Inequality Sensitive Optimal Treatment Assignment

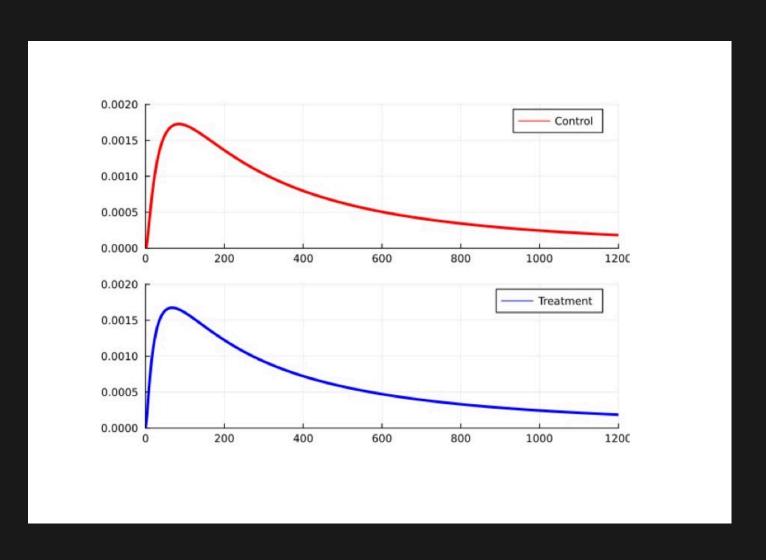
Eduardo Zambrano, Cal Poly

This material is based upon work supported by the NSF under Grant # 2313969

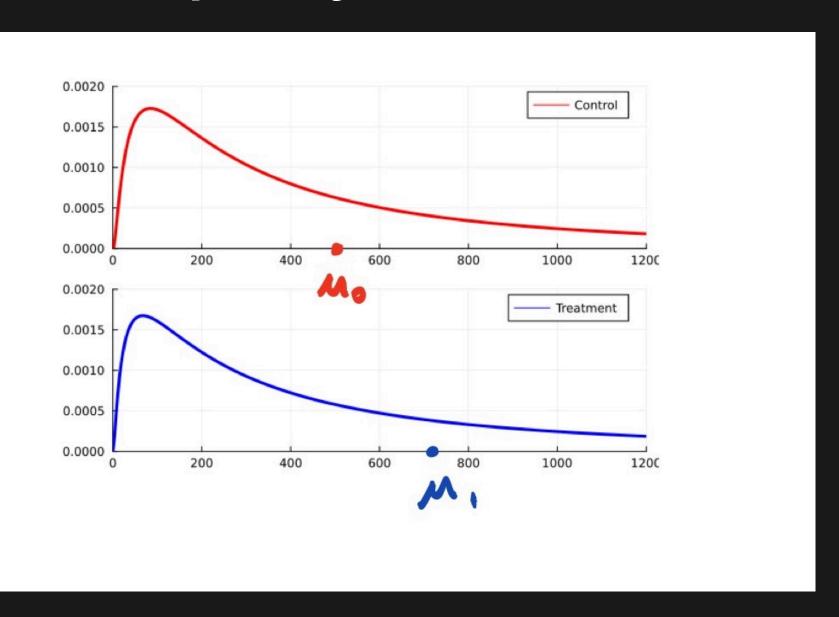
Preliminaries

- Project 1: Welfare Economics under Certainty (Theorem 1)
- Project 2: Welfare Economics under Uncertainty and Ambiguity (Theorem 2)

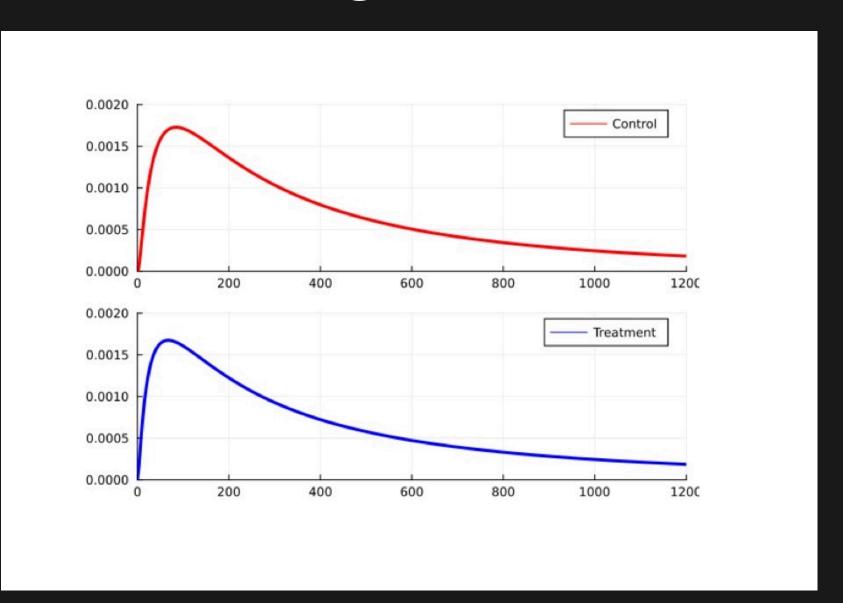
Two treatments and their outcome distributions



An inequality neutral evaluator



Incorporating inequality aversion



(From Project 1)

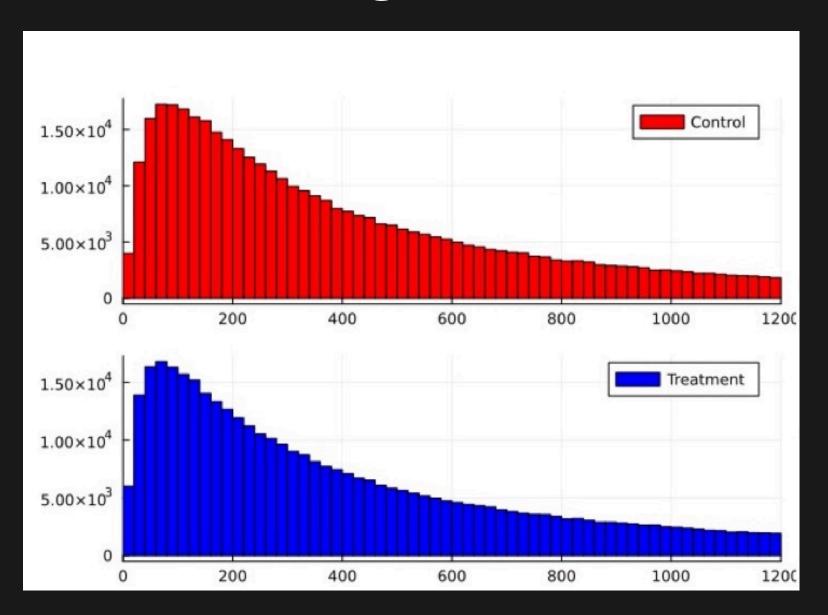
Theorem 1 (With Marc Fleurbaey) The Kolm-Atkinson class of SPs with $\gamma>1$ is (up to an affine transform) characterized by the requirement that there exist $\lambda,\mu\in(0,1)$ such that for all $y>0,\ddot{Y}(y)=\lambda y$ and $\mathring{Y}(y)=\mu y$.

One must then have $\lambda=2^{rac{1}{1-\gamma}}$ and $\mu=n^{rac{1}{1-\gamma}}$ for some $\gamma>1.$

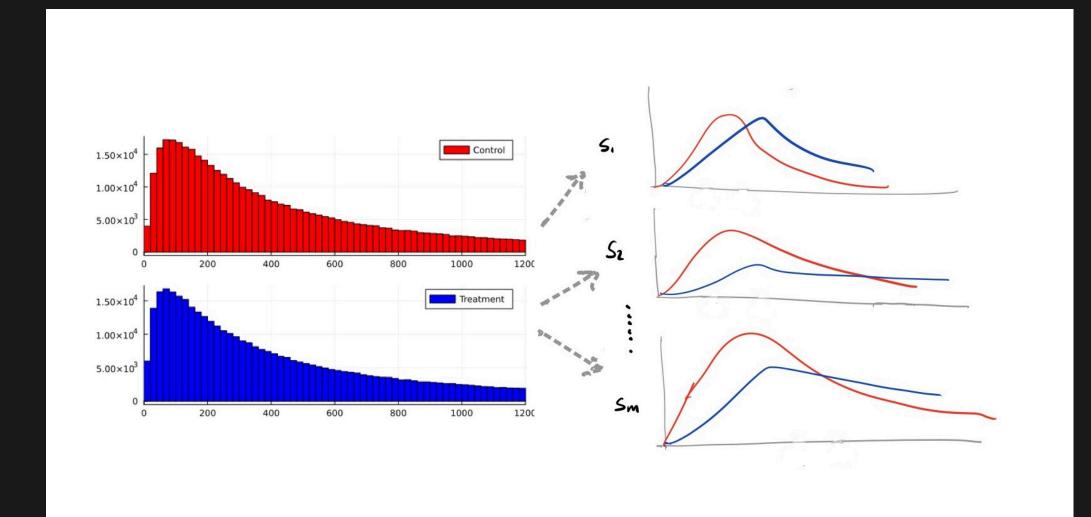
An illustration of Theorem 1

https://equistatlab.org/menu1/

Incorporating uncertainty



Incorporating uncertainty

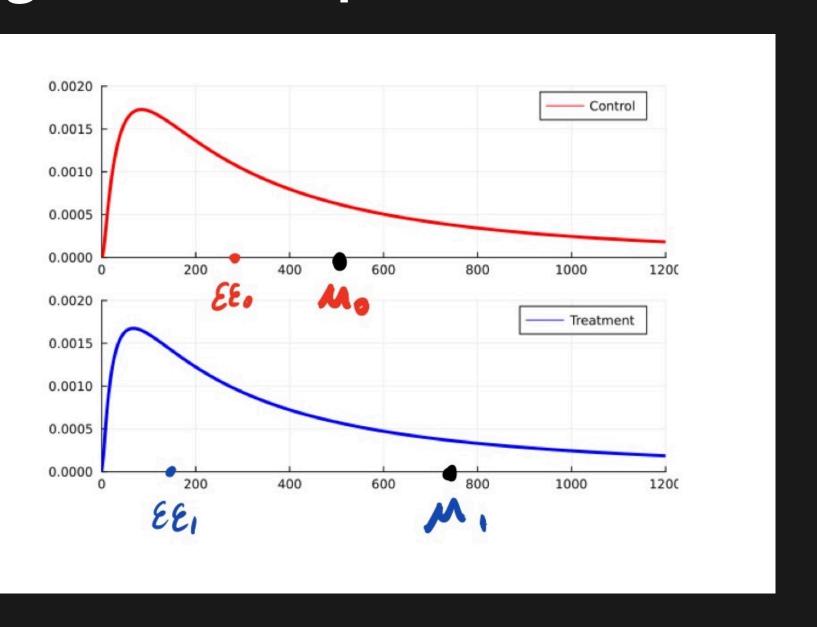


(From Project 2)

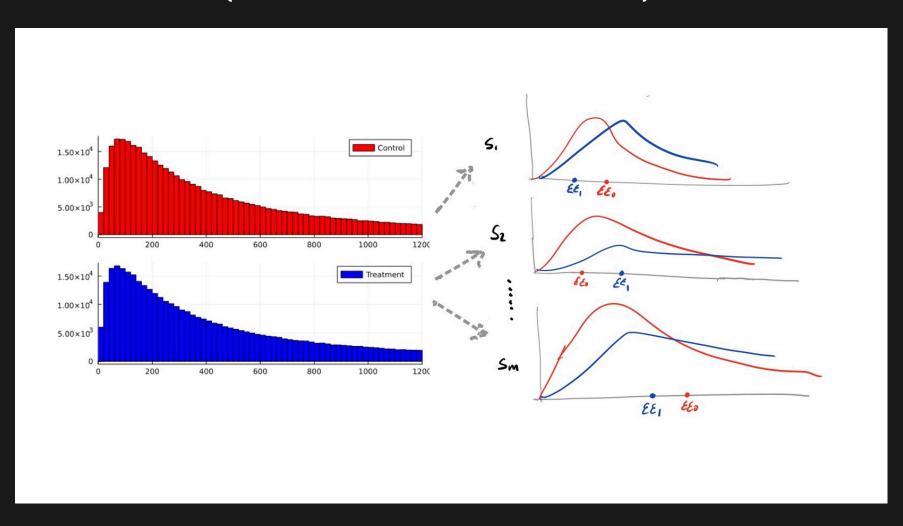
Theorem 2 Let $\Upsilon=\mathbb{R}^{nm}$ and R satisfy Continuity, Weak Dominance, Weak Pareto for No Risk, and Statistical Extension. For all $y,y^{'}\in\Upsilon$, one has

$$egin{aligned} yPy^{'} \ \iff \ & (\mathcal{EE}(y^{1}),\ldots,\mathcal{EE}(y^{m})) \succ (\mathcal{EE}(y^{'1}),\ldots,\mathcal{EE}(y^{'m})) \end{aligned}$$

Egalitarian Equivalent



Identify each treatment with the profile $(\mathcal{EE}^1,\dots,\mathcal{EE}^m)$



An illustration of Theorem 2

Table 1

This Paper (Project 3)

- Egalitarian Equivalent Optimal Statistical Decisions
 - Under point and partial identification
 - Small and large sample analysis
 - Applications: Microcredit, JobCorps

Statistical Decision Theory

• $d \in \{a, b\}$

$$ullet y^s(d) = egin{bmatrix} y^s_1(d) \ dots \ y^s_n(d) \end{bmatrix}$$

- $ullet ee^s(d) := ee(y^s(d)) = f^{-1}\left(rac{1}{n}\sum_{i=1}^n f(y^s_i(d))
 ight)$
- f is determined using Theorem 1

Egalitarian Equivalent Treatment Effect (EETE)

$$au(s) := ee^s(b) - ee^s(a)$$

Statistical Treatment Rules

ullet $\delta: \mathcal{P}
ightarrow [0,1]$

ullet $\delta: \mathcal{P}
ightarrow [0,1]$

$$ee^s(\delta(P_s)) =$$

$$\int_{0}^{\infty} f^{-1}\left[\delta(P_s)f(ee^s(b))+(1-\delta(P_s))f(ee^s(a))
ight]ds$$

Thanks to Theorem 2, the decision problem boils down to selecting δ in order to obtain the most favorable profile

$$(ee^1(\delta(P_s),\ldots,ee^m(\delta(P_s)))$$

according to the preferences \succeq over the restricted set Υ_1 , and given knowledge of the sampling distribution P_s .

- Let π be a prior on S
- for each P_s , let $S(P_s) \subseteq S$ denotes the state space obtained with knowledge of P_s .

The Bayes Decision Problem

$$\max_{\delta(P_s) \in [0,1]} E_\pi \left[ee^s(\delta(P_s)) | S(P_s)
ight]$$

$$(ee^1(\delta(P_s)),\ldots,ee^m(\delta(P_s)))$$

The Maximin Decision Problem

$$\max_{\delta(P_s) \in [0,1]} \min_{s \in S(P_s)} ee^s(\delta(P_s))$$

$$(ee^1(\delta(P_s)), \ldots, ee^m(\delta(P_s)))$$

The Minimax Regret Decision Problem

$$\min_{\delta(P_s) \in [0,1]} \max_{s \in S(P_s)} \left[\max\{ee^s(a), ee^s(b)\} - ee^s(\delta(P_s))
ight]$$

$$(ee^1(\delta(P_s)), \ldots, ee^m(\delta(P_s)))$$

Bayesian

$$\max_{\delta(P_s) \in [0,1]} E_{\pi} \left[\boxed{ee^s(\delta(P_s))} | S(P_s)
ight]$$

Maximin

$$\max_{\delta(P_s) \in [0,1]} \min_{s \in S(P_s)} \boxed{ee^s(\delta(P_s))}$$

Minimax Regret

$$\min_{\delta(P_s) \in [0,1]} \max_{s \in S(P_s)} \left[\max\{ee^s(a), ee^s(b)\} - \left[ee^s(\delta(P_s))
ight]
ight]$$

Let s_w , s_a and s_b be such that

$$ee^{s_w}(a)=ee^{s_b}(a)=\min_{s\in S(P_s)}ee^s(a)$$

$$ee^{s_w}(b)=ee^{s_a}(b)=\min_{s\in S(P_s)}ee^s(b)$$

and

$$ee^{s_d}(d) = \max_{s \in S(P_s)} ee^s(d)$$
 for $d \in \{a,b\}$.

Theorem 3 Assume that the true state s is partially identified, the set $\{(ee^s(a), ee^s(b))\}_{s \in S(P_s)}$ is bounded, and that $s_w, s_a, s_b \in S(P_s)$. Then

The solution to the Bayesian decision problem is

$$\delta^B(P_s) = 1 \left(E_\pi \left[au_{ee}(s) | S(P_s)
ight] > 0
ight)$$

Theorem 3 (cont.)

• The solution to the maximin decision problem is

$$\delta^M(P_s) = 1 \left(au_{ee}(s_w) > 0
ight)$$

Theorem 3 (cont.)

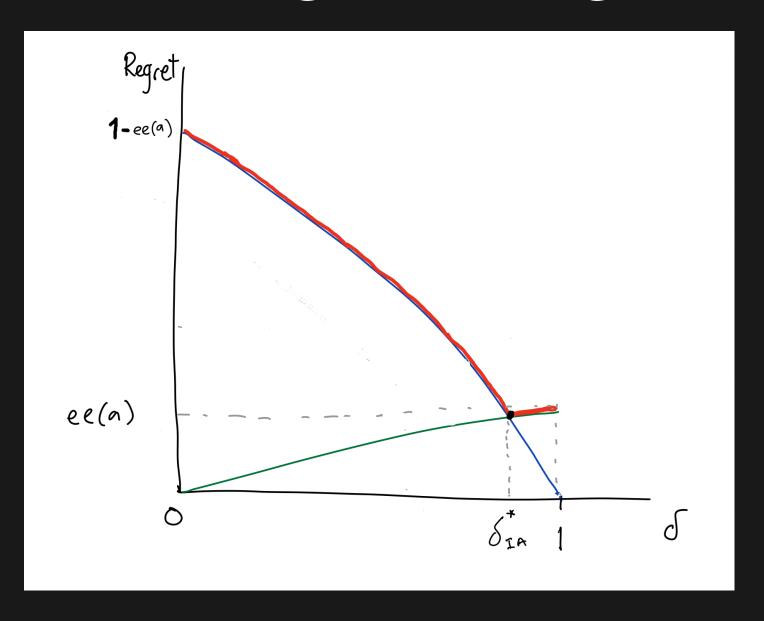
ullet The solution to the minimax regret decision problem is $\delta^R(P_s) \in (0,1)$ such that

$$(ee^{s_a}(a)-ee^{s_a}(\delta^R(P_s))=ee^{s_b}(b)-ee^{s_b}(\delta^R(P_s))$$

Side by Side

$$egin{align} \delta^B(P_s) &= 1 \left(E_\pi \left[ee(b) | S(P_s)
ight] - E_\pi \left[ee(a) | S(P_s)
ight] > 0
ight) \ \delta^M(P_s) &= 1 \left(ee^{s_w}(b) - ee^{s_w}(a) > 0
ight) \ ee^{s_a}(a) - ee^{s_a}(\delta^R(P_s)) = ee^{s_b}(b) - ee^{s_b}(\delta^R(P_s)) \ \end{aligned}$$

Minimizing Worst Regret



Application: A Bayesian Meta Analysis of the Microcredit Literature

 Meager 2022 estimates posterior distributions of the effect of microcredit interventions on consumption using data from randomized trials that expand access to microcredit in five countries.

Meager's results

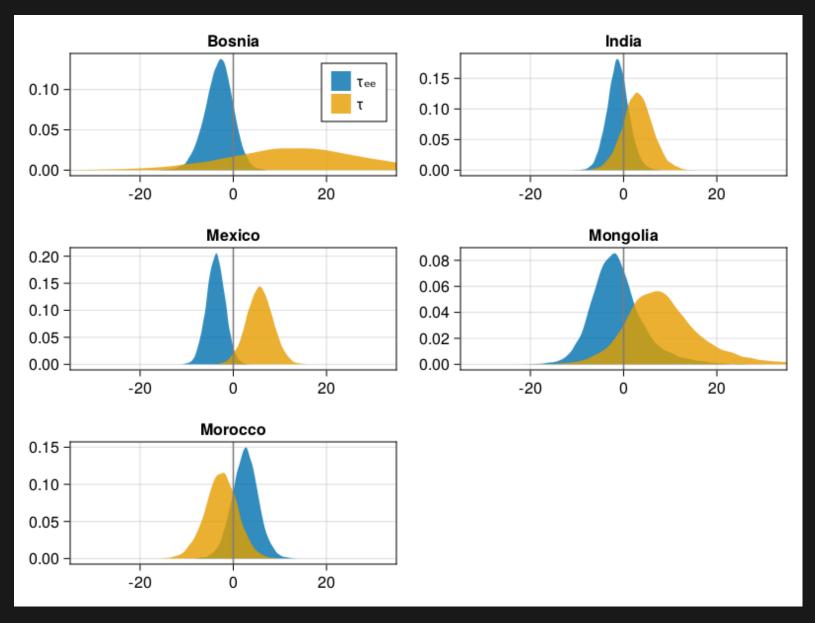
Meager's Bayesian hierarchical model

$$y_{ik}(T_{ik}) \sim exttt{LogNormal}(\mu_k + \zeta_k T_{ik}, \sigma_k \lambda_k^{T_{ik}}) ext{ for } k = 1, \dots, 5;$$

$$0.1\mu_k, 0.1\zeta_k, \log(\sigma_k), \log(\lambda_k) \sim \mathtt{MvN}(0, 10I) ext{ for } k=1,\ldots,5.$$

- I compute mean treatment effects and egalitarian equivalent treatment effects using Meager's Markov Chain Monte Carlo (MCMC) output, denoted $\hat{\pi}$, which contains three chains with four thousand draws per chain.
- The egalitarian equivalent at every draw can be computed using the expression

$$\mathcal{EE}(y_k(d)) = e^{\mu_k + \zeta_k d + rac{1}{2}(1-\gamma)\left(\sigma_k \lambda_k^d
ight)^2}.$$



Posterior distributions of the mean treatment effects (τ) and the egalitarian equivalent treatment effects (τ_{ee}). All units are 2009 USD PPP per two weeks.

Microcredit treatment effect Bayesian estimates

	$E_{\hat{\pi}}[au]$	$P_{\hat{\pi}}[au>0]$	$E_{\hat{\pi}}[au_{ee}]$	$P_{\hat{\pi}}[au_{ee}>0]$
Bosnia	13.82	82.5%	-3.11	14.3%
India	3.00	83.1%	-1.20	29.1%
Mexico	5.65	97.5%	-3.76	2.7%
Mongolia	8.58	88.1%	-1.35	35.0%
Morocco	-2.57	22.7%	2.55	82.4%

Bayesian quantile treatment effects on consumption (from Table 1 in Meager 2022)

Oceantiles	5+h	25th	25+h	15+h	55+h	65th	75+h	05th
Quantile:	5th	25th	35th	45th	55th	65th	75th	95th
Partial pooling								-
Bosnia	-5.2	-3.8	-3.9	-3.6	-2.8	-1.1	2.6	52.4
	(-9.6,2.6)	(-9,0.7)	(-9.9,1)	(-10.9,1.7)	(-12.4,3.6)	(-14.7, 8.6)	(-19.4,20.9)	(-75.8, 188.3)
India	-2	-1.2	-0.6	0.1	1.1	2.4	4.3	16
	(-5.3,1.4)	(-4.9,2.7)	(-4.7,3.6)	(-4.4,4.9)	(-4.2,6.6)	(-4,9)	(-4,12.8)	(-5.6,37.9)
Mexico	-4.7	-3.4	-2.2	-0.8	1.2	3.9	8	34.1
	(-7.3, -2.1)	(-6.5, -0.3)	(-5.8,1.2)	(-4.7,3)	(-3.5,5.6)	(-1.7,9.3)	(0.7,15)	(15.5,52.7)
Mongolia	-3	-1.7	-0.6	0.7	2.7	5.8	10.3	38.4
	(-11.4,5.3)	(-9.2,9.8)	(-7.5,12.2)	(-6,15.8)	(-5.3,20.2)	(-5.7,26.5)	(-7.3,36.2)	(-22.4,108)
Morocco	4.3	2.9	2	0.9	-0.4	-2.2	-4.6	-18.8
	(-0.5,9)	(-2.1,7.8)	(-3.1,7.1)	(-4.6,6.4)	(-6.7,5.6)	(-9.6,5)	(-14,4.4)	(-41.5,3.3)

Notes: All units are US\$ PPP per two weeks. Estimates are shown with their 95 percent uncertainty intervals below them in parentheses.

Discussion

- Interventions that exhibit considerable treatment effect heterogeneity can be difficult to evaluate and summarize in terms of welfare.
- The methods I propose aim to provide specific, quantitative guidance for how to do this evaluation.
- The microcredit empirical illustration I highlight above shows that performing the evaluation in an inequality sensitive way can make a difference in terms of how one ranks treatments, relative to when inequality considerations are put aside.

EETE estimation workflow

- 1. Use Theorem 1 to help determine which social preference under certainty to bring into the analysis.
- 2. Use Theorem 2 to decide how to incorporate that social preferences into a world where risk, uncertainty or ambiguity play a prominent role.
- 3. Use Theorem 3 to identify the correct optimal statistical treatment rule needed for the problem at hand.

Thank You For Coming!

Appendix 1: Protected Income and Inequality Aversion

- ullet Consider a population of n individuals
- ullet Each individual i has a known income $y_i \geq 0$.

Social Preferences

- $\overline{igl(ullet W \left(y_1, \ldots, y_n
 ight) = \sum_{i=1}^n f \left(y_i
 ight)}$
- ullet Let $\overline{y_1(y_2,y)}$ solve

$$f(y_1(y_2,y)) + f(y_2) = 2f(y)$$

Protected Income

$$\ddot{Y}\left(y
ight):=\inf\left\{ y_{1}\left(y_{2},y
ight)|\left(y_{2},y
ight)\in D_{2}
ight\} .$$

$$y_1(y_2,y)$$
 solves $f(y_1(y_2,y))+f(y_2)=2f(y)$

Kolm - Atkinson (K-A)

$$f(y) = egin{cases} rac{y^{1-\gamma}}{1-\gamma} & \gamma
eq 1, \gamma \geq 0 \ \ln y & \gamma = 1 \end{cases}$$

How much Inequality Aversion is not enough?

$$ullet$$
 $\ddot{Y}(y)=2^{rac{1}{1-\gamma}}y$ for $\gamma>1$

$$ullet$$
 $\ddot{Y}(y)=0$ for $\gamma\leq 1$

$$f(y) = egin{cases} rac{y^{1-\gamma}}{1-\gamma} & \gamma
eq 1, \gamma \geq 0 \ \ln y & \gamma = 1 \end{cases}$$

Protected Income (against n)

• $y_1(y_2,y_3,\ldots,y_n,y)$ is defined implicitly by the equation $f(y_1(y_2,y_3,\ldots,y_n,y))+f(y_2)+\cdots+f(y_n)=nf(y).$

In the case of the *K-A* class, one obtains

$$egin{aligned} \mathring{Y}(y) := \inf \left\{ y_1(y_2, y_3, \dots, y_n, y) | (y_2, y_3, \dots, y_n, y) \in D_n
ight\} \ &= n^{rac{1}{1-\gamma}} y \end{aligned}$$

Theorem 1 (With Marc Fleurbaey) The Kolm-Atkinson class of SPs with $\gamma>1$ is (up to an affine transform) characterized by the requirement that there exist $\lambda,\mu\in(0,1)$ such that for all $y>0, \ddot{Y}(y)=\lambda y$ and $\mathring{Y}(y)=\mu y$.

One must then have $\lambda=2^{\frac{1}{1-\gamma}}$ and $\mu=n^{\frac{1}{1-\gamma}}$ for some $\gamma>1$.

Appendix 2: Welfare Economics Under Uncertainty and Ambiguity

An extension of Theorem 1 in Fleurbaey (2010).

- ullet $N=\{1,\ldots,n\}$
- $ullet S = \{s_1, \ldots, s_m\}$
- $ullet y = (y_i^s)_{i \in N, s \in S}$
- $\Upsilon \subset \mathbb{R}^{nm}$
- ullet P,R,I: the evaluator's preferences over Υ
- ullet $[y^s]\in \Upsilon^c$
- ullet $(y_i)\in \Upsilon^e$
- \succ , \succeq , \sim : the evaluator's preferences over Υ^1

Continuity

Let $y,y^{'}\in\Upsilon$ and $(y(t))_{t\in\mathbb{N}}\in\Upsilon^{\mathbb{N}}$ be such that $y(t)\to y$. If $y(t)Ry^{'}$ for all $t\in\mathbb{N}$, then $yRy^{'}$. If $y^{'}Ry(t)$ for all $t\in\mathbb{N}$, then $y^{'}Ry$.

Weak Dominance

For all $y,y^{'}\in \Upsilon$, one has $yRy^{'}$ if for all $s\in S,[y^{s}]R[y^{s'}]$.

Weak Pareto for No Risk

For all $[y^s], [y^{'s}] \in \Upsilon^c$, one has $[y^s]P[y^{'s}]$ if, for all $i \in N, y^s_i > y^{'s}_i.$

Statistical Extension

For all $(y_i), (y_i^{'}) \in \Upsilon^e$, one has $(y_i)P(y_i^{'})$ if $y_i \succ y_i^{'}$.

Let $\mathcal{EE}(y^s)$ be the continuous function such that, for each $y^s \in \mathbb{R}^n$,

$$[y^s] \ I \ [(\mathcal{EE}(y^s), \dots, \mathcal{EE}(y^s))].$$

Theorem 2 Let $\Upsilon=\mathbb{R}^{nm}$ and R satisfy Continuity, Weak Dominance, Weak Pareto for No Risk, and Statistical Extension. For all $y,y^{'}\in\Upsilon$, one has

$$egin{aligned} yPy^{'}\ iff \ &(\mathcal{EE}(y^{1}),\ldots,\mathcal{EE}(y^{m}))\succ(\mathcal{EE}(y^{'1}),\ldots,\mathcal{EE}(y^{'m})) \end{aligned}$$

An Illustration

Table 1: Prospects 1 and 2

(a) Prospect 1					
	State 1	State 2			
y_1	4	2			
y_2	5	2			
y_3	6	2			

(b) Prospect 2

$$\pi = (1/3, 2/3)$$

Prospect 1	State 1	State 2
$oldsymbol{y_1}$	4	2
$oldsymbol{y_2}$	5	2
y_3	6	2
Geometric mean (across indiv.)	4.93	2
E[Geometric mean] (across states)		3.95

Prospect 2	State 1	Stat	e 2
y_1	3	1	
y_2	8	1	
y_3	10	1	
Geometric mean (across indiv.)	6.21	1	
E[Geometric mean] (across states)		4.48	

Prospect 1	State 1	State 2
$ln(y_1)$	1.31	0.69
$ln(y_2)$	1.61	0.69
$ln(y_3)$	1.79	0.69
Arithmetic mean of logs (across indiv.)	1.60	0.69
E[Arithmetic mean of logs] (across states)	1.2	29

Prospect 2	State 1	State 2
$ln(y_1)$	1.1	0
$ln(y_2)$	2.08	0
$ln(y_3)$	2.30	0
Arithmetic mean of logs (across indiv.)	1.83	0
E[Arithmetic mean of logs] (across states)	1	.22

Appendix 3: A Differences in Means comparison

- The point estimates (with standard errors in parentheses) of the differences in means E[y(1)]-E[y(0)] for the five sites are: Bosnia -1.59 (14.14), India 4.55 (3.85) (India), Mexico 5.51 (2.90), Mongolia 50.45 (15.67) and Morocco -2.93 (4.26).
- One would reject the hypothesis that the treatments have the same effect on average income at the 5% level in the case of Mongolia, and would not reject the hypothesis in the other four countries.

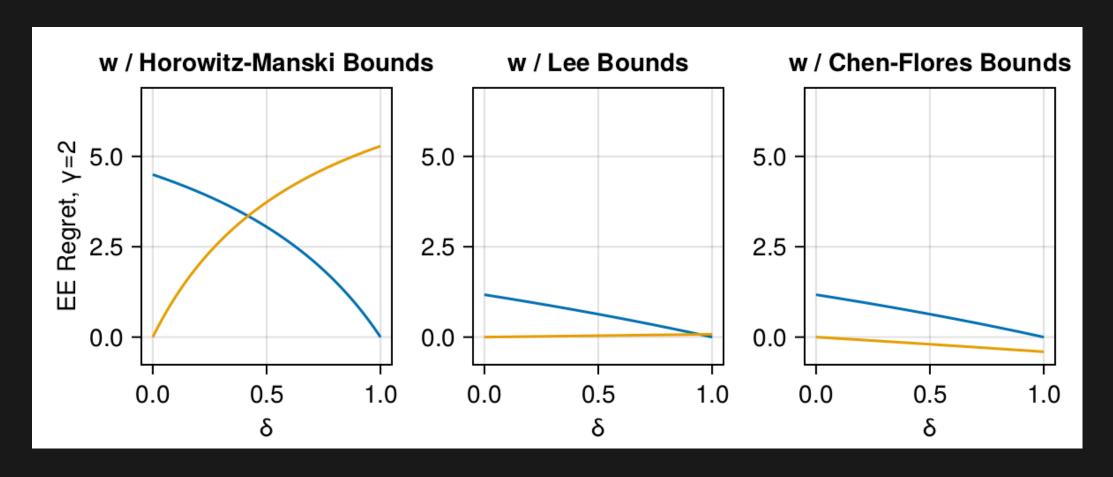
Appendix 4: Minimizing Egalitarian Equivalent Regret at JobCorps

- JobCorps is a widely studied RCT education and training program for disadvantaged youth.
- The evaluator only observes wages for those individuals who are employed.
- $ee^s(a)$ and $ee^s(b)$ are not point identified and therefore $\tau_{ee}(s)$ is also not point identified at s.
- Partial identification is achievable under relatively mild assumptions.

Egalitarian Equivalent Bounds

$\gamma=2$				
	$ee^L(a)$	$ee^U(a)$	$ee^L(b)$	$ee^U(b)$
Horowitz and Manski	4.1	9	3.8	8.6
Lee	6.6	6.6	6.5	7.7
Chen and Flores	6.6	6.6	6.8	7.7

Minimizing Egalitarian Equivalent Regret at JobCorps



Optimal treatment assignment

$$\gamma=2$$
Horowitz Lee Bounds Chen and and Manski Flores
Bounds Bounds δ^* 0.42 0.95 1

