STOCHASTIC PROCESSES

Fall 2017

Week 1

Solutions by

JINHONG DU

15338039

Show that Definition 2.1.1 of Poisson process implies Definition 2.1.2.

 \therefore $\{N(t), t \ge 0\}$ is a Poisson process by Definition 2.1.1

٠.

- (1) N(0) = 0;
- (2) The process has independent increments

(3)
$$\forall t > 0, s \geqslant 0, n \in \mathbb{N}, \mathbb{P}\{N(t+s) - N(s) = n\} = \frac{(\lambda t)^n}{n!}e^{-\lambda t}$$

- $(1^*) N(0) = 0;$
- (2*) From (2) the process has independent increments

From (3) we have $\forall t_2 > t_1 \ge 0$, the length of any interval $(t_1 + s, t_2 + s]$ is $\Delta t = t_2 - t_1$,

$$\mathbb{P}\{N(t_2+s) - N(t_1+s) = n\} = \frac{(\lambda \Delta t)^n}{n!} e^{-\lambda \Delta t}$$

i.e. the process has stationary increments

 $(3^*) \ \forall \ h > 0,$

$$\mathbb{P}\{N(h) = 1\} = \mathbb{P}\{N(h) - N(0) = 1\}$$

$$= \lambda h e^{-h\lambda}$$

$$= \lambda h \left(\sum_{k=0}^{\infty} \frac{(-\lambda h)^k}{k!}\right)$$

$$= \lambda h \left(1 - \lambda h + \lambda^2 h^2 - \cdots\right)$$

$$= \lambda h + o(h)$$

 $(4^*) \ \forall \ h > 0,$

$$\begin{split} \mathbb{P}\{N(h) \geqslant 2\} &= 1 - \mathbb{P}\{N(h) = 1\} - \mathbb{P}\{N(h) = 0\} \\ &= 1 - \mathbb{P}\{N(h) - N(0) = 1\} - \mathbb{P}\{N(h) - N(0) = 0\} \\ &= 1 - [\lambda h + o(h)] - \frac{(\lambda h)^0}{0!} e^{-\lambda h} \\ &= o(h) \end{split}$$

which satisfy Definition 2.1.2, i.e., Definition 2.1.1 of Poisson process implies Definition 2.1.2.

For another approach to proving that Definition 2.1.2 implies Definition 2.1.1.

(a) Prove, using Definition 2.1.2, that

$$P_0(t+s) = P_0(t)P_0(s)$$

 $\forall h > 0.$

$$\begin{split} P_0(t+s+h) &= \mathbb{P}\{N(t+s+h)=0\} \\ &= \mathbb{P}\{N(t+s+h)-N(t+h)=0, N(t+h)-N(t)=0, N(t)-N(0)=0\} \\ &= \mathbb{P}\{N(t+s+h)-N(t+h)=0\}\mathbb{P}\{N(t+h)-N(t)=0\} \\ &= \mathbb{P}\{N(t)-N(0)=0\} \\ &= P_0(s)[1-\lambda h+o(h)]P_0(t) \end{split}$$

 \therefore let $h \to 0+$, we have

$$P_0(t+s) = P_0(t)P_0(s)$$

(b) Use (a) to infer that the interarrival times $X_1, X_2 \cdots$ are independent exponential random variables with rate λ .

$$\therefore$$
 $\forall m, n \in \mathbb{N}^+, m > n, t > 0,$

$$\{X_m > t\} = \left\{ N \left(2t + \sum_{\substack{i=1\\i \neq n}}^{m-1} X_i \right) - N \left(t + \sum_{\substack{i=1\\i \neq n}}^{m-1} X_i \right) = 0 \right\}$$

and

$$\{X_n > t\} = \left\{ N\left(t + \sum_{i=1}^{n-1} X_i\right) - N\left(\sum_{i=1}^{n-1} X_i\right) = 0 \right\}$$

- \therefore by stationary and independent increments, X_m and X_n are independent identically distributed
- \therefore X_1, X_2, \cdots are independent identically distributed
- $\cdots \forall n \in \mathbb{N}^+, t > 0$

$$\mathbb{P}\{X_n > t\} = \mathbb{P}\left\{N\left(t + \sum_{k=1}^{n-1} X_k\right) - N\left(\sum_{k=1}^{n-1} X_k\right) = 0\right\}$$
$$= \mathbb{P}\{N(t) - N(0) = 0\}$$
$$= P_0(t)$$

 $\therefore \forall h > 0,$

$$\mathbb{P}\{X_n > t + h\} - \mathbb{P}\{X_n > t\} = P_0(t + h) - P_0(t)$$
$$= P_0(t)P_0(h) - P_0(t)$$
$$= P_0(t)[-\lambda h + o(h)]$$

٠.

$$P'_{0}(t) = \lim_{h \to 0+} \frac{P_{0}(t+h) - P_{0}(t)}{h}$$
$$= P_{0}(t)\lambda$$

Solution (cont.)

•.•

$$\begin{cases} P_0'(t) = P_0(t)\lambda \\ P_0(0) = 1 \end{cases}$$

٠.

$$P_0(t) = e^{-\lambda t}$$

i.e.

$$\mathbb{P}\{X_1 > t\} = e^{-\lambda t}$$

.:. $\forall n \in \mathbb{N}^+, X_n$ has exponential distribution with rate λ

Therefore, the interarrival times $X_1, X_2 \cdots$ are independent exponential random variables with rate λ .

(c) Use (b) to show that N(t) is Poisson distributed with mean λt .

Suppose that for $n \in \mathbb{N}$,

$$P_n(t) = \mathbb{P}\{N(t) = n\} = \frac{(\lambda t)^n}{n!} e^{-\lambda t} \tag{*}$$

Equation (*) holds when n = 1. Then for n + 1,

$$\begin{split} \mathbb{P}\{N(t) = n+1\} &= \int_0^t \mathbb{P}\{N(t) = n+1 | N(s) = n\} \mathbb{P}\{N(s) = n\} \mathrm{d}s \\ &= \int_0^t \mathbb{P}\{X_{n+1} = t - s\} \mathbb{P}\{N(s) = n\} \mathrm{d}s \\ &= \int_0^t \lambda e^{-\lambda(t-s)} \frac{(\lambda s)^n}{n!} e^{-\lambda s} \mathrm{d}s \\ &= \frac{\lambda^{n+1} e^{-\lambda t}}{n!} \int_0^t s^n \mathrm{d}s \\ &= \frac{(\lambda t)^{n+1}}{(n+1)!} e^{-\lambda t} \end{split}$$

Equation (*) holds for n+1. By induction, N(t) is Poisson distribution with mean λt .