

Spatial Analysis and Modeling (GIST 4302/5302)

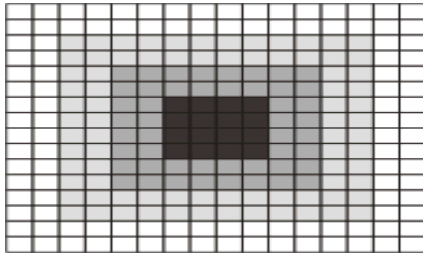
Guofeng Cao

Department of Geosciences

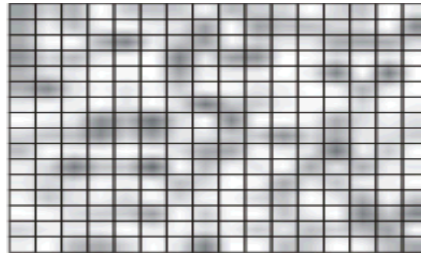
Texas Tech University

Spatial Autocorrelation

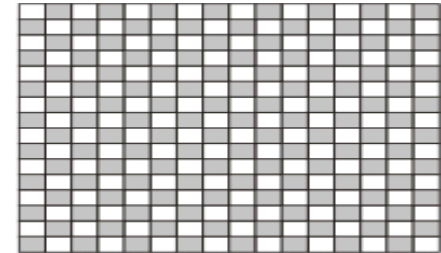
- Tobler's first law of geography
- Spatial auto/cross correlation



If like values tend to cluster together, then the field exhibits high **positive spatial autocorrelation**



If there is no apparent relationship between attribute value and location then there is **zero spatial autocorrelation**



If like values tend to be located away from each other, then there is **negative spatial autocorrelation**

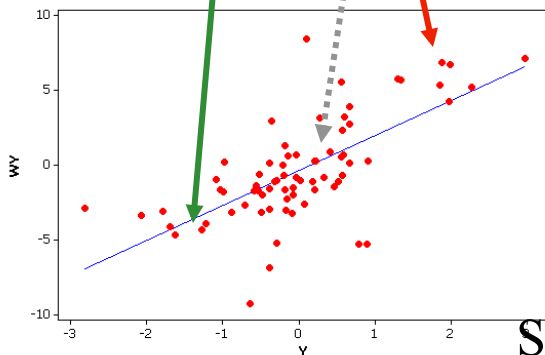
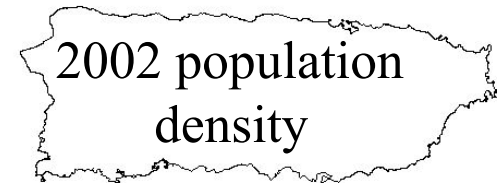
Spatial Autocorrelation

- Spatial autocorrelation is everywhere
 - Spatial point pattern
 - K, F, G functions
 - Kernel functions
 - Areal/lattice (this topic)
 - Geostatistical data (next topic)

Spatial Autocorrelation of Areal Data

Positive spatial autocorrelation

- high values surrounded by nearby high values
- intermediate values surrounded by nearby intermediate values
- low values surrounded by nearby low values

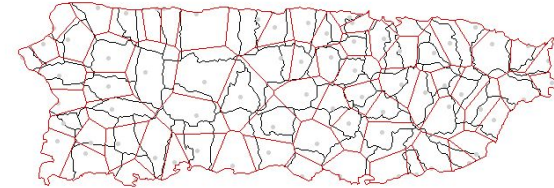


Source: Ron Briggs of UT Dallas

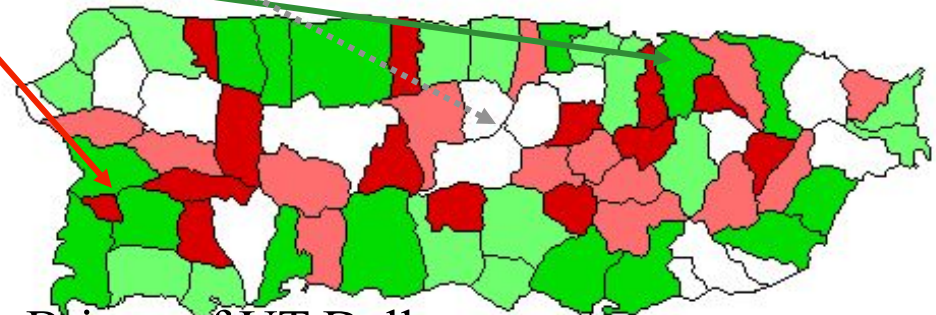
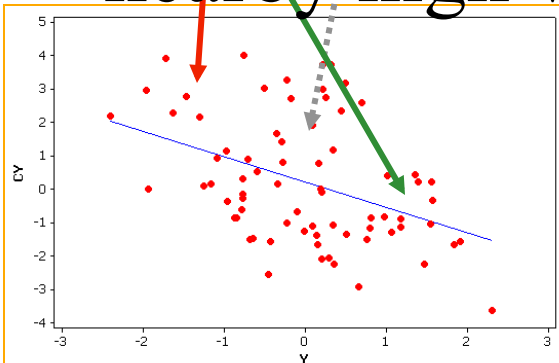
Negative spatial autocorrelation

- high values surrounded by nearby low values
- intermediate values surrounded by nearby intermediate values
- low values surrounded by nearby high values

competition for space



Grocery store density



Source: Ron Briggs of UT Dallas

Measuring Spatial Autocorrelation: the problem of measuring “nearness”

To measure spatial autocorrelation, we must know the “nearness” of our observations as we did for point pattern case

- Which points or polygons are “near” or “next to” other points or polygons?

– *Which states are near Texas?*

– How to measure this?

Seems simple and obvious,
but it is not!



Spatial Weight Matrix

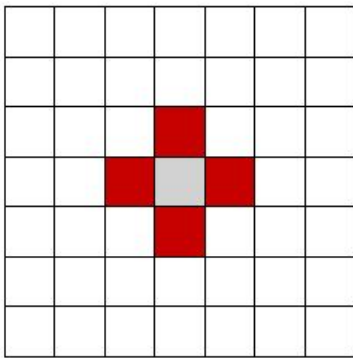
- **Core** concept in statistical analysis of areal data
- Two steps involved:
 - define which relationships between observations are to be given a nonzero weight, i.e., define spatial neighbors
 - assign weights to the neighbors

Spatial Neighbors

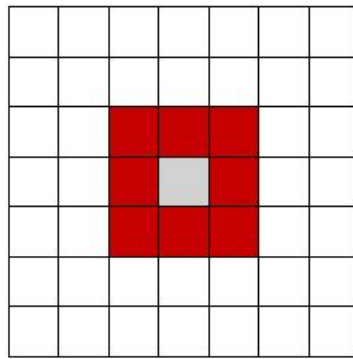
- **Contiguity-based neighbors**
 - Zone i and j are neighbors if zone i is contiguity or adjacent to zone j
 - But what constitutes contiguity?
- **Distance-based neighbors**
 - Zone i and j are neighbors if the distance between them are less than the threshold distance
 - But what distance do we use?

Contiguity-based Spatial Neighbors

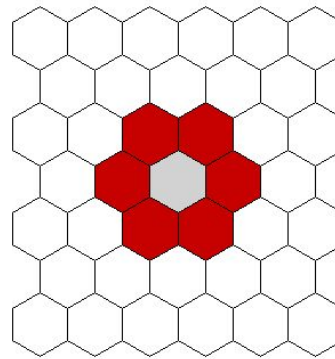
- Sharing a border or boundary
 - Rook: sharing a border
 - Queen: sharing a border or a point



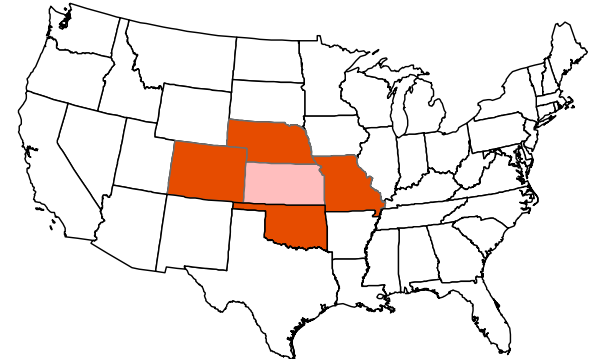
rook



queen



Hexagons



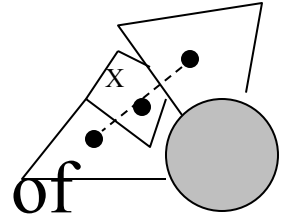
Irregular

Which use?

Problem Situations for Irregular Polygons

“Close” but no common border

- Include polygons which have a centroid within the “convex hull” for the centroids of polygons that do share a common border



Length of border

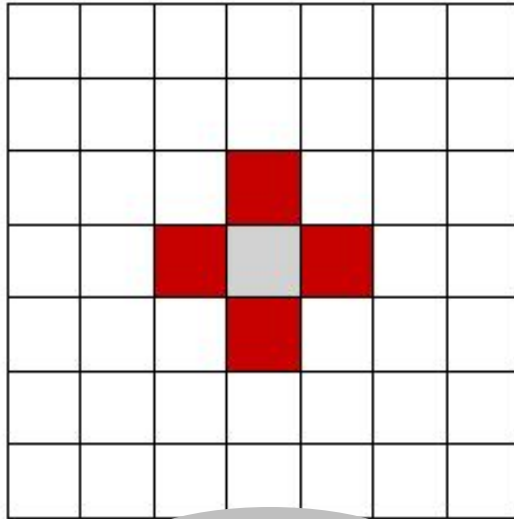
- Is Shanxi “as close to” Nei Mongol as to Henan?
- Base “closeness” on proportion of shared border, not just one (1) or zero (0)
- $w_{ij} = \text{border length}_{ij} / \text{border length}_j$



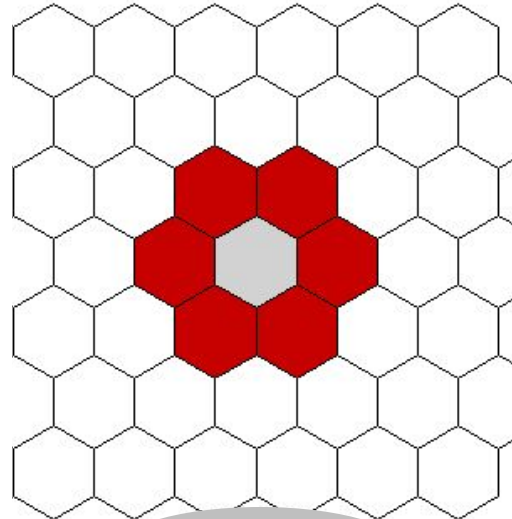
Higher-Order Contiguity

1st
order

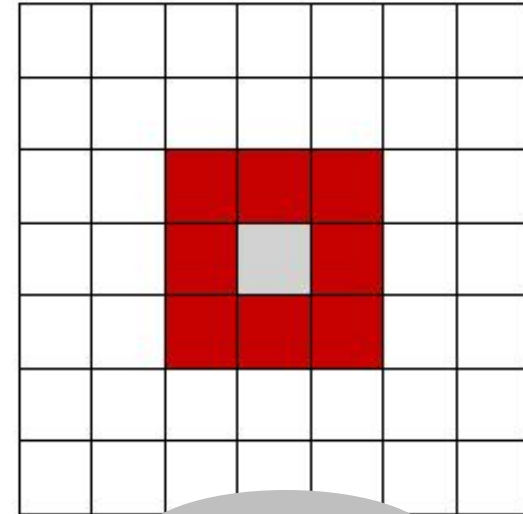
Nearest
neighbor



rook



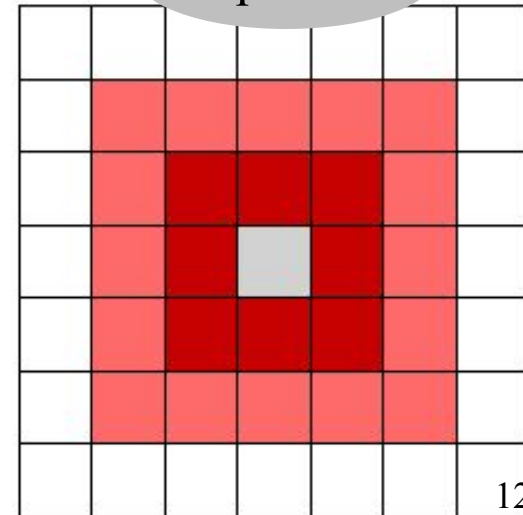
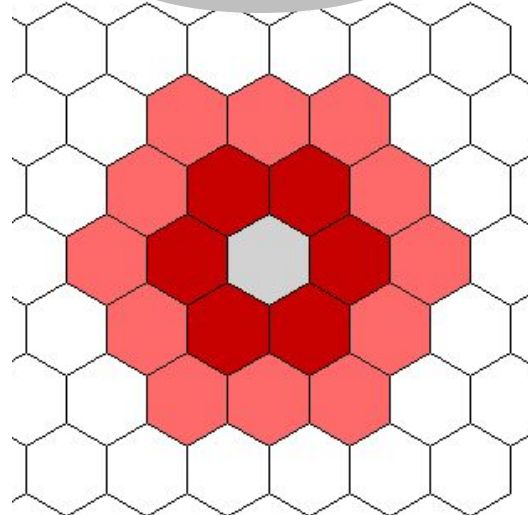
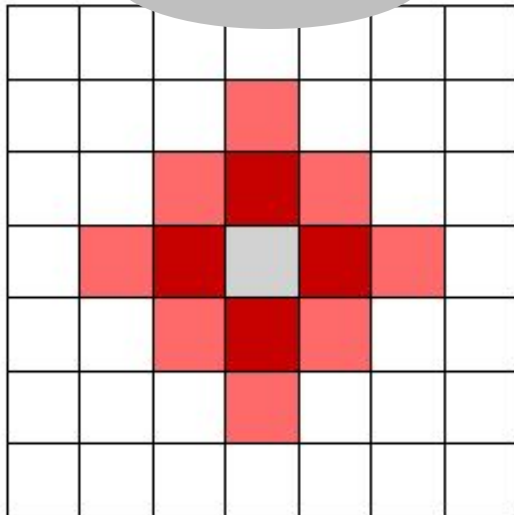
hexagon



queen

2nd
order

Next
nearest
neighbor



Distance-based Neighbors

- How to measure distance between polygons?
- Distance metrics
 - 2D Cartesian distance (projected data)
 - 3D spherical distance/great-circle distance (lat/long data)
 - Haversine formula

Haversine $a = \sin^2(\Delta\phi/2) + \cos(\phi_1) \cdot \cos(\phi_2) \cdot \sin^2(\Delta\lambda/2)$

formula: $c = 2 \cdot \text{atan2}(\sqrt{a}, \sqrt{1-a})$

$d = R \cdot c$

where ϕ is latitude, λ is longitude, R is earth's radius (mean radius = 6,371km)

Distance-based Neighbors

- k -nearest neighbors

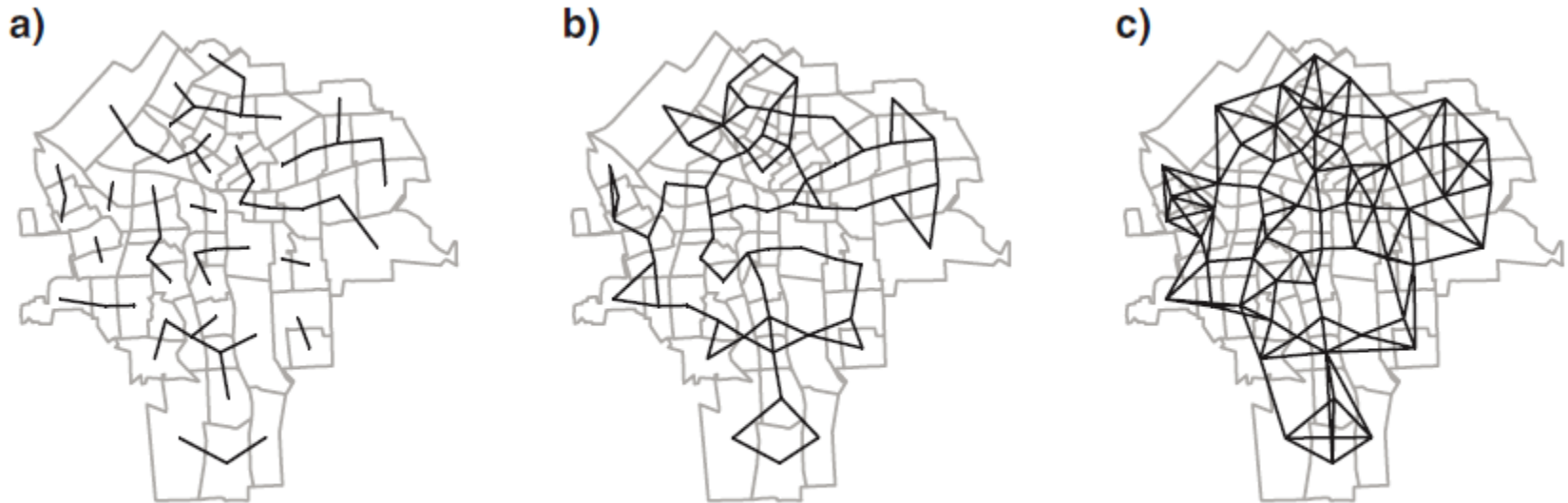


Fig. 9.5. (a) $k = 1$ neighbours; (b) $k = 2$ neighbours; (c) $k = 4$ neighbours

Source: Bivand and Pebesma and Gomez-Rubio

Distance-based Neighbors

- thresh-hold distance (buffer)

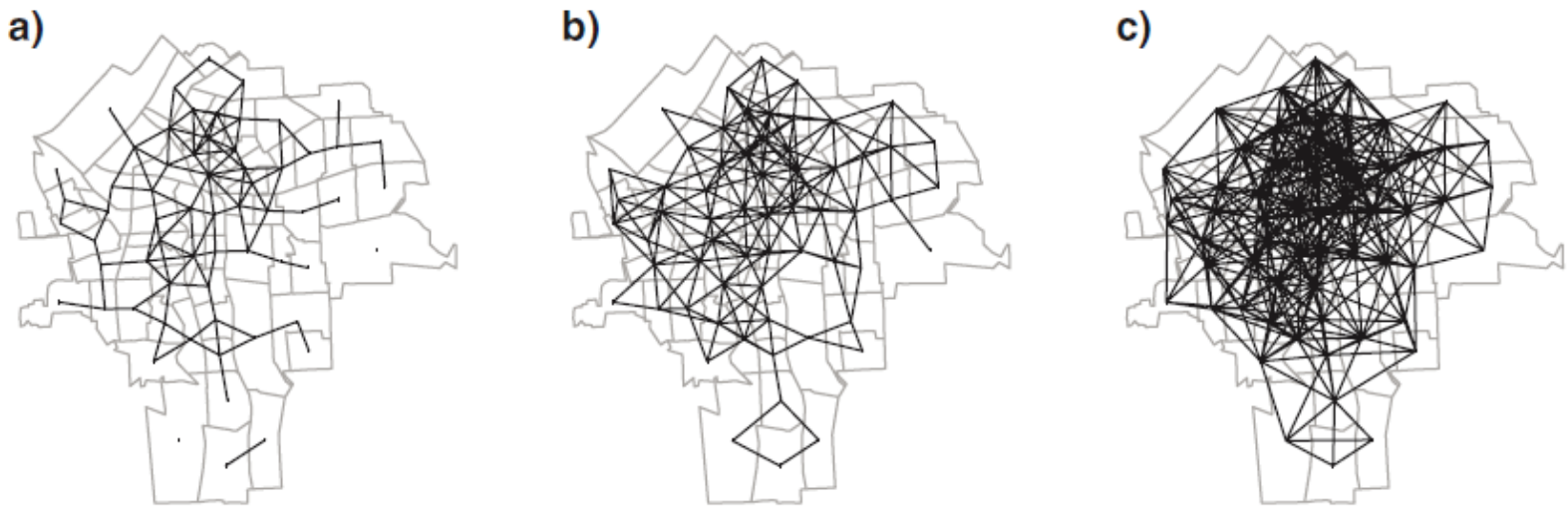
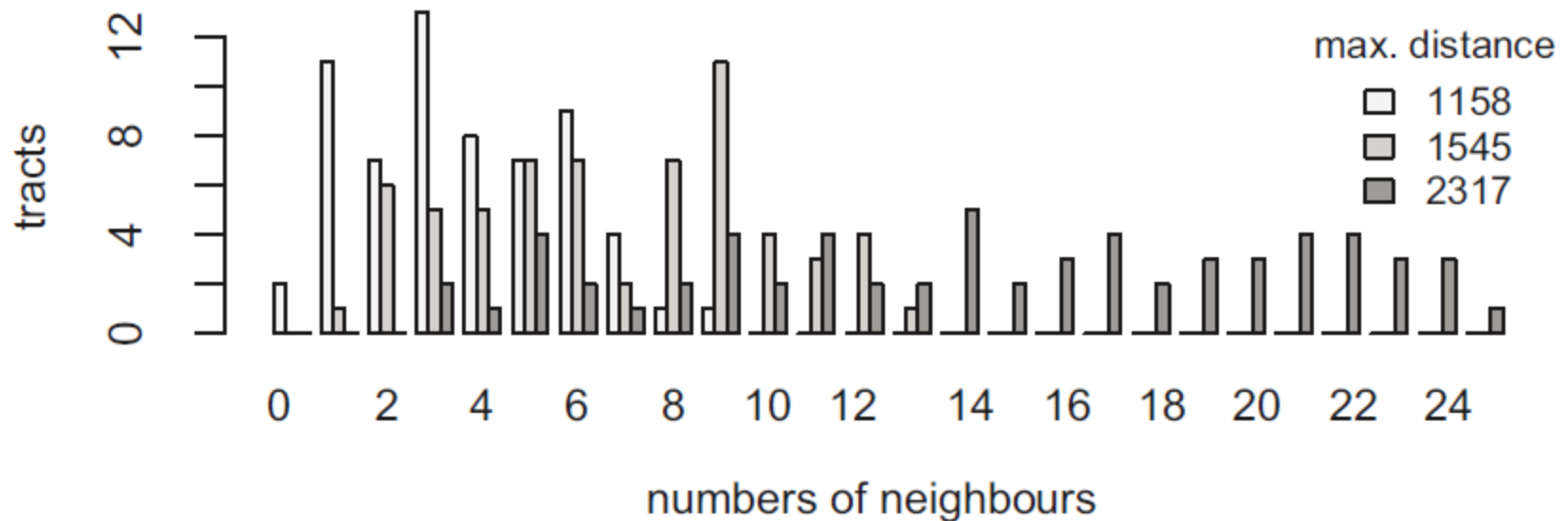


Fig. 9.6. (a) Neighbours within 1,158 m; (b) neighbours within 1,545 m; (c) neighbours within 2,317 m

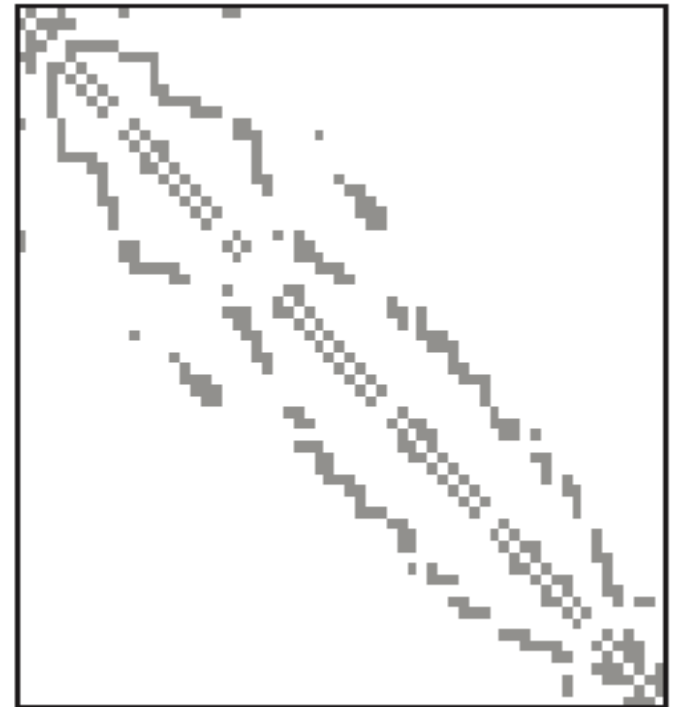
Neighbor/Connectivity Histogram



Source: Bivand and Pebesma and Gomez-Rubio

Spatial Weight Matrix

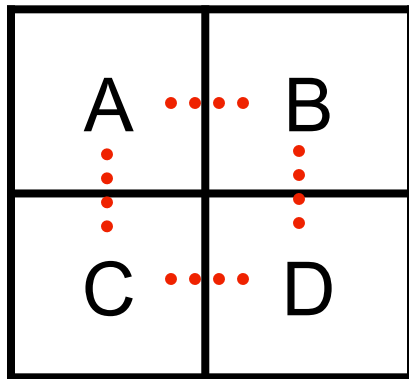
- Spatial weights can be seen as a list of weights indexed by a list of neighbors
- If zone j is not a neighbor of zone i , weights W_{ij} will set to zero
 - The weight matrix can be illustrated as an image
 - Sparse matrix



A Simple Example for Rook case

- Matrix contains a:
 - 1 if share a border
 - 0 if do not share a border

4 areal units

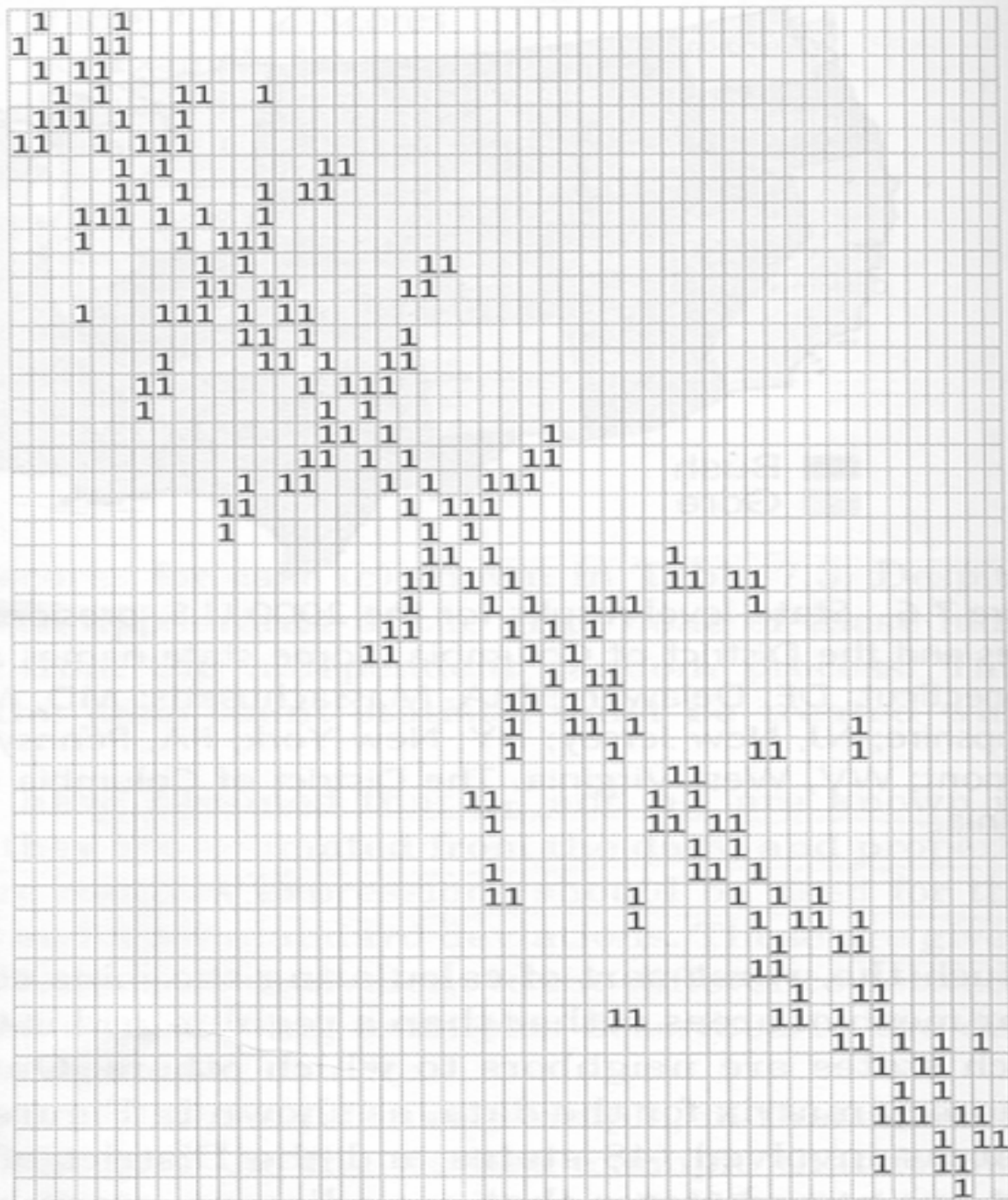


Common border

4x4 matrix

	A	B	C	D
A	0	1	1	0
B	1	0	0	1
C	1	0	0	1
D	0	1	1	0

1 Washington
 2 Oregon
 3 California
 4 Arizona
 5 Nevada
 6 Idaho
 7 Montana
 8 Wyoming
 9 Utah
 10 New Mexico
 11 Texas
 12 Oklahoma
 13 Colorado
 14 Kansas
 15 Nebraska
 16 South Dakota
 17 North Dakota
 18 Minnesota
 19 Iowa
 20 Missouri
 21 Arkansas
 22 Louisiana
 23 Mississippi
 24 Tennessee
 25 Kentucky
 26 Illinois
 27 Wisconsin
 28 Michigan
 29 Indiana
 30 Ohio
 31 West Virginia
 32 Florida
 33 Alabama
 34 Georgia
 35 South Carolina
 36 North Carolina
 37 Virginia
 38 Maryland
 39 Delaware
 40 District of Columbia
 41 New Jersey
 42 Pennsylvania
 43 New York
 44 Connecticut
 45 Rhode Island
 46 Massachusetts
 47 New Hampshire
 48 Vermont
 49 Maine



Sparse Contiguity Matrix for US States -- obtained from Anselin's web site (see powerpoint for link)

Name	Fips	Ncount	N1	N2	N3	N4	N5	N6	N7	N8
Alabama	1	4	28	13	12	47				
Arizona	4	5	35	8	49	6	32			
Arkansas	5	6	22	28	48	47	40	29		
California	6	3	4	32	41					
Colorado	8	7	35	4	20	40	31	49	56	
Connecticut	9	3	44	36	25					
Delaware	10	3	24	42	34					
District of Columbia	11	2	51	24						
Florida	12	2	13	1						
Georgia	13	5	12	45	37	1	47			
Idaho	16	6	32	41	56	49	30	53		
Illinois	17	5	29	21	18	55	19			
Indiana	18	4	26	21	17	39				
Iowa	19	6	29	31	17	55	27	46		
Kansas	20	4	40	29	31	8				
Kentucky	21	7	47	29	18	39	54	51	17	
Louisiana	22	3	28	48	5					
Maine	23	1	33							
Maryland	24	5	51	10	54	42	11			
Massachusetts	25	5	44	9	36	50	33			
Michigan	26	3	18	39	55					
Minnesota	27	4	19	55	46	38				
Mississippi	28	4	22	5	1	47				
Missouri	29	8	5	40	17	21	47	20	19	31
Montana	30	4	16	56	38	46				
Nebraska	31	6	29	20	8	19	56	46		
Nevada	32	5	6	4	49	16	41			
New Hampshire	33	3	25	23	50					
New Jersey	34	3	10	36	42					
New Mexico	35	5	48	40	8	4	49			
New York	36	5	34	9	42	50	25			
North Carolina	37	4	45	13	47	51				
North Dakota	38	3	46	27	30					
Ohio	39	5	26	21	54	42	18			
Oklahoma	40	6	5	35	48	29	20	8		
Oregon	41	4	6	32	16	53				
Pennsylvania	42	6	24	54	10	39	36	34		
Rhode Island	44	2	25	9						
South Carolina	45	2	13	37						
South Dakota	46	6	56	27	19	31	38	30		
Tennessee	47	8	5	28	1	37	13	51	21	29
Texas	48	4	22	5	35	40				
Utah	49	6	4	8	35	56	32	16		
Vermont	50	3	36	25	33					
Virginia	51	6	47	37	24	54	11	21		
Washington	53	2	41	16						
West Virginia	54	5	51	21	24	39	42			
Wisconsin	55	4	26	17	19	27				
Wyoming	56	6	49	16	31	8	46	30		

Style of Spatial Weight Matrix

- Row
 - a weight of unity for each neighbor relationship
- Row standardization
 - Symmetry not guaranteed
 - can be interpreted as allowing the calculation of average values across neighbors
- General spatial weights based on distances

Row vs. Row standardization

A	B	C
D	E	F

Divide each
number by the
row sum

Total number of neighbors
--some have more than others

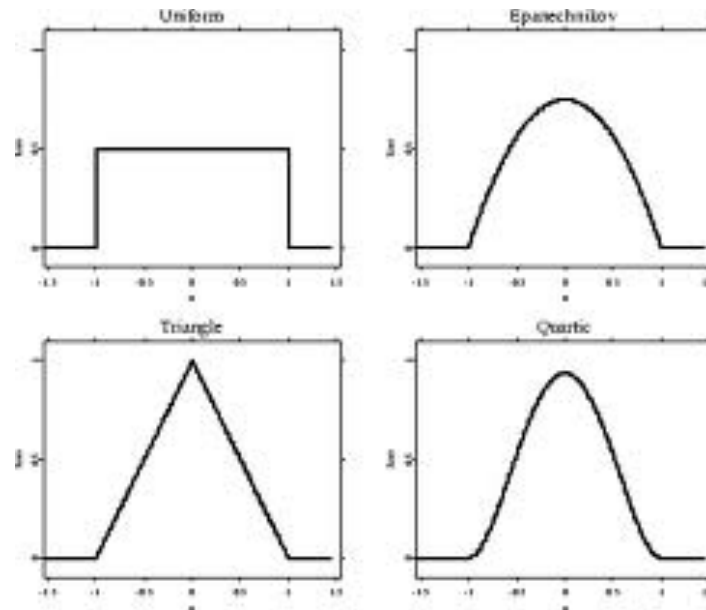
	A	B	C	D	E	F	Row Sum
A	0	1	0	1	0	0	2
B	1	0	1	0	1	0	3
C	0	1	0	0	0	1	2
D	1	0	0	0	1	0	2
E	0	1	0	1	0	1	3
F	0	0	1	0	1	0	2

Row standardized
--usually use this

	A	B	C	D	E	F	Row Sum
A	0.0	0.5	0.0	0.5	0.0	0.0	1
B	0.3	0.0	0.3	0.0	0.3	0.0	1
C	0.0	0.5	0.0	0.0	0.0	0.5	1
D	0.5	0.0	0.0	0.0	0.5	0.0	1
E	0.0	0.3	0.0	0.3	0.0	0.3	1
F	0.0	0.0	0.5	0.0	0.5	0.0	1

General Spatial Weights Based on Distance

- Decay functions of distance
 - Most common choice is the inverse (reciprocal) of the distance between locations i and j ($w_{ij} = 1/d_{ij}$)
 - Other functions also used
 - inverse of squared distance ($w_{ij} = 1/d_{ij}^2$), or
 - negative exponential ($w_{ij} = e^{-d}$ or $w_{ij} = e^{-d^2}$)



Measure of Spatial Autocorrelation

Global Measures and Local Measures

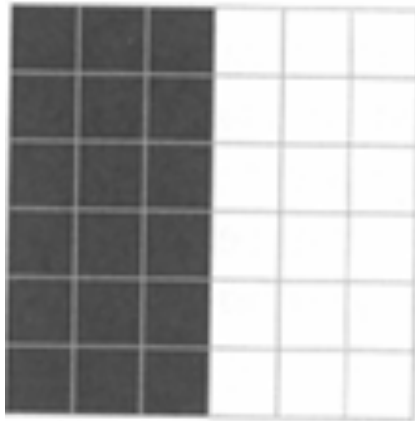
- Global Measures
 - A single value which applies to the entire data set
 - The same pattern or process occurs over the entire geographic area
 - An average for the entire area
- Local Measures
 - A value calculated for each observation unit
 - Different patterns or processes may occur in different parts of the region
 - A unique number for each location
- Global measures usually can be decomposed into a combination of local measures

Global Measures and Local Measures

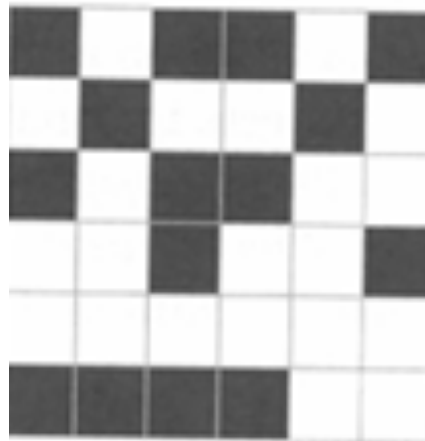
- Global Measures
 - Join Count
 - Moran's I, Getis-Ord's G
- Local Measures
 - Local Moran's I , Getis-Ord's G

Join (or Joint or Joins) Count Statistic

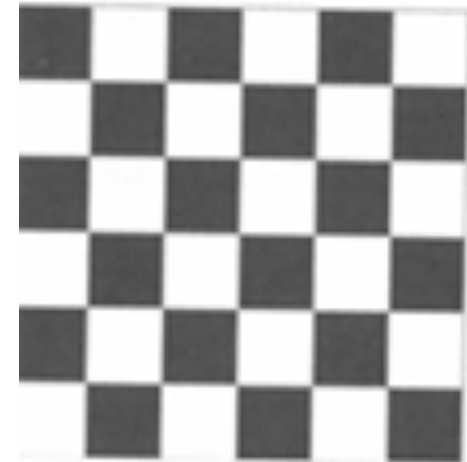
Positive autocorrelation



No autocorrelation



Negative autocorrelation



Rook's case

$$J_{BB} = 27$$

$$J_{WW} = 27$$

$$J_{BW} = 6$$

Queen's case

$$J_{BB} = 47$$

$$J_{WW} = 47$$

$$J_{BW} = 16$$

$$J_{BB} = 6$$

$$J_{WW} = 19$$

$$J_{BW} = 35$$

$$J_{BB} = 14$$

$$J_{WW} = 40$$

$$J_{BW} = 56$$

$$J_{BB} = 0$$

$$J_{WW} = 0$$

$$J_{BW} = 60$$

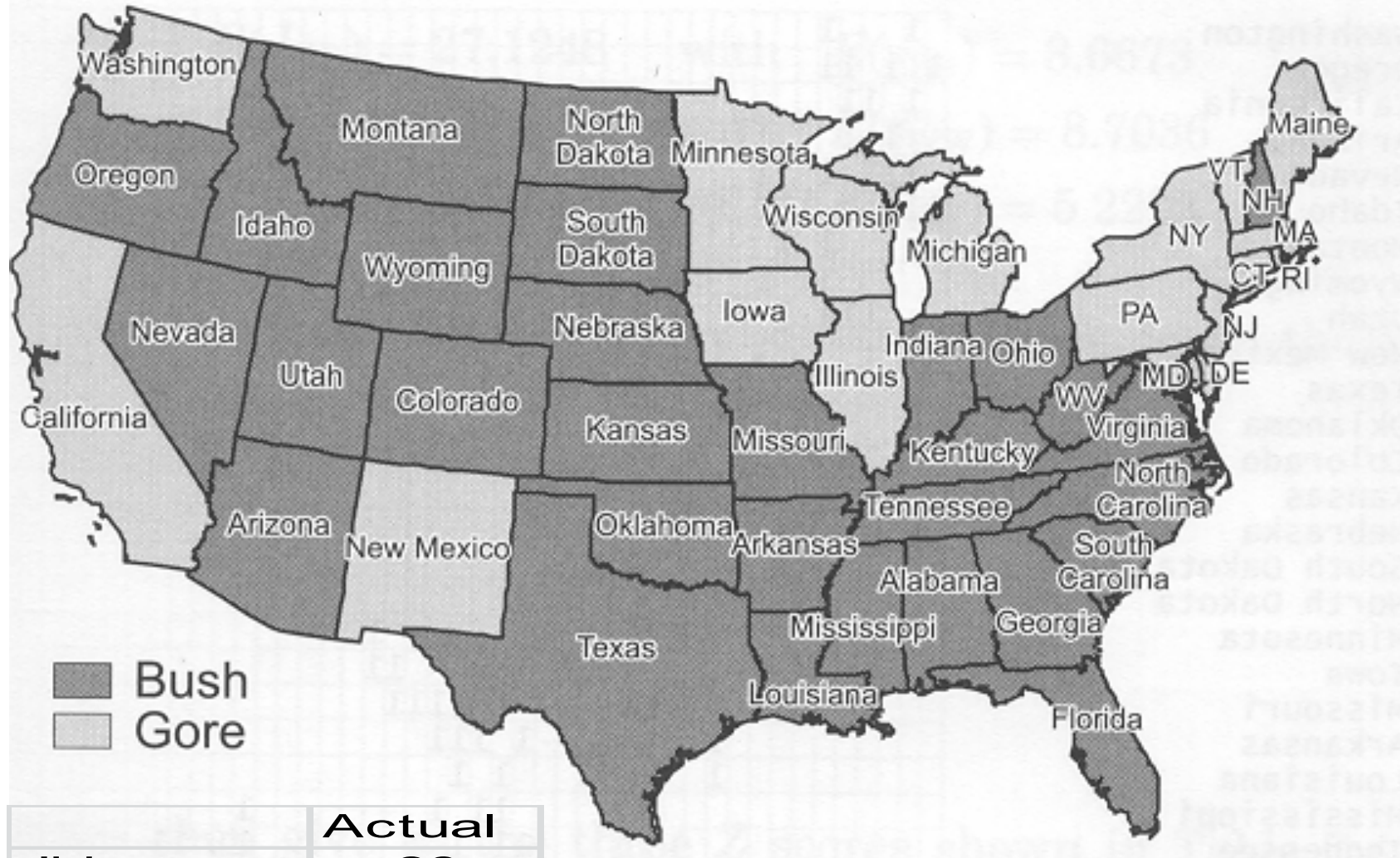
$$J_{BB} = 25$$

$$J_{WW} = 25$$

$$J_{BW} = 60$$

- 60 for Rook Case
- 110 for Queen Case

Gore/Bush Presidential Election 2000



	Actual
Jbb	60
Jgg	21
Jbg	28
Total	109

Join Count Statistic for Gore/Bush 2000 by State

candidates	probability
Bush	0.49885
Gore	0.50115

	Actual	Expected	Stan Dev	Z-score
Jbb	60	27.125	8.667	3.7930
Jgg	21	27.375	8.704	-0.7325
Jbg	28	54.500	5.220	-5.0763
Total	109	109.000		

- The expected number of joins is calculated based on the proportion of votes each received in the election (for Bush = $109 * .499 * .499 = 27.125$)
- There are far more Bush/Bush joins (actual = 60) than would be expected (27)
 - Positive autocorrelation
- There are far fewer Bush/Gore joins (actual = 28) than would be expected (54)
 - Positive autocorrelation
- No strong clustering evidence for Gore (actual = 21 slightly less than 27.375)

Moran's I

- The most common measure of Spatial Autocorrelation
- Use for points or polygons
 - Join Count statistic only for polygons
- Use for a continuous variable (any value)
 - Join Count statistic only for binary variable (1,0)



Patrick Alfred Pierce Moran (1917-1988)

Formula for Moran's I

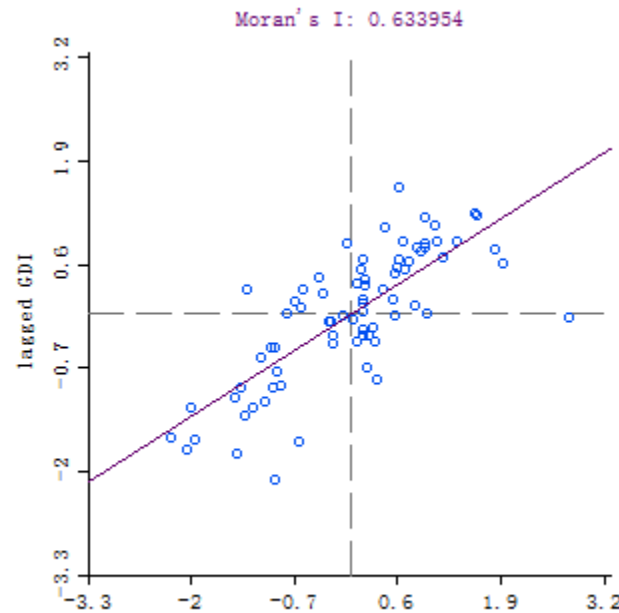
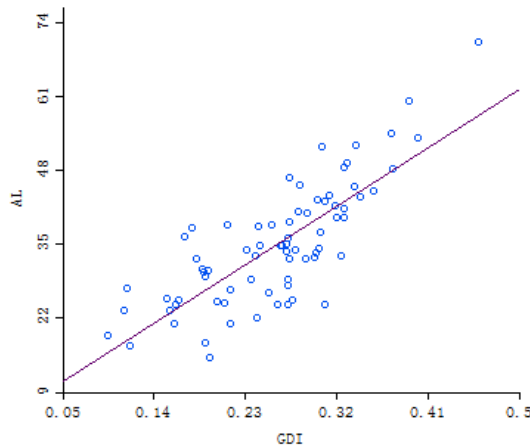
$$I = \frac{N \sum_{i=1}^n \sum_{j=1}^n w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{\left(\sum_{i=1}^n \sum_{j=1}^n w_{ij} \right) \sum_{i=1}^n (x_i - \bar{x})^2}$$

- Where:

- N is the number of observations (points or polygons)
- \bar{x} is the mean of the variable
- x_i is the variable value at a particular location
- x_j is the variable value at another location
- w_{ij} is a weight indexing location of i relative to j

Moran's I and Correlation Coefficient

- **Correlation Coefficient [-1, 1]**
 - Relationship between two different variables
- **Moran's I [-1, 1]**
 - Spatial autocorrelation and often involves one (spatially indexed) variable only
 - Correlation between observations of a spatial variable at location X and “spatial lag” of X formed by averaging all the observation at neighbors of X



Correlation Coefficient

Note the similarity of the numerator (top) to the measures of spatial association discussed earlier if we view Y_i as being the X_i for the neighboring polygon

(see next slide)

$$\frac{\sum_{i=1}^n (y_i - \bar{y})(x_i - \bar{x})/n}{\sqrt{\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n}} \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n}}}$$

$$\frac{N \sum_{i=1}^n \sum_{j=1}^n w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{(\sum_{i=1}^n \sum_{j=1}^n w_{ij}) \sum_{i=1}^n (x_i - \bar{x})^2}$$

Spatial
auto-correlation

=

$$\frac{\sum_{i=1}^n \sum_{j=1}^n w_{ij} (x_i - \bar{x})(x_j - \bar{x}) / \sum_{i=1}^n \sum_{j=1}^n w_{ij}}{\sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n}} \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n}}}$$

Correlation Coefficient

$$\frac{\sum_{i=1}^n 1(y_i - \bar{y})(x_i - \bar{x})/n}{\sqrt{\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{n}} \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n}}}$$

Spatial weights

Y_i is the X_i for the neighboring polygon

$$\frac{N \sum_{i=1}^n \sum_{j=1}^n w_{ij} (x_i - \bar{x})(x_j - \bar{x})}{(\sum_{i=1}^n \sum_{j=1}^n w_{ij}) \sum_{i=1}^n (x_i - \bar{x})^2} =$$

$$\frac{\sum_{i=1}^n \sum_{j=1}^n w_{ij} (x_i - \bar{x})(x_j - \bar{x}) / \sum_{i=1}^n \sum_{j=1}^n w_{ij}}{\sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n}} \sqrt{\frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n}}}$$

Moran's I

Statistical Significance Tests for Moran's I

- Based on the normal frequency distribution with

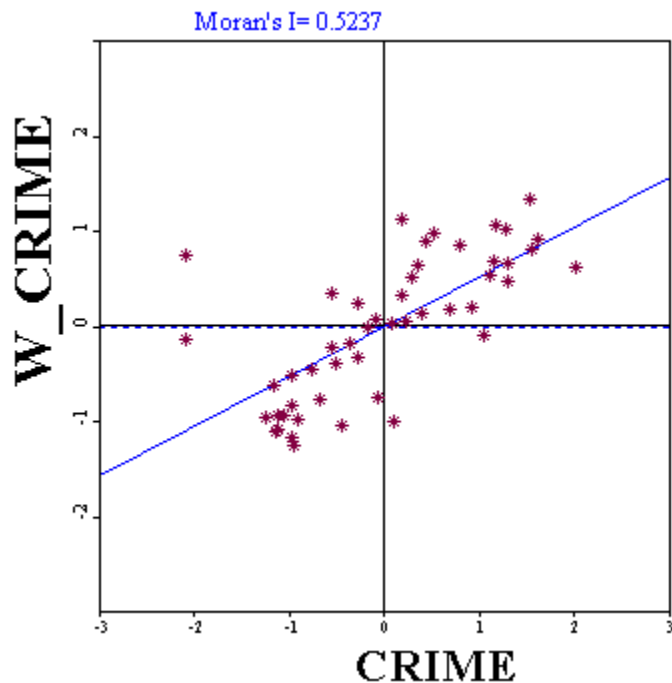
$$Z = \frac{I - E(I)}{S_{error}(I)}$$

Where: I is the calculated value for Moran's I
from the sample
E(I) is the expected value if random
S is the standard error

- Statistical significance test
 - Monte Carlo test, as we did for spatial pattern analysis
 - Permutation test
 - Non-parametric
 - Data-driven, no assumption of the data
 - Implemented in GeoDa

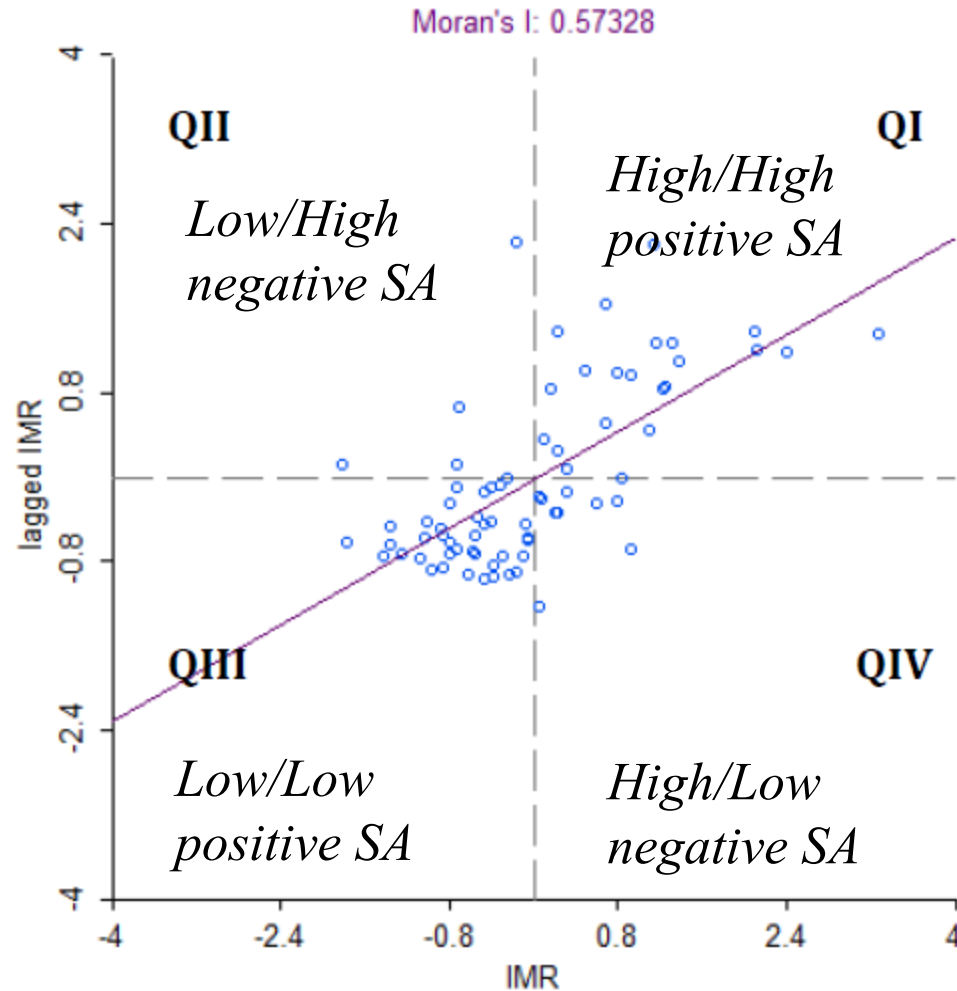
Moran Scatter Plots

We can draw a scatter diagram between these two variables (in standardized form): X and $\text{lag-}X$ (or W_X)

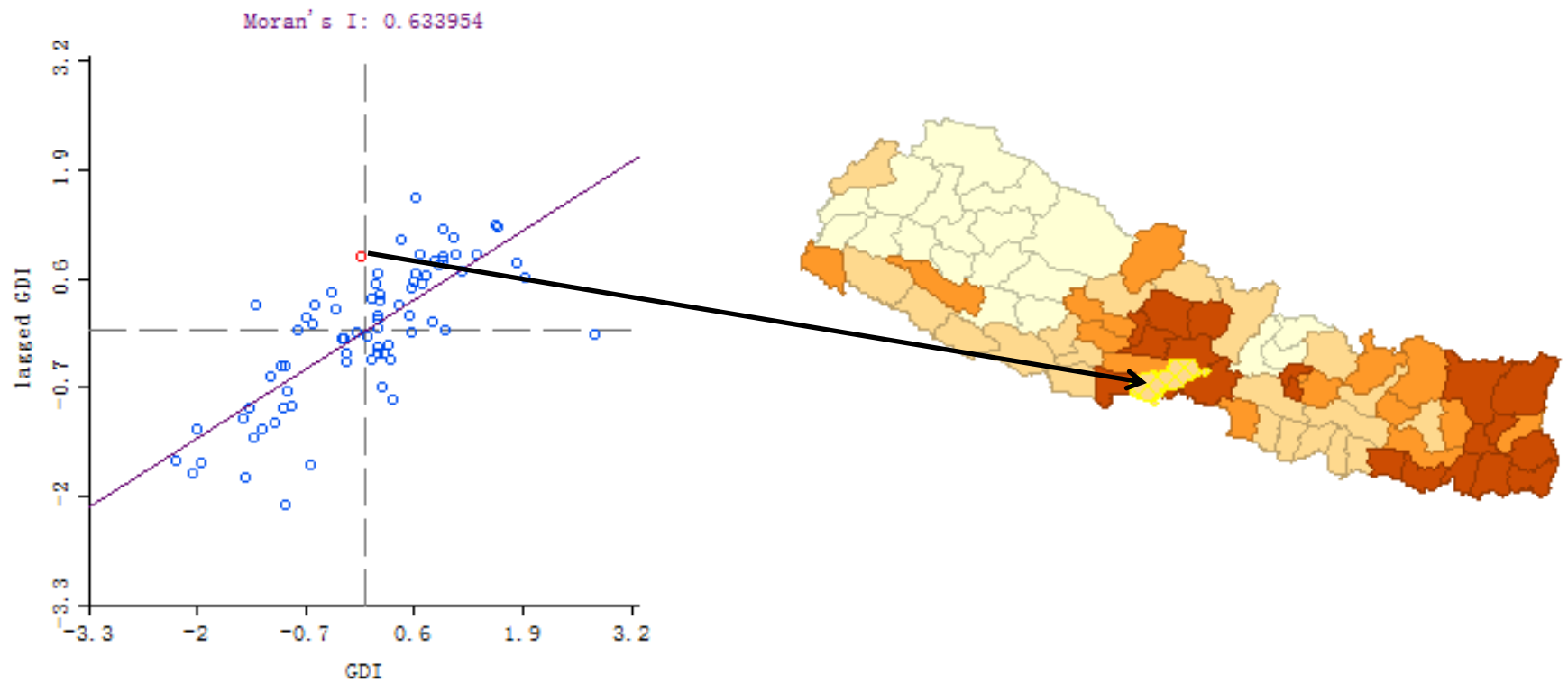


The slope of this *regression line* is
Moran's I

Moran Scatter Plots



Moran Scatterplot: Example

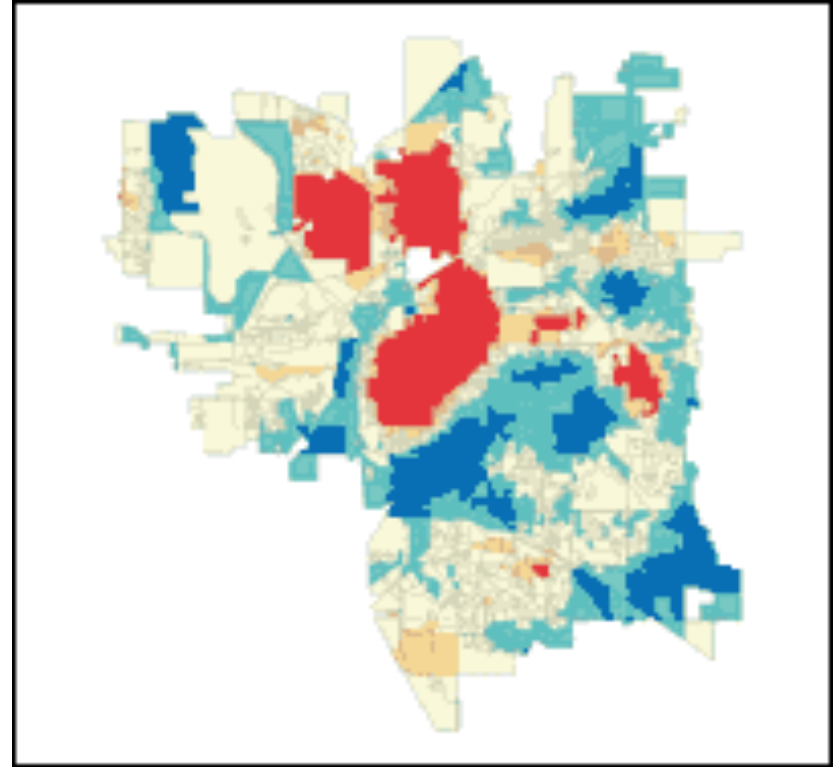


Moran's I for rate-based data

- Moran's I is often calculated for rates, such as crime rates (e.g. number of crimes per 1,000 population) or infant mortality rates (e.g. number of deaths per 1,000 births)
- An adjustment should be made, especially if the denominator in the rate (population or number of births) varies greatly (as it usually does)
- Adjustment is known as the *EB adjustment*:
 - see Assuncao-Reis *Empirical Bayes Standardization Statistics in Medicine*, 1999
- *GeoDA* software includes an option for this adjustment

Hot Spots and Cold Spots

- What is a *hot spot*?
 - A place where high values cluster together
- What is a *cold spot*?
 - A place where low values cluster together
- Moran's I and Geary's C cannot distinguish them
 - They only indicate clustering
 - Cannot tell if these are hot spots, cold spots, or both



Getis-Ord General/Global G-Statistic

- The G statistic distinguishes between hot spots and cold spots. It identifies *spatial concentrations*.
 - G is relatively large if high values cluster together
 - G is relatively low if low values cluster together
- The General G statistic is interpreted relative to its *expected value*
 - The value for which there is no spatial association
 - $G >$ (larger than) *expected value* → potential “hot spots”
 - $G <$ (smaller than) *expected value* → potential “cold spots”
- Comments:
 - General G will not show negative spatial autocorrelation
 - Should only be calculated for ratio scale data
 - data with a “natural” zero such as crime rates, birth rates
 - Although it was defined using a contiguity (0,1) weights matrix, any type of spatial weights matrix can be used
 - ArcGIS gives multiple options

Local Measures of Spatial Autocorrelation

Local Indicators of Spatial Association (LISA)

- Local versions of *Moran's I*, and the *Getis-Ord G statistic*
- Moran's I is most commonly used, and the local version is often called Anselin's LISA, or just LISA

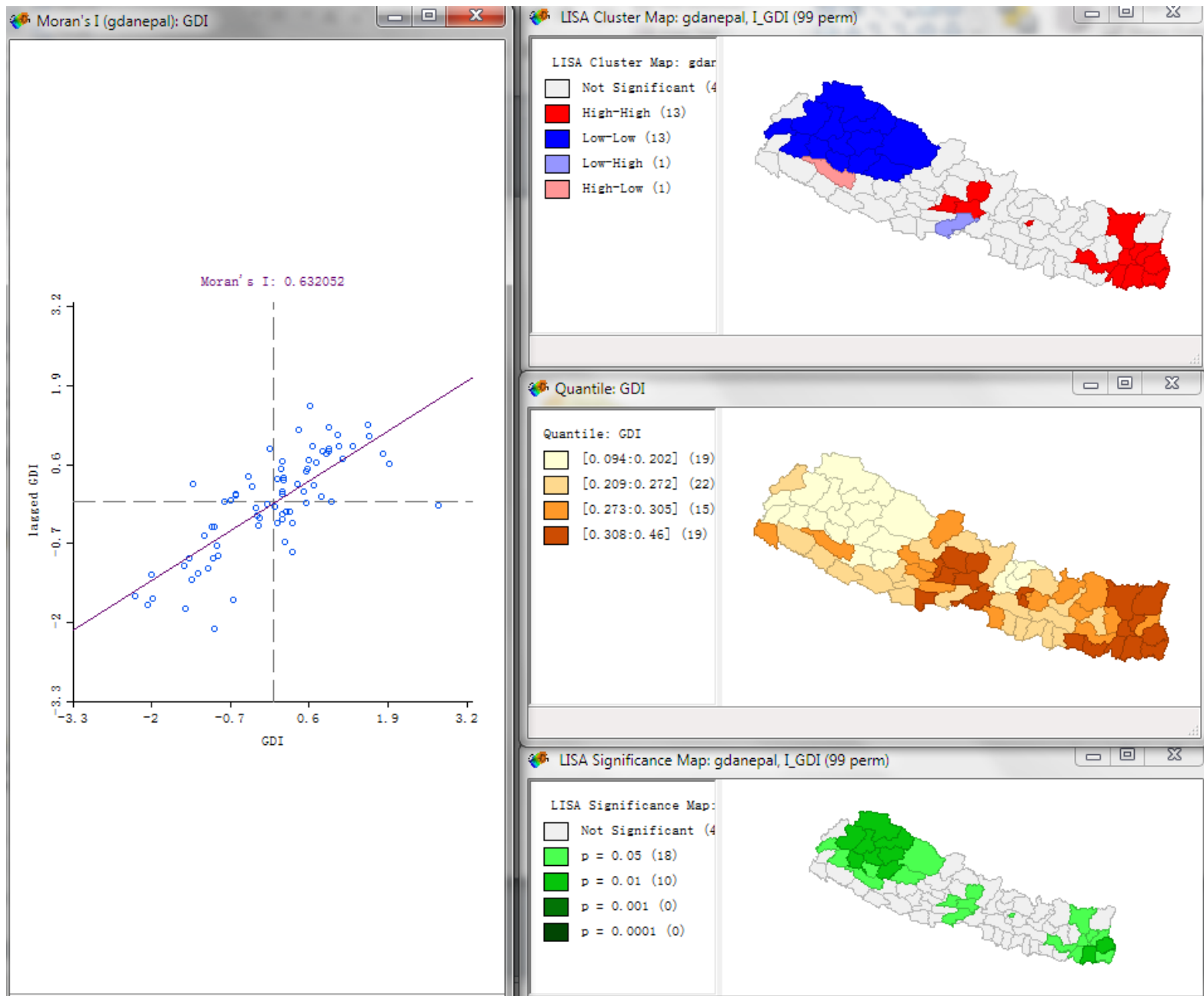
See:

Luc Anselin 1995 *Local Indicators of Spatial Association-LISA* Geographical Analysis 27: 93-115

Local Indicators of Spatial Association (LISA)

- The statistic is calculated for each areal unit in the data
- For each polygon, the index is calculated based on neighboring polygons with which it shares a border
- A measure is available for each polygon, these can be mapped to indicate how spatial autocorrelation varies over the study region
- Each index has an associated test statistic, we can also map which of the polygons has a statistically significant relationship with its neighbors, and show type of relationship

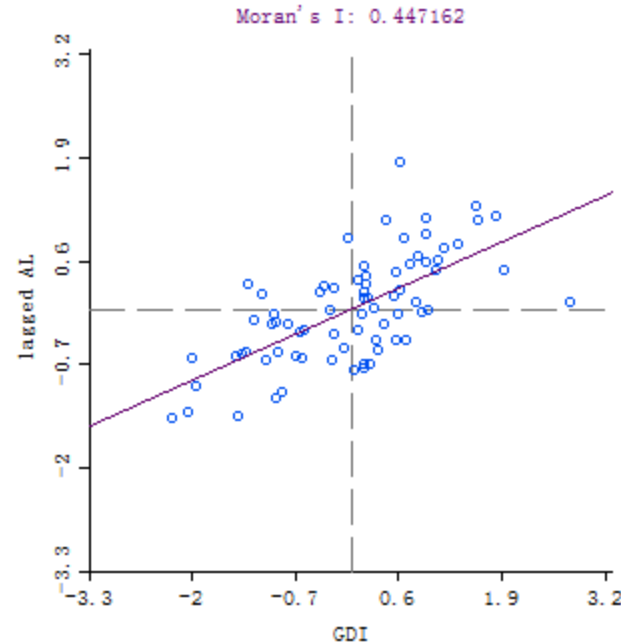
Example:



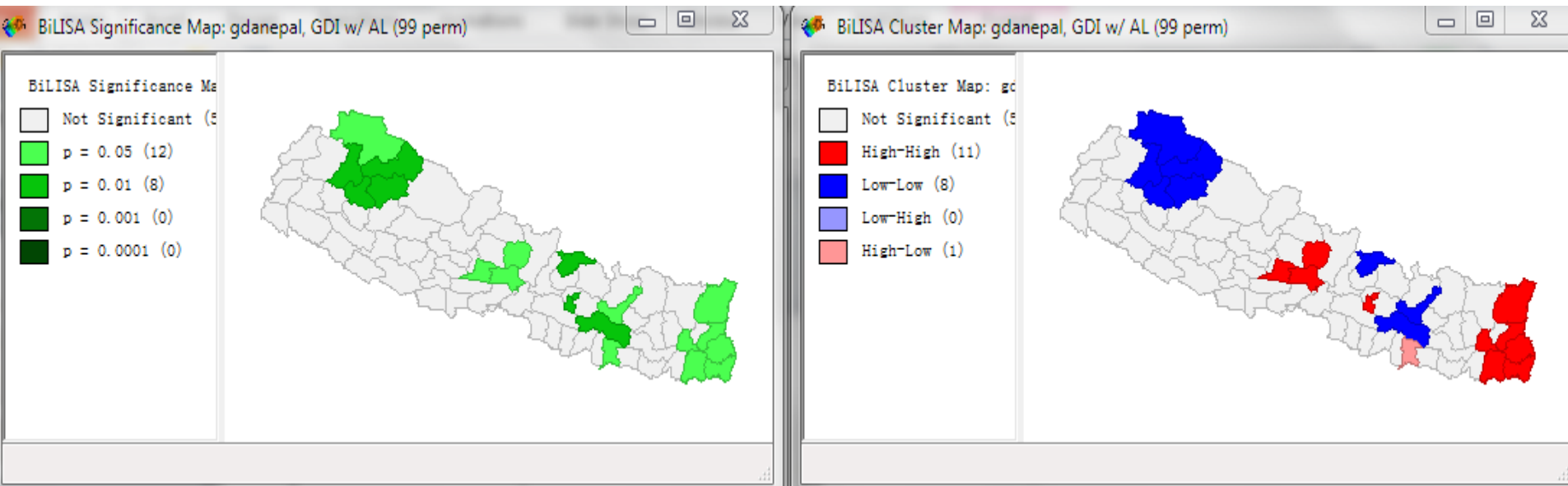
Bivariate LISA

- Moran's I is the correlation between X and Lag-X--the same variable but in nearby areas
 - Univariate Moran's I
- Bivariate Moran's I is a correlation between X and a different variable in nearby areas.

Moran Scatter Plot for GDI vs AL



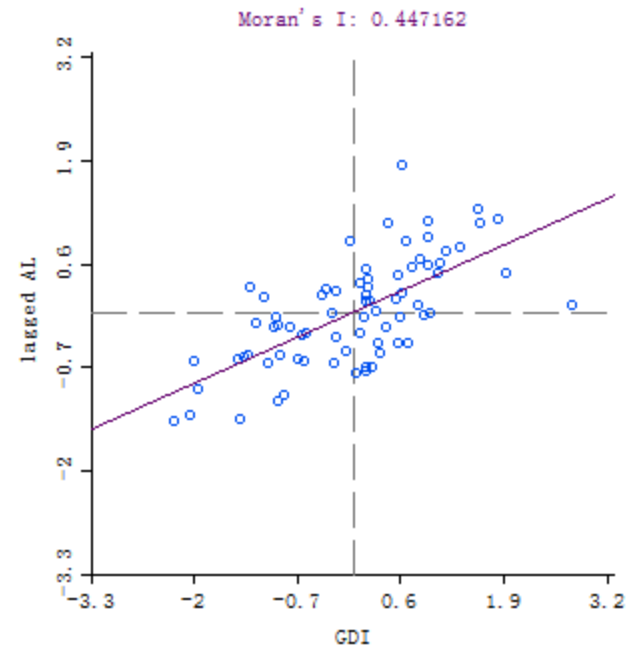
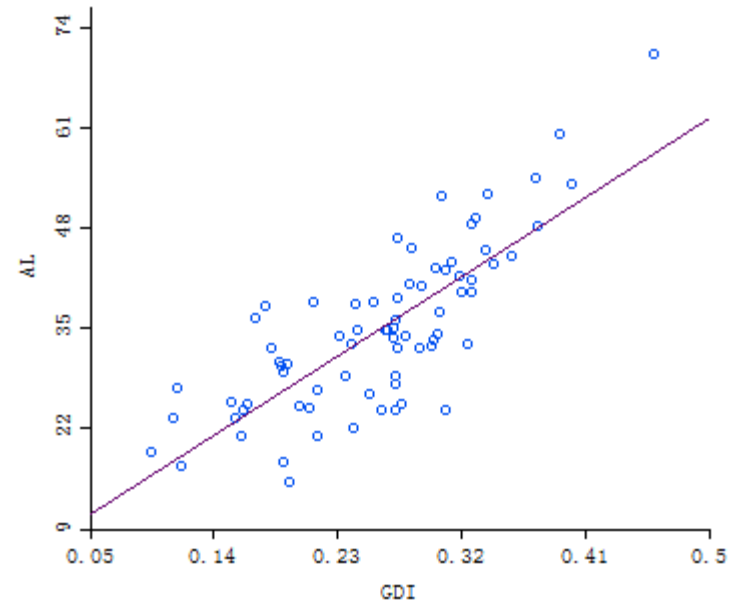
Moran Significance Map for GDI vs. AL



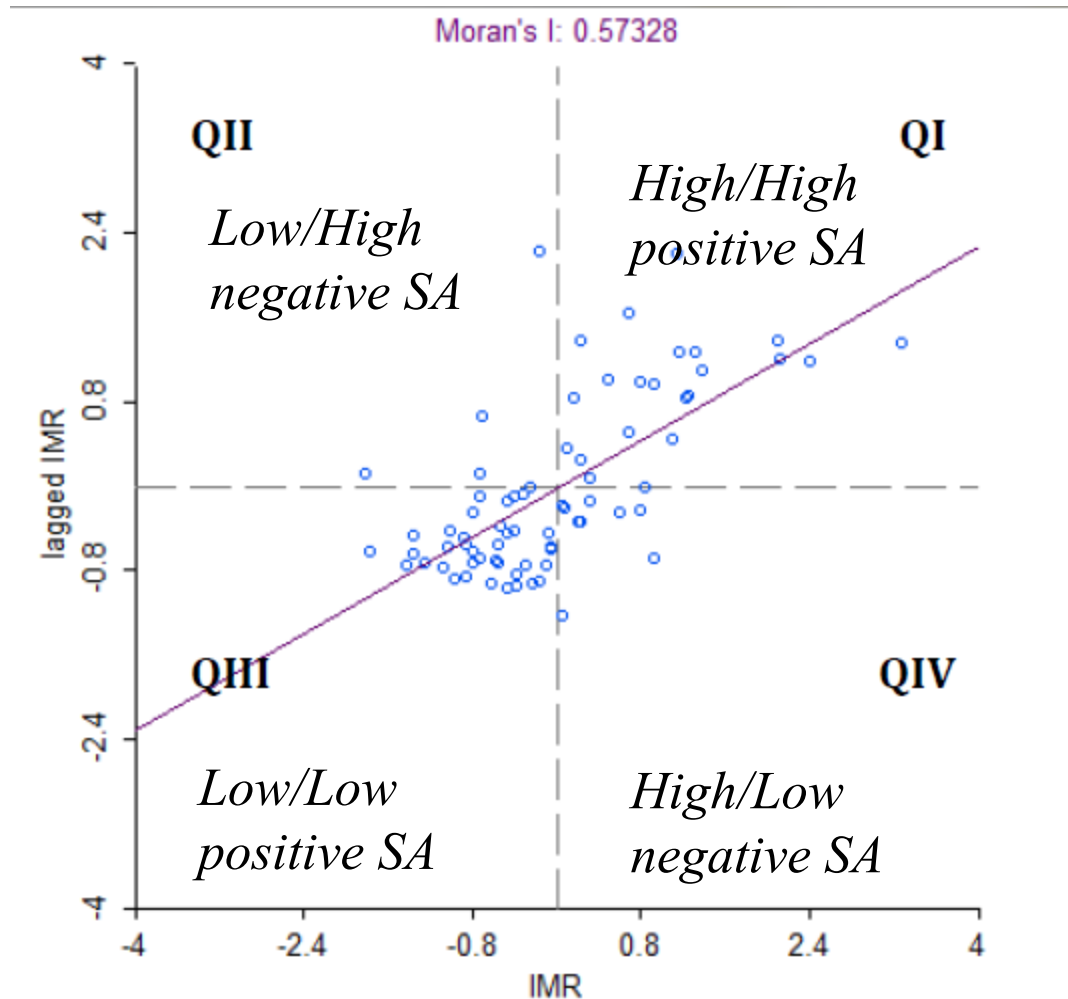
Bivariate LISA

and the Correlation Coefficient

- Correlation Coefficient is the relationship between two different variables in the same area
- Bivariate LISA is a correlation between two different variables in an area and in nearby areas.



Bivariate Moran Scatter Plot



Summary

- Spatial autocorrelation of areal data
- Spatial weight matrix
- Measures of spatial autocorrelation
- Global Measure
 - Moran's I/General G and G^*
- Local
 - LISA: Moran's I/General G and G^*
 - Bivariate LISA
 - Significance test