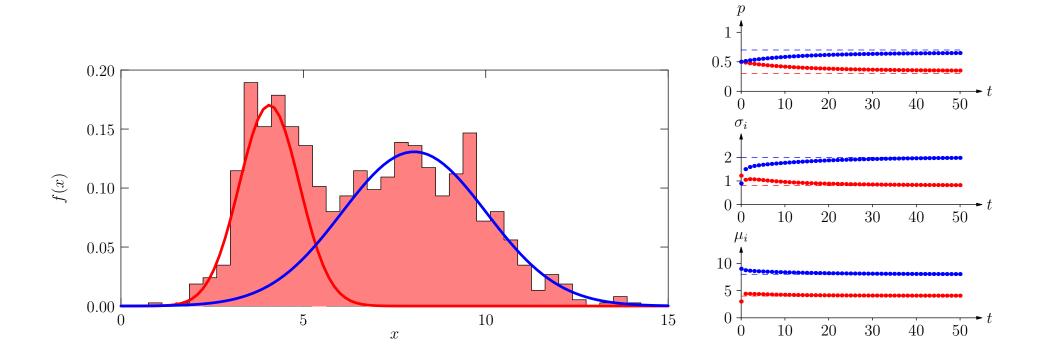
Advanced Machine Learning

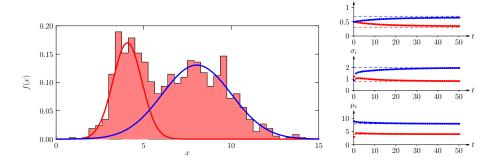
Probabilistic Inference



Hierarchical Models, Mixture of Gaussians, Expectation Maximisation

Outline

- 1. Building Probabilistic Models
- 2. Mixture of Gaussians
- 3. Expectation Maximisation



Building Probabilistic Models

- To describe a system with uncertainty we use random variables,
 X, Y, Z, etc.
- We use the convention of writing random variables in capitals (this is sometimes confusing as when you observe a random variables it is no longer random)
- The variables are described by probability mass function $\mathbb{P}(X,Y,Z)$ or if our variables are continuous, but probability densities $f_{X,Y,Z}(x,y,z)$.
- A major rule of probability is

$$\sum_{X} \mathbb{P}(X, Y, Z) = \mathbb{P}(Y, Z) \blacksquare$$

Conditional Probabilities

- When developing models it is often useful to consider conditional probabilities e.g. $\mathbb{P}(X,Y|Z)$ or $f_{X|Y,Z}(x|y,z)$.
- A second major rule in probabilistic modelling is

$$\mathbb{P}(X,Y) = \mathbb{P}(X|Y)\mathbb{P}(Y) = \mathbb{P}(Y|X)\mathbb{P}(X)$$

- This is a mathematical identity that does not imply causality (it defines conditional probability)
- It is the origins of Bayes' rule: $\mathbb{P}(X|Y) = \frac{\mathbb{P}(Y|X)\mathbb{P}(X)}{\mathbb{P}(Y)}$

Discriminative Models

- We often think of our observations as given and the predictions as random variables
- For example we might be given some features x and we wish to predict a class $C \in \mathcal{C}$
- ullet Our objective is then to find the probability $\mathbb{P}(C|oldsymbol{x})$
- This is known as a discriminative model
- E.g. in *foundations of machine learning* you learnt how to find the Bayes' optimal discrimination surface

Generative Models

- Sometimes it is easy to think about the joint process of generating the features and outputs together
- This leads to a joint distribution $\mathbb{P}(X,Y)$ where X are your features and Y is your output you are trying to predict!
- This is known as a generative model
- Generative models are often more natural to think about
- We can use them to do discrimination using

$$\mathbb{P}(Y|\boldsymbol{X}) = \frac{\mathbb{P}(\boldsymbol{X},Y)}{\mathbb{P}(\boldsymbol{X})} = \frac{\mathbb{P}(\boldsymbol{X},Y)}{\sum\limits_{Y} \mathbb{P}(\boldsymbol{X},Y)} \blacksquare$$

Latent Variables

- Sometimes we have models that involve random variables that we don't observe and we don't care about
- These are called latent variables
- ullet If we have a latent variable Z and observed variable X and we are predicting a variable Y then we would ullet are over the latent variable

$$\mathbb{P}(\boldsymbol{X},Y) = \sum_{Z} \mathbb{P}(\boldsymbol{X},Y,Z)$$

Modelling Virus

- Suppose we want to estimate the number of hospitalisation from Corona virus in the next month.
- Our observable is the number of reported cases
- In our model we might want to estimate the number of actual cases
- This would be a latent variable (it is not an observable or our final target, but it is very useful intermediate in our model)
- This will be a random variable (we are uncertain, but we can build a probabilistic model giving a distribution of number of actual cases)

Hierarchical Models

- Of course, if I was really modelling the spread of a disease I would care about the probability, f(C|A,V), of catching the disease, C, given the persons age A and the variant of the disease V
- I would want to know the distribution of ages f(A) and try to infer the probability of different variants $\mathbb{P}(V)$
- I would care about the probability, f(R|A,V), of cases being reported given age and variant
- ullet And the probability, f(H|A,V), of hospitalisation given A and V
- This would involve an elaborate (hierarchical) model with a large number of latent variables

Probabilistic Inference

ullet We can use Bayes' rules to learn a set of parameter ullet that occur in our likelihood function

$$\mathbb{P}(\mathbf{\Theta}|\mathcal{D}) = rac{\mathbb{P}(\mathcal{D}|\mathbf{\Theta})\,\mathbb{P}(\mathbf{\Theta})}{\mathbb{P}(\mathcal{D})}$$

- This provides us a full probabilistic description of the parameters
- It doesn't overfit (we are not choosing the best)
- Bayesian inference provides a description of its own uncertainty
- We need to specify a likelihood and prior, but this is usually not difficult

Problem with Bayes

- Bayes is problematic because it is often hard
- The posterior is often not expressible as a nice probability function
- ullet We need to compute the evidence or $margin\ likelihood$ we use

$$\mathbb{P}(\mathcal{D}) = \sum_{oldsymbol{\Theta}} \mathbb{P}(\mathcal{D}|oldsymbol{\Theta}) \mathbb{P}(oldsymbol{\Theta})$$

- ullet But sometimes the number of values that ullet can take are so large that we cannot easily compute this
- Nevertheless we can usually do this using Monte Carlo techniques

Maximum A Posteriori (MAP) Solution

One work around is to compute the mode of the posterior

$$\mathbf{\Theta}_{\mathsf{MAP}} = \operatorname*{argmax} f(\mathcal{D}|\mathbf{\Theta}) f(\mathbf{\Theta}) = \operatorname*{argmax} \log(f(\mathcal{D}|\mathbf{\Theta})) + \log(f(\mathbf{\Theta})) \mathbf{I}$$

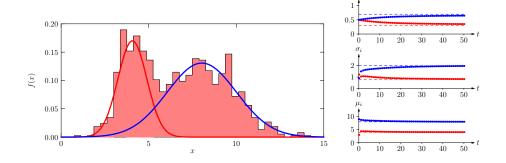
- We don't need to calculate $f(\mathcal{D})$ or explicitly calculate the posterior distribution
- But it is not Bayesian (despite what you are sometime told)—its
 not properly probabilistic
- You can overfit and you don't get an estimate of the error in your inference

Maximum Likelihood

- When we assume a uniform prior then the MAP solution is just maximising the likelihood
- Weirdly this hack was accepted as part of mainstream statistics even when Bayesian statistics was considered unscientific
- Maximum likelihood is often sufficient for government work, but it isn't the best you can do!
- In high-dimensional problems using a non-uniform prior can make a big difference
- And, of course, doing a full probabilistic calculation has real advantages

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- Building Probabilistic Models
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Mixture of Gaussians

- Suppose we were observing the decays from two types of short-lived particle, A or B^{\blacksquare}
- We observe the half life, X_i , but not the particle type
- We assume X_i is normally distributed with unknown means and variances: $\mathbf{\Theta} = \{\mu_A, \, \sigma_A^2, \, \mu_B, \, \sigma_B^2\}$
- Let $Z_i \in \{0,1\}$ be an indicator that particle i is of type A
- The probability of X_i is given by

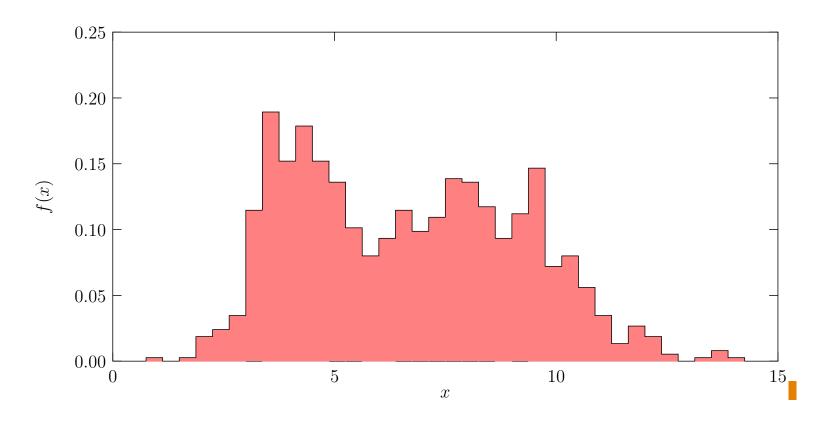
$$f(X_i|Z_i,\mathbf{\Theta}) = Z_i \mathcal{N}(X_i|\mu_A, \sigma_A^2) + (1 - Z_i) \mathcal{N}(X_i|\mu_B, \sigma_B^2) \blacksquare$$

Data

Note that

$$f(X_i|\mathbf{\Theta}) = \sum_{Z_i \in \{0,1\}} f(X_i, Z_i|\mathbf{\Theta}) = \sum_{Z_i \in \{0,1\}} f(X_i|Z_i, \mathbf{\Theta}) \mathbb{P}(Z_i)$$

$$= \mathbb{E}_{Z_i} [f(X_i|Z_i, \mathbf{\Theta})] = p \mathcal{N} (X_i | \mu_A, \sigma_A^2) + (1-p) \mathcal{N} (X_i | \mu_B, \sigma_B^2)$$



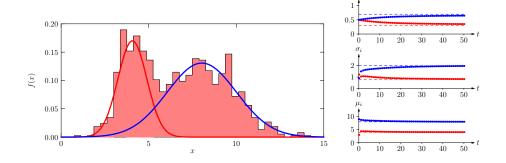
Maximum Likelihood

- To solve the model as a Bayesian we would have to assign priors to our parameters $\mathbf{\Theta} = (\mu_A, \sigma_A, \mu_B, \sigma_B, p)$
- This is doable, but complicated (we would also end up with a distribution for our parameters)
- Often we only want a reasonable estimate for some of our parameters (e.g. the half-lives μ_A and μ_B)
- A reasonable approach is to seek those parameters that maximise the likelihood of our observed data

$$f(\mathcal{D}|\mathbf{\Theta}) = \prod_{X \in \mathcal{D}} f(X|\mathbf{\Theta})$$

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Maximum Likelihood with Latent Variables

- The maximum likelihood is a non-linear function of the parameters so cannot be immediately maximised.
- If we knew which type of particle a data-point belongs to (Z_i) then it would be straightforward to maximise the likelihood
- As we don't we need to estimate $\mathbb{P}(Z_i=1)$, but this depends on μ_A , σ_A^2 , μ_B , σ_B^2 and p!
- We could use a standard optimiser, but this is slightly inelegant

EM Algorithm

- Instead we can use an expectation-maximisation algorithm usually known as an EM algorithm
- We proceed iteratively by maximising the expected log-likelihood with respect to the current set of parameters

$$\boldsymbol{\Theta}^{(t+1)} = \underset{\boldsymbol{\Theta}}{\operatorname{argmax}} \sum_{\boldsymbol{Z}} \mathbb{P} \left(\boldsymbol{Z} \big| \mathcal{D}, \boldsymbol{\Theta}^{(t)} \right) \log(f(\mathcal{D}|\boldsymbol{Z}, \boldsymbol{\Theta})) \mathbf{I}$$

It isn't obvious why this works

Why EM Algorithm Works

- The argument around why this works is quite involved
- Note that at each step we maximise

$$Q(\boldsymbol{\Theta}|\boldsymbol{\Theta}^{(t)}) = \sum_{\boldsymbol{Z} \in \{0,1\}^m} \mathbb{P}\Big(\boldsymbol{Z}\big|\mathcal{D}, \boldsymbol{\Theta}^{(t)}\Big) \log(f(\mathcal{D}|\boldsymbol{Z}, \boldsymbol{\Theta})) \mathbf{I}$$

• We can show that the maximum, $\mathbf{\Theta}^{(t+1)}$, is such that

$$\log \Big(f(\mathcal{D}|\mathbf{\Theta}^{(t+1)}) \Big) - \log \Big(f(\mathcal{D}|\mathbf{\Theta}^{(t)}) \Big) \ge Q(\mathbf{\Theta}^{(t+1)}|\mathbf{\Theta}^{(t)}) - Q(\mathbf{\Theta}^{(t)}|\mathbf{\Theta}^{(t)}) \ge 0$$

The details are given in the supplemental notes

Conditional Latent Variables

- We need to compute the distribution of latent variables conditioned on the data and current estimated parameters
- For our problem

$$\mathbb{P}\Big(oldsymbol{Z}ig|\mathcal{D},oldsymbol{\Theta}^{(t)}\Big) = \prod_{i=1}^{m} \mathbb{P}\Big(Z_iig|X_i,oldsymbol{\Theta}^{(t)}\Big)$$

where

$$\mathbb{P}\left(Z_{i}=1\big|X_{i},\boldsymbol{\Theta}^{(t)}\right) = \frac{p^{(t)}\mathcal{N}\left(X_{i}\big|\mu_{A}^{(t)},\sigma_{A}^{2(t)}\right)}{p^{(t)}\mathcal{N}\left(X_{i}\big|\mu_{A}^{(t)},\sigma_{A}^{2(t)}\right) + (1-p^{(t)})\mathcal{N}\left(X_{i}\big|\mu_{B}^{(t)},\sigma_{B}^{2(t)}\right)}$$

$$\mathbb{P}\left(Z_{i}=0\big|X_{i},\boldsymbol{\Theta}^{(t)}\right) = 1 - \mathbb{P}\left(Z_{i}=1\big|X_{i},\boldsymbol{\Theta}^{(t)}\right) \blacksquare$$

EM for Mixture of Gaussians

• Maximise with respect to parameters heta

$$Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)}) = \sum_{\boldsymbol{Z}} \mathbb{P}\left(\boldsymbol{Z}|\mathcal{D},\boldsymbol{\Theta}^{(t)}\right) \log(f(\mathcal{D}|\boldsymbol{Z},\boldsymbol{\Theta})) = \sum_{i=1}^{m} \sum_{Z_i} \mathbb{P}\left(Z_i|\mathcal{D},\boldsymbol{\Theta}^{(t)}\right) \log(f(X_i|Z_i,\boldsymbol{\Theta})) = \sum_{i=1}^{m} \sum_{Z_i \in \{0,1\}} \mathbb{P}\left(Z_i|X_i,\boldsymbol{\theta}_i^{(t)}\right) \left(Z_i\log(p) + (1-Z_i)\log(1-p)\right) - \frac{(X_i - \mu_{Z_i})^2}{2\sigma_{Z_i}^2} - \log\left(\sqrt{2\pi}\sigma_{Z_i}\right)$$

Compute update equations

$$\frac{\partial Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})}{\partial \mu_k} = 0, \qquad \frac{\partial Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})}{\partial \sigma_k} = 0, \qquad \frac{\partial Q(\boldsymbol{\theta}|\boldsymbol{\theta}^{(t)})}{\partial p} = 0$$

Update Equations

Means

$$\mu_{Z_i}^{(t+1)} = \frac{\sum_{i=1}^n \mathbb{P}(Z_i|X_i,\boldsymbol{\theta}^{(t)})X_i}{\sum_{i=1}^n \mathbb{P}(Z_i|X_i,\boldsymbol{\theta}^{(t)})}, \blacksquare$$

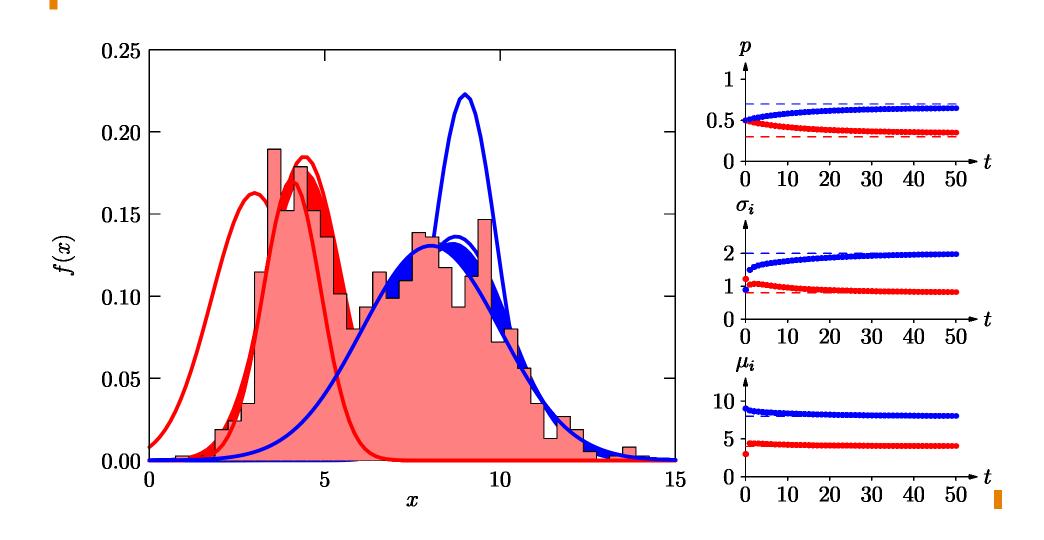
Variances

$$(\sigma_{Z_i}^{(t+1)})^2 = \frac{\sum_{i=1}^n \mathbb{P}(Z_i|X_i,\boldsymbol{\theta}^{(t)})(X_i - \mu_{Z_i}^{(t+1)})^2}{\sum_{i=1}^n \mathbb{P}(Z_i|X_i,\boldsymbol{\theta}^{(t)})}$$

Probability of being type 1

$$p^{(t+1)} = \frac{1}{n} \sum_{i=1}^{n} \mathbb{P}\left(Z_i = 1 \mid X_i, \boldsymbol{\theta}_i^{(t)}\right)$$

Example



Summary

- Building probabilistic models is an intricate process
- Identifying random variables that describe the system is the first step!
- Often we need to introduce variables that we don't observe and need to be marginalised out
- The EM algorithm provide one approach to maximising likelihoods or MAP solutions when we have latent variables
- It often gives nice update equations, but convergence can be slow!