balancing weights}}$

Sample: $\min_{\alpha_g, \beta_g} \text{regression loss} + \underbrace{\frac{\lambda}{2} \|\beta_g\|_2^2}_{\text{ridge penalty}}$

Global balance constraint \longleftrightarrow partial pooling in the dual problem

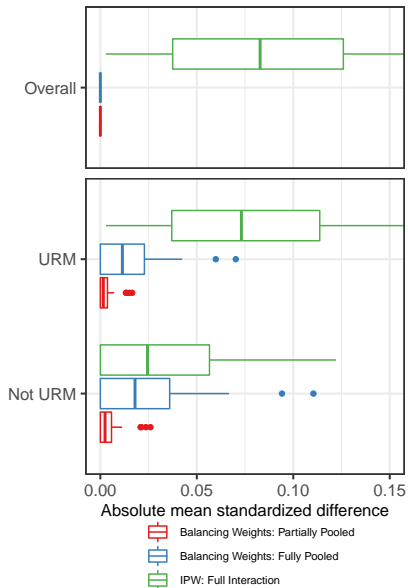
Dual for Partially Pooled Balancing Weights

$$\min_{\alpha_g, \beta_g, \mu_\beta} \text{regression loss} + \underbrace{\frac{\lambda_g}{2} \|\beta_g - \mu_\beta\|_2^2}_{\text{local} \rightarrow \text{global}}$$

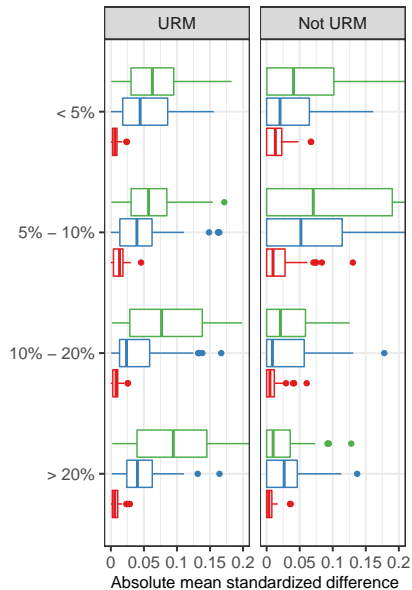
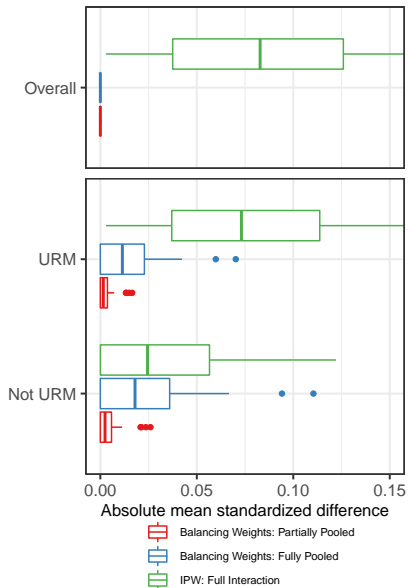
Partially pool $\text{local} \rightarrow \text{global}$ model: regularization *directly* related to imbalance

Differential impacts of letters of
recommendation

Partially pooled balancing weights \rightarrow improved balance

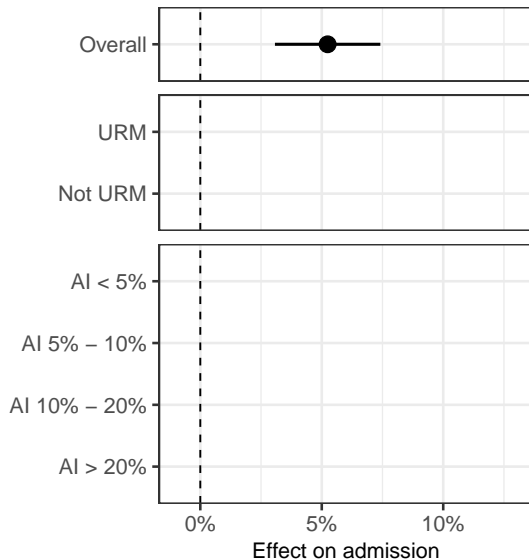


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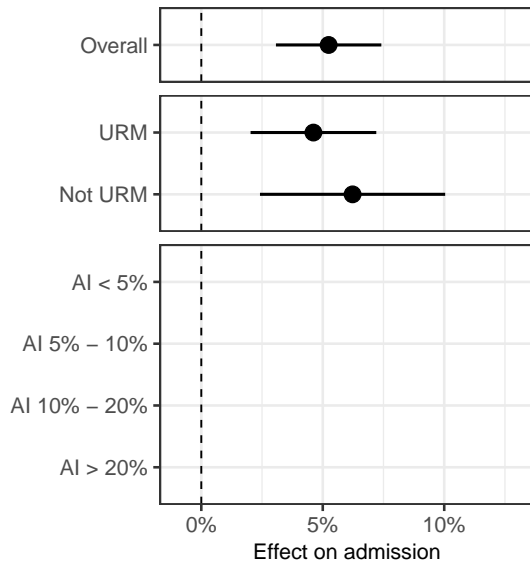


Large overall effect

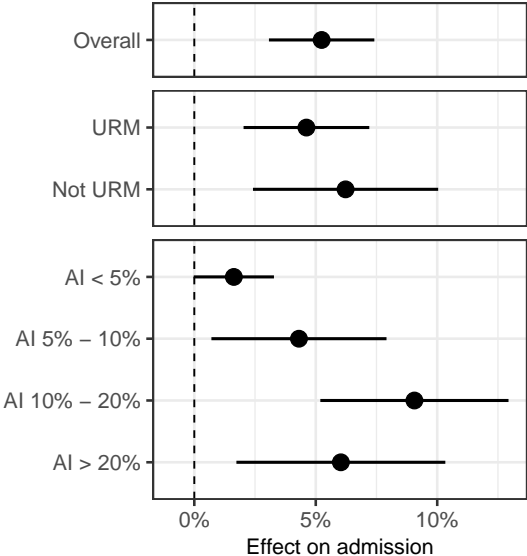
Baseline admissions rate $\sim 20\%$



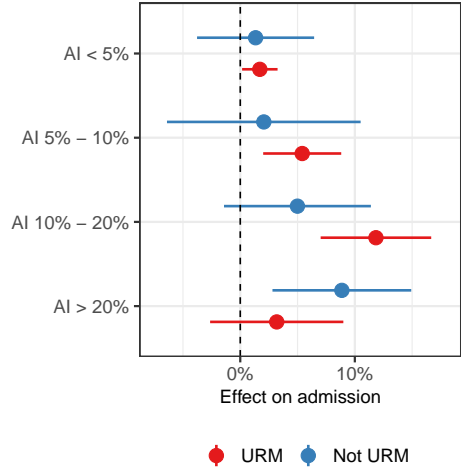
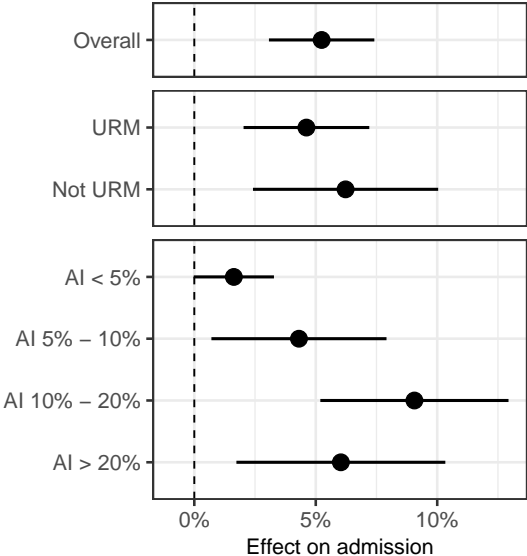
No differences by URM status



Large differences by Admissibility Index



Relative effect sizes flip



Recap: Varying Impacts of Letters of Recommendation

Partially pooled balancing weights

- Find weights that control both **Local Balance** and **Global Balance**
 - Dual relation to partially pooled IPW
 - R package **balancer**
- No evidence of different impacts by URM status

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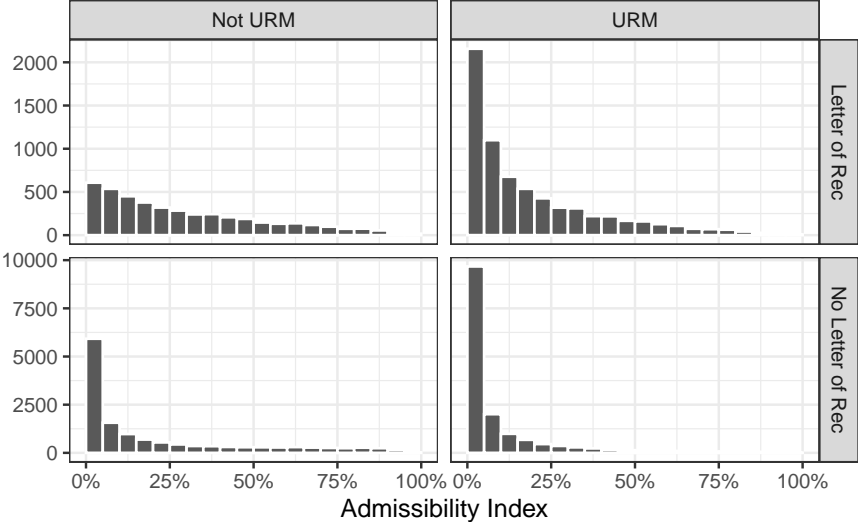
Thank you!

ebenmichael.github.io

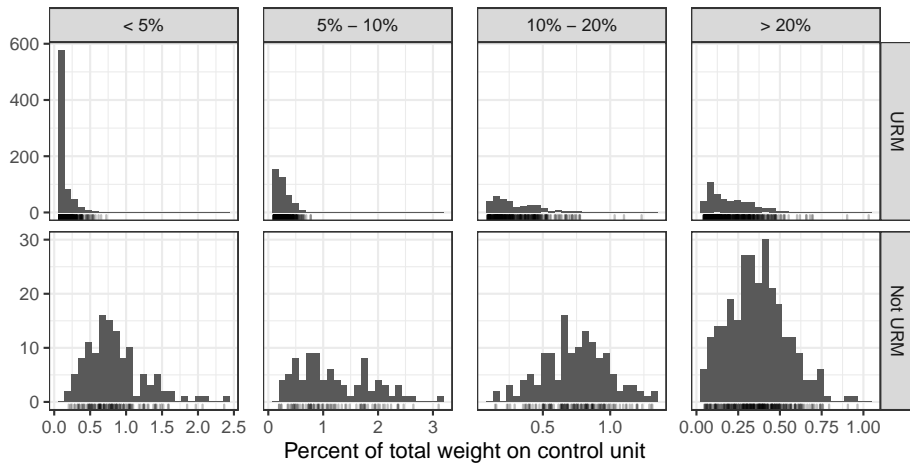


Appendix

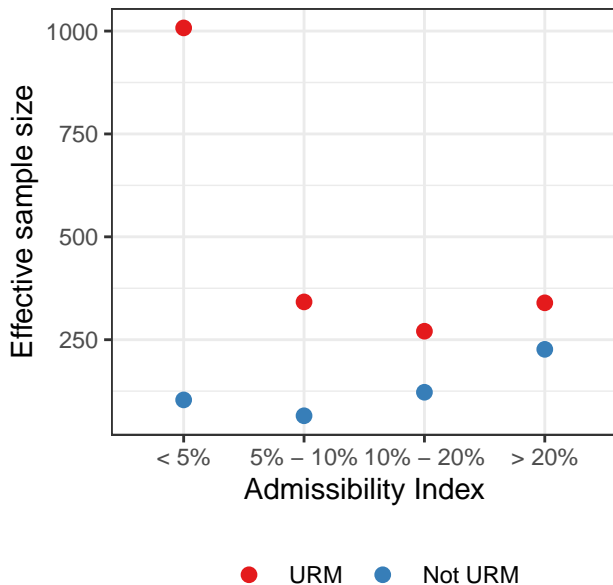
Heterogeneity across admissibility index



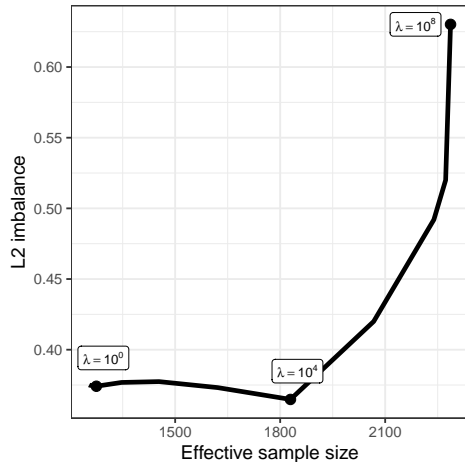
Distribution of the estimated weights



Effective sample sizes

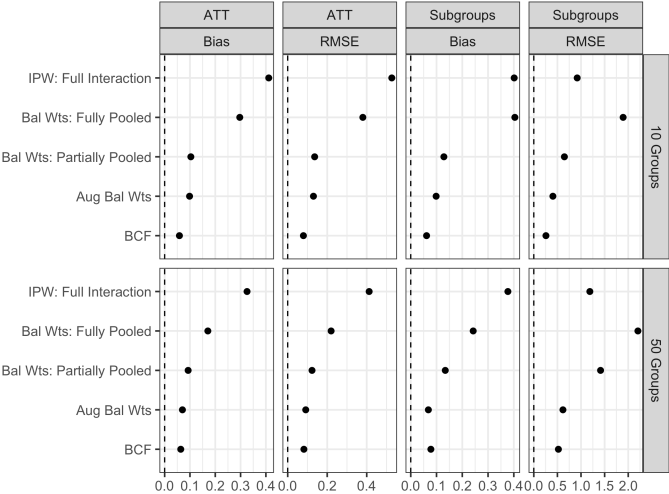


Hyperparameter tuning



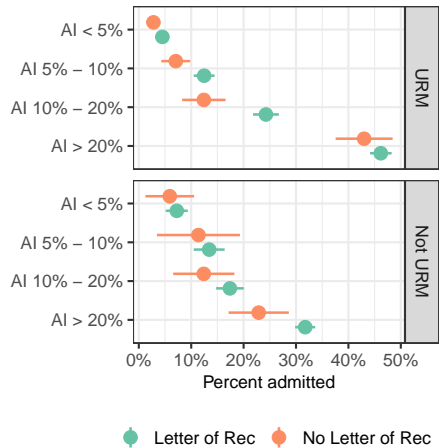
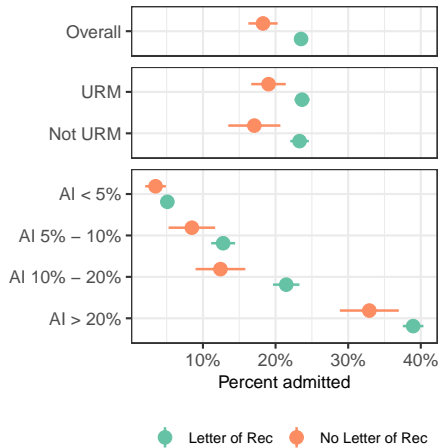
- Evaluate across a range of λ
- Gains in precision, comparable imbalance

Simulation Study



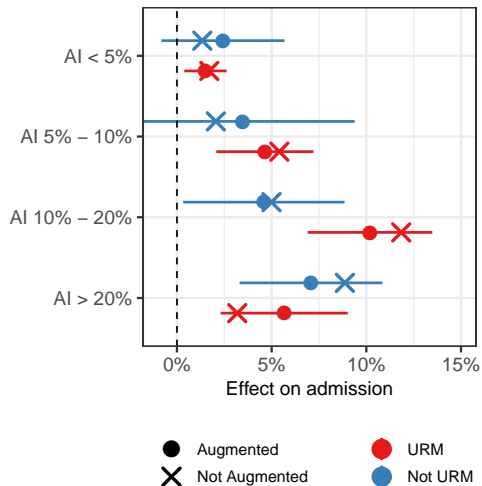
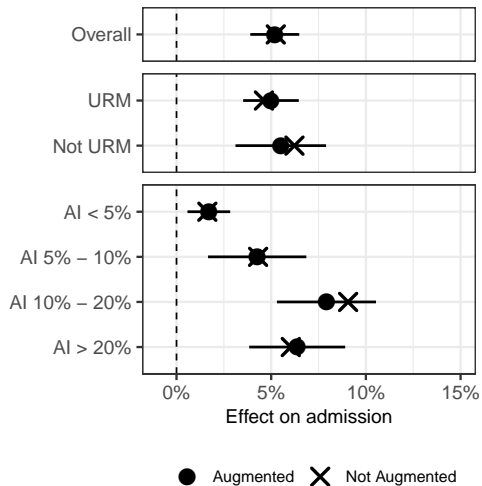
- Major gains relative to traditional IPW
- Comparable performance to ML methods; retain design-based advantages

Estimated group means



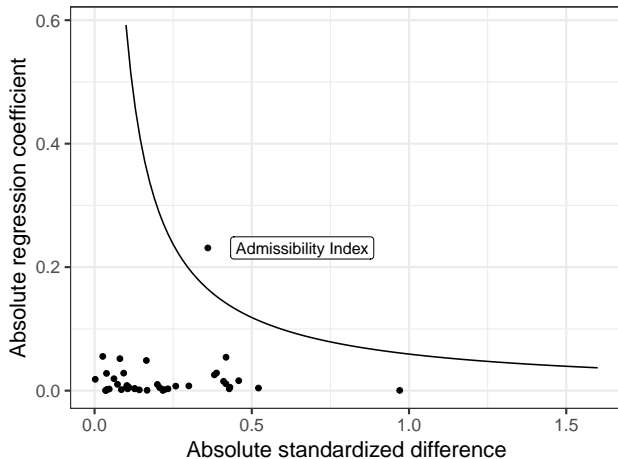
Augmentation diminishes differences

[Random forest outcome model]



Sensitivity to unmeasured confounding

- Adapt Soriano et al. [2021]
- Overall LOR effect still positive with $\Lambda = 1.1$
- Consistent with wide range of subgroup estimates



References I

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