Variational Inference for Bayesian Density Regression.

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1. Introduction

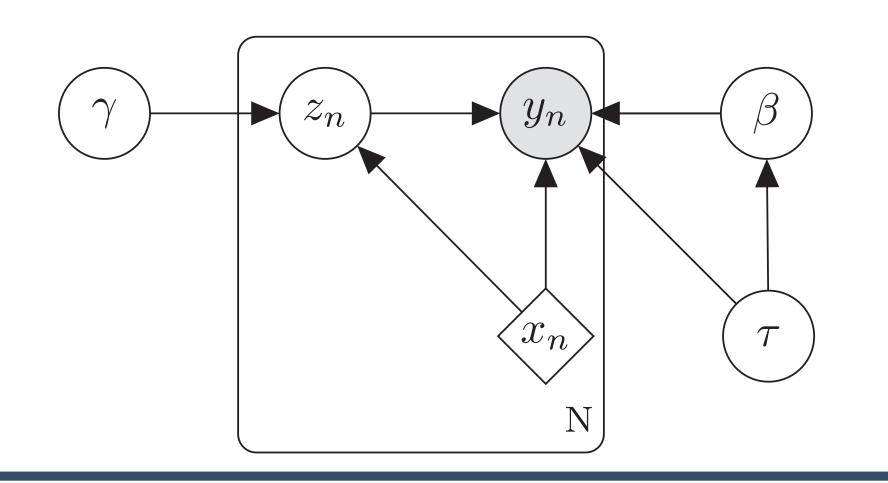
In the Bayesian density regression problem, we observe data (y_n, x_n) , n = 1, ..., N, and the goal is the estimate the conditional density of $y \mid x$. We model the density using a mixture of Gaussians for which covariates enter the weights through a logit link function.

$$f(y \mid x) = \sum_{k}^{K} \pi_k(x) \cdot \mathcal{N}\left(y \mid \mu_k(x), \tau_k^{-1}\right)$$

where $\mu_k(x) = x^{\mathsf{T}}\beta_k$ and $\pi_k(x) \propto \exp(x^{\mathsf{T}}\gamma_k)$. While this increases the flexibility of the model, it also increases the computational complexity. In order to perform scalable inference on the model parameters, we propose a variational approach that uses a tangential approximation of the softmax function to achieve fast, closed form updates for the coordinate ascent algorithm.

2. Notation

- Data: $\mathbf{y} = \{y_{1:N}\}, \mathbf{X} = \{x_{1:N}\} \subseteq \mathbb{R}^D$
- Coefficients: $\boldsymbol{\beta} = \{\beta_{1:K}\}, \boldsymbol{\gamma} = \{\gamma_{1:K}\}$
- Precision (Gaussian): $\tau = \{\tau_{1:K}\}$
- Cluster Indicator: $\mathbf{Z} = \{z_{1:N}\} \subseteq \mathbb{R}^K$



3. Model Setup

We use conjugate priors to ease computation.

$$p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\beta}, \boldsymbol{\tau}, \mathbf{Z}) = \prod_{n} \prod_{k} \mathcal{N} \left(y_{n} \mid x_{n}^{\mathsf{T}} \beta_{k}, \tau_{k}^{-1} \right)^{z_{nk}}$$

$$p(\mathbf{Z} \mid \mathbf{X}, \boldsymbol{\gamma}) = \prod_{n} \prod_{k} \left[\frac{e^{x_{n}^{\mathsf{T}} \gamma_{k}}}{\sum_{j=1}^{K} e^{x_{n}^{\mathsf{T}} \gamma_{j}}} \right]^{z_{nk}}$$

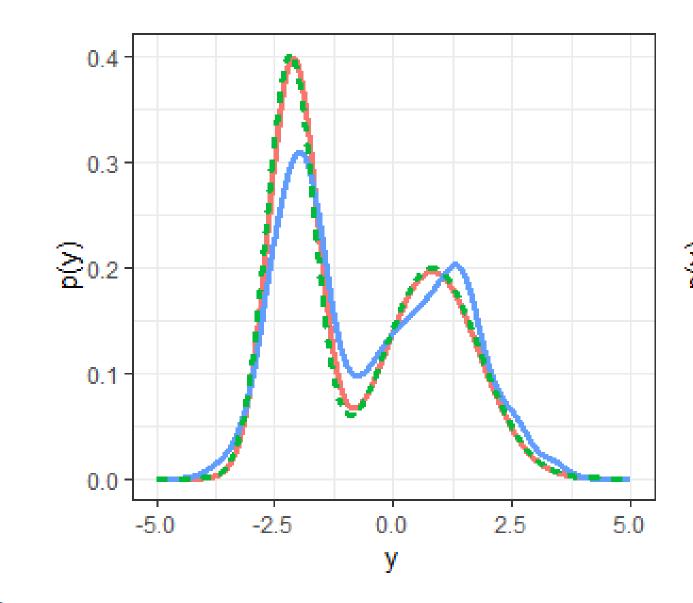
$$p(\boldsymbol{\gamma}) = \prod \mathcal{N} \left(\gamma_{k} \mid 0, I_{D} \right)$$

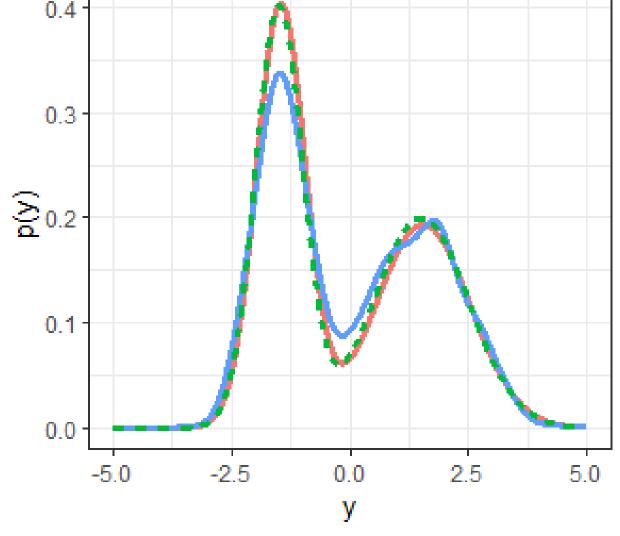
$$p(\boldsymbol{\beta}, \boldsymbol{\tau}) = \prod_{k}^{n} p(\beta_{k} | \tau_{k}) p(\tau_{k})$$

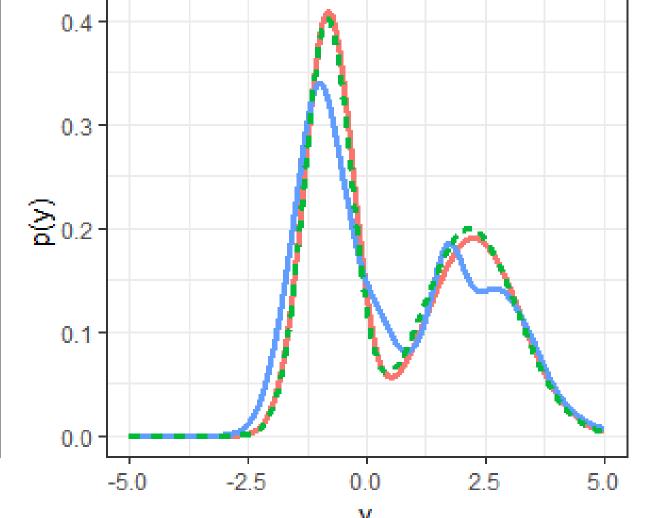
$$p(\beta_k | \tau_k) = \mathcal{N}\left(\beta_k | m_0, (\tau_k \Lambda_0)^{-1}\right)$$
$$p(\tau_k) = \text{Gamma}(\tau_k | a_0, b_0)$$

5. Application to Bimodal Conditional Densities

For the conditional density, $X \sim \mathcal{N}(0,1), Y \mid X \sim 0.5\mathcal{N}\left(X-1.5,0.5^2\right) + 0.5\mathcal{N}\left(X+1.5,1^2\right)$, we examine samples of size 1000 and look at the conditional density at the three quartiles of the predictor support: x = -0.6745 (left), x = 0 (center), x = 0.6475 (right). For a sample size of 1000, we plot the approximations below. The true density is a dashed green line, the variational approximation is in red, and the kernel density estimate is in blue.







4. Variational Approximation

We approximate the posterior distribution with:

$$q(\mathbf{Z}, \boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma}) = q(\mathbf{Z})q(\boldsymbol{\beta}, \boldsymbol{\tau}, \boldsymbol{\gamma})$$

Then the distribution for each of the variational parameters can be found by taking the expectation of the joint likelihood with respect the *other* variational parameters.

$$q^{\star}(z_{nk}) = r_{nk}^{z_{nk}}$$

$$q^{\star}(\gamma_k) = \mathcal{N}\left(\gamma_k \mid \mu_k, Q_k^{-1}\right)$$

$$q^{\star}(\beta_k \mid \tau_k) = \mathcal{N}\left(\beta_k \mid m_k, (\tau_k V_k)^{-1}\right)$$

$$q^{\star}(\tau_k) = \operatorname{Ga}\left(\tau_k \mid a_k, b_k\right)$$

An issue arises in calculating $q(z_{nk})$ because it requires computing:

$$arepsilon_n = \mathbb{E}_{q(oldsymbol{\gamma})} \left[\ln \sum_j \exp\{x_n^\intercal \gamma_j\}
ight]$$

which is not available in closed form. We resort to the following bound (Bouchard, 2007)

$$\varepsilon_n \le \alpha_n + \sum_{j=1}^{K} \frac{1}{2} \left(x_n^{\mathsf{T}} \mu_j - \alpha_n - \xi_{nj} \right) + \log(1 + e^{\xi_{nj}})$$

Two additional variational parameters:

$$\xi_{nj} = \sqrt{(\mu_{j}^{\mathsf{T}} x_{n} - \alpha_{n})^{2} + x_{n}^{\mathsf{T}} Q_{j}^{-1} x_{n}}$$

$$\alpha_{n} = \frac{\frac{1}{2} (\frac{K}{2} - 1) + \sum_{j=1}^{K} \lambda (\xi_{nj}) \mu_{j}^{\mathsf{T}} x_{n}}{\sum_{j=1}^{K} \lambda (\xi_{nj})}$$

$$\lambda(\xi) = \frac{1}{4\xi} \tanh\left(\frac{\xi}{2}\right)$$

6. Application to Speedflow Data

We consider the following bimodal conditional density, $X \sim \mathcal{N}(0,1)$, $Y \mid X \sim 0.5 \mathcal{N}\left(X-1.5,0.5^2\right) + 0.5 \mathcal{N}\left(X+1.5,0.5^2\right)$ In particular, we look at the conditional density at the three quartiles of the predictor support: x = -0.6745 (left), x = 0 (center), x = 0.6475 (right). For a sample size of 1000, we plot the approximations below, where the true conditional density is a dashed green line, the variational approximation is shown in red, and the kernel density estimate is shown in blue.

