

REGRESSÃO E CORRELAÇÃO



COEFICIENTE DE CORRELAÇÃO

Coeficiente de correlação de Pearson

$$R = \frac{\sum (X_i - \bar{X})(Y_i - \bar{Y})}{\sqrt{\sum (X_i - \bar{X})^2} \sqrt{\sum (Y_i - \bar{Y})^2}} = \frac{s_{XY}}{\sqrt{s_{xx}} \cdot \sqrt{s_{yy}}}$$



TESTES DE ASSOCIAÇÃO

Unilateral à direita

$$H_0 : \rho = 0$$

$$H_1 : \rho > 0$$

Unilateral à esquerda

$$H_0 : \rho = 0$$

$$H_1 : \rho < 0$$

Bilateral

$$H_0 : \rho = 0$$

$$H_1 : \rho \neq 0$$

Estatística de teste

$$t = \frac{r \cdot \sqrt{n - 2}}{\sqrt{1 - r^2}}$$

Região de Rejeição:

$$t > t_{n-2(\alpha)}$$

$$t < -t_{n-2(\alpha)}$$

$$|t| > t_{n-2(\alpha/2)}$$

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EXEMPLO

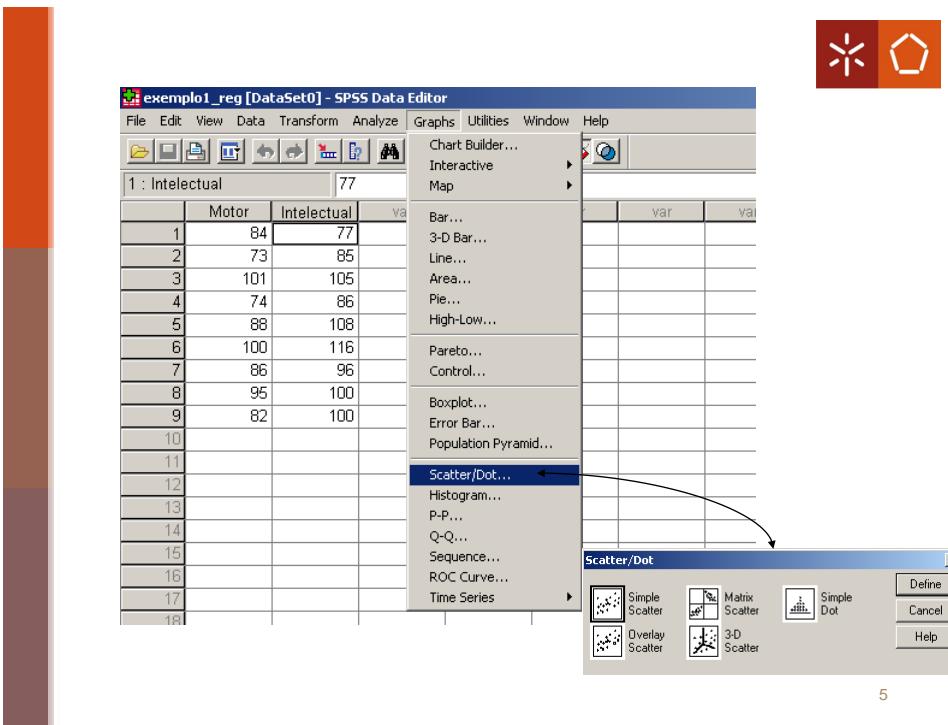


■ Índice de Desenvolvimento de Griffiths

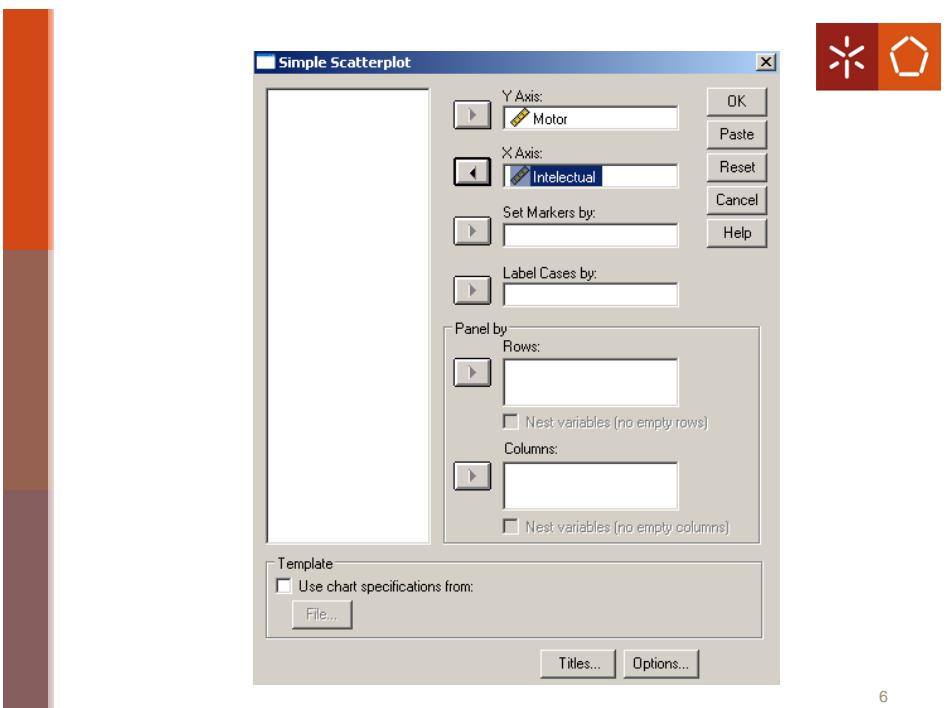
- avaliações motora e intelectual para 9 crianças com a idade de 4 anos

Motor	Intelectual
84	77
73	85
101	105
74	86
88	108
100	116
86	96
95	100
82	100

4



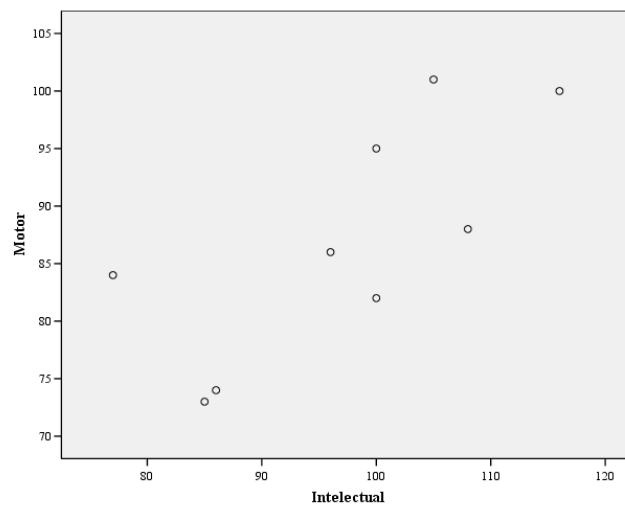
5



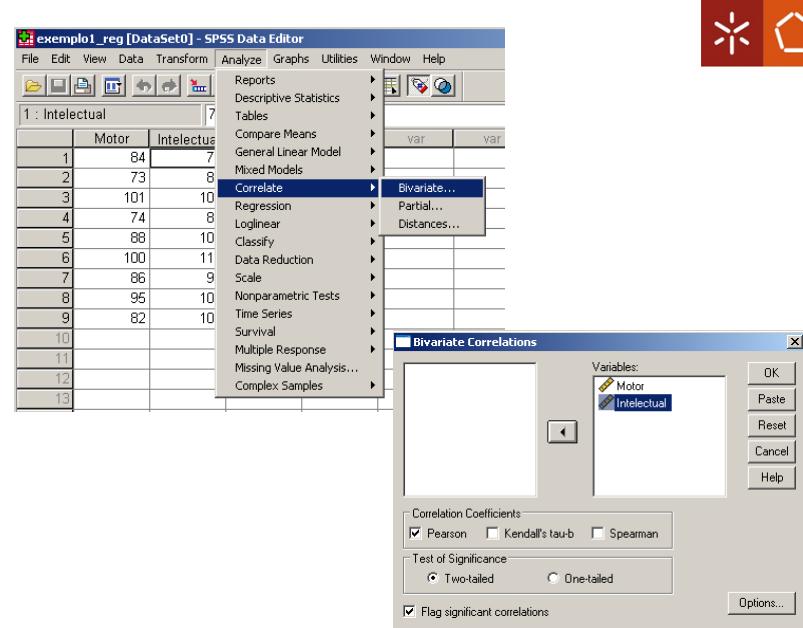
6



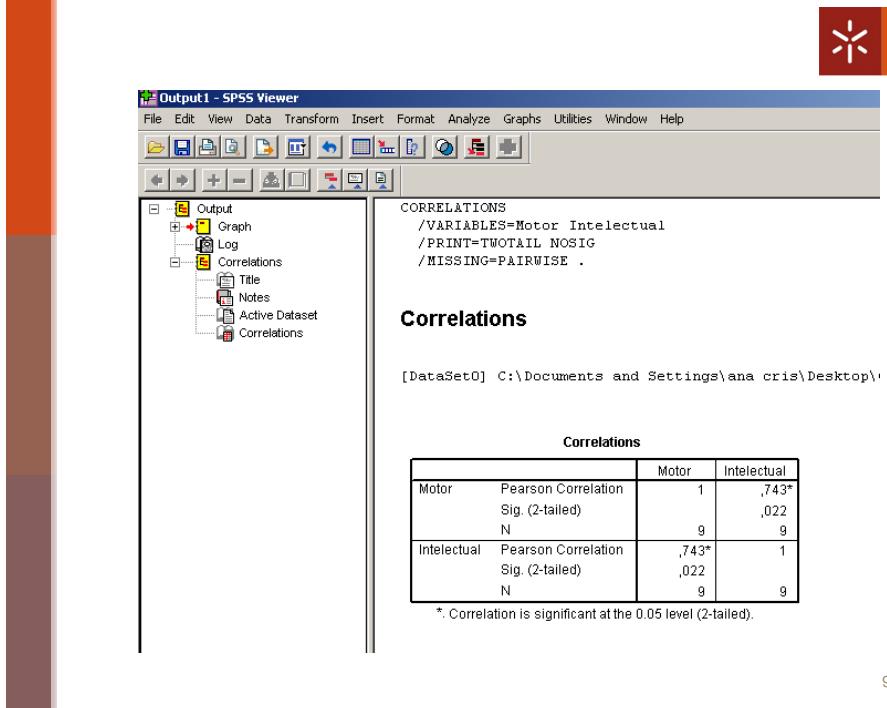
DIAGRAMA DE DISPERSÃO



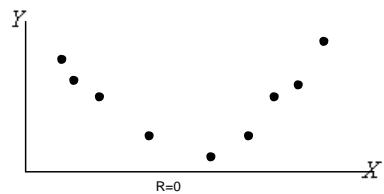
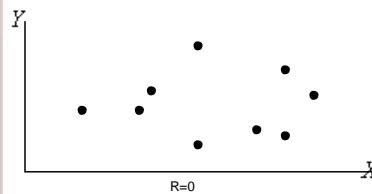
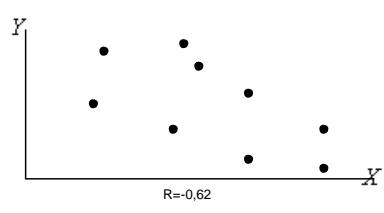
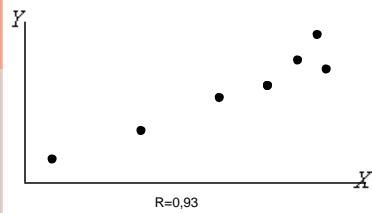
7



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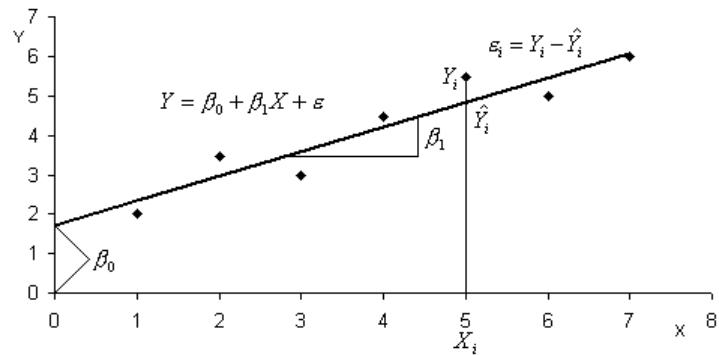
CORRELAÇÃO



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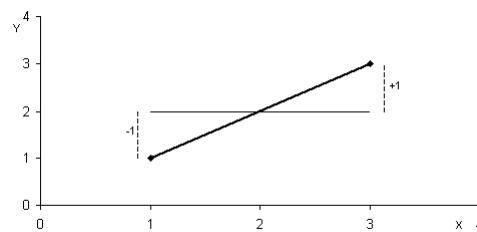
AJUSTE DE UMA RETA



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MINIMIZAÇÃO DOS DESVIOS

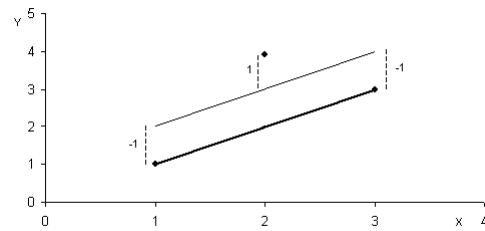


$$\sum(Y_i - \hat{Y}_i)$$

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MINIMIZAÇÃO DOS DESVIOS ABSOLUTOS



$$\sum |Y_i - \hat{Y}_i|$$

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EXEMPLO 1

- Consider the following set of points

X	Y
1	1
2	1
3	2
4	2
5	4

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RETAS DE AJUSTE

R1 $Y = -0,1 + 0,7X$

R2 $Y = 0,5 + 0,5X$

R3 $Y = -0,7 + 0,9X$

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RETAS



R1	R2	R3
0,6	1	0,2
1,3	1,5	1,1
2	2	2
2,7	2,5	2,9
3,4	3	3,8

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DESVIOS

Desv1	Desv2	Desv3
0,4	0	0,8
-0,3	-0,5	-0,1
0	0	0
-0,7	-0,5	-0,9
0,6	1	0,2
0	0	0

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DESVIOS ABSOLUTOS



Desv1	Desv2	Desv3
0,4	0	0,8
0,3	0,5	0,1
0	0	0
0,7	0,5	0,9
0,6	1	0,2
2	2	2

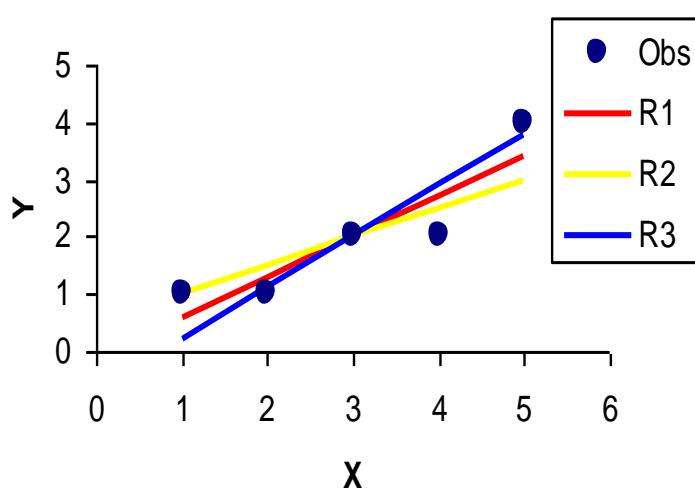
18



QUADRADO DOS DESVIOS

$(Desv1)^2$	$(Desv2)^2$	$(Desv3)^2$
0,16	0	0,64
0,09	0,25	0,01
0	0	0
0,49	0,25	0,81
0,36	1	0,04
1,10	1,50	1,50

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EXEMPLO 2

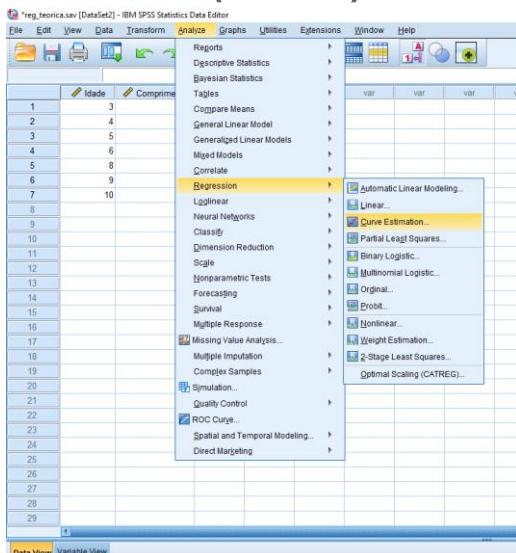
Comprimento alar (cm) em função da idade (dias) para andorinhas

Dias	Comp.
3	1,4
4	1,5
5	2,1
6	2,4
8	3,1
9	3,2
10	3,3

DPS

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EXEMPLO 2 (SPSS)

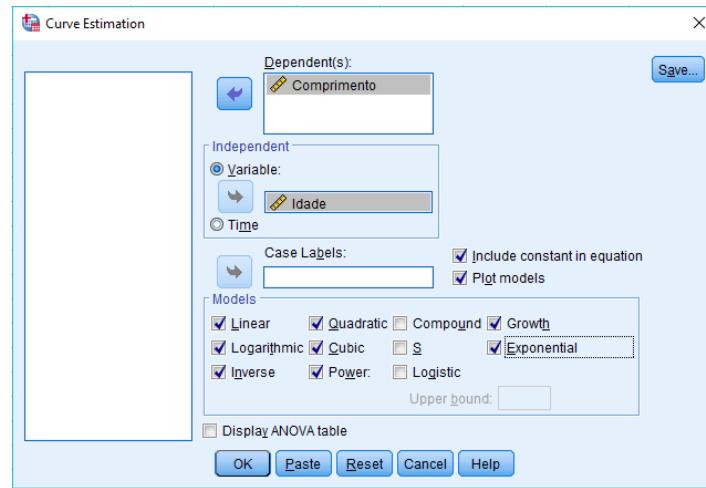


DPS

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EXEMPLO 2



DPS

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Models (curve estimation algorithms)

[Previous](#) [Next](#)

CURVEFIT allows the user to specify a model with or without a constant term designated by β_0 . If this constant term is excluded, simply set it zero or one depending upon whether it appears in an additive or multiplicative manner in the models listed below.

- | | |
|------------------|--|
| (1) Linear | $E(Y_t) = \beta_0 + \beta_1 t$ |
| (2) Logarithmic | $E(Y_t) = \beta_0 + \beta_1 \ln(t)$ |
| (3) Inverse | $E(Y_t) = \beta_0 + \beta_1 / t$ |
| (4) Quadratic | $E(Y_t) = \beta_0 + \beta_1 t + \beta_2 t^2$ |
| (5) Cubic | $E(Y_t) = \beta_0 + \beta_1 t + \beta_2 t^2 + \beta_3 t^3$ |
| (6) Compound | $E(Y_t) = \beta_0 \beta_1^t$ |
| (7) Power | $E(Y_t) = \beta_0 t^{\beta_1}$ |
| (8) S | $E(Y_t) = \exp(\beta_0 + \beta_1 / t)$ |
| (9) Growth | $E(Y_t) = \exp(\beta_0 + \beta_1 t)$ |
| (10) Exponential | $E(Y_t) = \beta_0 e^{\beta_1 t}$ |
| (11) Logistic | $E(Y_t) = \left(\frac{1}{u} + \beta_0 \beta_1^t \right)^{-1}$ |

DPS

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EXEMPLO 2



*Output1 [Document1] - IBM SPSS Statistics Viewer

File Edit View Data Transform Insert Format Analyze Graphs Utilities Extensions Window Help

Curve Fit

[DataSet2] C:\Users\ACB\OneDrive\Aulas2017_18\reg_teorica.sav

Model Description

Model Name	MOD_1
Dependent Variable	1 Comprimento
Equation	1 Linear
	2 Logarithmic
	3 Inverse
	4 Quadratic
	5 Cubic
	6 Power ^a
	7 Growth ^a
	8 Exponential ^a
Independent Variable	Idade
Constant	Included
Variable Whose Values Label Observations in Plots	Unspecified
Tolerance for Entering Terms in Equations	0,0001

a. The model requires all non-missing values to be positive.

DPS

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Variable Processing Summary

	Variables	
	Dependent	Independent
	Comprimento	Idade
Number of Positive Values	7	7
Number of Zeros	0	0
Number of Negative Values	0	0
Number of Missing Values	User-Missing	0
	System-Missing	0

Model Summary and Parameter Estimates

Dependent Variable: Comprimento

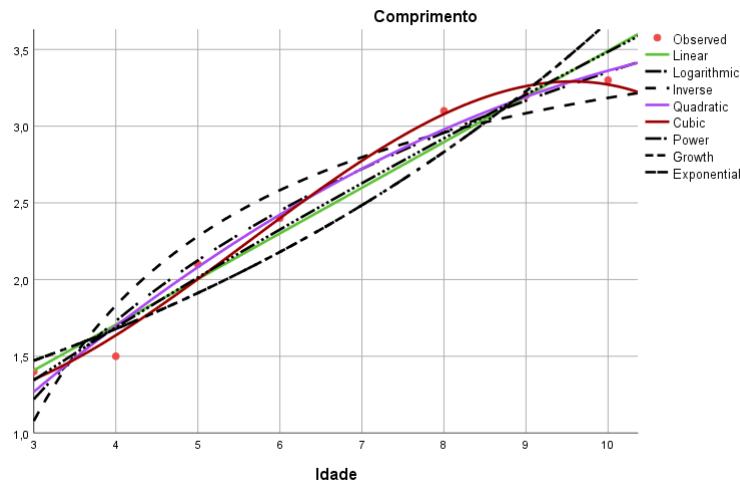
Equation	R Square	F	df1	df2	Sig.	Parameter Estimates			
						Constant	b1	b2	b3
Linear	0,964	132,174	1	5	0,000	0,515	0,298		
Logarithmic	0,971	165,753	1	5	0,000	-0,727	1,772		
Inverse	0,915	53,833	1	5	0,001	4,087	-9,026		
Quadratic	0,980	99,685	2	4	0,000	-0,274	0,579	-0,021	
Cubic	0,991	106,896	3	3	0,002	1,471	-0,387	0,141	-0,008
Power	0,968	149,638	1	5	0,000	0,563	0,792		
Growth	0,931	67,190	1	5	0,000	-0,006	0,131		
Exponential	0,931	67,190	1	5	0,000	0,994	0,131		

The independent variable is Idade.

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EXEMPLO 2

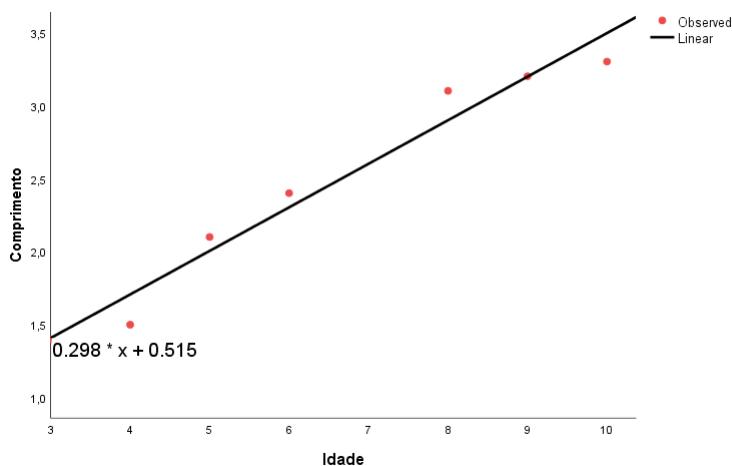


DPS

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RECTA DE MÍNIMOS QUADRADOS

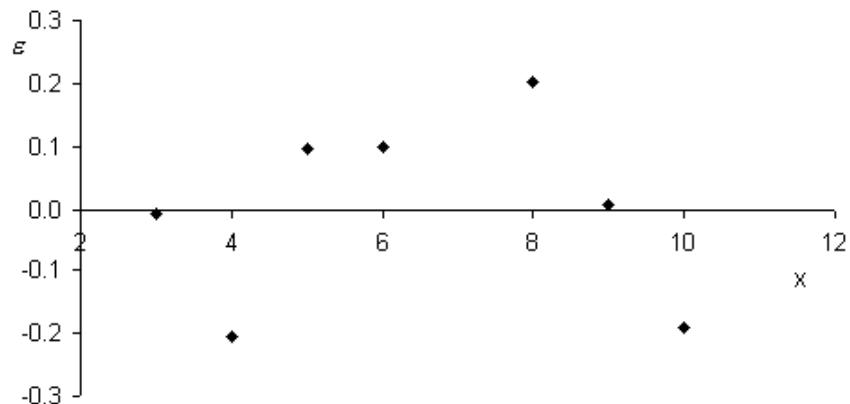


DPS

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RESÍDUOS



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Estimadores

$$Y_i = \beta_0 + \beta_1 \cdot (X_i - \bar{X}) + \varepsilon_i \quad i = 1, \dots, n$$

$$\hat{\beta}_0 = \frac{1}{n} \sum_i Y_i = \bar{Y}$$

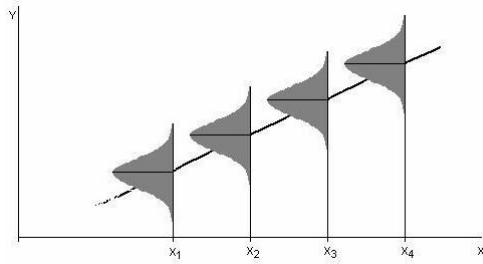
$$\hat{\beta}_1 = \frac{\sum_i (X_i - \bar{X})(Y_i - \bar{Y})}{\sum_i (X_i - \bar{X})^2} = \frac{s_{xy}}{s_{xx}}$$

$$\sigma^2 = \frac{1}{n-2} \sum_i \hat{\varepsilon}_i^2 = \frac{1}{n-2} \sum_i \left\{ Y_i - [\hat{\beta}_0 + \hat{\beta}_1 \cdot (X_i - \bar{X})] \right\}^2$$

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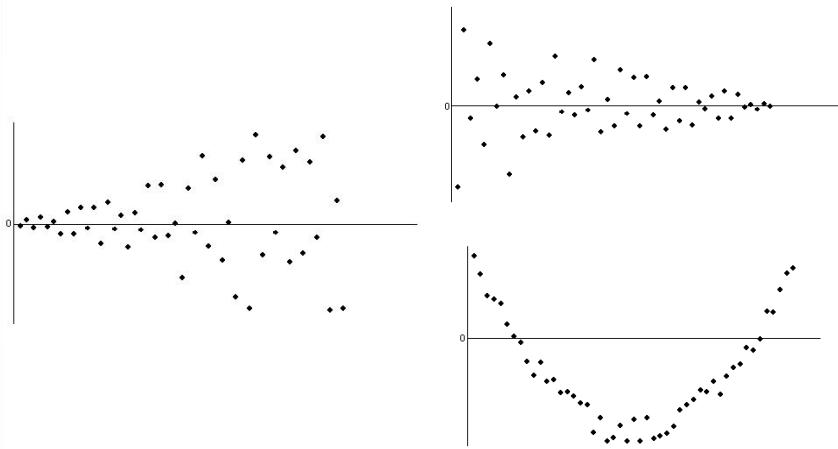


DISTRIBUIÇÃO DOS ERROS



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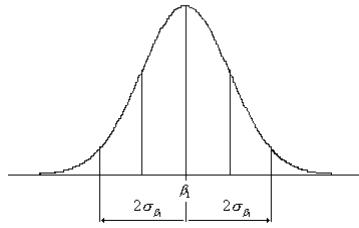
RESÍDUOS



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DISTRIBUIÇÃO DO DECLIVE



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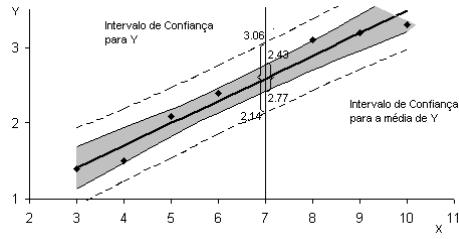
IC e Testes de hipóteses

	IC	TH
β_0	$\hat{\beta}_0 \pm t_{n-2,(\alpha/2)} \cdot \frac{s}{\sqrt{n}}$	$H_0: \beta_0 = b_0$ $H_1: \beta_0 \neq b_0, \beta_0 > b_0 \text{ ou } \beta_0 < b_0$ $ET = \frac{\hat{\beta}_0 - b_0}{s/\sqrt{n}}$ $H_0 \text{ verdadeira} \Rightarrow ET \sim t_{n-2}$
β_0'	$\left(\hat{\beta}_0 - \bar{X} \cdot \hat{\beta}_1 \right) \pm t_{n-2,(\alpha/2)} \cdot s \cdot \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{S_{XX}}}$	$H_0: \beta_0' = b_0'$ $H_1: \beta_0' \neq b_0', \beta_0' > b_0' \text{ ou } \beta_0' < b_0'$ $ET = \frac{\left(\hat{\beta}_0 - \bar{X} \cdot \hat{\beta}_1 \right) - b_0'}{s \cdot \sqrt{\frac{1}{n} + \frac{\bar{X}^2}{S_{XX}}}}$ $H_0 \text{ verdadeira} \Rightarrow ET \sim t_{n-2}$
β_1	$\hat{\beta}_1 \pm t_{n-2,(\alpha/2)} \cdot \frac{s}{\sqrt{S_{XX}}}$	$H_0: \beta_1 = b_{10}$ $H_1: \beta_1 \neq b_{10}, \beta_1 > b_{10} \text{ ou } \beta_1 < b_{10}$ $ET = \frac{\hat{\beta}_1 - b_{10}}{\frac{s}{\sum_i (x_i - \bar{X})^2}}$ $H_0 \text{ verdadeira} \Rightarrow ET \sim t_{n-2}$

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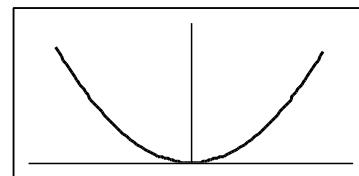
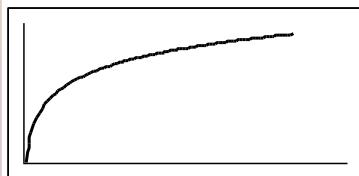
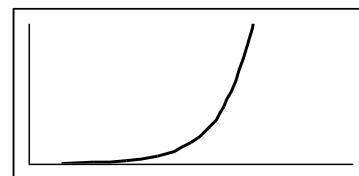
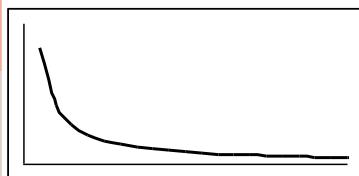
INTERVALO DE CONFIANÇA



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REGRESSÃO NÃO LINEAR



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REGRESSÃO NÃO LINEAR

Modelo	Transformação
<ul style="list-style-type: none"> • $Y_i = \alpha' + \frac{\beta}{X_i} + e_i$ 	$U_i = \frac{1}{X_i}$ $Y_i = \alpha' + \beta U_i + e_i$
<ul style="list-style-type: none"> • $Y_i = e^{\alpha' + \beta X_i + e_i}$ 	$Z_i = \ln Y_i$ $Z_i = \alpha' + \beta X_i + e_i$
<ul style="list-style-type: none"> • $Y_i = e^{\alpha' + \frac{\beta}{X_i} + e_i}$ com $\alpha' > 0, \beta < 0$ 	$U_i = \frac{1}{X_i}$ $Z_i = \ln Y_i$ $Z_i = \alpha' + \beta U_i + e_i$

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REGRESSÃO LINEAR E MÚLTIPLA

Um modelo de regressão linear múltipla descreve uma relação entre várias variáveis quantitativas **independentes**, X_1, X_2, \dots, X_J e uma variável quantitativa **dependente**, Y , nos termos seguintes:

$$Y_i = \beta_0 + \beta_1 \cdot (X_{1i} - \bar{X}_1) + \beta_2 \cdot (X_{2i} - \bar{X}_2) + \dots + \beta_j \cdot (X_{ji} - \bar{X}_j) + \varepsilon_i \quad i = 1, \dots, n$$

$$j = 1, \dots, J$$

onde:

- $(X_{1i}, X_{2i}, \dots, X_{ji}, Y_i)$ i-ésima observação das variáveis $X_{1i}, X_{2i}, \dots, X_{ji}$ e Y .
- \bar{X}_j média aritmética das observações X_{ji}
- $\beta_0, \beta_1, \beta_2, \dots, \beta_J$ parâmetros fixos da relação linear entre $X_{1i}, X_{2i}, \dots, X_{ji}$ e Y
- ε_i erro aleatório associado ao valor observado Y_i



RESÍDUOS

Pressupostos para ε_i

- Têm valor esperado nulo e variância constante, σ^2 ;
 - São mutuamente independentes;
 - São normalmente distribuídos.
- $\left. \begin{array}{l} \text{• Têm valor esperado nulo e variância constante, } \sigma^2; \\ \text{• São mutuamente independentes;} \\ \text{• São normalmente distribuídos.} \end{array} \right\} \varepsilon_i \sim IN(0, \sigma^2)$

Se estas hipóteses se verificarem então: $Y_i \sim IN(\mu_{Y_i}, \sigma^2)$

3
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ESTIMADORES MÍNIMOS QUADRADOS



$$\hat{\beta}_0 = \frac{1}{n} \sum_i Y_i = \bar{Y}$$

$$\begin{cases} \hat{\beta}_1 \cdot S_{X_1 X_1} + \hat{\beta}_2 \cdot S_{X_1 X_2} + \dots + \hat{\beta}_J \cdot S_{X_1 X_J} = S_{X_1 Y} \\ \hat{\beta}_1 \cdot S_{X_2 X_1} + \hat{\beta}_2 \cdot S_{X_2 X_2} + \dots + \hat{\beta}_J \cdot S_{X_2 X_J} = S_{X_2 Y} \\ \dots \\ \hat{\beta}_1 \cdot S_{X_J X_1} + \hat{\beta}_2 \cdot S_{X_J X_2} + \dots + \hat{\beta}_J \cdot S_{X_J X_J} = S_{X_J Y} \end{cases}$$

$$S_{X_{j_1} X_{j_2}} = \sum_n (X_{j_1 i} - \bar{X}_{j_1})(X_{j_2 i} - \bar{X}_{j_2})$$

$$S_{X_j Y} = \sum_n (X_{j i} - \bar{X}_j)(Y_i - \bar{Y})$$

$$\begin{aligned} s^2 &= \frac{1}{n - J - 1} \sum_i \hat{\varepsilon}_i^2 = \\ &= \frac{1}{n - J - 1} \sum_i \left\{ Y_i - [\hat{\beta}_0 + \hat{\beta}_1 \cdot (X_{1i} - \bar{X}_1) + \hat{\beta}_2 \cdot (X_{2i} - \bar{X}_2) + \dots + \hat{\beta}_J \cdot (X_{Ji} - \bar{X}_J)] \right\}^2 \end{aligned}$$

4
0

EXEMPLO 3



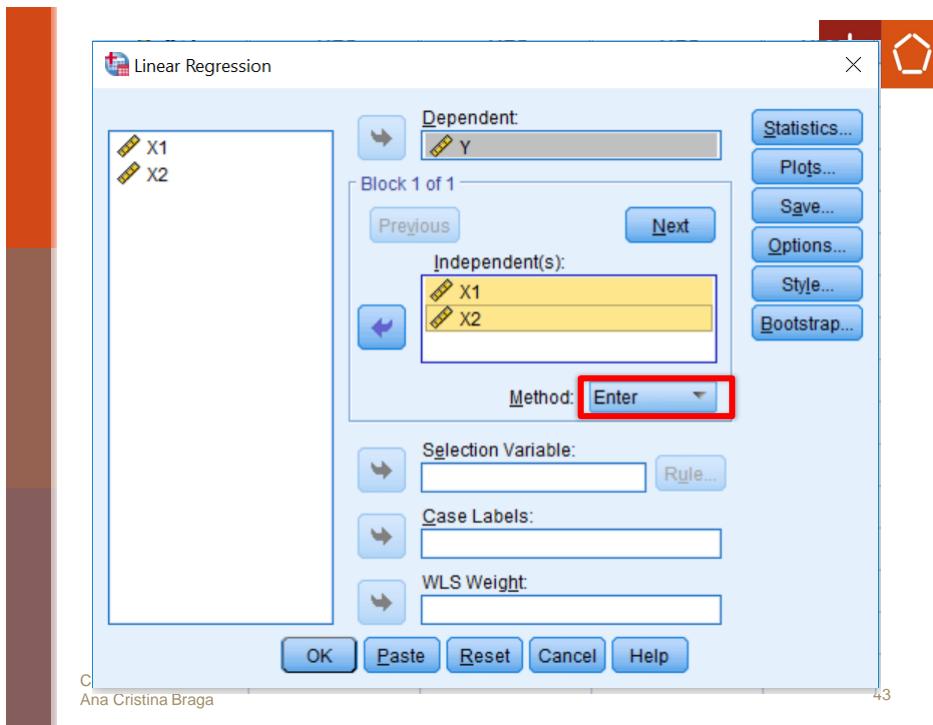
Determine a relação existente entre o calor envolvido no endurecimento, representado pela variável Y e os pesos de duas substâncias X_1 e X_2 , tendo em consideração os seguintes valores obtidos numa experiência:

Y	78,5	74,3	104,3	87,6	95,6	109,2	102,7	72,5	93,1	115,9
X_1	7	1	11	11	7	11	3	1	2	21
X_2	26	29	59	31	52	55	71	31	54	47

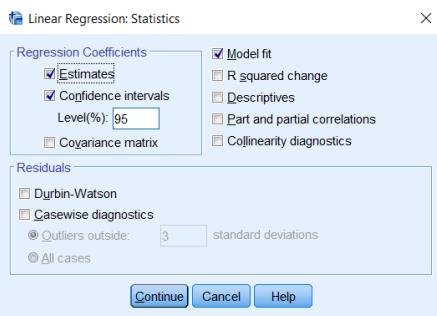
4
1

The screenshot shows the IBM SPSS Statistics Data Editor interface. The menu bar is visible at the top, with 'Analyze' being the active tab. A sub-menu for 'Regression' is open, listing various regression analysis options: Automatic Linear Modeling..., Linear..., Curve Estimation..., Partial Least Squares..., Binary Logistic..., Multinomial Logistic..., Ordinal..., Probit..., Nonlinear..., Weight Estimation..., 2-Stage Least Squares..., and Optimal Scaling (CATREG)... . The 'Linear...' option is currently selected. The main data view window shows a table with columns labeled 'Y' and 'var'. The bottom left corner of the window displays 'Data View' and 'Variable View' tabs, with 'Data View' being the active tab.

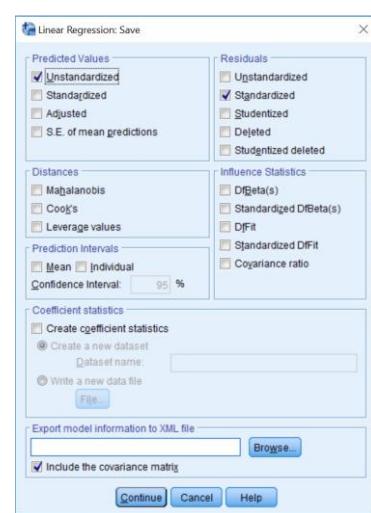
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Regression>Linear>Statistics



Regression>Linear>Save



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Variables Entered/Removed^a

Model	Variables Entered	Variables Removed	Method
1	X2, X1 ^b	.	Enter

a. Dependent Variable: Y

b. All requested variables entered.

Model Summary^b

Model	R	R Square	Adjusted R Square	Std. Error of the Estimate
1	,988 ^a	,977	,970	2,57617

a. Predictors: (Constant), X2, X1

b. Dependent Variable: Y



ANOVA (Modelo)

H_0 : O modelo de regressão considerado não serve

Decisão: Como valor $p < 0,05$, rejeita-se a H_0 , pelo que o modelo de regressão considerado é estatisticamente significativo

ANOVA^a

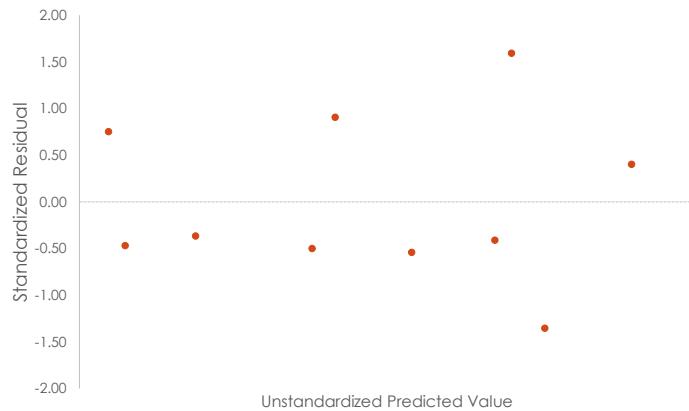
Model		Sum of Squares	df	Mean Square	F	Sig.
1	Regression	1976,924	2	988,462	148,940	,000 ^b
	Residual	46,457	7	6,637		
	Total	2023,381	9			

a. Dependent Variable: Y

b. Predictors: (Constant), X2, X1



RESÍDUOS (homoscedasticidade)



Complementos de Estatística, Profª
Ana Cristina Braga

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Resíduos (Normalidade)

- Teste analítico (KS com correção de Lilliefors)
- Método gráfico (P-P ou Q-Q plot)

NPar Tests

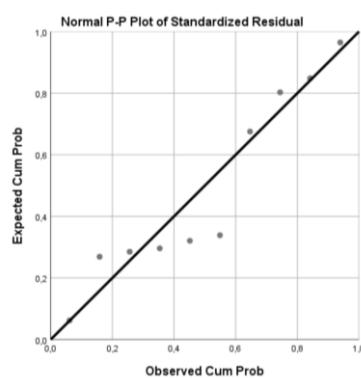
One-Sample Kolmogorov-Smirnov Test

		Standardized Residual
N		10
Normal Parameters ^{a,b}	Mean	,0000000
	Std. Deviation	,88191710
Most Extreme Differences	Absolute	,261
	Positive	,261
	Negative	-,169
Test Statistic		,261
Asymp. Sig. (2-tailed)		,051 ^c

a. Test distribution is Normal.

b. Calculated from data.

c. Lilliefors Significance Correction.



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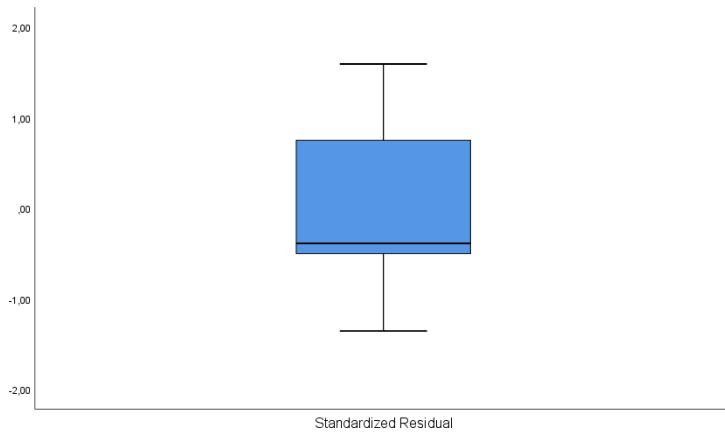
Resíduos (média zero)

T-Test

One-Sample Statistics					
	N	Mean	Std. Deviation	Std. Error Mean	
Standardized Residual	10	,0000000	,88191710	,27888668	
One-Sample Test					
				Test Value = 0	
	t	df	Sig. (2-tailed)	Mean Difference	95% Confidence Interval of the Difference
Standardized Residual	1,19E-014	9	1,000	3,32E-015	-.631 ,631

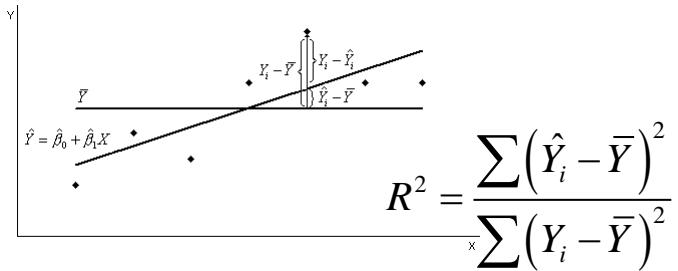


Verificação de outliers





COEFICIENTE DE DETERMINAÇÃO



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Coeficiente de determinação (r^2), representa a proporção da variação de Y que é explicada pela regressão

$$r^2 = \frac{\hat{\beta}_1^2 \cdot s_{xx}}{s_{yy}} = \frac{\hat{\beta}_1^2 \cdot \sum_i (X_i - \bar{X})^2}{\sum_i (Y_i - \bar{Y})^2} = \frac{\text{variação de } Y \text{ explicada pela regressão}}{\text{variação total de } Y}$$

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