

Innovation Diffusion in Heterogeneous Populations: Contagion, Social Influence, and Social Learning

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New ideas, products, and practices take time to diffuse, a fact that is often attributed to some form of heterogeneity among potential adopters. This paper examines three broad classes of diffusion models—contagion, social influence, and social learning—and shows how to incorporate heterogeneity into each at a high level of generality without losing analytical tractability. Each type of model leaves a characteristic “footprint” on the shape of the adoption curve which provides a basis for discriminating empirically between them. The approach is illustrated using the classic study of Ryan and Gross (1943) on the diffusion of hybrid corn. (JEL D83, O33, Q16, Z13)

I. Models of Innovation Diffusion

A basic puzzle posed by innovation diffusion is why long lags occur between an innovation's first appearance and its general acceptance by a population. Among the factors that have been suggested are delays in acting on information, a desire to conform, learning from others, and changes in external factors such as prices. There is an extensive literature on this phenomenon that spans economics, sociology, and marketing.¹ Although the models are highly varied, a common theme is that delay results in part from heterogeneity among potential adopters. People may realize different benefits and costs from the innovation, hear about it at different times, or have different amounts of information, different predispositions to conform, and so forth. Nevertheless, most of the extant models incorporate heterogeneity in a very restricted fashion, say by considering two homogeneous classes of agents, or by assuming that the heterogeneity is described by a particular class of distributions.²

The purpose of this paper is threefold. First, it will be shown that the benchmark models in economics, marketing, and sociology proceed from fundamentally different assumptions that lead to key differences in their dynamical structure. Second, each type of model can be formulated at a high level of generality that allows for essentially any distribution of heterogeneous characteristics without losing analytical tractability. Third, each model leaves a distinctive “footprint” in its pattern of acceleration that holds with few or no restrictions on the distribution of

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¹ For general overviews, see Vijay Mahajan and Robert A. Peterson (1985); Mahajan, Eitan Muller, and Frank M. Bass (1990); Paul A. Geroski (2000); Paul Stoneman (2002); Everett M. Rogers (2003); and Thomas W. Valente (1995, 1996, 2005).

² See, in particular, Abel Jeuland (1981); Richard Jensen (1982); Massoud Karshenas and Stoneman (1992); and Geroski (2000).

characteristics. The reason is that they have fundamentally different structures that details in the underlying distributions cannot overcome.

We shall consider three broad classes of models—contagion, social influence, and social learning—that are drawn from the literature on marketing, sociology, and economics, respectively.³ The focus of attention will be on situations where the dynamics are driven from within; that is, there are internal feedback effects from prior to future adopters. Diffusion models that are driven mainly by *external* factors, such as price and quality changes, obviously depend on the exogenous rates of change in these factors, so much less can be said about them from an *a priori* standpoint.⁴

1. *Contagion*. People adopt when they come in contact with others who have already adopted; that is, innovations spread much like epidemics.
2. *Social influence*. People adopt when enough other people in the group have adopted; that is, innovations spread by a conformity motive.
3. *Social learning*. People adopt once they see enough empirical evidence to convince them that the innovation is worth adopting, where the evidence is generated by the outcomes among prior adopters. Individuals may adopt at different times due to differences in their prior beliefs, amount of information gathered, and idiosyncratic costs.

Of the three categories of models, social learning is certainly the most plausible from an economic standpoint, because it has firm decision-theoretic foundations: agents are assumed to make rational use of information generated by prior adopters in order to reach a decision.⁵ By contrast, the contagion and social influence models are based on the notion of *exposure* rather than on utility maximization. Why, then, should economists be concerned with them? One reason is that they are the most prominent models in the literature outside of economics, so it is worth understanding how they work in a general setting. A second reason is that we would like to know whether adoption curves generated by social learning differ from those generated purely by exposure, that is, whether we can distinguish empirically between the two explanations.

We shall find that under certain conditions this is indeed the case. To make progress on this question, however, we need to formulate the aggregate dynamics of each class of model at a high level of generality without making restrictive assumptions about the distribution of individual characteristics. This is particularly challenging in the case of social learning models, where there are multiple sources of heterogeneity, including different costs of adoption, different prior beliefs, and different amounts of information-gathering. It turns out that all of these considerations can be folded together in a simple dynamical equation that is easy to analyze.

³ Some of these ideas have also been applied in anthropology; see, in particular, Robert Boyd and Peter J. Richerson (1985) and Joseph Henrich (2001).

⁴ These are sometimes called “moving equilibrium” models; see, in particular, David (1968, 1969, 1975, 2005); David and Foray (1984, 1986); and Stoneman (2002). There are also hybrid models in which the dynamics are driven by a combination of external and internal factors. The most common example is the Bass model, which will be considered in Section III.

⁵ There is a large literature on social learning, but the specific assumptions are highly varied. See, among others, Sushil Bikhchandani, David Hirshleifer, and Ivo Welch (1992); Abhijit Banerjee (1992); Alan Kirman (1993); Glenn Ellison and Drew Fudenberg (1993, 1995); Sandeep Kapur (1995); Venkatesh Bala and Sanjeev Goyal (1998); Lones Smith and Peter Sorensen (2000); Kalyan Chatterjee and Susan H. Xu (2004); Banerjee and Fudenberg (2004); and Charles F. Manski (2004). However, relatively little prior work has been done on the implications of social learning for the *shape* of the adoption curve. Notable exceptions are Henrich (2001); Jensen (1982); and Dunia Lopez-Pintado and Duncan J. Watts (2008).

One limitation of the analysis is the use of a “mean-field” approach in which the population is assumed to be infinitely large and encounters between individuals are purely random. If the population is finite and agents interact through a fixed social network, the aggregate dynamics are substantially more complex and may depend on the topology of the network. The extent to which our results can be generalized to this case is an interesting question that will be left to future work.⁶

The main results on shape and acceleration patterns can be summarized as follows.

1. A process that is driven purely by inertia must decelerate from start to finish, that is, *the adoption curve is strictly concave*. Acceleration—in particular S-shaped curves—cannot be explained solely by delay no matter how the rates of inertia are distributed in the population.⁷
2. A process that is driven purely by contagion typically accelerates initially and then decelerates as it approaches saturation (the curve is S-shaped), but it cannot accelerate beyond the 50 percent adoption level. When everyone is alike, the hazard rate (the rate at which nonadopters become adopters) is linear and increasing in the current number of adopters. From this it follows that *the hazard rate must be nonincreasing relative to the number of adopters*. Less obviously, this result holds for any joint distribution of “infection” rates from a combination of internal and external sources. This forms a significant and testable implication of the contagion model, as we shall see in Section VI.
3. A process that is driven purely by social influence can either decelerate or accelerate initially, but in the latter case it typically accelerates at a *superexponential rate* for some period of time.
4. A process that is driven purely by learning from the experiences of others (social learning with direct observation) *usually begins slowly and may even decelerate initially*. If such a process does eventually accelerate, it typically does so at a *superexponential rate* for some period of time. Moreover, the overall rate at which adoption proceeds is a monotone increasing function of the expected payoffs from adoption.

These results provide a general framework for assessing the relative plausibility of different mechanisms that could be driving a given diffusion process. It is not an identification strategy, which is virtually impossible with aggregate-level data and difficult enough with micro-level data. Rather, our results establish restrictions on the shape of adoption curves that are associated with particular explanatory mechanisms. The overall plausibility of a given explanation should be examined in conjunction with other information about the specific nature of the process, e.g., whether it was driven largely by internal communication among members of a population or by external sources such as advertising campaigns; whether payoffs played a prominent role (as in

⁶ For models of diffusion in social networks, see Lawrence E. Blume (1993, 1995); Bala and Goyal (1998); Young (2003); Chatterjee and Xu (2004); Robin Cowan and Nicholas Jonard (2004); Matthew O. Jackson and Brian W. Rogers (2007); Jackson and Leeat Yariv (2007); Benjamin Golub and Jackson (2008); and Fernando Vega-Redondo (2007).

⁷ This feature of inertia models distinguishes them from contagion models, as was pointed out in the classic study of medical innovation by James S. Coleman, Elihu Katz, and Herbert Menzel (1957, 1966). It is also a well-known feature of heterogeneous duration models (Tony Lancaster and Stephen Nickell 1980; James J. Heckman and Burt Singer 1982; Heckman, Richard Robb, and James R. Walker 1990). We draw attention to it here because it shows in a particularly transparent way how restrictions on shape can hold irrespective of the distribution of individual characteristics (in this case the inertia rates).

new production methods); or conformity was the main motive (as in new fashions). We outline how this framework can be applied empirically using Bryce Ryan and Neal C. Gross's classic study of the diffusion of hybrid corn in the 1920s and 1930s (Ryan and Gross 1943). While the data in this study are rather limited, it illustrates how the approach can be brought to bear on actual cases, and also highlights the kinds of data that one would need to conduct a full-fledged empirical analysis.

II. Inertia

Before launching into a discussion of the three main models, it will be useful to consider a much simpler reason why innovations take time to diffuse; namely, people sometimes *delay* in acting on new information. This hypothesis leads to a particularly tractable model that has been studied in other contexts, notably heterogeneous duration models (see, among others, Lancaster and Nickell 1980; Heckman and Singer 1982; Heckman, Robb, and Walker 1990). We include it here because it forms a useful foil for the models to come later; furthermore, all of the other models include inertia as one of the parameters, so this is by far the simplest case to analyze.

First, consider a model of delayed adoption in which there is no heterogeneity among the agents. Let $\lambda > 0$ be the instantaneous rate at which any given nonadopter first adopts. Let $p(t)$ be the proportion of adopters at time t , and let us set the clock so that $p(0) = 0$. The function $p(t)$ is called the *adoption curve*. Assume for simplicity that once agents have adopted the new technology, they do not switch back to the old technology within the time frame of the analysis. Then the expected motion is described by the ordinary differential equation $\dot{p}(t) = \lambda(1 - p(t))$, which has the solution $p(t) = 1 - e^{-\lambda t}$ given the initial condition $p(0) = 0$.

Notice that this curve is *concave* throughout; in particular, it is not *S-shaped*. It is a rather remarkable fact that this remains true when *any amount of heterogeneity* is introduced. To see this, suppose that $\nu(\lambda)$ is the distribution of λ in the population. Then the expected trajectory of the process is given by

$$(1) \quad p(t) = 1 - \int e^{-\lambda t} d\nu.$$

Differentiating (1) twice over, we see that $\ddot{p}(t) < 0$ *irrespective of the distribution* $\nu(\lambda)$. The intuition is straightforward: agents with high values of λ tend to adopt earlier than those with low values of λ . It follows that the average value of λ in the current population of nonadopters is nonincreasing, and the number of such individuals is strictly decreasing. Thus the flow of new adopters is strictly decreasing.

This simple example illustrates the kinds of results that hold in more complex situations: the structure of the model has implications for the *shape* of the curve that remain true even when an arbitrary amount of heterogeneity is introduced.

III. Contagion

Contagion refers to a process in which people adopt a new product or practice when they come in contact with others who have adopted it. An everyday example would be a new fashion that spreads because people imitate those who have already adopted. The resulting dynamics are similar to those of an epidemic; indeed, some of the models are borrowed directly from the epidemiology literature. In the context of innovation diffusion, it is common to use a two-parameter model that allows for contagion from within the group at one rate, and from sources outside the

group at a different rate. This is known as the *Bass model* of new product diffusion (Bass 1969, 1980) or the *mixed-influence diffusion model* (Mahajan and Peterson 1985). In the context of a new fashion in clothing, for example, these two rates would correspond to seeing other people on the street who are wearing it, and seeing television ads that promote it.⁸

Let us begin by describing the homogeneous version of the model, later we will introduce heterogeneity. Let λ be the instantaneous rate at which a current nonadopter “hears about” the innovation from a previous adopter *within* the group, and let γ be the instantaneous rate at which he hears about it from sources *outside* of the group. We shall assume that λ and γ are nonnegative, and that not both are zero. In the absence of heterogeneity, such a process is described by the ordinary differential equation $\dot{p}(t) = (\lambda p(t) + \gamma)(1 - p(t))$, and the solution is

$$(2) \quad p(t) = [1 - \beta\gamma e^{-(\lambda+\gamma)t}]/[1 + \beta\lambda e^{-(\lambda+\gamma)t}], \quad \lambda > 0.$$

When contagion is generated purely from internal sources ($\gamma = 0$), this boils down to the ordinary logistic function, which is of course *S-shaped*.⁹ When innovation is driven solely by an external source ($\gamma > 0$ and $\lambda = 0$), we obtain the pure inertia discussed earlier. When both γ and λ are positive, we can choose β in expression (2) so that $p(0) = 0$; namely, with $\beta = 1/\gamma$ we obtain

$$(3) \quad p(t) = [1 - e^{-(\lambda+\gamma)t}]/[1 + (\lambda/\gamma) e^{-(\lambda+\gamma)t}].$$

This model has spawned many variants, some of which assume a degree of heterogeneity, such as two groups with different contagion parameters (Karshenas and Stoneman 1992; Geroski 2000), while others employ distributions from a specific parametric family, such as gamma distributions (Jeuland 1981).

In fact, we can formulate a fully heterogeneous version that is analytically tractable and places virtually no restrictions on the joint distribution of the parameters. Specifically, let μ be the *joint distribution* of the contagion parameters λ and γ . Assume for analytical convenience that μ has bounded support, which we may take to be $\Omega = [0, 1]^2$. (Rescaling λ and γ by a common factor is equivalent to changing the time scale, so this involves no real loss of generality.) The only substantive restriction that we shall place on μ is that, averaged over the whole population, external influence is positive ($\bar{\gamma} = \int_{\Omega} \gamma d\mu > 0$), for otherwise the process cannot get out of the initial state $p(0) = 0$.

Let $p_{\lambda,\gamma}(t)$ be the *proportion* of all type- (λ, γ) individuals who have adopted by time t . Then the proportion of all individuals who have adopted by time t is

$$(4) \quad p(t) = \int p_{\lambda,\gamma}(t) d\mu.$$

(Hereafter integration over Ω is understood.) Each subpopulation of adopters $p_{\lambda,\gamma}(t)$ evolves according to the differential equation

$$(5) \quad \dot{p}_{\lambda,\gamma}(t) = (\lambda p(t) + \gamma)(1 - p_{\lambda,\gamma}(t)).$$

⁸ Some forms of Internet advertising attempt to exploit internal feedback effects by informing members of a target group how many other members of their group have already adopted.

⁹ The logistic model was common in the early work on innovation diffusion; see, for example, Zvi Griliches (1957) and Edwin Mansfield (1961).

This defines an infinite system of first-order differential equations coupled through the common term $p(t)$. We can reduce it to a single differential equation by the following device: let $x_{\lambda,\gamma}(t) = \ln(1 - p_{\lambda,\gamma}(t))$ and observe that (5) is equivalent to the system $\dot{x}_{\lambda,\gamma}(t) = -(\lambda p(t) + \gamma)$ for all (λ, γ) . From this and the initial condition $x_{\lambda,\gamma}(0) = 0$ we obtain

$$(6) \quad x_{\lambda,\gamma}(t) = - \int_0^t (\lambda p(s) + \gamma) ds = -\lambda \int_0^t (p(s) ds) - \gamma t.$$

From the definition of $x_{\lambda,\gamma}(t)$ it follows that

$$(7) \quad p(t) = 1 - \int e^{x_{\lambda,\gamma}(t)} d\mu,$$

that is, $p(t)$ satisfies the integral equation

$$(8) \quad p(t) = 1 - \int e^{-\gamma t - \lambda \int_0^t p(s) ds} d\mu.$$

In other words, the current number of nonadopters $1 - p(t)$ decays according to a weighted average of exponential functions, where the exponent depends on the elapsed time t and the average adoption level up to t . Starting from the initial condition $p(0) = 0$, one can solve for $p(t)$ by successive approximation.

In this model, the hazard rate $\dot{p}(t)/(1 - p(t))$ may be increasing or decreasing depending on the relative importance of the internal and external sources of contagion. It turns out, however, that the hazard rate cannot increase relative to the current number of adopters, that is, the ratio $\dot{p}(t)/[p(t)(1 - p(t))]$ must be nonincreasing. Furthermore, if contagion from external sources is present, the ratio must be strictly decreasing. We shall call this ratio the *relative hazard rate*.

When the parameters λ, γ are fixed, it is obvious that the relative hazard rate is nonincreasing; indeed, this follows at once from the equation $\dot{p}(t) = (\lambda p(t) + \gamma)(1 - p(t))$. The claim is not so obvious when the parameters are heterogeneously distributed, including the possibility that λ and γ are negatively correlated: people who are heavily influenced by outside sources (high γ) might be less susceptible to information from inside sources (low λ), and so forth.

PROPOSITION 1: *Suppose that diffusion is driven by heterogeneous contagion with positive external influence. Then the hazard rate decreases relative to the number of adopters, which implies that the process cannot accelerate beyond $p = 1/2$.*

The last statement follows immediately from the prior one, because if $p(t)$ were to accelerate beyond $1/2$, then at that point the numerator of $\dot{p}(t)/[p(t)(1 - p(t))]$ would be increasing and the denominator would be decreasing, contradicting the assertion that the hazard rate decreases relative to the number of adopters (this is proved in the Appendix).

We remark that there are perfectly reasonable S -shaped curves that do not satisfy the monotonicity condition in Proposition 1. Consider, for example, an adoption process described by the differential equation $\dot{p}(t) = p^a(t)(1 - p(t))$, which was first proposed as a model of innovation diffusion by Christopher J. Easingwood, Mahajan, and Muller (1981, 1983). When $a > 1$, $\dot{p}(t)/[p(t)(1 - p(t))] = p^{a-1}(t)$ is strictly increasing. Hence Proposition 1 shows that such a process cannot arise from a contagion model with any amount of heterogeneity. Nevertheless, it is possible to generate an S -shaped curve from a contagion model—in fact from a homogeneous contagion model—whose overall appearance is very similar (see Figure 1).

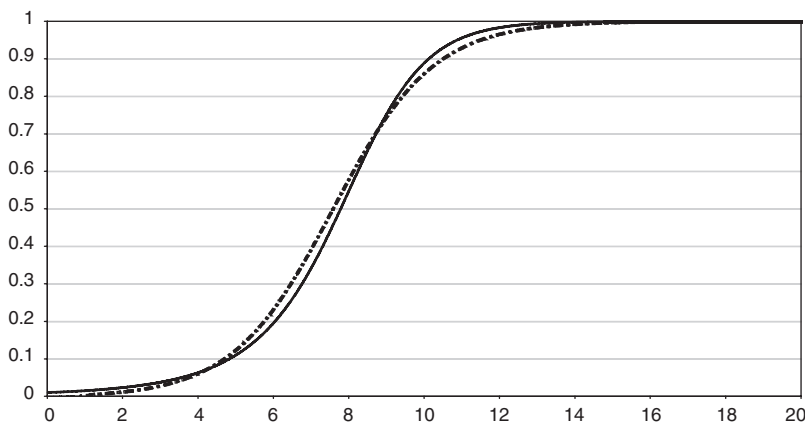


FIGURE 1. TWO ADOPTION CURVES
(The solid line is generated by $\dot{p}(t) = p^{1.1}(1-p)$ and $p(0) = 0.01$, the dashed line by a Bass model with $\lambda = 0.75$ and $\gamma = 0.0025$)

IV. Social Influence

The sociological literature on innovation stresses the importance of social pressure on individual decisions about whether to adopt. In the standard model individuals are assumed to have different “thresholds” that determine whether they will adopt as a function of the number (or proportion) of others in the population who have already adopted. (The dynamics of these models were first studied by Thomas C. Schelling 1971, 1978; Mark Granovetter 1978; and Granovetter and Roland Soong 1988; for more recent work in this vein, see Michael W. Macy 1991; Valente 1995, 1996, 2005; Peter Sheridan Dodds and Watts 2004, 2005; Damon Centola 2006; and Lopez-Pintado and Watts 2008.)

For each agent i , suppose that there exists a minimum proportion $r_i \geq 0$ such that i adopts as soon as r_i or more of the group has adopted. (If $r_i > 1$, the agent never adopts.) This is called the *social threshold* of agent i . The precise meaning of these thresholds varies from one context to another; broadly speaking we can think of them as *degrees of responsiveness to social influence*. A concrete example would be the transmission of rumors: some people would need to hear the rumor from many people to pass it on while others might need to hear it only once. Another is fashion: some people need to see only a few others wearing a new kind of clothing in order to try it, whereas others will wait until the fashion is widespread before jumping on the bandwagon. A key feature of such a model is that the adoption depends on the innovation’s current *popularity* rather than on how good or desirable the innovation has proven to be.¹⁰ The latter is the basis of *social learning models*, which we shall take up in the next section.

We wish to model the aggregate dynamics without assuming a specific parametric form for the distribution of thresholds. To this end, let $F(r)$ be the cumulative distribution function of thresholds in some given population. At time t the proportion of people whose thresholds have been crossed is $F(p(t))$. Of these, $p(t)$ have already adopted, so the proportion whose thresholds have been crossed but have not yet adopted is $F(p(t)) - p(t)$. Let $\lambda > 0$ be the instantaneous rate at which these people convert. Then the adoption process is described by the differential equation

¹⁰ Popularity is a factor in some of the economics models of diffusion; see, in particular, Ellison and Fudenberg (1993, 1995).

$$(9) \quad \dot{p}(t) = \lambda[F(p(t)) - p(t)], \quad \lambda > 0.$$

Assume that $F(0) > 0$ and let b be the *first fixed point* of F , that is, the smallest number in $(0, 1]$ such that $F(b) = b$, if any such exists; otherwise, let $b = 1$. We then have $F(p) > p$ for all $p \in [0, b)$. Since (9) is a separable ordinary differential equation, we obtain the following explicit solution for the inverse function $t = p^{-1}(x)$:

$$(10) \quad \forall x \in [0, b), \quad t = p^{-1}(x) = (1/\lambda) \int_0^x dr / (F(r) - r).$$

Notice that the right-hand side is integrable because $F(r)$ is monotone nondecreasing and $F(r) - r$ is bounded away from zero for all r in the interval $[0, x]$ whenever $x < b$. (The constant of integration is zero because of the initial condition $p(0) = 0$.) As $x \rightarrow b$, the right-hand side of (10) goes to infinity, which implies that the adoption curve approaches b asymptotically. In particular, *the adoption process peters out as it approaches the first fixed point of F .*

This phenomenon is illustrated in Figure 2. Here we assume that the thresholds are normally distributed to the right of the origin, and there is a point mass at the origin corresponding to the subset of innovators—the people who are willing to adopt even when no one else adopted. Notice that the adoption curve asymptotes to $p = 0.50$, which is the first fixed point of F . There is nothing special about the normal distribution in this regard; similar results hold for any cumulative distribution function (CDF) where it first crosses the 45° line.

The adoption curve in Figure 2 is concave, but this is by no means necessary or even typical for this family of models. Figure 3 illustrates an entirely different adoption curve generated by a truncated normal with smaller mean and variance. An interesting feature of this curve is that it accelerates very sharply in the early stages; indeed, it can be shown that it grows at a faster-than-exponential rate up to $p = 0.10$.

We claim that whenever the process starts in the left tail of the distribution, the curve must have one of two shapes: either it decelerates initially, or it accelerates initially at a superexponential rate. To see why this is so, consider the basic dynamic equation in (9). Assume that $F(0) > 0$ and that F has a continuous density $f(p)$ defined for all $p \in (0, 1]$. (Note that the density is not defined at the origin because there is a point mass there.) Differentiating (9) with respect to t and dividing through by $\dot{p}(t)$, which by assumption is positive, we obtain

$$(11) \quad \ddot{p}(t)/\dot{p}(t) = \lambda[f(p(t)) - 1].$$

In other words, the relative acceleration rate traces out a positive linear transformation of the underlying density. It follows that the process accelerates initially if and only if the initial density is large enough, that is, $f(p) > 1$ in a neighborhood of the origin. Suppose further that $f(p)$ is increasing in a neighborhood of the origin, that is, the process starts in the left tail of the distribution of thresholds. Then (11) shows that the relative acceleration rate $\ddot{p}(t)/\dot{p}(t)$ is also strictly increasing; that is, the adoption curve exhibits *superexponential growth*.¹¹ This phenomenon results from the compounding of two effects. First, as more and more people adopt, the amount of information available to the remainder of the population increases. Second, the number of people persuaded by each additional bit of information increases as the process moves up the left tail of the distribution.

These conclusions continue to hold when λ is heterogeneously distributed according to some distribution function $\nu(\lambda)$. Since λ is a scaling parameter, there is no real loss of generality in

¹¹ Growth is *exponential* if $\ddot{p}(t)/\dot{p}(t)$ is constant, *superexponential* if $\ddot{p}(t)/\dot{p}(t)$ is strictly increasing, and *subexponential* if $\ddot{p}(t)/\dot{p}(t)$ is strictly decreasing.

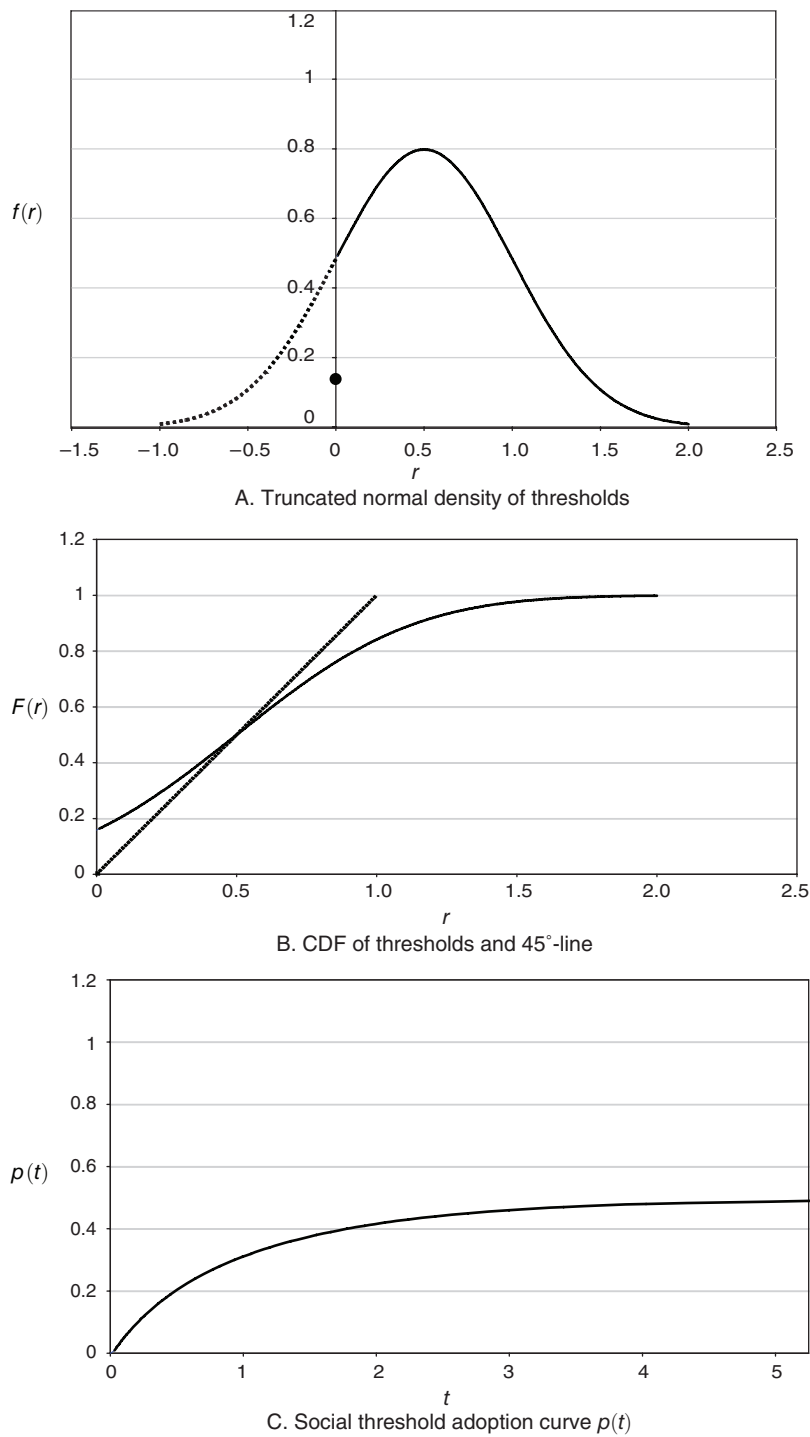


FIGURE 2. DENSITY, CDF, AND SOCIAL THRESHOLD ADOPTION CURVE GENERATED BY $N(0.50, 0.25)$ AND $\lambda = 4$

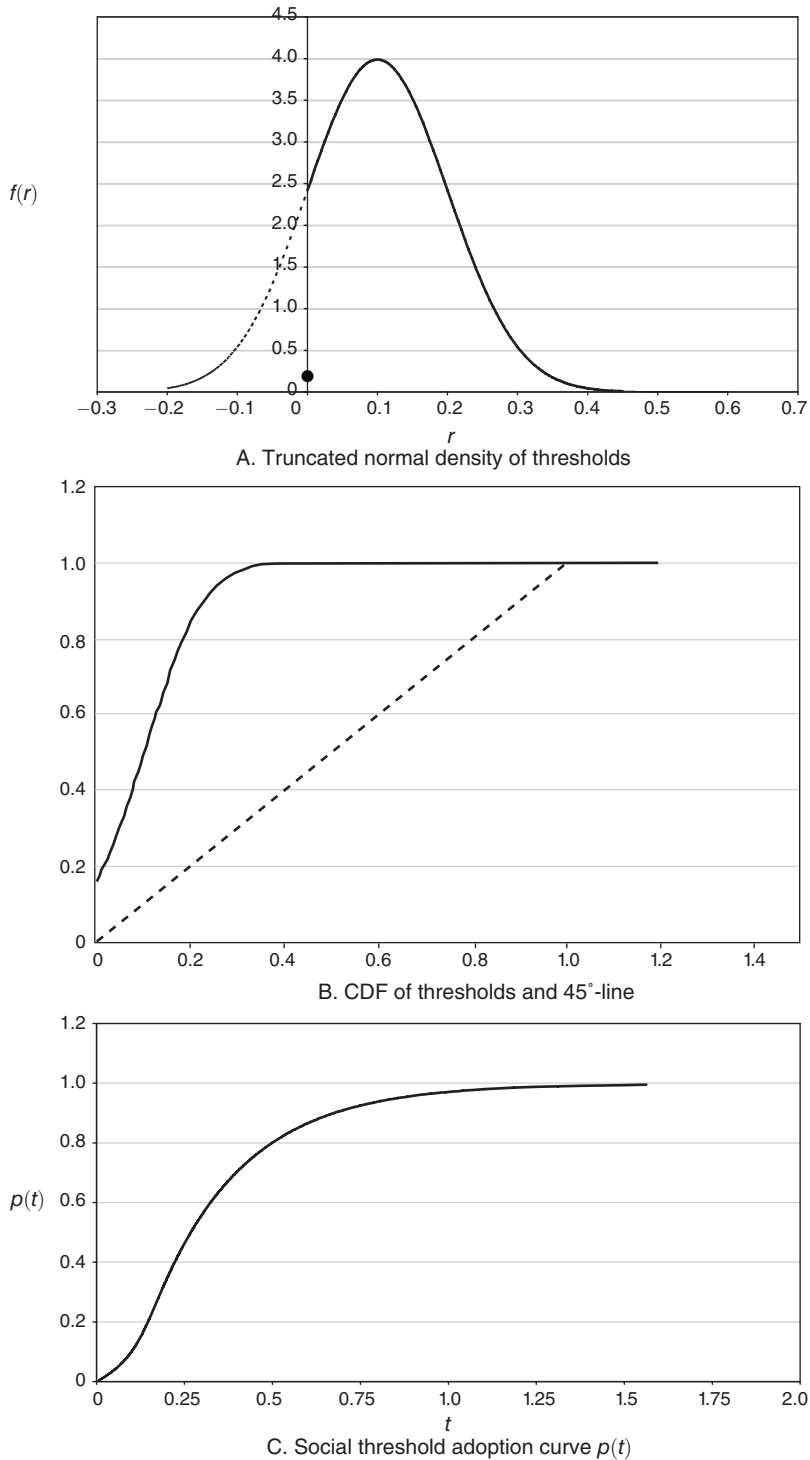


FIGURE 3. DENSITY, CDF, AND SOCIAL THRESHOLD ADOPTION CURVE GENERATED BY $N(0.10, 0.01)$ AND $\lambda = 4$

assuming that the support of ν lies in $(0, 1]$. Let $F_\lambda(p)$ be the cumulative distribution of thresholds p conditional on λ , and assume that the conditional density $f_\lambda(p)$ exists for every $p \in (0, 1]$. The following proposition is proved in the Appendix.

PROPOSITION 2: *Suppose that diffusion is driven by a heterogeneous social threshold model, where there is a positive mass of individuals who are prepared to adopt when no one else adopts, and the threshold densities $f_\lambda(p)$ are positive and increasing in a common neighborhood of zero for all λ . If the process accelerates initially, then it accelerates initially at a superexponential rate for some period of time.*

V. Social Learning

A notable limitation of the contagion and social influence models is that they provide no clear reason *why* people would adopt an innovation given that others have adopted it. In this section we consider a class of learning models in which the adoption decision flows directly from the rational evaluation of evidence. Specifically, let us suppose that an agent adopts if he has reason to believe the innovation is better than what he is doing now, where the evidence comes from directly observing the outcomes among prior adopters. For example, when a new product becomes available—e.g., a new type of medication (antibiotics) or a new communication technology (cell phones)—people will want to see how it works for others *over a period of time* before trying it themselves. These are variously known as *social learning models* or *social learning models based on direct observation*.

There is a sizable theoretical literature on social learning, but it is difficult to summarize due to the great diversity in behavioral and informational assumptions that different authors use. Some assume that payoff outcomes among prior adopters are fully observable, while others assume only that the act of adoption is observable (the latter are usually called herding models). Some assume that agents can recognize others' types, while others assume that types cannot be identified. There are also significant differences in the relevant *characteristics* that authors choose to focus on, including heterogeneity in risk aversion, discount rates, and amount of information.¹² There is considerable empirical evidence, however, that learning from the experience of others does in fact occur.¹³

Here we shall make a number of simplifying assumptions in order to get a handle on an issue that has not received much previous attention in this literature, namely, *what do the short-run aggregate dynamics of a heterogeneous learning model look like, and do they differ qualitatively from the dynamics generated by other classes of models?*

To make some progress on this question let us make the following assumptions:

- (i) Payoffs are observable;
- (ii) Payoffs generated by different individuals and/or at different points in time are independent and equally informative;

¹² See, among others, Bikchandani, Hirshleifer, and Welsh (1992); Banerjee (1992); Kirman (1993); Ellison and Fudenberg (1993, 1995); Kapur (1995); Bala and Goyal (1998); Smith and Sorensen (2000); Chatterjee and Xu (2004); Banerjee and Fudenberg (2004); Manski (2004); Golub and Jackson (2008).

¹³ Andrew Foster and Mark Rosenzweig (1995); Kaivan Munshi (2004); Timothy G. Conley and Christopher R. Udry (2005).

- (iii) Agents are risk-neutral and myopic;
- (iv) There is no idiosyncratic component to payoffs due to differences in agents' types, but agents may have different costs (not necessarily observable);
- (v) There are differences in agents' prior beliefs about how good the innovation is relative to the status quo;
- (vi) There are differences in the average number of people they observe and hence in the amount of information they have;
- (vii) The population is fully mixed.

Many other complicating factors could be introduced, such as discounting, one-time switching costs, risk aversion, imperfect observability, and restricted information acquisition (such as through social networks), but these extensions would distract from the main point, which is to identify the dynamic characteristics of a fairly general class of learning models without attempting to formulate the most general such model.

We claim that, under the assumptions above, the dynamical system has a surprisingly simple structure. One can reduce the various types of heterogeneity to a *composite index* that measures the *probability of a given agent adopting, conditional on the amount of information that has been generated thus far in the population*. The equation of motion of such a process turns out to be analogous to the one obtained for the social influence model, except that here the relevant state variable at time t turns out to be the *integral of the adoption curve* up through t , $\int_0^t p(s) ds$, rather than the *level of adoption* $p(t)$.¹⁴ The reason is that $\int_0^t p(s) ds$ measures the cumulative information generated by all prior adopters *from the time they first adopted*, which is the relevant variable in the learning context.

To illustrate how such a model works, let us walk through a particular example using a standard normal-normal updating framework. This is chosen mainly for its computational transparency; similar results hold under alternative assumptions.¹⁵ Assume that the payoff from the innovation is a normally distributed random variable X with mean $\mu > 0$ and variance σ^2 , which is i.i.d. among agents and time periods. We shall interpret μ as the *mean payoff gain per period from using the innovation as compared to the status quo technology*. Each agent is assumed to have an idiosyncratic variable cost c_i of using the innovation, so he adopts if and only if he believes that the mean payoff per period is at least c_i (thus, he may switch back if after receiving more information he believes that the mean payoff is less than c_i).

If everyone knew the true value of μ from the outset, then everyone would adopt for whom $c_i < \mu$. This is the *efficient outcome*. Ex ante, however, people do not know the true value of μ ; furthermore, they may start with substantially different beliefs (based on their prior private information) about what the true value is. As more information comes in, they update their beliefs. If this information is sufficiently favorable, more people will adopt, which creates a still-larger base of information, which causes even more people to adopt, and so on. This is the essential logic driving the learning dynamics.

¹⁴ Recall that the integral of the adoption curve also featured in the heterogeneous contagion model (see equation (8)).

¹⁵ Jensen (1982) and Lopez-Pintado and Watts (2008) study the case where the outcome variable is binomial (payoffs are "high" or "low"). They assume that agents pay attention only to current outcomes, not the cumulative amount of information generated from earlier periods.

To continue with our example, suppose that each agent i has a prior belief about the unknown mean μ and unknown precision $\rho = 1/\sigma^2$ such that: (i) the marginal of ρ is gamma-distributed, and (ii) for each value of ρ the conditional distribution of μ is normal with mean μ_{i0} and precision $\rho\tau_i$. (This is a standard normal-normal updating model; see Morris De Groot 1970.) Low values of τ_i reflect *flexibility* in beliefs, that is, relatively little evidence is needed to shift i 's belief about the mean by a given amount. Low values of μ_{i0} reflect *pessimism* about the payoffs from the innovation. In particular, if i initially believes that the mean is less than his costs, $\mu_{i0} < c_i$, he will not want to adopt. As more information comes in, however, his posterior estimate of the mean, μ_{it} , may increase sufficiently that he changes his mind. The point at which this happens depends, among other things, on how much information i collects and how flexible his beliefs are.

At a given point in time t , agent i will have seen (or heard about) a random assortment of prior outcomes depending on the chance encounters he had with other members of the population. In a continuous-time setting the total “number” or measure of prior outcomes up to time t is given by the integral under the adoption curve, namely, $r(t) = \int_0^t p(s) ds$. Assuming that i is equally likely to see any particular outcome, we can model i 's flow of information by a Poisson arrival process. In particular, let N_{it} be a Poisson random variable representing the number of i 's observations up to time t , where $E[N_{it}] = \beta_i r(t)$ and the parameter β_i is a measure of the extent to which i “gets around.”¹⁶

Let n_{it} denote the realization of N_{it} , and let \bar{x}_{it} denote the mean payoff among these n_{it} observations. Given our assumptions, \bar{x}_{it} is normal with mean μ and standard deviation $\sigma/\sqrt{n_{it}}$. In the normal-normal updating framework, i 's Bayesian posterior estimate of the mean, μ_{it} , can be expressed very simply as a convex combination of μ_{i0} and \bar{x}_{it} , namely,

$$(12) \quad \mu_{it} = \frac{n_{it}\bar{x}_{it} + \tau_i\mu_{i0}}{n_{it} + \tau_i}.$$

In other words, the posterior estimate is just a weighted average of the prior and the observed mean, where the weight on the mean is the number of independent observations that produced it.

Given our assumption that i is myopic, she is prepared to adopt once μ_{it} is at least equal to her variable c_i , which by (12) is equivalent to

$$(13) \quad (\bar{x}_{it} - c_i)n_{it} \geq \tau_i(c_i - \mu_{i0}).$$

By assumption, $(\bar{x}_{it} - \mu)(\sqrt{n_{it}}/\sigma) = z_{it}$ is $N(0, 1)$; hence (13) can be rewritten as

$$(14) \quad \left(\frac{\sigma z_{it}}{\sqrt{n_{it}}} + \mu - c_i \right) n_{it} \geq \tau_i(c_i - \mu_{i0}).$$

¹⁶ If agents were embedded in a fixed social network, the analogous parameter would be the number of other agents with whom a given agent is connected. In this case, however, the aggregate amount of information $r(t)$ will generally not be sufficient to describe the state of the system; the dynamics of the process will depend on the specific network topology. See Golub and Jackson (2008).

To make sense of this expression, let us focus on the subpopulation of agents for whom adoption is worthwhile, $P^0 = \{i: \mu > c_i\}$, and substitute the expected value $E[n_{it}] = \beta_i r(t)$ into (14). After rearranging terms we obtain

$$(15) \quad r(t) \geq \frac{\tau_i(c_i - \mu_{i0})}{\beta_i(\mu - c_i)} - \frac{\sigma \sqrt{r(t)} z_{it}}{(\mu - c_i) \sqrt{\beta_i}}.$$

Define i 's *resistance level* (or *information threshold*) to be the expected value of the right-hand side of this inequality, namely,

$$(16) \quad r_i = \frac{\tau_i(c_i - \mu_{i0})}{\beta_i(\mu - c_i)}.$$

We shall say that an agent is of *type* i if she has the vector of characteristics $(c_i, \tau_i, \beta_i, \mu_{i0})$. From (15) and (16) we see that an agent of type i is increasingly likely to adopt as $r(t)$ passes the threshold r_i . Moreover, this threshold has a natural interpretation: agents with high r_i are those who are initially *pessimistic* that the innovation will cover their costs ($c_i - \mu_{i0}$ is large), *inflexible* in their initial beliefs (τ_i is large), *marginally profitable* ($\mu - c_i$ is low), and *relatively uninformed* (β_i is small).

The precise form of expressions (15) and (16) is not essential for our purpose, however: the key point is that each agent i has a *response function* $\varphi_i(r)$, which represents the probability that i believes the innovation is worth adopting, given that the total amount of information generated by the prior adopters equals r .¹⁷

Let us now abstract from this particular example and take the notion of a response function as a primitive of the learning model. Suppose that each agent in the population has a "type" i that is characterized by a response function $\varphi_i: R_+ \rightarrow [0,1]$, where $\varphi_i(r)$ is the probability that i 's information threshold has been crossed when the total amount of information generated by the prior adopters is r . For ease of interpretation we shall assume that the functions $\varphi_i(r)$ are monotone nondecreasing, though this is not actually necessary for some of the results to follow. Notice that a given individual will typically know only a small fraction of the prior outcomes, that is, r , is a state variable that represents a common pool of information but it is not common knowledge.

Assume that there are countably many types in the population, and let p_i be the *proportion* of i -types.¹⁸ When the total information generated by prior adopters equals r , the proportion of the population whose thresholds have been crossed is given by the function

$$(18) \quad F(r) = \sum_i p_i \varphi_i(r).$$

¹⁷ In the present case, the response function has the following explicit representation. Given that $r(t) = r$, n_{it} is Poisson-distributed with mean $\beta_i r$. Given a realization $n_{it} = k > 0$, the mean observed payoff, \bar{x}_{it} , is normal with mean μ and variance σ^2/k . Let Φ denote the standard normal CDF. Then the probability that i 's posterior estimate exceeds i 's costs, i.e., the probability that (14) holds, is

$$(17) \quad \varphi_i(r) = \sum_{k=1}^{\infty} \frac{(\beta_i r)^k e^{-\beta_i r}}{k!} \Phi \left[\frac{(\mu - c_i) \sqrt{k}}{\sigma} - \frac{\tau_i(c_i - \mu_{i0})}{\sigma \sqrt{k}} \right].$$

¹⁸ Alternatively we could assume that the types are uncountable and distributed according to a density; this makes no material difference in the subsequent analysis.

$F(r)$ is a monotone nondecreasing function which we can interpret as a *notional distribution function* of agents' information thresholds. We allow for the possibility that some agents have an infinite threshold; hence, $\lim_{r \rightarrow \infty} F(r)$ may be less than one.¹⁹

Recalling that $r(t) = \int_0^t p(s) ds$, it follows that the aggregate dynamics is described by the differential equation

$$(19) \quad \dot{p}(t) = \lambda \left[F\left(\int_0^t p(s) ds\right) - p(t) \right], \quad \lambda > 0.$$

Observe that (19) is analogous to the dynamical equation (9) defining a social threshold model, except that in the present case the argument of F is the *integral of the adoption curve* rather than the adoption curve itself. This arises because agents use all past information generated by previous adopters rather than just the most recent information. Note that this formulation hinges on our assumption that payoff information is equally informative no matter when it was received. If this were not the case, for example, if payoffs received in earlier time periods were less valuable, the defining equation would have to be modified accordingly.²⁰

The cumulative feature of the social learning model has some important implications for the shape of the adoption curve. Compared to the social influence model, acceleration is initially quite weak; in fact, in a neighborhood of the origin the process strictly *decelerates* irrespective of the distribution generating it. To see why this is so, differentiate equation (19) to obtain the acceleration equation

$$(20) \quad (1/\lambda) \ddot{p}(t) = p(t)f(r(t)) - \dot{p}(t).$$

By assumption, $p(0) = 0$ and the solution $p(t)$ is continuous, so the first term is close to zero when t is close to zero. However, the second term is bounded away from zero when t is close to zero, because (19) implies that $\dot{p}(0) = \lambda F(0) > 0$. Hence,

$$(21) \quad \lim_{t \rightarrow 0^+} \ddot{p}(t) = -\lambda^2 F(0) + f(p(0))p(0) < 0.$$

More generally, when we compare (20) with the acceleration equation for the social influence model, namely, $(1/\lambda) \ddot{p}(t) = \dot{p}(t)f(p(t)) - \dot{p}(t)$, we see that the former has weaker initial acceleration than the latter because $\dot{p}(t)$ is initially positive, whereas $p(t)$ is initially zero. *The result may be an extended period of weak growth in the early phases of social learning.* The reason is that the initial block of optimists $F(0)$ exerts a decelerative drag on the process: they contribute at a decreasing rate as their numbers diminish, while the information generated by the new adopters gathers steam slowly because there are so few of them to begin with.²¹ Figure 4 illustrates this phenomenon for the same density that generated the social thresholds adoption curve in Figure 3.

¹⁹ The closest model to this one in the literature is due to Dodds and Watts (2004, 2005). They consider a cumulative-dose model of infection with heterogeneity in the thresholds at which agents become infected (including social as well as biological interpretations of infection), though their analysis of the dynamics emphasizes different features from the ones considered here.

²⁰ If payoff information is discounted at some rate $\delta > 0$, the dynamical equation takes the form $\dot{p}(t) = \lambda [F(\int_0^t e^{-\delta(t-s)} p(s) ds) - p(t)]$. Note that when δ is large, this process is similar to the social influence model. Another modification is to allow for joint heterogeneity in λ and r ; this extension is straightforward and is left to the reader.

²¹ Initial deceleration does not necessarily occur, however, if observations are *bunched* at particular points in time. An example would be an agricultural innovation (e.g., a new type of crop) that is tried once in each growing season, and whose outcomes farmers observe at the end of the season. This case will be discussed further in Section VI.

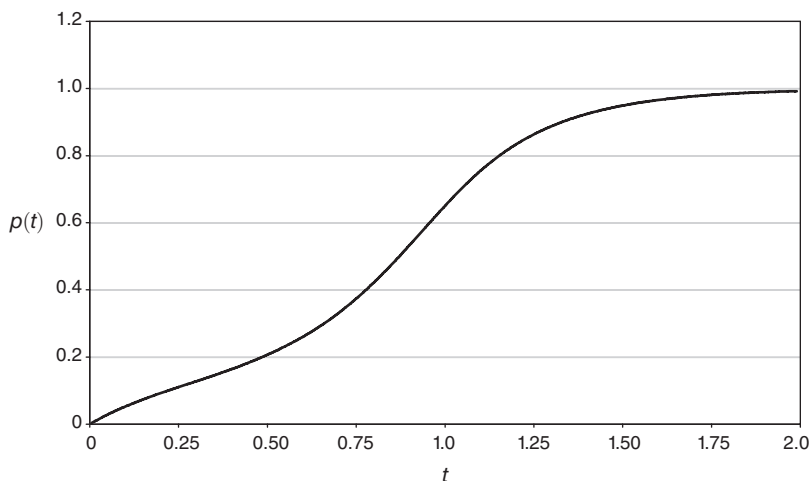


FIGURE 4. ADOPTION CURVE GENERATED BY SOCIAL LEARNING, A NORMAL DISTRIBUTION OF INFORMATION THRESHOLDS $N(0.10, 0.01)$, AND $\lambda = 4$

Next we shall investigate the behavior of the relative acceleration rate $\ddot{p}(t)/\dot{p}(t)$. To avoid uninteresting technical complications let us assume that the density is well behaved near the origin, that is, $\bar{f}(0) = \lim_{r \rightarrow 0^+} f(r) > 0$, and that f' is continuous and bounded. Define

$$(22) \quad \alpha(t) = (1/\lambda) \ddot{p}(t)/\dot{p}(t).$$

From (20) we deduce that

$$(23) \quad \alpha(t) = f(r(t))p(t)/\dot{p}(t) - 1.$$

Differentiating (23), we obtain

$$(24) \quad \dot{\alpha}(t) = f'(r(t))p^2(t)/\dot{p}(t) + f(r(t)) - f(r(t))\ddot{p}(t)p(t)/\dot{p}^2(t).$$

As $t \rightarrow 0^+$, the first term in (24) goes to zero, because by assumption f' is bounded, $p(t) \rightarrow 0$, and $\dot{p}(0) > 0$. The third term also goes to zero. However, $f(r(t)) \rightarrow \bar{f}(0) > 0$, so the second term is positive in the limit. It follows from continuity that $\dot{\alpha}(t)$ is strictly positive on some initial interval $0 \leq t \leq T$. In the region near the origin where $\dot{p}(t) < 0$, this says that the relative acceleration rate is becoming less negative.

Now, suppose that at some time t_0 , the process begins to accelerate. If at this point the density is increasing, then $f'(r(t_0)) > 0$, and it follows from (24) and continuity that $\dot{\alpha}(t)$ is positive for some interval of time beginning at t_0 . In other words, if the density is increasing when the process begins to accelerate (assuming it ever does accelerate), then the process undergoes a period of superexponential growth.

PROPOSITION 3: *Suppose that diffusion is driven by social learning, where there is a positive mass of individuals who are willing to adopt even when no one else adopts ($F(0) > 0$), and the distribution of resistances has a density $f(r)$ that is bounded away from zero near the origin and*

has a continuous bounded derivative. Then, initially the process decelerates whereas the relative acceleration rate strictly increases. If the process begins to accelerate, and if the density is increasing at that point, then the process undergoes a period of superexponential growth.

Propositions 2 and 3 show that both the social influence model and the social learning model can exhibit periods of superexponential growth. This is consistent with observations by various authors that adoption curves often exhibit a sudden take-off phase after a slow start (Ryan and Gross 1943; Henrich 2001; Maarten C. M. Vendrik 2003; Rogers 2003). The propositions above are more specific in predicting superexponential growth, and they clarify why one would expect it theoretically. Theory also suggests why the curves generated by social influence and social learning will differ in other respects. One difference is that, in a social learning model, the speed of adoption at a given point in time depends on adoption levels at earlier points in time, whereas in a social influence model this is not the case. Given sufficiently detailed adoption data, this effect could be tested by regression analysis with lagged variables.

A second distinction between the models is the way in which the resistance thresholds are constructed. In a social influence model the thresholds reflect differences in a single scalar parameter (people's tendency to conform), whereas in a social learning model they reflect differences in a variety of factors, including prior beliefs, amount of information gathered, and the payoff gain from the innovation. Some of these factors are unobservable and would be difficult to estimate even with micro-level adoption data. However, the payoff gain—the parameter μ in the social learning model—can sometimes be estimated from the data, as we shall see in the next section.

The key point to observe is that higher average payoffs from adoption shift the individual response functions upward: for a given “number” of prior outcomes r in the population, the probability $\varphi_i(r)$ that any given individual i adopts increases the higher the potential payoff gain μ . Thus, if we consider two distinct populations $\{1, 2\}$ such that one has a higher mean payoff than the other, say $\mu_2 > \mu_1$, then all else being equal the resistance distributions satisfy $F_2(r) \geq F_1(r)$ for all r . (F_1 first-order stochastically dominates F_2 .) Under these circumstances, adoption occurs more rapidly in population 2 than in population 1; that is, the solutions satisfy $p_2(t) \geq p_1(t)$ for all t .²² This is one way to evaluate the plausibility of the social learning model without knowing much about the distribution of the learning parameters.

VI. The Diffusion of Hybrid Corn

Diffusion data that permit definitive tests of the preceding propositions are not easy to come by. One reason is that an adoption curve must be charted at frequent time intervals in order to conduct tests of statistical significance. Consider, for example, a product or practice that took ten years to go from introduction to completion. If the adoption level is measured annually, there will be only ten data points, only nine first differences, and only eight second differences; furthermore, the errors in successive differences will not be independent. Of course, the situation would be different if we had monthly or weekly data over a period of ten years, but there appear to be very few datasets of this sort in the literature. Moreover, the data that do exist tend to be for new products where mass advertising played a critical role. In these cases it is difficult to sort out the

²² Consider the dynamical equations $\dot{p}_i(t) = \lambda[F_i(p_i(t)) - p_i(t)]$, $i = 1, 2$. Let $p_1(t)$, $p_2(t)$ be the solutions, which we assume to be continuous. Note that $\dot{p}_2(0) \geq \dot{p}_1(0)$, because by assumption $F_2(0) \geq F_1(0)$ and $p_1(0) = p_2(0) = 0$. Suppose, by way of contradiction, that $p_2(t) < p_1(t)$ at some time $t > 0$. By continuity there is a time $\tilde{t} > 0$ such that $p_2(s) \geq p_1(s)$ for all $s \leq \tilde{t}$ and $\dot{p}_2(\tilde{t}) < \dot{p}_1(\tilde{t})$. Therefore, $\int_0^{\tilde{t}} p_2(s) ds \geq \int_0^{\tilde{t}} p_1(s) ds$ and $F_2(\int_0^{\tilde{t}} p_2(s) ds) \geq F_1(\int_0^{\tilde{t}} p_1(s) ds)$, from which it follows that $\dot{p}_2(\tilde{t}) \geq \dot{p}_1(\tilde{t})$, a contradiction.

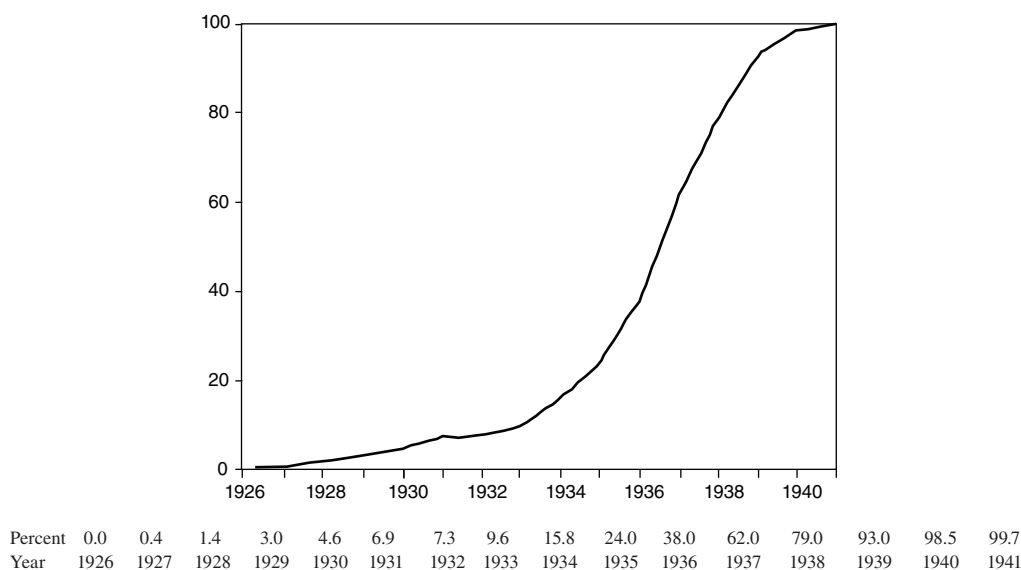


FIGURE 5. PERCENT OF ADOPTERS OF HYBRID CORN IN TWO IOWA COMMUNITIES, 1926–1941
(From Ryan and Gross 1943, Figure 4)

relative importance of internal and external sources of information, though the mixed-influence contagion model is certainly a step in that direction.

Fortunately, there exist some data that are sufficiently detailed to carry out a preliminary analysis of the ideas discussed above. The example we choose here is Ryan and Gross's classic study of the diffusion of hybrid corn in two Iowa communities (Ryan and Gross 1943). This work prompted a number of later studies on innovation diffusion, including Zvi Griliches's analysis of hybrid corn diffusion across the midwestern and southern United States (Griliches 1957). Although the latter is more comprehensive in its geographical coverage, the study by Ryan and Gross has several advantages for our purposes. First, it was carried out at the community level, whereas the adoption curves in Griliches's study were aggregated over several counties and involved scores of dispersed communities. Second, Ryan and Gross provided the proportion of hybrid corn adopters, whereas Griliches gave the proportion of *acreage* planted in hybrid corn, which (among other things) biases the results toward large landowners. Third, the Ryan and Gross data covered the critical start-up phase in the 1920s, which was absent in Griliches's analysis. Finally, Ryan and Gross's study included important related information, such as the year that farmers first learned about the new technology and the sources from which they learned about it. These data provide additional contextual details that, together with the shape of the curve itself, allow us to assess the relative plausibility of different diffusion mechanisms.

The adoption curve is shown in Figure 5. Particularly noteworthy features are the long choppy start, which lasted for some six years (1927–1932), followed by a sharp acceleration for about four years (1933–1936), and then a rapid deceleration as nearly 100 percent acceptance levels were reached in 1941.²³

²³ These features have been noted by other authors, notably Henrich (2001), who explains them by a combination of learning from direct experience and biased cultural transmission.

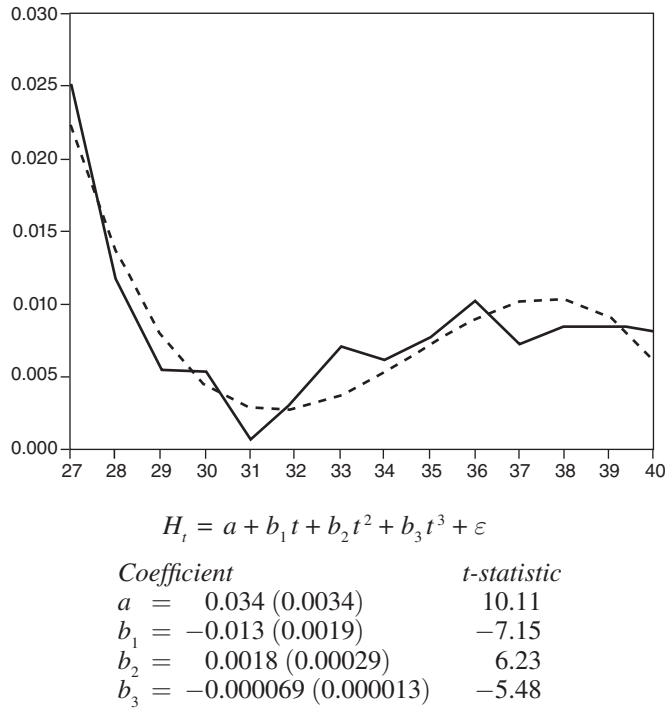


FIGURE 6. CUBIC POLYNOMIAL (*Smooth Line*) FITTED BY OLS TO THE RELATIVE HAZARD RATE $H_t = \Delta_t/p_t(1 - p_t)$ (*Jagged Line*)

Ryan and Gross repeatedly emphasize that natural *conservatism* was one of the key reasons why farmers delayed so long in adopting a technology that could have substantially increased their profits. Indeed, their survey data reveal that, although at least two-thirds of the farmers had heard about the advantages of hybrid corn as early as 1931, only about 8 percent had adopted by then. Thus, pure inertia could well have been the explanation for the long delay. Nevertheless, the shape of the curve suggests that this was not the case. The reason is that, in a pure inertia model with heterogeneous levels of inertia, the adoption curve must be concave throughout. This hypothesis can be rejected with reasonable confidence, since there was deceleration in only one of the first ten time periods, and sharp acceleration in most of them. Thus, although inertia may have been a contributing factor, it was unlikely to have been the sole factor.

Another interesting aspect of Ryan and Gross's survey is that farmers reported their *sources of information* about the new technology, and identified the sources they believed were most influential in their decision to adopt. Whereas seed salesmen were listed as the most common *original source* of knowledge, neighbors were listed as the *most influential in deciding whether to implement it*. This is broadly consistent with the three classes of models under consideration, all of which posit that adoption results from an internal feedback loop between earlier and later adopters.

Let us consider the contagion model first. This posits that adoption decisions are driven by a combination of internal and external influences, where external influence operates at a constant rate and internal influence operates at a rate proportional to the current number of adopters. We know that this model generally produces *S-shaped* curves, but we also know that the curve has

TABLE 1—RELATIVE ACCELERATION RATE MEASURED OVER THREE-YEAR INTERVALS

Interval	$\bar{\Delta}_i$	$\bar{\Delta}_{i+1}/\bar{\Delta}_i - 1$
1927–1929	1.0	—
1930–1932	1.4	1.4
1933–1935	5.6	4.0
1936–1938	18.3	3.3
1928–1930	1.4	—
1931–1933	1.7	1.2
1934–1936	9.5	5.6
1937–1939	18.3	1.9
1929–1931	1.8	—
1932–1934	3.0	1.7
1935–1937	15.4	5.1
1938–1940	12.2	0.8

Note: $\bar{\Delta}_i$ is the average annual change (in percent) over the i^{th} interval, and relative acceleration is $\bar{\Delta}_{i+1}/\bar{\Delta}_i - 1$.

a particular property: the hazard rate must be nonincreasing relative to the current number of adopters. (It can be verified that this property also holds in the discrete-time version.)

To check this, let p_t be the proportion of adopters at the end of period t , and let $\Delta_t = p_{t+1} - p_t$ be the rate of change at t . Then, $H_t = \Delta_t/p_t(1 - p_t)$ should be nonincreasing in t , and strictly decreasing if there is any outside influence. The piecewise linear curve shown in Figure 6 is represented by H_t .

To test whether it is monotone nonincreasing, let us fit a cubic polynomial using OLS (the smooth line in Figure 6). The overall fit is very good ($R^2 = 0.86$), and the quadratic and cubic coefficients *differ in sign* at a highly significant level. We therefore have reason to believe that H_t is not a monotone function, which is corroborating evidence against the contagion hypothesis. Of course this is not the same as saying that there is *no* contagion effect, merely that this effect *by itself* is unlikely to have produced the Ryan and Gross data.

Next let us consider the social influence and social learning models. Ryan and Gross argued that the balance of evidence from the survey suggested some form of learning by farmers: “In some sense the early acceptors provided a community laboratory from which neighbors could gain some vicarious experience with the new seed over a period of some years” (Ryan and Gross 1943, 16). This is precisely the effect posited in the social learning model. Is social learning consistent with the shape of the curve? The answer is affirmative. Indeed the curve has two features that are quite typical of social learning models: there is a long slow start-up followed by very rapid acceleration. We remind the reader why this is so: in the early phases, adoption is driven mainly by people who are already convinced that the new technology is a good idea but who may delay due to inertia; once this effect is overcome, there is rapid acceleration caused by the accumulation of information provided by prior adopters, combined with the rising number of people at the margin who are persuaded by the new information.

Unfortunately, the curve does not have enough data points to carry out tests of significance, but we can make a rough estimate of the relative acceleration rate in the following way. First, compute the average slope over successive three-year intervals in order to smooth out the data. Then, compute the relative acceleration rate as the ratio of successive slopes minus unity.²⁴ The results are shown in Table 1 for three different ways of grouping the data. In every case, the

²⁴ If Δ_t is the slope in period t , then $(\Delta_{t+1} - \Delta_t)/\Delta_t = \Delta_{t+1}/\Delta_t - 1$ is an estimate of the relative acceleration rate at t .

relative acceleration rate rises sharply from the second to the third period. (Recall that in an exponential growth model the relative acceleration rate would be constant.) While this is not proof of superexponential acceleration, it is certainly consistent with it.

These features provide corroborating evidence for social learning, but they certainly do not constitute identification. For example, we cannot rule out the possibility that in the mid-1930s there was an exogenous shock (e.g., a price war among seed companies) that caused the sudden uptick.²⁵ Moreover, the sharp rise in acceleration is equally consistent with the social influence model. In this case, however, there is evidence from another source—namely Griliches's study—that social influence was probably not the sole explanation. The reason is that, in a social learning model, the rate of adoption depends on the potential gain from adoption, whereas in a social influence model this is not the case. One of Griliches's main findings was that in the midwestern states generally (including Iowa), the rate of adoption in each region was positively correlated with the estimated monetary gain from adoption, just as the social learning model predicts.

Putting all of this information together, we can say that social learning is consistent with the observed pattern of diffusion of hybrid corn, although we cannot say that it was the sole explanatory factor. We can also say with some confidence, however, that inertia and contagion were probably *not* the sole explanatory factors, and given Griliches's findings neither was social influence.

While the Ryan and Gross study is quite limited in scope, it highlights some of the issues that would need to be addressed in a more complete empirical analysis. First, one would need a curve—or set of curves—that are measured at frequent time intervals. The reason is that one requires an accurate estimate of the acceleration pattern at precisely the point where acceleration is most rapid; hence, frequent observations in this part of the curve are crucial for running tests of significance. Second, one needs data that cover the whole life cycle of the adoption process, including the very early phases when adoption is just getting started. Many of the datasets in the literature are quite weak in this respect; the reason, of course, is that researchers are not likely to focus their attention on an innovation until it is already well underway.

Third, the framework developed here applies to situations where there is informational feedback between members of a group who interact more or less at random. If information flows through a fixed network, say between individuals who are near neighbors (in a geographical or social sense), then a different analysis will be required. (Ryan and Gross did not say whether farmers learned more or less uniformly from others in the community, or mainly from their immediate neighbors.) Fourth, we have restricted our attention to social learning models in which people can directly observe outcomes generated by other adopters. In the case of hybrid corn, this assumption seems reasonable; indeed Ryan and Gross specifically say that farmers looked at their neighbors' fields to assess the viability of the new technology (Ryan and Gross 1943, 16). In situations where direct observation is not possible, one would need an analog of the theory in which inferences are made indirectly.

Finally, the Ryan and Gross study highlights the importance of having supporting data about the agents' sources of information and the reasons they delayed in acting on it. These contextual details, combined with an analysis of the shape of the curve, have the potential to discriminate between alternative diffusion mechanisms even when detailed micro-level data are unavailable.

²⁵ It is worth noting, however, that Griliches investigated whether price changes in hybrid seed could explain the rates of adoption in various parts of the Midwest (including Iowa), and he found no significant effect (Griliches 1957, 503).

APPENDIX

PROOF OF PROPOSITION 1:

Define the function $H(t) = h(t)/p(t) = \dot{p}(t)/[p(t)(1 - p(t))]$; this is well defined for all $t > 0$ because, by assumption, $\dot{p}(0) = \bar{\gamma} > 0$; hence, $p(t) > 0$ when $t > 0$. We need to show that $\dot{H}(t) < 0$.

For each parameter pair (λ, γ) , let $q_{\lambda, \gamma}(t) = (1 - p_{\lambda, \gamma}(t))$ denote the *proportion of the (λ, γ) -population that has not yet adopted by time t* . The proportion of the total population that has not adopted by t is therefore

$$(A1) \quad q(t) = \int q_{\lambda, \gamma}(t) d\mu.$$

For each (λ, γ) , we have

$$(A2) \quad \dot{p}_{\lambda, \gamma}(t) = (\lambda p(t) + \gamma) q_{\lambda, \gamma}(t).$$

Integration with respect to μ yields

$$(A3) \quad \dot{p}(t) = [\lambda(t)p(t) + \gamma(t)]q(t),$$

where

$$(A4) \quad \lambda(t) = \frac{\int \lambda q_{\lambda, \gamma}(t) d\mu}{q(t)} \quad \gamma(t) = \frac{\int \gamma q_{\lambda, \gamma}(t) d\mu}{q(t)}.$$

Note that $\lambda(t)$ and $\gamma(t)$ are the *expected* values of λ and γ in the population of nonadopters at time t . It follows that

$$(A5) \quad H(t) = \dot{p}(t)/[p(t)q(t)] = \lambda(t) + \gamma(t)/p(t).$$

CLAIM:

$$(A6) \quad \text{For every } t > 0, \dot{\lambda}(t)p(t) + \dot{\gamma}(t) \leq 0.$$

PROOF OF CLAIM:

For every $t > 0$, we have

$$(A7) \quad \dot{\lambda}(t) = \frac{\int \lambda \dot{q}_{\lambda, \gamma}(t) d\mu}{\int q_{\lambda, \gamma}(t) d\mu} - \frac{[\int \lambda q_{\lambda, \gamma}(t) d\mu][\int \dot{q}_{\lambda, \gamma}(t) d\mu]}{[\int q_{\lambda, \gamma}(t) d\mu]^2},$$

and

$$(A8) \quad \dot{\gamma}(t) = \frac{\int \gamma \dot{q}_{\lambda, \gamma}(t) d\mu}{\int q_{\lambda, \gamma}(t) d\mu} - \frac{[\int \gamma q_{\lambda, \gamma}(t) d\mu][\int \dot{q}_{\lambda, \gamma}(t) d\mu]}{[\int q_{\lambda, \gamma}(t) d\mu]^2}.$$

To show that $\dot{\lambda}(t)p(t) + \dot{\gamma}(t) \leq 0$, multiply (A7) by $p(t)$ and add it to (A8); after simplifying, we obtain the equivalent condition

$$(A9) \quad \left[\int (\lambda p(t) + \gamma) \dot{q}_{\lambda,\gamma}(t) d\mu \right] \left[\int q_{\lambda,\gamma}(t) d\mu \right] - \left[\int (\lambda p(t) + \gamma) q_{\lambda,\gamma}(t) d\mu \right] \left[\int \dot{q}_{\lambda,\gamma}(t) d\mu \right] \leq 0.$$

We need to show that (A9) holds for every $t > 0$. (Notice that t does not vary in this expression; t is fixed and integration is taken with respect to λ and γ .) We know from (A2) that $\dot{q}_{\lambda,\gamma}(t) = -(\lambda p(t) + \gamma) q_{\lambda,\gamma}(t)$ for every λ, γ , and t . Substituting this into (A9) we obtain

$$(A10) \quad \left[\int (\lambda p(t) + \gamma)^2 q_{\lambda,\gamma}(t) d\mu \right] \left[\int q_{\lambda,\gamma}(t) d\mu \right] \geq \left[\int (\lambda p(t) + \gamma) q_{\lambda,\gamma}(t) d\mu \right]^2.$$

Fix $t > 0$ and define the random variables

$$(A11) \quad X = (\lambda p(t) + \gamma) \sqrt{q_{\lambda,\gamma}(t)} \text{ and } Y = \sqrt{q_{\lambda,\gamma}(t)}.$$

The realizations of X and Y are determined by random draws from μ . Thus, (A10) follows directly from Schwarz's inequality: $E[X^2]E[Y^2] \geq (E[XY])^2$. This establishes the claim.

We can now apply this result to show that $H(t)$ is strictly decreasing in t for all $t > 0$. Differentiating $H(t)$, we obtain

$$(A12) \quad \dot{H}(t) = \dot{\lambda}(t) + \dot{\gamma}(t)/p(t) - \gamma(t)\dot{p}(t)/p^2(t).$$

By the claim above, $\dot{\lambda}(t)p(t) + \dot{\gamma}(t) \leq 0$, so division by $p(t) > 0$ yields $\dot{\lambda}(t) + \dot{\gamma}(t)/p(t) \leq 0$. Thus, the sum of the first two terms on the right-hand side of (A12) is nonpositive. But the last term is strictly negative, because $\gamma(t) > 0$ for all $t > 0$, given the initial condition $\bar{\gamma} = \gamma(0) > 0$. Hence, $\dot{H}(t) < 0$ as was to be shown.

PROOF OF PROPOSITION 2:

The equations of motion are

$$(A13) \quad \dot{p}_\lambda(t) = \lambda[F_\lambda(p(t)) - p_\lambda(t)],$$

where the initial conditions are $p_\lambda(0) = 0$.

By assumption, $F_\lambda(0) > 0$; hence $\dot{p}_\lambda(t) > 0$ for all sufficiently small $t > 0$ and

$$(A14) \quad \ddot{p}_\lambda(t)/\dot{p}_\lambda(t) = \lambda[f_\lambda(p(t)) - 1].$$

By hypothesis, the functions $f_\lambda(p(t))$ are strictly increasing on some common neighborhood of the origin $0 < p(t) < \bar{p}$. It follows that, for every λ and all t in a suitable interval $(0, T]$,

$$(A15) \quad \ddot{p}_\lambda(t)/\dot{p}_\lambda(t) - (\dot{p}_\lambda(t)/\dot{p}_\lambda(t))^2 > 0.$$

Hence, we have

$$(A16) \quad \sqrt{\ddot{p}_\lambda(t)\dot{p}_\lambda(t)} > \dot{p}_\lambda(t),$$

$$(A17) \quad \int \sqrt{\ddot{p}_\lambda(t)\dot{p}_\lambda(t)} d\nu > \int p_\lambda(t) d\nu = \dot{p}(t).$$

By Schwarz's inequality,

$$(A18) \quad \sqrt{\ddot{p}(t)\dot{p}(t)} = \left(\int \ddot{p}_\lambda(t) d\nu \int \dot{p}_\lambda(t) d\nu \right)^{1/2} \geq \int \sqrt{\ddot{p}_\lambda(t)\dot{p}_\lambda(t)} d\nu.$$

Combining this with (A17), we conclude that $\sqrt{\ddot{p}(t)\dot{p}(t)} > \ddot{p}(t)$, which implies that $(d[\ddot{p}(t)/\dot{p}(t)])/dt > 0$, that is, $\ddot{p}(t)/\dot{p}(t)$ is strictly increasing on $(0, T]$. This concludes the proof of Proposition 2.

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