

Bayesian spatial modeling (contd.)

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Review of last lecture

- Using STAN package to run the MCMC for marginalized and latent spatial linear models
- MCMC diagnostics
- Spatial predictions from Bayesian output
- Model comparison using Bayesian output (DIC and WAIC)
- MCMC-free and fully Bayesian spatial analysis using spBayes package
- Hierarchical Bayesian spatial analysis beyond the spatial linear models
 - Example 1: Second-stage GP model for covariate imputation

Example 2: Discrete response

- In many applications, the spatial response of interest is binary or count-valued
- (Spatial LM for continuous data:) $E(y(s)) = x(s)' \beta + w(s)$
- (Spatial GLM for discrete data:) $g(E(y(s))) = x(s)' \beta + w(s)$
where g is a suitable link function (logit-link for binary data, log-link for count data)
- Other components stay the same, i.e., GP prior for $w(s)$, priors for β and covariance parameters

Spatial GLM for non-Gaussian response

- (Poisson Spatial GLM)

$$y(s) \stackrel{\text{iid}}{\sim} \text{Poisson}(\exp(x(s)'\beta + w(s)))$$

$$w = (w(s_1), \dots, w(s_n))' \sim N(0, C(\theta))$$

$$(\beta, \theta) \sim p(\beta, \theta)$$

- (Bernoulli Spatial GLM)

$$y(s) \stackrel{\text{iid}}{\sim} \text{Bernoulli}\left(\frac{\exp(x(s)'\beta + w(s))}{1 + \exp(x(s)'\beta + w(s))}\right)$$

$$w = (w(s_1), \dots, w(s_n))' \sim N(0, C(\theta))$$

$$(\beta, \theta) \sim p(\beta, \theta)$$

Eastern hemlock presence/absence data

- In many applications, the spatial response of interest is binary or count-valued

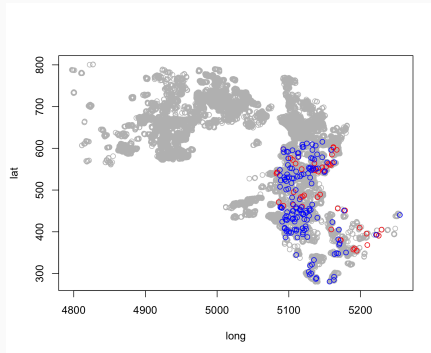


Figure: (Grey) All locations sampled for presence/absence of Eastern Hemlock species in Michigan. We will analyze a subset of locations with presence (red) and absence (blue).

- Covariates:
 - Minimum winter temperature (MIN)
 - Maximum summer temperature (MAX)
 - Total precipitation in the coldest quarter of the year (WIP)
 - Total precipitation in the warmest quarter of the year (SUP)
 - Annual actual evapotranspiration
 - Annual climatic water deficit (DEF) (AET)
- Model¹:

$$E(y(s)) = \text{logit} \left(\alpha + \beta_1 \text{MIN}(s) + \beta_2 \text{MAX}(s) + \right. \\ \left. \beta_3 \text{WIP}(s) + \beta_4 \text{SUP}(s) + \beta_5 \text{AET}(s) + \beta_6 \text{DEF}(s) + w(s) \right)$$

¹See <https://arxiv.org/pdf/2001.09111.pdf> for more details of the data and model. Dataset available in the spNNGP R-package.

Results using Stan

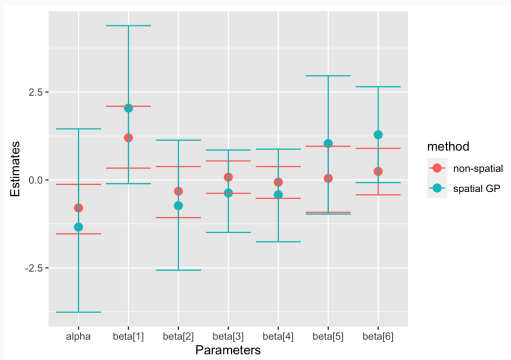


Figure: Regression coefficient estimates from spatial and non-spatial GLM.

Model	WAIC
Non-spatial GLM	206
Spatial GLM using GP	135

Table: Comparison of the spatial and non-spatial GLM.

Example 3: Spatially varying coefficient models

- The spatial linear model can be reparametrized as $y(s) = (\alpha + w(s)) + x(s)' \beta + \epsilon(s)$ where $\alpha + w(s)$ is the **spatially-varying intercept (SVI)** and the slope β is constant
- In many applications, the relationship between y and x can also vary over space and the assumption of a constant slope is inappropriate
- Such datasets can be modeled using **spatially varying coefficient (SVC)** models

Spatially varying coefficient models

- $y(s) = (\alpha + w_\alpha(s)) + x(s)'(\beta + w_\beta(s)) + \epsilon(s)$
- With p -covariates, $\beta + w_\beta(s)$ is a p -dimensional function modeling the **space-varying slope**
- $w(s) = (w_\alpha(s), w_\beta(s)')'$ is a $p + 1$ -dimensional function
- $w(s)$ can be modeled as either a collection of $p + 1$ -independent GPs or a $p + 1$ -dimensional **multivariate GP**
- `spSVC` function of the `spBayes` package offers fitting SVC model

Predicting air pollution in Central Europe

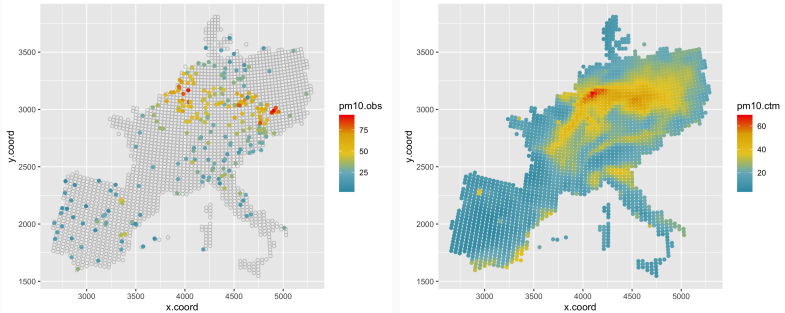


Figure: (Left) PM10 (in $\mu\text{g}/\text{m}^3$) measured at regulatory stations. (Right) PM10 output (in $\mu\text{g}/\text{m}^3$) from a chemistry transport model (CTM) at a grid over the area.

- Goal: Use the **high resolution** CTM output to predict PM10 at all the grid locations

Exploratory analysis

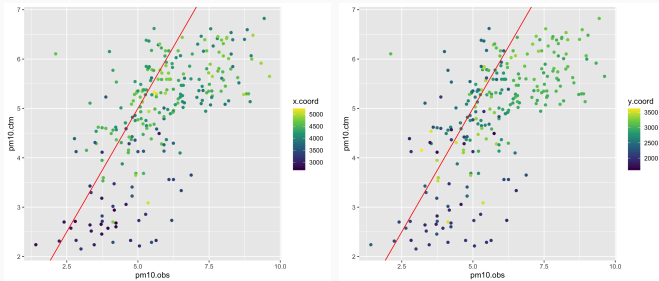


Figure: Scatterplots of **square-root transformed** PM10 measured by regulatory devices and predicted by CTM color-coded by longitude (left) and latitude (right).

Scatterplots reveal

- General under-prediction by CTM
- Variability in the relationship over space

$$\sqrt{PM10(s)} = \beta_0 + w_0(s) + \sqrt{CTM(s)}(\beta_1 + w_1(s)) + \epsilon(s)$$

$$w_i(\cdot) \sim GP(0, \sigma_i^2 R(\cdot, \cdot | \phi_i)), \text{ for } i = 0, 1$$

$$\epsilon(s) \stackrel{\text{iid}}{\sim} N(0, \tau^2)$$

$$\theta = (\beta_0, \sigma_0^2, \sigma_1^2, \phi_0, \phi_1, \tau^2)' \sim p(\theta)$$

Results

Parameters	50%	2.5%	97.5%
β_0	3.177	2.061	4.313
β_1	0.329	-0.085	0.737
σ_0^2	0.291	0.142	0.470
σ_1^2	0.104	0.070	0.151
ϕ_0	0.411	0.062	0.954
ϕ_1	0.002	0.001	0.002
τ^2	0.274	0.133	0.459

Results

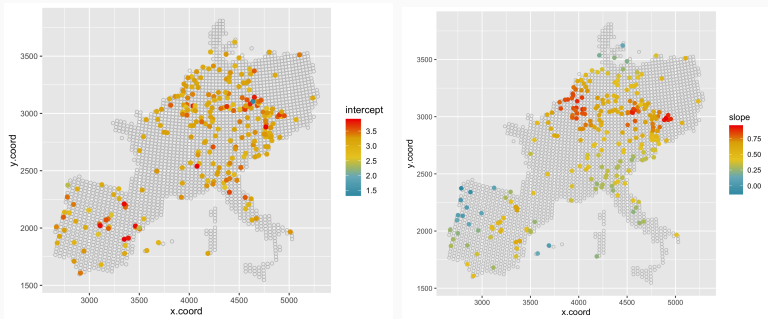


Figure: Estimates of the spatially-varying intercept (left) and spatially-varying slope (right) for the locations with regulatory data.

Results

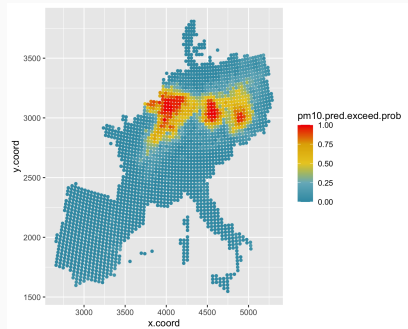
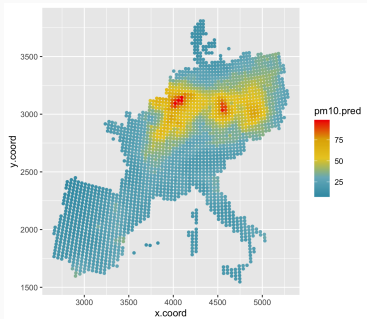


Figure: Predictions of PM10 (left) and predicted $P(\text{PM10} > 50 \mu\text{g}/\text{m}^3)$ (right) for the entire grid .

- More examples of hierarchical spatial modeling:
 - Example 2: Spatial GLM for binary or count spatial data
 - Example 3: Spatially-varying coefficient model for heterogenous covariate effects