Advanced Machine Learning



Bayes, Conjugate Priors, Uninformative Priors

Outline

- 1. Bayes' Rule
- 2. Conjugate Priors
- 3. Uninformative Priors



- In machine learning we are attempting to make inference under uncertainty
- The natural language for discussing uncertainty is probability
- The natural framework for making inferences is Bayesian statistics
- However, this requires that we encode our prior knowledge of the problem and specify a likelihood
- In consequence, probabilistic methods tend to be bespoke, rather then general purpose black boxes

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- $\star \mathbb{P}(\mathcal{H}_i|\mathcal{D})$ is the **posterior** probability of a hypothesis \mathcal{H}_i (i.e. the probability of \mathcal{H}_i after we see the data)
- $\star \mathbb{P}(\mathcal{D}|\mathcal{H}_i)$ is the **likelihood** of the data given the hypothesis. Note, that we calculated this from the forward problem
- $\star \mathbb{P}(\mathcal{H}_i)$ is the **prior** probability (i.e. the probability of \mathcal{H}_i before we see the data)
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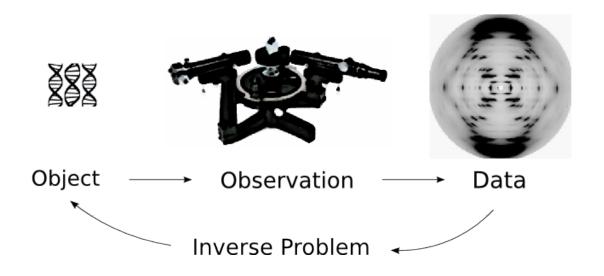
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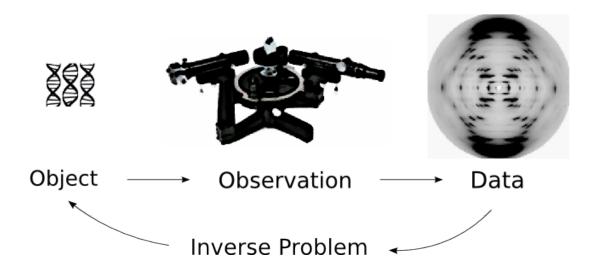
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- We consider the process of how the data is generated
- This uses the data we have (doesn't care about missing data)
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The normalisation term

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- It is called the marginal likelihood
- If we have two models M_1 and M_2 we can do **model selection** by choosing the model with the largest evidence $\mathbb{P}(\mathcal{D} \mid M_1)$ or $\mathbb{P}(\mathcal{D} \mid M_2)$
- This also allows us to select hyperparameters for a model

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Probability Density

 When we are working with continuous variables it is more natural to work with probability densities

$$f_X(x) = \lim_{\delta x \to 0} \frac{\mathbb{P}(x \le X < x + \delta x)}{\delta x}$$

- Note that densities are non-negative, but can be greater than 1 (they are not probabilities)
- However

$$\mathbb{P}(a \le X \le b) = \int_{a}^{b} f_X(x) \, \mathrm{d}x$$

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Bayes' rule also applies to densities

$$\mathbb{P}(x \le X < x + \delta x | Y) = \frac{\mathbb{P}(Y|x)\mathbb{P}(x \le X < x + \delta x)}{\mathbb{P}(Y)}$$

• Dividing by δx and taking the limit $\delta x \to 0$

$$f_{X|Y}(x|Y) = \frac{\mathbb{P}(Y|x) f_X(x)}{\mathbb{P}(Y)}$$

ullet Similarly if X is discrete and Y continuous

$$\mathbb{P}(X|y) = \frac{f_{Y|X}(y|X)\mathbb{P}(X)}{f_Y(y)}$$

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Practical Bayesian Inference

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$$p(\boldsymbol{\theta}|\mathcal{D}) = \frac{p(\mathcal{D}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathcal{D})}$$

- This can be hard for large data sets as the posterior, $p(\theta|\mathcal{D})$, is often a mess
- If we are lucky and have a simple likelihood then if we choose the right prior we end up with a posterior of the same form as the prior
- This occurs in some classic probabilistic inference problems, but as we will see soon it is also true for Gaussian Processes

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- Suppose we have a coin and we want to establish the probability of a head
- We want to learn this from a series of independent trials
- (Independent trials with two possible outcomes are known in probability theory as Bernoulli trials)
- Let X_i equal 1 if the i^{th} trial is a head and 0 otherwise
- If the probability of a head is p then the **likelihood** of a X_i is

$$\mathbb{P}(X_i|p) = p^{X_i}(1-p)^{1-X_i}$$

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$$\mathbb{P}(X_i|p) = p^{X_i}(1-p)^{1-X_i} = \begin{cases} p & \text{if } X_i = 1\\ (1-p) & \text{if } X_i = 0 \end{cases}$$

- We may have a prior belief (e.g. we have made a few trials or we see the coin looks like a normal penny)
- We will suppose we can model our prior belief in terms of a Beta distribution

$$f(p) = \text{Beta}(p|a,b) = \frac{p^{a-1}(1-p)^{b-1}}{B(a,b)}$$

• B(a,b) is just a normalisation constant

$$B(a,b) = \int_0^1 p^{a-1} (1-p)^{b-1} dp = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

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$$f(p) = \text{Beta}(p|a,b) = \frac{p^{a-1}(1-p)^{b-1}}{B(a,b)}$$

• B(a,b) is just a normalisation constant

$$B(a,b) = \int_0^1 p^{a-1} (1-p)^{b-1} dp = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

- Suppose we have no idea about p what should we do?
- Laplace (one of the first Bayesian's) suggested giving equal weighting to all values of \boldsymbol{p}
- This corresponds to a beta distribution with a=b=1
- (Surprisingly other arguments suggest using a=b=0 which provides a strong bias towards p=0 and p=1)
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Using Bayes' rule

$$f(p|\mathcal{D}) = \frac{\mathbb{P}(\mathcal{D}|p) f(p)}{\mathbb{P}(\mathcal{D})}$$

 Assuming the trials are independent (a reasonably fair assumption for tossing coins) then the likelihood factorises

$$\mathbb{P}(\mathcal{D}|p) = \prod_{i=1}^{n} p^{X_i} (1-p)^{1-X_i}$$

$$= p^{X_1} (1-p)^{1-X_1} p^{X_2} (1-p)^{1-X_2} \cdots p^{X_n} (1-p)^{1-X_n}$$

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Posterior

• Plugging in a prior $f(p) = \text{Beta}(p|a_0,b_0)$

$$f(p|\mathcal{D}) = \frac{\mathbb{P}(\mathcal{D}|p) f(p)}{\mathbb{P}(\mathcal{D})} = \frac{p^s (1-p)^{n-s} \times p^{a_0-1} (1-p)^{b_0-1}}{\mathbb{P}(\mathcal{D}) B(a_0, b_0)}$$

The denominator is a normalising factor

$$\mathbb{P}(\mathcal{D}) = \int_0^1 \mathbb{P}(\mathcal{D}|p) f(p) dp = \int_0^1 \frac{p^{s+a_0-1} (1-p)^{n-s+b_0-1}}{B(a_0,b_0)} dp$$
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- Something rather nice happened
- Starting with a beta distributed prior $f(p) = \text{Beta}(p|a_0,b_0)$ for a set of Bernoulli trials we obtain a beta distributed posterior $f(p|\mathcal{D}) = \text{Beta}(p|a_0 + s, b_0 + n s)$
- This is not always the case (often the posterior will be very complicated) but it happens for a few likelihoods and priors
- When the posterior is the same as the prior then the likelihood and prior distributions are said to be conjugate

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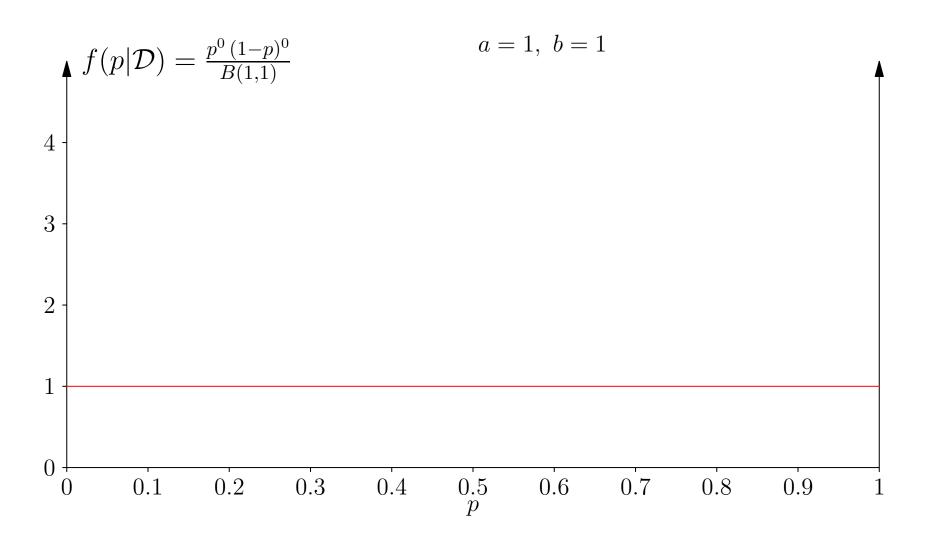
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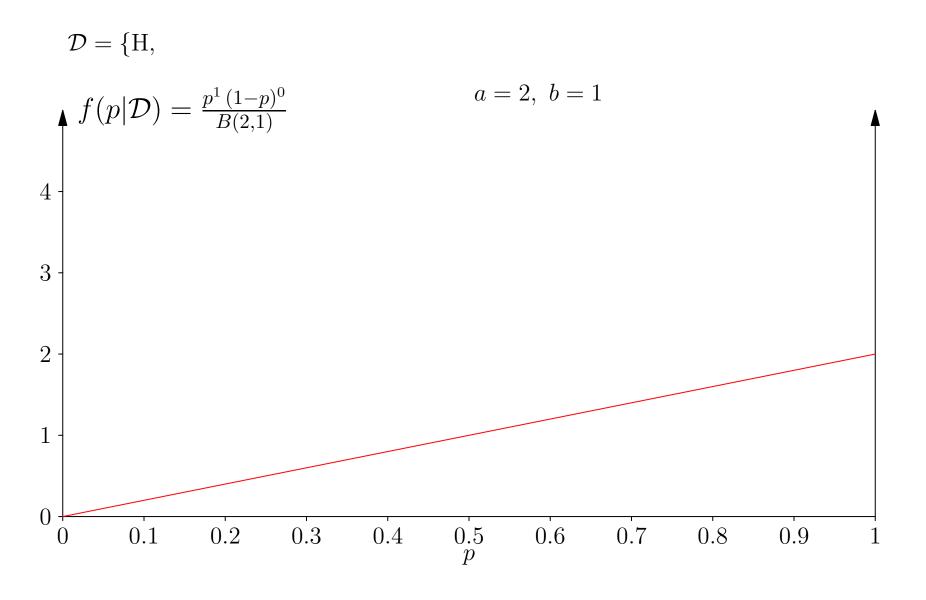
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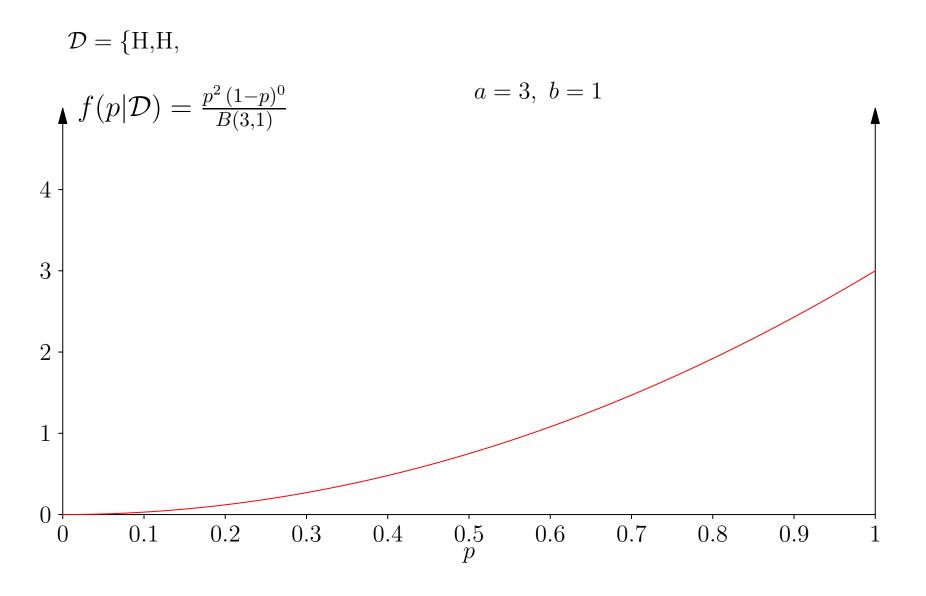
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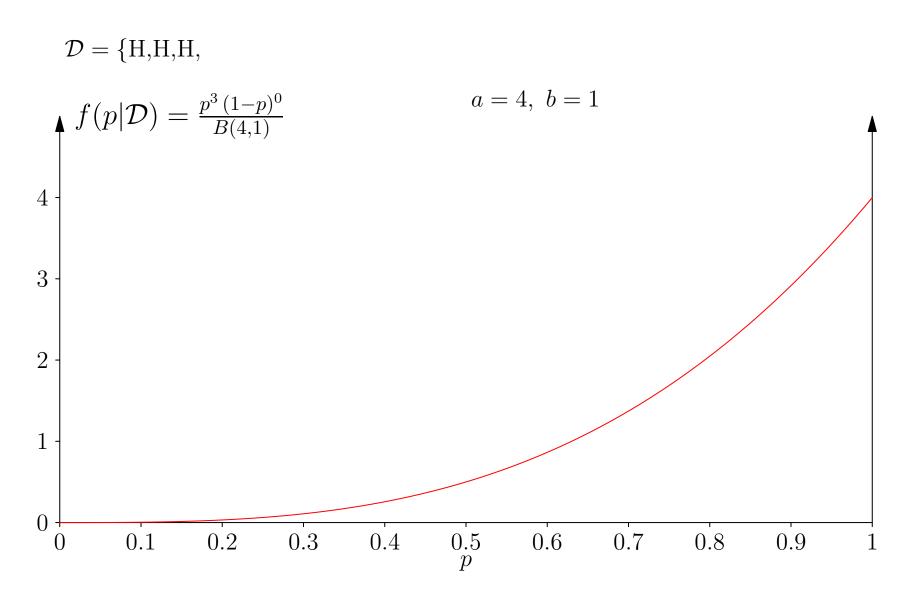
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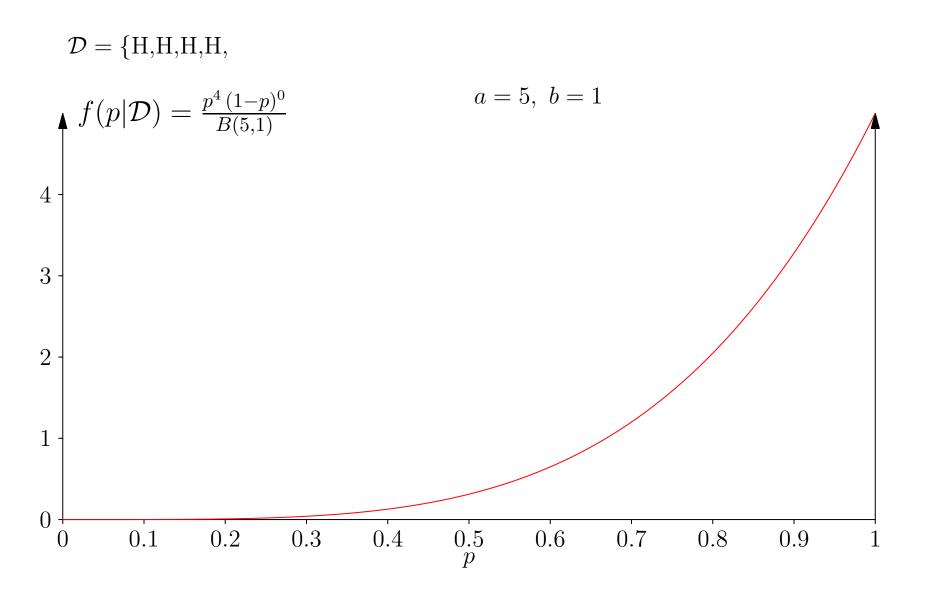
Example (p=0.7)

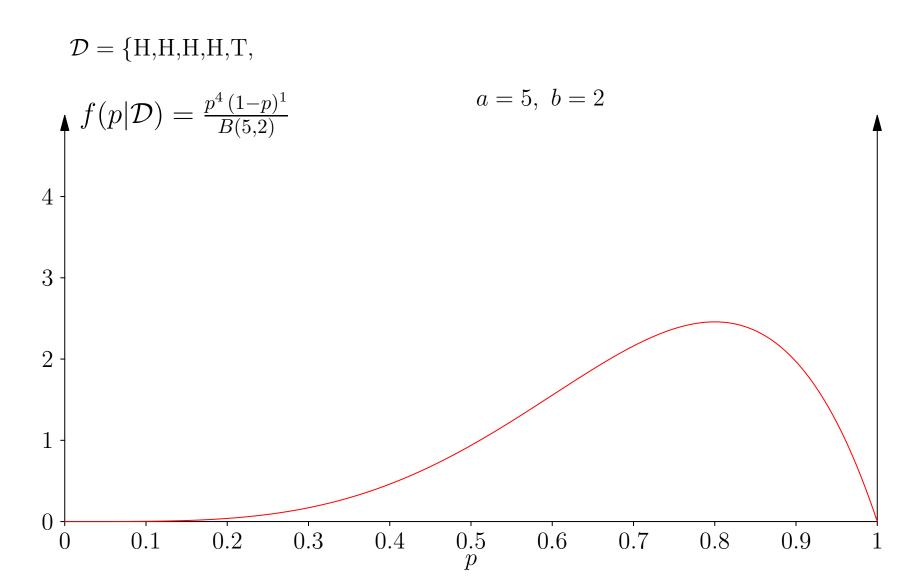




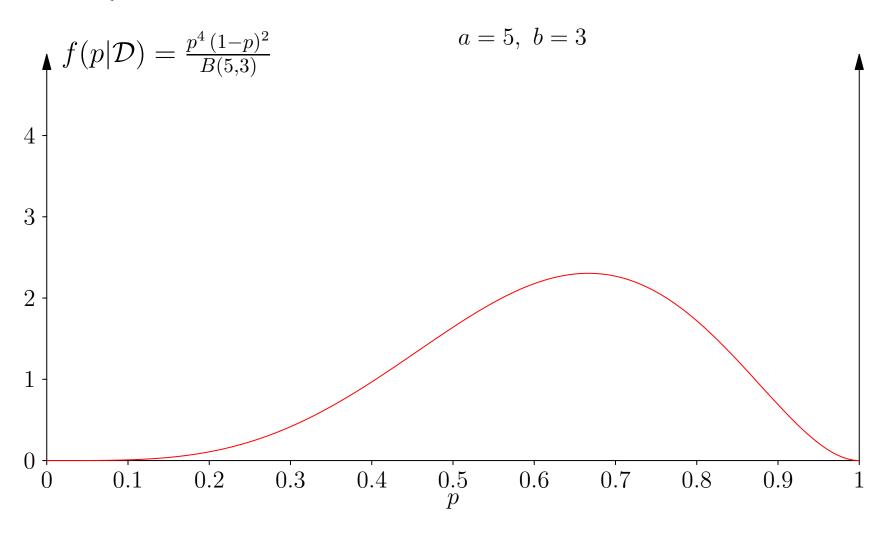


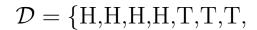


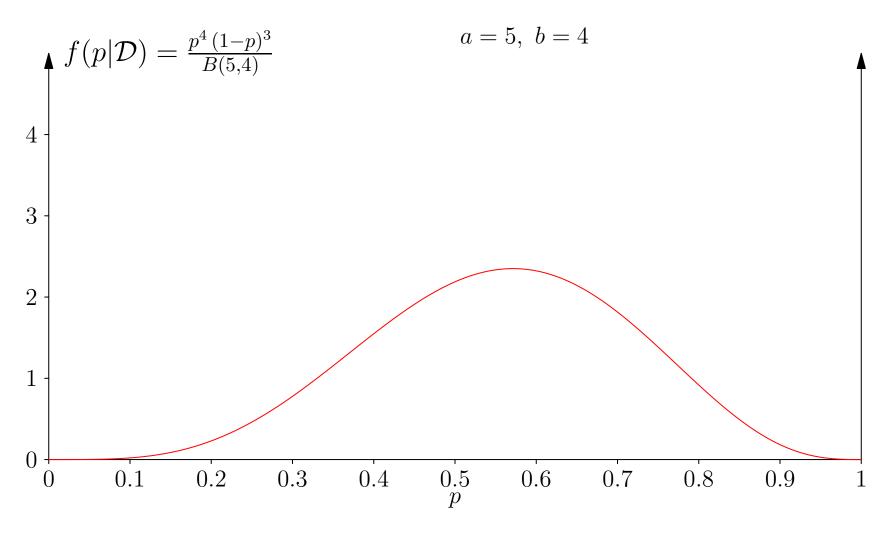




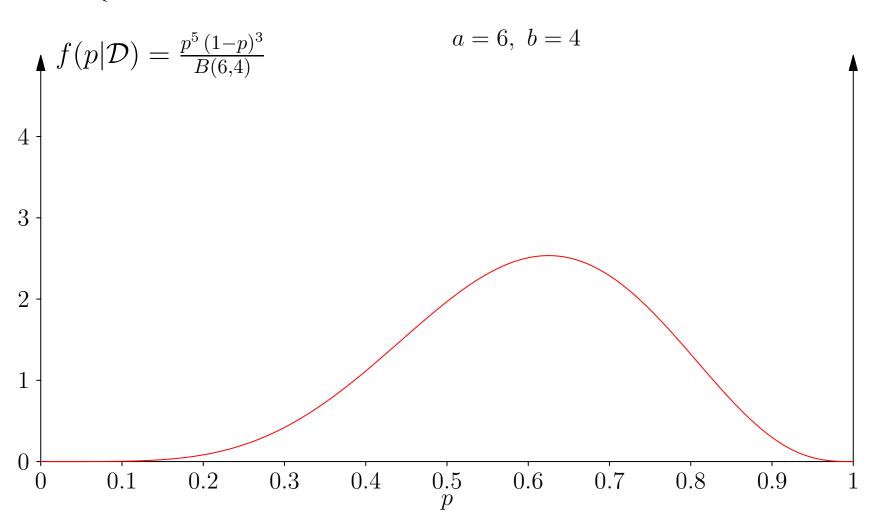




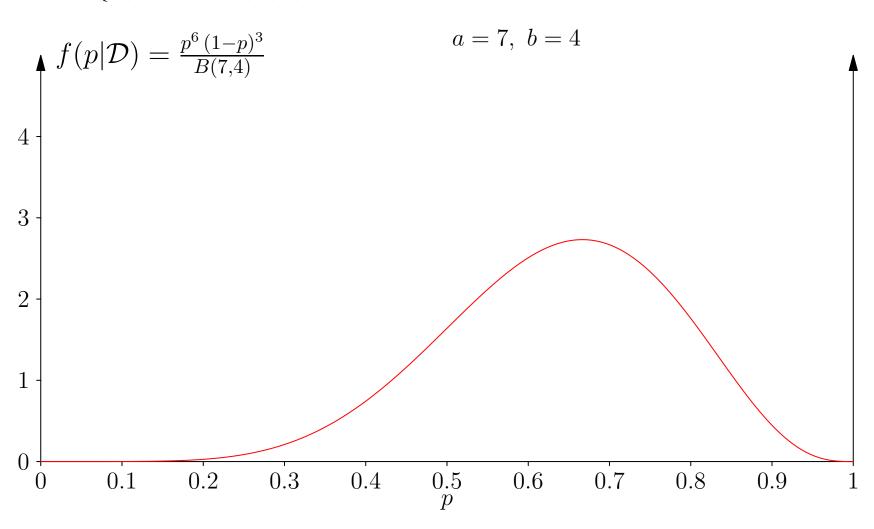


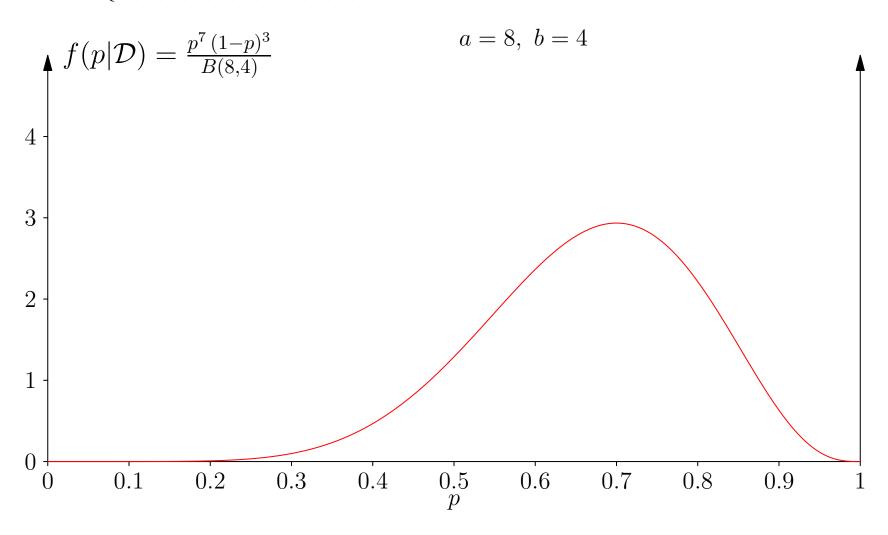


 $\mathcal{D} = \{\text{H,H,H,H,T,T,T,H},$

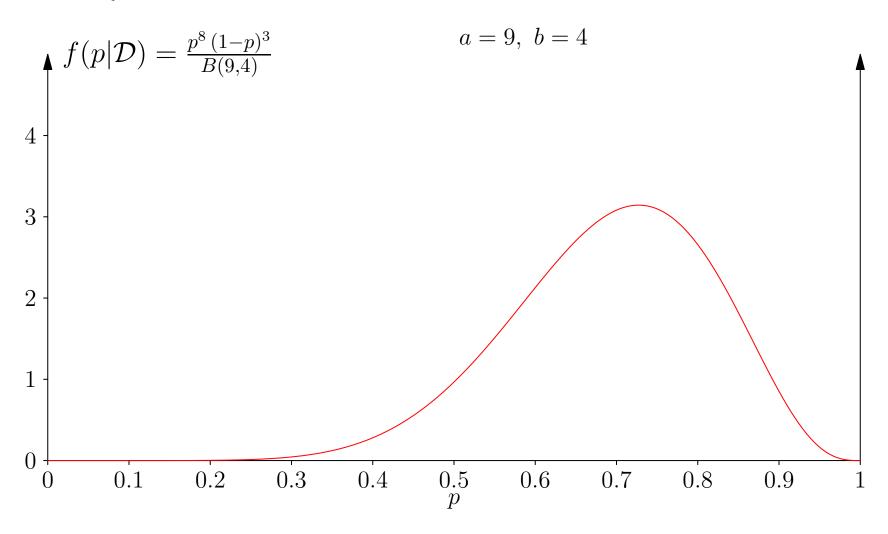


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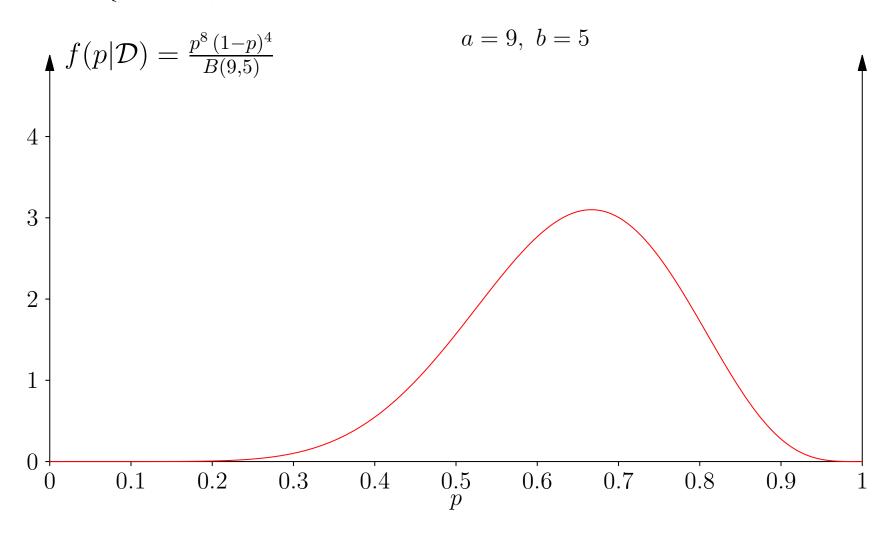




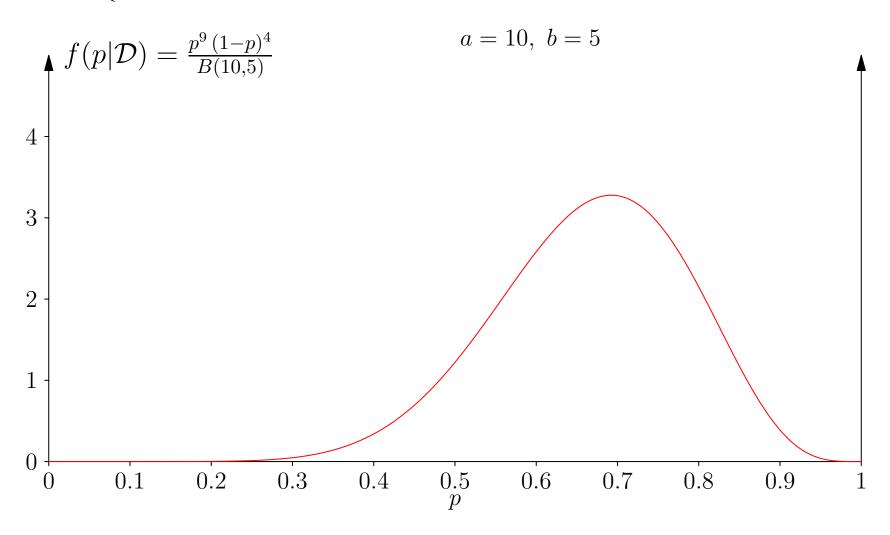
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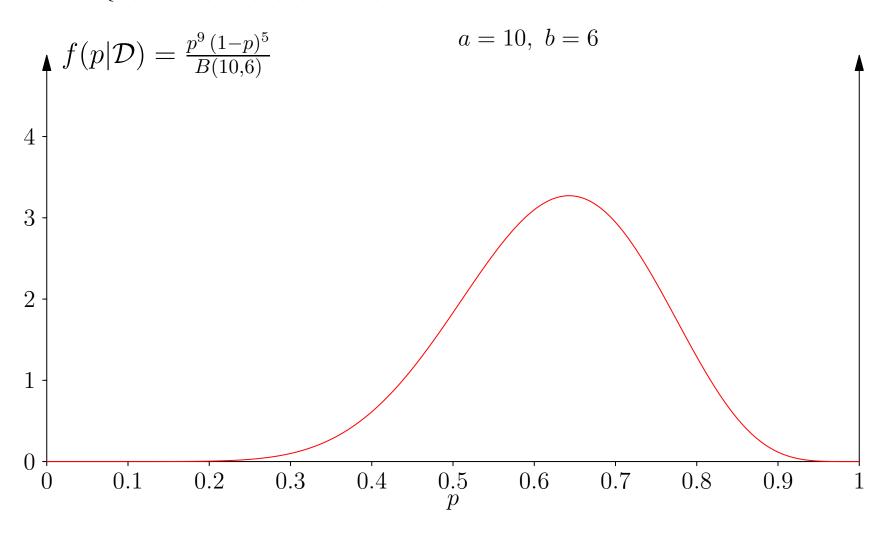


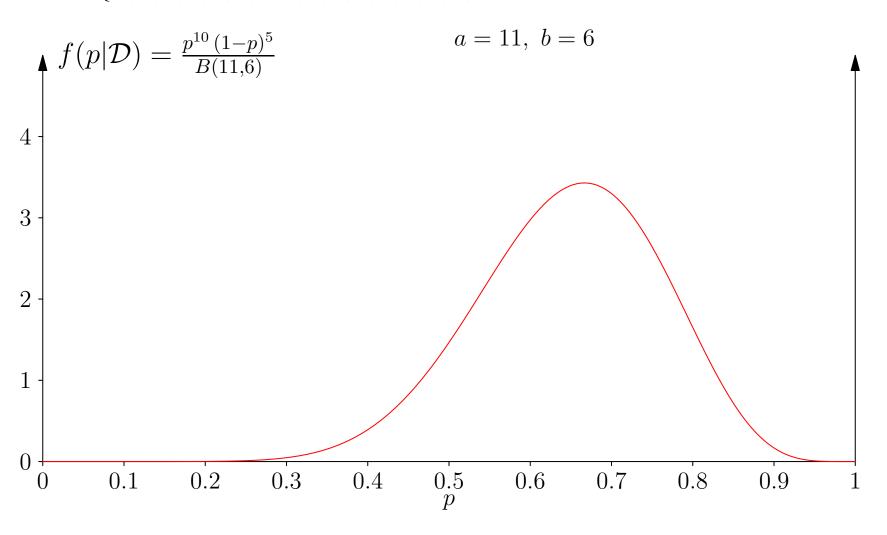
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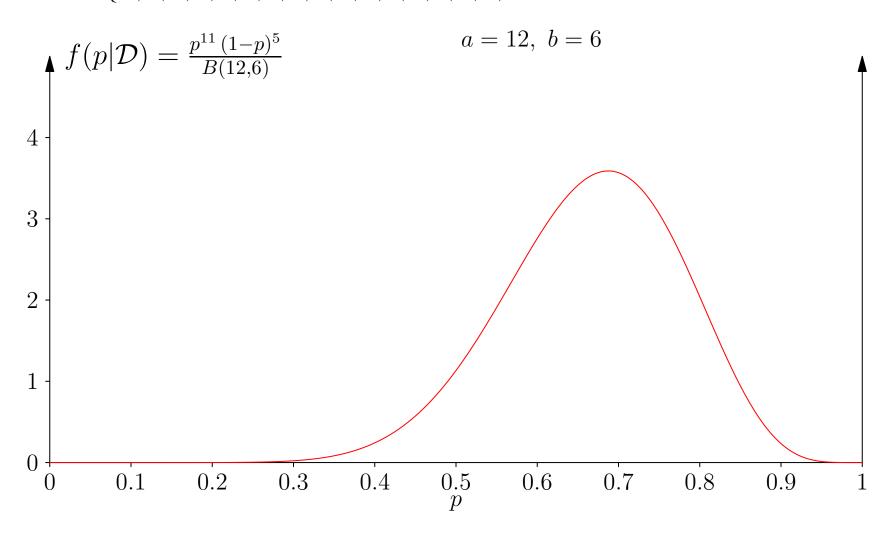
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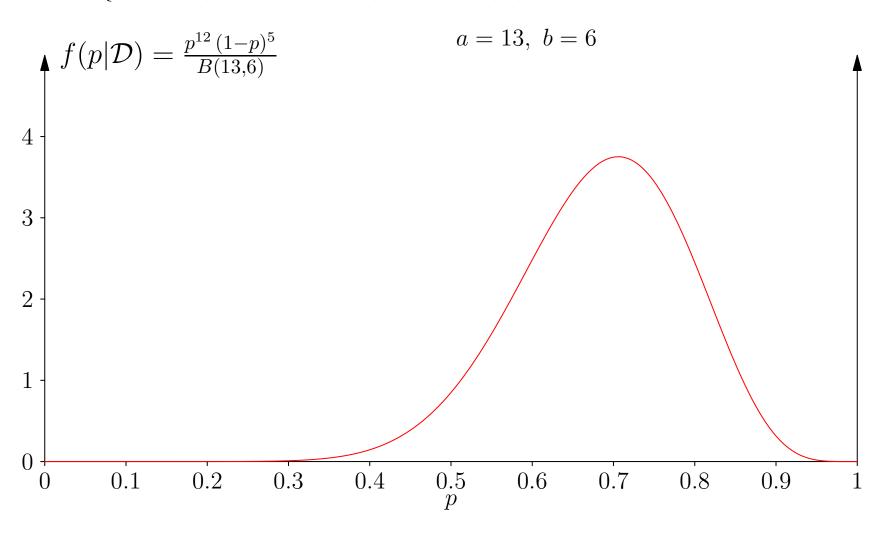


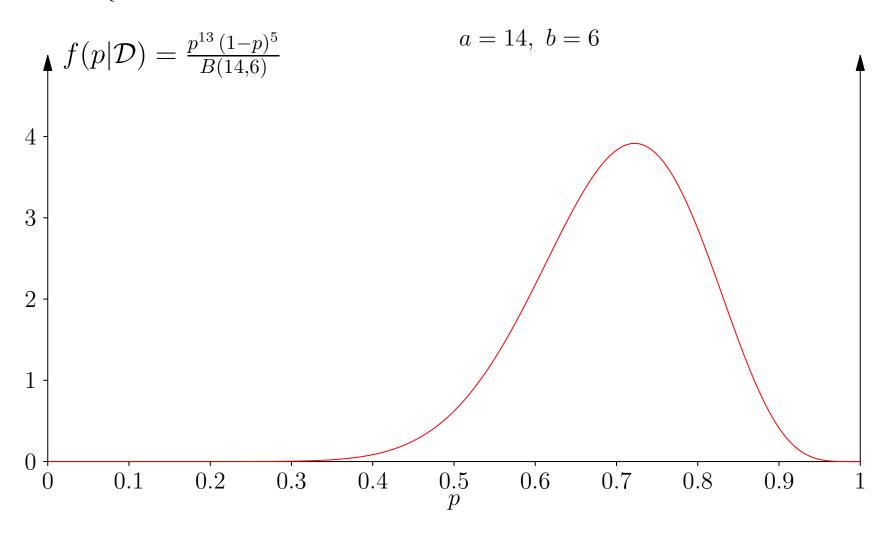


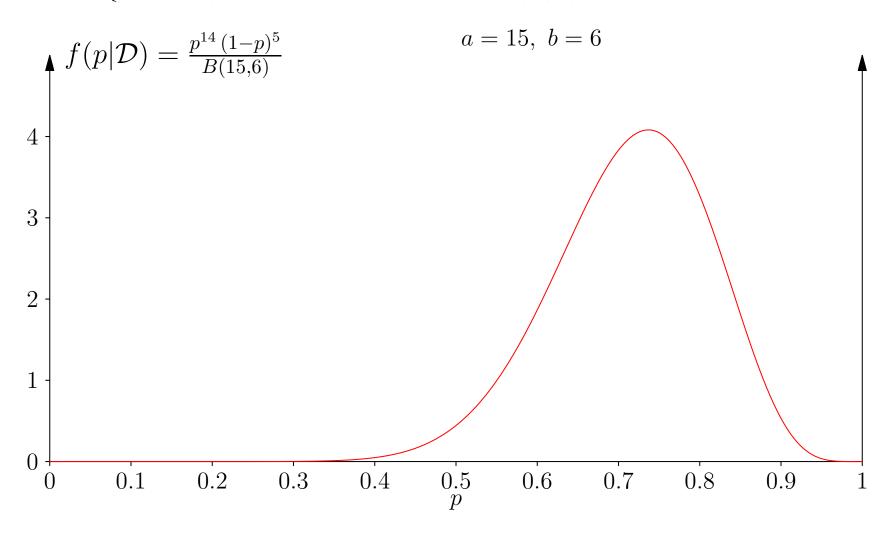


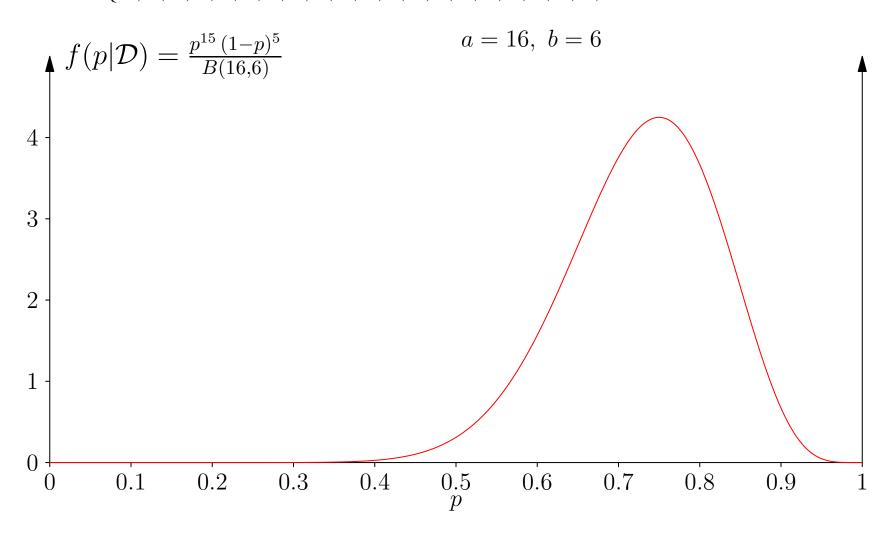
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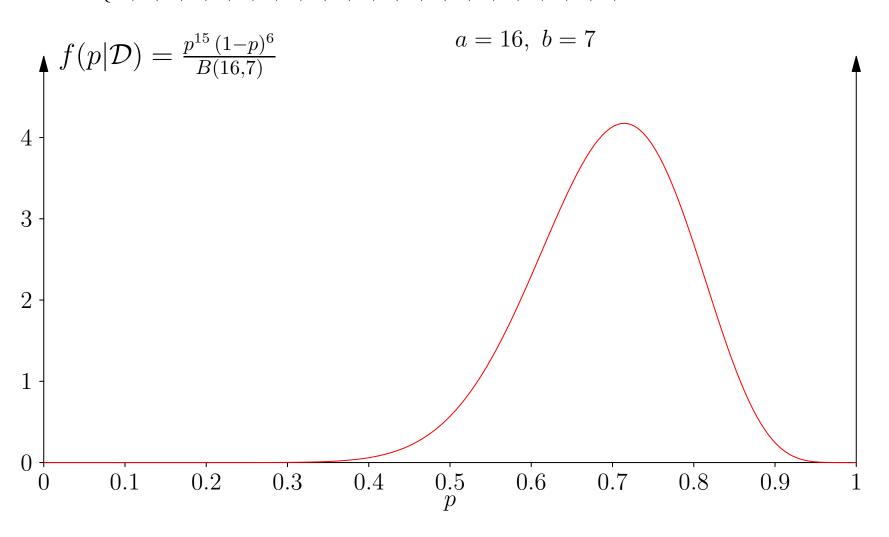


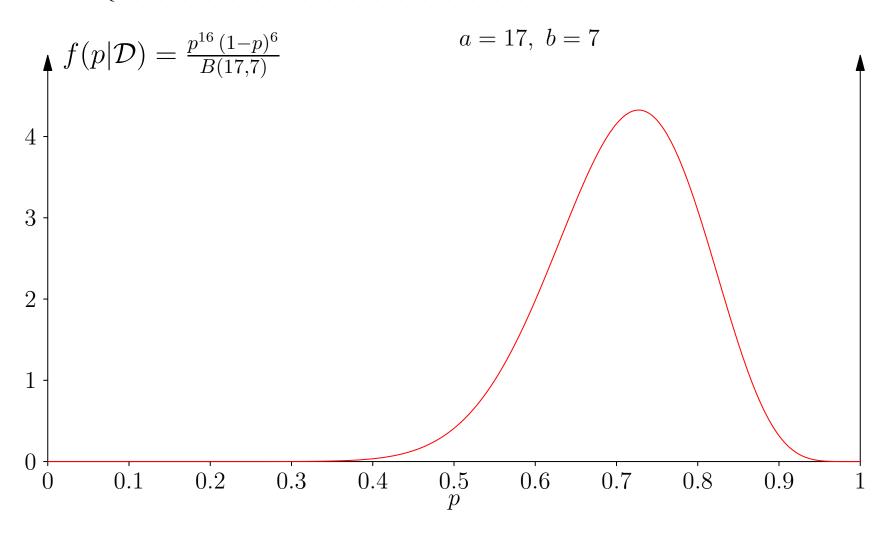


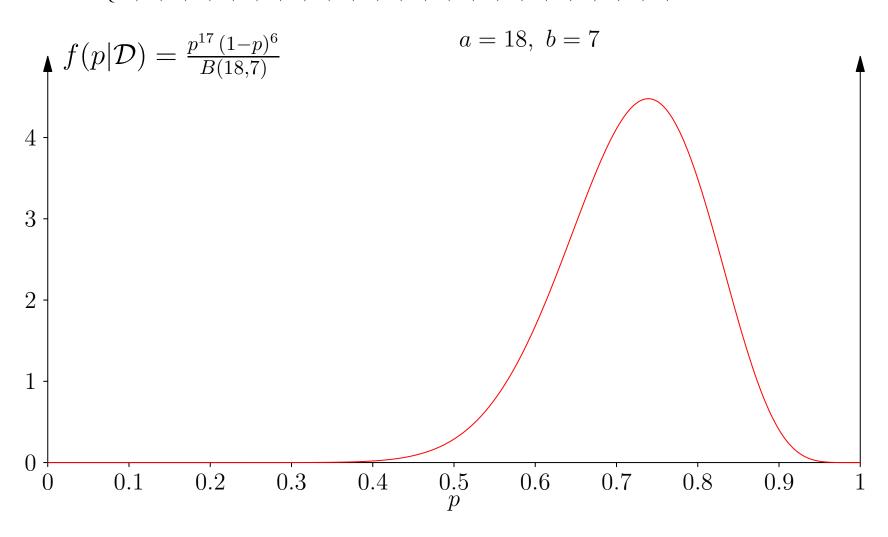


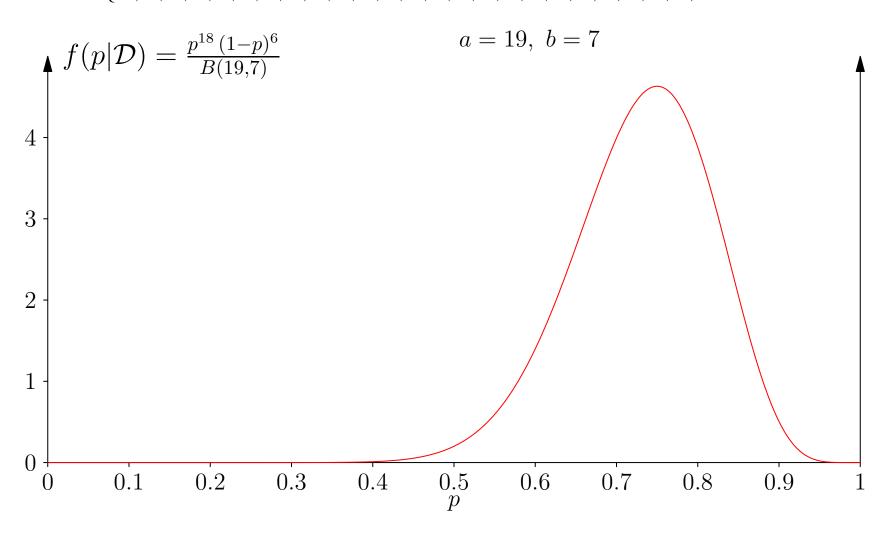


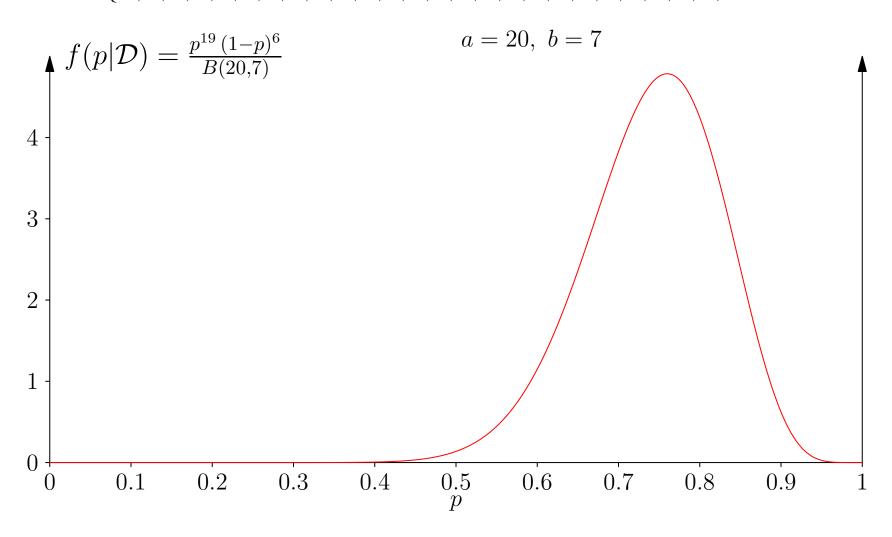


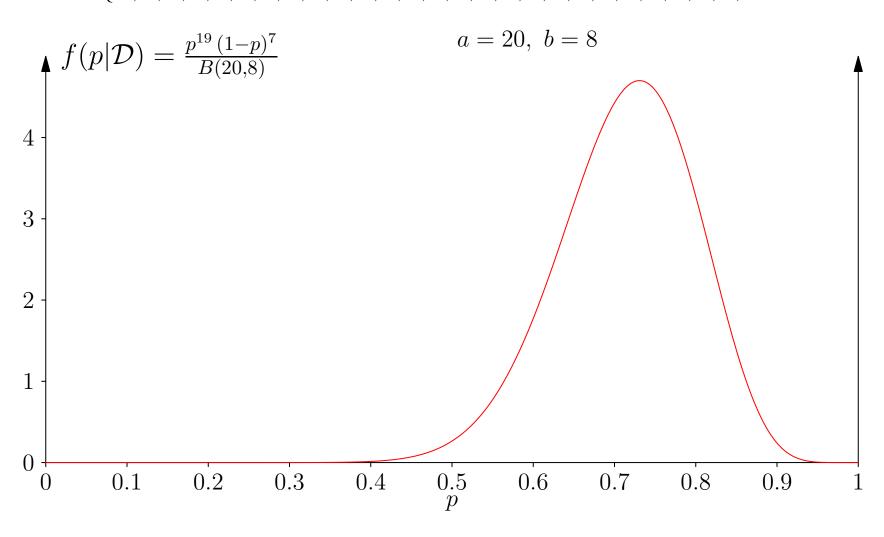


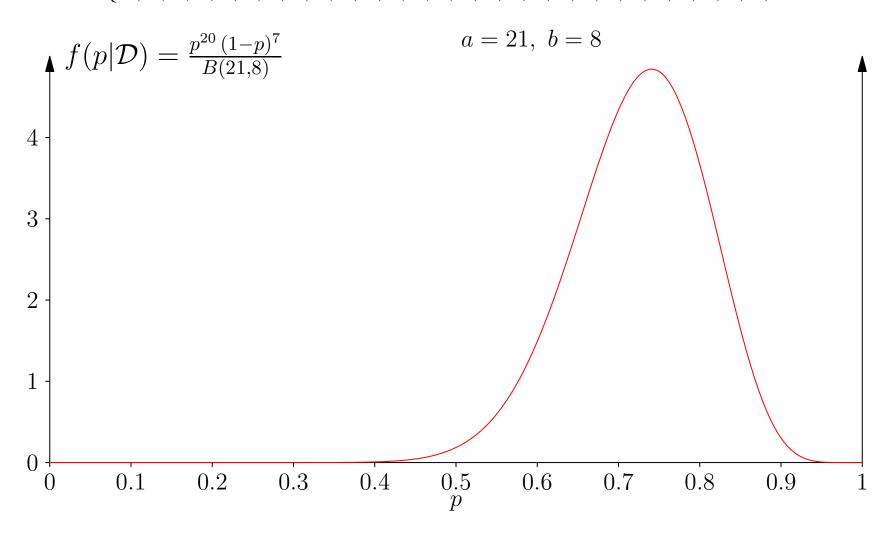


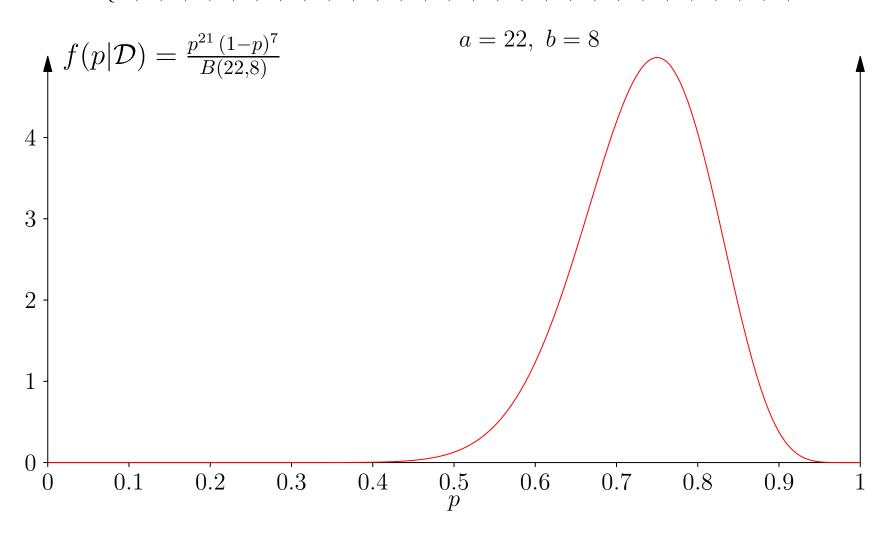


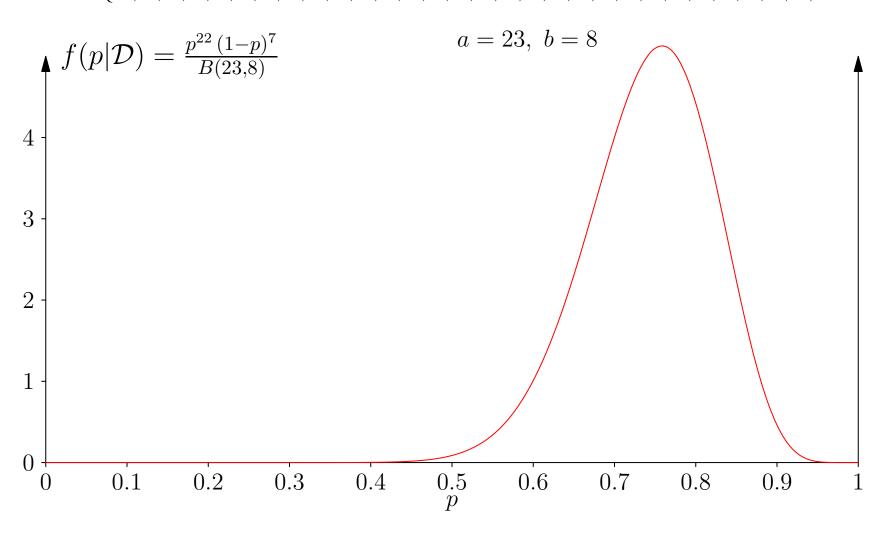


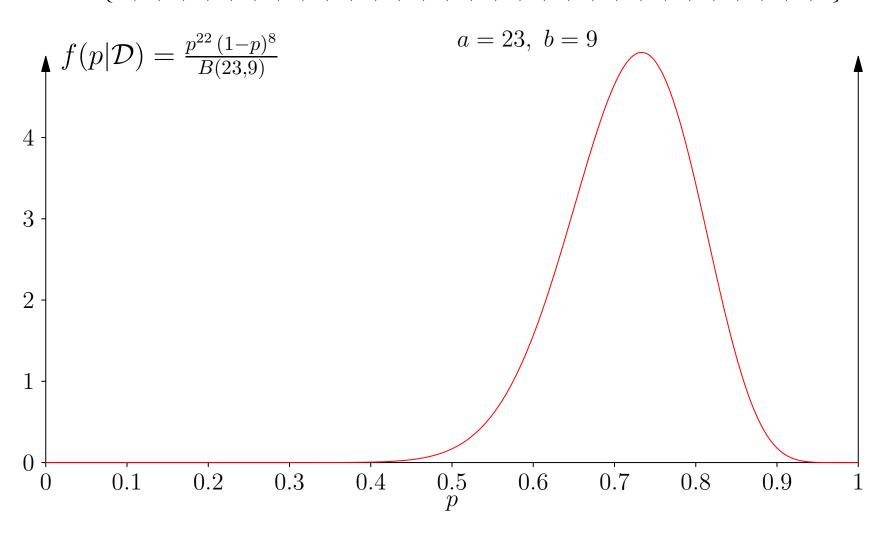












Estimating Prediction Errors

- A full Bayesian treatment gives a prediction of its own error
- Assuming $f(p|\mathcal{D}) = \text{Beta}(p|a,b)$
- The expected value of p is given by a/(a+b) = 23/32 = 0.719
- The standard deviation is

$$\sqrt{\frac{ab}{(a+b)^2(a+b+1)}} = 0.078$$

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Poisson Likelihoods

- Let's look at a second example of conjugate priors
- Suppose we want to find the rate of traffic along a road between 1:00pm and 2:00pm
- We assume the number of cars is given by a Poisson distribution

$$\mathbb{P}(N) = \operatorname{Pois}(N|\mu) = \frac{\mu^N}{N!} e^{-\mu}$$

ullet μ is the rate of traffic per hour which we want to infer from observation taken on different days

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$$p(\mu) = \Gamma(\mu|a_0, b_0) = \frac{b_0^{a_0} \mu^{a_0 - 1} e^{-b_0 \mu}}{\Gamma(a)}$$

- We will assume that we know nothing. The uninformative prior is $a_0 = b_0 = 0$
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Posterior

The posterior after seeing the first piece of data is

$$p(\mu|N_1) \propto \mathbb{P}(N_1|\mu) p(\mu)$$

 $\propto \frac{\mu^{N_1}}{N_1!} e^{-\mu} \mu^{a_0-1} e^{-b_0 \mu}$
 $\propto \mu^{N_1+a_0-1} e^{-(b_0+1)\mu}$

• The posterior is also a Gamma distribution $\Gamma(\mu|a_1,b_1)$ with $a_1=a_0+N_1,\ b_1=b_0+1$

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• The posterior is also a Gamma distribution $\Gamma(\mu|a_1,b_1)$ with $a_1=a_0+N_1,\ b_1=b_0+1$

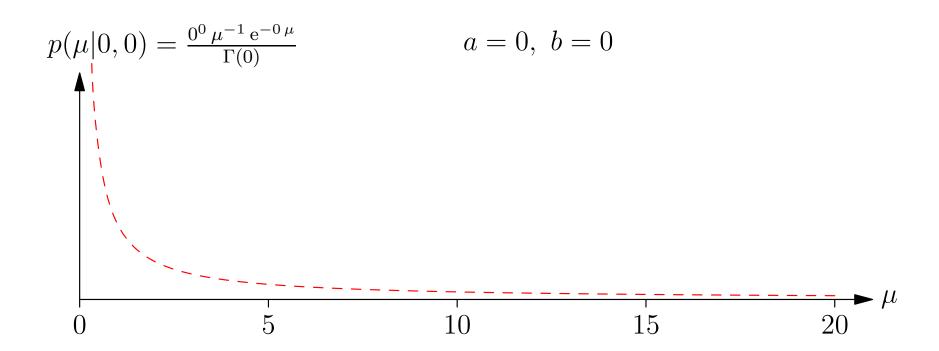
Posterior

The posterior after seeing the first piece of data is

$$p(\mu|N_1) \propto \mathbb{P}(N_1|\mu) p(\mu)$$

 $\propto \frac{\mu^{N_1}}{N_1!} e^{-\mu} \mu^{a_0-1} e^{-b_0 \mu}$
 $\propto \mu^{N_1+a_0-1} e^{-(b_0+1)\mu}$

• The posterior is also a Gamma distribution $\Gamma(\mu|a_1,b_1)$ with $a_1=a_0+N_1,\ b_1=b_0+1$



$$\mathcal{D} = \{4\}$$

$$p(\mu|4,1) = \frac{1^4 \,\mu^3 \,\mathrm{e}^{-1 \,\mu}}{\Gamma(4)} \qquad a = 4, \ b = 1$$

$$0 \quad 0 \quad 5 \quad 10 \quad 15 \quad 20$$

$$\mathcal{D} = \{4, 4\}$$

$$p(\mu|8, 2) = \frac{2^8 \,\mu^7 \,\mathrm{e}^{-2\,\mu}}{\Gamma(8)} \qquad a = 8, \ b = 2$$

$$\mathcal{D} = \{4, 4, 6\}$$

$$p(\mu|14, 3) = \frac{3^{14} \mu^{13} e^{-3 \mu}}{\Gamma(14)} \qquad a = 14, b = 3$$

$$\mathcal{D} = \{4, 4, 6, 4\}$$

$$p(\mu|18, 4) = \frac{4^{18} \mu^{17} e^{-4 \mu}}{\Gamma(18)}$$

$$a = 18, b = 4$$

$$1 - \frac{1}{5}$$

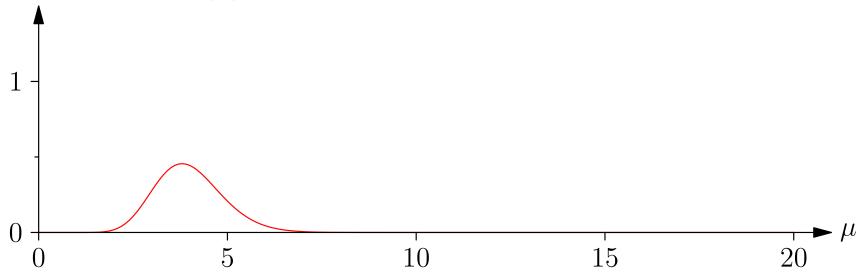
$$1 - \frac{1}{5}$$

$$1 - \frac{1}{5}$$

$$\mathcal{D} = \{4, 4, 6, 4, 2\}$$

$$p(\mu|20,5) = \frac{5^{20} \,\mu^{19} \,\mathrm{e}^{-5 \,\mu}}{\Gamma(20)}$$

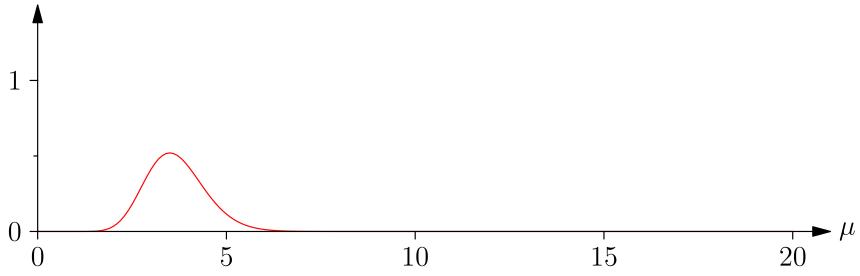
$$a = 20, b = 5$$



$$\mathcal{D} = \{4, 4, 6, 4, 2, 2\}$$

$$p(\mu|22,6) = \frac{6^{22} \mu^{21} e^{-6 \mu}}{\Gamma(22)}$$

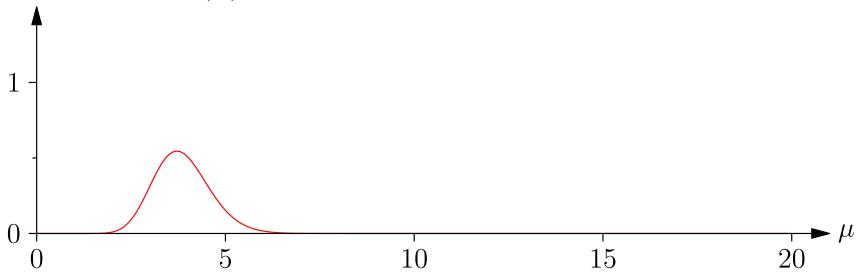
$$a = 22, b = 6$$



$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5\}$$

$$p(\mu|27,7) = \frac{7^{27} \mu^{26} e^{-7 \mu}}{\Gamma(27)}$$

$$a = 27, b = 7$$

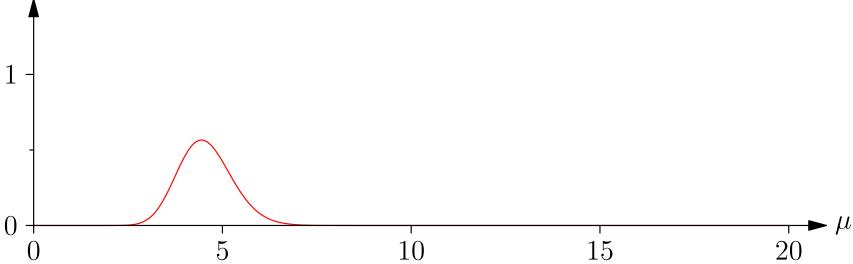


$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9\}$$

$$p(\mu|36, 8) = \frac{8^{36} \,\mu^{35} \,e^{-8 \,\mu}}{\Gamma(36)} \qquad a = 36, \ b = 8$$

$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5\}$$

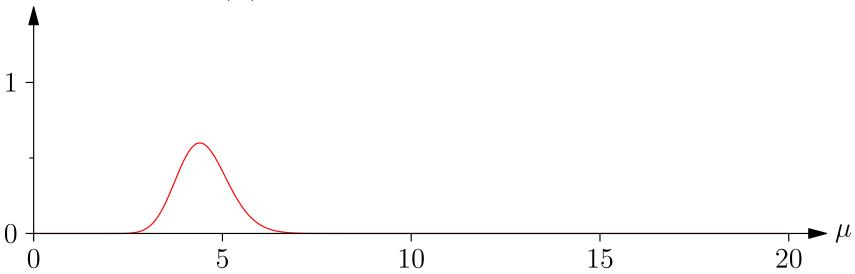
$$p(\mu|41,9) = \frac{9^{41} \,\mu^{40} \,\mathrm{e}^{-9 \,\mu}}{\Gamma(41)} \qquad \qquad a = 41, \ b = 9$$



$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4\}$$

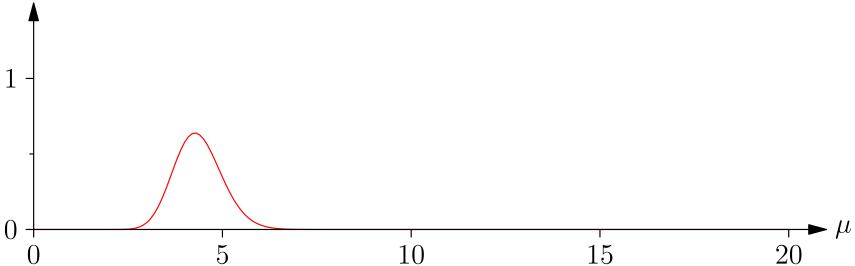
$$p(\mu|45, 10) = \frac{10^{45} \,\mu^{44} \,\mathrm{e}^{-10 \,\mu}}{\Gamma(45)}$$

$$a = 45, \ b = 10$$



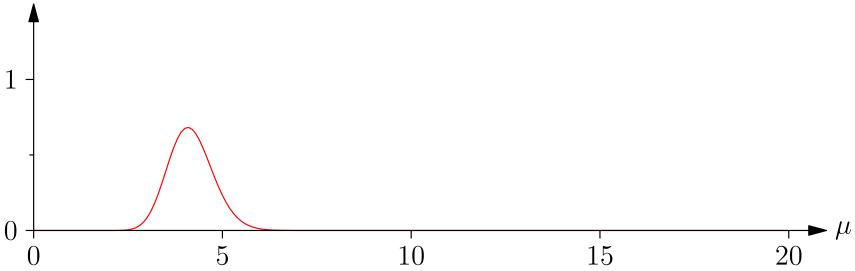
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3\}$$

$$p(\mu|48,11) = \frac{11^{48} \mu^{47} e^{-11 \mu}}{\Gamma(48)}$$
 $a = 48, b = 11$



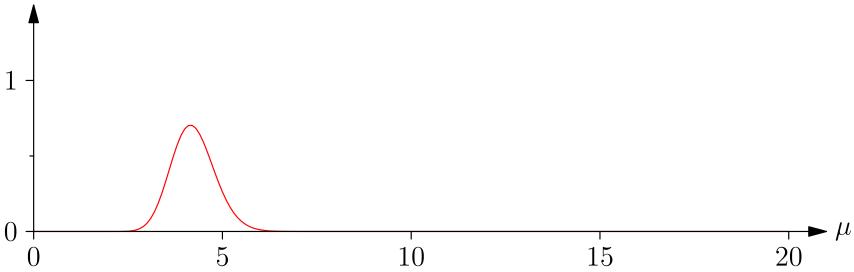
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2\}$$

$$p(\mu|50, 12) = \frac{12^{50} \mu^{49} e^{-12 \mu}}{\Gamma(50)}$$
 $a = 50, b = 12$



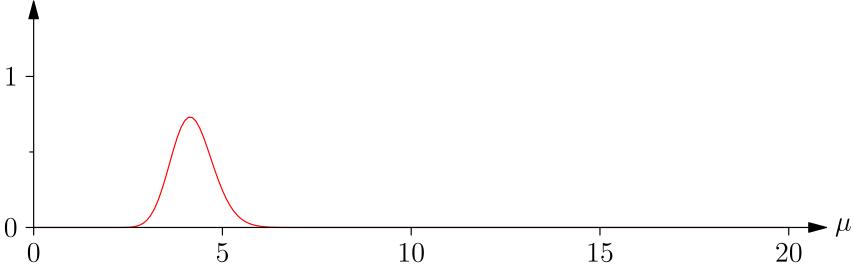
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5\}$$

$$p(\mu|55, 13) = \frac{13^{55} \mu^{54} e^{-13 \mu}}{\Gamma(55)}$$
 $a = 55, b = 13$



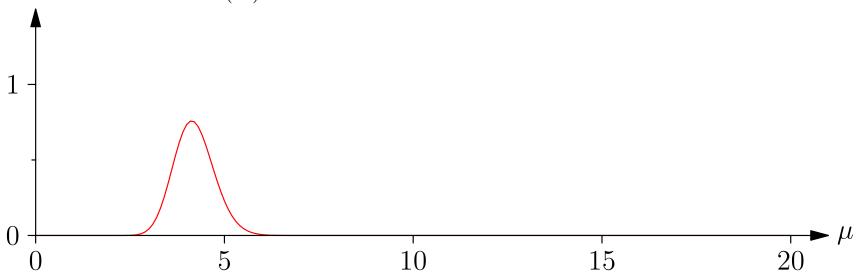
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4\}$$

$$p(\mu|59, 14) = \frac{14^{59} \mu^{58} e^{-14 \mu}}{\Gamma(59)}$$
 $a = 59, b = 14$



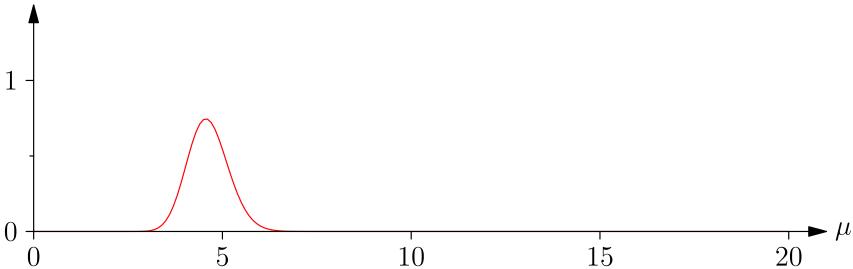
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4\}$$

$$p(\mu|63, 15) = \frac{15^{63} \mu^{62} e^{-15 \mu}}{\Gamma(63)}$$
 $a = 63, b = 15$



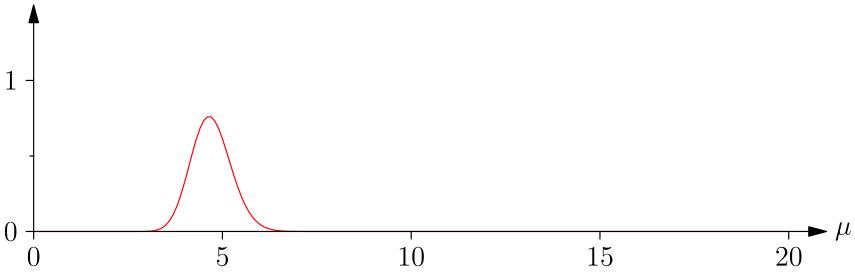
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11\}$$

$$p(\mu|74, 16) = \frac{16^{74} \mu^{73} e^{-16 \mu}}{\Gamma(74)}$$
 $a = 74, b = 16$



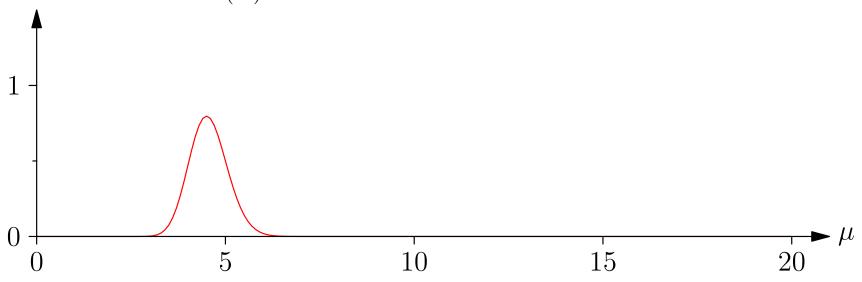
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11, 6\}$$

$$p(\mu|80, 17) = \frac{17^{80} \mu^{79} e^{-17 \mu}}{\Gamma(80)}$$
 $a = 80, b = 17$



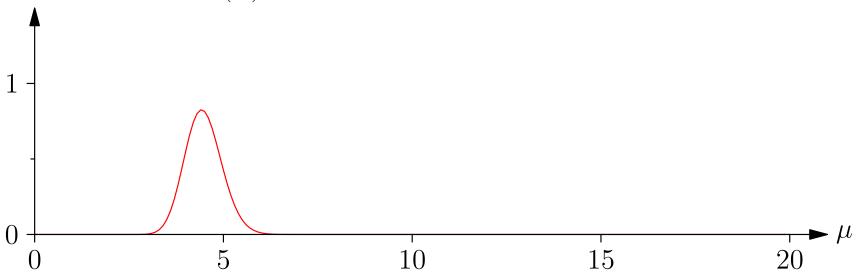
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11, 6, 2\}$$

$$p(\mu|82, 18) = \frac{18^{82} \mu^{81} e^{-18 \mu}}{\Gamma(82)}$$
 $a = 82, b = 18$



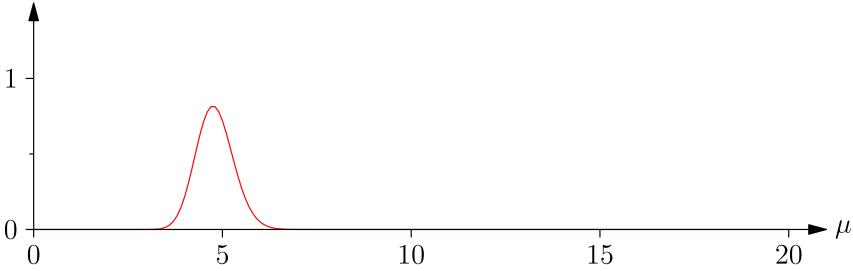
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11, 6, 2, 3\}$$

$$p(\mu|85, 19) = \frac{19^{85} \mu^{84} e^{-19 \mu}}{\Gamma(85)}$$
 $a = 85, b = 19$



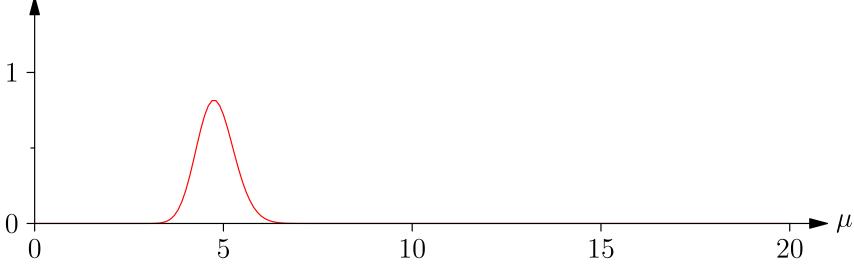
$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11, 6, 2, 3, 11\}$$

$$p(\mu|96, 20) = \frac{20^{96} \,\mu^{95} \,\mathrm{e}^{-20 \,\mu}}{\Gamma(96)} \qquad a = 96, \ b = 20$$



$$\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11, 6, 2, 3, 11\}$$

$$p(\mu|96, 20) = \frac{20^{96} \,\mu^{95} \,\mathrm{e}^{-20 \,\mu}}{\Gamma(96)} \qquad a = 96, \ b = 20$$



$$\mathbb{E}[\mu] = \frac{a}{b} = \frac{96}{20} = 4.8 \qquad \sqrt{\mathbb{V}\text{ar}(\mu)} = \sqrt{\frac{a}{b^2}} = 0.49$$

Outline

- 1. Bayes' Rule
- 2. Conjugate Priors
- 3. Uninformative Priors



- What if we have no prior knowledge, what should we do?
- OK usually we know whether we should make a measurement using a micrometer, ruler or car mileage, but we might still know almost nothing
- This led to Bayesian statistics being labelled as subjective
- However Ed. Jaynes (the greatest proponent of Bayesian methods) argued that we could answer this using symmetry arguments

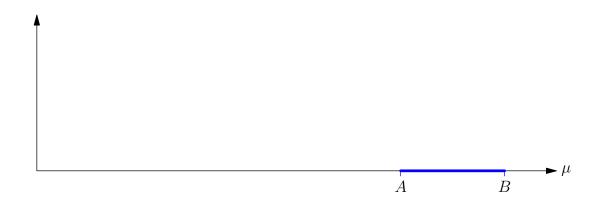
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Uninformative Priors for Scale Parameter

• Why did we choose $a_0 = b_0 = 0$ implying a prior $p(\mu) = 1/\mu$?

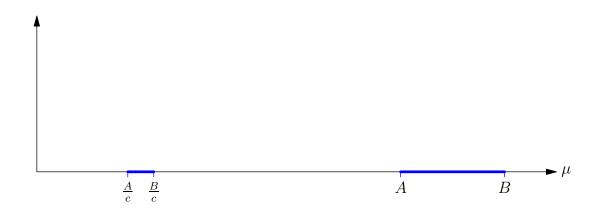


ullet That is, we have no idea on what scale to measure μ

$$\int_{A}^{B} p(\mu) d\mu = \int_{A/c}^{B/c} p(\mu) d\mu$$

• Or $p(\mu) = \frac{1}{c} p(\frac{\mu}{c})$ implying $p(\mu) \propto \frac{1}{\mu}$

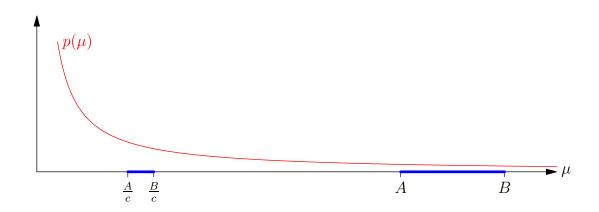
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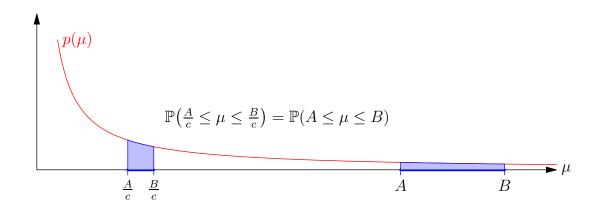
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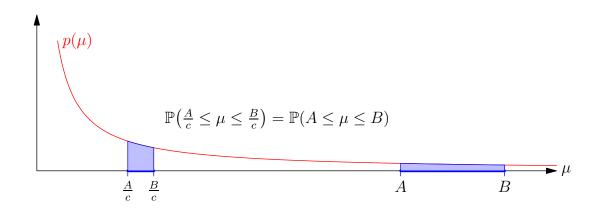
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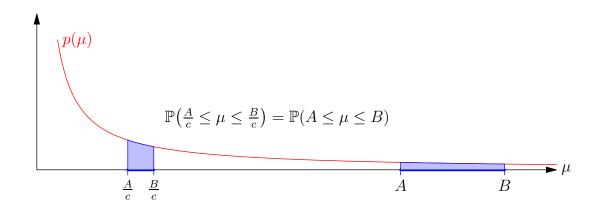
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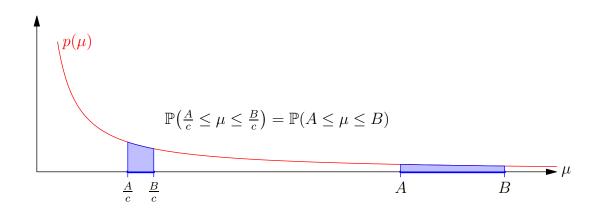


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$$\int_{A}^{B} p(\mu) d\mu = \int_{A/c}^{B/c} p(\mu) d\mu = \int_{A}^{B} \frac{1}{c} p(\frac{\nu}{c}) d\nu$$

making a change of variables $\mu = \nu/c$

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making a change of variables $\mu = \nu/c$

- Numbers occurring in life (physical constants, amounts of money) should not depend on the units (scale) measuring them
- They should then be distributed as $p(x) \propto 1/x$
- A curious consequence of this is that the significant figure has a distribution

$$\mathbb{P}(\text{most s.f. of } x = n) = \frac{\int_{n}^{n+1} \frac{1}{x} dx}{\int_{1}^{10} \frac{1}{x} dx}$$
$$= \frac{\log(n+1) - \log(n)}{\log(10)} = \log_{10}\left(\frac{n+1}{n}\right)$$

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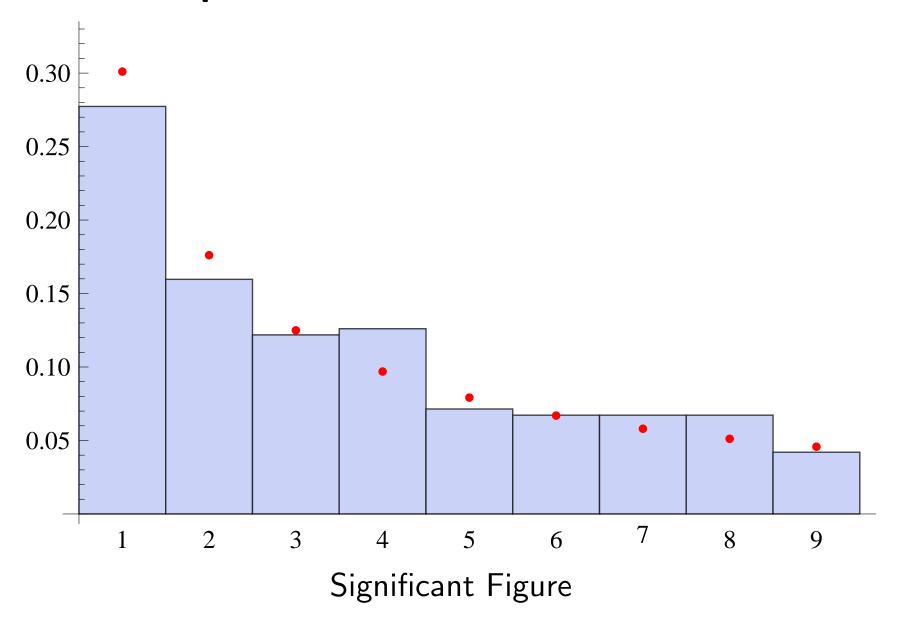
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Population Size of 238 Countries



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- However, it requires a model of what is happening
- In practice Bayesian methods are easy if the data is generated from a likelihood with a conjugate prior distribution—we have to be clever to choose the right prior
- We will see in the next lecture that much more frequently we will have likelihoods with no conjugate prior and we have to work much harder
- When we have no knowledge there are consistent ways to express our ignorance

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