

Predicting Final Exam Scores

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Type : Homework Problem

Course : Applied Statistics/Regression (MATH-564)

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Institution : Illinois Institute of Technology

Packages Used:

```
library(knitr)
library(kableExtra, warn.conflicts = F)
library(tidyverse, warn.conflicts = F)
```

The Data

```
table3.10 <- read_tsv("Table3.10.txt", show_col_types = F)[c(2,3,1)]
Table3.10 <- cbind(Index = 1:22, table3.10)
colnames(Table3.10) <- c("Index", "$P_1$", "$P_2$", "$F$")

kbl(cbind(Table3.10[1:11,], Table3.10[12:22,]), booktabs = T, escape = F,
     align = "c", linesep = "", valign = "c",
     caption = "$\\textbf{Table 3.10} - \\text{Examination Data}$") %>%
  kable_classic() %>%
  kable_styling(latex_options = c("condensed", "striped", "HOLD_position"), font_size = 11) %>%
  column_spec(c(2:4, 6:7), width = "1.25cm") %>%
  column_spec(c(1,5), width = "1.5cm", border_left = T, border_right = F) %>%
  column_spec(8, width = "1.25cm", border_right = T)
```

Table 3.10 – Examination Data

| Index | P_1 | P_2 | F | Index | P_1 | P_2 | F |
|-------|-------|-------|-----|-------|-------|-------|-----|
| 1 | 78 | 73 | 68 | 12 | 79 | 75 | 75 |
| 2 | 74 | 76 | 75 | 13 | 89 | 84 | 81 |
| 3 | 82 | 79 | 85 | 14 | 93 | 97 | 91 |
| 4 | 90 | 96 | 94 | 15 | 87 | 77 | 80 |
| 5 | 87 | 90 | 86 | 16 | 91 | 96 | 94 |
| 6 | 90 | 92 | 90 | 17 | 86 | 94 | 94 |
| 7 | 83 | 95 | 86 | 18 | 91 | 92 | 97 |
| 8 | 72 | 69 | 68 | 19 | 81 | 82 | 79 |
| 9 | 68 | 67 | 55 | 20 | 80 | 83 | 84 |
| 10 | 69 | 70 | 69 | 21 | 70 | 66 | 65 |
| 11 | 91 | 89 | 91 | 22 | 79 | 81 | 83 |

Exercise 3.3

Table 3.10 shows the scores in the final examination F and the scores in two preliminary examinations P_1 and P_2 for 22 students in a statistics course

(a) Fit each of the following models to the data:

$$\text{Model 1: } F = \beta_0 + \beta_1 P_1 + \varepsilon$$

$$\text{Model 2: } F = \beta_0 + \beta_2 P_2 + \varepsilon$$

$$\text{Model 3: } F = \beta_0 + \beta_1 P_1 + \beta_2 P_2 + \varepsilon$$

```
fit_P1 <- lm(`F` ~ P1, data = table3.10)
coefP1 <- round(coefficients(fit_P1), 3)

fit_P2 <- lm(`F` ~ P2, data = table3.10)
coefP2 <- round(coefficients(fit_P2), 3)

fit_P1P2 <- lm(`F` ~ ., data = table3.10)
coefP1P2 <- round(coefficients(fit_P1P2), 3)
```

$$\text{Fitted Model 1: } \hat{F} = -22.342 + 1.261 P_1$$

$$\text{Fitted Model 2: } \hat{F} = -1.854 + 1.004 P_2$$

$$\text{Fitted Model 3: } \hat{F} = -14.501 + 0.488 P_1 + 0.672 P_2$$

(b) Test whether $\beta_0 = 0$ in each of the three models.

I will use t-test hypothesis test for each model where $H_0 : \hat{\beta}_0 = 0$ and $H_A : \hat{\beta}_0 \neq 0$.

There are $n = 22$ rows in the dataset. Under the null, the critical t-value has $n - p$ degrees of freedom ($d.f.$), where p equals the number of coefficients in the alternative regression model. Equivalently, p equals the number of predictors in a regression model since the intercept term is removed under the null.

Thus, Model 1 and Model 2 both have 20 $d.f.$ and Model 3 has 19 $d.f.$

Using a significance level, $\alpha = 0.05$, then the critical t-values for a two-tailed test are the following:

$$t_{(\alpha/2, d.f.=20)} = \pm 2.086 \quad \text{and} \quad t_{(\alpha/2, d.f.=19)} = \pm 2.093$$

Next, the following equation is used to calculate the test statistic for $H_A : t^* = \frac{\hat{\beta}_0 - 0}{s.e.(\hat{\beta}_0)}$.

We reject H_0 in favor of H_A if $|t^*| > |t_{(\alpha/2, d.f.)}|$.

```
# Saving Model Summaries
sumP1 <- summary(fit_P1)
sumP2 <- summary(fit_P2)
sumP1P2 <- summary(fit_P1P2)
# Obtaining Standard Errors
seP1B0 <- sumP1$coefficients[1,2]
seP2B0 <- sumP2$coefficients[1,2]
seP1P2B0 <- sumP1P2$coefficients[1,2]
```

$$\begin{aligned}\text{Model 1: } |t^*| &= \left| \frac{-22.342}{11.564} \right| = 1.932 < 2.086 = |t_{(0.025, 20)}| \\ \text{Model 2: } |t^*| &= \left| \frac{-1.854}{7.562} \right| = 0.245 < 2.086 = |t_{(0.025, 20)}| \\ \text{Model 3: } |t^*| &= \left| \frac{-14.501}{9.236} \right| = 1.570 < 2.093 = |t_{(0.025, 19)}|\end{aligned}$$

Conclusion

In all three models $|t^*| < |t_{(\alpha/2, d.f.)}|$. As a result, we fail to reject the null hypothesis for all models. There is insufficient evidence in favor of the alternative hypothesis.

(c) Which Predictor is Better? P_1 or P_2 ? (Quick Model Selection)

The regression summaries for Model 1 and Model 2 are provided in the tables below:

```
coefTab1 <- cbind(IV = c("(Intercept)", "$P_1$"), as_tibble(signif(coef(sumP1), 5)))
coefTab1[,5] <- as.character(signif(coefTab1[,5], 2))
kbl(coefTab1, booktabs = T, align = "c", escape = F, digits = 2, valign = "c",
    col.names = c(" ", "$\\widehat{\\beta}_j$", "$s.e.$", "$t^{\\ast}$", "$\\textit{p-value}$")) %>%
  kable_classic() %>%
  kable_styling(latex_options = c("striped", "HOLD_position")) %>%
  add_header_above(c("Model 1" = 5), bold = T, font_size = 12) %>%
  column_spec(1:5, width = "2cm")
```

| Model 1 | | | | |
|-------------|-----------------|-------|-------|---------|
| | $\hat{\beta}_j$ | s.e. | t^* | p-value |
| (Intercept) | -22.34 | 11.56 | -1.93 | 0.068 |
| P_1 | 1.26 | 0.14 | 9.01 | 1.8e-08 |

| Model 2 | | | | |
|-------------|-----------------|------|-------|---------|
| | $\hat{\beta}_j$ | s.e. | t^* | p-value |
| (Intercept) | -1.85 | 7.56 | -0.25 | 0.81 |
| P_2 | 1.00 | 0.09 | 11.09 | 5.4e-10 |

Both predictors are statistically significant as their p -values are less than the desired level of significance, $\alpha = 0.05$.

A quick way too access which predictor is better is comparing the R^2 and Mean Squared Error (MSE) statistics of each model. R^2 measures a model's goodness-of-fit; MSE is statistic used to evaluate the prediction accuracy of a model.

$$R^2 = 1 - \frac{SSE}{SST} \quad \text{and} \quad MSE = \frac{SSE}{n} = \frac{1}{n} \sum_{i=1}^n (y_i - \hat{y}_i)^2,$$

where $SST = \sum_{i=1}^n (y_i - \bar{y})^2$ is the sum of squares in total, n is the number of rows in the data ($n = 22$), and b is the number of regression coefficients in a model.

For both Models 1 and 2, there is an intercept and one predictor so $b = 2$; Model 3 has one addition coefficient so $b = 3$

- $0 \leq R^2 \leq 1$ such that R^2 is optimized when it is maximized.
- $MSE \geq 0$ such that MSE is optimized when it is minimized.

```
ModStats <- tibble(" " = c("Model 1", "Model 2"),
  `R^2` = c(sumP1$r.squared, sumP2$r.squared),
  `MSE` = c(mean(sumP1$residuals^2), mean(sumP2$residuals^2)))
```

The values for both statistic are provided in the table below:

| | R^2 | MSE |
|----------------|--------|-------|
| Model 1 | 0.8023 | 23.47 |
| Model 2 | 0.8600 | 16.61 |

Conclusion:

This quick model selection method indicates that R^2 is maximized and MSE is minimized by Model 2. Therefore, I would prefer to use the second preliminary exam, P_2 , to predict final exam scores, F .

- (d) Which of the three models with intercepts would you use to predict the final examination scores for a student who scored 78 and 85 on the first and second preliminary examinations, respectively? (Quick Model Selection) What is your prediction in this case?

R^2 becomes larger as more predictors are used to fit a model. This means R^2 does not account the bias of larger models.

In contrast, adjusted R^2 denoted R_{adj}^2 accounts for bias by punishing models as they add predictors:

$$R_{adj}^2 = 1 - \frac{SSE/(n-b)}{SST/(n-1)} = 1 - \frac{SSE \cdot (n-1)}{SST \cdot (n-b)}.$$

The following properties of R^2 still hold: $0 \leq R_{adj}^2 \leq 1$ such that R_{adj}^2 is optimized when it is maximized. However, R_{adj}^2 decreases as predictors are added when all other values are fixed.

Because Model 3 has an additional coefficient, R_{adj}^2 should be used to compare it to the smaller models instead of the unadjusted R^2 .

```
ModStats2 <- tibble(" " = c("Model 1", "Model 2", "Model 3"),
  `R^2_{adj}` = c(sumP1$adj.r.squared, sumP2$adj.r.squared, sumP1P2$adj.r.squared),
  `\\textit{MSE}` = c(mean(sumP1$residuals^2), mean(sumP2$residuals^2), mean(sumP1P2$residuals^2))
```

The values of R_{adj}^2 and MSE for each model are displayed in the table below:

| | R_{adj}^2 | MSE |
|----------------|-------------|-------|
| Model 1 | 0.7924 | 23.47 |
| Model 2 | 0.8530 | 16.61 |
| Model 3 | 0.8744 | 13.49 |

Conclusion:

R_{adj}^2 and MSE were optimized by *Model 3*. Recall, the estimated coefficients for *Model 3* derived in Part (a):

- $\hat{F} = -14.501 + 0.488 P_1 + 0.672 P_2$

Accordingly, if a student had preliminary examination scores $P_1 = 78$ and $P_2 = 85$, *Model 3* predicts this student will have a final examination score of $\hat{F} = 80.713$.