Varying impacts of letters of recommendation on college admissions

Approximate balancing weights for subgroup effects in observational studies

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(joint work with Avi Feller and Jesse Rothstein)

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- 2016: UC Berkeley pilot program requests letters of recommendation (LORs) for undergrad admission
 - No LOR requirement at other UCs/CSUs

- Goal: Identify students from non-traditional backgrounds who might be overlooked
 - "Holistic review" of applicants

 Concern: Adverse impact on disadvantaged applicants, especially under-represented minority (URM) students



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→ LORs discontinued before study results released

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- Partially pooled balancing weights, control both local balance and global balance
- Dual: IPW with hierarchical propensity score

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No evidence of differential impacts on URM applicants

Letters of Recommendation:

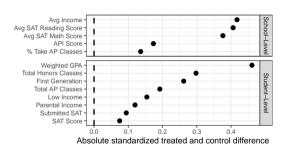
Pilot Study

LOR Pilot Study: Overview

- Total N = 40,541 applicants in 2016 [exclude athletes, other groups]
 - 14,596 invited to submit LORs
 - 11,143 submitted LoRs
- Two admissions readers
 - Scores of {No, Possible, Yes}
 - Admitted with 1-2 Yes votes
- Invitation to submit LORs:

[+ funkiness due to timing]

- First reader score of "possible"
- Predicted possible score of >50% [nearly all URM]



LoR Pilot Study: Subgroups

URM: Under-Represented Minority

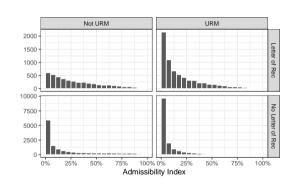
- Low-income or first-gen college
- Underrepresented racial/ethnic group
- Low-performing high school
 [~ 55% of all applicants]

Al: Admissibility Index

- Predicted prob. of admissions using 2015 data
- Strong prognostic indicator!

Define subgroups by URM × AI bin

+ First reader score; college applied to



Setup and

Background

For applicant i = 1, ..., N observe

- Outcome $Y_i \in \mathbb{R}$ (admission)
- Treatment status $W_i \in \{0, 1\}$ (submit LoRs)
- Covariates $X_i \in \mathcal{X}$
- Group indicator $G_i \in \{1, \dots, K\}$

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Estimands: Overall ATT and subgroup CATTs

$$\tau = \mathbb{E}[Y(1) - \underbrace{Y(0)}_{\widehat{\mu}_0 = \sum \widehat{\gamma}_i Y_i} \mid W = 1] \quad \text{and} \quad \tau_g = \mathbb{E}[Y(1) - \underbrace{Y(0)}_{\widehat{\mu}_{0g} = \sum_{G = g} \widehat{\gamma}_i Y_i} \mid W = 1, G = g]$$

Strong ignorability (sensitivity analysis in paper)

$$Y(1), Y(0) \perp W \mid X, G$$
 and $e(X, G) \equiv P(W = 1 \mid X, G) < 1$

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Many methods for subgroup effects under ignorability

- Outcome model and design-based approaches
- Review: 2018 ACIC data challenge [Carvalho et al., 2020]

Inverse Propensity Score Weighting identities, with known e(x, g)

$$\mu_0 = \mathbb{E}[\text{missing } Y(0) \mid \text{treated}] = \mathbb{E}\Big[\underbrace{\frac{e(x,g)}{1 - e(x,g)}}_{\text{weights}} Y^{\text{obs}} \mid \text{control}\Big]$$

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weights

→ How to estimate weights?

Background: Traditional IPW workflow

Goal: $\hat{e}(x,g)$ close to e(x,g)

- 1. Directly estimate $\hat{e}(x,g)$, via MLE, ML, etc.
- 2. Calculate weights $\hat{\gamma} = \frac{\hat{e}(x,g)}{1 \hat{e}(x,g)}$
- 3. Indirectly balance covariates

Probability of treatment



Weight units



Balance

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- 3. Indirectly balance covariates
- Poor finite sample performance, esp with many covariates

Probability of treatment



Weight units



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Background: Balancing weights workflow

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$$\hat{\gamma}$$
 close to $\frac{e(x,g)}{1-e(x,g)}$

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- 2. Indirectly estimate $\hat{\mathbf{e}}(x,g) = \frac{\hat{\gamma}}{1+\hat{\gamma}}$
 - Old history as raking and calibration in survey sampling with non-response
 [Deming and Stephan, 1940; Deville et al., 1993]
 - New causal inference literature
 [Hainmueller, 2011; Zubizarreta, 2015; Athey et al.,
 2018; Chernozhukov et al., 2018]

Probability of treatment Weight units Balance

Balancing weights to estimate

subgroup effects

Balancing weights for local balance only

Error for effect in subgroup g

For outcome model
$$m_0 = \eta_g \cdot \phi(x)$$
; weighting estimator $\hat{\mu}_{0g} = \sum \gamma Y_i$

$$\operatorname{error}_{g} \leq \|\eta_{g}\|_{2} \|\operatorname{Local Balance}_{g}\|_{2} + \|\gamma\|_{2}$$

Can generalize to flexible outcome models [Hirshberg et al., 2019; Hazlett, 2020]

Balancing local balance only

Balancing weights for subgroup g

$$\begin{aligned} & \min_{\gamma} & & \| \text{Local Balance}_g \|_2^2 & + & \frac{\lambda_g}{2} \| \gamma \|_2^2 \\ & \text{s.t.} & & \sum \gamma_i = 1, \quad \gamma_i \geq 0 \end{aligned}$$

Balancing local balance only

Balancing weights for subgroup g

$$\min_{\gamma} \quad \|\text{Local Balance}_g\|_2^2 \ + \ \frac{\lambda_g}{2} \|\gamma\|_2^2$$
 s.t.
$$\sum \gamma_i = 1, \quad \gamma_i \geq 0$$

Challenge:

- Small subgroups can be hard to balance well
- Balancing subgroups separately \rightarrow poor global balance

Balancing global balance and local balance

Partially Pooled Balancing Weights

$$\min_{\gamma} \quad \sum_{g} \|\text{Local Balance}_{g}\|_{2}^{2} \ + \ \frac{\lambda_{g}}{2} \|\gamma\|_{2}^{2}$$
 s.t.
$$\sum_{G_{i}=g} \gamma_{i} = n_{1g}, \quad \gamma_{i} \geq 0$$
 Global Balance $= 0$

Balancing global balance and local balance

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 s.t.
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 Global Balance = 0

- Overall errors depends on both global balance and local balance
- Further expand to control differences in local balance
- \sim Tuning parameter: Global parameter $\lambda \Rightarrow \lambda_g = \lambda/n_g$

Dual perspective: M estimation of treatment odds

Dual when optimizing for for local balance only

Population:
$$\underbrace{\frac{e(x,g)}{1-e(x,g)}}_{\text{inverse prop. score weights}} \sim \underbrace{\alpha_g + \beta_g' \phi(x)}_{\text{balancing weights}}$$

Dual perspective: M estimation of treatment odds

Dual when optimizing for for local balance only

Population:
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Sample:
$$\min_{\alpha_g, \beta_g}$$
 regression loss $+$ $\underbrace{\frac{\lambda}{2} \|\beta_g\|_2^2}_{\text{ridge penalty}}$

Global balance constraint ←→ partial pooling in the dual problem

Dual for Partially Pooled Balancing Weights

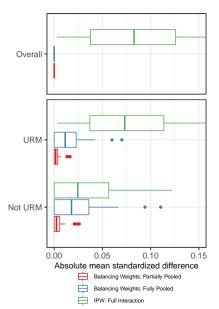
$$\min_{\alpha_g,\beta_g,\mu_{\beta}} \text{ regression loss } + \underbrace{\frac{\lambda_g}{2} \|\beta_g - \mu_{\beta}\|_2^2}_{\text{local} \rightarrow \text{global}}$$

Partially pool $local \rightarrow global$ model: regularization directly related to imbalance

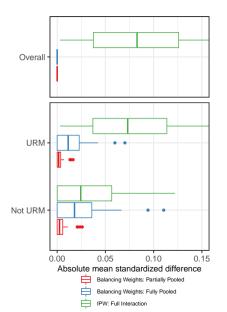
Differential impacts of letters of

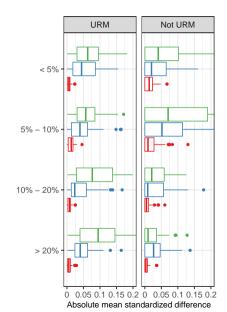
recommendation

Partially pooled balancing weights \rightarrow improved balance



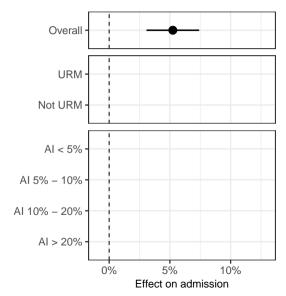
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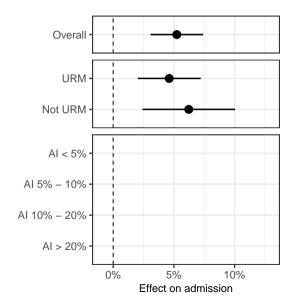


Large overall effect

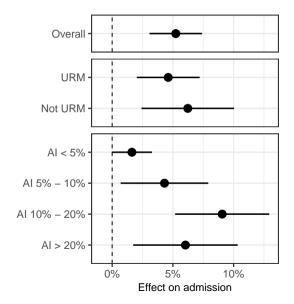
Baseline admissions rate ∼20%



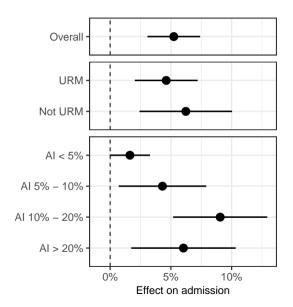
No differences by URM status

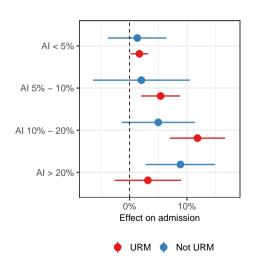


Large differences by Admissibility Index



Relative effect sizes flip





Recap: Varying Impacts of Letters of Recommendation

Partially pooled balancing weights

- Find weights that control both Local Balance and Global Balance
- Dual relation to partially pooled IPW
- R package balancer
- ightarrow No evidence of different impacts by URM status

Recap: Varying Impacts of Letters of Recommendation

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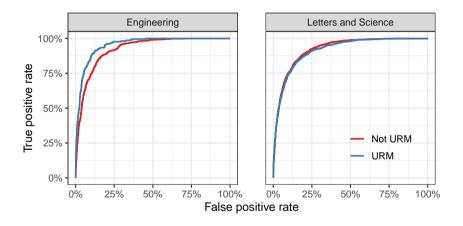
Thank you!

ebenmichael.github.io



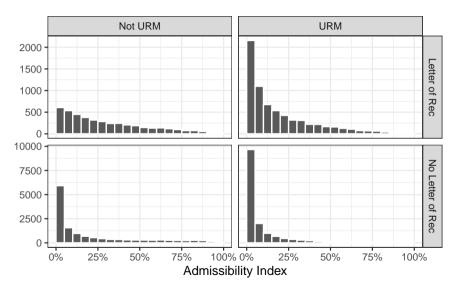
Appendix

Admissibility Index: a strong and simple prognostic

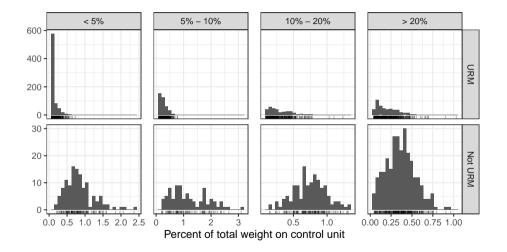


Use logisitic regression on 2015 applicant pool to predict admission for 2016 pool

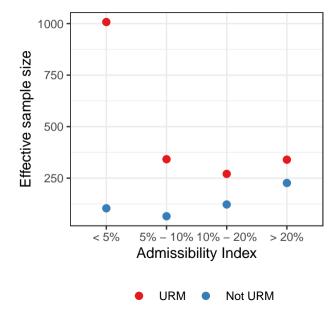
Heterogeneity across admissibility index



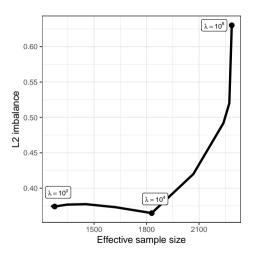
Distribution of the estimated weights



Effective sample sizes

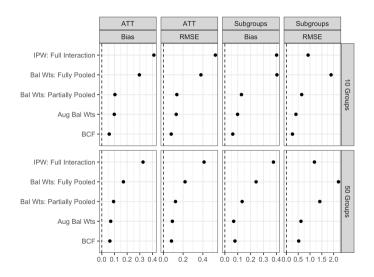


Hyperparameter tuning



- Evaluate across a range of λ
- Gains in precision, comparable imbalance

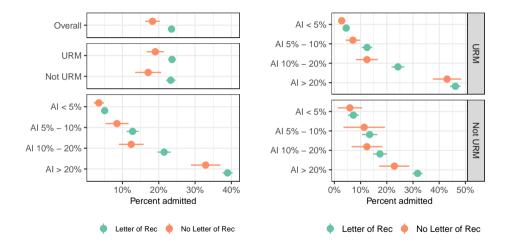
Obligatory Simulation Study



 Major gains relative to traditional IPW

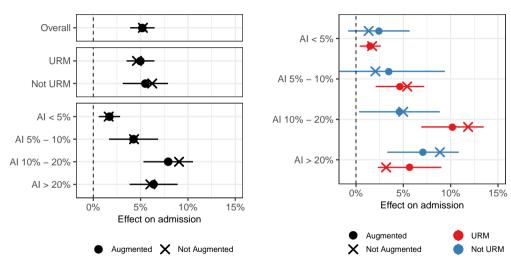
 Comparable performance to ML methods; retain design-based advantages

Estimated group means



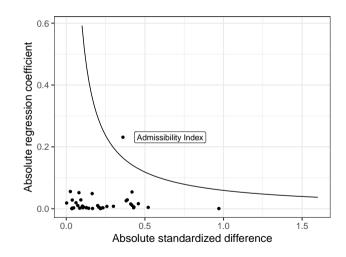
Augmentation diminishes differences

[Random forest outcome model]



Sensitivity to unmeasured confounding

- Adapt Soriano et al. [2021]
- Overall LOR effect still positive with $\Lambda=1.1$
- Consistent with wide range of subgroup estimates



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