The Influence of Individual Characterisitcs on Public Transportation Planning*

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Abstract

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```
## Warning: package 'tidyverse' was built under R version 3.5.3
## Warning: package 'ggplot2' was built under R version 3.5.3
## Warning: package 'tibble' was built under R version 3.5.3
## Warning: package 'tidyr' was built under R version 3.5.3
## Warning: package 'readr' was built under R version 3.5.3
## Warning: package 'purrr' was built under R version 3.5.3
## Warning: package 'dplyr' was built under R version 3.5.3
## Warning: package 'stringr' was built under R version 3.5.3
## Warning: package 'forcats' was built under R version 3.5.3
## Warning: package 'lubridate' was built under R version 3.5.3
## Warning: package 'stargazer' was built under R version 3.5.3
```

1 Literature Review

Allen et al. (2016) study the reasoning of the failure of a referendum on a congestion charging scheme in Edinburgh. Instead of using direct voting data, they conduct a survey after the referendum, which allows them to ask more specific questions. Researchers can gain detailed data by surveying, because the unit of measurement is each individual; however, a possible disadvantage of surveying is that respondents who turn in the questionnaire tend to have stronger attitudes towards the proposal, generating sampling bias. They conclude that people who use cars as the primary transportation mean, demonstrate a misconception of the pricing plan, or question the effectiveness of the scheme at reducing congestion are more likely to oppose it. Their findings can give insights to the similar failure in the Gwinnett referendum. Voters against the proposal could be those who rarely use public transportation and those who are not convinced by the effectiveness of expanding public transit in alleviating the traffic.

Another crucial factor is the accessibility of the proposed transit system. Kinsey et al. (2010) examine the relationship between the distance to the scheduled railway station and voter turnout by studying the Seattle monorail referendum. They introduce the concept of diffused and concentrated benefit/cost. People who live far from the monorail enjoy the diffused benefit of less traffic congestion, and bear the diffused cost of increased tax. People living close to the rail experience the same diffused benefit and cost, but they also gain the concentrated benefit of easily accessing the public good. Finally, those who live very close to the railway have the same benefits and costs, but they also face the concentrated cost such as inconvenience during construction. Since "people are more strongly motivated to avoid losses than to approach gains," they expect a higher turnout rate in farther places with votes for "no," which is verified from their analyses. Besides distance, they also find out precincts with a higher percentage of people of lower socioeconomic status or young people have a lower turnout rate. Interestingly, there is a significant interaction between partisanship and distance, which would be also tested in my study. In essence, the effect of distance on turnout is weakened by partisanship, and vanishes beyond a threshold of distance. Even though my dependent variable is voters' responses rather than turnout, it can be inferred from Kinsey et al.'s findings that people farther away from the transit system would vote against the referendum more. However, the relationship might be non-linear and requires some form of transformation. Regarding the methods, they utilize the spatial lag model to correct for autocorrelation, which is proper to use in my project as well since both studies use precinct-level data.

Table 1: Variable definitions

Variable name	Description			
GEOID	The geographic identifier of the census tract			
medage	The median age of the population in the tract			
medincome	The median income of the population in the tract			
$white_pct$	The percentage of white population in the tract			
$\operatorname{public} \operatorname{_pct}$	The percentage of people who go to work by public transportation (excluding taxi or cab)			
$time_pct$	The percentage of people who travel more than an hour to work			
$trump_pct$	The estimated percentage of votes for Donald Trump in that tract			
$voter_turnout$	The estimated percentage of voters who voted in this referendum in the tract			
yes_pct	The estimated percentage of voters who voted yes in this referendum in the tract			
plan_yes	Whether the tract is covered by short-range plan. 1 stands for yes, 0 stands for no			
_current_yes	Whether the tract is covered by the existing public transportation. 1 stands for yes, 0 stands for no			

2 Background

current transportation future plan referendum

3 Data & Methods

3.1 Conceptual model

3.2 Data

- **GEOID:** The geographic identifier of the census tract.
- **medage:** The median age of the population in the tract.
- medincome: The median income of the population in the tract.
- white_pct: The percentage of white population in the tract.
- public_pct: The percentage of people who go to work by public transportation (excluding taxi or cab).
- time_pct: The percentage of people who travel more than an hour to work.
- trump_pct: The estimated percentage of votes for Donald Trump in that tract.
- voter_turnout: The estimated percentage of voters who voted in this referendum in the tract.
- yes_pct: The estimated percentage of voters who voted yes in this referendum in the tract.

- plan_yes: Whether the tract is covered by the public transportation planned in the short-range (Y2020 2025), defined by whether any transportation is available within 500 meters. 1 stands for yes, 0 stands for no.
- **current_yes:** Whether the tract is covered by the existing public transportation, defined by whether any transportation is available within 500 meters. 1 stands for yes, 0 stands for no.

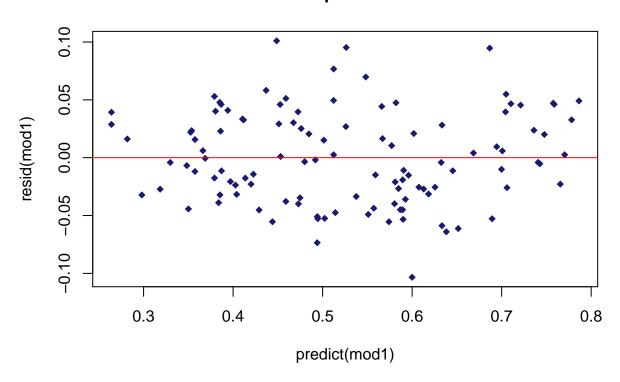
4 Results

model 1: all variables, no interaction, linear

```
load("processed_data/analysis_data.RData")
mod1 <- lm(data = final_data, yes_pct ~ medage + medincome + white_pct + public_pct + time_pct + trump_
summary(mod1)
Call: lm(formula = yes pct ~ medage + medincome + white pct + public pct + time pct + trump pct
+ voter_turnout + plan_yes + current_yes, data = final_data)
Residuals: Min 1Q Median 3Q Max -0.103384 -0.031723 -0.004037 0.030238 0.101018
Coefficients: Estimate Std. Error t value Pr(>|t|)
(\text{Intercept}) \ 8.016 \text{e-} 01 \ 4.677 \text{e-} 02 \ 17.138 < 2 \text{e-} 16 \quad \textit{medage} \ \textit{2.733e-} 03 \ \textit{1.477e-} 03 \ \textit{1.851} \ \textit{0.067054} \ .
medincome 1.458e-07 2.914e-07 0.500 0.617850
white_pct 1.177e-01 5.285e-02 2.226 0.028162
public_pct 8.393e-01 3.355e-01 2.501 0.013946 *
time\_pct -3.029e-01 9.013e-02 -3.361 0.001090 trump\_pct -8.884e-01 5.680e-02 -15.640 < 2e-16
voter\_turnout -3.285e-01 1.257e-01 -2.613 0.010323
plan_yes1 -2.967e-02 1.153e-02 -2.573 0.011511 *
current_yes1 4.627e-02 1.209e-02 3.825 0.000224 * — Signif. codes: 0 '' 0.001 '' 0.01 " 0.05 '' 0.1
, , 1
Residual standard error: 0.04185 on 103 degrees of freedom Multiple R-squared: 0.917, Adjusted R-squared:
0.9098 F-statistic: 126.5 on 9 and 103 DF, p-value: < 2.2e-16
model 1 assumption checking
```

plot(predict(mod1),resid(mod1),col="midnightblue",pch=18,main="Residual plot - Model 1")
abline(0,0,col="red")

Residual plot - Model 1



collinearity:

library(car)

```
## Warning: package 'car' was built under R version 3.5.3

## Loading required package: carData

## Warning: package 'carData' was built under R version 3.5.3

##

## Attaching package: 'car'

## The following object is masked from 'package:dplyr':

##

## recode
```

```
\mbox{\tt \#\#} The following object is masked from 'package:purrr':
```

##

some

vif(mod1)

time_pct	<pre>public_pct</pre>	white_pct	medincome	medage
1.559837	1.151309	3.863617	3.221504	2.920936
	current_yes	plan_yes	voter_turnout	trump_pct
	2.350032	1.741403	3.368751	4.512575

all below 5: good, no collinearity problem

model 2: all variables, no interaction, logistic

```
mod2 <- glm(data = final_data, yes_pct ~ medage + medincome + white_pct + public_pct + time_pct + trump</pre>
```

Warning in eval(family\$initialize): non-integer #successes in a binomial glm!

summary(mod2)

Call: glm(formula = yes_pct ~ medage + medincome + white_pct + public_pct + time_pct + trump_pct + voter_turnout + plan_yes + current_yes, family = "binomial", data = final_data)

Deviance Residuals: Min 1Q Median 3Q Max

 $\hbox{-0.224114} \hbox{ -0.065598} \hbox{ -0.003135} \hbox{ 0.067243} \hbox{ 0.206940}$

Coefficients: Estimate Std. Error z value $\Pr(>|z|)$ (Intercept) 1.315e+00 2.311e+00 0.569 0.569 medage 1.068e-02 7.223e-02 0.148 0.882 medincome 6.880e-07 1.432e-05 0.048 0.962 white_pct 5.666e-01 2.627e+00 0.216 0.829 public_pct 3.642e+00 1.706e+01 0.213 0.831 time_pct -1.376e+00 4.462e+00 -0.308 0.758 trump_pct -3.790e+00 2.839e+00 -1.335 0.182 voter_turnout -1.453e+00 6.216e+00 -0.234 0.815 plan_yes1 -1.276e-01 5.671e-01 -0.225 0.822 current_yes1 1.891e-01 5.911e-01 0.320 0.749

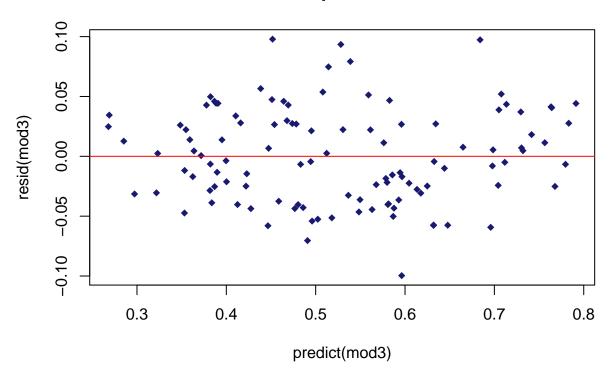
(Dispersion parameter for binomial family taken to be 1)

Null deviance: 9.05063 on 112 degrees of freedom

```
Residual deviance: 0.83282 on 103 degrees of freedom AIC: 133.92
Number of Fisher Scoring iterations: 4
model 3: all variables, no interaction, some transformation, linear
step 1: find the skewed variables
library(dlookr)
## Warning: package 'dlookr' was built under R version 3.5.3
## Loading required package: mice
## Warning: package 'mice' was built under R version 3.5.3
##
## Attaching package: 'mice'
## The following objects are masked from 'package:base':
##
       cbind, rbind
##
##
## Attaching package: 'dlookr'
## The following object is masked from 'package:base':
##
##
       transform
find_skewness(final_data)
[1] 3 5 8
medincome, public_pct, voter_turnout
step 2: transform them
```

```
data_tf <- final_data %>%
   mutate(medincome = log(medincome),
            public_pct = log(public_pct + 0.01),
            voter_turnout = (voter_turnout)^0.5)
find_skewness(data_tf)
integer(0)
step 3: model them
mod3 <- lm(data = data_tf, yes_pct ~ medage + medincome + white_pct + public_pct + time_pct + trump_pct
summary(mod3)
Call: lm(formula = yes_pct \sim medage + medincome + white_pct + public_pct + time_pct + trump_pct
+ voter_turnout + plan_yes + current_yes, data = data_tf)
Residuals: Min 1Q Median 3Q Max -0.099662 -0.030760 -0.004331 0.027626 0.097867
Coefficients: Estimate Std. Error t value Pr(>|t|)
(Intercept) 1.022033 0.217430 4.701 8.07e-06 medage 0.003595 0.001506 2.386 0.01884
\mathbf{medincome} \ \textbf{-0.007268} \ \ \textbf{0.022156} \ \ \textbf{-0.328} \ \ \textbf{0.74356}
white_pct 0.119648 0.050332 2.377 0.01929 *
public_pct 0.022875 0.008179 2.797 0.00616 time pct -0.279306 0.089340 -3.126 0.00230 **
 \text{trump\_pct -} 0.859270 \ 0.058269 \ \text{-} 14.747 < 2\text{e-} 16 \quad \textit{voter\_turnout -} \textbf{0.321784} \ \textbf{0.103026 -} \textbf{3.123} \ \textbf{0.00232} 
plan_yes1 -0.031542 0.011317 -2.787 0.00633 current_yes1 0.045541 0.011848 3.844 0.00021 —
Signif. codes: 0 '' 0.001 '' 0.01 '' 0.05 '.' 0.1 '' 1
Residual standard error: 0.04112 on 103 degrees of freedom Multiple R-squared: 0.9199, Adjusted R-squared:
0.9129 \text{ F-statistic: } 131.4 \text{ on } 9 \text{ and } 103 \text{ DF, p-value: } < 2.2\text{e-}16
model 3 assumption checking
plot(predict(mod3),resid(mod3),col="midnightblue",pch=18,main="Residual plot - Model 3")
abline(0,0,col="red")
```

Residual plot - Model 3



collinearity:

vif(mod3)

time_pct	<pre>public_pct</pre>	white_pct	medincome	medage
1.587723	1.133155	3.629803	3.713809	3.149466
	current_yes	plan_yes	voter_turnout	trump_pct
	2.336190	1.737295	3.621155	4.919508

all below 5: no collinearity

model 4: all variables, interaction, no transformation, linear

```
mod4 <- lm(data = final_data, yes_pct ~ medage + medincome + white_pct + public_pct + time_pct + trump_summary(mod4)</pre>
```

Call: $lm(formula = yes_pct \sim medage + medincome + white_pct + public_pct + time_pct + trump_pct + voter_turnout + plan_yes * current_yes, data = final_data)$

Residuals: Min 1Q Median 3Q Max -0.103714 -0.032077 -0.002399 0.030226 0.101740

Coefficients: Estimate Std. Error t value Pr(>|t|)

 $(\text{Intercept}) \ 8.041 \text{e-} 01 \ 4.707 \text{e-} 02 \ 17.085 < 2 \text{e-} 16 \quad \textit{medage 2.738e-} 03 \ \textit{1.481e-} 03 \ \textit{1.849 0.06739} \ .$

medincome 1.326e-07 2.930e-07 0.453 0.65183

white_pct 1.143e-01 5.326e-02 2.147 0.03420

public_pct 8.173e-01 3.382e-01 2.417 0.01743 *

— Signif. codes: 0 '' **0.001** '' 0.01 " 0.05 '' 0.1 ' '1

Residual standard error: 0.04197 on 102 degrees of freedom Multiple R-squared: 0.9174, Adjusted R-squared: 0.9093 F-statistic: 113.2 on 10 and 102 DF, p-value: < 2.2e-16

anova(mod1,mod4)

Analysis of Variance Table

Model 1: yes_pct \sim medage + medincome + white_pct + public_pct + time_pct + trump_pct + voter_turnout + plan_yes + current_yes Model 2: yes_pct \sim medage + medincome + white_pct + public_pct + time_pct + trump_pct + voter_turnout + plan_yes * current_yes Res.Df RSS Df Sum of Sq F Pr(>F) 1 103 0.18044

 $2\ 102\ 0.17970\ 1\ 0.00073903\ 0.4195\ 0.5186$

Model w/ interaction doesn't differ significantly from the one w/o interaction.