

# Inference and Simulation

EC 607, Set 04

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# Prologue

# Schedule

## Last time

The *CEF* and least-squares regression

## Today

Inference

*Read MHE 3.1*

## Upcoming

Lab: TBD

Problem set 002 coming soon.

Class project, step 1 due on April 27

# Inference

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Why?

Q What's the big deal with inference?

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**A** We rarely know the CEF or the population (and its regression vector).

We *can* draw statistical inferences about the population using samples.

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## Why?

**Q** What's the big deal with inference?

**A** We rarely know the CEF or the population (and its regression vector).

We *can* draw statistical inferences about the population using samples.

**Important** The issue/topic of *statistical inference* is separate from *causality*.

Separate questions

1. How do we interpret the estimated coefficient  $\hat{\beta}$ ?
2. What is the sampling distribution of  $\hat{\beta}$ ?

# Inference

## Moving from population to sample

**Recall** The population-regression function gives us the best linear approximation to the CEF.



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$$\beta = E [\mathbf{X}_i \mathbf{X}_i']^{-1} E[\mathbf{X}_i \mathbf{Y}_i]$$

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We're interested in the (unknown) population-regression vector

$$\beta = E [\mathbf{X}_i \mathbf{X}_i']^{-1} E[\mathbf{X}_i \mathbf{Y}_i]$$

which we estimate via the ordinary least squares (OLS) estimator<sup>†</sup>

$$\hat{\beta} = \left( \sum_i \mathbf{X}_i \mathbf{X}_i' \right)^{-1} \left( \sum_i \mathbf{X}_i \mathbf{Y}_i \right)$$

<sup>†</sup> MHE presents a method-of-moments motivation for this derivation, where  $\frac{1}{n} \sum_i \mathbf{X}_i \mathbf{X}_i'$  is our sample-based estimated for  $E[\mathbf{X}_i \mathbf{X}_i']$ . You've also seen others, e.g., minimizing MSE of  $\mathbf{Y}_i$  given  $\mathbf{X}_i$ .

# Inference

## A classic

However you write it, this OLS estimator

$$\begin{aligned}\hat{\beta} &= (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{y} \\ &= \left(\sum_i \mathbf{x}_i \mathbf{x}_i'\right)^{-1} \left(\sum_i \mathbf{x}_i Y_i\right) \\ &= \beta + \left[\sum_i \mathbf{x}_i \mathbf{x}_i'\right]^{-1} \sum_i \mathbf{x}_i e_i\end{aligned}$$

is the same estimator you've been using since undergrad.

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is the same estimator you've been using since undergrad.

*Note* I'm following *MHE* in defining  $e_i = Y_i - \mathbf{x}_i'\beta$ .

# Inference

## A classic

As you've learned, the OLS estimator

$$\hat{\beta} = \left( \sum_i \mathbf{X}_i \mathbf{X}_i' \right)^{-1} \left( \sum_i \mathbf{X}_i Y_i \right) = \beta + \left[ \sum_i \mathbf{X}_i \mathbf{X}_i' \right]^{-1} \sum_i \mathbf{X}_i e_i$$

has asymptotic covariance

$$E \left[ \mathbf{X}_i \mathbf{X}_i' \right]^{-1} E \left[ \mathbf{X}_i \mathbf{X}_i' e_i^2 \right] E \left[ \mathbf{X}_i \mathbf{X}_i' \right]^{-1}$$

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which we estimate by **(1)** replacing  $e_i$  with  $\hat{e}_i = Y_i - \mathbf{X}_i' \hat{\beta}$  and **(2)** replacing expectations with sample means, e.g.,  $E[\mathbf{X}_i \mathbf{X}_i' e_i^2]$  becomes  $\frac{1}{n} \sum [\mathbf{X}_i \mathbf{X}_i' \hat{e}_i^2]$ .

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Standard errors of this flavor are known as heteroskedasticity-consistent (or -robust) standard errors (or Eicker-White).

# Inference

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 $E[e_i^2 \mid \mathbf{X}_i] = \sigma^2$  for all  $i$ .



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$$E[\mathbf{X}_i \mathbf{X}_i' e_i^2] = E[E[\mathbf{X}_i \mathbf{X}_i' e_i^2 | \mathbf{X}_i]] = E[\mathbf{X}_i \mathbf{X}_i' E[e_i^2 | \mathbf{X}_i]] = \sigma^2 E[\mathbf{X}_i \mathbf{X}_i']$$

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Now, returning to to the asym. covariance matrix of  $\hat{\beta}$ ,

$$\begin{aligned} E[\mathbf{X}_i \mathbf{X}_i']^{-1} E[\mathbf{X}_i \mathbf{X}_i' e_i^2] E[\mathbf{X}_i \mathbf{X}_i']^{-1} &= E[\mathbf{X}_i \mathbf{X}_i']^{-1} \sigma^2 E[\mathbf{X}_i \mathbf{X}_i'] E[\mathbf{X}_i \mathbf{X}_i']^{-1} \\ &= \sigma^2 E[\mathbf{X}_i \mathbf{X}_i']^{-1} \end{aligned}$$

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Thus, even if  $Y_i \mid \mathbf{X}_i$  has constant variance,  $e_i \mid \mathbf{X}_i$  is heteroskedastic. Unless you want to assume the CEF is *linear*.



# Inference

## Two notes

1. Heteroskedasticity is **not our biggest concern** in inference.

...as an empirical matter, heteroskedasticity may matter very little... If heteroskedasticity matters a lot, say, more than a 30 percent increase or any marked decrease in standard errors, you should worry about possible programming errors or other problems. (*MHE*, p.47)

2. Notice that we've **avoided "standard" stronger assumptions**, e.g., normality, fixed regressors, linear CEF, homoskedasticity.

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1. Heteroskedasticity is **not our biggest concern** in inference.

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2. Notice that we've **avoided "standard" stronger assumptions**, e.g., normality, fixed regressors, linear CEF, homoskedasticity.

Following (2): We only have large-sample, asymptotic results (consistency) rather than finite-sample results (unbiasedness).

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One practical way we can study the behavior of an estimator: **simulation**.

*Note* You need to make sure your simulation can actually test/respond to the question you are asking (e.g., bias vs. consistency).

# Inference

## Simulation

Let's compare false- and true-positive rates<sup>†</sup> for

1. **Homoskedasticity-assuming standard errors** ( $\text{Var}[e_i | \mathbf{X}_i] = \sigma^2$ )
2. **Heteroskedasticity-robust standard errors**

<sup>†</sup> The false-positive rate goes by many names; another common name: *type-I error rate*.

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### Simulation outline

1. Define data-generating process (DGP).
2. Choose sample size  $n$ .
3. Set seed.
4. Run 10,000 iterations of
  - a. Draw sample of size  $n$  from DGP.
  - b. Conduct inference.
  - c. Record inferences' outcomes.

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Let's keep the disturbances well behaved.

$$Y_i = 1 + e^{0.5X_i} + \varepsilon_i$$

where  $X_i \sim \text{Uniform}(0, 10)$  and  $\varepsilon_i \sim N(0, 1)$ .

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## Data-generating process

$$Y_i = 1 + e^{0.5X_i} + \varepsilon_i$$

where  $X_i \sim \text{Uniform}(0, 10)$  and  $\varepsilon_i \sim N(0, 15^2)$ .

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$$Y_i = 1 + e^{0.5X_i} + \varepsilon_i$$

where  $X_i \sim \text{Uniform}(0, 10)$  and  $\varepsilon_i \sim N(0, 15^2)$ .

```
library(pacman)
p_load(dplyr)
# Choose a size
n = 1000
# Generate data
dgp_df = tibble(
  ε = rnorm(n, sd = 15),
  x = runif(n, min = 0, max = 10),
  y = 1 + exp(0.5 * x) + ε
)
```

# Simulation

## Data-generating process

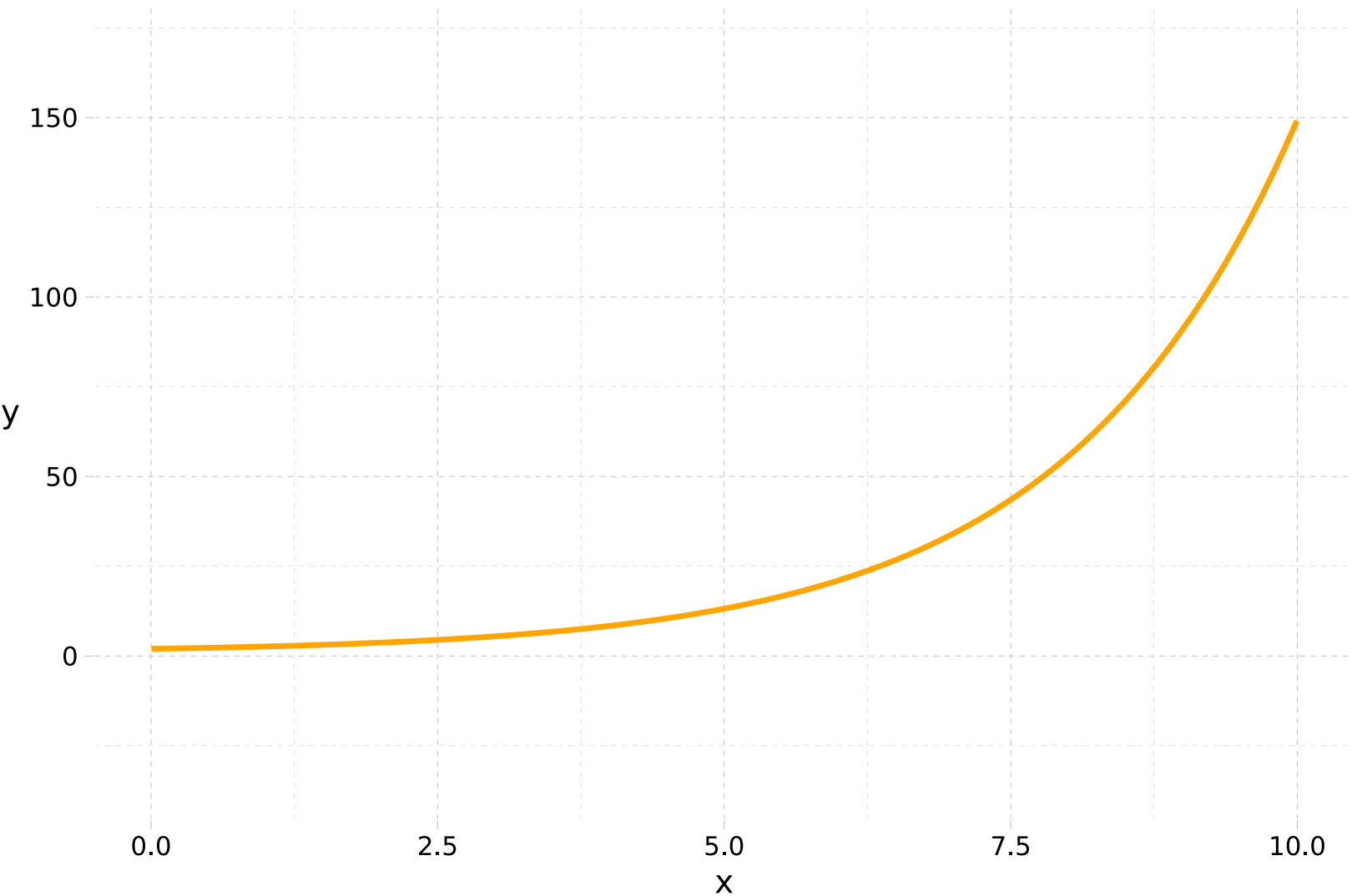
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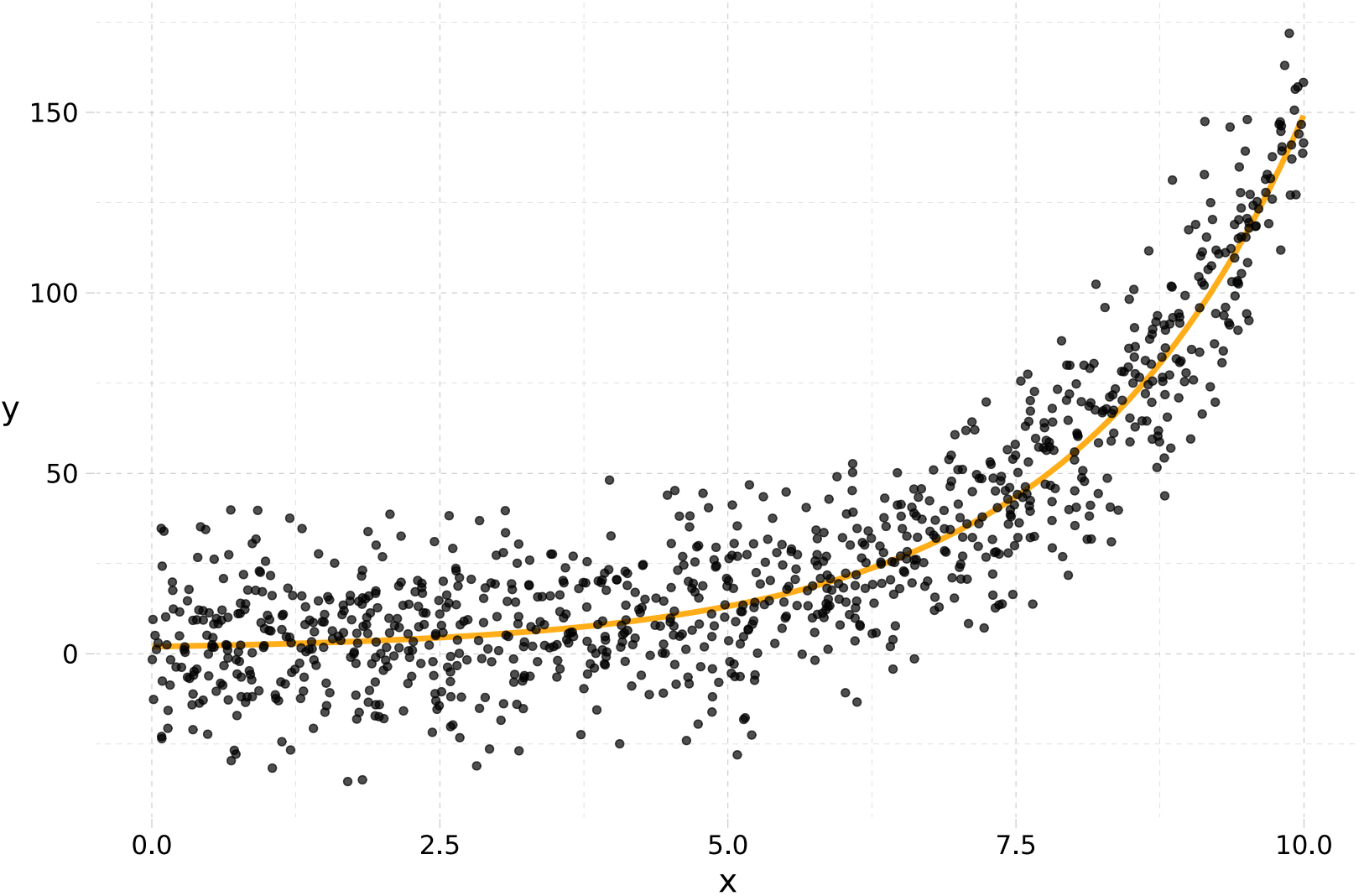
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)
```

```
#> # A tibble: 1,000 × 3
#>       ε      x      y
#>   <dbl> <dbl> <dbl>
#> 1   8.78  9.53 127.
#> 2  10.6   6.22  34.0
#> 3  -1.64  5.32  13.6
#> 4  -6.80  8.92  80.7
#> 5   9.09  1.96  12.8
#> 6 -27.3   8.84  57.0
#> 7   9.45  2.18  13.4
#> 8  -4.14  3.78   3.47
#> 9  -4.26  3.52   2.54
#> 10 -13.8   9.88 127.
#> # ... with 990 more rows
```

Our CEF

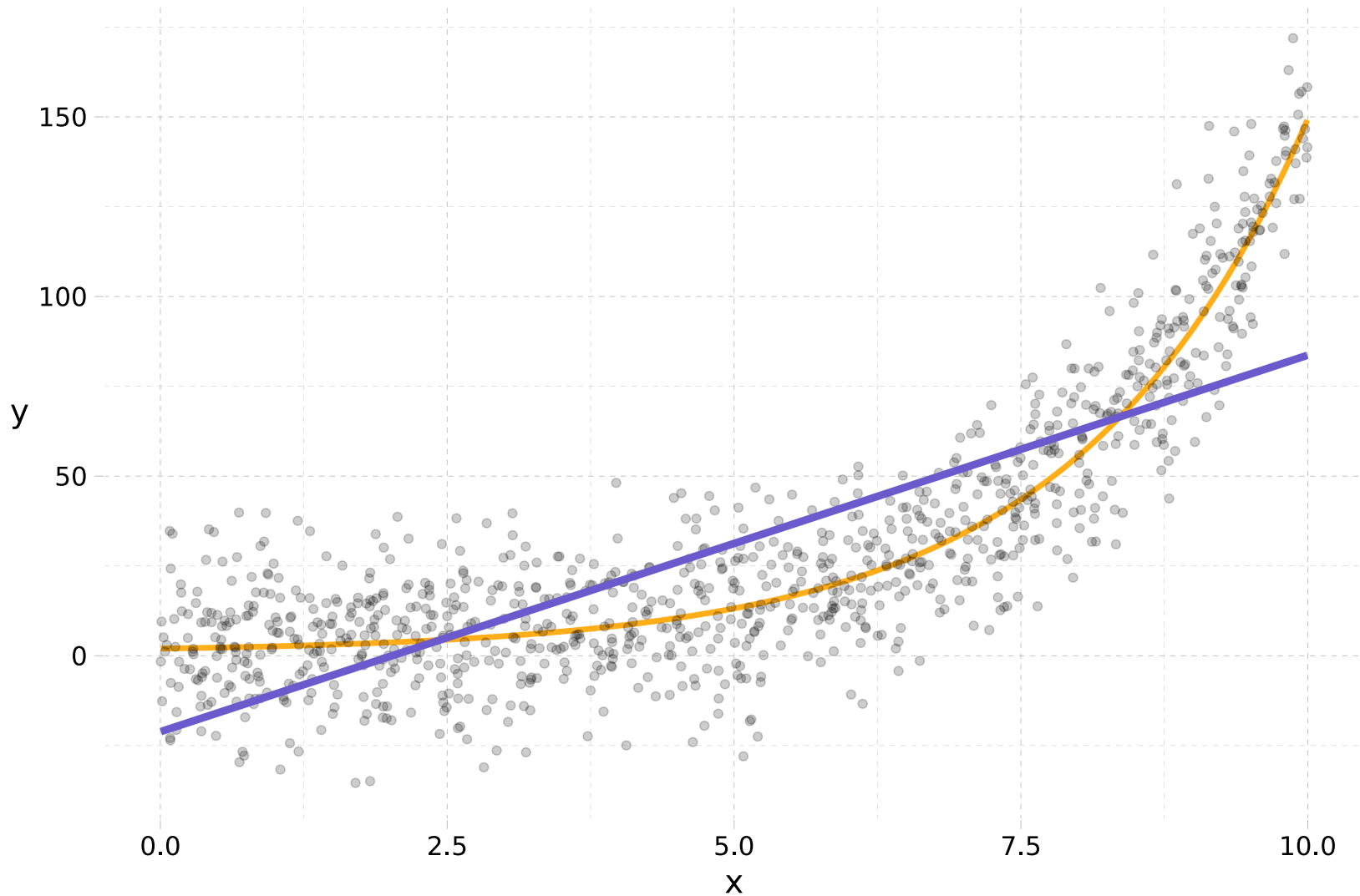


Our population





## The population least-squares regression line



# Simulation

## Iterating

To make iterating easier, let's wrap our DGP in a function.

```
fun_iter = function(iter, n = 30) {  
  # Generate data  
  iter_df = tibble(  
     $\epsilon$  = rnorm(n, sd = 15),  
    x = runif(n, min = 0, max = 10),  
    y = 1 + exp(0.5 * x) +  $\epsilon$   
  )  
}
```

We still need to run a regression and draw some inferences.

*Note* We're defaulting to size-30 samples.

# Simulation

We will use `lm_robust()` from the `estimatr` package for OLS and inference.<sup>†</sup>

- `se_type = "classical"` provides homoskedasticity-assuming SEs
- `se_type = "HC2"` provides heteroskedasticity-robust SEs

```
lm_robust(y ~ x, data = dgp_df, se_type = "classical") %>% tidy() %>% select(1:5)
```

```
#>           term estimate std.error statistic      p.value
#> 1 (Intercept) -21.14183  1.473496 -14.34807  1.383951e-42
#> 2           x   10.48074  0.257810  40.65294  6.560626e-214
```

```
lm_robust(y ~ x, data = dgp_df, se_type = "HC2") %>% tidy() %>% select(1:5)
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```
#>           term estimate std.error statistic      p.value
#> 1 (Intercept) -21.14183  1.4335274 -14.74812  1.112039e-44
#> 2           x   10.48074  0.3097606  33.83495  8.788638e-168
```

<sup>†</sup> `lm()` works for "spherical" standard errors but cannot calculate het.-robust standard errors.

# Simulation

## Inference

Now add these estimators to our iteration function...

```
fun_iter = function(iter, n = 30) {  
  # Generate data  
  iter_df = tibble(  
    ε = rnorm(n, sd = 15),  
    x = runif(n, min = 0, max = 10),  
    y = 1 + exp(0.5 * x) + ε  
  )  
  # Estimate models  
  lm1 = lm_robust(y ~ x, data = iter_df, se_type = "classical")  
  lm2 = lm_robust(y ~ x, data = iter_df, se_type = "HC2")  
  # Stack and return results  
  bind_rows(tidy(lm1), tidy(lm2)) %>%  
    select(1:5) %>% filter(term == "x") %>%  
    mutate(se_type = c("classical", "HC2"), i = iter)  
}
```

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## Run it

Now we need to actually run our `fun_iter()` function 10,000 times.

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There are a lot of ways to run a single function over a list/vector of values.

- `lapply()`, *e.g.*, `lapply(X = 1:3, FUN = sqrt)`
- `for()`, *e.g.*, `for (x in 1:3) sqrt(x)`
- `map()` from `purrr`, *e.g.*, `map(1:3, sqrt)`

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We're going to go with `map()` from the `purrr` package because it easily parallelizes across platforms using the `furrr` package.

# Simulation

## Run it!

Run our function 10,000 times

```
# Packages  
p_load(purrr)  
# Set seed  
set.seed(12345)  
# Run 10,000 iterations  
sim_list = map(1:1e4, fun_iter)
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Parallelized 10,000 iterations

```
# Packages
p_load(purrr, furrr)
# Set options
set.seed(123)
# Tell R to parallelize
plan(multiprocess)
# Run 10,000 iterations
sim_list = future_map(
  1:1e4, fun_iter,
  .options = future_options(seed = T)
)
```

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*Note* Use `multisession` or `multicore` instead of `multiprocess`.

# Simulation

## Run it!!

Our `fun_iter()` function returns a `data.frame`, and `future_map()` returns a `list` (of the returned objects).

So `sim_list` is going to be a `list` of `data.frame` objects. We can bind them into one `data.frame` with `bind_rows()`.

```
# Bind list together  
sim_df = bind_rows(sim_list)
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# Simulation

## Run it!!

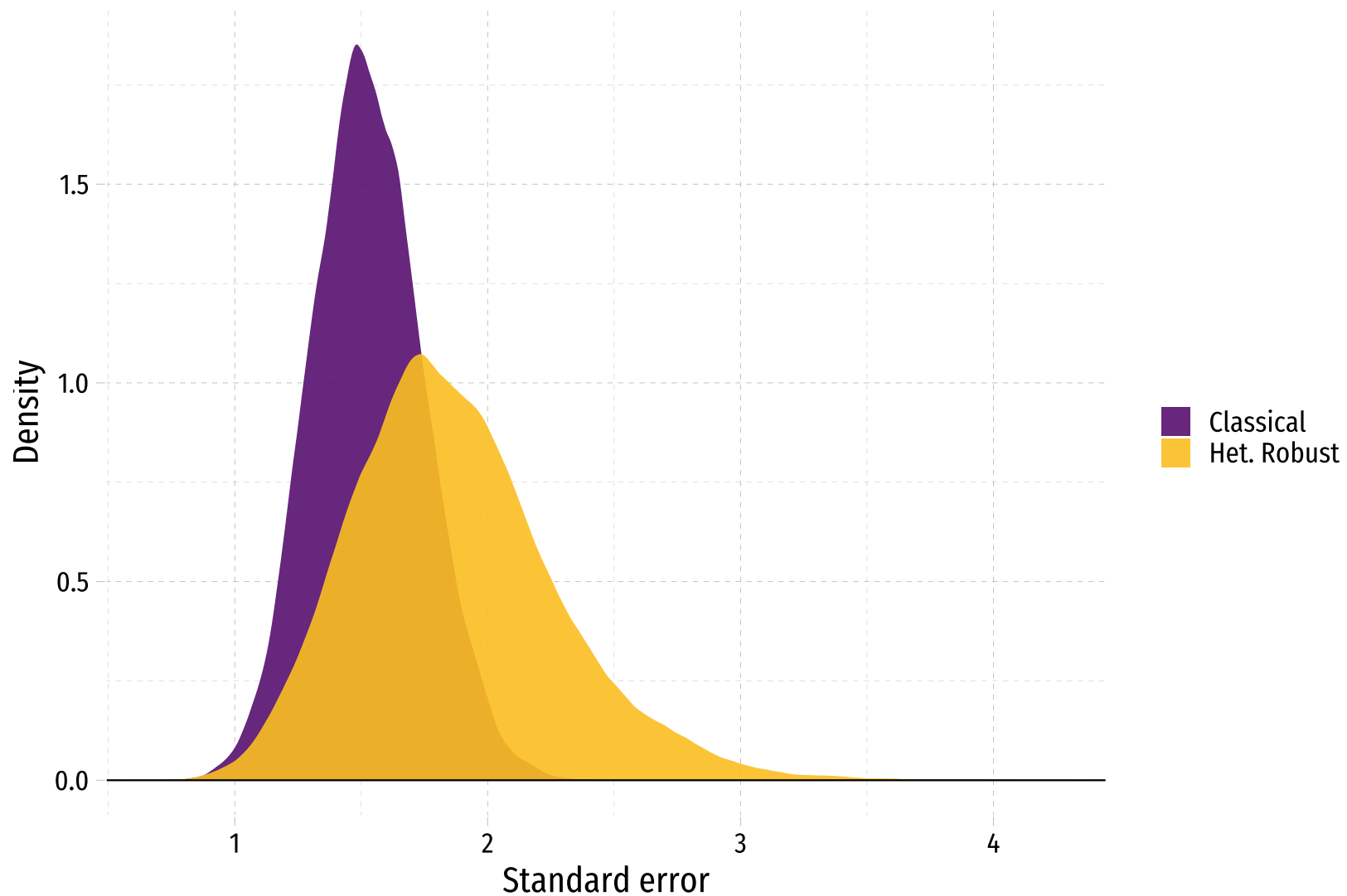
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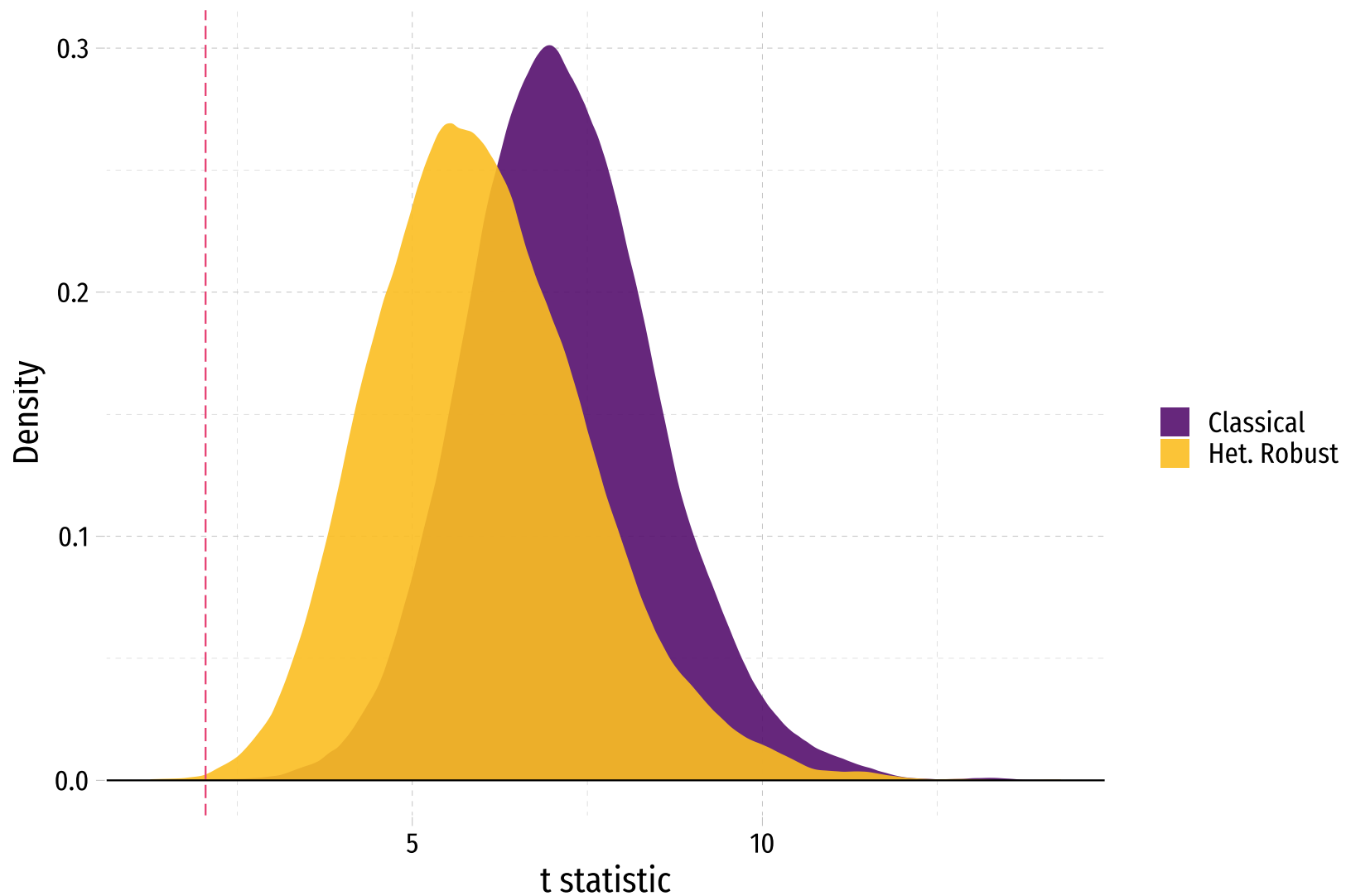
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sim_df = bind_rows(sim_list)
```

So what are the results?

Comparing the distributions of standard errors for the coefficient on  $x$



Comparing the distributions of  $t$  statistics for the coefficient on  $x$



Q All of these test are for a false  $H_0$ . How would the simulation change to enforce a *true* null hypothesis?



# Simulation

## Updating to enforce the null

Let's update our simulation function to take arguments  $\gamma$  and  $\delta$  such that

$$Y_i = 1 + e^{\gamma X_i} + \varepsilon_i$$

where  $\varepsilon_i \sim N(0, \sigma^2 X_i^\delta)$ .

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$$Y_i = 1 + e^{\gamma X_i} + \varepsilon_i$$

where  $\varepsilon_i \sim N(0, \sigma^2 X_i^\delta)$ .

In other words,

- $\gamma = 0$  implies no relationship between  $Y_i$  and  $X_i$ .
- $\delta = 0$  implies homoskedasticity.

# Simulation

## Updating to enforce the null

Updating the function...

```
flex_iter = function(iter,  $\gamma$  = 0,  $\delta$  = 1, n = 30) {  
  # Generate data  
  iter_df = tibble(  
    x = runif(n, min = 0, max = 10),  
     $\epsilon$  = rnorm(n, sd = 15 * x $\delta$ ),  
    y = 1 + exp( $\gamma$  * x) +  $\epsilon$   
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  # Estimate models  
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  # Stack and return results  
  bind_rows(tidy(lm1), tidy(lm2)) %>%  
    select(1:5) %>% filter(term == "x") %>%  
    mutate(se_type = c("classical", "HC2"), i = iter)  
}
```

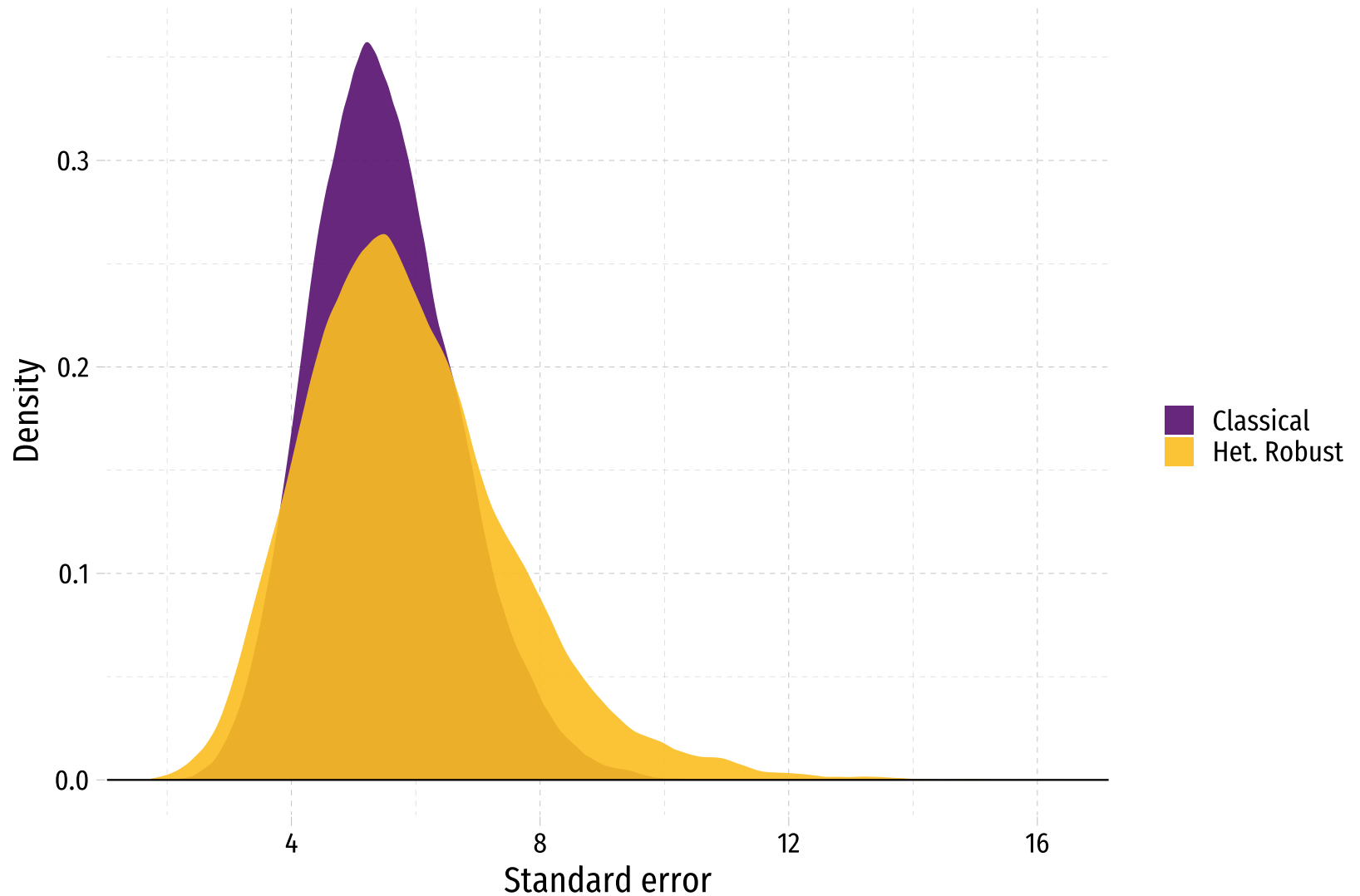
# Simulation

## Run again!

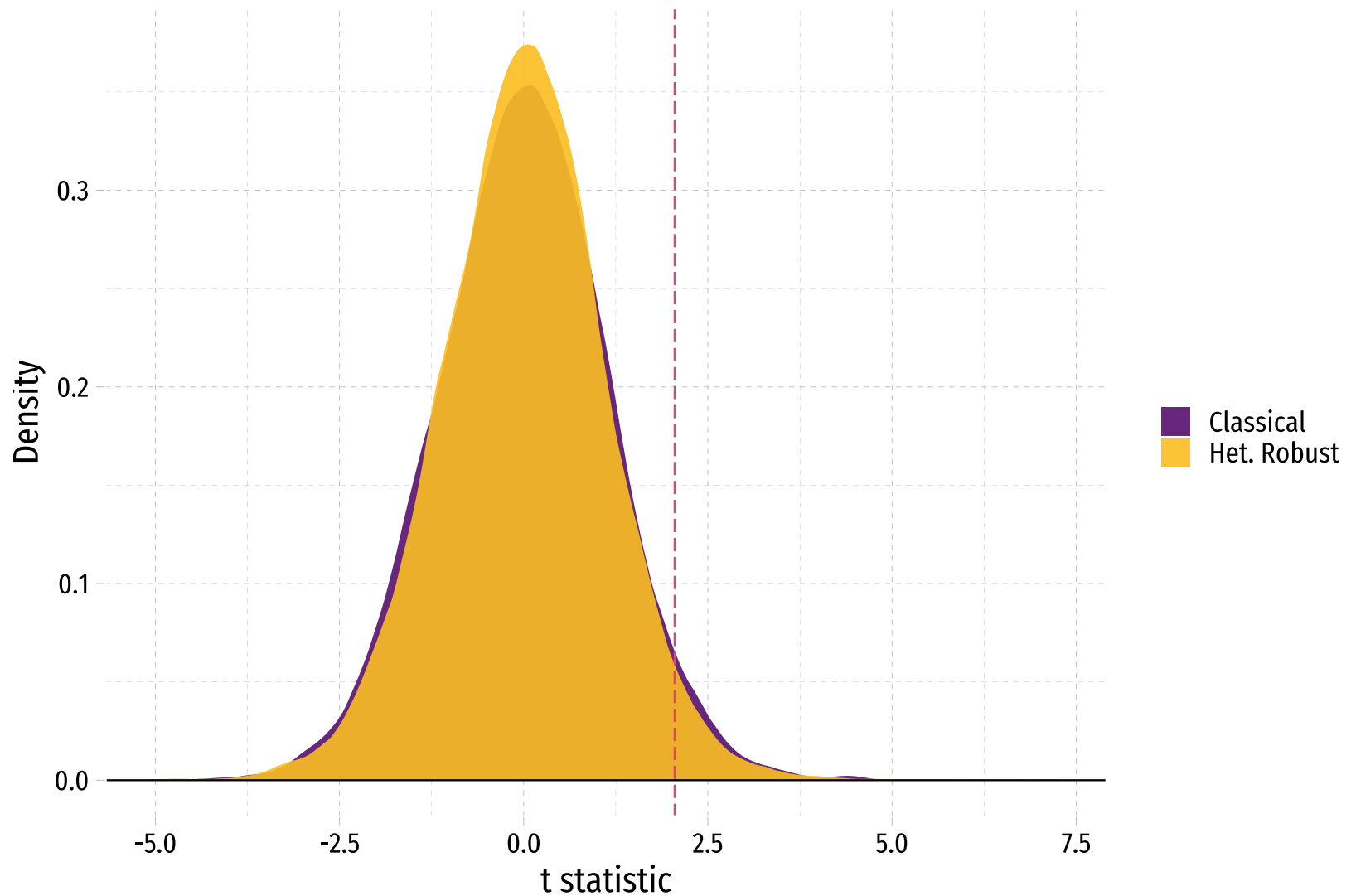
Now we run our new function `flex_iter()` 10,000 times

```
# Packages
p_load(purrr, furrr)
# Set options
set.seed(123)
# Tell R to parallelize
plan(multiprocess)
# Run 10,000 iterations
null_df = future_map(
  1:1e4, flex_iter,
  # Enforce the null hypothesis
   $\gamma = 0$ ,
  # Specify heteroskedasticity
   $\delta = 1$ ,
  .options = future_options(seed = T)
) %>% bind_rows()
```

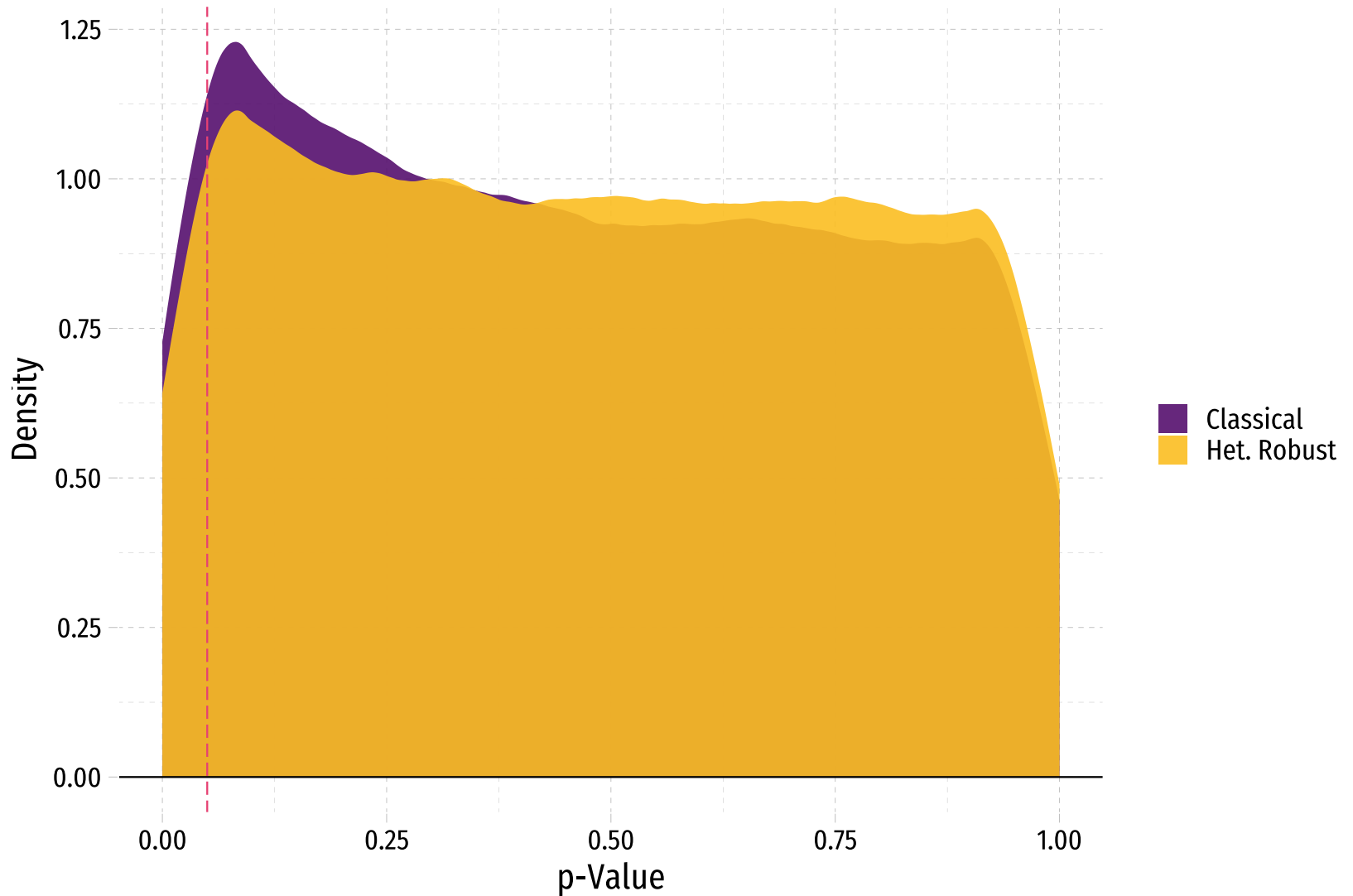
Comparing the distributions of standard errors for the coefficient on  $x$



Comparing the distributions of  $t$  statistics for the coefficient on  $x$



Distributions of  $p$ -values: both methods slightly over-reject the (true) null



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