  - Sample $\boldsymbol{x}_{\eta_2}^{(k)} \sim T_{\eta_1}(\mathbf{x}_{\eta_2}^{(k)} \mid \boldsymbol{x}_{\eta_1}^{(k)})$
  - ...
  - Sample $\boldsymbol{x}_{\eta_{n-1}}^{(k)} \sim T_{\eta_{n-2}}(\mathbf{x}_{\eta_{n-1}}^{(k)} \mid \boldsymbol{x}_{\eta_{n-2}}^{(k)})$
  - Sample $\boldsymbol{x}_{\eta_n}^{(k)} \sim T_{\eta_{n-1}}(\mathbf{x}_{\eta_n}^{(k)} \mid \boldsymbol{x}_{\eta_{n-1}}^{(k)})$
- end

For sample k, we can derive the importance weight by chaining together the importance weights for the jumps between the intermediate distributions given in equation 18.49:

$$w^{(k)} = \frac{\tilde{p}_{\eta_1}(\boldsymbol{x}_{\eta_1}^{(k)})}{\tilde{p}_0(\boldsymbol{x}_{\eta_1}^{(k)})} \frac{\tilde{p}_{\eta_2}(\boldsymbol{x}_{\eta_2}^{(k)})}{\tilde{p}_{\eta_1}(\boldsymbol{x}_{\eta_2}^{(k)})} \cdots \frac{\tilde{p}_1(\boldsymbol{x}_1^{(k)})}{\tilde{p}_{\eta_{n-1}}(\boldsymbol{x}_{\eta_n}^{(k)})}. \tag{18.52}$$

To avoid numerical issues such as overflow, it is probably best to compute $\log w^{(k)}$ by adding and subtracting log probabilities, rather than computing $w^{(k)}$ by multiplying and dividing probabilities.

With the sampling procedure thus defined and the importance weights given in equation 18.52, the estimate of the ratio of partition functions is given by:

$$\frac{Z_1}{Z_0} \approx \frac{1}{K} \sum_{k=1}^{K} w^{(k)} \tag{18.53}$$

In order to verify that this procedure defines a valid importance sampling scheme, we can show (Neal, 2001) that the AIS procedure corresponds to simple importance sampling on an extended state space with points sampled over the product space $[\boldsymbol{x}_{\eta_1}, \ldots, \boldsymbol{x}_{\eta_{n-1}}, \boldsymbol{x}_1]$. To do this, we define the distribution over the extended space as:

$$\tilde{p}(\boldsymbol{x}_{\eta_1}, \ldots, \boldsymbol{x}_{\eta_{n-1}}, \boldsymbol{x}_1) \tag{18.54}$$

$$= \tilde{p}_1(\boldsymbol{x}_1) \tilde{T}_{\eta_{n-1}}(\boldsymbol{x}_{\eta_{n-1}} \mid \boldsymbol{x}_1) \tilde{T}_{\eta_{n-2}}(\boldsymbol{x}_{\eta_{n-2}} \mid \boldsymbol{x}_{\eta_{n-1}}) \ldots \tilde{T}_{\eta_1}(\boldsymbol{x}_{\eta_1} \mid \boldsymbol{x}_{\eta_2}), \tag{18.55}$$

where $\tilde{T}_a$ is the reverse of the transition operator defined by T_a (via an application of Bayes' rule):

$$\tilde{T}_a(\boldsymbol{x}' \mid \boldsymbol{x}) = \frac{p_a(\boldsymbol{x}')}{p_a(\boldsymbol{x})} T_a(\boldsymbol{x} \mid \boldsymbol{x}') = \frac{\tilde{p}_a(\boldsymbol{x}')}{\tilde{p}_a(\boldsymbol{x})} T_a(\boldsymbol{x} \mid \boldsymbol{x}'). \tag{18.56}$$

Plugging the above into the expression for the joint distribution on the extended state space given in equation 18.55, we get:

$$\tilde{p}(\boldsymbol{x}_{\eta_1}, \ldots, \boldsymbol{x}_{\eta_{n-1}}, \boldsymbol{x}_1) \tag{18.57}$$