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proof of Bennett inequality

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By http://planetmath.org/ChernoffCramerBoundChernoff-Cramer inequality, we have:

$$\Pr\left\{\sum_{i=1}^{n} \left(X_{i} - E[X_{i}]\right) > \varepsilon\right\} \leq \exp\left[-\sup_{t \geq 0} \left(t\varepsilon - \psi(t)\right)\right]$$

where

$$\psi(t) = \sum_{i=1}^{n} \left(\ln E \left[e^{tX_i} \right] - tE \left[X_i \right] \right).$$

Keeping in mind that the condition

$$\Pr\{|X_i| \le M\} = 1 \ \forall i$$

implies that, for all i,

$$E[|X_i|^k] \le M^k \ \forall k \ge 0$$

(see http://planetmath.org/RelationBetweenAlmostSurelyAbsolutelyBoundedRandomVariab for a proof) and since $\ln x \le x - 1 \ \forall x > 0$, and

$$E[\left|X\right|^{k}] \leq M^{k} \quad \Longrightarrow \quad E\left[\left|X\right|^{k}\right] \leq E\left[X^{2}\right]M^{k-2} \qquad \quad \forall k \geq 2, k \in N$$

 $(see \, \mathtt{http://planetmath.org/AbsoluteMomentsBoundingNecessaryAndSufficientCondition here on the absolute and the second of t$

for a proof), one has:

$$\psi(t) = \sum_{i=1}^{n} \left(\ln E \left[e^{tX_i} \right] - tE \left[X_i \right] \right) \\
\leq \sum_{i=1}^{n} E \left[e^{tX_i} \right] - tE \left[X_i \right] - 1 \\
= \sum_{i=1}^{n} E \left[\sum_{k=0}^{\infty} \frac{(tX_i)^k}{k!} \right] - tE \left[X_i \right] - 1 \\
= \sum_{i=1}^{n} \left(\sum_{k=0}^{\infty} \frac{t^k E \left[X_i^k \right]}{k!} \right) - tE \left[X_i \right] - 1 \\
= \sum_{i=1}^{n} \left(\sum_{k=2}^{\infty} \frac{t^k E \left[X_i^k \right]}{k!} \right) \\
\leq \sum_{i=1}^{n} \left(\sum_{k=2}^{\infty} \frac{t^k E \left[X_i^2 \right] M^{k-2}}{k!} \right) \\
= \sum_{i=1}^{\infty} \left(\sum_{k=2}^{\infty} \frac{t^k M^{k-2} \sum_{i=1}^{n} E \left[X_i^2 \right]}{k!} \right) \\
= \sum_{k=2}^{\infty} \frac{t^k M^{k-2} \sum_{i=1}^{n} E \left[X_i^2 \right]}{k!} \\
= \frac{v^2}{M^2} \sum_{k=2}^{\infty} \frac{(tM)^k}{k!} \\
= \frac{v^2}{M^2} \left[\exp(tM) - tM - 1 \right]$$

One can now write

$$\sup_{t\geq 0}\left(t\varepsilon-\psi(t)\right)\geq \sup_{t\geq 0}\left(t\varepsilon-\frac{v^2}{M^2}\left(e^{tM}-tM-1\right)\right)=\sup_{t>0}\left[\frac{v^2}{M^2}\left(\frac{M^2\varepsilon}{v^2}t-\left(e^{tM}-tM-1\right)\right)\right].$$

By elementary calculus, we obtain the value of t that maximizes the expression in round brackets:

$$t_{opt} = \frac{1}{M} \ln \left(1 + \frac{M\varepsilon}{v^2} \right)$$

which, once plugged into the bound, yields

$$\Pr\left\{\sum_{i=1}^{n} \left(X_{i} - E[X_{i}]\right) > \varepsilon\right\} \leq \exp\left[-\frac{v^{2}}{M^{2}}\left(\left(1 + \frac{M\varepsilon}{v^{2}}\right)\ln\left(1 + \frac{M\varepsilon}{v^{2}}\right) - \frac{M\varepsilon}{v^{2}}\right)\right].$$

Observing that $(1+x)\ln(1+x)-x \ge \frac{x}{2}\ln(1+x) \, \forall x \ge 0$ (see http://planetmath.org/ASimple one gets the sub-optimal yet more easily manageable formula:

$$\Pr\left\{\sum_{i=1}^{n} \left(X_{i} - E[X_{i}]\right) > \varepsilon\right\} \leq \exp\left[-\frac{\varepsilon}{2M} \ln\left(1 + \frac{\varepsilon M}{v^{2}}\right)\right].$$