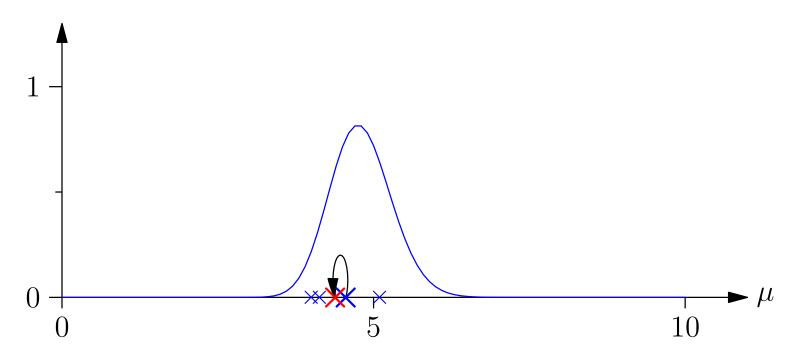
Advanced Machine Learning

MCMC

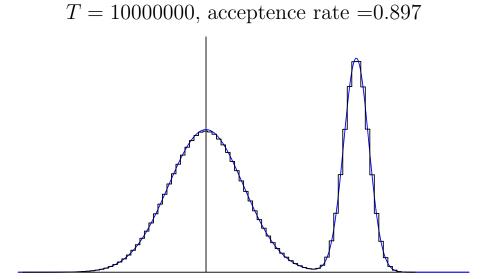
 $\mathcal{D} = \{4, 4, 6, 4, 2, 2, 5, 9, 5, 4, 3, 2, 5, 4, 4, 11, 6, 2, 3, 11\}$



Monte Carlo methods, MCMC, Variational Methods

Outline

- 1. Sampling
- 2. Random Number Generation
- 3. MCMC



- We saw that in some cases if we had a simple likelihood (normal, binomial, Poisson, multinomial) you can choose a conjugate prior (gamma-normal/Wishart, beta, gamma, Dirchlet) so that the posterior has the same form as the prior
- Very often we are working with more complex models where no conjugate prior exists
- The posterior is not described by a known distribution
- We have to work a lot harder—particularly with multivariate distributions

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- ullet Recall our problem is that we are given some data ${\mathcal D}$
- Our posterior is given by

$$\mathbb{P}(\boldsymbol{\theta}|\mathcal{D}) = \frac{\mathbb{P}(\mathcal{D}|\boldsymbol{\theta})\,\mathbb{P}(\boldsymbol{\theta})}{\mathbb{P}(\mathcal{D})} \qquad \text{or} \qquad f(\boldsymbol{\theta}|\mathcal{D}) = \frac{f(\mathcal{D}|\boldsymbol{\theta})\,f(\boldsymbol{\theta})}{f(\mathcal{D})}$$

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Histograms, Samples and Means

- We could represent our posterior as a histogram, although for multivariate distributions (i.e. when we are modelling more than one variable) a histogram can be unwieldy
- A sample from the posterior distribution is often sufficient e.g. in our topic models (LDA) a typical set of topics is what we are after
- However, when samples vary a lot, often the most useful quantities are expectation, e.g.

$$\mathbb{E}[\boldsymbol{\Theta}] \qquad \qquad \mathbb{E}\left[\Theta_i^2\right] - \mathbb{E}[\boldsymbol{\Theta}_i]^2$$

$$\mathbb{E}[\boldsymbol{\Theta}_i\boldsymbol{\Theta}_j] - \mathbb{E}[\boldsymbol{\Theta}_i]\mathbb{E}[\boldsymbol{\Theta}_j] \qquad \qquad \mathbb{E}\left[\boldsymbol{\Theta}\boldsymbol{\Theta}^\mathsf{T}\right] - \mathbb{E}[\boldsymbol{\Theta}]\mathbb{E}[\boldsymbol{\Theta}]^\mathsf{T}$$

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• If we can draw independent **deviates** (aka **variates**), Θ_i , from our posterior distribution then we can obtain an estimate of our expectation

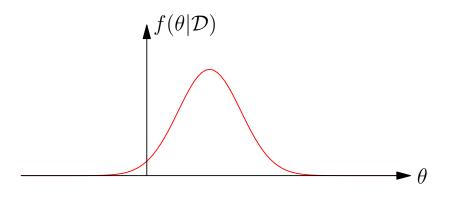
$$\mathbb{E}[g(\mathbf{\Theta})] \approx \frac{1}{n} \sum_{i=1}^{n} g(\mathbf{\Theta}_i)$$

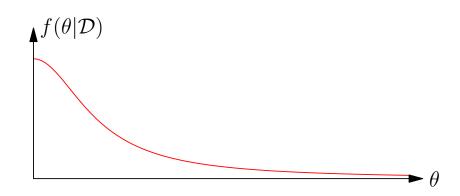
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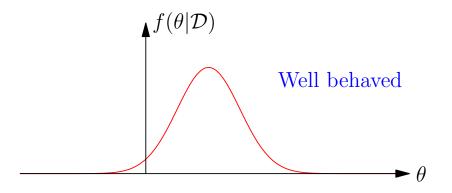
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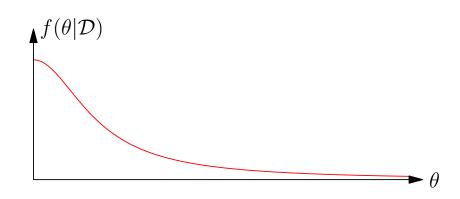




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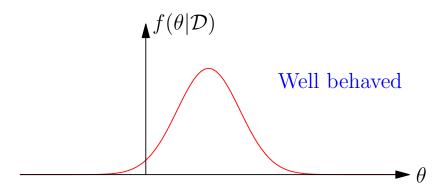
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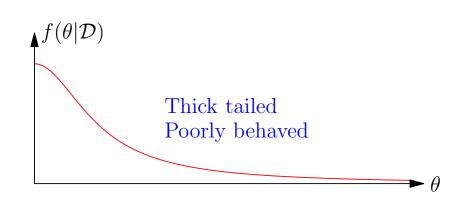




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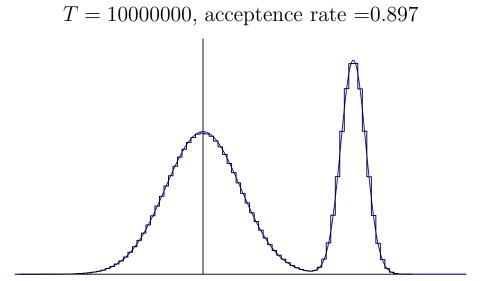
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Outline

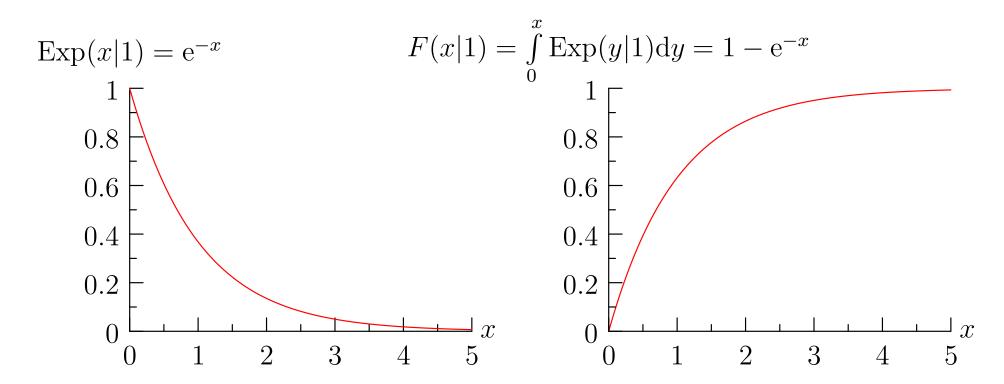
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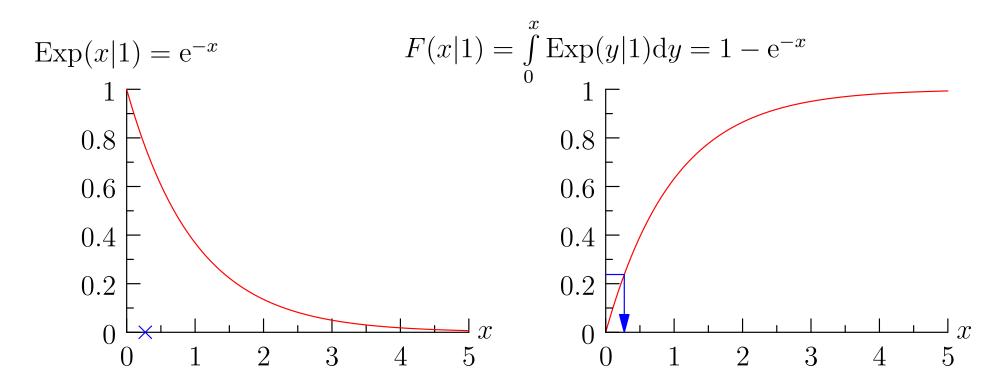
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- For some very simple distributions we can use the transformation methods to transform a uniform distribution

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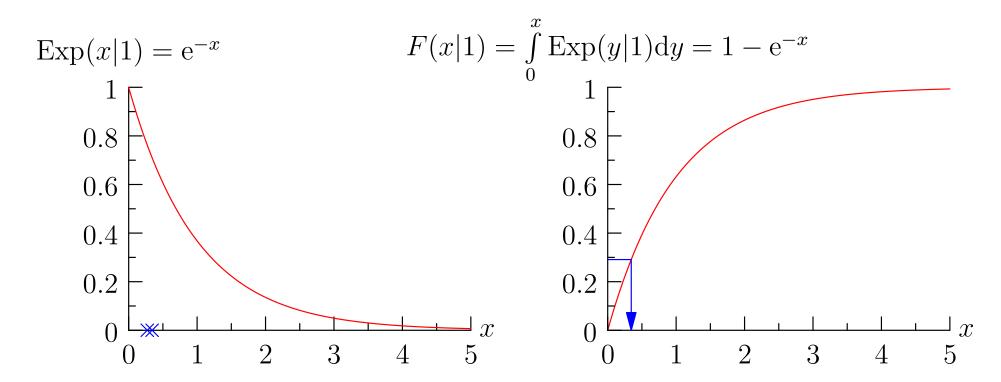
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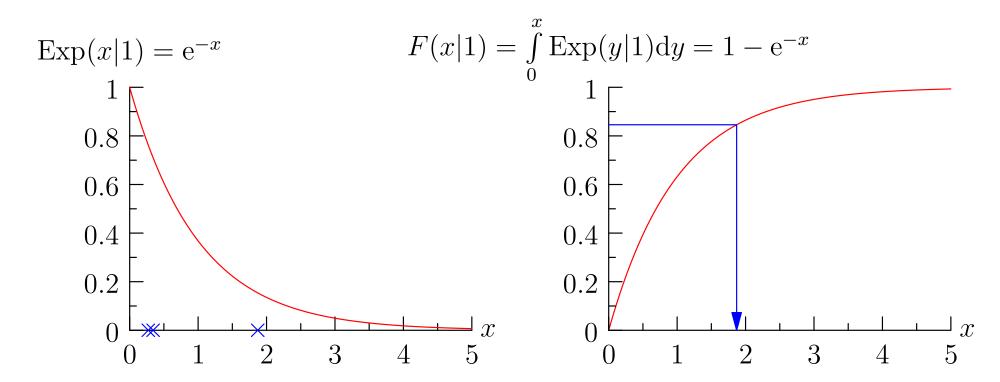
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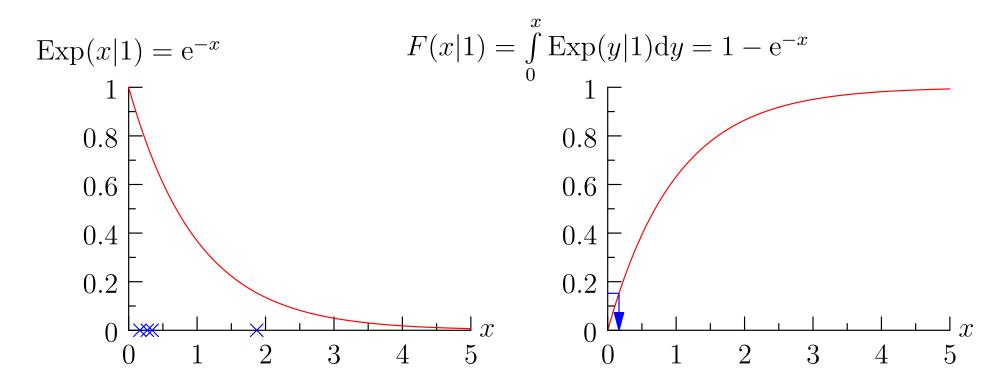
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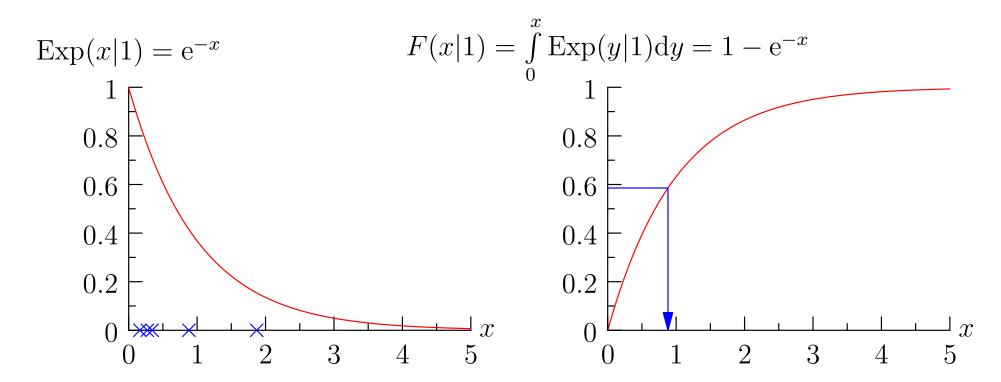
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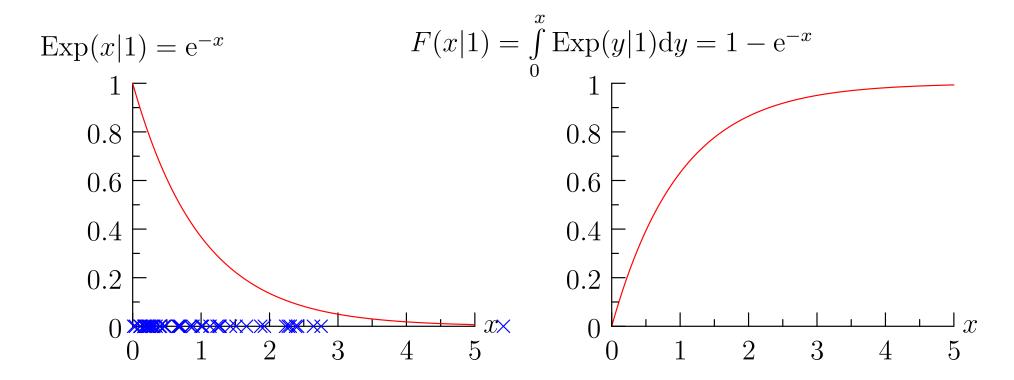
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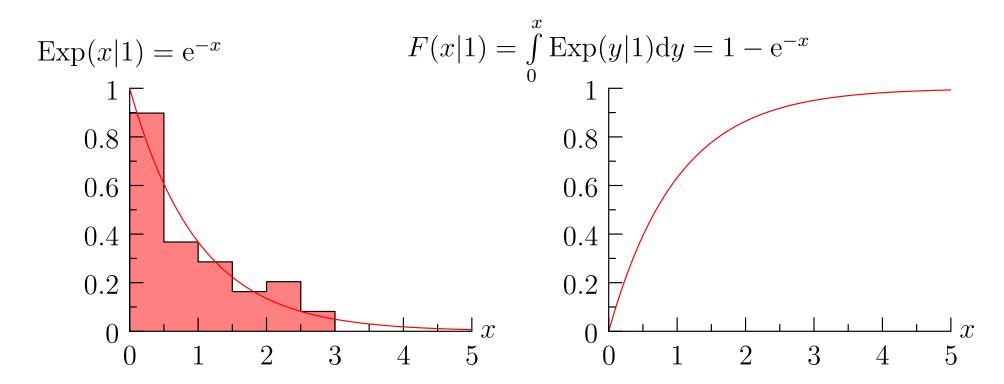
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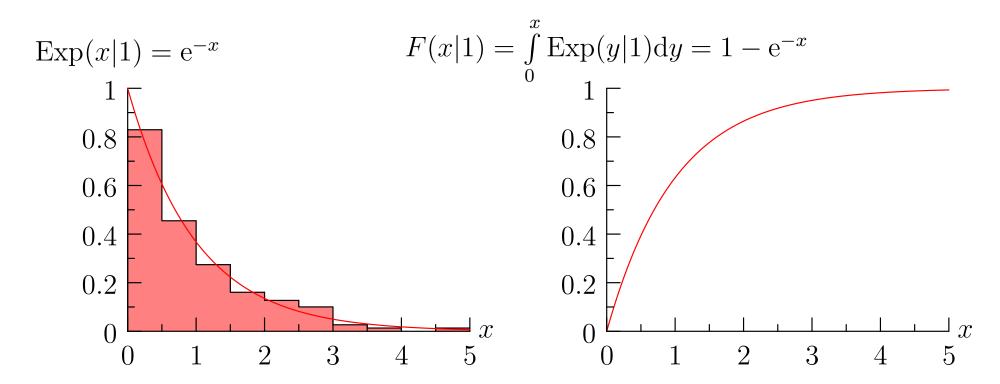
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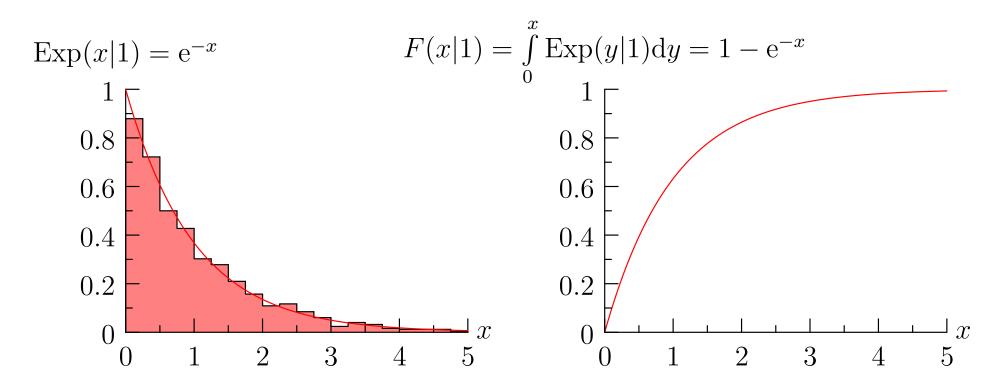
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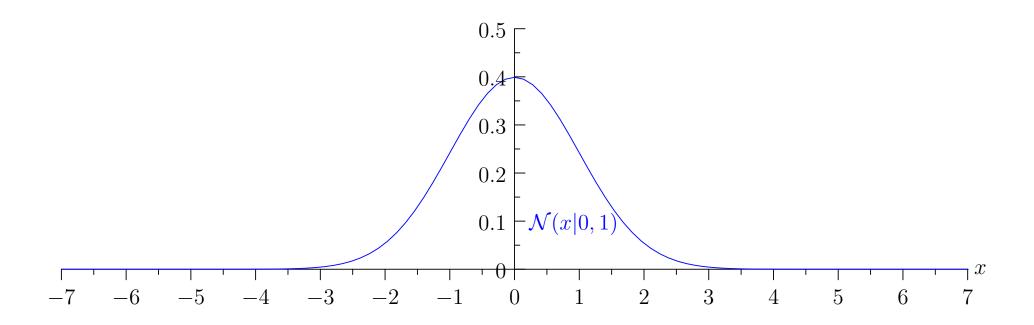
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- A more general technique is the **rejection method** where we generate deviates from $g_Y(y)$ such that $cg_Y(x) \ge f_X(x)$
- To draw deviates from $f_X(x)$ we draw a deviate $Y \sim g_Y$ and then accept the deviate with probability $f_X(Y)/(cg_Y(Y))$
- The expected rejection rate is c-1
- Need to choose a good distribution $g_Y(y)$

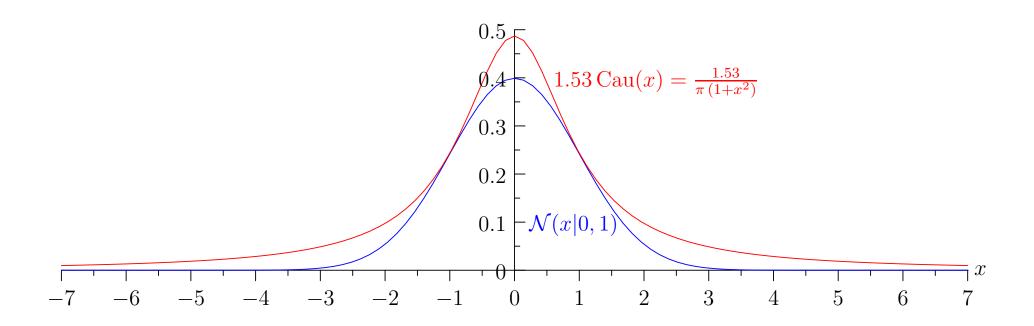
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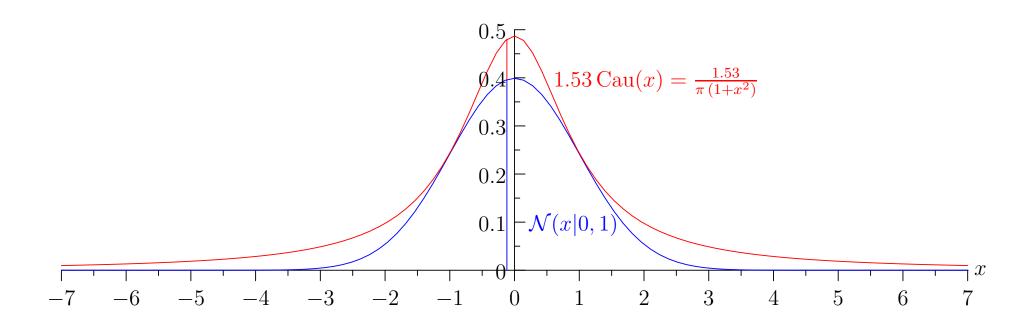
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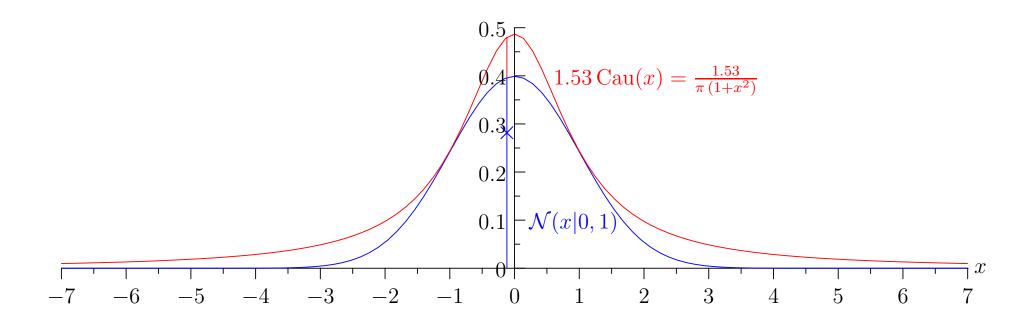
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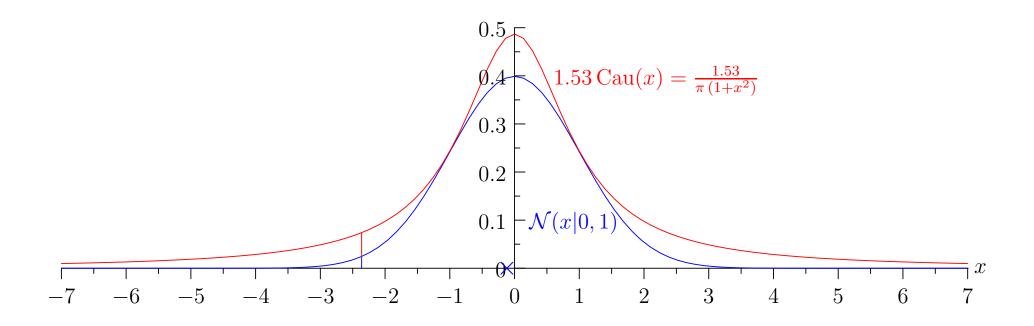
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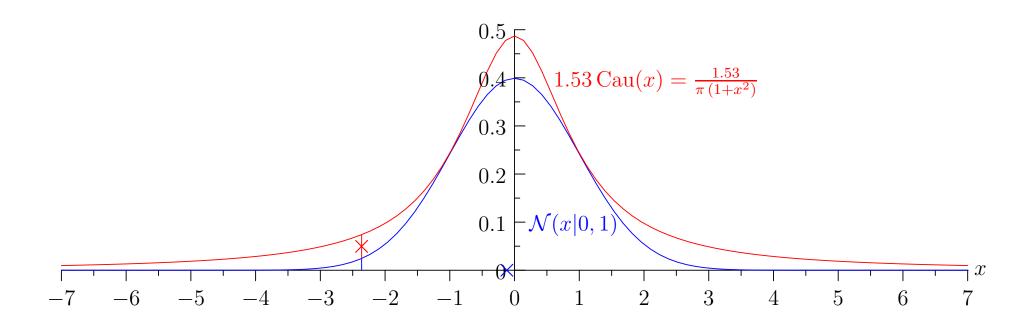


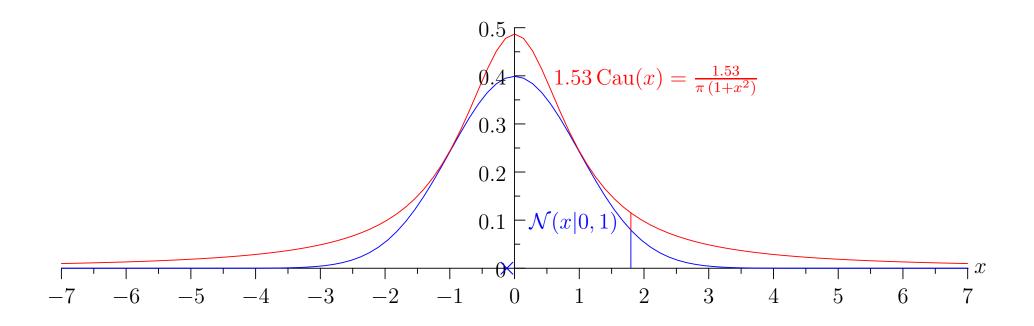


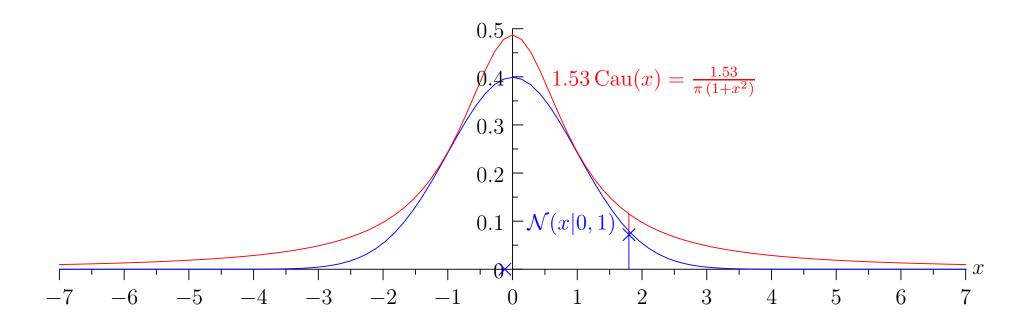


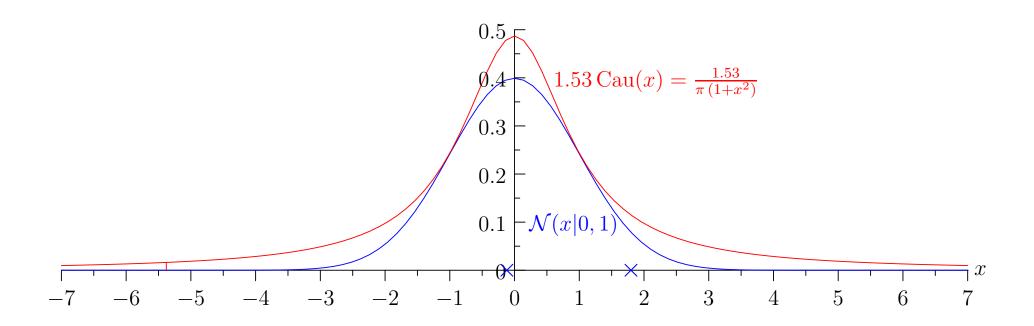


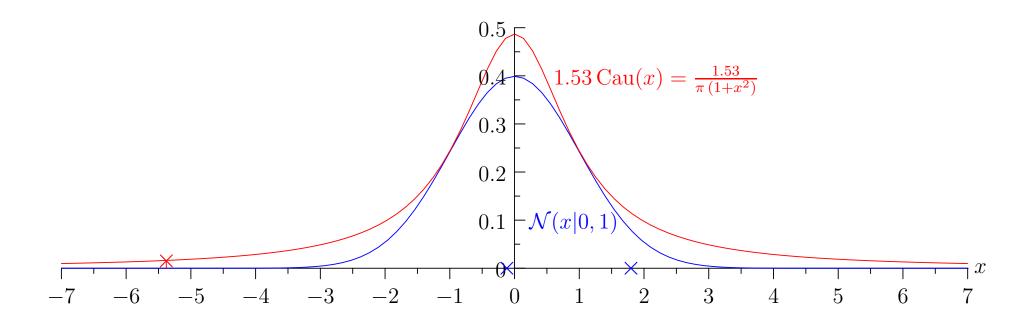


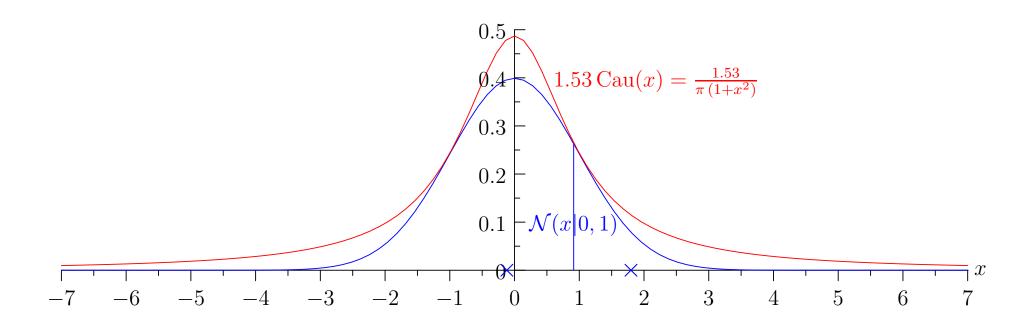


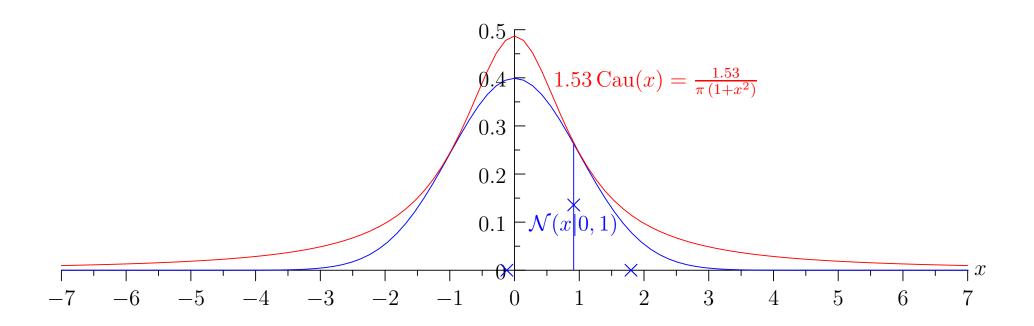


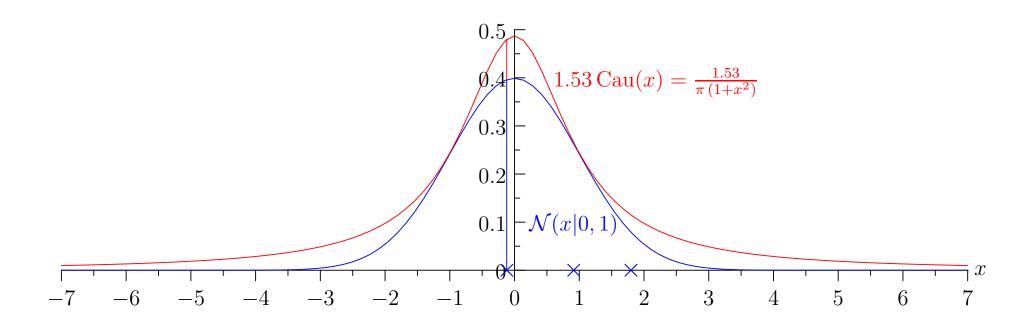


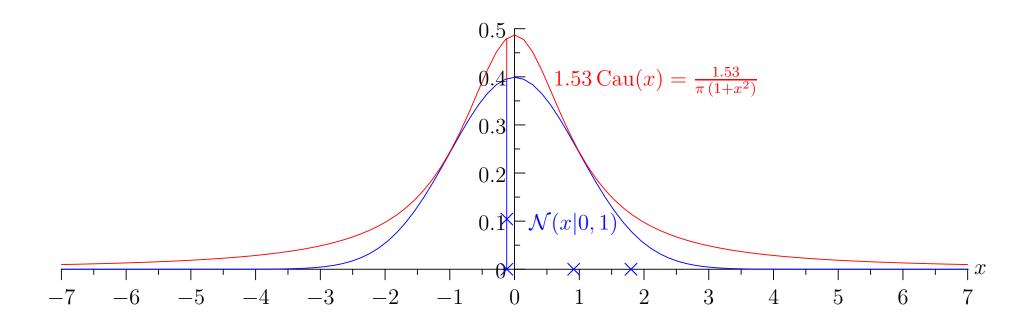


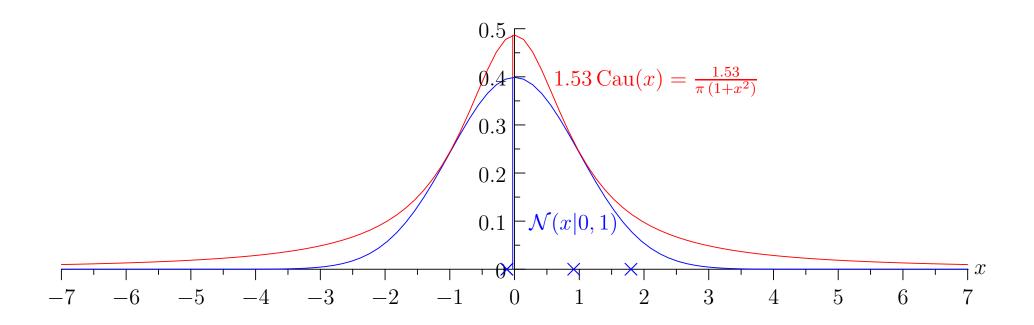


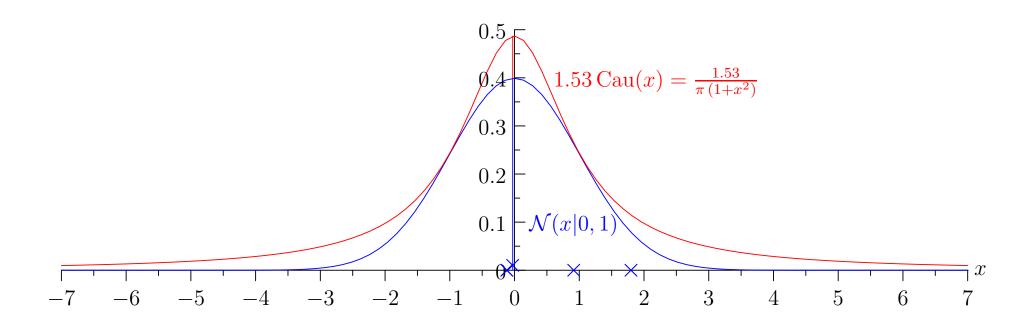


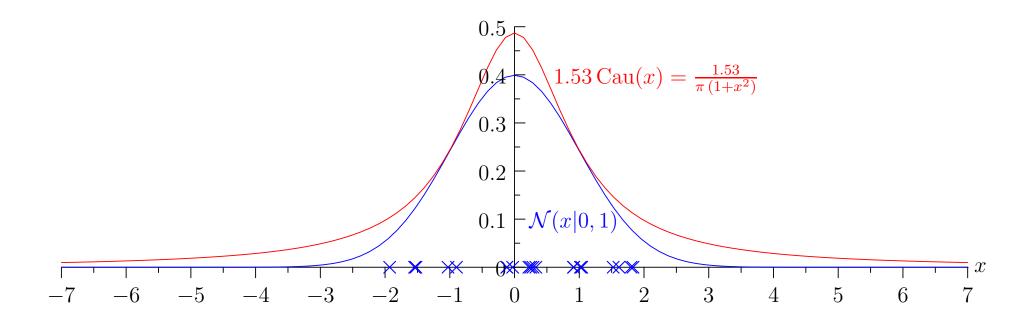


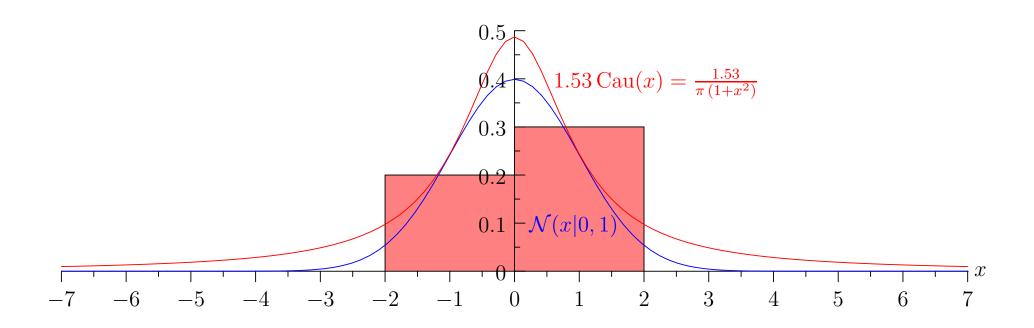


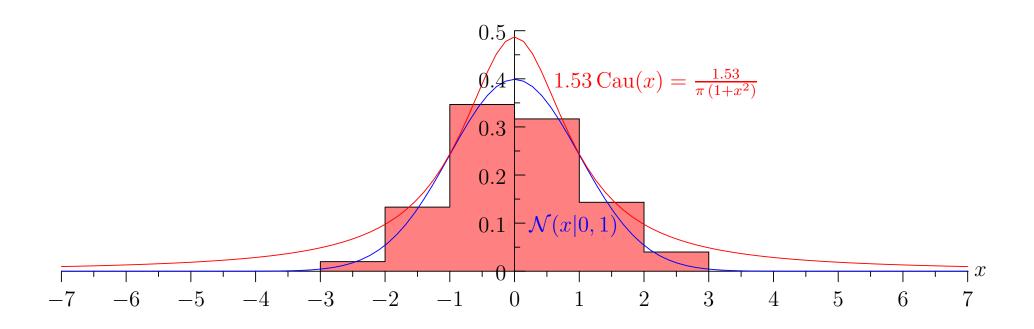


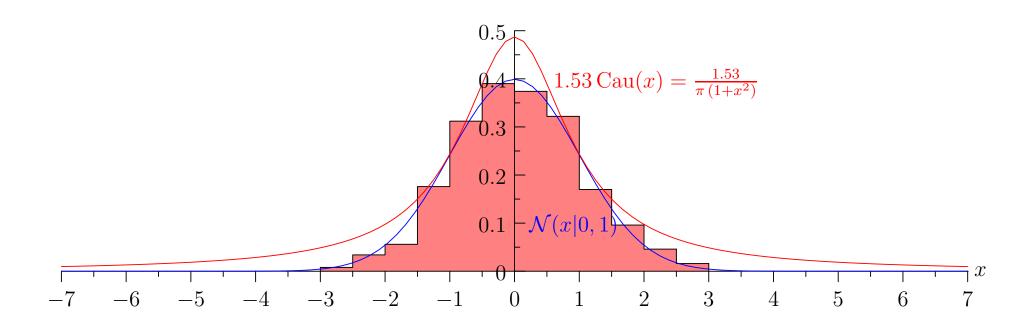












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- However, for complicated probability distributions it can be difficult to find a good proposal distribution $g_Y(y)$
- This is particular true for multivariate distributions
- ullet If the proposal distribution is poor c might be very high and the number of rejections is stupidly high

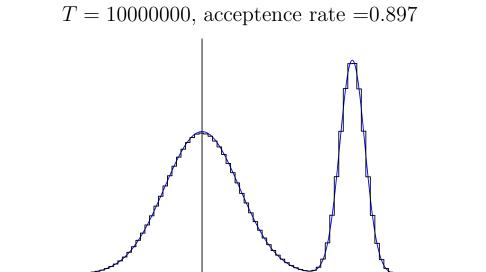
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- 3. **MCMC**



- Suppose we have a set of states $\mathcal S$ and want to draw sample from a probability distribution $\boldsymbol{\pi} = (\pi_i | i \in \mathcal S)$
- ullet We invent a dynamical system with a transition probability M_{ij} from state j to state i such that

$$M_{ij}\pi_j = M_{ji}\pi_i$$

- This is known as detailed balance
- Summing both sides over j

$$\sum_{j} M_{ij} \pi_j = \sum_{j} M_{ji} \pi_i$$

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- Suppose we start from a state $x(0) = \sum_i c_i v^{(i)}$ where the $v^{(i)}$'s are an eigenvectors of the transition matrix M with eigenvalues λ_i
- It I apply M many times then

$$\boldsymbol{x}(t) = \mathbf{M}^t \boldsymbol{x}(0)$$

• And $\lim_{t \to \infty} {m x}(t) = {m v}^*$ where ${m v}^*$ is the eigenvector with the maximum eigenvalue

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- Now $\|\mathbf{M}v\|_1 \leq \|\mathbf{M}\|_1 \|v\|_1 = \|v\|_1$ so the maximum eigenvalue is 1 with eigenvector π (\mathbf{M} is known as a stochastic matrix)

- ullet A very easy way to achieve detailed balance is starting from state j choose a "neighbouring" state, i with equal probability
- We accept the move if either
 - \star $\pi_i > \pi_j$ or
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$$W(\boldsymbol{\theta} \to \boldsymbol{\theta}')\pi(\boldsymbol{\theta}) = W(\boldsymbol{\theta}' \to \boldsymbol{\theta})\pi(\boldsymbol{\theta}')$$

- ullet where $\pi(oldsymbol{ heta})$ is the probability distribution we wish to sample from
- The update rule is to choose a nearby value θ' , compute $r = \pi(\theta')/\pi(\theta)$ and accept the update with probability $\min(1,r)$
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- Because we are free to choose where we move (and choose close by neighbours) $\pi(\theta') \approx \pi(\theta)$ so that moves are not too infrequent
- Also very importantly the updates depend only on the ratio $\pi(\pmb{\theta}')/\pi(\pmb{\theta})$
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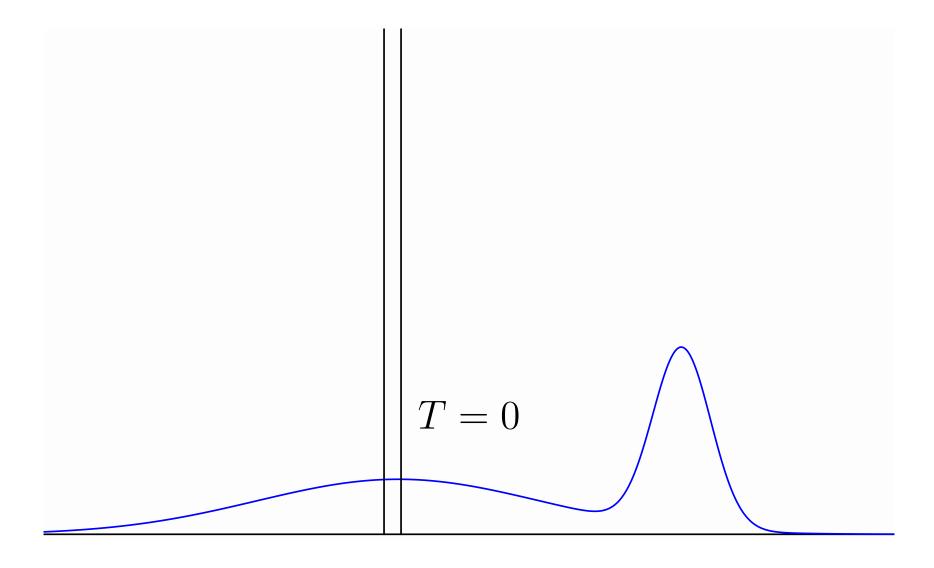
- It can take a long time until our states occur with the probability π (i.e. we have forgotten our initial state)
- We don't even know how long we have to wait
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- To get a good approximation to the posterior expectation requires running for many times the equilibration time
- Note, if we are just finding sample averages then we can use all samples after equilibrating even if they are not independent

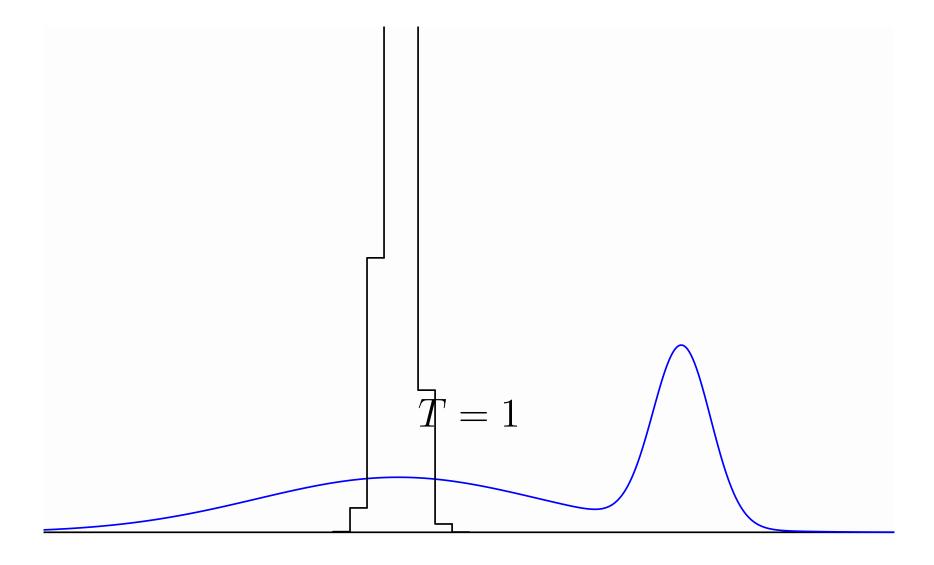
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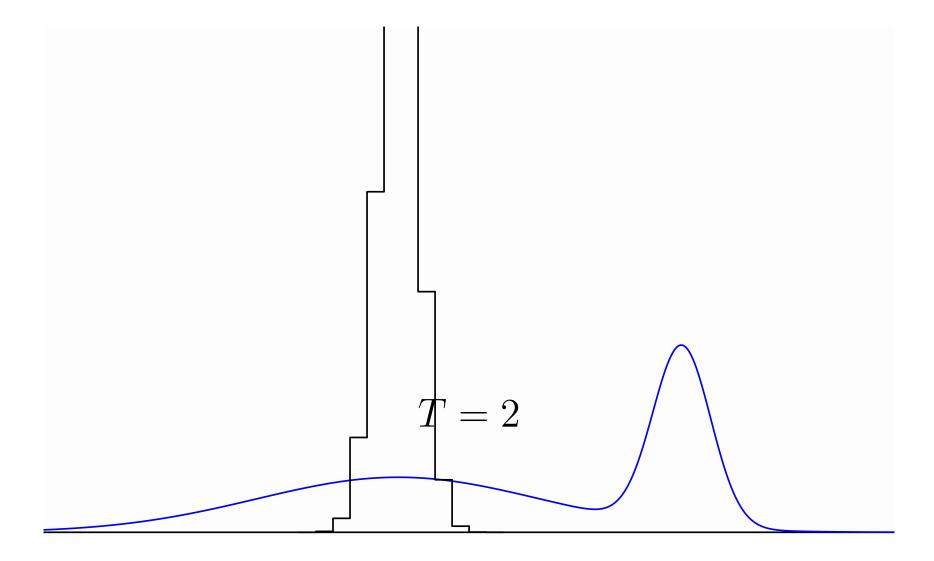
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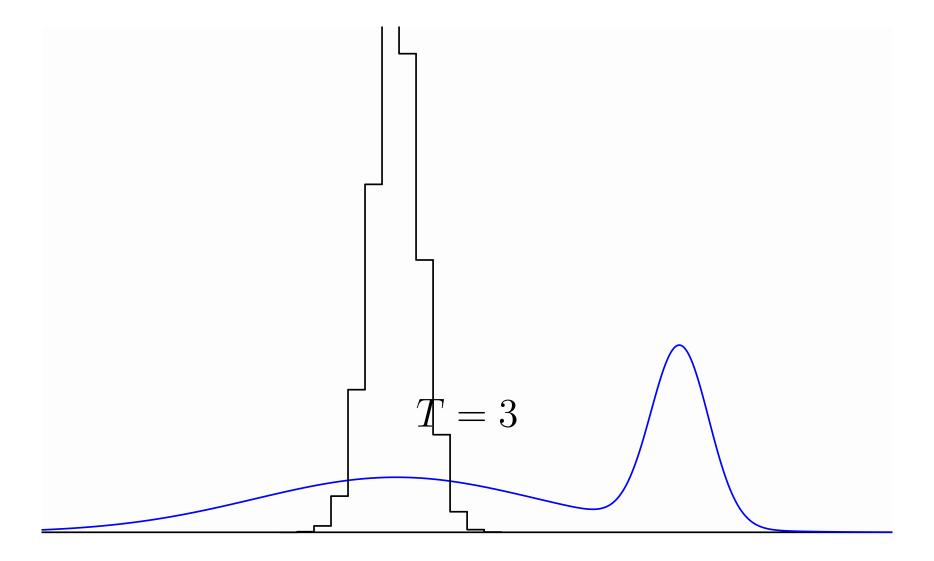
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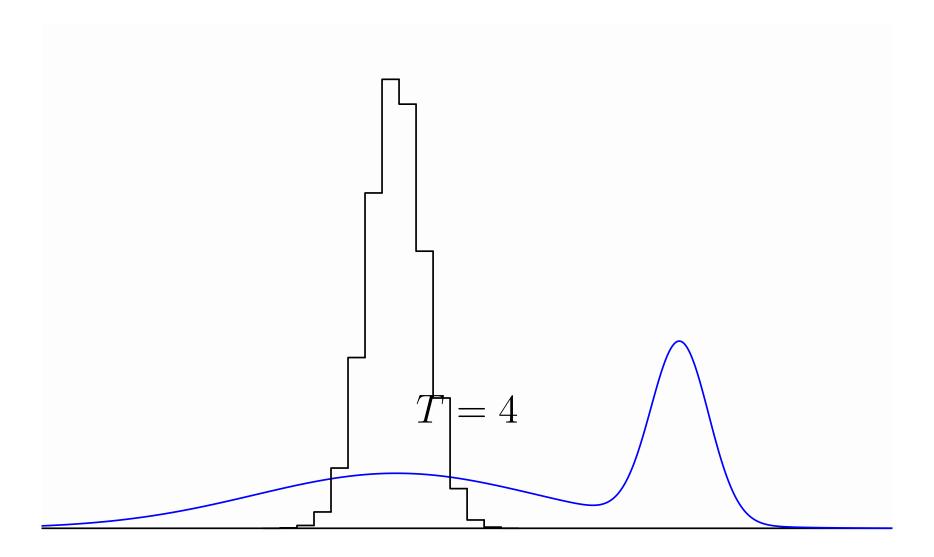
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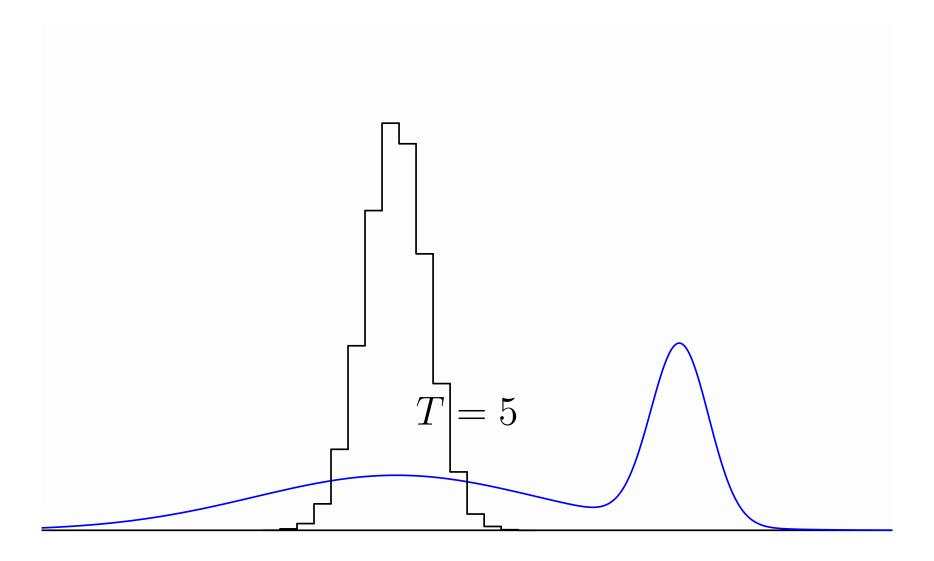


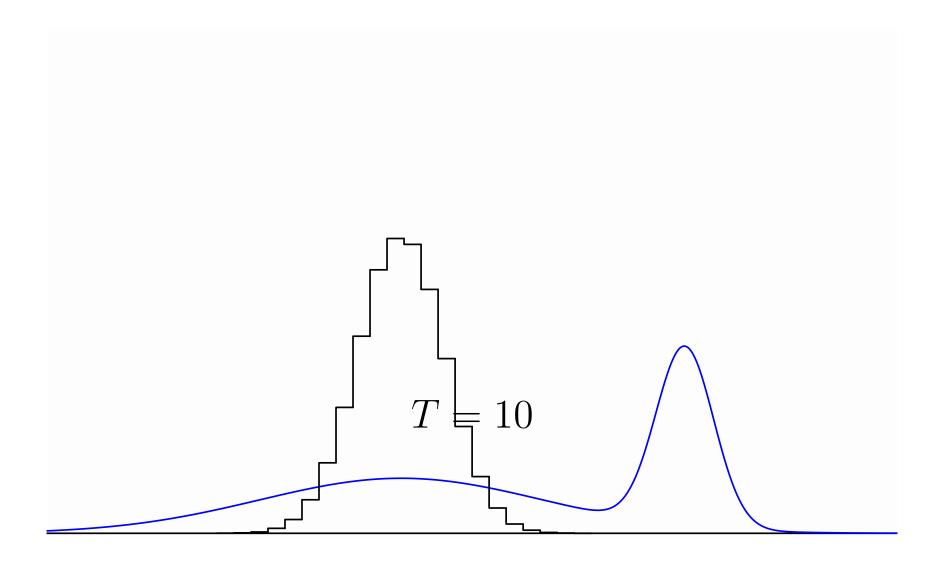


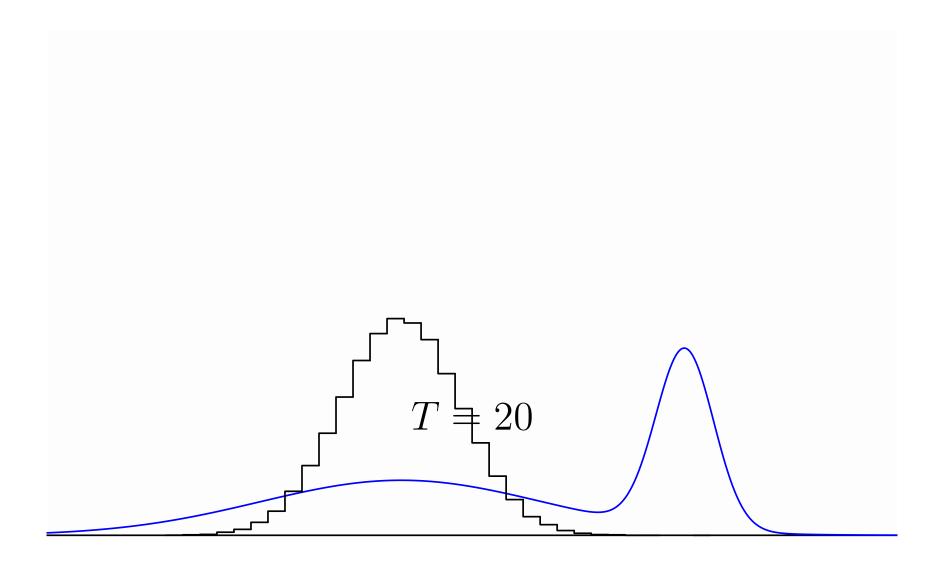


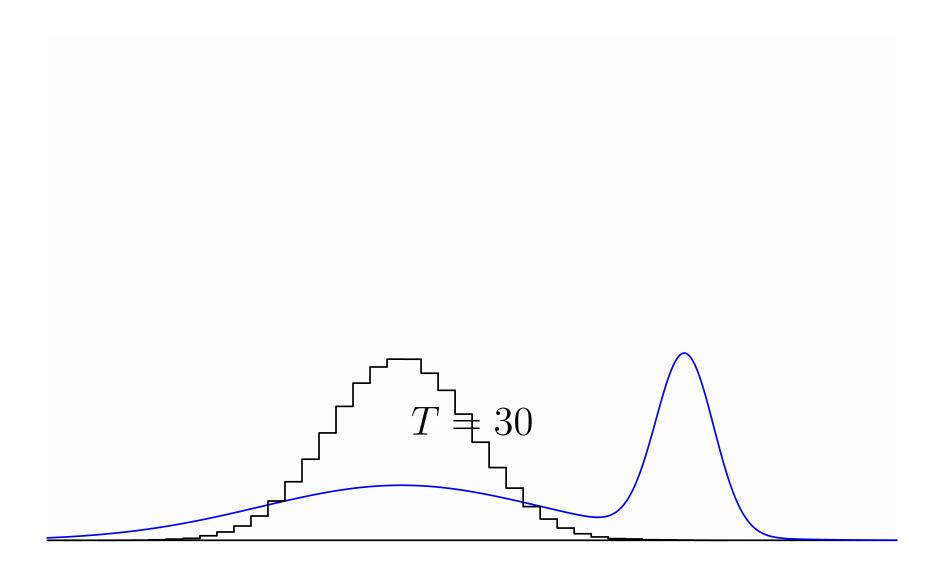


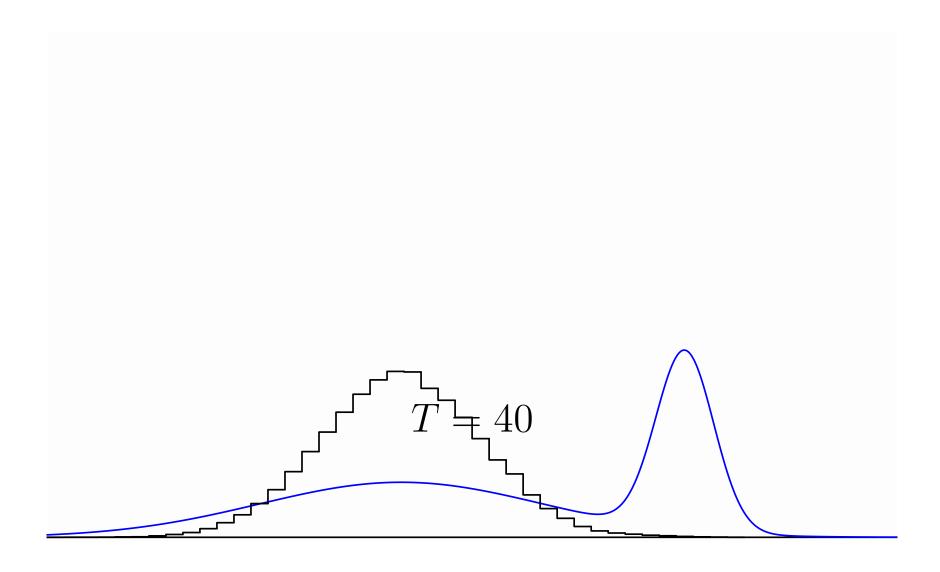


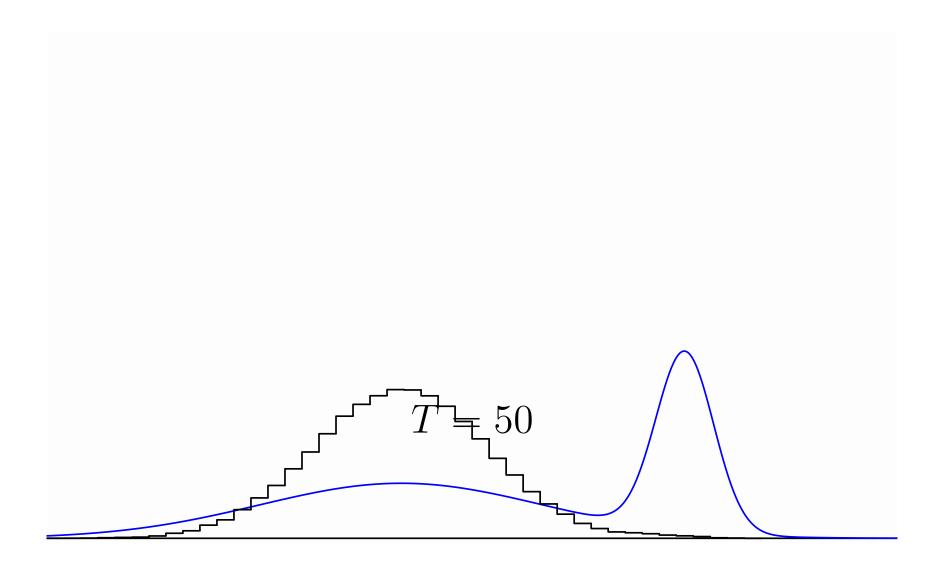


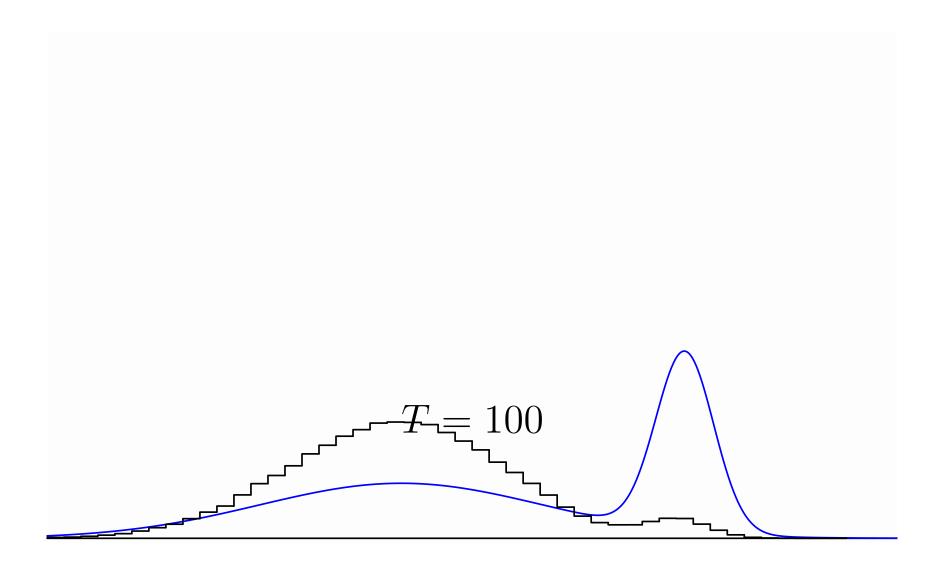


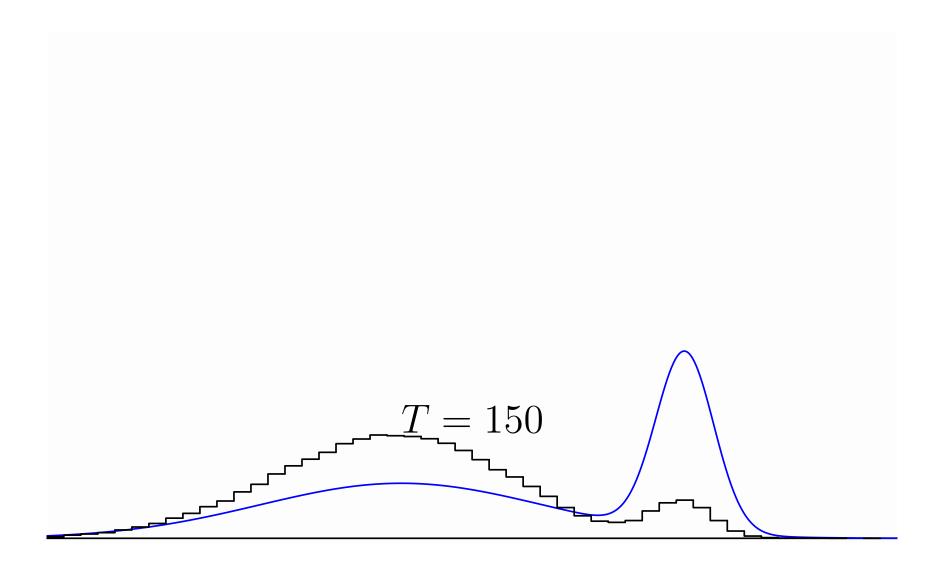


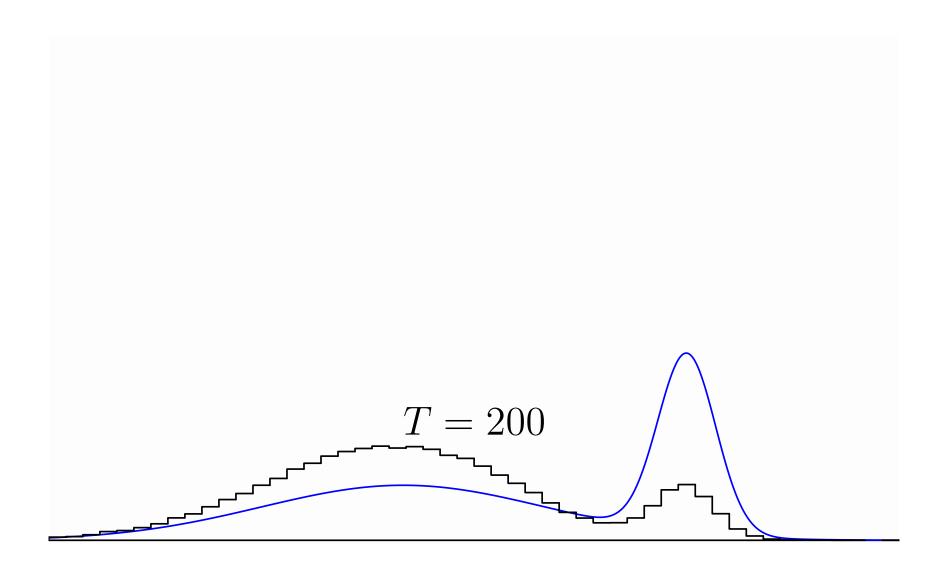


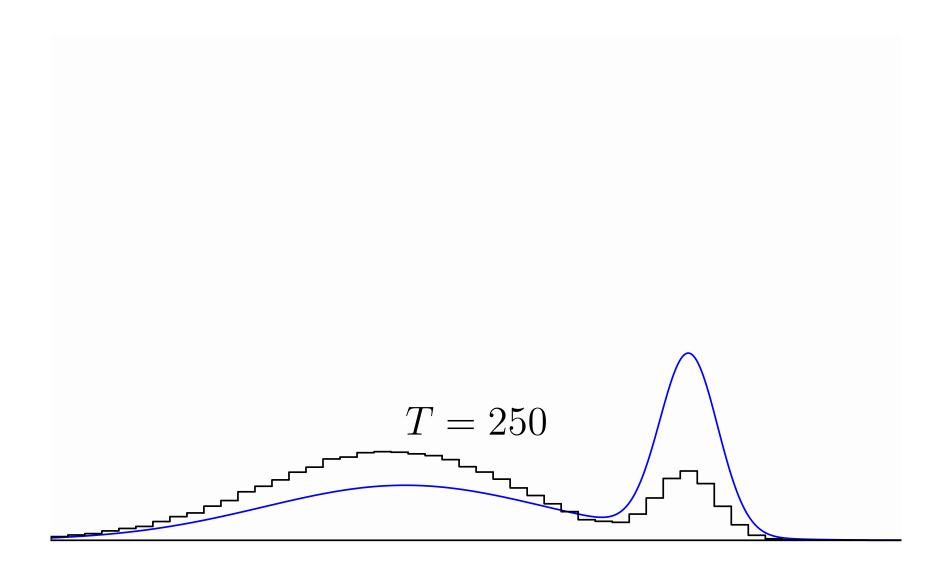


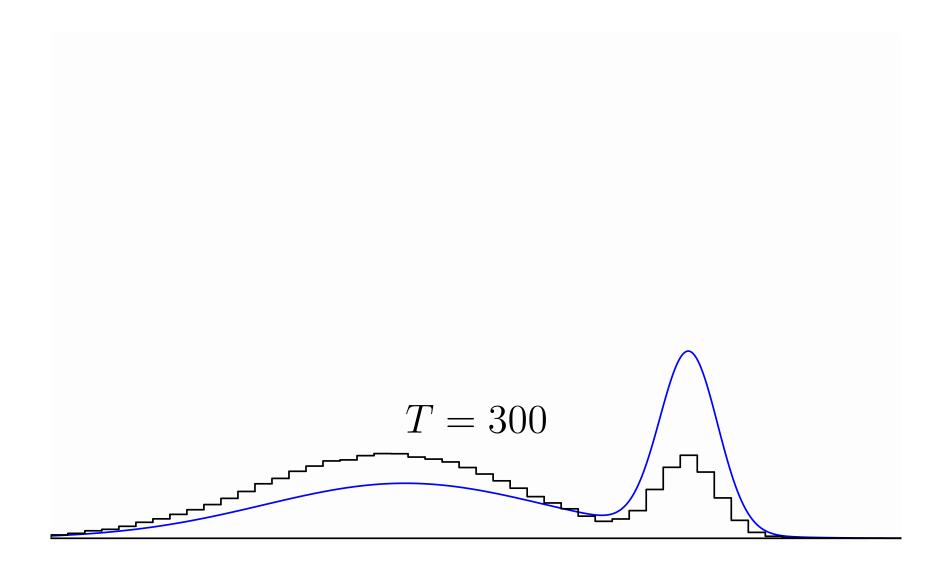


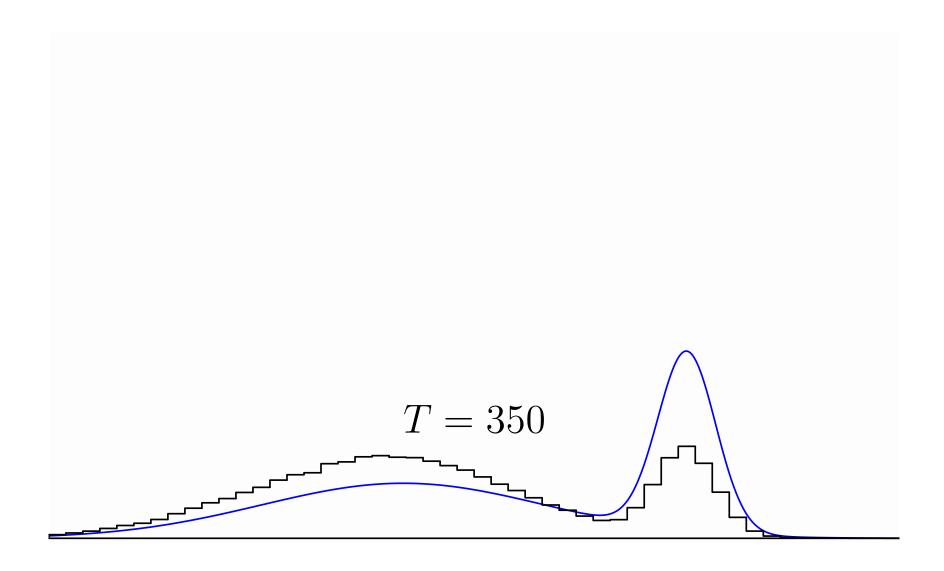


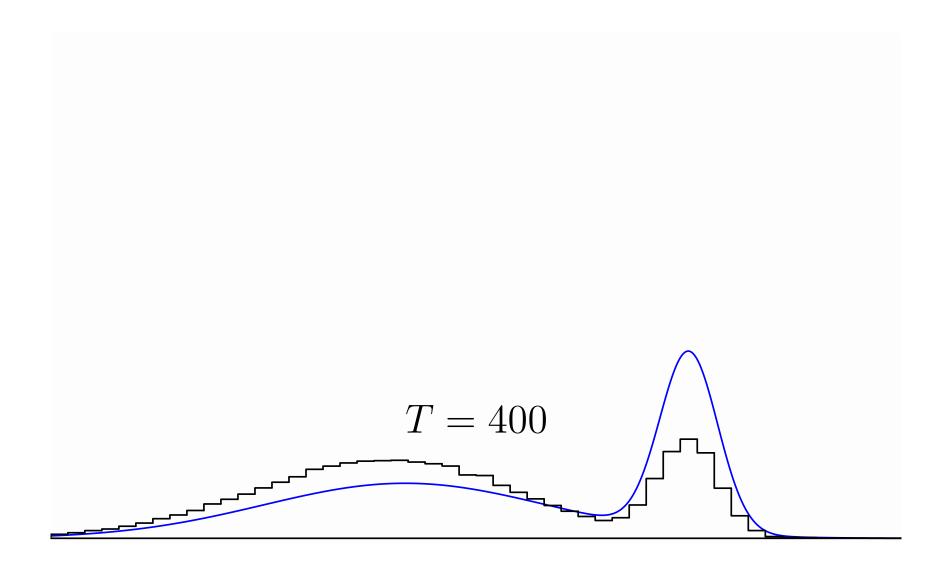


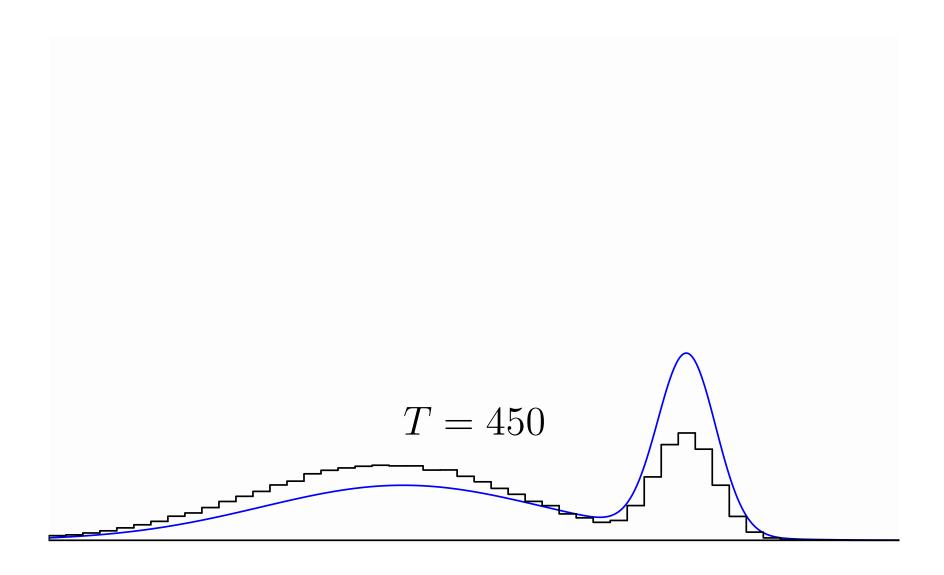


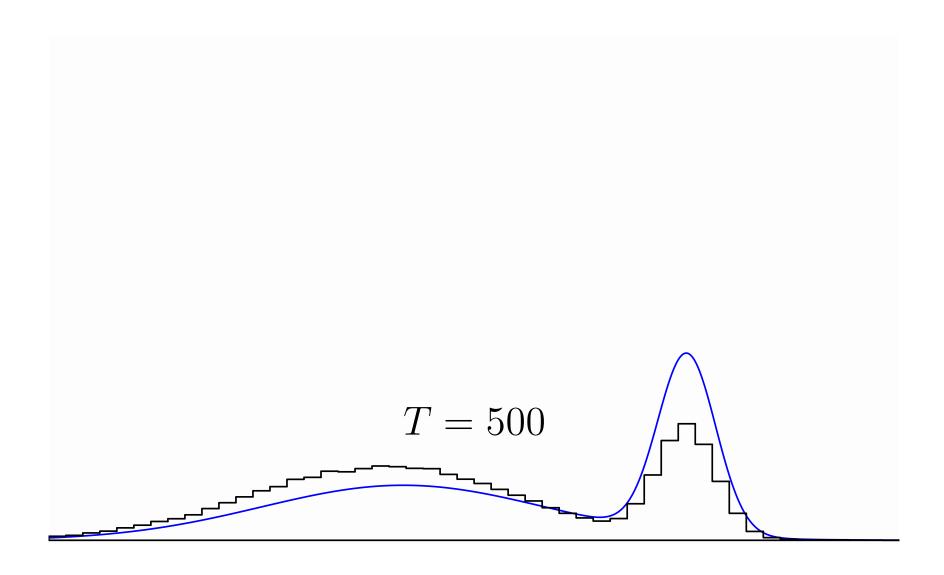


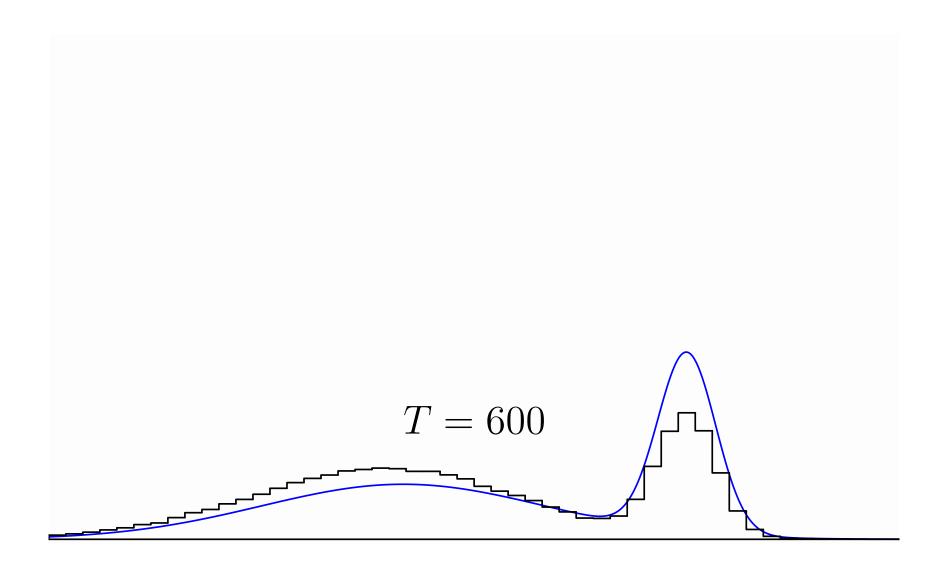


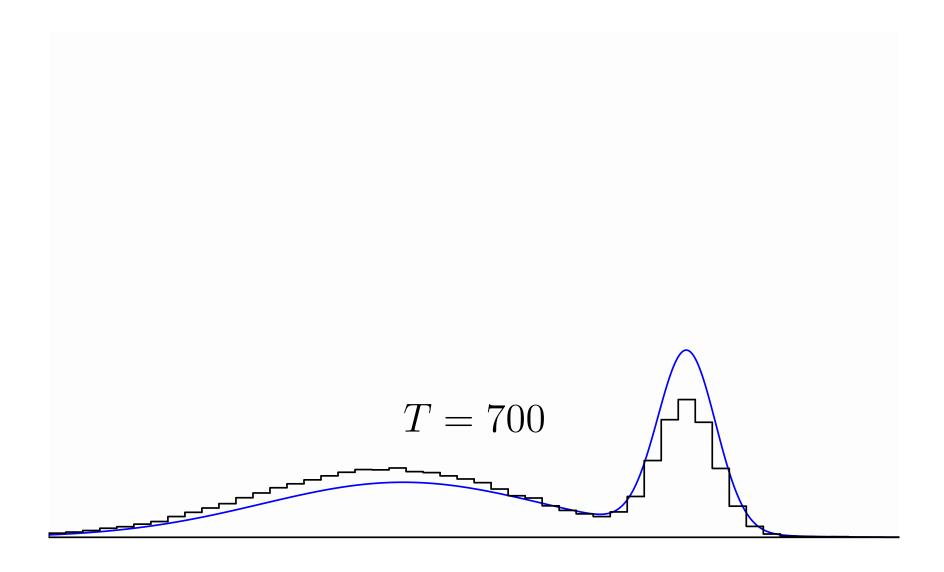


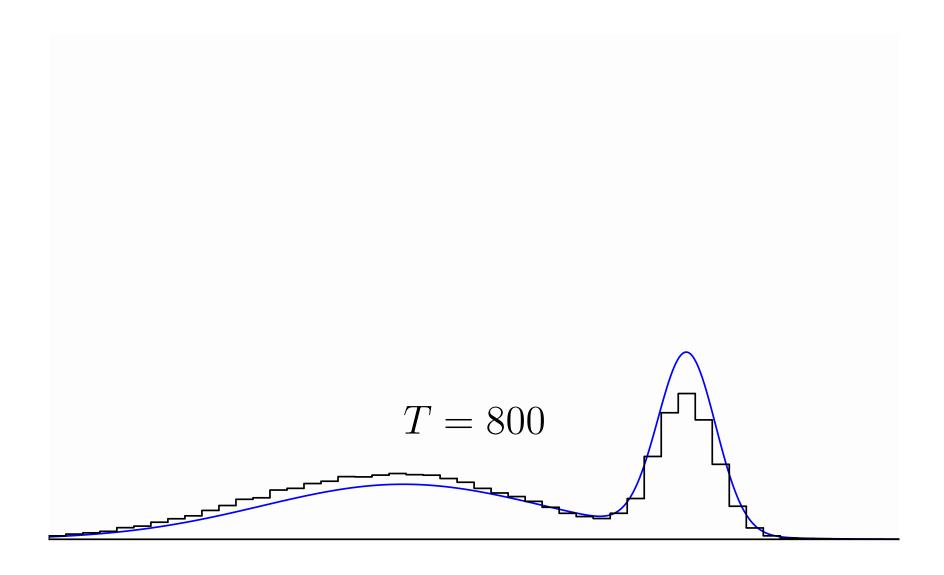


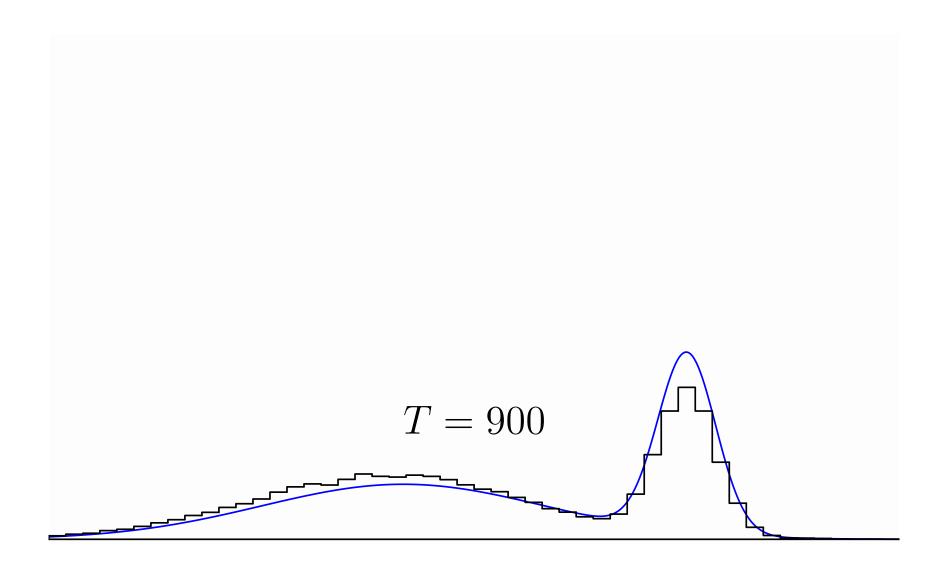


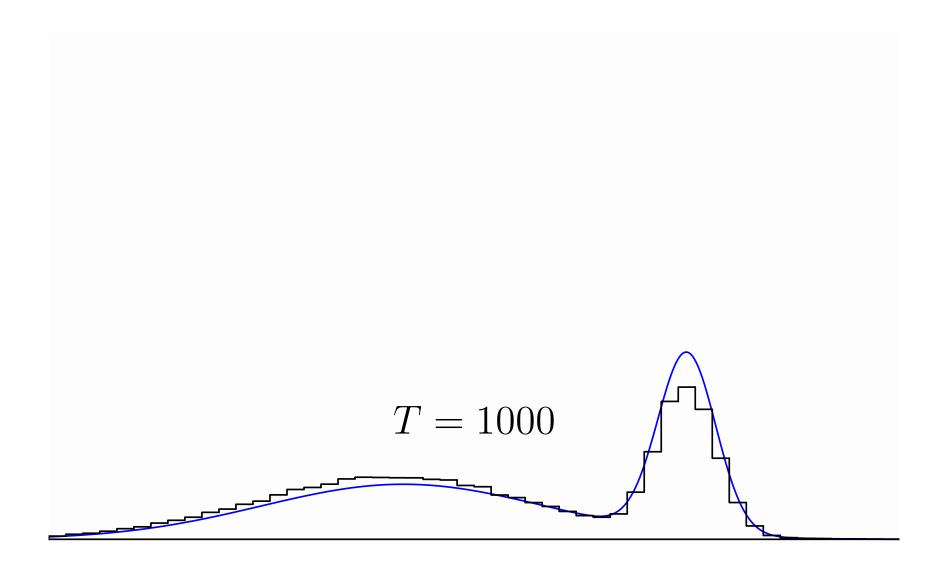


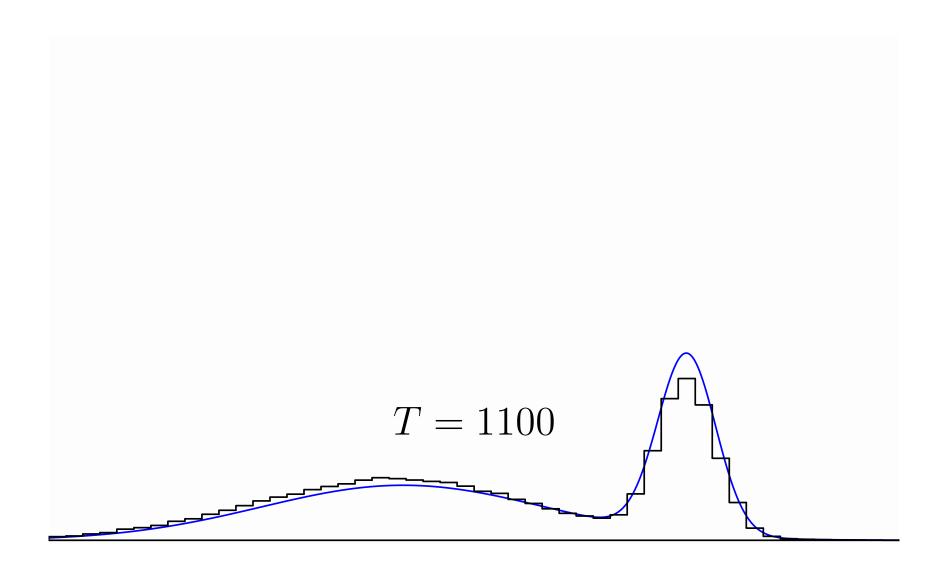


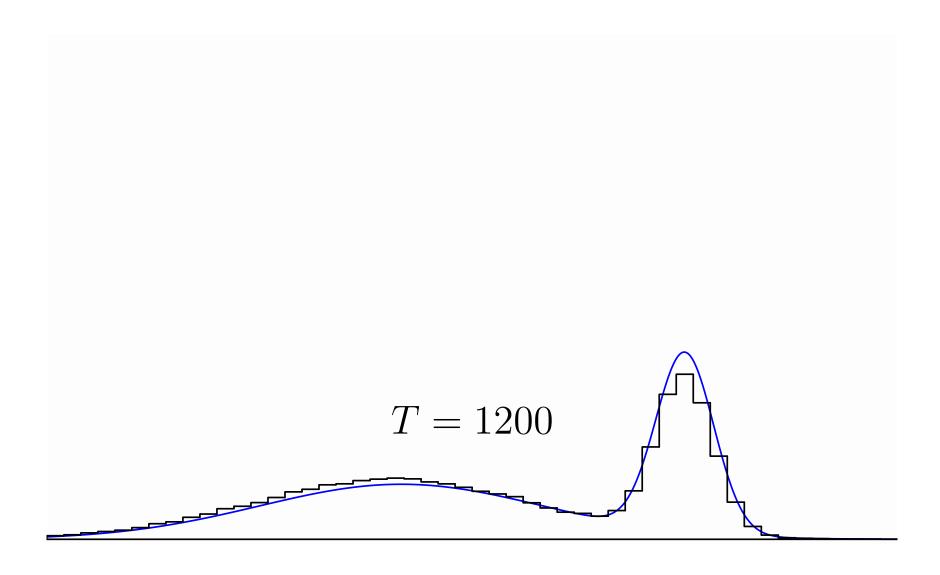


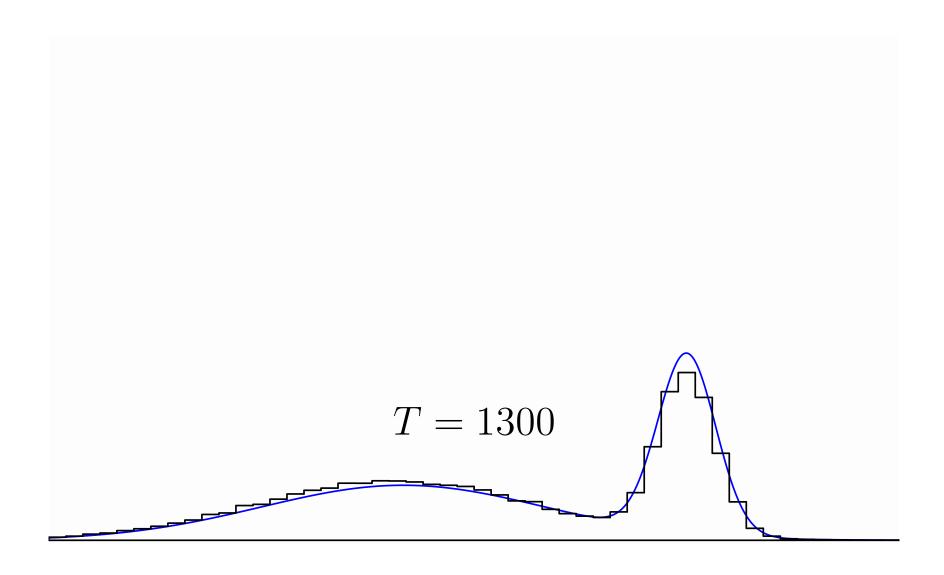


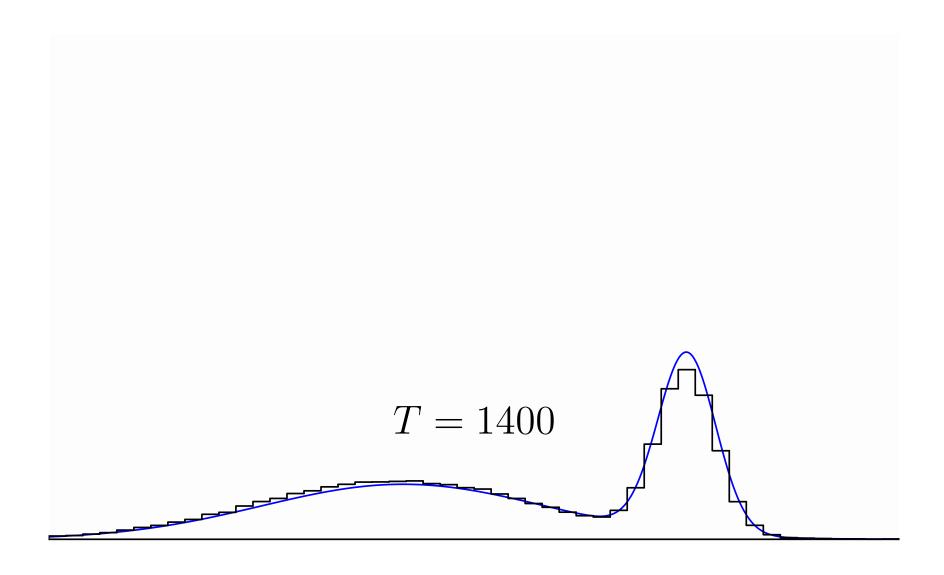


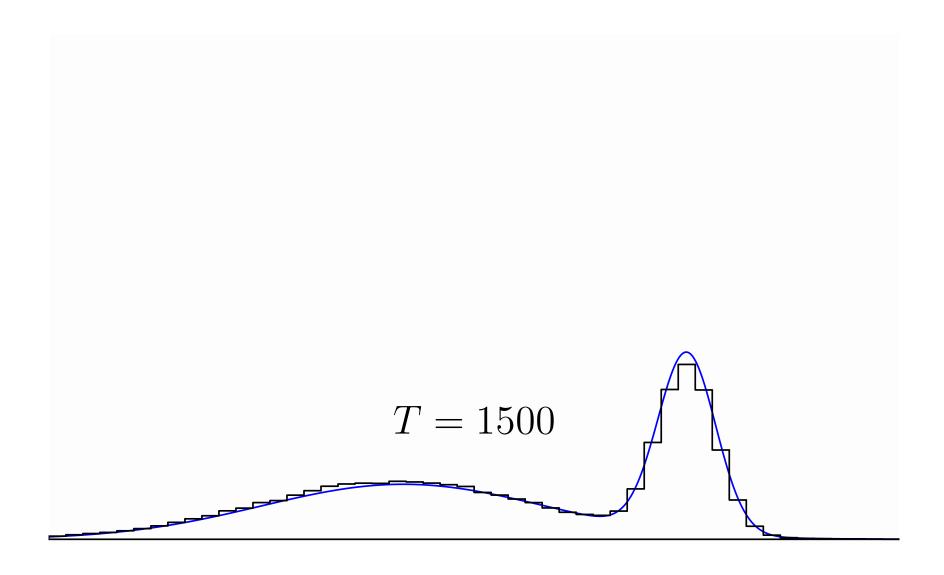


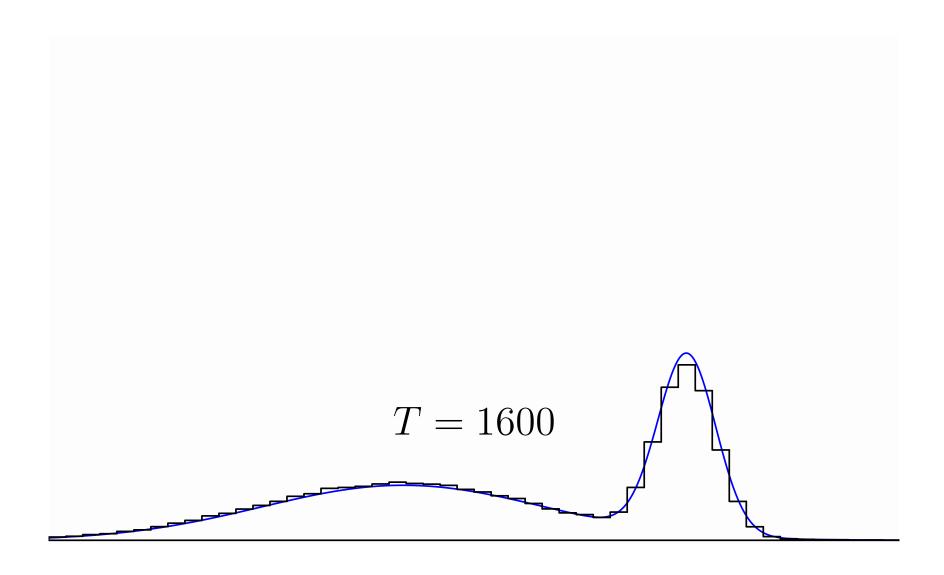


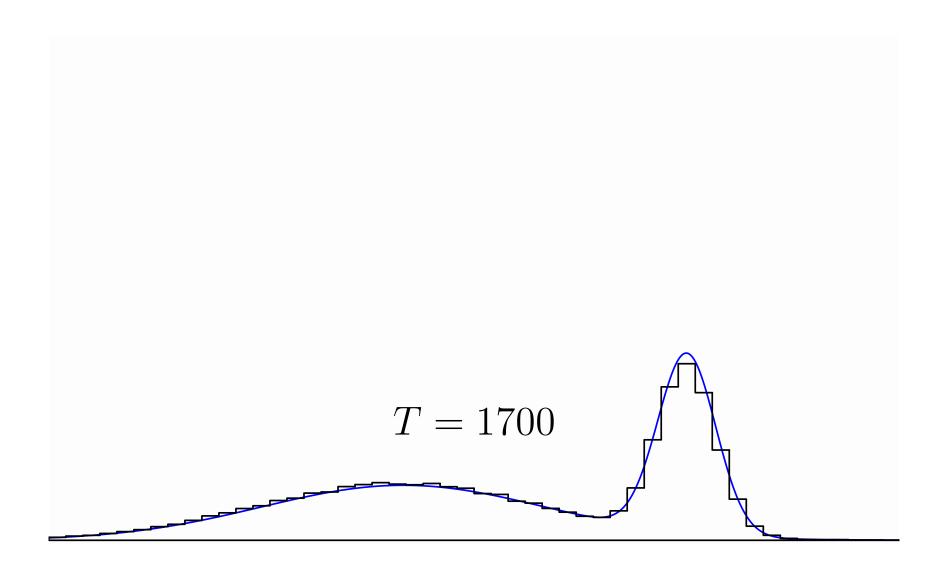


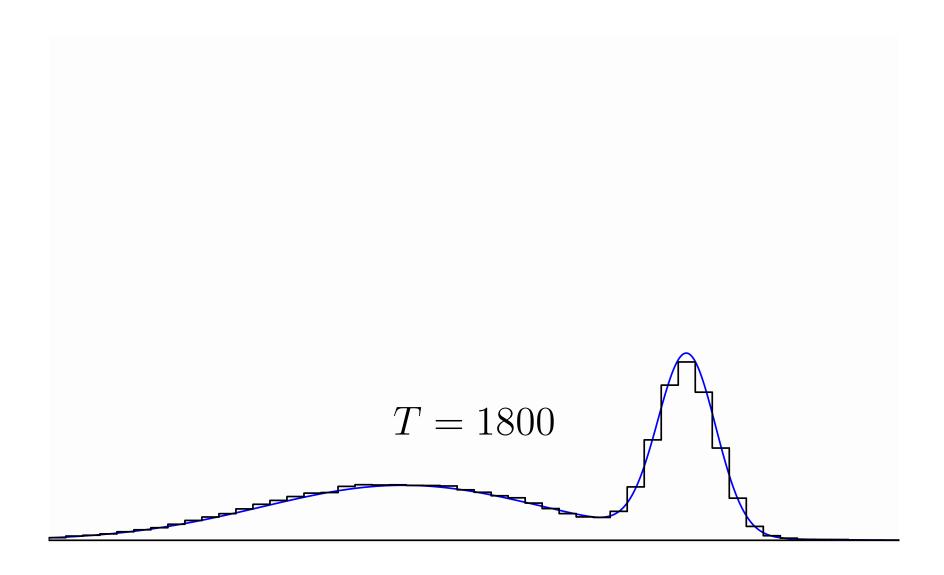


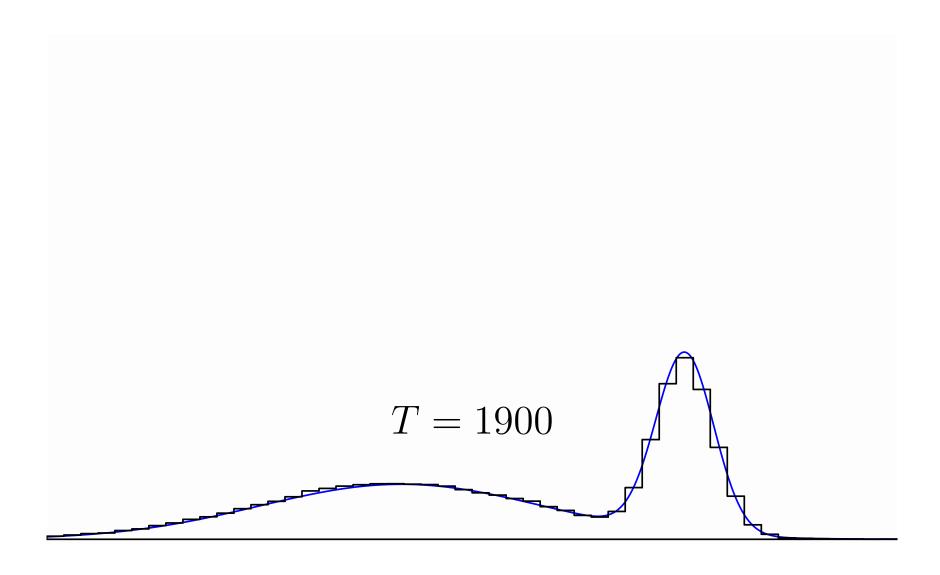


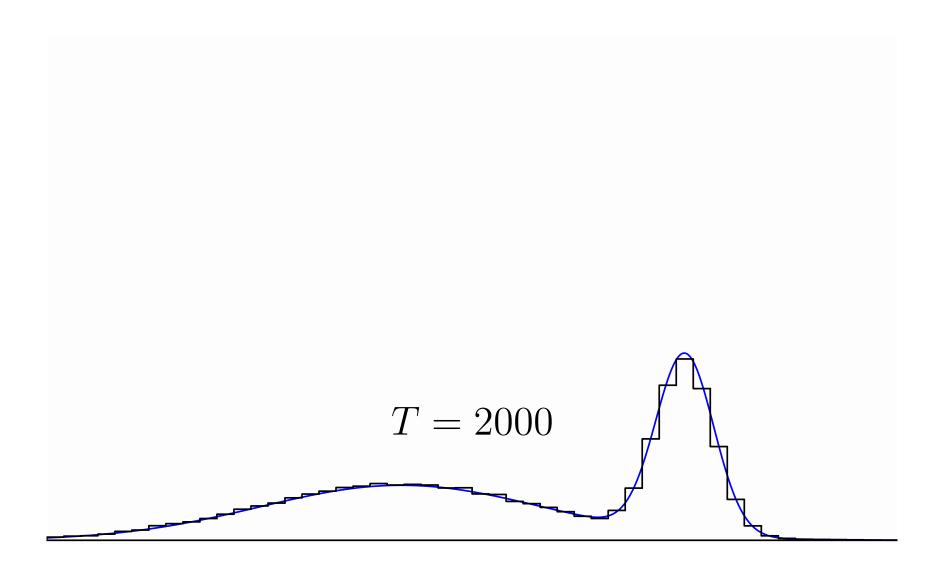












- We have some freedom in choosing a new proposal θ' from our current position θ —a good choice can increase the acceptance rate making the MCMC more efficient
- We define the proposal distribution $p(\boldsymbol{\theta}'|\boldsymbol{\theta})$
- For the standard Metropolis algorithm to work we require $p(\theta'|\theta) = p(\theta|\theta')$
- In some cases (e.g when $\theta_i \ge 0$) this can be hard to achieve
- We can modify our update rule to accept a move with probability

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Consider monitoring the flow of traffic where we have data

$$\mathcal{D} = (N_1, N_2, \dots, N_n)$$

- We assume $N_i \sim \operatorname{Poi}(\mu)$ and want to infer μ
- The Poisson distribution has a beta conjugate prior
- We don't have any prior knowledge on μ so we use a non-informative prior $\operatorname{Gam}(\mu|0,0)=1/\mu$
- Note that we can solve this problem exactly—however, lets compare with MCMC

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- Note that we can solve this problem exactly—however, lets compare with MCMC

Consider monitoring the flow of traffic where we have data

$$\mathcal{D} = (N_1, N_2, \dots, N_n)$$

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Proposal Distribution

- If we can choose our proposal distribution $p(\mu'|\mu)$ to be close to the posterior distribution then our acceptance rate would be close to 1
- We choose $p(\mu'|\mu) = \mathrm{Gam}(\mu'|\mu,\mu^2)$ which has $\mathbb{E}[\mu'] = \mu$ and variance 1
- We update with probability min(1,r) where

$$r = \frac{\operatorname{Gam}(\mu | \mu'^{2}, \mu') \frac{1}{\mu'} \prod_{i=1}^{n} \operatorname{Poi}(N_{i} | \mu')}{\operatorname{Gam}(\mu' | \mu^{2}, \mu) \frac{1}{\mu} \prod_{i=1}^{n} \operatorname{Poi}(N_{i} | \mu)}$$
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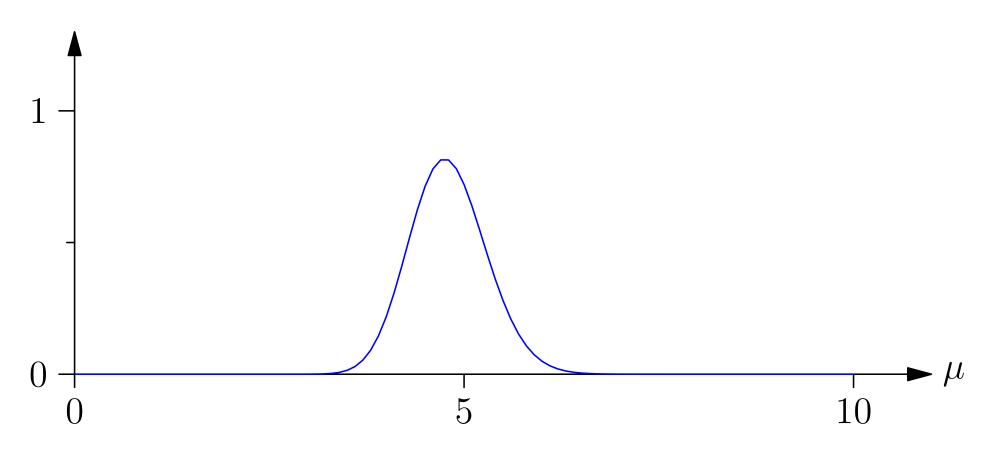
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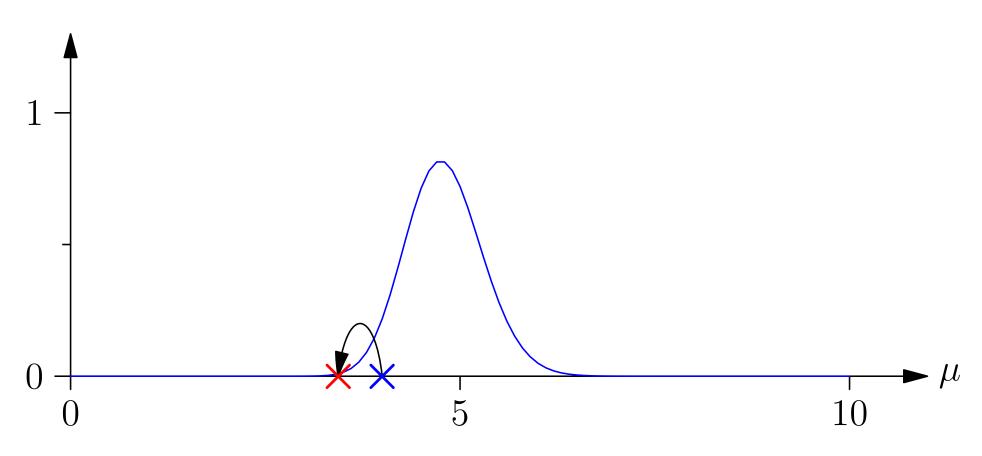
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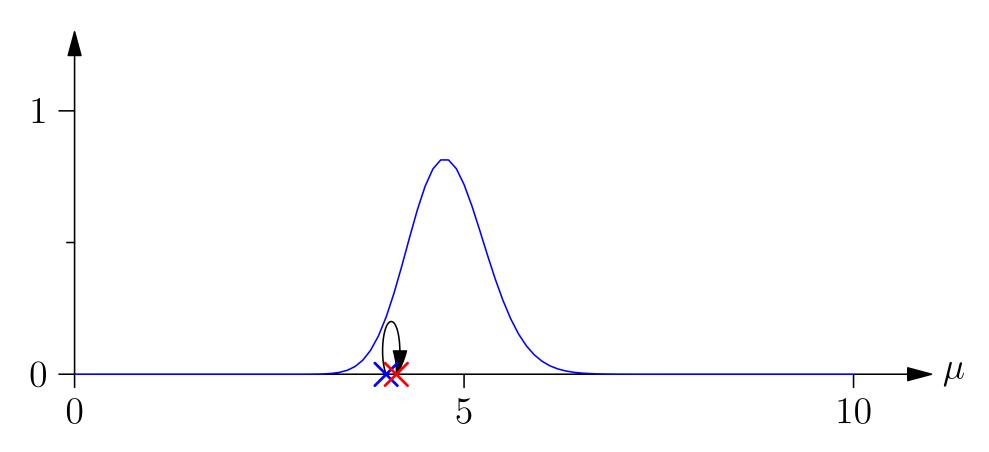
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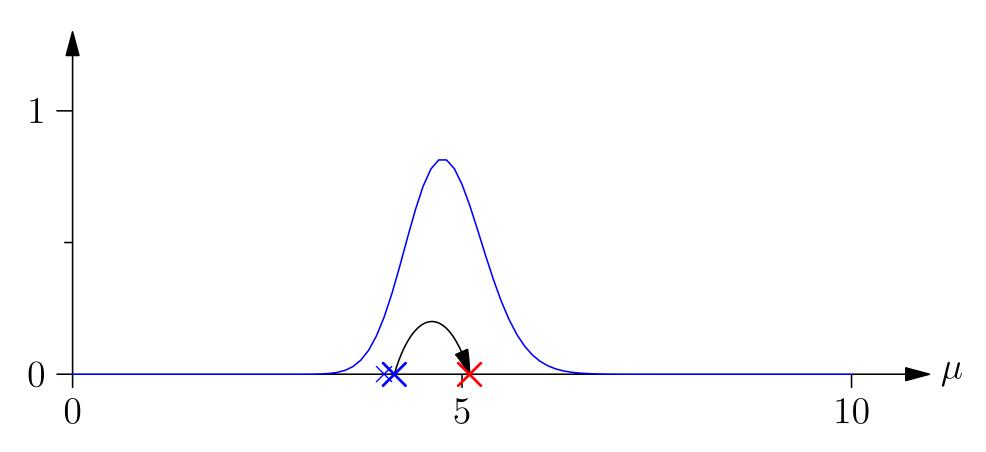
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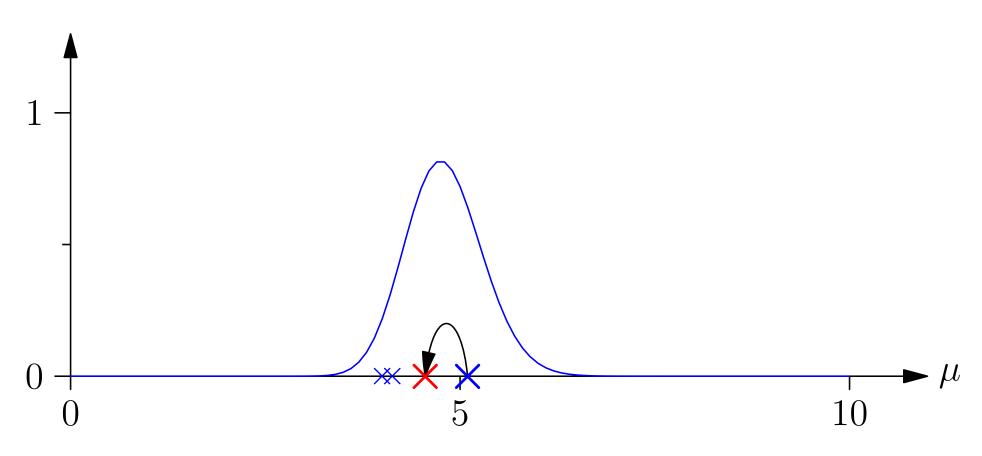
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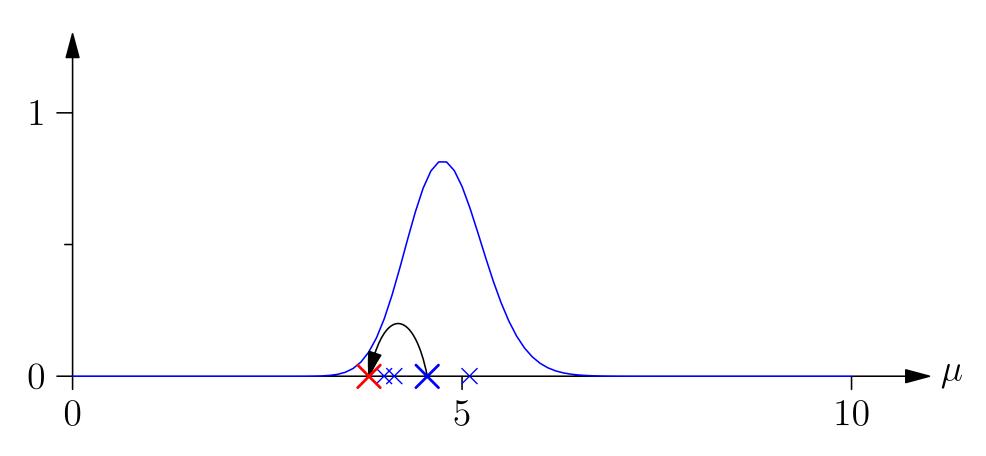


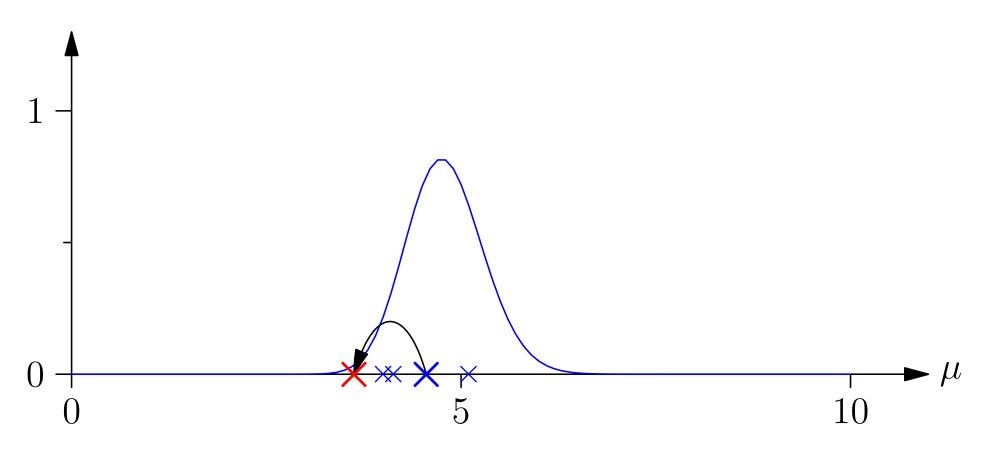


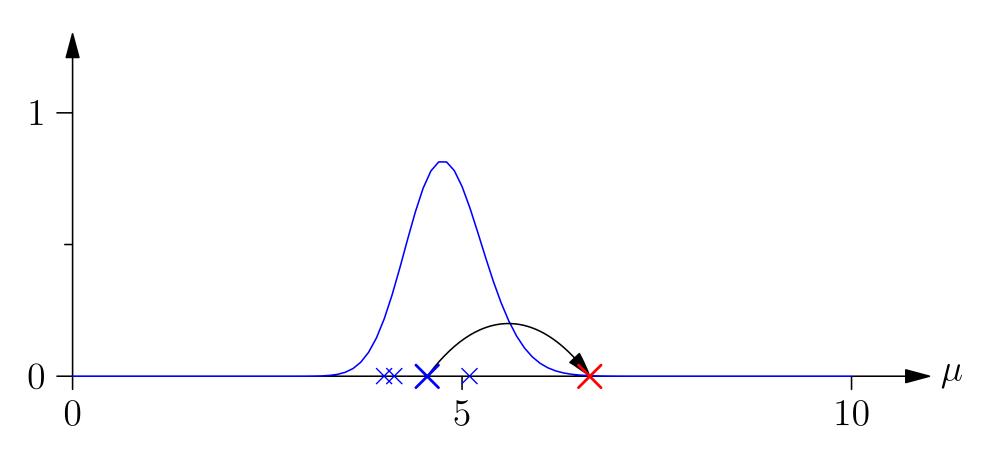


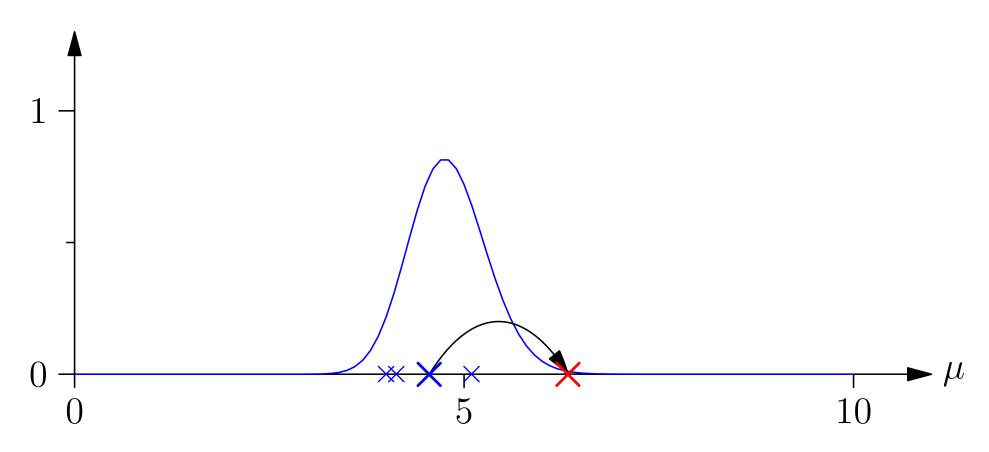


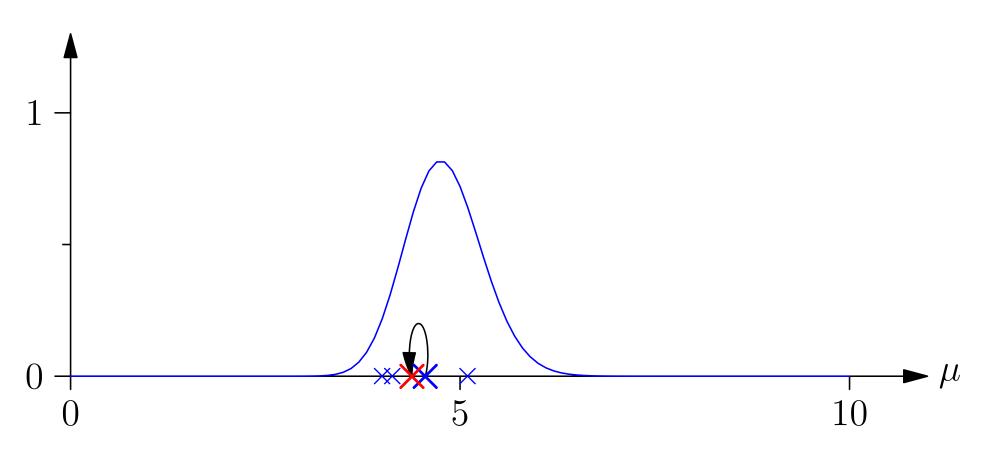


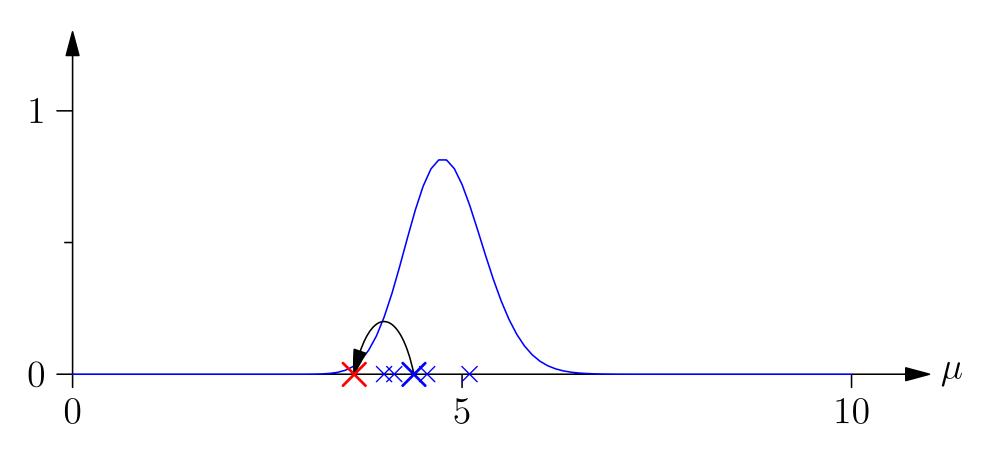


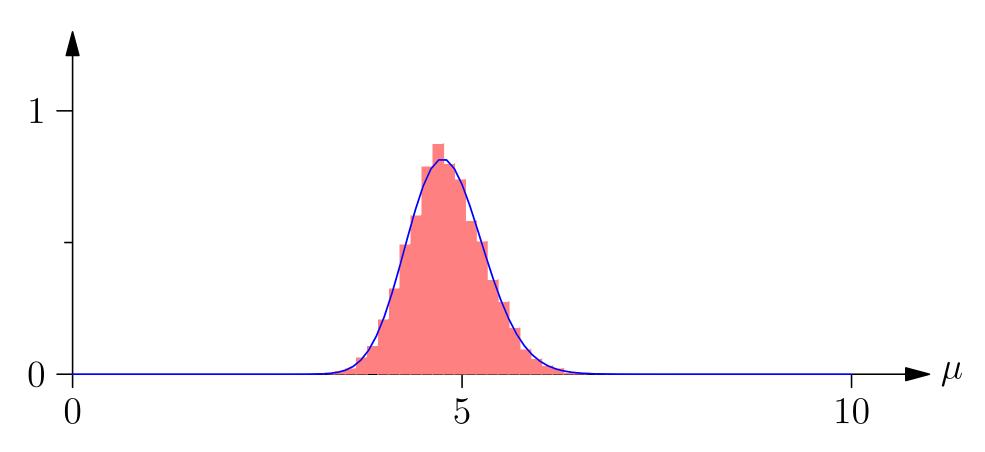












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