

(1) The Poisson distribution has probability mass function

$$p(y_i|\theta) = \frac{\theta^{y_i} e^{-\theta}}{y_i!}, \quad \theta > 0, \quad y_i = 0, 1, \dots \quad (1)$$

and let  $y_1, \dots, y_n$  be random sample from this distribution.

1. Show that the gamma distribution  $\mathcal{G}(\alpha, \beta)$  is a conjugate prior distribution for the Poisson distribution.

**Solution:**

(a) The Gamma prior is given by

$$\pi(\theta|\alpha, \beta) \propto \theta^{\alpha-1} \cdot e^{-\theta/\beta} \quad (2)$$

The likelihood function is given by

$$f(y|\theta) \propto \theta^{\sum y_i} \cdot e^{-n\theta} \quad (3)$$

Multiplying these two equations gives

$$\pi(\theta|y) \propto \theta^{\alpha-1+\sum y_i} \cdot e^{-\theta(n/\beta)} \quad (4)$$

We see that

$$\theta|y \sim \mathcal{G}(\alpha', \beta') \quad \text{where } \alpha' = \alpha + \sum y_i, \quad \beta' = n + (1/\beta) \quad (5)$$

2. Show that  $\bar{y}$  is the MLE for  $\theta$ .

**Solution:**

(a) The likelihood function is given by

$$f(y|\theta) \propto \theta^{\sum y_i} \cdot e^{-n\theta} \quad (6)$$

Taking the log gives us

$$\log f(y|\theta) = \sum y_i \log(\theta) - n\theta - \log(\sum y_i) \quad (7)$$

Taking the derivative with respect to  $\theta$  gives

$$\log f(y|\theta)' = \frac{\sum y_i}{\theta} - n \quad (8)$$

Setting this equal to 0 and solving for  $\theta$  gives

$$\theta = \frac{\sum y_i}{n} = \bar{y} \quad (9)$$

$$Q.E.D \quad (10)$$

3. Write the mean of the posterior distribution as a weighted average of the mean of the prior distribution and the MLE.

Solution:

- (a) We want to find weights ( $w$  and  $(1-w)$ ) so that

$$E(\theta|y) = w \cdot E[0] + (1 - w) \cdot \bar{y} \quad (11)$$

Substituting answers from parts (1) and (2) gives

$$\frac{\alpha + \sum y_i}{\beta + n} = w \cdot (\alpha/\beta) + (1 - w) \cdot \frac{\sum y_i}{n} \quad (12)$$

Using WolframAlpha to solve for  $w$  gives

$$w = \frac{\beta}{\beta + n} \quad (13)$$

This means the posterior distribution can be expressed as

$$E(\theta|y) = \frac{\beta}{\beta + n} \cdot E[0] + \left(1 - \frac{\beta}{\beta + n}\right) \cdot \bar{y} \quad (14)$$

4. What happens to the weight on the prior mean as  $n$  becomes large?

As  $n$  becomes larger, the weight on the prior mean decreases. If  $n$  becomes sufficiently large, the weight will tend to 0.

(2) Consider the following two sets of data obtained after tossing a die 100 and 1000 times, respectively:

$n$	1	2	3	4	5	6
100	19	12	17	18	20	14
1000	190	120	170	180	200	140

Suppose you are interested in  $\theta_1$ , the probability of obtaining a one spot. Assume your prior for all the probabilities is a Dirichlet distribution, where each  $\alpha_i = 2$ . Compute the posterior distribution for  $\theta_1$  for each of the sample sizes in the table. Plot the resulting distribution and compare the results. Comment on the effect of having a larger sample.

**Solution:**

1. The prior distribution is expressed as

$$\text{Dirichlet}(2, 2, 2, 2, 2, 2) \quad (15)$$

The posterior distribution for  $n = 100$  is expressed as

$$\text{Dirichlet}(21, 14, 19, 20, 22, 16) \quad (16)$$

The posterior distribution for  $n = 1000$  is expressed as

$$\text{Dirichlet}(192, 122, 172, 182, 202, 142) \quad (17)$$

The marginal distribution can be expressed as  $\theta_1 \sim \text{Beta}(\alpha_1, \sum \alpha)$ , or specifically

$$\text{for } n = 100 : \theta_1 \sim \text{Beta}(21, 91) \quad (18)$$

$$\text{for } n = 1000 : \theta_1 \sim \text{Beta}(192, 820) \quad (19)$$

Plotting the graphs:

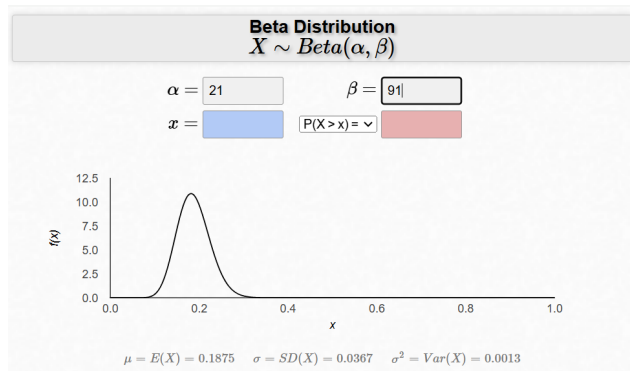


Figure 1: Beta(21,91)

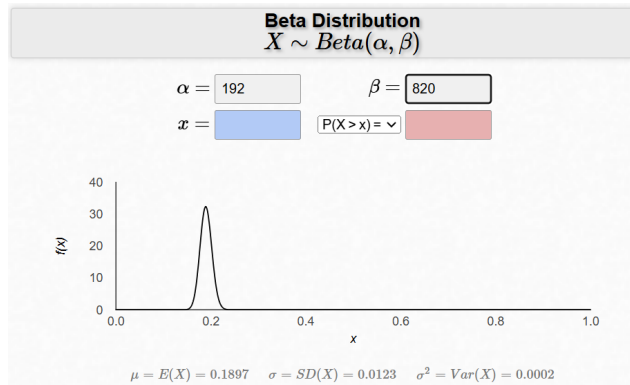


Figure 2: Beta(192,820)

Comparing these two figures, we clearly see the larger sample size gives a more accurate and robust estimation range. The estimation seems to converge closer to the actual parameter value. This also makes intuitive sense- more data results in more accurate estimations.