# Variational Approach for Bayesian Density Regression

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## Introduction and Setup

### **Problem Setup**

- Given data  $(y_n, x_n)_{n=1}^N$  and we want to estimate the density of  $y \mid x$
- Gaussian mixture models are a common choice due to flexibility

$$f(y \mid x) = \sum_{k} \pi_{k} \mathcal{N}(y \mid \mu_{k}(x), \tau_{k}^{-1})$$

• Covariate-independent weights have limited flexibility in practice

## Introduction and Setup

### **Covariate Dependent Weights**

Allow for covariate dependence in the mixing weights

$$f(y \mid x) = \sum_{k} \pi_{k}(x) \mathcal{N}(y \mid \mu_{k}(x), \tau_{k}^{-1})$$

- Existing methods that use kernel stick breaking process (Dunson and Park, 2008), logit stick breaking prior (Rigon and Durante, 2017)
- We allow covariates to enter the weights via a logit link function

$$\pi_k(x) \propto \exp\{x^{\mathsf{T}}\gamma_k\}$$

### Notation

#### **Notation**

- Observed Data:  $\mathbf{y} = \{y_1, \dots, y_N\}, \mathbf{X} = \{x_1, \dots, x_N\} \subseteq \mathbb{R}^D$
- Guassian Componenets:  $\beta = \{\beta_1, \dots, \beta_K\}, \tau = \{\tau_1, \dots, \tau_k\}$

$$f(y \mid x) = \sum_{k} \pi_{k}(x) \, \mathcal{N}(y \mid x^{\mathsf{T}} \beta_{k}, \tau_{k}^{-1})$$

- coefficient vectors in the weights:  $\gamma = \{\gamma_1, \dots, \gamma_k\}$
- Latent variables:  $\mathbf{Z} = \{z_1, \dots, z_N\}$ 
  - ullet  $z_n \in \mathbb{R}^K$  and  $z_{nk} = 1$  iff  $y_n$  belongs to the k-th cluster

## **Prior Specification**

With the introduction of latent variables, we have the simplified conditional distribution,

$$p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\tau}, \mathbf{Z}) = \prod_{n} \prod_{k} \mathcal{N} \left( y_n \mid \mathbf{x}_n^{\mathsf{T}} \boldsymbol{\beta}_k, \tau_k^{-1} \right)^{\mathbf{z}_{nk}}$$
(1)

The conditional distribution of **Z** given  $X, \gamma$ 

$$p(\mathbf{Z} \mid \mathbf{X}, \gamma) = \prod_{n} \prod_{k} \pi_{k}(x_{n})^{z_{nk}} = \prod_{n} \prod_{k} \left( \frac{\exp\{x_{n}^{\mathsf{T}} \gamma_{k}\}}{\sum_{j=1}^{K} \exp\{x_{n}^{\mathsf{T}} \gamma_{j}\}} \right)^{z_{nk}}$$
(2)

Priors over  $\beta, \tau, \gamma$ 

- $p(\gamma) = \prod_k p(\gamma_k) = \mathcal{N}(\gamma_k \mid 0, I_D)$
- $p(\beta, \tau) = p(\beta \mid \tau) p(\tau) = \prod_k \mathcal{N}(\beta_k \mid m_0, (\tau_k \Lambda_0)^{-1}) \operatorname{Ga}(\tau_k \mid a_0, b_0)$

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# General Approach

### Full joint distribution of all random variables

$$p\left(\mathbf{y},\mathbf{X},\boldsymbol{\beta},\boldsymbol{\tau},\mathbf{Z}\right)=p\left(\mathbf{y}\mid\mathbf{X},\boldsymbol{\beta},\boldsymbol{\tau},\mathbf{Z}\right)p\left(\mathbf{Z}\mid\mathbf{X},\boldsymbol{\gamma}\right)p\left(\boldsymbol{\gamma}\right)p\left(\boldsymbol{\beta},\boldsymbol{\tau}\right)$$

### Formulate an approximating distribution

$$q\left(\mathsf{Z},eta,oldsymbol{ au},oldsymbol{\gamma}
ight)=q(\mathsf{Z})q(eta,oldsymbol{ au},oldsymbol{\gamma})$$

### Derive update equations for the variational distributions

$$\begin{split} & \ln q^*(\mathbf{Z}) = \mathbb{E}_{-q(\mathbf{Z})} \Big[ \ln \big\{ p\left(\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\tau}, \mathbf{Z}, \boldsymbol{\gamma}\right) \big\} \Big] \\ & \ln q^*(\boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma}) = \mathbb{E}_{-q(\boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma})} \Big[ \ln \big\{ p\left(\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\tau}, \mathbf{Z}, \boldsymbol{\gamma}\right) \big\} \Big] \end{split}$$

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## Variational Approximation

#### Joint Distribution of all random variables

$$\begin{split} & \ln p \left( \mathbf{y}, \mathbf{X}, \beta, \tau, \mathbf{Z} \right) = \sum_{n} \sum_{k} z_{nk} \bigg\{ -\frac{1}{2} \ln (2\pi) + \frac{1}{2} \ln \tau_{k} - \frac{\tau_{k}}{2} \left( y_{n} - x_{n}^{\mathsf{T}} \beta_{k} \right)^{2} \bigg\} \\ & + \sum_{n} \sum_{k} z_{nk} \bigg\{ x_{n}^{\mathsf{T}} \gamma_{k} - \ln \left( \sum_{j=1}^{K} \exp\{x_{n}^{\mathsf{T}} \gamma_{j}\} \right) \bigg\} \\ & + \sum_{k} \bigg\{ -\frac{D}{2} \ln (2\pi) + \frac{D}{2} \ln \tau_{k} + \ln |\Lambda_{0}| - \frac{\tau_{k}}{2} (\beta_{k} - m_{0})^{\mathsf{T}} \Lambda_{0} (\beta_{k} - m_{0}) \bigg\} \\ & + \sum_{k} \bigg\{ (a_{0} - 1) \ln \tau_{k} - b_{0} \tau_{k} \bigg\} \end{split}$$

### **Approximating Distribution**

$$q(\mathbf{Z}, \beta, \tau, \gamma) = q(\mathbf{Z})q(\beta, \tau, \gamma) \tag{3}$$

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# Example calculation for $q(\mathbf{Z})$

$$\begin{split} \ln q^*(\mathbf{Z}) &= \sum_n \sum_k z_{nk} \Bigg\{ -\frac{1}{2} \ln(2\pi) + \frac{1}{2} E_{q(\tau)} [\ln \tau_k] \\ &- \frac{1}{2} E_{q(\beta,\tau)} [\tau_k (y_n - x_n^\mathsf{T} \beta_k)^2] \\ &+ x_n^\mathsf{T} E_{q(\gamma)} [\gamma_k] - E_{q(\gamma)} \Bigg[ \ln \Bigg( \sum_j \exp\{x_n^\mathsf{T} \gamma_j\} \Bigg) \Bigg] \Bigg\} \\ &= \sum_n \sum_k z_{nk} \ln \rho_{nk} \end{split}$$

Exponentiating and normalizing, we arrive at the optimal variational distribution. Then,  $\mathbb{E}[z_{nk}] = r_{nk}$ .

$$q^*(\mathbf{Z}) = \prod_{n} \prod_{k} r_{nk}^{z_{nk}}, \quad r_{nk} = \frac{\rho_{nk}}{\sum_{j} \rho_{nj}}$$
(4)

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# Example calculation for $q(\mathbf{Z})$

Consider the quantity  $\ln \rho_{nk}$ :

- $E_{q(\tau)}[\ln \tau_k]$
- $E_{q(\beta,\tau)}[\tau_k(y_n-x_n^{\mathsf{T}}\beta_k)^2]$
- $E_{q(\gamma)}[\gamma_k]$
- $E_{q(\gamma)} \left[ \ln \left( \sum_{j} \exp\{x_n^{\mathsf{T}} \gamma_j\} \right) \right]$

# Variational Updates

Our choice of conjugate families tells us that we will have the following form of the variational distributions, for k = 1, ..., K,

$$q^*(\gamma_k) = \mathcal{N}\left(\gamma_k \mid \mu_k, \mathcal{Q}_k^{-1}\right) \tag{5}$$

$$q^*(\beta_k \mid \tau_k) = \mathcal{N}\left(\beta_k \mid m_k, (\tau_k \mathbf{V}_k)^{-1}\right) \tag{6}$$

$$q^*(\tau_k) = \operatorname{Ga}(\tau_k \mid a_k, b_k) \tag{7}$$

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# Example calculation for $q(\mathbf{Z})$

Consider the quantity  $\ln \rho_{nk}$ :

• 
$$E_{q(\tau)}[\ln \tau_k] = \psi(a_k) - \psi(b_k)$$

• 
$$E_{q(\beta,\tau)}[\tau_k(y_n-x_n^{\mathsf{T}}\beta_k)^2] = \frac{a_k}{b_k}(y_n+m_k^{\mathsf{T}}x_n)^2 + x_n^{\mathsf{T}}V_k^{-1}x_n$$

• 
$$E_{q(\gamma)}[\gamma_k] = \mu_k$$

• 
$$E_{q(\gamma)} \left| \ln \left( \sum_{j} \exp\{x_n^{\mathsf{T}} \gamma_j\} \right) \right| = \dots \text{ oh no!}$$

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## Bouchard's Bound for Problematic Quantity

### Upper bound on the sum of exponentials

$$\sum_{j=1}^K \mathrm{e}^{t_j} \leq \prod_{j=1}^K (1+\mathrm{e}^{t_j})$$

### Tangential Bound (Jaakkola and Jordan, 1996)

For  $x \in \mathbb{R}, \alpha \in \mathbb{R}, \xi \in [0, \infty)$ 

$$\log(1 + e^x) \le \lambda(\xi)(x^2 - \xi^2) + \frac{x - \xi}{2} + \log(1 + e^\xi)$$

Taking log the first bound and setting  $t_j = x_n^T \gamma_j - \alpha_n$ , we have

$$\log \sum_{j=1}^{K} \exp\{x_n^{\mathsf{T}} \gamma_j\} \le \alpha_n + \sum_{j=1}^{K} \frac{x_n^{\mathsf{T}} \gamma_j - \alpha_n + \xi_{nj}}{2} + \lambda(\xi_{nj}) \left( (x_n^{\mathsf{T}} \gamma_j - \alpha_n)^2 - \xi_{nj}^2 \right) + \log \left( 1 + e^{\xi_{nj}} \right)$$

## Bouchard's Bound for Problematic Quantity

### Upper Bound for the problematic expectation

$$\begin{split} \mathrm{E}_{q(\gamma)} \Big[ \ln \left( \sum_{j}^{K} \exp\{x_{n}^{\mathsf{T}} \gamma_{j}\} \right) \Big] \\ & \leq \alpha_{n} + \sum_{j}^{K} \lambda(\xi_{nj}) \left( (x_{n}^{\mathsf{T}} \mu_{j} - \alpha_{k})^{2} - \xi_{nj}^{2} + x_{n}^{\mathsf{T}} \mathrm{Q}_{k}^{-1} x_{n} \right) + \log(1 + \mathrm{e}^{\xi_{nj}}) \\ & + \frac{1}{2} \left( x_{n}^{\mathsf{T}} \mu_{j} - \alpha_{n} + \xi_{nj} \right) \end{split}$$

### Two new variational parameters that need to be updated

The additional parameters introduced in the two upper bounds can be updated using the following equations

$$\xi_{nk} = \sqrt{\left(\mu_k^{\mathsf{T}} x_n - \alpha_n\right)^2 + x_n^{\mathsf{T}} Q_k^{-1} x_n} \quad \forall k$$

$$\alpha_{n} = \frac{\frac{1}{2} \left( \frac{K}{2} - 1 \right) + \sum_{j=1}^{K} \lambda \left( \xi_{nj} \right) \mu_{j}^{\mathsf{T}} x_{n}}{\sum_{j=1}^{K} \lambda \left( \xi_{nj} \right)}$$

# Variational Algorithm

#### Input

- Number of components, K
- **3** Prior mean, precision for coefficients vectors,  $\beta_{1:K}$
- **1** Prior shape, rate parameters for precision parameters,  $\tau_{1:K}$

### Output

A variational density,

$$q(\mathbf{Z}, \beta, \tau, \gamma) = q(\mathbf{Z})q(\beta, \tau, \gamma) = q(\mathbf{Z})\prod_{k}q(\beta_{k}, \tau_{k})q(\gamma_{k})$$

Fully specified by the variational parameters

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# Variational Algorithm

### **Algorithm**. CAVI for Conditional Density Estimation

```
while the ELBO has not converged do
      for n \in \{1, \ldots, N\} do
             for k \in \{1, ..., K\} do
                   Set r_{nk} \propto \exp \left\{ -\frac{1}{2} \ln(2\pi) + \frac{1}{2} \mathbb{E}[\ln \tau_k] + x_n^{\mathsf{T}} \mathbb{E}[\gamma_k] \right\}
                                     -rac{1}{2}\mathbb{E}[	au_k(y_n-x_n^\intercaleta_k)^2]-\mathbb{E}\Big[\ln\Big(\sum_{j=1}^K\exp\{x_n^\intercal\gamma_j\}\Big)\Big]\Big\}
             end
      end
      for n \in \{1, \ldots, N\} do
             for k \in \{1, ..., K\} do
                  Set \xi_{nk} \leftarrow \sqrt{(x_n^{\mathsf{T}}\mu_k - \alpha_n)^2 + x_n^{\mathsf{T}}Q_k^{-1}x_n}
             end
            Set \alpha_n \leftarrow \left[\frac{1}{2}\left(\frac{K}{2}-1\right) + \sum_k \lambda(\xi_{nk})\mu_k^{\mathsf{T}} x_n\right] / \left[\sum_k \lambda(\xi_{nk})\right]
      end
         /** Remaining Variational Updates on Next Slide **/
end
```

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# Variational Algorithm (cont.)

# Algorithm. CAVI for Conditional Density Estimation

while the ELBO has not converged do

```
for k \in \{1, ..., K\} do
      Q_k \leftarrow I_D + 2 \sum_{n} r_{nk} \lambda(\xi_{nk}) x_n x_n^{\mathsf{T}}
                                                                                           /* gamma_k cov */
      \eta_k \leftarrow \sum_n r_{nk} \left[ \frac{1}{2} + 2\lambda(\xi_{nk})\alpha_n \right] x_n
      \mu_k \leftarrow Q_k^{-1} \eta_k
                                                                                            /* gamma_k mean */
     V_k \leftarrow \sum_n r_{nk} x_n x_n^{\mathsf{T}} + \Lambda_0
                                                                                            /* beta k cov */
      \zeta_k \leftarrow \sum_n r_{nk} y_n x_n + \Lambda_0 m_0
      m_k \leftarrow V_k^{-1} \zeta_k
                                                                                            /* beta_k mean */
      a_k \leftarrow a_0 + N_k
                                                                                            /* tau_k shape */
      b_k \leftarrow b_0 + \frac{1}{2} \left[ \sum_n r_{nk} y_n^2 + m_0^{\mathsf{T}} \Lambda_0 m_0 - \zeta_k^{\mathsf{T}} V_k^{-1} \zeta_k \right] / * \text{ tau_k rate } */
end
```

Compute ELBO using updated parameters

end return  $q(\mathbf{Z}, \boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma})$ 

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# Monitoring Convergence

At the end of each iteration, we compute the ELBO:

$$\mathcal{L}(q) = \sum_{\mathbf{z}} \int \int \int q(\boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma}, \mathbf{Z}) \ln \Big\{ \frac{p(\mathbf{y}, \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma}, \mathbf{Z})}{q(\boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma}, \mathbf{Z})} \Big\} d\boldsymbol{\beta} d\boldsymbol{\tau} d\boldsymbol{\gamma}$$

#### 7 quantities to calculate:

- $\mathbb{E}[\ln p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\tau}, \mathbf{Z})]$
- $\mathbb{E}[\ln p(\mathbf{Z} \mid \mathbf{X}, \boldsymbol{\gamma})]$
- $\mathbb{E}[\ln p(\gamma)]$
- $\mathbb{E}[\ln p(\beta, \tau)]$
- E[In q(Z)]
- $\bullet$   $\mathbb{E}[\ln q(oldsymbol{eta},oldsymbol{ au})]$
- ullet  $\mathbb{E}[\ln q(\gamma)]$



### References

Please see report submission for updated set of references.

# Thank you!